

Homework 2, STAT 5205

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Question 1

Consider the Bayesian model: G is a categorical random variable which takes values in $\{1, 2\}$ and satisfies

$$\mathbb{P}(G = 1) = \pi_1, \quad \mathbb{P}(G = 2) = \pi_2$$

with $\pi_1 + \pi_2 = 1$. \mathbf{X} is a d -dimensional random vector ($\mathbf{X} \in \mathbb{R}^d$) with conditional Gaussian distributions

$$\mathbf{X} | G = 1 \sim N(\boldsymbol{\mu}_1, \boldsymbol{\Sigma}) \quad \mathbf{X} | G = 2 \sim N(\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

Compute the following items (write out the formulas in terms of $\pi_1, \pi_2, \boldsymbol{\mu}_1, \boldsymbol{\mu}_2, \boldsymbol{\Sigma}$).

Part i

The joint distributions $P\{G = 1, \mathbf{X} = \mathbf{x}\} = ?$, $P(G = 2, \mathbf{X} = \mathbf{x}) = ?$

By the law of conditional probability:

$$P(G = g, X = x) = P(G = g)P(X = x | G = g)$$

For $G = 1$:

$$P(G = 1, X = x) = \pi_1 \cdot \frac{1}{(2\pi)^{d/2}(\det(\boldsymbol{\Sigma}))^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_1)^T \boldsymbol{\Sigma}^{-1}(x - \mu_1)\right)$$

For $G = 2$:

$$P(G = 2, X = x) = \pi_2 \cdot \frac{1}{(2\pi)^{d/2}(\det(\boldsymbol{\Sigma}))^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_2)^T \boldsymbol{\Sigma}^{-1}(x - \mu_2)\right)$$

Part ii

The marginal distributional density of \mathbf{X} , $d(\mathbf{x}) = ?$

The marginal density of X , denoted as $d(x)$, can be found by summing the joint densities over all possible values of G :

$$d(x) = P(X = x) = P(G = 1, X = x) + P(G = 2, X = x)$$

Using the joint distributions derived earlier:

$$d(x) = \pi_1 f_1(x) + \pi_2 f_2(x)$$

Thus the marginal density of X is:

$$d(x) = \pi_1 \frac{1}{(2\pi)^{d/2}(\det(\boldsymbol{\Sigma}))^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_1)^T \boldsymbol{\Sigma}^{-1}(x - \mu_1)\right) + \pi_2 \frac{1}{(2\pi)^{d/2}(\det(\boldsymbol{\Sigma}))^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_2)^T \boldsymbol{\Sigma}^{-1}(x - \mu_2)\right)$$

Part iii

The conditional distribution $P\{G = 1 | \mathbf{X} = \mathbf{x}\} = ?$

We can solve this problem with the help of Baye's Rule

$$P(G = 1 | X = x) = \frac{P(G = 1, X = x)}{P(X = x)}$$

From our previous derivations:

The joint distribution:

$$P(G = 1, X = x) = \pi_1 f_1(x)$$

where

$$f_1(x) = \frac{1}{(2\pi)^{d/2}(\det(\Sigma))^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_1)^T \Sigma^{-1}(x - \mu_1)\right)$$

And the marginal density of X :

$$P(X = x) = d(x) = \pi_1 f_1(x) + \pi_2 f_2(x)$$

where

$$f_2(x) = \frac{1}{(2\pi)^{d/2}(\det(\Sigma))^{1/2}} \exp\left(-\frac{1}{2}(x - \mu_2)^T \Sigma^{-1}(x - \mu_2)\right)$$

So the conditional distribution is

$$P(G = 1 \mid X = x) = \frac{\pi_1 f_1(x)}{\pi_1 f_1(x) + \pi_2 f_2(x)}$$

Question 2

$f(\mathbf{x})$ is a function of four variables/arguments, and has the formula below

$$f(x_1, x_2, x_3, x_4) = x_4 e^{-(x_1-1)^2/2} + \log(x_2^2 + 1) + \cos(x_3) \sin(x_4^2) + \frac{\sqrt{x_2}}{x_3^2 + 2}$$

Compute the first and second order gradients of $f(\mathbf{x})$, i.e.,

$$\begin{aligned} \frac{\partial f(\mathbf{x})}{\partial \mathbf{x}} &=? \\ \frac{\partial^2 f(\mathbf{x})}{\partial \mathbf{x} \partial \mathbf{x}^\top} &=? \end{aligned}$$

Part i For the first order problem, we have to compute

$$\nabla f = \left(\frac{\partial f}{\partial x_1}, \frac{\partial f}{\partial x_2}, \frac{\partial f}{\partial x_3}, \frac{\partial f}{\partial x_4} \right).$$

Partial derivative with respect to x_1 :

$$\frac{\partial f}{\partial x_1} = x_4 \cdot e^{-\frac{(x_1-1)^2}{2}} \cdot \left(-\frac{x_1-1}{1} \right) = -x_4(x_1-1)e^{-\frac{(x_1-1)^2}{2}}$$

Partial derivative w.r.t. x_2 :

$$\frac{\partial f}{\partial x_2} = \frac{2x_2}{x_2^2 + 1} + \frac{1}{2(x_3^2 + 2)\sqrt{x_2}}$$

Partial derivative w.r.t. x_3 :

$$\frac{\partial f}{\partial x_3} = -\sin(x_3) \sin(x_4^2) - \frac{2x_3\sqrt{x_2}}{(x_3^2 + 2)^2}$$

Partial derivative w.r.t. x_4 :

$$\frac{\partial f}{\partial x_4} = e^{-\frac{(x_1-1)^2}{2}} + \cos(x_3) \cos(x_4^2) \cdot 2x_4$$

Thus the gradient vector is:

$$\nabla f = \begin{bmatrix} -x_4(x_1-1)e^{-\frac{(x_1-1)^2}{2}} \\ \frac{2x_2}{x_2^2+1} + \frac{1}{2(x_3^2+2)\sqrt{x_2}} \\ -\sin(x_3) \sin(x_4^2) - \frac{2x_3\sqrt{x_2}}{(x_3^2+2)^2} \\ e^{-\frac{(x_1-1)^2}{2}} + \cos(x_3) \cos(x_4^2) \cdot 2x_4 \end{bmatrix}$$

Part ii The Hessian matrix H_f is a 4×4 symmetric matrix whose entries are the second-order partial derivatives of $f(x)$:

$$H_f = \begin{bmatrix} \frac{\partial^2 f}{\partial x_1^2} & \frac{\partial^2 f}{\partial x_1 \partial x_2} & \frac{\partial^2 f}{\partial x_1 \partial x_3} & \frac{\partial^2 f}{\partial x_1 \partial x_4} \\ \frac{\partial^2 f}{\partial x_2 \partial x_1} & \frac{\partial^2 f}{\partial x_2^2} & \frac{\partial^2 f}{\partial x_2 \partial x_3} & \frac{\partial^2 f}{\partial x_2 \partial x_4} \\ \frac{\partial^2 f}{\partial x_3 \partial x_1} & \frac{\partial^2 f}{\partial x_3 \partial x_2} & \frac{\partial^2 f}{\partial x_3^2} & \frac{\partial^2 f}{\partial x_3 \partial x_4} \\ \frac{\partial^2 f}{\partial x_4 \partial x_1} & \frac{\partial^2 f}{\partial x_4 \partial x_2} & \frac{\partial^2 f}{\partial x_4 \partial x_3} & \frac{\partial^2 f}{\partial x_4^2} \end{bmatrix}$$

1. $\frac{\partial^2 f}{\partial x_1^2} = x_4 e^{-\frac{(x_1-1)^2}{2}} \left(\frac{(x_1-1)^2}{1} - 1 \right)$
2. $\frac{\partial^2 f}{\partial x_2^2} = \frac{2(1-x_2^2)}{(x_2^2+1)^2} - \frac{1}{4(x_3^2+2)x_2^{3/2}}$
3. $\frac{\partial^2 f}{\partial x_3^2} = -\cos(x_3) \sin(x_4^2) + \frac{2\sqrt{x_2}(3x_3^2-2)}{(x_3^2+2)^3}$
4. $\frac{\partial^2 f}{\partial x_4^2} = -2x_4 e^{-\frac{(x_1-1)^2}{2}} + \cos(x_3)(-2\sin(x_4^2) \cdot 2x_4^2 + 2\cos(x_4^2))$
5. $\frac{\partial^2 f}{\partial x_1 \partial x_2} = \frac{\partial^2 f}{\partial x_2 \partial x_1} = 0$
6. $\frac{\partial^2 f}{\partial x_1 \partial x_3} = \frac{\partial^2 f}{\partial x_3 \partial x_1} = 0$:
7. $\frac{\partial^2 f}{\partial x_1 \partial x_4} = \frac{\partial^2 f}{\partial x_4 \partial x_1} = -(x_1-1)e^{-\frac{(x_1-1)^2}{2}}$
8. $\frac{\partial^2 f}{\partial x_2 \partial x_3} = \frac{\partial^2 f}{\partial x_3 \partial x_2} = -\frac{x_3}{(x_3^2+2)^2\sqrt{x_2}}$
9. $\frac{\partial^2 f}{\partial x_2 \partial x_4} = \frac{\partial^2 f}{\partial x_4 \partial x_2} = 0$
10. $\frac{\partial^2 f}{\partial x_3 \partial x_4} = \frac{\partial^2 f}{\partial x_4 \partial x_3} = -\cos(x_3) \cos(x_4^2) \cdot 2x_4$

So final Hessian Matrix is

$$H_f = \begin{pmatrix} x_4 e^{-\frac{(x_1-1)^2}{2}} & \frac{2(1-x_2^2)}{(x_2^2+1)^2} - \frac{1}{4(x_3^2+2)x_2^{3/2}} & -\frac{x_3}{(x_3^2+2)^2 \sqrt{x_2}} & -(x_1-1) e^{-\frac{(x_1-1)^2}{2}} \\ -\frac{x_3}{(x_3^2+2)^2 \sqrt{x_2}} & -\cos(x_3) \sin(x_4^2) + \frac{2\sqrt{x_2}(3x_3^2+2)}{(x_3^2+2)^2} & -\cos(x_3) \cos(x_4^2) 2x_4 & -2x_4 e^{-\frac{(x_1-1)^2}{2}} + \cos(x_3) \cdot (-2\sin(x_4^2) \cdot 2x_4^2 + 2\cos(x_4^2)) \\ -(x_1-1) e^{-\frac{(x_1-1)^2}{2}} & -\cos(x_3) \cos(x_4^2) 2x_4 & -2x_4 e^{-\frac{(x_1-1)^2}{2}} + \cos(x_3) \cdot (-2\sin(x_4^2) \cdot 2x_4^2 + 2\cos(x_4^2)) & \end{pmatrix}$$

Question 3

Let $A \in \mathbb{R}^{p \times q}$ ($p > q$) be a matrix and has singular value decomposition

$$A = UDV^\top$$

where $U \in \mathbb{R}^{p \times q}$, $V \in \mathbb{R}^{q \times q}$ are orthonormal matrices, and D is a diagonal matrix with nonnegative diagonal elements, i.e.,

$$D = \begin{bmatrix} \lambda_1 & & \\ & \ddots & \\ & & \lambda_q \end{bmatrix}$$

The diagonal elements are in a descending order $\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_q \geq 0$. Prove that

$$\sup_{x \in \mathbb{R}^p, \|x\|=1, y \in \mathbb{R}^q, \|y\|=1} |x^\top A y| = \lambda_1.$$

Answer

We can use the singular decomposition on A :

$$A = UDV^\top,$$

for any vectors $x \in \mathbb{R}^p$ and $y \in \mathbb{R}^q$, we can rewrite as $x^\top A y = x^\top U D V^\top y$. Define new vectors $u = U^\top x \in \mathbb{R}^q$ and $v = V^\top y \in \mathbb{R}^q$. Since U and V are orthonormal matrices, we have: $\|u\| \leq \|x\| = 1$ (because U is not necessarily a square), $\|v\| = \|y\| = 1$. Since we are proving the sup, we will choose the value of $\|u\|$ to be its maximum, which is 1.

Thus, we can rewrite $x^\top A y = u^\top D v = \sum_{i=1}^q \lambda_i u_i v_i$. Since D is diagonal. By the Cauchy-Schwarz inequality, $|\sum_{i=1}^q \lambda_i u_i v_i| \leq \sum_{i=1}^q \lambda_i |u_i v_i|$.

Using the property that $\sum_{i=1}^q u_i^2 \leq 1$ and $\sum_{i=1}^q v_i^2 \leq 1$, the maximum value occurs when all the weight is concentrated on the largest singular value λ_1 , this means that $u_1 = v_1 = 1$ and $u_i = v_i = 0$ for all $i \geq 2$. And thus we know that $\sup_{\|u\|=1, \|v\|=1} \sum_{i=1}^q \lambda_i u_i v_i = \lambda_1$.

Then we choose $x = u_1$, $y = v_1$, where u_1 and v_1 are the first columns of U and V , respectively. Then we have $x^\top A y = u_1^\top U D V^\top v_1 = \lambda_1$.

Thus, the supremum is attained at λ_1 , proving:

$$\sup_{x \in \mathbb{R}^p, \|x\|=1, y \in \mathbb{R}^q, \|y\|=1} |x^\top A y| = \lambda_1.$$

□