T-testing: A Linear-time Adaptive Nonparametric Technique

Zoltán Szabó (CMAP, École Polytechnique)

Wittawat Jitkrittum

Kacper Chwialkowski

Arthur Gretton

Machine Learning Seminar
Télécom ParisTech
February 2, 2017

Contents

- Motivating examples: NLP, computer vision.
- Two-sample test: t-test $\rightarrow$ distribution features.
- Linear-time, interpretable, high-power, nonparametric t-test.
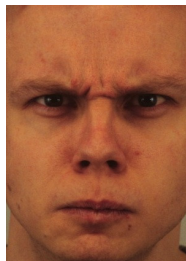- Numerical illustrations.

Motivating examples

- Given: two categories of documents (Bayesian inference, neuroscience).
- Task:
  - test their distinguishability,
  - most discriminative words $\rightarrow$ interpretability.
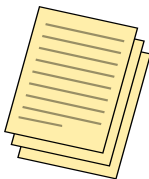
- Given: two sets of faces (happy, angry).
- Task:
    - check if they are different,
    - determine the most discriminative features/regions.

Contribution:

- We propose a nonparametric t-test.
- It gives a reason why H_0 is rejected.
- It has high test power.
- It runs in linear time.

One-page summary

Contribution:

- We propose a nonparametric t-test.
- It gives a reason why H_0 is rejected.
- It has high test power.
- It runs in linear time.

Dissemination, code:

- NIPS-2016 [Jitkrittum et al., 2016]: full oral = top 1.84%.
- https://github.com/wittawatj/interpretable-test.

Two-sample test, distribution features

- Given:
  - $X = \{\mathbf{x}_i\}_{i=1}^n \overset{i.i.d.}{\sim} \mathbb{P}$, $Y = \{\mathbf{y}_j\}_{j=1}^n \overset{i.i.d.}{\sim} \mathbb{Q}$.
  - Example: $\mathbf{x}_i = i^{th}$ happy face, $\mathbf{y}_j = j^{th}$ sad face.

What is a two-sample test?

- Given:
  - $X = \{\mathbf{x}_i\}_{i=1}^n \overset{i.i.d.}{\sim} \mathbb{P}$, $Y = \{\mathbf{y}_j\}_{j=1}^n \overset{i.i.d.}{\sim} \mathbb{Q}$.
  - Example: $\mathbf{x}_i = i^{th}$ happy face, $\mathbf{y}_j = j^{th}$ sad face.
- Problem: using X, Y test

$$H_0 : \mathbb{P} = \mathbb{Q}, \text{ vs}$$
$$H_1 : \mathbb{P} \neq \mathbb{Q}.$$

What is a two-sample test?

- Given:
    - $X = \{\mathbf{x}_i\}_{i=1}^n \overset{i.i.d.}{\sim} \mathbb{P}$, $Y = \{\mathbf{y}_j\}_{j=1}^n \overset{i.i.d.}{\sim} \mathbb{Q}$.
    - Example: $\mathbf{x}_i = i^{th}$ happy face, $\mathbf{y}_j = j^{th}$ sad face.
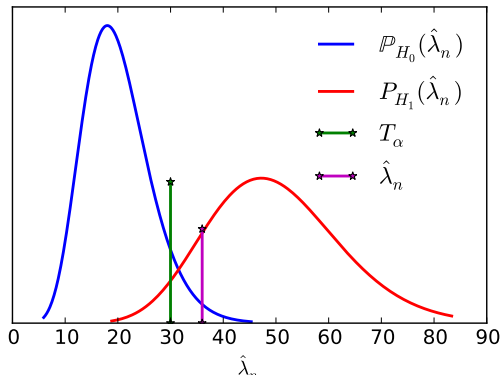- Problem: using X, Y test

$$H_0 : \mathbb{P} = \mathbb{Q}, \text{ vs}$$
$$H_1 : \mathbb{P} \neq \mathbb{Q}.$$

- Assume $X, Y \subset \mathbb{R}^d$.

Ingredients of two-sample test

- Test statistic: $\hat{\lambda}_n = \hat{\lambda}_n(X, Y)$, random.
- Significance level: $\alpha = 0.01$.
- Under H_0: $P_{H_0}(\underbrace{\hat{\lambda}_n \leqslant T_\alpha}_{\text{correctly accepting } H_0}) = 1 - \alpha$.
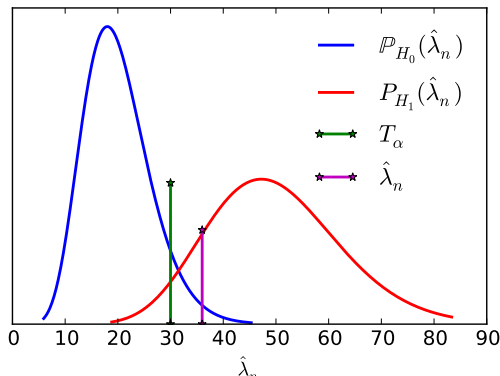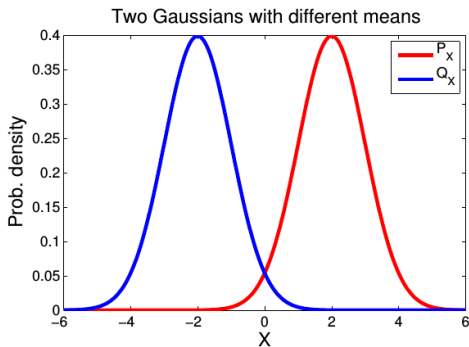
Ingredients of two-sample test

- Test statistic: $\hat{\lambda}_n = \hat{\lambda}_n(X, Y)$, random.
- Significance level: $\alpha = 0.01$.
- Under H_0: $P_{H_0}(\underbrace{\hat{\lambda}_n \leqslant T_\alpha}_{\text{correctly accepting } H_0}) = 1 - \alpha$.
- Under H_1: $P_{H_1}(T_\alpha < \hat{\lambda}_n) = P(\text{correctly rejecting } H_0) =: $ power.

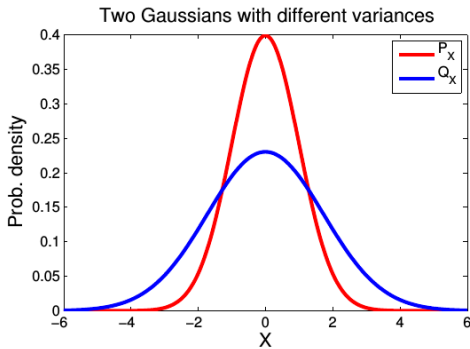- Given: 2 Gaussians with (possibly) different means.
- Solution: t-test.

Two Gaussians with different means

- Setup: 2 Gaussians; same means, different variances.
- Idea: look at 2nd-order features of RVs.

Two Gaussians with different variances
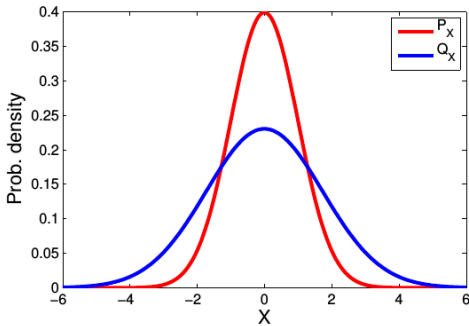
- Setup: 2 Gaussians; same means, different variances.
- Idea: look at 2nd-order features of RVs.
- $\varphi_x = x^2 \Rightarrow$ difference in $\mathbb{E}X^2$.
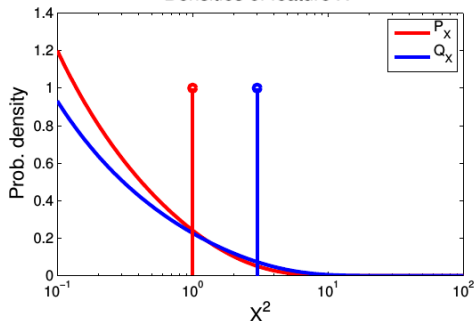
- Setup: a Gaussian and a Laplacian distribution.
- Challenge: their means *and* variances are the same.
- Idea: look at higher-order features.

Gaussian and Laplace densities
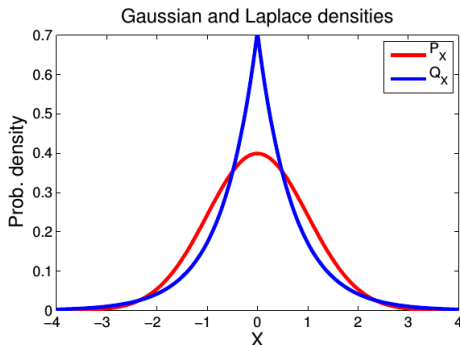
Let us consider feature/distribution representations!

- Given: $\mathbf{x}$ and $\mathbf{x}'$ objects (images or texts).

Kernel: similarity between features

- Given: $\mathbf{x}$ and $\mathbf{x}'$ objects (images or texts).
- Question: how similar they are?

Kernel: similarity between features

- Given: $\mathbf{x}$ and $\mathbf{x}'$ objects (images or texts).
- Question: how similar they are?
- Define features of the objects:

$$\varphi_{\mathbf{x}} : \text{features of } \mathbf{x},$$
$$\varphi_{\mathbf{x}'} : \text{features of } \mathbf{x}'.$$

- Kernel: inner product of these features

$$k(\mathbf{x}, \mathbf{x}') := \langle \varphi_{\mathbf{x}}, \varphi_{\mathbf{x}'} \rangle .$$

- Polynomial kernel:

$$k(\mathbf{x}, \mathbf{y}) = (\langle \mathbf{x}, \mathbf{y} \rangle + \gamma)^p.$$

- Gaussian kernel:

$$k(\mathbf{x}, \mathbf{y}) = e^{-\gamma \|\mathbf{x} - \mathbf{y}\|_2^2}.$$

Towards distribution features
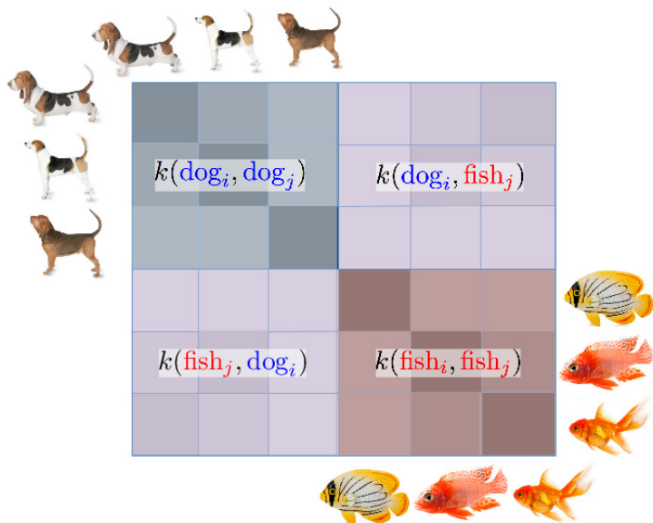
$$\widehat{MMD}^2(\mathbb{P}, \mathbb{Q}) = \overline{K_{\mathbb{P},\mathbb{P}}} + \overline{K_{\mathbb{Q},\mathbb{Q}}} - 2\overline{K_{\mathbb{P},\mathbb{Q}}} \text{ (without diagonals in } \overline{K_{\mathbb{P},\mathbb{P}}}, \overline{K_{\mathbb{Q},\mathbb{Q}}})$$

[†]$\widehat{MMD}$ illustration credit: Arthur Gretton

- Kernel recall: $k(\mathbf{x}, \mathbf{x}') = \langle \varphi_{\mathbf{x}}, \varphi_{\mathbf{x}'} \rangle$.

- Kernel recall: $k(\mathbf{x}, \mathbf{x}') = \langle \varphi_{\mathbf{x}}, \varphi_{\mathbf{x}'} \rangle$.
- Feature of $\mathbb{P}$ (mean embedding):

$$\mu_{\mathbb{P}} := \mathbb{E}_{\mathbf{x} \sim \mathbb{P}}[\varphi_{\mathbf{x}}].$$

- Kernel recall: $k(\mathbf{x}, \mathbf{x}') = \langle \varphi_{\mathbf{x}}, \varphi_{\mathbf{x}'} \rangle$.
- Feature of $\mathbb{P}$ (mean embedding):

$$\mu_{\mathbb{P}} := \mathbb{E}_{\mathbf{x} \sim \mathbb{P}}[\varphi_{\mathbf{x}}].$$

- Previous quantity: unbiased estimate of

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|^2.$$

Kernel → distribution feature

- Kernel recall: $k(\mathbf{x}, \mathbf{x}') = \langle \varphi_\mathbf{x}, \varphi_{\mathbf{x}'} \rangle$.
- Feature of $\mathbb{P}$ (mean embedding):

$$\mu_\mathbb{P} := \mathbb{E}_{\mathbf{x} \sim \mathbb{P}}[\varphi_\mathbf{x}].$$

- Previous quantity: unbiased estimate of

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_\mathbb{P} - \mu_\mathbb{Q}\|^2.$$

- Valid test [Gretton et al., 2012]. Challenges:
  1. Threshold choice: 'ugly' asymptotics of $n\widehat{MMD^2}(\mathbb{P}, \mathbb{P})$.
  2. Test statistic: quadratic time complexity.
  3. Witness $\in \mathcal{H}(k)$: can be hard to interpret.

Linear-time tests

- Recall:

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|^2_{\mathcal{H}(k)}.$$

- Changing [Chwialkowski et al., 2015] this to

$$\rho^2(\mathbb{P}, \mathbb{Q}) := \frac{1}{J} \sum_{j=1}^{J} [\mu_{\mathbb{P}}(\mathbf{v}_j) - \mu_{\mathbb{Q}}(\mathbf{v}_j)]^2$$

  with random $\{\mathbf{v}_j\}_{j=1}^{J}$ test locations.

ρ is a metric (a.s.). How do we estimate it? Distribution under H_0?

Compute

$$\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} = \frac{1}{J} \sum_{j=1}^{J} [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2,$$

where $\hat{\mu}_{\mathbb{P}}(\mathbf{v}) = \frac{1}{n} \sum_{i=1}^{n} k(\mathbf{x}_i, \mathbf{v})$. Example using $k(\mathbf{x}, \mathbf{v}) = e^{-\frac{\|\mathbf{x}-\mathbf{v}\|^2}{2\sigma^2}}$:

Legend:
- $\hat{\mu}_P(\mathbf{v})$
- $\hat{\mu}_Q(\mathbf{v})$
- $(\hat{\mu}_P(\mathbf{v}) - \hat{\mu}_Q(\mathbf{v}))^2$

$$\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} = \frac{1}{J} \sum_{j=1}^{J} [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2$$

$$\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} = \frac{1}{J} \sum_{j=1}^{J} [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2$$

$$= \frac{1}{J} \sum_{j=1}^{J} \left[\frac{1}{n} \sum_{i=1}^{n} k(\mathbf{x}_i, \mathbf{v}_j) - \frac{1}{n} \sum_{i=1}^{n} k(\mathbf{y}_i, \mathbf{v}_j) \right]^2$$

$$\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} = \frac{1}{J} \sum_{j=1}^{J} [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2$$

$$= \frac{1}{J} \sum_{j=1}^{J} \left[\frac{1}{n} \sum_{i=1}^{n} k(\mathbf{x}_i, \mathbf{v}_j) - \frac{1}{n} \sum_{i=1}^{n} k(\mathbf{y}_i, \mathbf{v}_j) \right]^2 = \frac{1}{J} \sum_{j=1}^{J} (\bar{\mathbf{z}}_n)_j^2 = \frac{1}{J} \bar{\mathbf{z}}_n^T \bar{\mathbf{z}}_n,$$

where $\bar{\mathbf{z}}_n = \frac{1}{n} \sum_{i=1}^{n} \underbrace{\left[k(\mathbf{x}_i, \mathbf{v}_j) - k(\mathbf{y}_i, \mathbf{v}_j) \right]_{j=1}^{J}}_{=: \mathbf{z}_i} \in \mathbb{R}^J.$

$$\widehat{\rho^2(\mathbb{P}, \mathbb{Q})} = \frac{1}{J} \sum_{j=1}^{J} [\hat{\mu}_{\mathbb{P}}(\mathbf{v}_j) - \hat{\mu}_{\mathbb{Q}}(\mathbf{v}_j)]^2$$

$$= \frac{1}{J} \sum_{j=1}^{J} \left[\frac{1}{n} \sum_{i=1}^{n} k(\mathbf{x}_i, \mathbf{v}_j) - \frac{1}{n} \sum_{i=1}^{n} k(\mathbf{y}_i, \mathbf{v}_j) \right]^2 = \frac{1}{J} \sum_{j=1}^{J} (\bar{\mathbf{z}}_n)_j^2 = \frac{1}{J} \bar{\mathbf{z}}_n^T \bar{\mathbf{z}}_n,$$

where $\bar{\mathbf{z}}_n = \frac{1}{n} \sum_{i=1}^{n} \underbrace{\left[k(\mathbf{x}_i, \mathbf{v}_j) - k(\mathbf{y}_i, \mathbf{v}_j) \right]_{j=1}^{J}}_{=: \mathbf{z}_i} \in \mathbb{R}^J$.

- Good news: estimation is linear in n!
- Bad news: intractable null distr. $= \sqrt{n} \widehat{\rho^2(\mathbb{P}, \mathbb{P})} \xrightarrow{w}$ sum of J correlated χ^2.

- Modified test statistic:

$$\hat{\lambda}_n = n\bar{\mathbf{z}}_n^T \boldsymbol{\Sigma}_n^{-1} \bar{\mathbf{z}}_n,$$

  where $\boldsymbol{\Sigma}_n = cov\left(\{\mathbf{z}_i\}_{i=1}^n\right)$.
- Under H_0:
  - $\hat{\lambda}_n \xrightarrow{w} \chi^2(J)$. $\Rightarrow$ Easy to get the $(1-\alpha)$-quantile!

Our idea

- Until this point: test locations ($\mathcal{V}$) are fixed.
- Instead: choose $\theta = \{\mathcal{V}, \sigma\}$ to

   maximize lower bound on the test power.

- Until this point: test locations ($\mathcal{V}$) are fixed.
- Instead: choose $\theta = \{\mathcal{V}, \sigma\}$ to

  maximize lower bound on the test power.

Theorem (Lower bound on power, for large n)

Test power $\geqslant L(\lambda_n)$; L: explicit function, increasing.

- Here,
  - $\lambda_n = n\boldsymbol{\mu}^T \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu}$: population version of $\hat{\lambda}_n$.
  - $\boldsymbol{\mu} = \mathbb{E}_{\mathbf{xy}}[\mathbf{z}_1]$, $\boldsymbol{\Sigma} = \mathbb{E}_{\mathbf{xy}}\left[(\mathbf{z}_1 - \boldsymbol{\mu})(\mathbf{z}_1 - \boldsymbol{\mu})^T\right]$.

But λ_n is unknown. Split (X, Y) into (X_{tr}, Y_{tr}) and (X_{te}, Y_{te}).

1. Locations, kernel parameter: $\hat{\theta} = \arg\max_\theta \hat{\lambda}^{tr}_{\frac{n}{2}}(\theta)$.

Convergence of the λ_n estimator

But λ_n is unknown. Split (X, Y) into (X_{tr}, Y_{tr}) and (X_{te}, Y_{te}).

1. Locations, kernel parameter: $\hat{\theta} = \arg\max_\theta \hat{\lambda}_{\frac{n}{2}}^{tr}(\theta)$.
2. Test statistic: $\hat{\lambda}_{\frac{n}{2}}^{te}(\hat{\theta})$.

Convergence of the λ_n estimator

Theorem (Guarantee on objective approximation, $\gamma_n \to 0$)

$$\sup_{\mathcal{V},\mathcal{K}} \left| \bar{\mathbf{z}}_n^T (\mathbf{\Sigma}_n + \gamma_n)^{-1} \bar{\mathbf{z}}_n - \boldsymbol{\mu}^T \mathbf{\Sigma}^{-1} \boldsymbol{\mu} \right| = \mathcal{O}\left(n^{-\frac{1}{4}}\right).$$

Theorem (Guarantee on objective approximation, $\gamma_n \to 0$)

$$\sup_{\mathcal{V},\mathcal{K}} \left| \bar{\mathbf{z}}_n^T (\mathbf{\Sigma}_n + \gamma_n)^{-1} \bar{\mathbf{z}}_n - \boldsymbol{\mu}^T \mathbf{\Sigma}^{-1} \boldsymbol{\mu} \right| = \mathcal{O}\left(n^{-\frac{1}{4}}\right).$$

Examples:

$$\mathcal{K} = \left\{ k_\sigma(\mathbf{x}, \mathbf{y}) = e^{-\frac{\|\mathbf{x}-\mathbf{y}\|^2}{2\sigma^2}} : \sigma > 0 \right\},$$

$$\mathcal{K} = \left\{ k_{\mathbf{A}}(\mathbf{x}, \mathbf{y}) = e^{-(\mathbf{x}-\mathbf{y})^T \mathbf{A}(\mathbf{x}-\mathbf{y})} : \mathbf{A} > 0 \right\}.$$

Numerical demos

Parameter settings

- Gaussian kernel (σ). $\alpha = 0.01$. $J = 1$. Repeat 500 trials.
- Report

$$P(\text{reject } H_0) \approx \frac{\#\text{times } \hat{\lambda}_n > T_\alpha \text{ holds}}{\#\text{trials}}.$$

- Compare 4 methods
    - **ME-full**: Optimize $\mathcal{V}$ and Gaussian bandwidth σ.
    - **ME-grid**: Optimize σ. Random $\mathcal{V}$ [Chwialkowski et al., 2015].
    - **MMD-quad**: Test with quadratic-time MMD [Gretton et al., 2012].
    - **MMD-lin**: Test with linear-time MMD [Gretton et al., 2012].
- Optimize kernels to power in MMD-lin, MMD-quad.

NLP: discrimination of document categories

- 5903 NIPS papers (1988-2015).
- Keyword-based category assignment into 4 groups:
    - Bayesian inference, Deep learning, Learning theory, Neuroscience
- $d = 2000$ nouns. TF-IDF representation.

Problem	n^{te}	**ME-full**	ME-grid	MMD-quad	MMD-lin
1. Bayes-Bayes	215	.012	.018	.022	.008
2. Bayes-Deep	216	.954	.034	.906	.262
3. Bayes-Learn	138	.990	.774	1.00	.238
4. Bayes-Neuro	394	1.00	.300	.952	.972
5. Learn-Deep	149	.956	.052	.876	.500
6. Learn-Neuro	146	.960	.572	1.00	.538

- Performance of ME-full $[\mathcal{O}(n)]$ is comparable to MMD-quad $[\mathcal{O}(n^2)]$.

- Aggregating over trials; example: 'Bayes-Neuro'.
- Most discriminative words:

  spike, markov, cortex, dropout, recurr, iii, gibb.

  - learned test locations: highly interpretable,
  - 'markov', 'gibb' ($\Leftarrow$ Gibbs): Bayesian inference,
  - 'spike', 'cortex': key terms in neuroscience.

- Aggregating over trials; example: 'Bayes-Neuro'.

- Least dicriminative ones:
  circumfer, bra, dominiqu, rhino, mitra, kid, impostor.

Distinguish positive/negative emotions

- Karolinska Directed Emotional Faces (KDEF) [Lundqvist et al., 1998].
- 70 actors = 35 females and 35 males.
- $d = 48 \times 34 = 1632$. Grayscale. Pixel features.

+ :   happy   neutral   surprised   − :   afraid   angry   disgusted
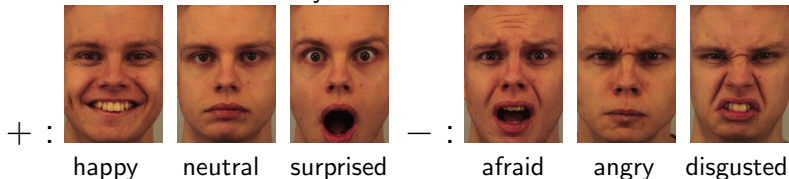
Problem	n^{te}	**ME-full**	ME-grid	MMD-quad	MMD-lin
± vs. ±	201	.010	.012	.018	.008
+ vs. −	201	.998	.656	1.00	.578

- Learned test location (averaged) =

Summary

- We proposed a nonparametric t-test:
  - linear time,
  - high-power ($\approx$ 'MMD-quad'),
- 2 demos: discriminating
  - documents of different categories,
  - positive/negative emotions.

- Extension (independence testing):

  https://arxiv.org/abs/1610.04782

  https://github.com/wittawatj/fsic-test

# Thank you for the attention!

Contents

- Non-convexity, informative features.
- Number of locations (J).
- MMD: IPM representation.
- Estimation of MMD^2.
- Proof idea.
- Computational complexity: (J, n, d)-dependence.
- Why do analytic features work?

Non-convexity, informative features

- 2D problem:

  $$\mathbb{P} := \mathcal{N}(\mathbf{0}, \mathbf{I}), \quad \mathbb{Q} := \mathcal{N}(\mathbf{e}_1, \mathbf{I}).$$

- $\mathcal{V} = \{\mathbf{v}_1, \mathbf{v}_2\}$. Fix $\mathbf{v}_1$ to ▲.
- $\mathbf{v}_2 \mapsto \hat{\lambda}_n(\{\mathbf{v}_1, \mathbf{v}_2\})$: contour plot.

$\mathbf{v}_2 \mapsto \hat{\lambda}_{n/2}^{tr}(\mathbf{v}_1, \mathbf{v}_2)$

$\mathbf{v}_2 \mapsto \hat{\lambda}_{n/2}^{tr}(\mathbf{v}_1, \mathbf{v}_2)$

Non-convexity, informative features

$$\mathbf{v}_2 \mapsto \hat{\lambda}_{n/2}^{tr}(\mathbf{v}_1, \mathbf{v}_2)$$
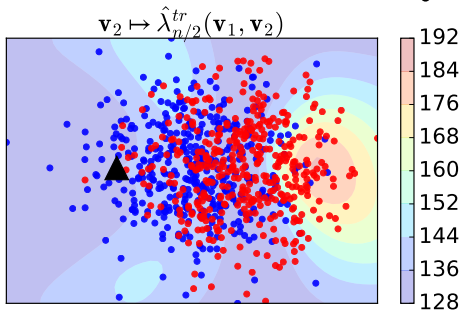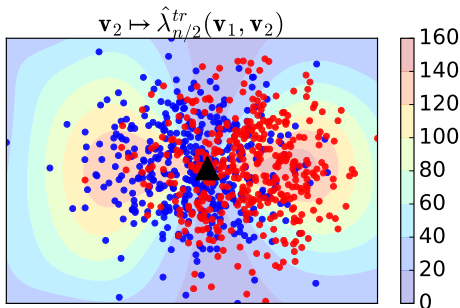
- Nearby locations: do not increase discrimininability.
- Non-convexity: reveals multiple ways to capture the difference.
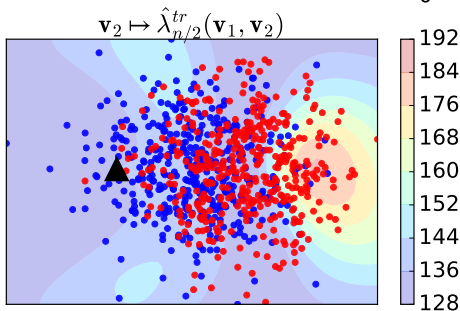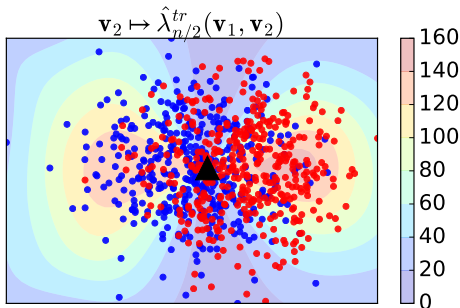
- Small J:
    - often enough to detect the difference of $\mathbb{P}$ & $\mathbb{Q}$.
    - few distinguishing regions to reject H_0.
    - faster test.

- Very large J:
    - test power need not increase monotonically in J (more locations $\Rightarrow$ statistic can gain in variance).
    - defeats the purpose of a linear-time test.

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}(k)}^2$$

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|^2_{\mathcal{H}(k)} = \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \langle \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}, f \rangle_{\mathcal{H}(k)} \right]^2$$

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|^2_{\mathcal{H}(k)} = \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \langle \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}, f \rangle_{\mathcal{H}(k)} \right]^2$$

$$\stackrel{(*)}{=} \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \mathbb{E}_{x \sim \mathbb{P}} f(x) - \mathbb{E}_{y \sim \mathbb{Q}} f(y) \right]^2.$$

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_\mathbb{P} - \mu_\mathbb{Q}\|^2_{\mathcal{H}(k)} = \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \langle \mu_\mathbb{P} - \mu_\mathbb{Q}, f\rangle_{\mathcal{H}(k)}\right]^2$$

$$\overset{(*)}{=} \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \mathbb{E}_{x\sim\mathbb{P}} f(x) - \mathbb{E}_{y\sim\mathbb{Q}} f(y)\right]^2.$$

$(*)$ in details:

$$\langle \mu_\mathbb{P}, f\rangle_{\mathcal{H}(k)} = \left\langle \int k(\cdot, x)\mathrm{d}\mathbb{P}(x), f\right\rangle_{\mathcal{H}(k)}$$

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}(k)}^2 = \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \langle \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}, f \rangle_{\mathcal{H}(k)}\right]^2$$

$$\overset{(*)}{=} \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \mathbb{E}_{x \sim \mathbb{P}} f(x) - \mathbb{E}_{y \sim \mathbb{Q}} f(y)\right]^2.$$

$(*)$ in details:

$$\langle \mu_{\mathbb{P}}, f \rangle_{\mathcal{H}(k)} = \left\langle \int k(\cdot, x) \mathrm{d}\mathbb{P}(x), f \right\rangle_{\mathcal{H}(k)} = \int \underbrace{\langle k(\cdot, x), f \rangle_{\mathcal{H}(k)}}_{=f(x)} \mathrm{d}\mathbb{P}(x)$$

$$MMD^2(\mathbb{P}, \mathbb{Q}) = \|\mu_\mathbb{P} - \mu_\mathbb{Q}\|^2_{\mathcal{H}(k)} = \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \langle \mu_\mathbb{P} - \mu_\mathbb{Q}, f \rangle_{\mathcal{H}(k)} \right]^2$$

$$\stackrel{(*)}{=} \left[\sup_{\|f\|_{\mathcal{H}(k)} \leqslant 1} \mathbb{E}_{x \sim \mathbb{P}} f(x) - \mathbb{E}_{y \sim \mathbb{Q}} f(y) \right]^2.$$

$(*)$ in details:

$$\langle \mu_\mathbb{P}, f \rangle_{\mathcal{H}(k)} = \left\langle \int k(\cdot, x) \mathrm{d}\mathbb{P}(x), f \right\rangle_{\mathcal{H}(k)} = \int \underbrace{\langle k(\cdot, x), f \rangle_{\mathcal{H}(k)}}_{=f(x)} \mathrm{d}\mathbb{P}(x)$$

$$= \mathbb{E}_{x \sim \mathbb{P}} f(x).$$

Estimation of MMD^2

Squared difference between feature means:

$$
\begin{aligned}
MMD^2(\mathbb{P}, \mathbb{Q}) &= \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}}^2 = \langle \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}, \mu_{\mathbb{P}} - \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} \\
&= \langle \mu_{\mathbb{P}}, \mu_{\mathbb{P}} \rangle_{\mathcal{H}} + \langle \mu_{\mathbb{Q}}, \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} - 2 \langle \mu_{\mathbb{P}}, \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} \\
&= \mathbb{E}_{\mathbb{P}, \mathbb{P}} k(x, x') + \mathbb{E}_{\mathbb{Q}, \mathbb{Q}} k(y, y') - 2 \mathbb{E}_{\mathbb{P}, \mathbb{Q}} k(x, y).
\end{aligned}
$$

Estimation of MMD^2

Squared difference between feature means:

$$
\begin{aligned}
MMD^2(\mathbb{P}, \mathbb{Q}) &= \|\mu_{\mathbb{P}} - \mu_{\mathbb{Q}}\|_{\mathcal{H}}^2 = \langle \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}, \mu_{\mathbb{P}} - \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} \\
&= \langle \mu_{\mathbb{P}}, \mu_{\mathbb{P}} \rangle_{\mathcal{H}} + \langle \mu_{\mathbb{Q}}, \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} - 2 \langle \mu_{\mathbb{P}}, \mu_{\mathbb{Q}} \rangle_{\mathcal{H}} \\
&= \mathbb{E}_{\mathbb{P}, \mathbb{P}} k(x, x') + \mathbb{E}_{\mathbb{Q}, \mathbb{Q}} k(y, y') - 2 \mathbb{E}_{\mathbb{P}, \mathbb{Q}} k(x, y).
\end{aligned}
$$

Unbiased empirical estimate for $\{x_i\}_{i=1}^n \sim \mathbb{P}$, $\{y_j\}_{j=1}^n \sim \mathbb{Q}$:

$$
\widehat{MMD^2}(\mathbb{P}, \mathbb{Q}) = \overline{K_{\mathbb{P}, \mathbb{P}}} + \overline{K_{\mathbb{Q}, \mathbb{Q}}} - 2\overline{K_{\mathbb{P}, \mathbb{Q}}}.
$$

1. Lower bound on the test power:
   1. $|\hat{\lambda}_n - \lambda_n| \lesssim \|\bar{\mathbf{z}}_n - \boldsymbol{\mu}\|_2 + \|\boldsymbol{\Sigma}_n - \boldsymbol{\Sigma}\|_F$.
   2. Bound the r.h.s. by Hoeffding inequality $\Rightarrow P(|\hat{\lambda}_n - \lambda_n| \geq t)$.
   3. By reparameterization: $P(\hat{\lambda}_n \geq T_\alpha)$ bound.

Proof idea

1. Lower bound on the test power:
   1. $|\hat{\lambda}_n - \lambda_n| \lesssim \|\bar{\mathbf{z}}_n - \boldsymbol{\mu}\|_2 + \|\boldsymbol{\Sigma}_n - \boldsymbol{\Sigma}\|_F$.
   2. Bound the r.h.s. by Hoeffding inequality $\Rightarrow P(|\hat{\lambda}_n - \lambda_n| \geqslant t)$.
   3. By reparameterization: $P(\hat{\lambda}_n \geqslant T_\alpha)$ bound.

2. Uniformly $\hat{\lambda}_n \approx \lambda_n$:
   - Reduction to bounding $\sup\limits_{\mathcal{V},\mathcal{K}} \|\bar{\mathbf{z}}_n - \boldsymbol{\mu}\|_2$, $\sup\limits_{\mathcal{V},\mathcal{K}} \|\boldsymbol{\Sigma}_n - \boldsymbol{\Sigma}\|_F$.
   - Empirical processes, Dudley entropy bound.

Computational complexity

- Optimization & testing: linear in n.
- Testing: $\mathcal{O}\left(ndJ + nJ^2 + J^3\right)$.
- Optimization: $\mathcal{O}\left(ndJ^2 + J^3\right)$ per gradient ascent.

Why do analytic features work?

1. k: bounded, analytic $\Rightarrow$ elements of $\mathcal{H}(k)$: analytic.

1. k: bounded, analytic $\Rightarrow$ elements of $\mathcal{H}(k)$: analytic.
2. k: characteristic, bounded $\Rightarrow \mu = \mu_k$: well-defined, injective.

Why do analytic features work?

1. k: bounded, analytic $\Rightarrow$ elements of $\mathcal{H}(k)$: analytic.
2. k: characteristic, bounded $\Rightarrow \mu = \mu_k$: well-defined, injective.
3. Since μ is injective to analytic functions

$$d^2(\mathbb{P}, \mathbb{Q}) := \sum_{j=1}^{J} \left[\mu_{\mathbb{P}}(\mathbf{v}_j) - \mu_{\mathbb{Q}}(\mathbf{v}_j) \right]^2$$

is a metric, a.s. w.r.t. $m \ll \lambda$ ($\mathbf{v}_j \sim m$). Reason: $f := \mu_{\mathbb{P}} - \mu_{\mathbb{Q}}$ analytic $\Rightarrow$

$$m\{\mathbf{v} : f(\mathbf{v}) = 0\} = 0.$$

Chwialkowski, K., Ramdas, A., Sejdinovic, D., and Gretton, A. (2015).
Fast Two-Sample Testing with Analytic Representations of Probability Measures.
In *Neural Information Processing Systems (NIPS)*, pages 1981–1989.

Gretton, A., Borgwardt, K., Rasch, M., Schölkopf, B., and Smola, A. (2012).
A kernel two-sample test.
Journal of Machine Learning Research, 13:723–773.

Jitkrittum, W., Szabó, Z., Chwialkowski, K., and Gretton, A. (2016).
Interpretable distribution features with maximum testing power.
In *Neural Information Processing Systems (NIPS)*.

Lundqvist, D., Flykt, A., and Öhman, A. (1998).
The Karolinska directed emotional faces-KDEF.