

# HW 3

Ex. 2

$\Rightarrow:$

Assume  $\text{corr}(X, Y) = 1$ .

Let  $X' = \frac{X}{\sigma_X}$ ,  $Y' = \frac{Y}{\sigma_Y}$ . Then:

$$\begin{aligned}\text{Var}(X' - Y') &= \text{Var}(X') + \text{Var}(-Y') + 2 \text{Cov}(X', -Y') \\ &= \frac{1}{\sigma_X^2} \text{Var}(X) + \frac{1}{\sigma_Y^2} \text{Var}(Y) - \frac{2}{\sigma_X \sigma_Y} \text{Cov}(X, Y) \\ &= 2 - 2 \text{corr}(X, Y) \\ &= 0\end{aligned}$$

We have:

$$\text{Var}(X' - Y') = 0$$

$$\Leftrightarrow \mathbb{E}[(X' - Y')^2] = 0$$

$$\Leftrightarrow X' - Y' - \mathbb{E}[X'] + \mathbb{E}[Y'] = 0 \quad |P\text{-a.s.}$$

$$\Leftrightarrow X = \underbrace{\frac{\sigma_X}{\sigma_Y} Y}_{\geq 0} + \underbrace{\mathbb{E}[X] - \frac{\sigma_X}{\sigma_Y} \mathbb{E}[Y]}_{\in \mathbb{R}} \quad |P\text{-a.s.}$$

since  $\sigma_X, \sigma_Y > 0$

otherwise corr would not be well-defined

$\Leftarrow:$

Assume  $\exists a > 0, b \in \mathbb{R}$  st.  $X = aY + b$   $|P\text{-a.s.}$

Then we have  $|P\text{-a.s.}$ :

$$\begin{aligned}\text{corr}(X, Y) &= \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y} \\ &= \frac{\text{Cov}(aY + b, Y)}{\sigma_{aY+b} \sigma_Y} \quad (*)\end{aligned}$$

We have:

$$\begin{aligned}\text{Cov}(aY + b, Y) &= a \text{Cov}(Y, Y) \\ &= a \text{Var}(Y)\end{aligned}$$

$$\sigma_{aY+b} = \sqrt{\text{Var}(aY+b)}$$

$$= |a| \sqrt{\text{Var}(Y)}$$

$$= a_{\theta Y}$$

Thus:

$$(*) = \frac{a_{\theta Y} \text{Var}(Y)}{a_{\theta Y} a_Y} = 1$$

Ex. 3

Let  $\theta_0$  be the true parameter and  $\hat{\theta}_{MLE}$  the MLE of  $\theta_0 \in \Theta$ . Assume  $f: \Theta \rightarrow \tilde{\Theta}$  is a bijection. Then  $\exists \lambda_0 \in \tilde{\Theta}$  st.  $f(\theta_0) = \lambda_0$ .

We know by definition of the MLE that:

$$\hat{\theta}_{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} (L(\theta; X))$$

Let us denote  $\tilde{L}(\cdot, X)$  the likelihood function wrt.  $\lambda = f(\theta) \in \tilde{\Theta}$ , for  $\theta \in \Theta$ .

Since  $f$  is a bijection, we know that it is invertible. Thus, we can find the corresponding value  $\theta \in \Theta$  associated to  $\lambda \in \tilde{\Theta}$  st.:

$$\theta = f^{-1}(\lambda)$$

This allows us to write for  $f(\theta) = \lambda$ :

$$\tilde{L}(\lambda; X) = \tilde{L}(f(\theta); X) = L(f^{-1}(\lambda); X) = L(\theta; X)$$

By the definition of  $\hat{\theta}_{MLE}$ , we have  $\forall \theta \in \Theta$ :

$$L(\hat{\theta}_{MLE}; X) \geq L(\theta; X)$$

$$\Rightarrow \tilde{L}(f(\hat{\theta}_{MLE}); X) \geq \tilde{L}(\underbrace{f(\theta)}_{\in \tilde{\Theta}}; X)$$

$$\Rightarrow \tilde{L}(f(\hat{\theta}_{MLE}); X) \geq \tilde{L}(\lambda; X) \quad \forall \lambda \in \tilde{\Theta}$$

$$\Rightarrow f(\hat{\theta}_{MLE}) = \underset{\lambda \in \tilde{\Theta}}{\operatorname{argmax}} (\tilde{L}(\lambda; X))$$

$\Rightarrow f(\hat{\theta}_{MLE})$  is the MLE of  $\lambda_0 = f(\theta_0) \in \tilde{\Theta}$