

# Gradient algorithms for quadratic optimization with fast convergence rates

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**Abstract** We propose a family of gradient algorithms for minimizing a quadratic function  $f(x) = (Ax, x)/2 - (x, y)$  in  $\mathbb{R}^d$  or a Hilbert space, with simple rules for choosing the step-size at each iteration. We show that when the step-sizes are generated by a dynamical system with ergodic distribution having the arcsine density on a subinterval of the spectrum of  $A$ , the asymptotic rate of convergence of the algorithm can approach the (tight) bound on the rate of convergence of a conjugate gradient algorithm stopped before  $d$  iterations, with  $d \leq \infty$  the space dimension.

**Keywords** Chebyshev polynomials · Conjugate gradient · Krylov space · Logistic map · Quadratic operator · Steepest descent

## 1 Introduction

Consider the problem of minimizing a quadratic function  $f(\cdot)$  defined either on  $\mathbb{R}^d$  or a Hilbert space by

$$f(x) = \frac{1}{2}(Ax, x) - (x, y), \quad (1)$$

where  $(\cdot, \cdot)$  denotes the inner product. We assume that  $A$  is either a symmetric positive-definite matrix or a self-adjoint operator, with

$$0 < m = \inf_{(x,x)=1} (Ax, x) < M = \sup_{(x,x)=1} (Ax, x) < \infty.$$

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If  $A$  is a matrix, then  $m$  and  $M$  are the smallest and largest eigenvalues of  $A$ , respectively.

Consider a general gradient algorithm with iterations of the form

$$x_{k+1} = x_k - \gamma_k g_k, \quad k = 0, 1, 2, \dots \quad (2)$$

where  $g_k = \nabla f(x_k)$  is the gradient of the objective function  $f(\cdot)$  at point  $x_k$ . For the objective function (1),  $\nabla f(x) = Ax - y$ . The iteration (2) can be rewritten in terms of the gradients as

$$g_{k+1} = g_k - \gamma_k A g_k. \quad (3)$$

In a series of papers [10–12] and the monograph [9] many gradient algorithms have been shown to be equivalent to special algorithms for updating measures on the interval  $[m, M]$ . The central idea is that of renormalization applied to the gradient. For simplicity the presentation is made for the finite dimensional case where  $A$  is a matrix, which can be assumed, without loss of generality, to be diagonal  $A = \text{diag}(\lambda_1, \dots, \lambda_d)$  with eigenvalues  $m = \lambda_1 \leq \lambda_2 \leq \dots \leq \lambda_d = M$ . Extension to the Hilbert-space case will be considered in Sect. 5.

Write  $z_k = g_k / \sqrt{(g_k, g_k)}$  for the normalized gradient at  $x_k$  and define

$$p_i^{(k)} = \{z_k\}_i^2 = \frac{\{g_k\}_i^2}{\sum_{j=1}^d \{g_k\}_j^2}, \quad i = 1, \dots, d,$$

as the  $i$ -th probability corresponding to vector  $z_k$ , where  $\{v\}_i$  denotes the  $i$ -th component of vector  $v$ . Let  $\nu_k$  denote the probability measure on the spectrum of  $A$  defined by the  $p_i^{(k)}$ 's, that is,  $\nu_k(\lambda_i) = p_i^{(k)}$ . The probability measure  $\nu_{k+1}$  is defined by

$$p_i^{(k+1)} = \frac{\{g_{k+1}\}_i^2}{(g_{k+1}, g_{k+1})} \quad \text{for } i = 1, \dots, d.$$

Note that (3) gives

$$(g_{k+1}, g_{k+1}) = (g_k, g_k) - 2\gamma_k (A g_k, g_k) + \gamma_k^2 (A^2 g_k, g_k), \quad (4)$$

so that

$$\begin{aligned} p_i^{(k+1)} &= \frac{(1 - \gamma_k \lambda_i)^2}{(g_k, g_k) - 2\gamma_k (A g_k, g_k) + \gamma_k^2 (A^2 g_k, g_k)} \{g_k\}_i^2 \\ &= \frac{(1 - \gamma_k \lambda_i)^2}{1 - 2\gamma_k \mu_1^{(k)} + \gamma_k^2 \mu_2^{(k)}} p_i^{(k)}, \end{aligned} \quad (5)$$

where

$$\mu_\alpha^{(k)} = \mu_\alpha(\nu_k) = \frac{(A^\alpha g_k, g_k)}{(g_k, g_k)} \quad (6)$$

is the  $\alpha$ -th moment of the measure  $\nu_k$ . When two eigenvalues of  $A$  are equal, say  $\lambda_j = \lambda_{j+1}$ , the updating rules for  $p_j^{(k)}$  and  $p_{j+1}^{(k)}$  are identical so that the analysis of

the behaviour of the algorithm remains the same when  $p_j^{(k)}$  and  $p_{j+1}^{(k)}$  are confounded. We may thus assume that all eigenvalues of  $A$  are distinct. Also, a zero weight remains equal to zero at all subsequent iterations, we thus assume that  $v_0(\lambda_i) > 0$  for all  $i$ .

A common definition for the rate of convergence of the algorithm at iteration  $k$  is  $r_k = (g_{k+1}, g_{k+1}) / (g_k, g_k)$ . The rate for  $n$  iterations is

$$\prod_{k=0}^{n-1} r_k = \frac{(g_n, g_n)}{(g_0, g_0)};$$

therefore, the asymptotic rate of the algorithm can naturally be defined as

$$R = \lim_{n \rightarrow \infty} R_n, \quad \text{with } R_n = \left( \prod_{k=0}^{n-1} r_k \right)^{1/n}. \quad (7)$$

Of course, this rate may depend on the initial point  $x_0$  or, equivalently, on  $g_0$ . Other rates which are asymptotically equivalent to  $\{r_k\}$  can be considered as well, see [12] and Remark 6.

The most familiar gradient algorithm is the steepest-descent algorithm, for which the step-size  $\gamma_k$  at iteration  $k$  is chosen so as to minimize  $f(x_k - \gamma g_k)$  with respect to  $\gamma$ , which gives  $\gamma_k = (g_k, g_k) / (Ag_k, g_k) = 1 / \mu_1^{(k)}$ . Its asymptotic behaviour is well-known, see [1, 10]. In particular, its convergence is slow: the asymptotic rate  $R$  depends on the starting point but is never far from its worst value given by the Kantorovich bound

$$R_{\max} = \left( \frac{\rho - 1}{\rho + 1} \right)^2,$$

where  $\rho = M/m$ , the condition number of  $A$ . The asymptotic behaviour of the family of algorithms defined by  $\gamma_k = \mu_\alpha^{(k)} / \mu_{\alpha+1}^{(k)}$  (which includes the method of minimum residues for  $\alpha = 1$ ) is shown in [12] to be similar.

Obtaining a faster asymptotic rate of convergence for gradient algorithms requires to extend the possible choices for the step-size  $\gamma_k$ . Rewrite the updating rule (5) in terms of iteration applied to the probability measure  $v_k$ ,

$$v_{k+1}(\lambda) = \frac{(1 - \gamma_k \lambda)^2}{1 - 2\gamma_k \mu_1^{(k)} + \gamma_k^2 \mu_2^{(k)}} v_k(\lambda) = \frac{(\lambda - \beta_k)^2}{\beta_k^2 - 2\beta_k \mu_1^{(k)} + \mu_2^{(k)}} v_k(\lambda), \quad (8)$$

where  $\beta_k = 1/\gamma_k$  and  $v_k(\lambda)$  is the weight assigned by the measure  $v_k$  to the point  $\lambda$ . The roots  $\beta_k$  in (8) are the key control variables for a gradient algorithm. Different strategies for choosing  $\beta_k$  give different families of algorithms. Note that the only information about  $v_k$  one has access to corresponds to its moments  $\mu_\alpha^{(k)}$ ,  $\alpha = 1, 2, \dots$ . Many of the examples of algorithms presented in [6], with  $\beta_k$  a function of  $\mu_1^{(k)}$  and  $\mu_2^{(k)}$ , exhibit a much faster asymptotic rate of convergence than  $R_{\max}$  (it seems that allowing  $\beta_k$  to depend on more moments  $\mu_\alpha^{(k)}$  does not yield further improvement in the rate of convergence). Fast convergence (small  $R$ ) is observed for algorithms that

exhibit a chaotic-type behaviour in  $\mathbb{R}^d$ , which makes their theoretical study difficult. The same is true for some algorithms for which  $\beta_k$  is allowed to depend on moments of several previous measures  $\nu_{k-i}$ ,  $i = 1, \dots, u$ . For instance, in the Barzilai-Borwein algorithm [2],  $\beta_k$  is either  $\mu_1^{(k-1)}$  or  $\mu_2^{(k-1)}/\mu_1^{(k-1)}$ .

Conjugate gradient,  $s$ -step optimal, MINRES and other algorithms based on Krylov spaces do not use gradient directions for their successive iterations, see, e.g., [8]. However, when analyzing their behaviour, one can construct an equivalent sequence of iterations following the gradient directions with control variables  $\beta_k$  depending on  $k$  and on moments of previous measures  $\nu_{k-i}$ ,  $i = 0, 1, 2, \dots$ . The conjugate gradient algorithm in  $\mathbb{R}^d$  converges in  $d$  iterations. When  $d$  is large, preserving the conjugacy of successive directions is difficult and restarting the algorithm after each sequence of  $s$  iterations is recommended. This corresponds to the  $s$ -step optimal gradient algorithm, see [5, 13], which does not have finite convergence but whose guaranteed asymptotic rate of convergence is

$$R_s^* = \left( \frac{R_\infty^{s/2} + R_\infty^{-s/2}}{2} \right)^{-2/s} = T_s^{-2/s} \left( \frac{\rho + 1}{\rho - 1} \right) \quad (9)$$

where

$$R_\infty = \lim_{s \rightarrow \infty} R_s^* = \left( \frac{\sqrt{\rho} - 1}{\sqrt{\rho} + 1} \right)^2$$

and  $T_s(\cdot)$  is the  $s$ -th Chebyshev polynomial:

$$T_s(t) = \cos[s \arccos(t)] = \frac{(t + \sqrt{t^2 - 1})^s + (t - \sqrt{t^2 - 1})^s}{2}.$$

In this paper we propose a family of gradient algorithms based on simple rules for choosing the sequence of control variables  $\beta_k$ . The main idea is to force  $\nu_k(\lambda_j)$ ,  $j = 2, \dots, d - 1$ , to tend to zero as  $k \rightarrow \infty$ . The measure  $\nu_k$ , which summarizes the state of the iterates at step  $k$ , is then almost fully characterized by  $\nu_k(m)$ , which facilitates the analysis of the asymptotic behaviour. Furthermore, we show that the sequence  $\{\beta_k\}$  can be chosen independently of  $\{\nu_k\}$  while ensuring that the asymptotic rate of convergence is arbitrarily close to  $R_\infty$ . This independence of  $\{\beta_k\}$  on  $\{\nu_k\}$  makes the algorithms at the same time simple and robust with respect to the precision of calculations. Also, the step-sizes  $\gamma_k = 1/\beta_k$ ,  $k = 1, 2, \dots$  are simpler to calculate than those of the steepest-descent algorithm. Convergence rates close to  $R_\infty$  are obtained when the  $\beta_k$ 's are constructed so that their asymptotic distribution is close to a distribution with the arcsine density.

The worst-case rate  $R_s^*$  can be reached for the  $s$ -step optimal gradient when  $d > s$ , in the sense that there exist eigenvalues  $\lambda_i$  and initial point  $x_0$  for the algorithm such that the rate of convergence after  $s$  iterations is exactly  $R_s^*$  (and the behavior in terms of renormalized gradient  $z_k$  is then periodic with period  $s$ ), see [5, 13]. The same is true for the conjugate gradient algorithm: for  $s < d$  there exist eigenvalues  $\lambda_i$  and a starting point  $x_0$  such that the convergence rate after  $s$  iterations is exactly  $R_s^*$ .

If  $d$  is large (relative to the total number of iterations),  $s$  is not very large and the eigenvalues of  $A$  are well-spread in the spectral interval  $[m, M]$ , then the actual rates

(per one matrix-vector multiplication) of the MINRES and other optimal methods based on the use of  $s$ -dimensional Krylov spaces are very close to  $R_s^*$  and are often larger than  $R_\infty$ . Bearing in mind that the asymptotic rates of the algorithms suggested below can be arbitrarily close to  $R_\infty$  and these algorithms are extremely simple and robust, these algorithms may be preferable to MINRES and other Krylov space based methods for large-scale quadratic optimization problems.

The paper is organized as follows. In Sect. 2 we show that for a suitable choice of the sequence  $\{\beta_k\}$  the algorithm attracts to the plane spanned by the eigenvectors associated with  $\lambda_1 = m$  and  $\lambda_d = M$ . In Sect. 3, we assume that the values of  $m$  and  $M$  are known and give the expression of the asymptotic rate of convergence of the algorithm in the case where the  $\beta_k$ 's are generated by pairs symmetric with respect to  $(m + M)/2$ . Several examples are presented, some with a rate arbitrarily close to  $R_\infty$ . The case where  $m$  and  $M$  are unknown is considered in Sect. 4 where a practical algorithm is suggested and some simulation results are presented. Finally, the infinite dimensional situation where  $f(\cdot)$  is defined on a Hilbert space is considered Sect. 5.

## 2 Attraction of the sequence $\{v_k\}$ to the set of measures supported at $m$ and $M$

**Theorem 1** Assume that  $\beta_k > 0$ ,  $\beta_k \notin \{m, M\}$  for all  $k$  and that the sequence  $\{\beta_k\}$  has asymptotic distribution function  $F(\beta)$  which is supported on an interval  $[m', M']$  with  $0 < m' \leq M' < \infty$ . Suppose, moreover, that the limiting distribution satisfies

$$\int \log(\beta - \lambda)^2 dF(\beta) < \max \left\{ \int \log(M - \beta)^2 dF(\beta), \int \log(\beta - m)^2 dF(\beta) \right\},$$

$$\forall \lambda \in \{\lambda_2, \dots, \lambda_{d-1}\}. \quad (10)$$

Then, the gradient algorithm associated with the sequence  $\{\beta_k\}$  is such that  $\lim_{k \rightarrow \infty} v_k(\lambda_i) = 0$  for all  $i = 2, \dots, d - 1$ . Furthermore, there exist constants  $C > 0$ ,  $k_0 > 0$  and  $0 \leq \theta < 1$  such that

$$\sum_{i=2}^{d-1} v_k(\lambda_i) \leq C\theta^k \quad \text{for } k > k_0. \quad (11)$$

*Proof* The fact that the sequence  $\{\beta_k\}$  has asymptotic distribution function  $F(\beta)$  implies

$$\lim_{k \rightarrow \infty} \frac{1}{k} \sum_{j=0}^{k-1} h(\beta_j) = \int h(\beta) dF(\beta) \quad (12)$$

for any continuous function  $h(\cdot)$  such that  $\int |h(\beta)| dF(\beta) < \infty$ , see [7]. Define

$$H_k(\lambda) = C_k(\lambda - \beta_0)^2(\lambda - \beta_1)^2 \cdots (\lambda - \beta_{k-1})^2, \quad (13)$$

with  $C_k$  a normalizing constant such that  $v_k(\lambda) = H_k(\lambda)v_0(\lambda)$  in (8), and assume that

$$\int \log(M - \beta)^2 dF(\beta) \leq \int \log(\beta - m)^2 dF(\beta) \quad (14)$$

(if this inequality is not met,  $m$  should be replaced with  $M$  in all considerations below). Define the sum

$$S_k(\lambda, m) = \frac{1}{k} \log \frac{H_k(\lambda)}{H_k(m)} = -\frac{1}{k} \sum_{j=0}^{k-1} \log(\beta_j - m)^2 + \frac{1}{k} \sum_{j=0}^{k-1} \log(\lambda - \beta_j)^2 \quad (15)$$

and consider the first sum  $I_k(m) = (1/k) \sum_{j=0}^{k-1} \log(\beta_j - m)^2$  in the right-hand side of (15) and the related integral  $I(m) = \int \log(\beta - m)^2 dF(\beta)$ . Since the c.d.f.  $F(\cdot)$  is supported on a bounded interval  $[m', M']$  we have  $I(m) < \infty$ . The assumptions (10) and (14) imply  $I(m) > -\infty$  and the property (12) then gives the convergence  $I_k(m) \rightarrow I(m)$  as  $k \rightarrow \infty$ .

Consider now the second sum  $I_k(\lambda) = (1/k) \sum_{j=0}^{k-1} \log(\beta_j - \lambda)^2$  in the right-hand side of (15) and the related integral  $I(\lambda) = \int \log(\beta - \lambda)^2 dF(\beta)$ . Since the c.d.f.  $F(\cdot)$  is supported on a bounded interval, the integral  $I(\lambda)$  is properly defined but may equal  $-\infty$  (for example, if the c.d.f.  $F(\cdot)$  has a discontinuity at the point  $\lambda$ ). If  $I(\lambda) = -\infty$  then as  $k \rightarrow \infty$  the sum  $I_k(\lambda)$  tends to  $-\infty$  too. If  $I(\lambda) > -\infty$  then either  $I_k(\lambda) = -\infty$  for all  $k$  large enough (when at least one  $\beta_j$  is equal to  $\lambda$ ) or (12) implies that  $I_k(\lambda)$  tends to  $I(\lambda)$  as  $k \rightarrow \infty$ .

Therefore, from (10),  $S_k(\lambda, m)$  tends to a negative value (possibly  $-\infty$ ) as  $k \rightarrow \infty$ . This implies that there exists  $k_0 \geq 0$  and  $\delta > 0$  such that for all  $k \geq k_0$  and  $\lambda \in \{\lambda_2, \dots, \lambda_{d-1}\}$

$$S_k(\lambda, m) = \frac{1}{k} \log \frac{H_k(\lambda)}{H_k(m)} \leq -\delta; \quad (16)$$

that is,  $H_k(\lambda)/H_k(m) \leq \theta^k$ , where  $\theta = \exp(-\delta) < 1$ . This yields  $\sum_{i=2}^{d-1} v_k(\lambda_i) \leq \theta^k (\sum_{i=2}^{d-1} v_0(\lambda_i))/v_0(m)$  for  $k > k_0$ , hence (11). The result  $\lim_{k \rightarrow \infty} v_k(\lambda_i) = 0$  for  $i = 2, \dots, d-1$  obviously follows from (11).  $\square$

**Remark 1** The sequence  $\{\beta_k\}$  can be assumed random, for instance formed by independent and identically distributed random variables. In this case, all the statements are true with probability one. When the  $\beta_k$ 's are simply independent, with  $\{F_k\}$  the sequence of corresponding distribution functions and  $(1/k) \sum_{j=0}^{k-1} F_j$  converging weakly to  $F$  as  $k$  tends to infinity, one may refer to [3, Th. 2.5.3, p. 36] for a property similar to (12).

**Remark 2** Typically, the spectrum of  $A$  is unknown. In that case, the condition (10) can be replaced with the more restrictive one

$$\int \log(\beta - \lambda)^2 dF(\beta) < \max \left\{ \int \log(M - \beta)^2 dF(\beta), \int \log(\beta - m)^2 dF(\beta) \right\},$$

$$\forall \lambda \in (m, M). \quad (17)$$

**Remark 3** If the distribution with c.d.f.  $F(\cdot)$  is symmetric with respect to  $(m + M)/2$ , then we have  $\int \log(M - \beta)^2 dF(\beta) = \int \log(\beta - m)^2 dF(\beta)$  and therefore the con-

dition (17) simplifies to

$$\int \log(\beta - \lambda)^2 dF(\beta) < \int \log(\beta - m)^2 dF(\beta), \quad \forall \lambda \in (m, M). \quad (18)$$

*Remark 4* Note that the support  $[m', M']$  of the distribution with c.d.f.  $F(\cdot)$  could be different from  $[m, M]$  and does not have to be a subset of  $[m, M]$ .

*Remark 5* The results of Theorem 1 also apply when  $\beta_k$  depends on the moments of previous measures  $\nu_{k-i}$ ,  $i = 0, 1, 2, \dots$

*Example 1* For the steepest-descent algorithm with  $\beta_k = \mu_1^{(k)}$ , the limiting measure for  $\{\beta_k\}$  is the two-point measure assigning weights  $1/2$  at  $z$  and  $m + M - z$  for some  $z \in (m, M)$ . The condition (17) then simply expresses the property that two successive iterations (8) of the algorithm asymptotically give a larger increase of the weights at the endpoints  $m$  and  $M$  than at any other point in the interval  $(m, M)$ ; that is,

$$(z - m)^2(M - z)^2 > (z - \lambda)^2(m + M - z - \lambda)^2, \quad \forall \lambda \in (m, M). \quad (19)$$

Since for all  $z$  the only maximum of  $(z - \lambda)^2(m + M - z - \lambda)^2$  with respect to  $\lambda \in (m, M)$  is at  $\lambda^* = (m + M)/2$ , the inequality (19) can be rewritten as  $(z - m)^2(M - z)^2 > (z - \lambda^*)^2(m + M - z - \lambda^*)^2$ , which gives

$$z \in \left( \frac{1}{2}(m + M) - \frac{1}{2\sqrt{2}}(M - m), \frac{1}{2}(m + M) + \frac{1}{2\sqrt{2}}(M - m) \right). \quad (20)$$

This corresponds to the definition of the stability interval for the attractor in [10, 12]. A similar result holds for all gradient-type algorithms from the family considered in [12].

*Example 2* If we choose  $\beta_k = \sqrt{\mu_2^{(k)}}$ , then the limiting measure for  $\{\beta_k\}$  is the delta-measure concentrated at the point  $\lambda^* = (m + M)/2$ ; as a consequence, the asymptotic rate for the related gradient algorithm is  $R_{\max}$ . Proof of these facts can be found in [4] and [6], Sect. 2.7.

### 3 Asymptotic rate for symmetrically placed control variables

#### 3.1 Main result

**Theorem 2** Assume that the conditions of Theorem 1 are satisfied and that, moreover, the control variables  $\beta_k$  are generated by symmetric pairs for large  $k$ ; that is,  $\beta_{2j+1} = M + m - \beta_{2j}$  for all  $j \geq j_0$ , with  $\beta_{2j} \in [m + \varepsilon, M - \varepsilon]$  for some  $\varepsilon \in (0, (M - m)/2)$ . Then, the asymptotic rate  $R$  satisfies

$$\log R = \int \log \left| \frac{(M - \beta)(\beta - m)}{\beta(m + M - \beta)} \right| dF(\beta) = \int \log \frac{(\beta - m)^2}{\beta^2} dF(\beta). \quad (21)$$

*Proof* First note that dividing (4) through by  $(g_k, g_k)$  gives the following expression for the rate  $r_k$ ,

$$r_k = 1 - 2\gamma_k \frac{(Ag_k, g_k)}{(g_k, g_k)} + \gamma_k^2 \frac{(A^2 g_k, g_k)}{(g_k, g_k)} = 1 - 2\mu_1^{(k)} / \beta_k + \mu_2^{(k)} / \beta_k^2. \quad (22)$$

Consider a measure  $\nu$  with weights  $p$  and  $1 - p$  at  $m$  and  $M$  respectively,  $0 < p < 1$ . Apply two successive iterations (8) with control parameters  $\beta$  and  $\beta' = m + M - \beta$  to this measure. The product of the two successive rates does not depend on  $p$  and is equal to  $R_2^2(\beta) = (M - \beta)^2(\beta - m)^2 / [\beta(m + M - \beta)]^2$ .

According to Theorem 1,  $\nu_k$  tends to be supported at  $m$  and  $M$  and the rate of convergence is exponential. We thus obtain for two successive iterations with control variables  $\beta_{2j}$  and  $\beta_{2j+1} = m + M - \beta_{2j}$

$$R_2^2(\beta_{2j}) \left[ 1 - \frac{A\theta^{2j}}{R_2^2(\beta_{2j})} \right] < r_{2j}r_{2j+1} < R_2^2(\beta_{2j}) \left[ 1 + \frac{A\theta^{2j}}{R_2^2(\beta_{2j})} \right]$$

for some  $A > 0$  and  $j > k_0/2$ , see Theorem 1. Since  $\beta_{2j} \in [m + \varepsilon, M - \varepsilon]$ , we have  $R_2^2(\beta_{2j}) \geq R_2^2(m + \varepsilon) = \varepsilon(M - m - \varepsilon) / [(m + \varepsilon)(M - \varepsilon)] > 0$ . Therefore,

$$\log R_2(\beta_{2j}) - B\theta^{2j} < \log \sqrt{r_{2j}r_{2j+1}} < \log R_2(\beta_{2j}) + B\theta^{2j},$$

with  $B = A/R_2^2(m + \varepsilon)$ , for  $j$  large enough. Since  $\sum_{j=0}^{\infty} \theta^{2j} = 1/(1 - \theta^2) < \infty$ , we obtain from (12),

$$\log R = \lim_{k \rightarrow \infty} \frac{1}{k} \sum_{j=0}^{k-1} \log \sqrt{r_{2j}r_{2j+1}} = \int \log R_2(\beta) dF(\beta),$$

hence the first expression in (21). The second expression follows from the fact that the c.d.f.  $F(\cdot)$  is symmetric with respect to  $(m + M)/2$ .  $\square$

*Example 3* (Uniform density) Let the distribution with c.d.f.  $F(\cdot)$  be uniform with density  $p(\beta) = 1/(M' - m')$ ,  $\beta \in [m', M']$ , with  $m' = m + \varepsilon$ ,  $M' = M - \varepsilon$  and  $0 < \varepsilon < (M - m)/2$ . Then the asymptotic rate of convergence is

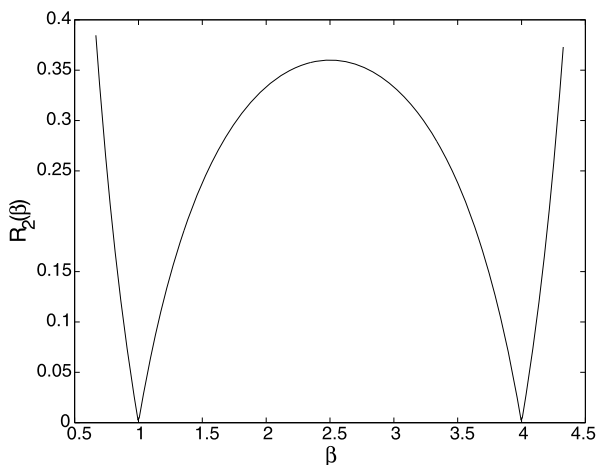
$$\begin{aligned} R_{\text{uniform}, \varepsilon} &= \exp \left\{ \frac{1}{M' - m'} \int_{m'}^{M'} \log \frac{(\beta - m')^2}{\beta^2} d\beta \right\} \\ &= (M' - m')^2 \exp \left\{ -2 \frac{M' \log M' - m' \log m'}{M' - m'} \right\}. \end{aligned} \quad (23)$$

*Remark 6* One can easily check that the result stated in Theorem 2 holds for other definitions for the rate of convergence, see, e.g., [12, Th. 6]. For instance, the rate

$$r'_k = \frac{f(x_{k+1}) - f^*}{f(x_k) - f^*} = \frac{(A^{-1}g_{k+1}, g_{k+1})}{(A^{-1}g_k, g_k)},$$



**Fig. 1**  $R_2(\beta)$  for  $m = 1, M = 4$



where  $f^* = \min_x f(x)$ , can be written as

$$r'_k = 1 - 2/(\mu_{-1}^{(k)}\beta_k) + \mu_1^{(k)}/(\mu_{-1}^{(k)}\beta_k^2) \quad (24)$$

and the corresponding asymptotic rate  $R' = \lim_{n \rightarrow \infty} (\prod_{k=0}^{n-1} r'_k)^{1/n}$  is equal to  $R$  which can be computed by (21).

*Remark 7* The shape of  $R_2(\beta)$  as a function of  $\beta$  shows that fast convergence is obtained for  $\beta$  close to  $m$  or  $M$ , see Fig. 1, hence the interest of taking  $\varepsilon$  small in Theorem 2.

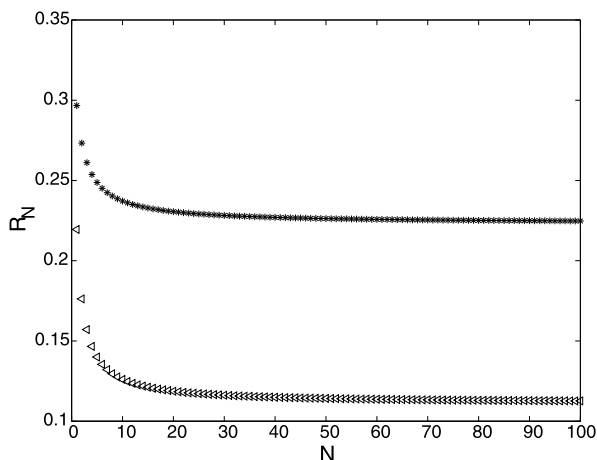
*Remark 8* When  $\nu_k$  is a two-point measure supported at  $m$  and  $M$ , two iterations of (8) with  $\beta_{k+1} = M + m - \beta_k$  give  $\nu_{k+2} = \nu_k$ . Under the conditions of Theorem 2 the measure  $\nu_k$  thus converges to a measure  $\bar{\nu}_k = p_k\delta_m + (1 - p_k)\delta_M$  supported at  $m$  and  $M$ , with  $p_{2j}$  tending to a constant  $p_\infty$  as  $j$  tends to infinity. The limiting distribution of the sequence  $\{p_{2j+1}\}$  depends on  $F(\cdot)$  and  $p_\infty$ , while the value of  $p_\infty$  depends on the initial measure  $\nu_0$  and the spectrum of  $A$ .

### 3.2 Finite collection of control variables

Assume that the points  $\beta_0, \beta_1, \dots$  are generated in repeated groups  $B = \{\beta_0, \dots, \beta_N\}$  of  $N + 1$  points in  $(m, M)$ ,  $N \geq 0$ . Additionally, the points in  $B$  are symmetric with respect to  $(m + M)/2$ . We may always assume that  $\beta_0 \leq \dots \leq \beta_N$ . In this case, if  $N$  is even then  $\beta_{N/2} = (m + M)/2$ . The condition (18) now becomes

$$\sum_{j=0}^N \log(\beta_j - \lambda)^2 < \sum_{j=0}^N \log(\beta_j - m)^2, \quad \forall \lambda \in (m, M). \quad (25)$$

**Fig. 2** Asymptotic rate of convergence (26) for  $m = 1$  and  $M = 4$  when the  $\beta_j$ 's are on the uniform grid (stars) and when they correspond to Chebyshev points (triangles,  $\varepsilon = 10^{-6}$ )



If this condition is met then the asymptotic rate is

$$R = R_N = \left[ \prod_{j=0}^N \frac{(\beta_j - m)^2}{\beta_j^2} \right]^{1/(N+1)}. \quad (26)$$

**Example 4 (Uniform grid)** Assume that for some integer  $N \geq 0$ ,

$$B = \{\beta_0, \dots, \beta_N\} \quad \text{with} \quad \beta_i = m + \frac{i + \frac{1}{2}}{N+1}(M - m), \quad i = 0, 1, \dots, N. \quad (27)$$

It is easy to see that the condition (25) is met. The rate  $R_N$  computed by (26) is given by

$$R_N = \left( \frac{\Gamma^2(N + 3/2) \Gamma^2(\frac{m+M+2Nm}{2(M-m)})}{\pi \Gamma^2(\frac{2NM+3M-m}{2(M-m)})} \right)^{1/(N+1)},$$

where  $\Gamma(\cdot)$  is the gamma-function. The value of  $R_N$  for  $m = 1$ ,  $M = 4$  is plotted in Fig. 2 as a function of  $N$ . Asymptotically, as  $N \rightarrow \infty$ ,  $R_N$  approaches  $R_{\text{uniform},0}$  defined in (23) (with  $R_{\text{uniform},0} \simeq 0.2232$  for  $m = 1$ ,  $M = 4$ ). Instead of using the  $\beta_i$ 's according to (27) for large  $N$ , one can generate the sequence  $\{\beta_i\}$  using, for example, the Bernoulli shift:

$$H_B(t) = 2t[\text{mod } 1], \quad t \in (0, 1), \quad (28)$$

with  $\beta_0$  randomly chosen in  $(m', M')$ , and for all  $j = 0, 1, 2, \dots$

$$\beta_{2j+1} = M' + m' - \beta_{2j}, \quad \beta_{2j+2} = m' + (M' - m')H_B\left(\frac{\beta_{2j} - m'}{M' - m'}\right),$$

with  $m' = m + \varepsilon$ ,  $M' = M - \varepsilon$  and  $0 < \varepsilon < (M - m)/2$ .

**Example 5** (Nearly optimal  $N + 1$  points) Consider first the case  $N = 1$ . When the condition (25) is satisfied, the asymptotic rate is  $R_2 = |M - \beta| |\beta - m| / [\beta |m + M - \beta|]$ , see the proof of Theorem 2. For  $\beta \in [m, M]$ ,  $R_2$  improves when  $|\beta - (m + M)/2|$  increases and reaches its minimum value, zero, at  $\beta \in \{m, M\}$ , see Remark 7. Condition (25) imposes that  $\beta$  belongs to the interval (20), by choosing  $\beta$  sufficiently close to  $(m + M)/2 \pm (M - m)/(2\sqrt{2})$  one makes the rate arbitrarily close to  $R_2^*$ , with  $R_s^*$  defined by (9).

Take now  $N = 2$ , with  $\beta_0 = \beta$ ,  $\beta_1 = (m + M)/2$  and  $\beta_2 = m + M - \beta$ . Similarly to the previous case, condition (25) imposes that  $\beta$  belongs to the interval  $((m + M)/2 - \sqrt{3}(M - m)/4, (m + M)/2 + \sqrt{3}(M - m)/4)$ , with the rate  $R_3$  getting close to  $R_3^*$  for  $\beta$  close to  $(m + M)/2 \pm \sqrt{3}(M - m)/4$ .

By induction, one can show that the rate  $R_N$  can be made arbitrarily close to the value  $R_s^*$  defined by (9), with  $s = N + 1$ , when the  $N + 1$  points  $\beta_i$  are suitably chosen and are constructed from the roots of Chebyshev polynomials. This construction is considered in the next example. (Note that the fact that  $R_N$  can be made arbitrarily close to  $R_{N+1}^*$  is not a coincidence: the worst case analysis of the  $s$ -step optimal gradient algorithm, which yields the rate  $R_s^*$ , corresponds to the situation where the  $\beta_i$ 's are rescaled roots of the  $s$ -th order Chebyshev polynomial, see [5].)

**Example 6** (Chebyshev points) Chebyshev points are defined by

$$t_k = \cos\left(\frac{\pi}{2} \frac{2k + 1}{N + 1}\right), \quad k = 0, \dots, N$$

and correspond to the roots of  $T_{N+1}(x) = \cos((N + 1) \arccos x)$ , the Chebyshev polynomials of the first kind. These points are symmetric on  $(-1, 1)$ . The asymptotic density of the points  $\{t_k\}_0^N$ , as  $N \rightarrow \infty$ , is  $p(t) = 1/(\pi\sqrt{1 - t^2})$ ,  $t \in (-1, 1)$ .

Define

$$\beta_k = \frac{m + M}{2} + \frac{M - m - 2\varepsilon}{2} t_k, \quad k = 0, \dots, N,$$

where  $0 < \varepsilon < (M - m)/2$ . These points belong to the interval  $(m + \varepsilon, M - \varepsilon)$  and are symmetric with respect to  $(m + M)/2$ . As  $\varepsilon > 0$ , the condition (25) holds. The rate  $R_N$  computed by (26) is plotted in Figure 2 as a function of  $N$  for  $m = 1$ ,  $M = 4$  and  $\varepsilon = 10^{-6}$ . Asymptotically, as  $N \rightarrow \infty$ ,  $R_N$  approaches  $R_{\arcsin, \varepsilon}$  defined below in (31).

### 3.3 Control variables with arcsine density on a subinterval of $[m, M]$

Let us assume that the distribution with c.d.f.  $F(\cdot)$  has the density

$$p_\varepsilon(\beta) = \frac{1}{\pi \sqrt{(\beta - m') (M' - \beta)}}, \quad m' \leq \beta \leq M', \quad (29)$$

where  $m' = m + \varepsilon$ ,  $M' = M - \varepsilon$  and  $0 < \varepsilon < (M - m)/2$ . The density (29) is called the arcsine density on the interval  $[m', M']$ .

The sequence of points  $\{\beta_i\}$  can be generated using, for example, the logistic map

$$H_L(x) = 4x(1-x), \quad x \in (0, 1), \quad (30)$$

with  $\beta_0$  randomly chosen in  $(m', M')$ , and for all  $j = 0, 1, 2, \dots$

$$\beta_{2j+1} = M' + m' - \beta_{2j}, \quad \beta_{2j+2} = m' + (M' - m')H_L\left(\frac{\beta_{2j} - m'}{M' - m'}\right).$$

Note that the control variables  $\beta_j$  are placed symmetrically in the interval  $[m, M]$ . We show below that the condition (17) holds for each  $\varepsilon > 0$ . According to (21), the asymptotic rate of convergence is then

$$R_{\arcsin, \varepsilon} = \exp \left\{ \int_{m'}^{M'} \log \frac{(\beta - m)^2}{\beta^2} p_\varepsilon(\beta) d\beta \right\} \quad (31)$$

and we show below that

$$R_{\arcsin, \varepsilon} = \left( \frac{M - m + 2\sqrt{\varepsilon(M - m - \varepsilon)}}{M + m + 2\sqrt{(M - \varepsilon)(m + \varepsilon)}} \right)^2. \quad (32)$$

For  $\varepsilon = 0$  this gives  $R_{\arcsin, 0} = R_\infty = (\sqrt{\rho} - 1)^2 / (\sqrt{\rho} + 1)^2$  where  $\rho = M/m$ . However, we cannot choose  $\varepsilon = 0$  as the condition (17) does not hold (we also show below that  $I(\lambda) = \int \log(\beta - \lambda)^2 dF(\beta) = 2 \log(M - m) - 4 \log 2$  for  $\lambda \in [m, M]$ ). Since the condition does hold for any  $\varepsilon > 0$ , the rate of the algorithm can be made arbitrarily close to  $R_\infty$ : for small  $\varepsilon > 0$ , we have

$$R_{\arcsin, \varepsilon} = R_\infty(1 + 4\sqrt{\varepsilon(M - m)}) + O(\varepsilon), \quad \varepsilon \rightarrow 0.$$

The rest of this section is devoted to the verification of (17) for  $\varepsilon > 0$  and to the derivation of the formula (32) for the rate  $R_{\arcsin, \varepsilon}$ . Define the integral

$$J(z, m', M') = \int_{m'}^{M'} \frac{\log(\beta - z)^2}{\pi \sqrt{(\beta - m')(M' - \beta)}} d\beta, \quad (33)$$

where  $-\infty < z < \infty$ . The changes of variables  $t = -1 + 2(\beta - m')/(M' - m')$  and  $x = -1 + 2(z - m')/(M' - m')$  in the integral (33) give

$$J(z, m', M') = 2 \log \frac{M' - m'}{2} + \frac{1}{\pi} I_x, \quad \text{where } I_x = \int_{-1}^1 \frac{\log(t - x)^2}{\sqrt{1 - t^2}} dt. \quad (34)$$

Assume first that  $|x| \leq 1$ . By changing the variable  $t = \cos \phi$  in the integral  $I_x$ , we obtain

$$I_x = \int_0^\pi \frac{\log(\cos \phi - x)^2}{\sin \phi} \sin \phi d\phi = \int_0^\pi \log(\cos \phi - x)^2 d\phi.$$

As  $\cos(\phi) = \cos(2\pi - \phi) \forall \phi$ , we have  $\int_0^\pi \log(\cos \phi - x)^2 d\phi = \int_\pi^{2\pi} \log(\cos \phi - x)^2 d\phi$ , which implies  $I_x = \frac{1}{2} \int_0^{2\pi} \log(\cos \phi - x)^2 d\phi$ . As we assume  $-1 \leq x \leq 1$  we

can set  $\psi = \arccos x$  (so that  $x = \cos \psi$ ). Using now the identity  $\cos \phi - \cos \psi = 2 \sin \frac{\psi - \phi}{2} \sin \frac{\phi + \psi}{2}$ , we obtain

$$\begin{aligned} I_x &= \frac{1}{2} \int_0^{2\pi} \log(\cos \phi - \cos \psi)^2 d\phi = \frac{1}{2} \int_0^{2\pi} \log \left( 2 \sin \frac{\phi - \psi}{2} \sin \frac{\phi + \psi}{2} \right)^2 d\phi \\ &= \frac{1}{2} \left[ \int_0^{2\pi} 2 \log 2 d\phi + \int_0^{2\pi} \log \left( \sin \frac{\phi - \psi}{2} \right)^2 d\phi \right. \\ &\quad \left. + \int_0^{2\pi} \log \left( \sin \frac{\phi + \psi}{2} \right)^2 d\phi \right] \\ &= 2\pi \log 2 + \left[ \int_0^\pi \log(\sin^2(\phi - \psi/2)) d\phi + \int_0^\pi \log(\sin^2(\phi + \psi/2)) d\phi \right]. \end{aligned}$$

The function  $t \rightarrow \sin^2 t$  is  $\pi$ -periodic and therefore for any  $\psi'$  we get

$$\int_0^\pi \log(\sin^2(\phi + \psi')) d\phi = \int_0^\pi \log(\sin^2(\phi)) d\phi = 2 \int_0^\pi \log(\sin \phi) d\phi.$$

This implies

$$\begin{aligned} I_x &= 2\pi \log 2 + 4 \int_0^\pi \log(\sin \phi) d\phi \\ &= 2\pi \log 2 - 4\pi \log 2 = -2\pi \log 2, \quad \forall x \in [-1, 1]. \end{aligned} \quad (35)$$

Assume now that  $|x| \geq 1$ . From (35) we have  $I_1 = -2\pi \log 2$  and differentiating  $I_x$  we get

$$I'_x = \left( \int_{-1}^1 \frac{\log(x-t)^2}{\sqrt{1-t^2}} dt \right)' = \frac{2\pi}{\sqrt{x^2-1}}.$$

Therefore, for  $x > 1$ ,

$$\begin{aligned} I_{-x} &= I_x = I_1 + \int_1^x I'_z dz = -2\pi \log 2 + \int_1^x \frac{2\pi}{\sqrt{z^2-1}} dz \\ &= -2\pi \log 2 + 2\pi \log \left( \frac{x + \sqrt{x^2-1}}{2} \right). \end{aligned} \quad (36)$$

Combining (35) and (36) we obtain

$$I_x = \int_{-1}^1 \frac{\log(t-x)^2}{\sqrt{1-t^2}} dt = \begin{cases} -2\pi \log 2 & \text{if } |x| \leq 1 \\ 2\pi \log(|x| + \sqrt{x^2-1}) - 2\pi \log 2 & \text{if } |x| \geq 1, \end{cases}$$

together with (34), it gives

$$J(z, m', M') = \begin{cases} 2\log(M' - m') - 4\log 2 & \text{if } m' \leq z \leq M' \\ 2\log(M' - m') + 2\log(|t_z| + \sqrt{t_z^2-1}) - 4\log 2 & \text{if } z < m' \text{ or } z > M', \end{cases} \quad (37)$$

where  $t_z = -1 + 2(z - m')/(M' - m')$ . Therefore,  $J(\lambda, m', M') < J(m, m', M') = J(M, m', M')$  for all  $\lambda$  in  $(m, M)$  and (17) is satisfied. The expression (32) for the rate  $R_{\arcsin, \varepsilon}$  easily follows from (37) and the representation  $R_{\arcsin, \varepsilon} = \exp[J(M, m', M') - J(0, m', M')]$  with  $m' = m + \varepsilon$  and  $M' = M - \varepsilon$ .

## 4 Estimation of $m$ , $M$ and a practical algorithm

### 4.1 Estimation of $m$ , $M$ and asymptotic behavior in the non symmetric case

The values of  $m$  and  $M$  can be easily estimated in the first iterations of the algorithm (3), for instance by computing the first moment  $\mu_1^{(j)}$  for several values of  $j = 0, 1, 2, \dots$  and taking

$$m_k = \min\{\mu_1^{(j)}, j = 0, \dots, k\}, \quad M_k = \max\{\mu_1^{(j)}, j = 0, \dots, k\} \quad (38)$$

as estimates. We then necessarily have  $m < m_k < M_k < M$  for  $k \geq 1$ .

Suppose that the estimation is stopped at some  $k_0$ , that is,  $m_k = m_{k_0}$  and  $M_k = M_{k_0}$  for all  $k > k_0$ . Then, under the conditions of Theorem 1 with  $m' = m_{k_0}$  and  $M' = M_{k_0}$  we have  $\sum_{i=2}^{d-1} v_k(\lambda_i) \leq C\theta^k$  for  $k$  larger than some  $k_1$  and constants  $C > 0$  and  $0 \leq \theta < 1$ . Suppose that the control variables are generated by pairs for  $k > k_0$ , as in Theorem 2, but with  $\beta_{2k+1} = M_{k_0} + m_{k_0} - \beta_{2k}$ , for all  $k > k_0$ .

If  $M_{k_0} + m_{k_0} = M + m$ , Theorem 2 applies and the asymptotic rate  $R$  satisfies (21). For instance, if the  $\beta_k$ 's are generated as in Sect. 3.3, and have the arcsine density on  $[m_{k_0}, M_{k_0}]$ , the asymptotic rate is  $R_{\arcsin, \varepsilon}$  with  $\varepsilon = m_{k_0} - m = M - M_{k_0}$ . Consider now the standard situation where  $M_{k_0} + m_{k_0} \neq M + m$  and suppose that  $M - M_{k_0} > m_{k_0} - m$ . The asymptotic distribution of the  $\beta_k$ 's, symmetric in  $[m_{k_0}, M_{k_0}]$ , is then biased towards  $m$  and  $v_k(m)$  tends to zero when  $k \rightarrow \infty$ . Following the same line as in the proof of Theorem 2, we obtain that the product of rates at two successive iterations for the delta measure at  $M$ , with control parameters respectively  $\beta$  and  $\beta' = M_{k_0} + m_{k_0} - \beta$ , is  $R_2^2 = (M - \beta)^2(M - \beta')^2/(\beta\beta')^2$ . The asymptotic rate then satisfies

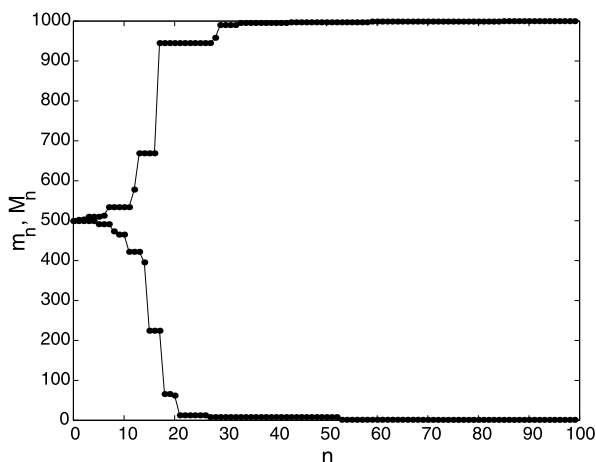
$$\log R = \int \log \left| \frac{(M - \beta)(M + \beta - M_{k_0} - m_{k_0})}{\beta(M_{k_0} + m_{k_0} - \beta)} \right| dF(\beta).$$

Similarly, supposing that  $M - M_{k_0} < m_{k_0} - m$  gives an asymptotic distribution of the  $\beta_k$ 's biased towards  $M$ , so that  $v_k(m)$  tends to 1 as  $k \rightarrow \infty$ , and the asymptotic rate satisfies

$$\log R = \int \log \left| \frac{(\beta - m)(M_{k_0} + m_{k_0} - \beta - m)}{\beta(M_{k_0} + m_{k_0} - \beta)} \right| dF(\beta).$$

Now, note that  $v_k(m) \rightarrow 0$  implies that  $\mu_1^{(k)} \rightarrow M$  and  $v_k(m) \rightarrow 1$  implies that  $\mu_1^{(k)} \rightarrow m$ ,  $k \rightarrow \infty$ , so that maintaining the adaptation of the estimation of  $m_k$  and  $M_k$  by (38) ensures that  $m_k \rightarrow m$  and  $M_k \rightarrow M$  as  $k \rightarrow \infty$ . This permits to recover the same asymptotic rates as Sect. 3.3, even in situations where  $m$  and  $M$  are unknown. Since the estimated values  $m_k$  and  $M_k$  quickly converge to  $m$  and  $M$ , see for instance Fig. 3, we need to generate the control variable  $\beta_k$  in  $[m_k + \varepsilon, M_k - \varepsilon]$  at iteration  $k$ . A practical algorithm is given below.

**Fig. 3** Convergence of the estimates  $m_n$  and  $M_n$  as functions of  $n$  ( $m = 1$ ,  $M = 1000$ ,  $d = 1000$ )



## 4.2 An algorithm based on the arcsine density

A possible algorithm is then as follows.

- Choose  $\tau$  as a small positive number (e.g.,  $\tau = 10^{-6}$ ), set  $z_0 = 0$ ;
- for  $k = 0, 1$ , set  $\beta_k = \mu_1^{(k)}$  (steepest-descent) and set  $m_1 = \min\{\mu_1^{(0)}, \mu_1^{(1)}\}$ ,  $M_1 = \max\{\mu_1^{(0)}, \mu_1^{(1)}\}$ ;
- for  $k > 1$ , set  $\varepsilon_k = \tau(M_{k-1} - m_{k-1})$  and generate the  $\beta_k$ 's by pairs:
  - for  $k = 2j$ , set  $z_j = \{\varphi + z_{j-1}\}$  and  $\beta_{2j} = m_k + \varepsilon_k + (\cos(\pi z_j) + 1)(M_k - m_k - 2\varepsilon_k)/2$ , where  $\{t\}$  denotes the fractional part of  $t$  and  $\varphi = (\sqrt{5} - 1)/2 \simeq 0.61803$ ;
  - for  $k = 2j + 1$ , set  $\beta_{2j+1} = M_k + m_k - \beta_{2j}$ ;
 set  $m_k = \min\{m_{k-1}, \mu_1^{(k)}\}$ ,  $M_k = \max\{M_{k-1}, \mu_1^{(k)}\}$ .

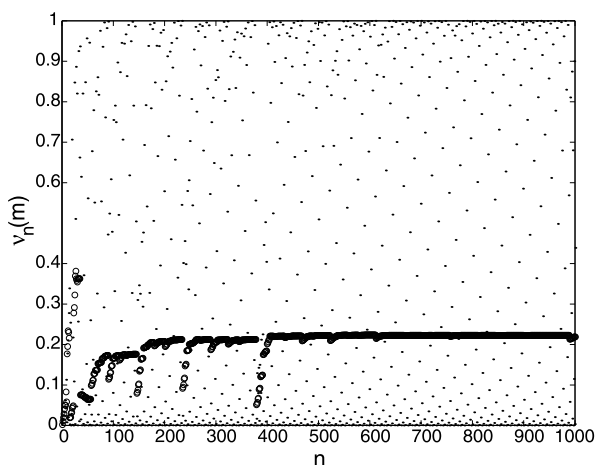
The sequence  $z_1, z_2, \dots$  is such that  $z_j = \{j\varphi\}$  so that the sequence is asymptotically uniform on  $[0, 1]$ , see, e.g., [7]. This implies that the asymptotic distribution of the sequence  $\beta_k$  has the arcsine density on  $[m + \varepsilon, M - \varepsilon]$  where  $\varepsilon = \tau(M - m)$ . From (32), the rate of the algorithm satisfies

$$\lim_{n \rightarrow \infty} R_n = R_{\arcsin, \tau(M-m)} = R_\infty(1 + 4\sqrt{\tau}) + \mathcal{O}(\tau), \quad \tau \rightarrow 0.$$

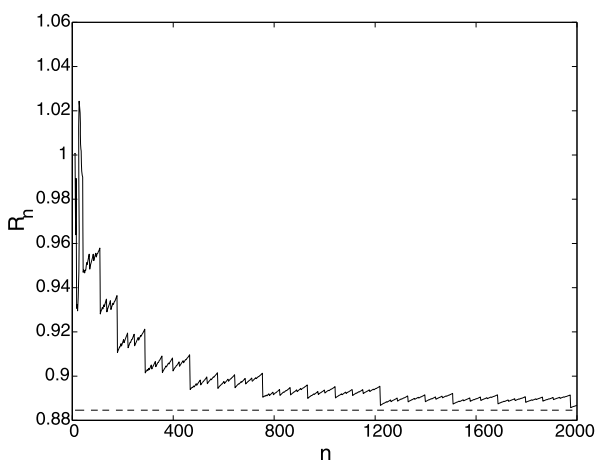
The dynamical system  $z_j = \{j\varphi\}$  generates a sequence in  $[0, 1]$  with much better uniformity characteristics than sequences generated by the Bernoulli shift (28). Since the logistic map (30) corresponds to a transformation of the Bernoulli shift, the construction above, based on  $z_j = \{j\varphi\}$ , produces a sequence of control variables  $\beta_k$  with better distribution characteristics than sequences generated with (30).

Figures 3, 4 and 5 illustrate the typical behavior of the algorithm above in a large-dimensional badly conditioned problem. In the example presented,  $d = 1000$ ,  $m = 1$ ,  $M = \rho = 1000$  and the eigenvalues  $\lambda_i$  are random and uniformly distributed on the interval  $[m, M]$  (one would obtain exactly the same plots if the eigenvalues were equally-spaced on  $[m, M]$ ).

**Fig. 4** Values  $v_n(m)$  as functions of  $n$ ; circles for  $n = 2j$ , dots for  $n = 2j + 1$  ( $\rho = 1000$ ,  $d = 1000$ )



**Fig. 5** Rate of convergence  $R_n$ , see (7), as a function of  $n$ ; the limiting value  $R_{\arcsin, \tau(M-m)}$  is indicated by the dashed line ( $m = 1$ ,  $M = 1000$ ,  $d = 1000$ ,  $\tau = 10^{-6}$ )



In terms of complexity of calculations, only the multiplications of  $d$ -dimensional vectors by the  $d \times d$  matrix  $A$  are expensive. The steepest descent algorithm requires the calculation of  $\beta_k = \mu_2^{(k)} / \mu_1^{(k)} = (Ag_k, Ag_k) / (Ag_k, g_k)$  at iteration  $k$ . Having computed  $g_k$  and  $Ag_k$ , one may notice that next gradient  $g_{k+1}$  can be obtained as  $g_{k+1} = g_k - (1/\beta_k)Ag_k$ , so that only the computation of  $Ag_{k+1}$  is expensive at iteration  $k + 1$ . However, a long sequence of iterations of this type may produce an accumulation of rounding errors, and it is rather recommended to recalculate  $g_{k+1}$  from  $x_{k+1}$  by  $g_{k+1} = Ax_{k+1} - y$ , see (1). This then requires two multiplications by  $A$  at each steepest-descent iteration.

In the algorithm above, iteration  $k$  only requires the calculation of the gradient  $g_k = Ax_k - y$ , and thus only one multiplication by  $A$ . Notice that the estimation of  $m_k$  and  $M_k$  through the moments  $\mu_1^{(j)} = (Ag_j, g_j) / (g_j, g_j)$ , see (38), does not require the calculation of  $Ag_j$  at step  $k$ . Indeed, allowing a delay of one step in the estimation, we have  $(g_j, g_{j+1}) = (g_j, g_j - (1/\beta_j)Ag_j)$  so that  $\mu_1^{(j)}$  is obtained at



next step from

$$\mu_1^{(j)} = \beta_j \left[ 1 - \frac{(g_j, g_{j+1})}{(g_j, g_j)} \right].$$

Also, one may observe in Fig. 3 that the convergence of  $m_n$  and  $M_n$  to  $m$  and  $M$  respectively is very fast, so that the estimation can be stopped after a few iterations. On the whole, it makes iterations with the algorithm above about twice simpler than steepest-descent iterations (even when  $m$  and  $M$  are estimated), with much faster convergence.

## 5 Hilbert space case

In the Hilbert-space case,  $A$  is a self-adjoint operator and its spectrum  $\mathcal{S}_A$  is a closed subset of the interval  $[m, M]$  of the real line, with  $m, M \in \mathcal{S}_A$ . Let  $E_\lambda$  be the spectral family associated with  $A$  and define the measure  $\nu_k = d(E_\lambda z_k, z_k)$ ,  $m \leq \lambda \leq M$ , with  $z_k = g_k / \sqrt{(g_k, g_k)}$  the normalized gradient at  $x_k$ . We have  $(z_k, z_k) = 1 = \int_m^M \nu_k(d\lambda)$  and  $\nu_k$  is a probability measure on the Borel sets of  $(0, \infty)$ , satisfying  $\nu_k([m, M]) = 1$  for all  $k$  and with moments still defined by (6). One iteration of a gradient algorithm with control variable  $\beta_k$  thus gives in terms of  $\nu_k$

$$\nu_{k+1}(\mathcal{A}) = \frac{\int_{\mathcal{A}} (\lambda - \beta_k)^2 \nu_k(d\lambda)}{\beta_k^2 - 2\beta_k \mu_1^{(k)} + \mu_2^{(k)}},$$

for  $\mathcal{A}$  any measurable subset of  $[m, M]$ , see (8). The properties obtained for the finite dimensional case remain valid and only a few adaptations are required.

**Theorem 3** Assume that the sequence  $\{\beta_k\}$  has asymptotic distribution function  $F(\beta)$  which is supported on an interval  $[m', M'] = [m + \varepsilon, M - \varepsilon]$  with  $0 < \varepsilon < (M - m)/2$ . Suppose, moreover, that  $I(\lambda) = \int \log(\beta - \lambda)^2 dF(\beta)$  is a continuous function of  $\lambda$  on  $(m', M')$  and that

$$I(\lambda) < \max \left\{ \int \log(M - \beta)^2 dF(\beta), \int \log(\beta - m)^2 dF(\beta) \right\}, \quad \forall \lambda \in (m', M'), \quad (39)$$

and that  $\nu_0\{[m, m + \gamma]\} > 0$  and  $\nu_0\{(M - \gamma, M]\} > 0$  for all  $\gamma > 0$ . Then, the measure  $\nu_k$  converges to a two-point measure supported at  $m$  and  $M$ , in the sense that there exists  $k_0$  such that, for any function  $g(\lambda)$  continuous on  $[m, M]$  and any  $\delta > 0$ , there exists  $\gamma > 0$  such that

$$\begin{aligned} \max \left\{ \left| \int_m^C g(\lambda) \nu_k(d\lambda) - g(m) \int_m^C \nu_k(d\lambda) \right|, \right. \\ \left. \left| \int_C^M g(\lambda) \nu_k(d\lambda) - g(M) \int_C^M \nu_k(d\lambda) \right| \right\} < \delta + C_\gamma \alpha_\gamma^k, \quad k > k_0, \end{aligned}$$

where  $C = (m + M)/2$  and  $C_\gamma > 0$ ,  $\alpha_\gamma \in (0, 1)$  are constants depending on  $\gamma$ . If, moreover, the control variables  $\beta_k$  are generated by symmetric pairs for large  $k$ , that is,  $\beta_{2j+1} = M + m - \beta_{2j}$  for all  $j \geq j_0$ , then the asymptotic rate  $R$  satisfies (21).

*Proof* The proof of convergence of  $\nu_k$  to a two-point measure follows the same arguments as for Theorem 1. Suppose that  $F(\cdot)$  satisfies (14). We still have for the first term of the sum  $S_k(\lambda, m)$  defined by (15)

$$I_k(m) = \frac{1}{k} \sum_{j=0}^{k-1} \log(\beta_j - m)^2 \rightarrow I(m) = \int \log(\beta - m)^2 dF(\beta), \quad k \rightarrow \infty.$$

Concerning the second term  $I_k(\lambda) = (1/k) \sum_{j=0}^{k-1} \log(\beta_j - \lambda)^2$  we need now a bound uniform in  $\lambda$ , that is, we need to show that

$$\forall \epsilon > 0, \exists K_0 \text{ such that: } \sup_{\lambda \in (m', M')} I_k(\lambda) - I(\lambda) < \epsilon, \quad \forall k > K_0. \quad (40)$$

Take a ball  $\mathcal{B}(\lambda_1, \delta) = \{\lambda : |\lambda - \lambda_1| \leq \delta\}$  and consider  $\bar{a}_\delta(\beta) = \sup_{\lambda \in \mathcal{B}(\lambda_1, \delta)} \log(\beta - \lambda)^2$ , which is an increasing function of  $\delta$ ,  $\bar{a}_\delta(\beta) = 2 \log(|\beta - \lambda_1| + \delta)$ . We have

$$\lim_{\delta \rightarrow 0} \int \bar{a}_\delta(\beta) dF(\beta) = \int \left[ \lim_{\delta \rightarrow 0} \bar{a}_\delta(\beta) \right] dF(\beta) = I(\lambda_1)$$

and therefore, there exists  $\delta_1 = \delta_1(\lambda_1)$  such that  $\int \bar{a}_\delta(\beta) dF(\beta) < I(\lambda_1) + \epsilon/3$  for  $\delta < \delta_1$ . Now,

$$\sup_{\lambda \in \mathcal{B}(\lambda_1, \delta)} I_k(\lambda) \leq (1/k) \sum_{j=0}^{k-1} 2 \log(|\beta_j - \lambda| + \delta) < \int \bar{a}_\delta(\beta) dF(\beta) + \epsilon/3$$

for all  $k$  larger than some  $K_1 = K_1(\lambda_1, \delta)$ . Also, from the continuity of  $I(\lambda)$ , there exists  $\delta_2 = \delta_2(\lambda_1)$  such that  $\inf_{\lambda \in \mathcal{B}(\lambda_1, \delta)} I(\lambda) > I(\lambda_1) - \epsilon/3$  for  $\delta < \delta_2$ . Altogether it gives  $\sup_{\lambda \in \mathcal{B}(\lambda_1, \delta)} I_k(\lambda) - I(\lambda) < \epsilon$  for  $\delta < \delta_0(\lambda_1) = \min(\delta_1, \delta_2)$  and  $k > K_1$ . It only remains to cover  $[m', M']$  with a finite number of such balls  $\mathcal{B}(\lambda_i, \delta)$ , with  $\delta < \min_i \delta_0(\lambda_i)$  to obtain the result (40). Since  $\log(\beta - \lambda)^2$  is a decreasing (resp. increasing) function of  $\lambda$  in  $[m, m']$  (resp. in  $[M', M]$ ), together with the condition (39) it implies that for any set  $\mathcal{S} \subset (m, M)$ ,  $\limsup_{k \rightarrow \infty} \sup_{\lambda \in \mathcal{S}} S_k(\lambda, m) \leq -\delta$  for some  $\delta = \delta(\mathcal{S}) > 0$ . Therefore, there exists  $k_0$  such that,  $\forall k > k_0$ ,  $\sup_{\lambda \in (m', M')} H_k(\lambda) / H_k(m) \leq \theta_\epsilon^k$  where  $\theta_\epsilon = \exp(-\delta_\epsilon) < 1$ .

Consider now a function  $g(\lambda)$  continuous on  $[m, M]$  and define

$$\Delta_k = \left| \int_m^C g(\lambda) \nu_k(d\lambda) - g(m) \int_m^C \nu_k(d\lambda) \right|,$$

where  $C = (m + M)/2$ . We show below that

$$\forall \delta > 0, \exists \gamma > 0 \text{ such that } \Delta_k < \delta + 2 \frac{D_g}{\int_m^{m+\gamma} \nu_0(d\lambda)} \alpha_\gamma^k \text{ for all } k > k_0, \quad (41)$$

for some  $\alpha_\gamma < 1$ , where  $D_g = \max_{\lambda \in [m, C]} |g(\lambda) - g(m)|$ . We have  $\Delta_k < \int_m^C |g(\lambda) - g(m)| \nu_k(d\lambda) = \Delta_{k,1} + \Delta_{k,2} + \Delta_{k,3}$ , with

$$\Delta_{k,1} = \int_m^{m+2\gamma} |g(\lambda) - g(m)| \nu_k(d\lambda),$$

$$\Delta_{k,2} = \int_{m+2\gamma}^{m'} |g(\lambda) - g(m)| v_k(d\lambda),$$

$$\Delta_{k,3} = \int_{m'}^C |g(\lambda) - g(m)| v_k(d\lambda),$$

$\gamma < \varepsilon/2$ . From the continuity of  $g(\lambda)$ , we can take  $\gamma$  small enough to have  $\Delta_{k,1} < \delta \int_m^{m+2\gamma} v_k(d\lambda) \leq \delta$ . Next,  $\Delta_{k,2} < D_g \int_{m+2\gamma}^{m'} v_k(d\lambda) = D_g \int_{m+2\gamma}^{m'} H_k(\lambda) v_0(d\lambda)$  with  $H_k(\lambda)$  defined by (13). Since  $\beta_k \in [m', M']$  for all  $k$ ,  $H_k(\lambda)$  is a decreasing function of  $\lambda$  for  $\lambda \in [m, m']$ , and for  $m + 2\gamma < \lambda < m'$  it satisfies

$$H_k(\lambda) < H_k(m + 2\gamma) < H_k(m + \gamma) \left( \frac{M - m - \varepsilon - 2\gamma}{M - m - \varepsilon - \gamma} \right)^{2k}.$$

Since  $\int_m^{m+\gamma} v_k(d\lambda) = \int_m^{m+\gamma} H_k(\lambda) v_0(d\lambda) \geq H_k(m + \gamma) \int_m^{m+\gamma} v_0(d\lambda)$ , we obtain

$$\Delta_{k,2} < \frac{D_g}{\int_m^{m+\gamma} v_0(d\lambda)} \left[ \frac{M - m - \varepsilon - 2\gamma}{M - m - \varepsilon - \gamma} \right]^{2k}.$$

We also obtain for the last term,

$$\begin{aligned} \Delta_{k,3} &< D_g \int_{m'}^C H_k(\lambda) v_0(d\lambda) \\ &< D_g \theta_\varepsilon^k H_k(m) \int_{m'}^C v_0(d\lambda) < D_g \theta_\varepsilon^k H_k(m) \quad \text{for } k > k_0. \end{aligned}$$

For  $\lambda \in [m, m']$  we have  $H_k(\lambda)/H_k(m) \geq (m' - \lambda)^{2k}/\varepsilon^{2k}$  so that

$$\begin{aligned} 1 &\geq \int_m^{m+\gamma} v_k(d\lambda) \geq H_k(m) \int_m^{m+\gamma} [(m' - \lambda)/\varepsilon]^{2k} v_0(d\lambda) \\ &> H_k(m) [(\varepsilon - \gamma)/\varepsilon]^{2k} \int_m^{m+\gamma} v_0(d\lambda). \end{aligned}$$

Therefore, for  $k > k_0$ ,

$$\Delta_{k,3} < \frac{D_g}{\int_m^{m+\gamma} v_0(d\lambda)} \left[ \frac{\theta_\varepsilon \varepsilon^2}{(\varepsilon - \gamma)^2} \right]^k.$$

We have  $\theta_\varepsilon \varepsilon^2/(\varepsilon - \gamma)^2 < 1$  for  $\gamma < \varepsilon(1 - \sqrt{\theta_\varepsilon})$  so that (41) is satisfied for  $\alpha_\gamma = \max\{\theta_\varepsilon \varepsilon^2/(\varepsilon - \gamma)^2, (M - m - \varepsilon - 2\gamma)^2/(M - m - \varepsilon - \gamma)^2\}$  and  $\alpha_\gamma < 1$  for  $\gamma$  small enough. One can show a similar property for  $\Delta'_k = |\int_C^M g(\lambda) v_k(d\lambda) - g(M) \int_C^M v_k(d\lambda)|$ .

Finally, we apply the property above to  $g(\lambda) = \lambda$  and  $g(\lambda) = \lambda^2$  and, following the same line as in the proof of Theorem 2, we then obtain for the product of rates at

two successive iterations with control variables  $\beta_{2j}$  and  $\beta_{2j+1} = m + M - \beta_{2j}$ :

$$R_2^2(\beta_{2j}) \left[ 1 - \frac{A_\gamma \alpha_\gamma^{2j} + B\delta}{R_2^2(\beta_{2j})} \right] < r_{2j} r_{2j+1} < R_2^2(\beta_{2j}) \left[ 1 + \frac{A_\gamma \alpha_\gamma^{2j} + B\delta}{R_2^2(\beta_{2j})} \right],$$

for some  $A_\gamma > 0$ ,  $B > 0$  and  $j > k_0/2$ . Therefore,

$$\log R_2(\beta_{2j}) - A'_\gamma \alpha_\gamma^{2j} - B'\delta < \log \sqrt{r_{2j} r_{2j+1}} < \log R_2(\beta_{2j}) + A'_\gamma \alpha_\gamma^{2j} + B'\delta,$$

with  $A'_\gamma = A_\gamma/R_2^2(m + \varepsilon)$  and  $B' = B/R_2^2(m + \varepsilon)$ , for  $j$  large enough. Since  $\sum_{j=0}^\infty \alpha_\gamma^{2j} = 1/(1 - \alpha_\gamma^2) < \infty$ , we obtain from (12),

$$\begin{aligned} & \left| \log R - \int \log R_2(\beta) dF(\beta) \right| \\ &= \left| \lim_{k \rightarrow \infty} \frac{1}{k} \sum_{j=0}^{k-1} \log \sqrt{r_{2j} r_{2j+1}} - \int \log R_2(\beta) dF(\beta) \right| < B'\delta. \end{aligned}$$

Since  $\delta$  is arbitrary, the asymptotic rate of convergence is thus the same as in the finite dimensional case.  $\square$

**Remark 9** Note that the condition  $I(\lambda)$  being a continuous function of  $\lambda$  is satisfied for all examples considered in Sect. 3. It is also satisfied when the distribution function  $F(\cdot)$  has density  $\phi(\cdot)$  with derivative  $\phi'(\cdot)$  uniformly bounded on  $(m', M')$ . Indeed, one can write  $I(\lambda) = \int_{\lambda-M'}^{\lambda-m'} \phi(\lambda - t) \log t^2 dt$  which has derivative  $I'(\lambda) = \phi(m') \log(\lambda - m')^2 - \phi(M') \log(\lambda - M')^2 + \int_{m'}^{M'} \phi'(t) \log(\lambda - t)^2 dt$ ; this derivative is bounded, which implies the continuity of  $I(\lambda)$ .

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