

Descent Algorithm | Finding Local Minimum

Choose x_{start} . Descent algorithm generates a sequence $(x_k)_{k \in \mathbb{N}}$ defined by

$$x_{k+1} = x_k + s_k d_k$$

$$f(x_{k+1}) \leq f(x_k)$$

- d_k is the descent **direction**
- s_k is the descent **step** or **learning rate**

General Algorithm

Input: f differentiable on \mathbb{R}^d , x_{start} initial point, $\varepsilon > 0 \approx 0$ **precision**. *Output:* approximation of $\min_{x \in \mathbb{R}^d} f(x)$

Algorithm DA($f, x_{\text{start}}, \varepsilon$) with fixed learning rate:
 - Initialize $x_0 = x_{\text{start}}, k := 0$ - While $\|\nabla f(x_k)\| > \varepsilon$
 (**stop condition**) - Find d_k s.t. $\nabla f(x_k)^T d_k < 0$
 ($< \nabla f(x_k), d_k > < 0$) - Choose $s_k > 0$ in direction d_k s.t. $f(x_k + s_k d_k) < f(x_k)$ (**descent property**) -
 $x_{k+1} := x_k + s_k d_k$ (**update**) - $k := k + 1$ - Return x_k

We find d_k s.t. $< \nabla f(x_k), d_k > < 0$ because by Taylor:

$$\begin{aligned} f(x_{k+1}) &= f(x_k + s_k d_k) = \\ &= f(x_k) + s_{k>0} < \nabla f(x_k), d_k >_{<0} + \text{remainder} \\ &\Rightarrow f(x_{k+1}) \leq f(x_k) \end{aligned}$$

More Stop Conditions

$$\begin{aligned} \nabla f(x_k) \rightarrow \nabla f(x^*) = 0 &\Rightarrow \|\nabla f(x_k)\| \leq \varepsilon_1 \\ x_k \rightarrow x^* &\Rightarrow \|x_k - x_{k+1}\| \leq \varepsilon_2 \\ f(x_k) \rightarrow f(x^*) &\Rightarrow \|f(x_k) - f(x_{k+1})\| \leq \varepsilon_3 \end{aligned}$$

Gradient Algorithm

Take $d_k = -\nabla f(x_k)$
 $< \nabla f(x_k), -\nabla f(x_k) > = -\|\nabla f(x_k)\|^2 < 0$
 $x_{k+1} = x_k - s \cdot \nabla f(x_k)$

Input: f differentiable on \mathbb{R}^d , x_{start} initial point, $\varepsilon > 0 \approx 0$ **precision**. *Output:* approximation of $\min_{x \in \mathbb{R}^d} f(x)$

Algorithm GD(f, x_{start}) with fixed learning rate: -
 Initialize $x_0 = x_{\text{start}}, k := 0$ - While $\|\nabla f(x_k)\| > \varepsilon$
 (**stop condition**) - $x_{k+1} := x_k - s \nabla f(x_k)$ (**update**)
 - $k := k + 1$ - Return x_k

Note: choosing good starting point x_{start} and learning rate s_k is very important. If otherwise, the algorithm

might not converge if the learning rate too large. The algorithm might converge slowly, if learning rate is too small. If starting point is chosen wisely, the algorithm might even converge to a global minimum, not local.

Gradient Descent Convergence

Lipschitz Gradient

Let $f : \mathbb{R}^d \rightarrow \mathbb{R}$ differentiable function. If $\exists L \in \mathbb{R}_+^*$ s.t.

$$\frac{\|\nabla f(x) - \nabla f(y)\|}{\|x - y\|} \leq L, \quad x \neq y, \forall (x, y) \in \mathbb{R}^d$$

then f is differentiable with continuous L -Lipschitz gradient.

Let $f : \mathbb{R}^d \rightarrow \mathbb{R}$ differentiable function with L -Lipschitz continuous gradient. Then

$$f(y) \leq f(x) + \langle \nabla f(x), y - x \rangle + \frac{L}{2} \|y - x\|^2, \quad \forall (x, y) \in \mathbb{R}^d$$

Optimal Step

$$s^* = \operatorname{argmin}_{s>0} f(x - s \nabla f(x))$$

Convergence of GD

We have convergent GD $\iff \|\nabla f(x_k)\| < \varepsilon$ at some point x_k .

Let a function f bounded below, of class C1 with gradient Lipschitz of constant $L > 0$. Then:

1. If $\forall k, s_k = s < \frac{2}{L}$ the GD algorithm with constant step **converges globally** and

$$0 < s \|\nabla f(x_k)\|^2 - \frac{L s^2}{2} \|\nabla f(x_k)\|^2 \leq f(x_k) - f(x_{k+1})$$

2. The **optimal constant step** is defined by $s_k = s^* = \frac{1}{L} \forall k \in \mathbb{N}$ and

$$\frac{1}{2L} \|\nabla f(x_k)\|^2 \leq f(x_k) - f(x_{k+1})$$

Wolfe's Condition

The direction d and the step s of a descent algorithm with linear search satisfy Wolfe's conditions if there exists $(\varepsilon_1, \varepsilon_2) \in (0, 1)^2$ such that $\varepsilon_1 < \varepsilon_2$ and

1. $f(x + sd) \leq f(x) + \varepsilon_1 s (\nabla f(x)^T d)_{\text{linear}}$
2. $\nabla f(x + sd)^T d \geq \varepsilon_2 (\nabla f(x)^T d)$

The GD algorithm with Wolfe's step globally converges.

Newton's Method

By Taylor:

$$f(x) \approx f(x_k) + f'(x_k)(x - x_k) + \frac{f''(x_k)}{2}(x - x_k)^2 =: q(x)$$

Here, $f(x_k) + f'(x_k)(x - x_k)$ - linear approximation of the function in a neighborhood of $x + k$.

The necessary condition for optimization is $q'(x) = 0$, $f'(x) \approx q'(x)$.

$$q'(x) = f'(x_k) + f''(x_k)(x - x_k)$$

If $f''(x_k) > 0$, q achieves a minimum at

$$x = x_k - \frac{f'(x_k)}{f''(x_k)}$$

Newton's Algorithm

If f is smooth enough (atleast twice differentiable), then Newton's method is faster (quadratic convergence) than GD.

- input: f **twice differentiable** on R^d , x_{start} initial point, $\varepsilon > 0$
- output: approximation of $\min_{x \in R^d} f(x)$

Algorithm: 1. initialize $x_0 = x_{start}$, $k = 0$ 2. While $\|\nabla f(x_k)\| > \varepsilon$ (stop criterion) 1. $x_{k+1} = x_k - [Hf(x_k)]^{-1} \nabla f(x_k)$ 2. $k = k + 1$ 3. return x_k

Newton's direction is: $d_k = [Hf(x_k)]^{-1} \nabla f(x_k)$

Gauss-Newton Method

Approximate the Hessian, since it is and its inverse are hard to estimate.

$r : R^d \rightarrow R^d$, differentiable

$$f(x) = \frac{1}{2} \|r(x)\|_2^2$$

sum of squared residuals

Idea: replace at each iteration the non-linear LS problem by an approximate linear LS problem (by

Taylor). Let x_k be the current iterate. Replace in a neighbourhood of x_k , the problem by

$$\min_{y \in R^d} \tilde{f}(y) = \frac{1}{2} \|r(x_k) + J_r(x_k)(y - x_k)\|_2^2$$

$$y = x_{k+1}$$

$$x_{k+1} = (J_r^T(x_k) J_r(x_k))^{-1} (J_r^T(x_k) J_r(x_k) x_k - J_r^T(x_k) r(x_k))$$

$r(x_k)$ - residual vector

$$J_r^T(x_k) r(x_k) = \nabla f(x_k)$$

$$J_r^T(x_k) J_r(x_k) \approx Hf(x_k) - \text{approximation of Hessian}$$

The direction $d_k = x_{k+1} - x_k$ is called Gauss-Newton direction.

Gauss-Newton Algorithm

- input: r differentiable on R^d , x_{start} initial point, $\varepsilon > 0$ precision.
- output: approximation of $\min_{x \in R^d} f(x)$

Algorithm: 1. initialize $x_0 = x_{start}$, $k = 0$ 2. while $\|\nabla f(x_k)\| > \varepsilon$ 1. $x_{k+1} = x_k - [J_r(x_k)^T J_r(x_k)]^{-1} J_r(x_k)^T r(x_k)$ 2. $k = k + 1$ 3. return x_k

Pseudo-random Numbers Generation

Generalized Inverse

Let F be a non-decreasing function on R . The generalized inverse of F , denoted by F^- is the function defined by

$$F^-(u) = \inf\{x \in R : F(x) \geq u\}$$

The function F^- is non-decreasing, left-continuous and satisfies

$$F^-(u) \leq x \iff u \leq F(x), \forall u \in (0, 1)$$

If F is **increasing and continuous** on R , then F has an inverse function defined by F^{-1} s.t. $F * F^{-1} = Id_{(0,1)}$ and $F^{-1} * F = Id_R$

Proposition

Let $U \sim U((0, 1))$ and $F = \mu((-\infty, x])$ where μ is a probability distribution on $(R, B(R))$. Then $F^-(U) \sim \mu$.

In other words, $F^-(U)$ follows the law of X .

Example

We have that X follows an exponential distribution and want to simulate pseudo-random numbers.

$$f_X(x) = \lambda e^{-\lambda x} 1_{[0, +\infty]}(x)$$

$$F_X(x) = \int_{-\infty}^x \lambda e^{-\lambda x} 1_{[0, +\infty]}(x) dx$$

- if $t < 0 : F_X(t) = 0$
- if $t > 0 : F_X(t) = 1 - e^{-\lambda t}$

$$u \in (0, 1), u = F_X(t)$$

$$u = 1 - e^{-\lambda t}$$

$$\Rightarrow t = -\frac{1}{\lambda} \log(1 - u)$$

To simulate pseudo-random numbers:

$$u \sim U((0, 1))$$

$$F_x^-(u) = -\frac{1}{\lambda} \log(1 - u) \sim E(\lambda)$$

Acceptance-Rejection

Method

1. Sample Y with density g
2. Generate $U \sim U((0, 1))$ independently from Y
3. If $U \leq \frac{f(Y)}{cg(Y)}$ (accept with probability $\frac{f(Y)}{cg(Y)}$) (notice that it is bounded between 0 and 1)
 - $X = Y$ (acceptance)
 - else go back to step 1 (rejection)
4. Result: samples following target density (not $f!$).

Optimal

$$c = \sup_{x \in R^d} \frac{f(x)}{g(x)}$$

Algorithm

1. Generate $U_1 \sim g \sim \mathcal{U}([0, 1]), U_2 \sim \mathcal{U}([0, 1])$ independently ($U_1 = Y$).
2. Acceptance/rejection:
 - If $U_2 \leq \frac{f(U_1)}{cg(U_1)}$: Set $X = U_1$
 - Else: Go back to step 1

Example

Simulation of a Beta distribution.

$$\forall x \in (0, 1), f(x) = bx^n(1 - x)^n$$

$$b = \left[\int_0^1 x^n(1 - x)^n dx \right]^{-1} (b \cdot \int f(x) dx = 1)$$

$$\text{Let } g = 1_{(0,1)}$$

$$f(x) \leq cg(x)$$

Aim: choose optimal c .

$$f(x) \leq bg(x), c = b \text{ but not optimal}$$

$$c = \sup \frac{f(x)}{g(x)} = \sup_{x \in (0,1)} bx^n(1 - x)^n$$

$$f'(x) = nx^{n-1}(1 - x)^n - nx^n(1 - x)^{n-1} = bnx^{n-1}(1 - x)^{n-1}(1 - 2x)$$

$f(\frac{1}{2})$ - critical point, for $x < 1/2$, f increases, $x > 1/2$ f decreases.

\Rightarrow the maximum is reached at $x = 1/2$.

$$f(1/2) = b \frac{1}{2^n} (1 - 1/2)^n = b \frac{1}{4^n} = c$$

$$u_1 \sim U((0, 1)), u_2 \sim U((0, 1)) \text{ independently}$$

$$Y = u_1$$

$$\text{Acceptance/rejection: } - \text{ if } u_2 \leq \frac{f(u_1)}{cg(u_1)} =$$

$$\frac{bu_1^n(1-u_1)^n}{b} = 4^n u_1^n (1 - u_1)^n: \text{ set } X = u_1 - \text{ else go back to step 1}$$

Box-Müller Algorithm

Let X_1, X_2 independent r.v.. Denote by R, θ the polar coordinates of (X_1, X_2) .

$$R^2 = X_1^2 + X_2^2 \sim \mathcal{E}(1/2), \theta \sim \mathcal{U}([0, \pi])$$

If U_1 and U_2 are i.i.d. $\mathcal{U}([0, 1])$, the variables X_1 and X_2 defined by

$$X_1 = R \cos(\theta) = \sqrt{-2 \log(U_1)} \cos(2\pi U_2)$$

$$X_2 = R \sin(\theta) = \sqrt{-2 \log(U_1)} \sin(2\pi U_2)$$

are then i.i.d. of distribution $\mathcal{N}(0, 1)$.

Algorithm

1. Generate independently $U_1, U_2 \sim \mathcal{U}([0, 1])$.
2. Define $R_2 = \sqrt{-2 \log(U_1)}$ and $\theta = 2\pi U_2$.
3. Return $(X, Y) = (\sqrt{R_2} \cos(\theta), \sqrt{R_2} \sin(\theta))$

Markov Chains

Metropolis-Hastings Algorithm

This algorithm allows us to move closer to the area of high probability, while simple accept-reject method

did not do, since samples were independent of the last.

Goal: sample from $f \cdot \lambda$, but we know only f .

$f \cdot \lambda$, λ – normalizing constant that is difficult to compute. So, we only use $f(x)$ that is easy to compute, e.g. Normal.

Main idea: $X_0 \rightarrow X_1 \rightarrow \dots \rightarrow X_B \rightarrow X_{B+1} \rightarrow \dots$. Here, $X_0 \rightarrow X_1 \rightarrow \dots$ until some X_i are **burn-in**, they are typically not follow the distribution that we are trying to simulate, while starting from X_B are samples from $f \cdot \lambda$.

- Input: transition matrix P (conditional density $q(\cdot|\cdot)$)
- Output: HMC X (dependent sample distributed according to $f \cdot \lambda$).

1. Initialize $x_0 \in E$
2. For all $n \in \{0, \dots, N-1\}$: (at each n : given $X_n = x_n$)
 1. Generate $Y_n \sim q(\cdot|X_n) = q(\cdot|x_n)$ (Y_n – next state after x_n)
 2. Compute acceptance probability $A(a \rightarrow b) = p(x_n, Y_n) = \min\left(1, \frac{f(Y_n)q(x_n|Y_n)}{f(x_n)q(Y_n|x_n)}\right)$
 - here $x_n = a, Y_n = b$
 3. Draw $U \sim \mathcal{U}([0, 1])$
 4. If $U < p(x_n, Y_n)$
 - Set $X_{n+1} = Y_n$
 - Else set $X_{n+1} = x_n$
 5. $n = n + 1$

- The accept/reject step only depends on the ratios $\frac{f(Y_n)}{f(x_n)}, \frac{q(x_n|Y_n)}{q(Y_n|x_n)}$

Steps explained in simpler terms:

1. You are currently in some state X_t
2. Sample next candidate from some easier distribution g (e.g. Normal)
3. Should accept this candidate? We accept the candidate based on acceptance probability $A(a \rightarrow b) = p(x_n, Y_n) = \min\left(1, \frac{f(Y_n)q(x_n|Y_n)}{f(x_n)q(Y_n|x_n)}\right)$
4. If accept, then next state in the Markov chain is indeed $Y_n = b$
5. If do not accept, then the next state is previous state carried forward, i.e. stay where we at, take the same previous sample again, and then try again for the next sample.
6. Eventually it leads to $f \cdot \lambda$ being a stationary distribution for the Markov Chain.

7. So we computed $X \sim f \cdot \lambda$ just from f .

Random Walk Metropolis-Hastings

What this algorithm is aiming to achieve is that no matter where we are right now, no matter where we are proposed to go, we are going to accept that move if that is a more likely place to be (we want more samples from high density areas) and we might accept the move if it is low area density area (it is still important to sample from tails).

It is the same as previous algorithm, but we have that q is symmetric, meaning $q(x_n|Y_n)$ and $q(Y_n|x_n)$ cancel each other, thus the acceptance probability of moving from a to b becomes $\min\left(1, \frac{f(b)}{f(a)}\right)$. We also now know that $\frac{f(b)}{f(a)}$ is actually $\frac{f(b) \cdot \lambda}{f(a) \cdot \lambda}$, so it follow the distribution we want.

Assume that q is symmetric i.e. $q(x|y) = q(y|x), x, y \in E$. Take $q(x, y) = g(y - x)$ and $p(x_n, y) = \min(1, f(y)/f(x_n))$.

- Input: g density.
- Output: HMC X (dependent sample distributed according to $f \cdot \lambda$).

1. Initialize $x_0 \in E$
2. For all $n \in \{0, \dots, N-1\}$: (at each n : given $X_n = x_n$)
 1. Generate $Y_n \sim g(\cdot - x_n) \cdot \lambda$
 2. Compute $p(x_n, Y_n) = \min\left(1, \frac{f(Y_n)}{f(x_n)}\right)$
 3. Draw $U \sim \mathcal{U}([0, 1])$
 4. If $U < p(x_n, Y_n)$
 - Set $X_{n+1} = Y_n$
 - Else set $X_{n+1} = x_n$
 5. $n = n + 1$

This algorithm tells us: - If the probability of being at b is bigger than of being at a , then the acceptance probability of going from a to b is 1: $p(b) > p(a) \Rightarrow A(a \rightarrow b) = 1$ - Else, if probability of being at b is less than that of a , then the acceptance probability is: $p(b) < p(a) \Rightarrow A(a \rightarrow b) = \frac{p(b)}{p(a)}$ - Essentially, if are moving up the curve if the next simulated state is above the current state on the curve, while we are moving downwards if the next state if below on the curve of the current state. - The further the state b from a downwards, the less likely that we go from a to b .