## **Chapter 7 Observational Study**

(Optional topic)

### 7.1 Causality in observational studies

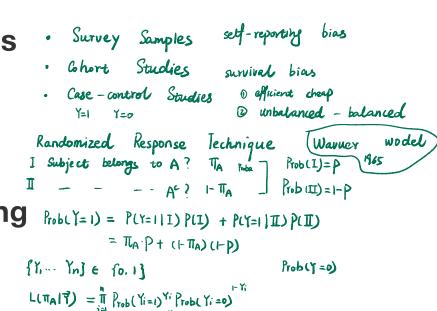
- Features of observational studies
- Selection bias

# 7.2 Analysis with no latent confouding | Prob(Y=1) = P(Y=1|I) P(I) + P(Y=1|II) P(II)

- Assumptions
  - 。 I.I.D.
  - Ignorability
  - Overlap
- Estimation
  - Stratification
  - Outcome regression
- Propensity score
  - Definition and key properties
  - Propensity scores: matching
  - Propensity scores: weighting
  - Doubly-robust regression
- Covariance balancing

#### 7.3 Instrumental variable

- Definition and assumptions
- Key properties of IV



unbiased estimator

Logistic Regression [case-controlled studies] not bec. a binary treatment  $\mathbb{R}$  should we use a logistic Regression to analyze data from an RCT? Data(Xi, Yi), Yi  $\in$  [0,1]  $\times$  i  $\in$  IRP

logistic regression

$$logit(\Pi_i) = X_i^T \beta \quad where \quad \Pi_i = |P(Y_i = 1 \mid X_i)| = E[Y_i \mid X_i] \in [0, 1]$$
and 
$$logit(a) = log \left(\frac{a}{1-a}\right) : log \quad odd$$

$$\beta_i = log \left(\frac{\Pi(X_i = b_i + 1, X_2, \dots, X_p)}{\frac{\Pi(X_i = b_i + 1, X_i, \dots, X_p)}{\frac{\Pi(X_i = b_i + 1, X_i$$

MA, MLE = n-(+p)

= 
$$n_i \log \pi_i + (n-n_i) \log u - \pi_i$$
 where  $\log \pi = \beta_0$   $\pi_i = \exp(\beta_0)$ 

Estimation

## 7.4 Missing data

- Missing mechanisms
- Multiple imputation

$$= n_{1} \left[ \beta_{0} - \log \left[ \exp(\beta_{0}) + 1 \right] \right] + (n - n_{1}) \left[ 0 - \log \left[ \exp(\beta_{0}) + 1 \right] \right]$$

$$\frac{\partial \log L}{\partial \beta_{0}} = 0 \qquad n_{1} - n \frac{\exp(\hat{\beta}_{0})}{\exp(\hat{\beta}_{0}) + 1} = 0 \implies \hat{\beta}_{0} : \log it \left( \frac{n_{1}}{n_{1}} \right)$$

$$U(\beta) = \frac{\partial \log L}{\partial \beta_{0}} \qquad \text{score function } \text{ if } \beta \in \mathbb{R}^{p}$$

$$At \quad \text{MLE}(\hat{\beta}_{0}), \quad U(\hat{\beta}) = 0$$

$$I(\beta) = \mathbb{E} \left[ -\frac{\partial^{2} \log L}{\partial \beta^{2}} \right] \quad \text{Fisher information } \in \mathbb{R}^{p \times p} \quad \text{var}(\hat{\beta}) = \text{I}^{1}(\hat{\beta})$$

$$In(\hat{\beta} - \beta) \longrightarrow N(0, \quad \text{I}^{-1}(\beta))$$

$$Confidence \quad \text{Interval for } \hat{\beta}_{1}$$

$$\hat{Se}^{2}(\hat{\beta}_{1}) = \text{Ist diagnos (entry in } \text{I}^{-1}(\hat{\beta}))$$

$$Ho: \beta_{1} = 0 \quad \text{vs. Ha} \quad \beta_{1} \neq 0$$

$$t = \frac{\hat{\beta}_{1}}{\hat{Se}(\beta_{1})} \quad \text{$\sim N(0, 1)$} \quad \text{wald test}$$

$$t = \frac{\widehat{\beta}_{1}}{\widehat{\operatorname{Se}(\beta_{1})}} \rightsquigarrow N(0,1) \quad | \text{ wald test}$$

$$S = \frac{U(\beta_{1}=0)^{2}}{\left[I^{-1}(\beta_{1}=0)\right]_{1}} \rightsquigarrow X^{2} \quad \text{with } df = 1 \quad | \text{ score test}$$

$$LR = -2[log L(residual) - log L(full)] \times \chi^2 \circ df$$

$$Likelihood ratio test$$

$$F + test is a special case F(a, a_2) \xrightarrow{a_2 \to a_2} \chi^2_{a_1}$$

Model Diagnostics

Pearson test

Deviance test