Functional Data Analysis (Lecture 4) - PCA

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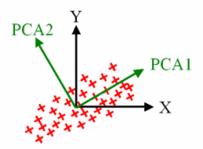
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One-page summary

- Covered topics:
 - PEN_L-regularized least squares,
 - smoothing with constraints,
 - curve registration.

One-page summary

- Today:
 - 1 dimensionality reduction,
 - ② principal component analysis (PCA; in \mathbb{R}^d first):
 - see continuous registration.



PCA

PCA - intuition

- Given: a set of observations $X = \{\mathbf{x}_i\}_{i=1}^N \subset \mathbb{R}^D$.
- Goal: find the best d-dimensional subspace approximating X.
- $d \ll D$: compression (images, music, ...).



PCA formulation: d = 1

• We are looking for the best one-dimensional projection.



- \mathbb{E} := empirical expectation (population: similarly).
- In other words: $\mathbb{E}\mathbf{x} = \frac{1}{N} \sum_{i=1}^{N} \mathbf{x}_i$.
- Assumption: $\mathbb{E}\mathbf{x} = \mathbf{0}$

PCA formulation: d = 1

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- Assumption: $\mathbb{E}\mathbf{x} = \mathbf{0}$
 - centering: $\mathbf{x} \to \mathbf{x} \mathbb{E}\mathbf{x}$.

PCA: projection

• One-dimensional projection:

• w with
$$\|\mathbf{w}\|_2 = 1$$
,
• $\hat{\mathbf{x}} = \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w}$.
=score

• Mean of the projection is zero:

$$\mathbf{0} \stackrel{?}{=} \mathbb{E} \hat{\mathbf{x}} = \mathbb{E} \left[\langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w} \right]$$

PCA: projection

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 - **w** with $\|\mathbf{w}\|_2 = 1$,

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- $\hat{\mathbf{x}} = \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w}.$
- Mean of the projection is zero:

$$\mathbf{0} \stackrel{?}{=} \mathbb{E} \hat{\mathbf{x}} = \mathbb{E} \left[\left\langle \mathbf{w}, \mathbf{x} \right\rangle \mathbf{w} \right] = \left\langle \mathbf{w}, \underbrace{\mathbb{E} \mathbf{x}}_{=\mathbf{0}} \right\rangle \mathbf{w}.$$

$$\|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 = \|\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w}\|_2^2$$

$$\begin{split} \|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 &= \|\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w}\|_2^2 \\ &= \left(\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w}\right)^T \left(\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w}\right) \end{split}$$

$$\|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} = \|\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w}\|_{2}^{2}$$

$$= (\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w})^{T} (\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w})$$

$$= \|\mathbf{x}\|_{2}^{2} - 2 \langle \mathbf{w}, \mathbf{x} \rangle^{2} + \langle \mathbf{w}, \mathbf{x} \rangle^{2} \|\mathbf{w}\|_{2}^{2}$$

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$$= \|\mathbf{x}\|_{2}^{2} - \langle \mathbf{w}, \mathbf{x} \rangle^{2}.$$

Residual:

$$\|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} = \|\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w}\|_{2}^{2}$$

$$= (\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w})^{T} (\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \mathbf{w})$$

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MSE:

$$\min_{\mathbf{w}} \leftarrow \mathbb{E} \left\| \mathbf{x} - \hat{\mathbf{x}} \right\|_{2}^{2} = \mathbb{E} \left[\left\| \mathbf{x} \right\|_{2}^{2} - \left\langle \mathbf{w}, \mathbf{x} \right\rangle^{2} \right]$$

Residual:

$$\begin{aligned} \left\| \mathbf{x} - \hat{\mathbf{x}} \right\|_{2}^{2} &= \left\| \mathbf{x} - \left\langle \mathbf{w}, \mathbf{x} \right\rangle \mathbf{w} \right\|_{2}^{2} \\ &= \left(\mathbf{x} - \left\langle \mathbf{w}, \mathbf{x} \right\rangle \mathbf{w} \right)^{T} \left(\mathbf{x} - \left\langle \mathbf{w}, \mathbf{x} \right\rangle \mathbf{w} \right) \\ &= \left\| \mathbf{x} \right\|_{2}^{2} - 2 \left\langle \mathbf{w}, \mathbf{x} \right\rangle^{2} + \left\langle \mathbf{w}, \mathbf{x} \right\rangle^{2} \underbrace{\left\| \mathbf{w} \right\|_{2}^{2}}_{=1} \\ &= \left\| \mathbf{x} \right\|_{2}^{2} - \left\langle \mathbf{w}, \mathbf{x} \right\rangle^{2}. \end{aligned}$$

MSE:

$$\begin{split} \min_{\mathbf{w}} \leftarrow \mathbb{E} \left\| \mathbf{x} - \hat{\mathbf{x}} \right\|_2^2 &= \mathbb{E} \left[\left\| \mathbf{x} \right\|_2^2 - \left\langle \mathbf{w}, \mathbf{x} \right\rangle^2 \right] \\ &= \underbrace{\mathbb{E} \left\| \mathbf{x} \right\|_2^2}_{\text{independent of } \mathbf{w}} - \mathbb{E} \left\langle \mathbf{w}, \mathbf{x} \right\rangle^2. \end{split}$$

Residual:

$$\begin{aligned} \|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} &= \|\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w}\|_{2}^{2} \\ &= \left(\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w}\right)^{T} \left(\mathbf{x} - \langle \mathbf{w}, \mathbf{x} \rangle \, \mathbf{w}\right) \\ &= \|\mathbf{x}\|_{2}^{2} - 2 \, \langle \mathbf{w}, \mathbf{x} \rangle^{2} + \langle \mathbf{w}, \mathbf{x} \rangle^{2} \underbrace{\|\mathbf{w}\|_{2}^{2}}_{=1} \\ &= \|\mathbf{x}\|_{2}^{2} - \langle \mathbf{w}, \mathbf{x} \rangle^{2} \, . \end{aligned}$$

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⇔ Maximize the mean squared projection.

By using
$$\mathbb{E}y^2 = (\mathbb{E}y)^2 + var(y)$$
:

$$\max_{\mathbf{w}} \leftarrow \mathbb{E} \langle \mathbf{w}, \mathbf{x} \rangle^2 = \left(\mathbb{E} \langle \mathbf{w}, \mathbf{x} \rangle \right)^2 + var(\langle \mathbf{w}, \mathbf{x} \rangle).$$

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To sum up:

Minimize MSE of the residual :
$$\mathbb{E} \|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 \rightarrow \min_{\mathbf{w}} \Leftrightarrow$$

By using
$$\mathbb{E}y^2 = (\mathbb{E}y)^2 + var(y)$$
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To sum up:

Minimize MSE of the residual :
$$\mathbb{E} \|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 \to \min_{\mathbf{w}} \Leftrightarrow$$
 Maximize mean squared projection : $\mathbb{E} \langle \mathbf{w}, \mathbf{x} \rangle^2 \to \max_{\mathbf{w}} \Leftrightarrow$

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To sum up:

Minimize MSE of the residual :
$$\mathbb{E} \|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 \to \min_{\mathbf{w}} \Leftrightarrow$$
 Maximize mean squared projection : $\mathbb{E} \langle \mathbf{w}, \mathbf{x} \rangle^2 \to \max_{\mathbf{w}} \Leftrightarrow$ Maximize variance of the projection : $var(\langle \mathbf{w}, \mathbf{x} \rangle) \to \max_{\mathbf{w}}$.

Using the bilinearity of covariance:

$$var\left(\left\langle \mathbf{w},\mathbf{x}\right
angle
ight) = cov\left(\left\langle \mathbf{w},\mathbf{x}
ight
angle ,\left\langle \mathbf{w},\mathbf{x}
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$$var(\langle \mathbf{w}, \mathbf{x} \rangle) = cov(\langle \mathbf{w}, \mathbf{x} \rangle, \langle \mathbf{w}, \mathbf{x} \rangle)$$
$$= cov(\mathbf{w}^T \mathbf{x}, \mathbf{w}^T \mathbf{x})$$

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Lagrange function, solving for 'derivatives = 0':

•

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$$L(\mathbf{w}, \lambda) = \underbrace{\mathbf{w}^T \mathbf{\Sigma} \mathbf{w}}_{\text{=objective}} - \lambda(\underbrace{\mathbf{w}^T \mathbf{w} - 1}_{\text{=condition}}) \Rightarrow$$

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To sum up: $\Sigma \mathbf{w} = \lambda \mathbf{w}$, $\|\mathbf{w}\|_2 = 1$. $\Rightarrow \mathbf{w}^*$: eigenvector associated to $\lambda_{\max}(\Sigma)$.

PCA: $d \geqslant 1$

PCA $(d \ge 1)$: basis, approximation

- Goal: approximate with a d-dimensional subspace.
- ONB in the subspace: $\{\mathbf{w}_i\}_{i=1}^d$, $\langle \mathbf{w}_i, \mathbf{w}_j \rangle = \delta_{ij}$.
- Approximation ($\mathbf{W} = [\mathbf{w}_1, \dots, \mathbf{w}_d]$):

$$\hat{\mathbf{x}} = \sum_{i=1}^{d} \underbrace{\langle \mathbf{w}_i, \mathbf{x} \rangle}_{=\text{scores}} \mathbf{w}_i = \mathbf{W} \mathbf{W}^T \mathbf{x}.$$

$$\|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 = \|\mathbf{x} - \mathbf{W}\mathbf{W}^T\mathbf{x}\|_2^2$$

$$\|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} = \|\mathbf{x} - \mathbf{W}\mathbf{W}^{\mathsf{T}}\mathbf{x}\|_{2}^{2} = (\mathbf{x} - \mathbf{W}\mathbf{W}^{\mathsf{T}}\mathbf{x})^{\mathsf{T}}(\mathbf{x} - \mathbf{W}\mathbf{W}^{\mathsf{T}}\mathbf{x})$$

Using $\mathbf{W}^T\mathbf{W} = \mathbf{I}$

$$\|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} = \|\mathbf{x} - \mathbf{W}\mathbf{W}^{T}\mathbf{x}\|_{2}^{2} = (\mathbf{x} - \mathbf{W}\mathbf{W}^{T}\mathbf{x})^{T} (\mathbf{x} - \mathbf{W}\mathbf{W}^{T}\mathbf{x})$$

$$= \mathbf{x}^{T} \underbrace{(\mathbf{I} - \mathbf{W}\mathbf{W}^{T}) (\mathbf{I} - \mathbf{W}\mathbf{W}^{T})}_{=\mathbf{I} - 2\mathbf{W}\mathbf{W}^{T} + \mathbf{W}\mathbf{W}^{T}\mathbf{W}\mathbf{W}^{T} = \mathbf{I} - \mathbf{W}\mathbf{W}^{T}}_{=\mathbf{I} - 2\mathbf{W}\mathbf{W}^{T}}$$

Using $\mathbf{W}^T\mathbf{W} = \mathbf{I}$

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Using $\mathbf{W}^T\mathbf{W} = \mathbf{I}$

$$\begin{split} \|\mathbf{x} - \hat{\mathbf{x}}\|_2^2 &= \left\|\mathbf{x} - \mathbf{W}\mathbf{W}^T\mathbf{x}\right\|_2^2 = \left(\mathbf{x} - \mathbf{W}\mathbf{W}^T\mathbf{x}\right)^T \left(\mathbf{x} - \mathbf{W}\mathbf{W}^T\mathbf{x}\right) \\ &= \mathbf{x}^T \underbrace{\left(\mathbf{I} - \mathbf{W}\mathbf{W}^T\right) \left(\mathbf{I} - \mathbf{W}\mathbf{W}^T\right)}_{=\mathbf{I} - 2\mathbf{W}\mathbf{W}^T + \mathbf{W}\mathbf{W}^T\mathbf{W}\mathbf{W}^T = \mathbf{I} - \mathbf{W}\mathbf{W}^T} \\ &= \left\|\mathbf{x}\right\|_2^2 - \left\|\mathbf{W}^T\mathbf{x}\right\|_2^2, \\ \mathbb{E} \left\|\mathbf{x} - \hat{\mathbf{x}}\right\|_2^2 = \mathbb{E} \left\|\mathbf{x}\right\|_2^2 - \mathbb{E} \left\|\mathbf{W}^T\mathbf{x}\right\|_2^2. \end{split}$$

Using $\mathbf{W}^T\mathbf{W} = \mathbf{I}$

$$\begin{split} \|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} &= \left\|\mathbf{x} - \mathbf{W}\mathbf{W}^{T}\mathbf{x}\right\|_{2}^{2} = \left(\mathbf{x} - \mathbf{W}\mathbf{W}^{T}\mathbf{x}\right)^{T} \left(\mathbf{x} - \mathbf{W}\mathbf{W}^{T}\mathbf{x}\right) \\ &= \mathbf{x}^{T} \underbrace{\left(\mathbf{I} - \mathbf{W}\mathbf{W}^{T}\right) \left(\mathbf{I} - \mathbf{W}\mathbf{W}^{T}\right)}_{=\mathbf{I} - 2\mathbf{W}\mathbf{W}^{T} + \mathbf{W}\mathbf{W}^{T}\mathbf{W}\mathbf{W}^{T} = \mathbf{I} - \mathbf{W}\mathbf{W}^{T}} \mathbf{x} \\ &= \|\mathbf{x}\|_{2}^{2} - \left\|\mathbf{W}^{T}\mathbf{x}\right\|_{2}^{2}, \\ \mathbb{E} \left\|\mathbf{x} - \hat{\mathbf{x}}\right\|_{2}^{2} = \mathbb{E} \left\|\mathbf{x}\right\|_{2}^{2} - \mathbb{E} \left\|\mathbf{W}^{T}\mathbf{x}\right\|_{2}^{2}. \end{split}$$

Thus $\min_{\mathbf{w}} \mathbb{E} \|\mathbf{x} - \hat{\mathbf{x}}\|_{2}^{2} \Leftrightarrow \max_{\mathbf{w}} \mathbb{E} \|\mathbf{W}^{T}\mathbf{x}\|_{2}^{2}$.

PCA $(d \ge 1)$: mean squared projection \rightarrow variance

Let
$$\mathbf{y} = \mathbf{W}^T \mathbf{x}$$
:

$$\mathbb{E} \|\mathbf{y}\|_{2}^{2} - \|\mathbb{E}\mathbf{y}\|_{2}^{2} = var(\mathbf{y})?$$

PCA $(d \ge 1)$: mean squared projection \rightarrow variance

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$$= \mathbb{E} \left[\sum_{i} y_{i}^{2} \right] - \sum_{i} (\mathbb{E}y_{i})^{2} = \sum_{i} var(y_{i}) \Rightarrow$$

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$$\mathbb{E} \|\mathbf{W}^T \mathbf{x}\|_2^2 - \|\mathbb{E} [\mathbf{W}^T \mathbf{x}]\|_2^2 = \sum_i var((\mathbf{W}^T \mathbf{x})_i) \to \max_{\mathbf{W}}.$$

 $=W^T\mathbb{E}x=0$

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- Variance decomposition: $cov(\mathbf{x}) = \sum_{i=1}^{D} \lambda_i \mathbf{w}_i \mathbf{w}_i^T$.

PCA: $d \geqslant 1$

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- Variance decomposition: $cov(\mathbf{x}) = \sum_{i=1}^{D} \lambda_i \mathbf{w}_i \mathbf{w}_i^T$.
- Energy preserved using d components: $\sum_{i=1}^{d} \lambda_i \Rightarrow$

$$R^2 = R^2(d) := \frac{\sum_{i=1}^d \lambda_i}{\sum_{i=1}^D \lambda_i} \in [0, 1].$$

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- Energy preserved using d components: $\sum_{i=1}^{d} \lambda_i \Rightarrow$

$$R^2 = R^2(d) := \frac{\sum_{i=1}^d \lambda_i}{\sum_{i=1}^D \lambda_i} \in [0, 1].$$

• In practice: choose d such that $R^2 \approx 0.8 - 0.9$.

PCA: alternative view/definition

Recursive formulation (deflation approach):

- Assume: we have \mathbf{w}_0 .
- Ask for the first PC of the residual (\mathbf{w}_1) .
- Iterate.

This leads to an equivalent definition.

Summary

- Dimensionality reduction.
- Approximation with a 'small' dimensional subspace.
- Top eigenvectors/-values of the observation covariance.

We covered the ' \mathbb{R}^d part' of Chapter 8 in [1], and [3].