Beyond Mean Embedding: Cumulants in RKHSs

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Learning with Kernels and Non-linear Transformations session, CMStatistics Dec. 17, 2023

Moments and cumulants on $\mathbb{R} \ni X \sim \gamma$

• Moments $\mu(\gamma) := \left(\mu^{(i)}(\gamma)\right)_{i \in \mathbb{N}}$:

$$\mu^{(i)}(\gamma) := \mathbb{E}\left(X^i\right) \in \mathbb{R}, \qquad \qquad \mu^{(0)}(\gamma) := 1.$$

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$$\sum_{i\in\mathbb{N}} \kappa^{(i)}(\gamma) \frac{\theta^i}{i!} = \log\left(\sum_{i\in\mathbb{N}} \mu^{(i)}(\gamma) \frac{\theta^i}{i!}\right).$$

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$$\kappa^{(1)}(\gamma) = \mathbb{E}(X) \qquad \text{mean}$$

$$\kappa^{(2)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^2 \qquad \text{variance}$$

$$\kappa^{(3)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^3 \qquad \text{3rd central moment}$$

$$\kappa^{(4)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^4 - 3\left[\mathbb{E}(X - \mathbb{E}X)^2\right]^2$$

$$\kappa^{(5)}(\gamma) = \mathbb{E}(X - \mathbb{E}X)^5 - 10\mathbb{E}(X - \mathbb{E}X)^3\mathbb{E}(X - \mathbb{E}X)^2$$

$$\kappa^{(1)}(\gamma) = \mathbb{E}(X), \tag{\{1\}}$$

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{{1}}}

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$$\kappa^{(2)}(\gamma) = \mathbb{E}(X^2) - \widetilde{\mathbb{E}^2(X)}, \qquad \{\{1,2\}\}, \{\{1\}, \{2\}\}\}$$

where $X, X' \sim \gamma$, independent.

$$\kappa^{(1)}(\gamma) = \mathbb{E}(X),$$

$$\kappa^{(2)}(\gamma) = \mathbb{E}\left(X^2\right) - \underbrace{\mathbb{E}^{(XX')}}_{\mathbb{E}(X)},$$

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Question

. . .

What are the weights in front of the moments?

Unzipping cumulants on \mathbb{R} : the weights

m	elements of $\pi \in P(m)$	$ \pi $	c_{π}
1	{1}	1	1
2	{1,2}	1	1
	{1},{2}	2	-1
3	{1,2,3}	1	1
	{1,2}, {3}	2	-1
	{1,3}, {2}	2	-1
	{2,3}, {1}	2	-1
	{1}, {2}, {3}	3	2

with P(m) := all partitions of [m], $c_{\pi} = (-1)^{|\pi|-1}(|\pi|-1)!$

Motivation, i.e. one reason why one likes cumulants

Moment and cumulants on \mathbb{R}^d

Change
$$\mathbb{E}\left(X^i\right)\in\mathbb{R}$$
 to $\mathbb{E}\left[X_1^{i_1}\cdots X_d^{i_d}\right]\in\mathbb{R}$ ($\mathbf{i}\in\mathbb{N}^d$). $(\mathbf{og},P(m):\checkmark)$

Known theorem [Billingsley, 2012]

Let γ be a probability measure on a bounded subset of \mathbb{R}^d with cumulants $\kappa(\gamma)$ and let $(X_1, \ldots, X_d) \sim \gamma$. Then

- \bullet $\gamma \mapsto \kappa(\gamma)$ is injective.
- X_1, \ldots, X_d are independent $\Leftrightarrow \kappa^{\mathbf{i}}(\gamma) = 0$ for all $\mathbf{i} \in \mathbb{N}_+^d$.

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- X_1, \ldots, X_d are independent $\Leftrightarrow \kappa^{\mathbf{i}}(\gamma) = 0$ for all $\mathbf{i} \in \mathbb{N}^d_+$.

Motivation

- Various data types, nonlinear features: kernels.
- 2 Linear: not even characteristic (see MMD and HSIC).
- 3 Computable estimators.

Idea

Lifting

$$(X_1,\ldots,X_d)\in\times_{j=1}^d\mathcal{X}_j\to \left(\Phi_1(X_1),\ldots,\Phi_d(X_d)\right)\in\times_{j=1}^d\mathcal{H}_{k_j}.$$

Lifting

$$(X_1,\ldots,X_d)\in imes_{j=1}^d \mathcal{X}_j o (\Phi_1(X_1),\ldots,\Phi_d(X_d))\in imes_{j=1}^d \mathcal{H}_{k_j}.$$

Ingredients:

$$\mathbb{E}\left[\left[\Phi_1(X_1)\right]^{\otimes i_1}\otimes\cdots\otimes\left[\Phi_d(X_d)\right]^{\otimes i_d}\right]\in\mathcal{H}_{k_1}^{\otimes i_1}\otimes\cdots\otimes\mathcal{H}_{k_d}^{\otimes i_d}.$$

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- From moments to cumulants:
 - log on tensor algebras, or
 - combinatorial description of cumulants (\leftarrow a bit simpler, but \Leftrightarrow).

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- From moments to cumulants:
 - log on tensor algebras, or
 - combinatorial description of cumulants (\leftarrow a bit simpler, but \Leftrightarrow).
- 3 Computation: by the 'expected kernel trick' (V-statistics).

Kernel (generalization of $\mathbf{a}^{\mathsf{T}}\mathbf{b}$), RKHS

• Def-1 (feature space):

$$k(a,b) = \langle \Phi(a), \Phi(b) \rangle_{\mathcal{H}}.$$

• Def-2 (reproducing kernel):

$$k(\cdot, b) \in \mathcal{H},$$
 $f(b) = \langle f, k(\cdot, b) \rangle_{\mathcal{H}}.$

• Def-3 (Gram matrix): $\mathbf{G} = [k(x_i, x_j)]_{i,j=1}^n \in \mathbb{R}^{n \times n} \succeq \mathbf{0}$.

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Notes

- $k \stackrel{1:1}{\leftrightarrow} \mathcal{H}_k = \overline{\mathsf{Span}}(k(\cdot, x) : x \in \mathcal{X})$: Fourier analysis, polynomials, splines, . . .
- Examples: $k_p(\mathbf{x}, \mathbf{y}) = (\langle \mathbf{x}, \mathbf{y} \rangle + \gamma)^p$, $k_G(\mathbf{x}, \mathbf{y}) = e^{-\gamma \|\mathbf{x} \mathbf{y}\|_2^2}$.
- Kernels exist on various domains!

Mean embedding

• Mean embedding (Bochner integral):

$$\mu_k(\mathbb{P}) := \int_{\mathcal{X}} \underbrace{k(\cdot, x)}_{\Phi(x) \in \mathcal{H}_k} d\mathbb{P}(x) \in \mathcal{H}_k.$$

Mean embedding, MMD

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• Maximum mean discrepancy:

$$\mathsf{MMD}_{k}(\mathbb{P},\mathbb{Q}) := \|\mu_{k}(\mathbb{P}) - \mu_{k}(\mathbb{Q})\|_{\mathcal{H}_{k}}.$$

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• Hilbert-Schmidt independence criterion, $k := \bigotimes_{j=1}^{d} k_j$:

$$\mathsf{HSIC}_k\left(\mathbb{P}\right) := \mathsf{MMD}_k\left(\mathbb{P}, \otimes_{j=1}^d \mathbb{P}_j\right)$$

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$$\begin{aligned} \mathsf{HSIC}_{\pmb{k}}\left(\mathbb{P}\right) &:= \mathsf{MMD}_{\pmb{k}}\left(\underline{\mathbb{P}}, \otimes_{j=1}^{\pmb{d}} \mathbb{P}_{j}\right) \\ &= \left\| \underbrace{\mu_{\otimes_{j=1}^{\pmb{d}} k_{j}}(\mathbb{P}) - \otimes_{j=1}^{\pmb{d}} \mu_{k_{j}}\left(\mathbb{P}_{j}\right)}_{\mathsf{cross-covariance operator}} \right\|_{\mathcal{H}_{\pmb{k}}}. \end{aligned}$$

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Clarification of what $\bigotimes_{j=1}^d k_j$ and $\bigotimes_{j=1}^d \mu_{k_j}(\mathbb{P}_j)$ are follows.

Tensor product: $\bigotimes_{j=1}^{d} a_j$

• If $\mathbf{a} \in \mathbb{R}^{n_1}$, $\mathbf{b} \in \mathbb{R}^{n_2}$:

$$\mathbb{R}\ni \mathbf{v}^\top \begin{pmatrix} \mathbf{a} \mathbf{b}^\top \end{pmatrix} \mathbf{w} = \begin{pmatrix} \mathbf{v}^\top \mathbf{a} \end{pmatrix} \begin{pmatrix} \mathbf{b}^\top \mathbf{w} \end{pmatrix} = \langle \mathbf{a}, \mathbf{v} \rangle_{\mathbb{R}^{n_1}} \langle \mathbf{b}, \mathbf{w} \rangle_{\mathbb{R}^{n_2}},$$

 $\mathbf{a} \otimes \mathbf{b} := \mathbf{a} \mathbf{b}^{\top}$ is an $\mathbb{R}^{n_1} \times \mathbb{R}^{n_2} \to \mathbb{R}$ bilinear form.

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- For $a \in \mathcal{H}_1$, $b \in \mathcal{H}_2$ Hilbert spaces, i.e. for d = 2:

$$(a \otimes b)(v, w) := \langle a, v \rangle_{\mathcal{H}_1} \langle b, w \rangle_{\mathcal{H}_2}.$$

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• For $d \ge 2$ and $a_i \in \mathcal{H}_i$,

$$\left(\otimes_{j=1}^d a_j\right)(b_1,\ldots,b_d) := \prod_{i=1}^d \langle a_i,b_i \rangle_{\mathcal{H}_j}.$$

Tensor product: $\bigotimes_{j=1}^{d} \mathcal{H}_{j}$

$$\otimes_{j=1}^d \mathcal{H}_j := \overline{\mathsf{Span}} (\otimes_{j=1}^d \mathsf{a}_j \,:\, \mathsf{a}_j \in \mathcal{H}_j), \; \langle \otimes_{j=1}^d \mathsf{a}_j, \otimes_{j=1}^d b_j \rangle := \prod_{j=1}^d \langle \mathsf{a}_j, b_j
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 $\xrightarrow{\text{spec.}}$ The tensor product of RKHSs is an RKHS

$$\mathcal{H}_k = \bigotimes_{j=1}^d \mathcal{H}_{k_j},$$

$$k(x, x') := (\bigotimes_{j=1}^d k_j)(x, x') := \prod_{j=1}^d \underbrace{k_j(x_j, x'_j)}_{\text{coordinate-wise similarity}}$$

Validness:

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Properties:

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Properties:

- Injectivity of μ_k on probability / finite signed measures, so universal \Rightarrow characteristic.
- 2 Easy-to-estimate: expected kernel trick

$$\langle \mu_k(\mathbb{P}), \mu_k(\mathbb{Q}) \rangle_{\mathcal{H}_k} = \int_{\mathcal{X}} \int_{\mathcal{X}} k(x, y) d\mathbb{P}(x) d\mathbb{Q}(y).$$

Kernelized moments – towards kernelized cumulants

- From now:
 - $X = (X_i)_{i=1}^d \in \times_{i=1}^d \mathcal{X}_i, X \sim \gamma$,
 - kernels $k_i : \mathcal{X}_i \times \mathcal{X}_i \to \mathbb{R}, j \in [d],$
 - lifting $\Phi(X) = (\Phi_j(X_j))_{j=1}^d$ with $\Phi_j(x_j) := k_j(\cdot, x_j)$,
 - RKHS $\mathcal{H}^{\otimes i} := \mathcal{H}_{k_1}^{\otimes i_1} \otimes \cdots \otimes \mathcal{H}_{k_d}^{\otimes i_d}$ with kernel $k^{\otimes i} := k_1^{\otimes i_1} \otimes \cdots \otimes k_d^{\otimes i_d}$, and feature

$$\Phi^{\otimes i}(X) := \left[\Phi_1(X_1)\right]^{\otimes i_1} \otimes \cdots \otimes \left[\Phi_d(X_d)\right]^{\otimes i_d}.$$

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 - RKHS $\mathcal{H}^{\otimes \mathbf{i}} := \mathcal{H}_{k_1}^{\otimes i_1} \otimes \cdots \otimes \mathcal{H}_{k_d}^{\otimes i_d}$ with kernel $k^{\otimes \mathbf{i}} := k_1^{\otimes i_1} \otimes \cdots \otimes k_d^{\otimes i_d}$, and feature

$$\Phi^{\otimes i}(X) := [\Phi_1(X_1)]^{\otimes i_1} \otimes \cdots \otimes [\Phi_d(X_d)]^{\otimes i_d}.$$

• Moment sequence:

$$\mu(\gamma) = \left(\mu^{\mathsf{i}}(\gamma)\right)_{\mathsf{i} \in \mathbb{N}^d}, \qquad \quad \mu^{\mathsf{i}}(\gamma) := \mathbb{E}\left[\Phi^{\otimes \mathsf{i}}(X)\right] \in \mathcal{H}^{\otimes \mathsf{i}}.$$

Kernelized cumulants: examples first, analogous to $\mathbb R$

• $d = 1, m \in [3]: X \sim \gamma$,

$$\kappa_k^{(1)}(\gamma) = \mathbb{E}\big[\Phi(X)\big]$$

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•
$$d = 2$$
, $m = 2$: $(X_1, X_2) \sim \gamma$,

$$\kappa_{k_1,k_2}^{(2,0)}(\gamma) = \mathbb{E}\left[\Phi_1^{\otimes 2}(X_1)\right] - \mathbb{E}^{\otimes 2}\left[\Phi_1(X_1)\right],$$

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\kappa_{k_1,k_2}^{(1,1)}(\gamma) = \mathbb{E}\left[\Phi_1(X_1) \otimes \Phi_2(X_2)\right] - \mathbb{E}\left[\Phi_1(X_1)\right] \otimes \mathbb{E}\left[\Phi_2(X_2)\right]$$

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Wanted: repetition and partitioning. Weights: as before (c_{π}) .

Kernelized cumulants: $X \sim \gamma$ prob. measure on $\times_{j=1}^{d} \mathcal{X}_{j}$

• Repetition (diagonal measure): $\mathbf{i} \in \mathbb{N}^d$,

$$\gamma^{\mathbf{i}} \coloneqq \mathsf{Law}(\underbrace{X_1, \dots, X_1}_{i_1 \text{ times}}, \underbrace{X_2, \dots, X_2}_{i_2 \text{ times}}, \dots, \underbrace{X_d, \dots, X_d}_{i_d \text{ times}}).$$

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• Kernelized cumulants: $m = \deg(\mathbf{i}) := \sum_{j=1}^{d} i_j \xrightarrow{\mathsf{OK}} \gamma_{\pi}^{\mathbf{i}} = (\gamma^{\mathbf{i}})_{\pi}$

$$\kappa_{k_1,\dots,k_d}(\gamma) := \left(\kappa_{k_1,\dots,k_d}^{\mathbf{i}}(\gamma)\right)_{\mathbf{i}\in\mathbb{N}^d},$$

$$\kappa_{k_1,\dots,k_d}^{\mathbf{i}}(\gamma) := \sum_{\pi\in P(m)} c_{\pi} \mathbb{E}_{\gamma_{\pi}^{\mathbf{i}}} k^{\otimes \mathbf{i}}(\cdot, (X_1,\dots,X_m)).$$

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⇒ expected kernel trick is applicable

Point-separating k := injectivity of $\Phi \Leftarrow$ characteristic $k \Leftarrow$ universal k.

Point-separating k := injectivity of Φ \Leftarrow characteristic k \Leftarrow universal k.

- Assume:
 - γ , η : probability measures on $\times_{j=1}^{d} \mathcal{X}_{j}$,
 - $(\mathcal{X}_j)_{j=1}^d$ are Polish spaces,
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- Then, $\gamma = \eta \Leftrightarrow \kappa_{k_1,\dots,k_d}(\gamma) = \kappa_{k_1,\dots,k_d}(\eta)$, and

$$\begin{split} d^{\mathbf{i}}(\gamma,\eta) &:= \|\kappa_{k_{1},\dots,k_{d}}^{\mathbf{i}}(\gamma) - \kappa_{k_{1},\dots,k_{d}}^{\mathbf{i}}(\eta)\|_{\mathcal{H}^{\otimes \mathbf{i}}}^{2} \\ &= \sum_{\pi,\tau \in P(m)} c_{\pi} c_{\tau} \Big[\mathbb{E}_{\gamma_{\pi}^{\mathbf{i}} \otimes \gamma_{\tau}^{\mathbf{i}}} k^{\otimes \mathbf{i}} ((X_{1},\dots,X_{m}),(Y_{1},\dots,Y_{m})) \\ &+ \mathbb{E}_{\eta_{\pi}^{\mathbf{i}} \otimes \eta_{\tau}^{\mathbf{i}}} k^{\otimes \mathbf{i}} ((X_{1},\dots,X_{m}),(Y_{1},\dots,Y_{m})) \\ &- 2 \mathbb{E}_{\gamma_{\pi}^{\mathbf{i}} \otimes \eta_{\tau}^{\mathbf{i}}} k^{\otimes \mathbf{i}} ((X_{1},\dots,X_{m}),(Y_{1},\dots,Y_{m})) \Big]. \end{split}$$

Cumulants characterize independence

- Assume:
 - ullet γ : probability measure on $imes_{j=1}^d \mathcal{X}_j$,
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Theorem

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$$\|\kappa_{k_1,\ldots,k_d}^{\mathbf{i}}(\gamma)\|_{\mathcal{H}^{\otimes \mathbf{i}}}^2 = \sum_{\pi,\tau\in P(m)} c_{\pi} c_{\tau} \mathbb{E}_{\gamma_{\pi}^{\mathbf{i}}\otimes\gamma_{\tau}^{\mathbf{i}}} k^{\otimes \mathbf{i}}((X_j)_{j=1}^m, (Y_j)_{j=1}^m),$$

where $m = \deg(\mathbf{i})$.

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- Then, $\gamma = \gamma|_{\mathcal{X}_1} \otimes \cdots \otimes \gamma|_{\mathcal{X}_d} \Leftrightarrow \kappa^{\mathbf{i}}_{k_1,\dots,k_d}(\gamma) = 0$ for every $\mathbf{i} \in \mathbb{N}_+^d$, and

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where $m = \deg(\mathbf{i})$.

Estimation in both cases

$$\mathbb{E} k^{\otimes i}((X_1,\ldots,X_m),(Y_1,\ldots,Y_m)) \Rightarrow V$$
-statistics \checkmark

Distance between kernel variance embeddings

- By our theorem: if $\gamma = \eta$, then $d^{(2)}(\gamma, \eta) = 0$.
- V-statistic estimator of $d^{(2)}(\gamma, \eta)$:

$$\frac{1}{N^2}\mathrm{Tr}\!\left[(\boldsymbol{\mathsf{K}}_{x}\boldsymbol{\mathsf{J}}_{N})^2\right] + \frac{1}{M^2}\mathrm{Tr}\!\left[(\boldsymbol{\mathsf{K}}_{y}\boldsymbol{\mathsf{J}}_{M})^2\right] - \frac{2}{NM}\mathrm{Tr}\!\left[\boldsymbol{\mathsf{K}}_{xy}\boldsymbol{\mathsf{J}}_{M}\boldsymbol{\mathsf{K}}_{xy}^{\top}\boldsymbol{\mathsf{J}}_{N}\right],$$

with
$$(x_n)_{n=1}^N \overset{\text{i.i.d.}}{\sim} \gamma$$
, $(y_m)_{m=1}^M \overset{\text{i.i.d.}}{\sim} \eta$, $\mathbf{K}_x = [k(x_i, x_j)]_{i,j=1}^N$, $\mathbf{K}_y = [k(y_i, y_j)]_{i,j=1}^M$, $\mathbf{K}_{x,y} = [k(x_i, y_j)]_{i,j=1}^{N,M}$, $\mathbf{J}_n = \mathbf{I}_n - \mathbf{H}_n = \mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^\top$.

Distance between kernel variance/skewness embeddings

- By our theorem: if $\gamma = \eta$, then $d^{(2)}(\gamma, \eta) = 0$.
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with
$$(x_n)_{n=1}^{N} \stackrel{\text{i.i.d.}}{\sim} \gamma$$
, $(y_m)_{m=1}^{M} \stackrel{\text{i.i.d.}}{\sim} \eta$, $\mathbf{K}_x = [k(x_i, x_j)]_{i,j=1}^{N}$, $\mathbf{K}_y = [k(y_i, y_j)]_{i,j=1}^{M}$, $\mathbf{K}_{x,y} = [k(x_i, y_j)]_{i,j=1}^{N,M}$, $\mathbf{J}_n = \mathbf{I}_n - \mathbf{H}_n = \mathbf{I}_n - \frac{1}{n} \mathbf{1}_n \mathbf{1}_n^{\top}$.

Time complexity

Quadratic as MMD.

• $d^{(3)}(\gamma, \eta)$: similarly; quadratic time.

Cross-skewness independence criterion (CSIC)

- By our theorem: if $\gamma = \gamma|_{\mathcal{X}_1} \otimes \gamma|_{\mathcal{X}_2}$, then $\kappa_{k,\ell}^{(2,1)}(\gamma) = 0$ and $\kappa_{k,\ell}^{(1,2)}(\gamma) = 0$.
- V-statistic estimator of $\|\kappa_{k,\ell}^{(1,2)}(\gamma)\|_{\mathcal{H}_k^{\otimes 1}\otimes\mathcal{H}_\ell^{\otimes 2}}^2$:

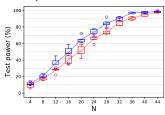
$$\begin{split} &\frac{1}{N^2} \left\langle \mathbf{K} \circ \mathbf{K} \circ \mathbf{L} - 4\mathbf{K} \circ \mathbf{K} \mathbf{H} \circ \mathbf{L} - 2\mathbf{K} \circ \mathbf{K} \circ \mathbf{L} \mathbf{H} + 4\mathbf{K} \mathbf{H} \circ \mathbf{K} \circ \mathbf{L} \mathbf{H} \right. \\ &+ 2\mathbf{K} \circ \mathbf{L} \left\langle \frac{\mathbf{K}}{N^2} \right\rangle + 2\mathbf{K} \mathbf{H} \circ \mathbf{H} \mathbf{K} \circ \mathbf{L} + 4\mathbf{K} \circ \mathbf{H} \mathbf{K} \circ \mathbf{L} \mathbf{H} + \mathbf{K} \circ \mathbf{K} \left\langle \frac{\mathbf{L}}{N^2} \right\rangle \\ &- 8\mathbf{K} \circ \mathbf{L} \mathbf{H} \left\langle \frac{\mathbf{K}}{N^2} \right\rangle - 4\mathbf{K} \circ \mathbf{H} \mathbf{K} \left\langle \frac{\mathbf{L}}{N^2} \right\rangle + 4 \left\langle \frac{\mathbf{K}}{N^2} \right\rangle^2 \mathbf{L} \right\rangle, \end{split}$$

with kernels $k: \mathcal{X}_1^2 \to \mathbb{R}$, $\ell: \mathcal{X}_2^2 \to \mathbb{R}$, $\mathbf{K}:=\mathbf{K}_{\mathsf{x}}, \mathbf{L}:=\mathbf{L}_{\mathsf{y}}$, $\langle \mathbf{A} \rangle := \sum_{i,j} A_{i,j}$.

• Time complexity: quadratic.

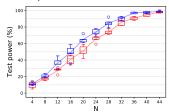
Numerical illustration: improved power

- Seoul bicycle rental data:
 - two-sample test (MMD, $d^{(2)}$): winter vs fall, d = 11,

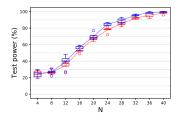


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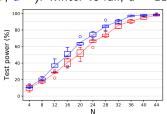


- Brazilian traffic data:
 - independence test (HSIC, CSIC); (blockage, fire, ...) vs slowness of traffic; $d_1 = 16$, $d_2 = 1$; l.h.s.

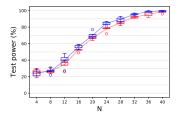


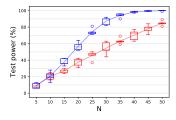
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 - two-sample test (MMD, $d^{(3)}$): slow vs fast moving traffic, d=16; r.h.s.





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- leveraging a combinatorial route (and tensor algebras).

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- MMD $\stackrel{\mathsf{m}=\mathsf{d}=1}{\longleftarrow} k$ -cumulants $\stackrel{\mathsf{i}=\mathbf{1}_2}{\longrightarrow} \mathsf{HSIC}\ (d=2).$
- k-Lancaster interaction $\stackrel{d=3}{\longleftarrow}$ k-Streitberg interaction $\stackrel{i=1_d}{\longleftarrow}$ k-cumulants.

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- Details @ NeurIPS [Bonnier et al., 2023], code.

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Appendix

- Bell numbers
- Characteristic kernels
- Universal kernels
- Moments and cumulants on \mathbb{R}^d
- Estimator for $d^{(3)}(\gamma, \eta)$.

Bell numbers

- B(m) := number of elements in P(m).
- $B_0 = B_1 = 1$, $B_2 = 2$, $B_3 = 5$, $B_4 = 15$, $B_5 = 52$, $B_6 = 203$, $B_7 = 877$, $B_8 = 4140$, ...

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- Recursion:

$$|B_{m+1}| = |P(m+1)| = \sum_{k=0}^{m} {m \choose k} B_k.$$

Bell numbers - continued

• Easy computation by the Bell triangle

```
1 2 2 3 5 5 7 10 15 15 20 27 37 52 52 ...
```

Bell numbers - continued

Easy computation by the Bell triangle

Asymptotics:

$$\frac{\ln B_n}{n} = \ln n - \ln \ln n - 1 + \frac{\ln \ln n}{\ln n} + \frac{1}{\ln n} + \frac{1}{2} \left(\frac{\ln \ln n}{\ln n} \right)^2 + \mathcal{O}\left(\frac{\ln \ln n}{\ln^2 n} \right)$$

as $n \to \infty$.



Description of characteristic kernels on \mathbb{R}^d

For continuous bounded shift-invariant kernels on \mathbb{R}^d :

$$k(\mathbf{x}, \mathbf{x}') = k_0(\mathbf{x} - \mathbf{x}') \stackrel{(*)}{=} \int_{\mathbb{R}^d} e^{-i\langle \mathbf{x} - \mathbf{x}', \boldsymbol{\omega} \rangle} d\boldsymbol{\Lambda}(\boldsymbol{\omega})$$

(*): Bochner's theorem.

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Description of characteristic kernels on \mathbb{R}^d

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Theorem ([Sriperumbudur et al., 2010])

k is characteristic iff. $supp(\Lambda) = \mathbb{R}^d$.

Examples on \mathbb{R} ; similarly \mathbb{R}^d [Sriperumbudur et al., 2010]

For Poisson kernel: $\sigma \in (0,1)$.

kernel name	e^{-k_0}	$\widehat{k}_0(\omega)$	$suppig(\widehat{k_0}ig)$
Gaussian	$e^{-\frac{x^2}{2\sigma^2}}$	$\sigma e^{-\frac{\sigma^2 \omega^2}{2}}$	\mathbb{R}
Laplacian		$\sqrt{rac{2}{\pi}} rac{\sigma}{\sigma^2 + \omega^2}$	\mathbb{R}
B_{2n+1} -spline	$e^{x^{2n+2}}\chi_{\left[-\frac{1}{2},\frac{1}{2}\right]}(x)$	$\frac{4^{n+1}}{\sqrt{2\pi}} \frac{\sin^{2n+2}\left(\frac{\omega}{2}\right)}{\omega^{2n+2}}$	\mathbb{R}
Sinc	$\frac{\sin(\sigma x)}{x}$	$\sqrt{\frac{\pi}{2}}\chi_{[-\sigma,\sigma]}(\omega)$	$[-\sigma,\sigma]$
Poisson	$\frac{1-\sigma}{\sigma^2-2\sigma\cos(x)+1}$	$\sqrt{2\pi} \sum_{j=-\infty}^{\infty} \sigma^{ j } \delta(\omega - j)$	\mathbb{Z}
Dirichlet	$\frac{\sin\left(\frac{(2n+1)x}{2}\right)}{\sin\left(\frac{x}{2}\right)}$	$\sqrt{2\pi} \sum_{j=-\infty}^{\infty} \delta(\omega - j)$	$\{0,\pm 1,\pm 2,\ldots,\pm n\}$
Fejér	$\frac{1}{n+1} \frac{\sin^2(\frac{(n+1)x}{2})}{\sin^2(\frac{x}{2})}$	$\sqrt{2\pi} \sum_{j=-n}^{n} \left(1 - \frac{ j }{n+1}\right) \delta(\omega - j)$	$\{0,\pm 1,\pm 2,\ldots,\pm n\}$
Cosine	$\cos(\sigma x)^{(2)}$	$\sqrt{\frac{\pi}{2}} \left[\delta(\omega - \sigma) + \delta(\omega + \sigma) \right]$	$\{-\sigma,\sigma\}$

Examples on \mathbb{R} ; similarly \mathbb{R}^d [Sriperumbudur et al., 2010]

For Poisson kernel: $\sigma \in (0,1)$.

Contents

Properties of universal kernels [Steinwart, 2001, Steinwart and Christmann, 2008]

If k is universal, then

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$$k(x,x) > 0$$
 for all $x \in \mathcal{X}$.

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- $\Phi(x) = k(\cdot, x)$ is injective, i.e.

$$\rho_k(x,y) = \|\Phi(x) - \Phi(y)\|_{\mathcal{H}_k}$$

is a metric.

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• The normalized kernel (like corr)

$$\tilde{k}(x,y) := \frac{k(x,y)}{\sqrt{k(x,x)k(y,y)}}$$

is universal.

Universal Taylor kernels [Steinwart, 2001, Steinwart and Christmann, 2008]

• For an $C^{\infty} \ni f: (-r,r) \to \mathbb{R}$

$$f(t) = \sum_{n=0}^{\infty} a_n t^n \quad t \in (-r, r), \ r \in (0, \infty].$$

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$$f(t) = \sum_{n=0}^{\infty} a_n t^n \quad t \in (-r, r), \ r \in (0, \infty].$$

• If $a_n > 0 \ \forall n$, then

$$k(\mathbf{x}, \mathbf{y}) = f(\langle \mathbf{x}, \mathbf{y} \rangle)$$

is universal on
$$\mathcal{X} := \left\{\mathbf{x} \in \mathbb{R}^d : \left\|\mathbf{x}\right\|_2 \le \sqrt{r}\right\}$$
.

Universal kernels on compact subsets of \mathbb{R}^d , $\alpha > 0$

• $k(\mathbf{x}, \mathbf{y}) = e^{\alpha(\mathbf{x}, \mathbf{y})}$: previous result with $f(t) = e^{\alpha t} \Rightarrow a_n = \frac{\alpha^n}{n!}$.

Universal kernels on compact subsets of \mathbb{R}^d , $\alpha > 0$

- $k(\mathbf{x}, \mathbf{y}) = e^{\alpha(\mathbf{x}, \mathbf{y})}$: previous result with $f(t) = e^{\alpha t} \Rightarrow a_n = \frac{\alpha^n}{n!}$.
- $k(\mathbf{x}, \mathbf{y}) = e^{-\alpha \|\mathbf{x} \mathbf{y}\|_2^2}$: exp. kernel & normalization.

Universal kernels on compact subsets of \mathbb{R}^d , $\alpha > 0$

- $k(\mathbf{x}, \mathbf{y}) = (1 \langle \mathbf{x}, \mathbf{y} \rangle)^{-\alpha}$ binomial kernel
 - $\quad \text{on } \mathcal{X} \text{ compact } \subset \{\mathbf{x} \in \mathbb{R}^d: \|\mathbf{x}\|_2 < 1\}.$

•
$$f(t) = (1-t)^{-\alpha} = \sum_{n=0}^{\infty} \frac{\binom{-\alpha}{n} (-1)^n}{\binom{n}{n}} t^n \quad (|t| < 1),$$

where
$$\binom{b}{n} = \sum_{i=1}^{n} \frac{b-i+1}{i}$$
.

Contents

Moments and cumulants on $\mathbb{R}^d \ni X \sim \gamma$, $\mathbf{i} \in \mathbb{N}^d$

$$\begin{array}{ll} d=1 & d\geq 1 \\\\ \text{moment sequence} & \mu(\gamma):=\left(\mu^{(i)}(\gamma)\right)_{i\in\mathbb{N}} & \mu(\gamma):=\left(\mu^{\mathbf{i}}(\gamma)\right)_{\mathbf{i}\in\mathbb{N}^d} \\\\ \text{moments} & \mu^{(i)}(\gamma):=\mathbb{E}\left(X^i\right)\in\mathbb{R} & \mu^{\mathbf{i}}(\gamma):=\mathbb{E}\left[X_1^{i_1}\cdots X_d^{i_d}\right]\in\mathbb{R} \end{array}$$

Moments and cumulants on $\mathbb{R}^d \ni X \sim \gamma$, $\mathbf{i} \in \mathbb{N}^d$

	d = 1	$d \ge 1$
moment sequence	$\mu(\gamma) := \left(\mu^{(i)}(\gamma)\right)_{i \in \mathbb{N}}$	$\mu(\gamma) := \left(\mu^{f i}(\gamma) ight)_{{f i} \in \mathbb{N}^d}$
moments	$\mu^{(i)}(\gamma) := \mathbb{E}(X^i) \in \mathbb{R}$	$\mu^{i}(\gamma) := \mathbb{E}\left[X_1^{i_1}\cdots X_d^{i_d} ight] \in \mathbb{R}$
<i>m</i> -th moment	$\mu^{(m)}(\gamma)$	$\mu^{m}(\gamma) := \left(\mu^{\mathbf{i}}(\gamma)\right)_{\mathrm{deg}(\mathbf{i})=m}$

where
$$\deg(\mathbf{i}) := i_1 + \cdots + i_d$$
, $\mu^0(\gamma) = 1$

Moments and cumulants on $\mathbb{R}^d \ni X \sim \gamma$, $\mathbf{i} \in \mathbb{N}^d$

$$\begin{array}{ll} d=1 & d\geq 1 \\ \\ \text{moment sequence} & \mu(\gamma):=\left(\mu^{(i)}(\gamma)\right)_{i\in\mathbb{N}} & \mu(\gamma):=\left(\mu^{\mathbf{i}}(\gamma)\right)_{\mathbf{i}\in\mathbb{N}^d} \\ \\ \text{moments} & \mu^{(i)}(\gamma):=\mathbb{E}\left(X^i\right)\in\mathbb{R} & \mu^{\mathbf{i}}(\gamma):=\mathbb{E}\left[X_1^{i_1}\cdots X_d^{i_d}\right]\in\mathbb{R} \\ \\ \textit{m-th moment} & \mu^{(m)}(\gamma) & \mu^m(\gamma):=\left(\mu^{\mathbf{i}}(\gamma)\right)_{\deg(\mathbf{i})=m} \end{array}$$

and cumulants $\kappa(\gamma) = (\kappa^{\mathbf{i}}(\gamma))_{\mathbf{i} \in \mathbb{N}^d}$

$$\sum_{\mathbf{i} \in \mathbb{N}^d} \kappa^{\mathbf{i}}(\gamma) \frac{\theta^{\mathbf{i}}}{\mathbf{i}!} = \log \left(\sum_{\mathbf{i} \in \mathbb{N}^d} \mu^{\mathbf{i}}(\gamma) \frac{\theta^{\mathbf{i}}}{\mathbf{i}!} \right), \quad \boldsymbol{\theta} \in \mathbb{R}^d,$$

where $\deg(\mathbf{i}) := i_1 + \dots + i_d$, $\mu^0(\gamma) = 1$, $\mathbf{i}! = i_1! \dots i_d!$, $\theta^{\mathbf{i}} = \theta_1^{i_1} \dots \theta_d^{i_d}$.

Contents \mathbb{R} moments and cumulants on \mathbb{R} motivation of cumulants

Estimator for $d^{(3)}(\gamma,\eta) = \|\kappa_k^{(3)}(\gamma) - \kappa_k^{(3)}(\eta)\|_{\mathcal{H}_k^{\otimes 3}}^2$, N = M

$$d^{(3)}(\gamma,\eta) = \|\kappa_k^{(3)}(\gamma)\|_{\mathcal{H}_k^{\otimes 3}}^2 + \|\kappa_k^{(3)}(\eta)\|_{\mathcal{H}_k^{\otimes 3}}^2 - 2\langle\kappa_k^{(3)}(\gamma),\kappa_k^{(3)}(\eta)\rangle_{\mathcal{H}_k^{\otimes 3}}$$

Estimator for $d^{(3)}(\gamma,\eta) = \|\kappa_k^{(3)}(\gamma) - \kappa_k^{(3)}(\eta)\|_{\mathcal{H}^{\otimes 3}}^2$, N = M

$$d^{(3)}(\gamma,\eta) = \|\kappa_{k}^{(3)}(\gamma)\|_{\mathcal{H}_{k}^{\otimes 3}}^{2} + \|\kappa_{k}^{(3)}(\eta)\|_{\mathcal{H}_{k}^{\otimes 3}}^{2} - 2\langle\kappa_{k}^{(3)}(\gamma),\kappa_{k}^{(3)}(\eta)\rangle_{\mathcal{H}_{k}^{\otimes 3}}$$

$$\langle\kappa_{k}^{(3)}(\gamma),\kappa_{k}^{(3)}(\eta)\rangle_{\mathcal{H}_{k}^{\otimes 3}} \approx \frac{1}{N^{2}} \left\langle \mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy} - 3\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \right.$$

$$- 3\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy}$$

$$+ 3\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \left\langle \frac{\mathbf{K}_{xy}}{N^{2}} \right\rangle + 2\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy} \circ \mathbf{H}\mathbf{K}_{xy}$$

$$+ 2\mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} + 6\mathbf{K}_{xy} \circ \mathbf{K}_{xy} \circ$$

Note: Matrix multiplication takes precedence over the Hadamard one.

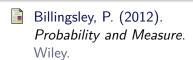
Estimator for $d^{(3)}(\gamma, \eta)$ – continued

$$\begin{split} \|\kappa_{k}^{(3)}(\gamma)\|_{\mathcal{H}_{k}^{\otimes 3}}^{2} &\approx \frac{1}{N^{2}} \left\langle \mathbf{K}_{x} \circ \mathbf{K}_{x} \circ \mathbf{K}_{x} - 6\mathbf{K}_{x} \circ \mathbf{K}_{x} \mathbf{H} \circ \mathbf{K}_{x} \right. \\ &+ 4\mathbf{K}_{x} \mathbf{H} \circ \mathbf{K}_{x} \circ \mathbf{K}_{x} \mathbf{H} + 3\mathbf{K}_{x} \circ \mathbf{K}_{x} \left\langle \frac{\mathbf{K}_{x}}{N^{2}} \right\rangle \\ &+ 6\mathbf{K}_{x} \mathbf{H} \circ \mathbf{H} \mathbf{K}_{x} \circ \mathbf{K}_{x} - 12\mathbf{K}_{x} \circ \mathbf{H} \mathbf{K}_{x} \left\langle \frac{\mathbf{K}_{x}}{N^{2}} \right\rangle \\ &+ 4 \left\langle \frac{\mathbf{K}_{x}}{N^{2}} \right\rangle^{2} \mathbf{K}_{x} \right\rangle. \end{split}$$

 $\|\kappa_k^{(3)}(\eta)\|_{\mathcal{H}_{\nu}^{\otimes 3}}^2$: similarly (change \mathbf{K}_x to \mathbf{K}_y).



Contents $d^2(\gamma, \eta)$ estimation



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