

Introduction

The **score bootstrap** perturbs the score contributions of an M-estimator using IID, mean zero, unit variance weights. The distribution of the perturbed score statistics is then used for inference on the observed test statistic.

Motivation

Consider an M-estimator $\hat{\theta}$ obtained by solving the estimating equation:

$$\Psi_n(\theta) = \sum_{i=1}^n \psi_i(\theta) \stackrel{\text{Set}}{=} \mathbf{0}$$

By first order Taylor expansion:

$$\begin{aligned} (\hat{\theta} - \theta_0) &= -\dot{\Psi}_n(\theta_0)^{-1} \Psi_n(\theta_0) + \mathcal{O}_p(n^{-1}) \\ \sqrt{n}(\hat{\theta} - \theta_0) &= -\left[\frac{1}{n} \sum_{i=1}^n \frac{\partial \psi_i}{\partial \theta'}(\theta_0) \right]^{-1} \frac{1}{\sqrt{n}} \sum_{i=1}^n \psi_i(\theta_0) + \mathcal{O}_p(n^{-1/2}) \end{aligned}$$

Define \mathbf{A}_0 as the expected Jacobian of ψ_i w.r.t. θ :

$$\mathbf{A}_0 \equiv -E \left[\frac{\partial \psi_i}{\partial \theta'}(\theta_0) \right]$$

By LLN:

$$-\frac{1}{n} \sum_{i=1}^n \frac{\partial \psi_i}{\partial \theta'}(\theta_0) = \mathbf{A}_0 + o_p(1)$$

Thus the quantity $\sqrt{n}(\hat{\theta} - \theta_0)$ is expressible as:

$$\sqrt{n}(\hat{\theta} - \theta_0) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \mathbf{A}_0^{-1} \psi_i(\theta_0) + o_p(1)$$

The limiting distribution is:

$$\sqrt{n}(\hat{\theta} - \theta_0) \xrightarrow{\mathcal{L}} N(\mathbf{0}, \mathbf{A}_0^{-1} \mathbf{B}_0 \mathbf{A}_0^{-T})$$

Here \mathbf{B}_0 is the expected outer product of the score equation:

$$\mathbf{B}_0 = E[\psi_i \psi_i'(\theta_0)]$$

Let (ω_i) denote IID mean zero, unit variance weights. Form the perturbed score:

$$\tilde{\Psi}_n(\theta) \equiv \sum_{i=1}^n \omega_i \psi_i(\theta)$$

Denote by $\tilde{\theta}$ a solution to the equation:

$$\Psi_n(\theta) - \tilde{\Psi}_n(\tilde{\theta}) \stackrel{\text{Set}}{=} \mathbf{0}$$

By first order Taylor expansion:

$$\Psi_n(\tilde{\theta}) = \Psi_n(\theta_0) + \dot{\Psi}_n(\theta_0)(\tilde{\theta} - \theta_0) + \mathcal{O}_p(n^{-1})$$

Subtract $\tilde{\Psi}_n(\hat{\theta})$ from each side, then:

$$\mathbf{0} = \Psi_n(\theta_0) - \tilde{\Psi}_n(\hat{\theta}) + \dot{\Psi}_n(\theta_0)(\tilde{\theta} - \theta_0) + \mathcal{O}_p(n^{-1})$$

Introduce and remove $\hat{\theta}$ from the equation:

$$\mathbf{0} = \Psi_n(\theta_0) - \tilde{\Psi}_n(\hat{\theta}) + \dot{\Psi}_n(\theta_0)(\tilde{\theta} - \hat{\theta}) + \dot{\Psi}_n(\theta_0)(\hat{\theta} - \theta_0) + \mathcal{O}_p(n^{-1})$$

The quantity $\dot{\Psi}_n(\theta_0)(\hat{\theta} - \theta_0)$ is expressible as:

$$\dot{\Psi}_n(\theta_0)(\hat{\theta} - \theta_0) = -\Psi_n(\theta_0) + \mathcal{O}_p(n^{-1})$$

Therefore:

$$\dot{\Psi}_n(\theta_0)(\tilde{\theta} - \hat{\theta}) = \tilde{\Psi}_n(\hat{\theta}) + \mathcal{O}_p(n^{-1})$$

Scaling by \sqrt{n} :

$$\sqrt{n}(\tilde{\theta} - \hat{\theta}) = \frac{1}{\sqrt{n}} \sum_{i=1}^n \omega_i \mathbf{A}_0^{-1} \psi_i(\hat{\theta})$$

Taking the limit conditional on $\hat{\theta}$:

$$\sqrt{n}(\tilde{\theta} - \hat{\theta}) \xrightarrow{\mathcal{L}} N(\mathbf{0}, \mathbf{A}_0^{-1} \mathbf{B}(\hat{\theta}) \mathbf{A}_0^{-T})$$

This argument motivates approximation of the distribution of $\sqrt{n}(\hat{\theta} - \theta_0)$ as:

$$\sqrt{n}(\hat{\theta} - \theta_0) \stackrel{d}{\approx} \sqrt{n}(\tilde{\theta} - \hat{\theta}) \stackrel{d}{\approx} \frac{1}{\sqrt{n}} \sum_{i=1}^n \omega_i \mathbf{A}_0^{-1} \psi_i(\hat{\theta})$$

Remark 2.1. Although $\tilde{\theta}$ was introduced as a solution to the equation $\Psi_n(\theta) - \tilde{\Psi}_n(\hat{\theta}) \stackrel{\text{Set}}{=} \mathbf{0}$, **actually obtaining $\tilde{\theta}$ is unnecessary.** ♦

Perturbation Distribution

3.1 Mammen Distribution

The Mammen two-point distribution takes the form:

$$\omega_i = \begin{cases} \frac{1-\sqrt{5}}{2} & , \text{ w.p. } \frac{1+\sqrt{5}}{2\sqrt{5}} \\ \frac{1+\sqrt{5}}{2} & , \text{ w.p. } 1 - \frac{1+\sqrt{5}}{2\sqrt{5}} \end{cases}$$

The Mammen distribution has expectation $E[\omega_i] = 0$, variance $E[\omega_i^2] = 1$, and third (central) moment $E[\omega_i^3] = 1$. Thus, Mammen perturbations are able to match this first three moments of a distribution, providing an asymptotic refinement over the normal approximation.

3.2 Rademacher Distribution

The Rademacher distribution is defined as:

$$\omega_i = \begin{cases} -1 & , \text{ w.p. } \frac{1}{2} \\ 1 & , \text{ w.p. } \frac{1}{2} \end{cases}$$

The Rademacher distribution has $E[\omega_i] = E[\omega_i^3] = 0$ and $E[\omega_i^2] = E[\omega_i^4] = 1$. Thus, Rademacher perturbations are able to match the first four moments of a **symmetric** distribution, providing an asymptotic refinement over the normal approximation.

Procedure

4.1 General Case

1. Obtain $\hat{\theta}$ satisfying $\Psi_n(\theta) = \mathbf{0}$.

i. Calculate the individual score contributions evaluated at $\hat{\theta}$:

$$\hat{\psi}_i = \psi_i(\hat{\theta})$$

ii. Estimate the Jacobian at $\hat{\theta}$:

$$\hat{A}_n = -\frac{1}{n} \sum_{i=1}^n \frac{\partial \psi_i}{\partial \theta'}(\hat{\theta})$$

iii. Calculate the normalized score:

$$\hat{U}_n = \frac{1}{\sqrt{n}} \sum_{i=1}^n \hat{A}_n^{-1} \hat{\psi}_i$$

iv. Estimate the outer product:

$$\hat{B}_n = \frac{1}{n} \sum_{i=1}^n \hat{\psi}_i \hat{\psi}_i'$$

v. Construct the observed test statistic:

$$T_{\text{obs}} = \hat{U}_n' (\hat{A}_n^{-1} \hat{B}_n \hat{A}_n^{-T})^{-1} \hat{U}_n$$

2. For $b = 1, \dots, B$:

i. Draw the perturbation weights $(\omega_i^{(b)})$.

ii. Generate the perturbed score:

$$\mathbf{U}_n^{(b)} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \omega_i^{(b)} \hat{A}_n^{-1} \hat{\psi}_i$$

iii. Estimate the perturbed outer product:

$$\mathbf{B}_n^{(b)} = \frac{1}{n} \sum_{i=1}^n (\omega_i^{(b)})^2 \hat{\psi}_i \hat{\psi}_i'$$

iv. Calculate the perturbed test statistic:

$$T^{(b)} = (\mathbf{U}_n^{(b)})' (\hat{A}_n^{-1} \mathbf{B}_n^{(b)} \hat{A}_n^{-T})^{-1} \mathbf{U}_n^{(b)}$$

3. Use the empirical distribution of $\{T^{(b)}\}$ to approximate that of T .

- To estimate of p -value for T_{obs} :

$$\hat{p} = \frac{1}{B} \sum_{b=1}^B I[T^{(b)} \geq T_{\text{obs}}]$$

- To construct a confidence interval, use the empirical $\alpha/2$ and $1 - \alpha/2$ percentiles of $\{T^{(b)}\}$.

4.2 Maximum Likelihood

1. Obtain the MLE $\hat{\boldsymbol{\theta}}$ satisfying $\mathbf{S}_n(\boldsymbol{\theta}) = \mathbf{0}$.

i. Calculate the individual score contributions evaluated at $\hat{\boldsymbol{\theta}}$, and the overall score:

$$\hat{\mathbf{s}}_i = \frac{\partial \ell_i}{\partial \boldsymbol{\theta}}(\hat{\boldsymbol{\theta}})$$

ii. Aggregate the individual score contributions:

$$\hat{\mathbf{S}}_n = \sum_{i=1}^n \hat{\mathbf{s}}_i$$

iii. Estimate the observed information at $\hat{\boldsymbol{\theta}}$:

$$\hat{\mathbf{J}}_n = - \sum_{i=1}^n \frac{\partial^2 \ell}{\partial \boldsymbol{\theta} \partial \boldsymbol{\theta}'}(\hat{\boldsymbol{\theta}})$$

iv. Calculate the test statistic:

$$T_{\text{obs}} = \hat{\mathbf{S}}_n' \hat{\mathbf{J}}_n^{-1} \hat{\mathbf{S}}_n$$

2. For $b = 1, \dots, B$:

i. Draw the perturbation weights $(\omega_i^{(b)})$.

ii. Generate the perturbed score:

$$\mathbf{S}_n^{(b)} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \omega_i^{(b)} \hat{\mathbf{s}}_i$$

iii. Calculate the perturbed test statistic:

$$T^{(b)} = (\hat{\mathbf{S}}_n^{(b)})' \hat{\mathbf{J}}_n^{-1} \hat{\mathbf{S}}_n^{(b)}$$

3. Use the empirical distribution of $\{T^{(b)}\}$ to approximate that of T .