

Recap: Simple Linear Regression in Matrix Form

$$\mathbf{Y} = \mathbf{X} \underbrace{\boldsymbol{\beta}}_{n \times 1} + \underbrace{\boldsymbol{\epsilon}}_{1 \times 2} \underbrace{\boldsymbol{\beta}}_{2 \times 1} + \underbrace{\boldsymbol{\epsilon}}_{1 \times 1}.$$

•
$$\mathbf{E}\{\epsilon\} = \mathbf{0}_n, \ \sigma^{\mathbf{2}}\{\epsilon\} = \sigma^2 \mathbf{I}_n.$$

- Normal error model: $\epsilon \sim \text{Normal}_n(\mathbf{0}_n, \sigma^2 \mathbf{I}_n)$.
- LS estimators:

$$\hat{\boldsymbol{\beta}} = \begin{bmatrix} \hat{\beta}_0 \\ \hat{\beta}_1 \end{bmatrix} = (\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y}.$$

Fitted values and residuals:

$$\widehat{\mathbf{Y}} = \mathbf{X}(\mathbf{X}'\mathbf{X})^{-1}\mathbf{X}'\mathbf{Y} = \mathbf{H}\mathbf{Y}, \quad \mathbf{e} = (\mathbf{I}_n - \mathbf{H})\mathbf{Y}.$$

→ Hat matrix: H := X(X'X)⁻¹X', is a projection matrix.



(1)

Recap: Column Space of the Design Matrix X

$$\mathbf{X} = (\mathbf{1}_n, \mathbf{x}).$$

• $\langle X \rangle = \{c_0 \mathbf{1}_n + c_1 \mathbf{x} = \mathbf{X} \mathbf{c} : c_0, c_1 \in R, \mathbf{c} = (c_0, c_1)^T\}$, is the linear subspace of \mathbf{R}^n generated by the columns of \mathbf{X} .

Geometric Interpretation of Linear Regression

The hat matrix **H** projects a vector in \mathbb{R}^n to the column space $\langle X \rangle$ of the design matrix **X**: for any $\mathbf{w} \in \mathbb{R}^n$

- Hw $\in \langle X \rangle$, i.e., there exists $c_0, c_1 \in \mathbb{R}$ such that Hw = $c_0 \mathbf{1}_n + c_1 \mathbf{x}$.
- $\mathbf{w} \mathbf{H}\mathbf{w} \perp \langle X \rangle$, i.e., for any $\mathbf{v} \in \langle X \rangle$, the inner product $\langle \mathbf{w} \mathbf{H}\mathbf{w}, \mathbf{v} \rangle = (\mathbf{w} \mathbf{H}\mathbf{w})^T \mathbf{v} = 0$.

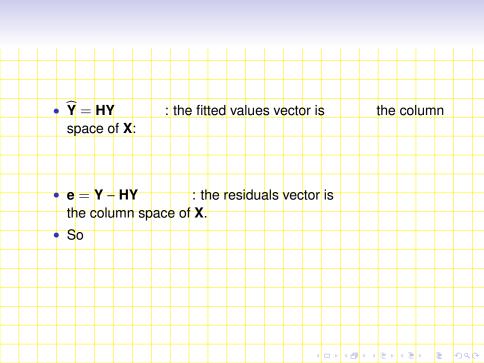
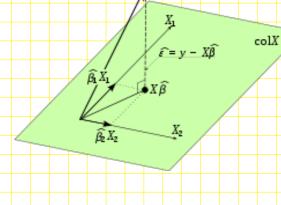


Figure Orthogonal projection of response vector \mathbf{Y} onto the linear subspace of \mathbb{R}^n generated by the columns of the design matrix \mathbf{X} .



LS Estimators: Expectations

LS estimators are unbiased estimators : Expectation of the fitted values: Expectation of the residuals:

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LS Estimators: Variance-covariance Matrices

Variance-covariance of the LS estimators:

What is the covariance between
$$\hat{\beta}_0$$
 and $\hat{\beta}_1$? What happens if $\overline{X}=0$?

Variance-covariance of fitted values:

Variance-covariance of residuals:

Are residuals uncorrelated? Do they have the same variance?

Sum of Squares in Matrix Form

Error sum of squares:
$$SSE = \sum_{i=1}^{n} e_i^2.$$
• Matrix form:
• Recall that $\mathbf{I}_n - \mathbf{H}$ is a matrix.
• $df(SSE) =$

Total sum of squares:

$$SSTO = \sum_{i=1}^{n} (Y_i - \overline{Y})^2 = \sum_{i=1}^{n} Y_i^2 - n(\overline{Y})^2.$$
• Matrix form:

• Note $\mathbf{I}_n - \frac{1}{n} \mathbf{J}_n$ is a projection matrix:

$$\mathbf{J}_n = \mathbf{1}_n \mathbf{1}_n' = \begin{bmatrix} 1 & 1 & \cdots & 1 \\ 1 & 1 & \cdots & 1 \\ \vdots & \vdots & \ddots & \vdots \\ 1 & 1 & \cdots & 1 \end{bmatrix}$$
• $df(SSTO) = \mathbf{J}_n \mathbf{1}_n' = \mathbf{J}_n \mathbf{1}_n' \mathbf{J}_n' \mathbf{$

Regression sum of squares :
$$SSR = \sum_{i=1}^{n} (\widehat{Y}_i - \overline{Y})^2$$
.

• Matrix form: $\overline{Y} =$

• Note $\mathbf{H} - \frac{1}{n} \mathbf{J}_n$ is a projection matrix:

• $df(SSR) =$

Recap: Sum of Squares in Matrix Form

$$SSTO = Y' \left(I_n - \frac{1}{n}J_n\right)Y.$$

$$SSE = Y'(I_n - H)Y.$$

$$SSR = Y'(H - \frac{1}{n}J_n)Y.$$

E(SSE)

Properties of Projection Matrices

Optional Reading material.

They have eigen-decomposition of the form:

$$Q\Lambda Q^T$$
,

where Q is an orthogonal matrix of eigenvectors and Λ is a diagonal matrix of eigenvalues.

- Their eigenvalues are either 1 or 0.
- The number of nonzero eigenvalues equals to trace of the matrix equals to the rank.
- For simple linear regression:

$$rank(\mathbf{H}) = 2$$
, $rank(\mathbf{I}_n - \mathbf{H}) = n - 2$.

Sampling Distribution of SSE

$$\mathbf{I}_n - \mathbf{H} = \mathbf{Q}^T \mathbf{\Lambda} \mathbf{Q},$$

where $\Lambda = \text{diag}\{1, \dots, 1, 0, 0\}$ and \mathbf{Q} is an orthogonal matrix.

•
$$(I_n - H)X = 0 \Longrightarrow$$

$$\mathbf{e} = (\mathbf{I}_n - \mathbf{H})\mathbf{Y} = (\mathbf{I}_n - \mathbf{H})(\mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}) = (\mathbf{I}_n - \mathbf{H})\boldsymbol{\epsilon}$$

Optional Reading material (cont'd).

•
$$SSE = \mathbf{e}^T \mathbf{e} = \boldsymbol{\epsilon}^T (\mathbf{I}_n - \mathbf{H}) \boldsymbol{\epsilon} = (\mathbf{Q} \boldsymbol{\epsilon})^T \mathbf{\Lambda} (\mathbf{Q} \boldsymbol{\epsilon}).$$

• Let
$$\mathbf{z} = \mathbf{Q}\boldsymbol{\epsilon}$$
, then

$$SSE = \sum_{i=1}^{n-2} z_i^2.$$

$$\mathsf{E}(\mathsf{z}) = \mathsf{Q}\mathsf{E}\{\epsilon\} = \mathsf{0}, \quad \sigma^2\{\mathsf{z}\} = \mathsf{Q}\sigma^2\{\epsilon\}\mathsf{Q}^\mathsf{T} = \sigma^2\mathsf{Q}\mathsf{Q}^\mathsf{T} = \sigma^2\mathsf{I}_n.$$

So under Normal error model, z_i s are i.i.d. $N(0, \sigma^2)$.

• So SSE
$$\sim \sigma^2 \chi^2_{(n-2)}$$
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