

HW 3

Ex. 2

\Rightarrow :

Assume $\text{corr}(X, Y) = 1$.

Let $X' = \frac{X}{\sigma_X}$, $Y' = \frac{Y}{\sigma_Y}$. Then:

$$\begin{aligned}\text{Var}(X' - Y') &= \text{Var}(X') + \text{Var}(-Y') + 2\text{Cov}(X', -Y') \\ &= \frac{1}{\sigma_X^2} \text{Var}(X) + \frac{1}{\sigma_Y^2} \text{Var}(Y) - \frac{2}{\sigma_X \sigma_Y} \text{Cov}(X, Y) \\ &= 2 - 2 \text{corr}(X, Y) \\ &= 0\end{aligned}$$

We have:

$$\text{Var}(X' - Y') = 0$$

$$\Leftrightarrow \underbrace{E[(X' - Y' - E[X'] + E[Y'])^2]}_{\geq 0} = 0$$

$$\Leftrightarrow X' - Y' - E[X'] + E[Y'] = 0 \quad \text{IP-a.s.}$$

$$\Leftrightarrow X = \underbrace{\frac{\sigma_X}{\sigma_Y} Y}_{>0} + \underbrace{E[X] - \frac{\sigma_X}{\sigma_Y} E[Y]}_{\in \mathbb{R}} \quad \text{IP-a.s.}$$

since $\sigma_X, \sigma_Y > 0$
otherwise corr would
not be well-defined

\Leftarrow :

Assume $\exists a > 0, b \in \mathbb{R}$ s.t. $X = aY + b$ IP-a.s.

Then we have IP-a.s.:

$$\begin{aligned}\text{corr}(X, Y) &= \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y} \\ &= \frac{\text{Cov}(aY + b, Y)}{b a Y + b \sigma_Y} \quad (*)\end{aligned}$$

We have:

$$\begin{aligned}\text{Cov}(aY + b, Y) &= a \text{Cov}(Y, Y) \\ &= a \text{Var}(Y)\end{aligned}$$

$$\begin{aligned}b a Y + b \sigma_Y &= \sqrt{\text{Var}(aY + b)} \\ &= |a| \sqrt{\text{Var}(Y)}\end{aligned}$$

$$= a \sigma_Y$$

Thus:

$$(*) = \frac{a \text{Var}(Y)}{a \sigma_Y \sigma_Y} = 1$$

Ex. 3

Let θ_0 be the true parameter and $\hat{\theta}_{MLE}$ the MLE of $\theta_0 \in \Theta$. Assume $f: \Theta \rightarrow \tilde{\Theta}$ is a bijection. Then $\exists \lambda_0 \in \tilde{\Theta}$ st.: $f(\theta_0) = \lambda_0$.

We know by definition of the MLE that:

$$\hat{\theta}_{MLE} = \underset{\theta \in \Theta}{\operatorname{argmax}} (L(\theta; X))$$

Let us denote $\tilde{L}(\cdot, X)$ the likelihood function wrt. $\lambda = f(\theta) \in \tilde{\Theta}$, for $\theta \in \Theta$.

Since f is a bijection, we know that it is invertible. Thus, we can find the corresponding value $\theta \in \Theta$ associated to $\lambda \in \tilde{\Theta}$ st.:

$$\theta = f^{-1}(\lambda)$$

This allows us to write for $f(\theta) = \lambda$:

$$\tilde{L}(\lambda; X) = \tilde{L}(f(\theta); X) = L(f^{-1}(\lambda); X) = L(\theta; X)$$

By the definition of $\hat{\theta}_{MLE}$, we have $\forall \theta \in \Theta$:

$$L(\hat{\theta}_{MLE}; X) \geq L(\theta; X)$$

$$\Rightarrow \tilde{L}(f(\hat{\theta}_{MLE}); X) \geq \tilde{L}(\underbrace{f(\theta)}_{\in \tilde{\Theta}}; X)$$

$$\Rightarrow \tilde{L}(f(\hat{\theta}_{MLE}); X) \geq \tilde{L}(\lambda; X) \quad \forall \lambda \in \tilde{\Theta}$$

$$\Rightarrow f(\hat{\theta}_{MLE}) = \underset{\lambda \in \tilde{\Theta}}{\operatorname{argmax}} (\tilde{L}(\lambda; X))$$

$$\Rightarrow f(\hat{\theta}_{MLE}) \text{ is the MLE of } \lambda_0 = f(\theta_0) \in \tilde{\Theta}$$