

given simple random sample

$\{X_1, \dots, X_n\}$ from F_θ

We learned how to construct point estimator $\hat{\theta}$ of θ
interval estimator (or confidence interval (CI)) of θ
 $[\hat{\theta}_L, \hat{\theta}_U]$

Symmetric $[\hat{\theta} - Mx SE_{\hat{\theta}}, \hat{\theta} + Mx SE_{\hat{\theta}}$

need to look at sampling distribution of $\hat{\theta}$
(distribution of $\hat{\theta}$) not easy

Normally you can simplify

Simulation can always be done

3 fundamental sampling distributions

$Z_1, \dots, Z_n \text{ iid } \sim N(0, 1)$

$X \equiv Z_1^2 + \dots + Z_n^2$ is χ^2 with df n ($X \sim \chi_n^2$)

$Z \sim N(0, 1), X \sim \chi_n^2$

$Z \perp\!\!\!\perp X$

Then $T \equiv \frac{Z}{\sqrt{X/n}}$ is called t distribution with df n ($T \sim t_n$)
if $X_1 \sim \chi_{n_1}^2, X_2 \sim \chi_{n_2}^2$

$X_1 \perp\!\!\!\perp X_2$

$F \equiv \frac{X_1/n_1}{X_2/n_2}$ is called F distribution with df n_1 and n_2 ($F \sim F_{n_1, n_2}$)

$X_1, \dots, X_n \text{ iid } \sim N(\mu, \sigma^2)$

$\bar{X} \sim N(\mu, \frac{\sigma^2}{n})$

$\frac{(n-1)s^2}{\sigma^2} \sim \chi_{n-1}^2$

$s^2 \equiv \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$

$\bar{X} \perp\!\!\!\perp s^2$

$\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \sim N(0, 1)$

$\frac{\sqrt{n}(\bar{X} - \mu)}{s} \sim t_{n-1}$

$$Y_1, \dots, Y_n \text{ iid } N(\mu, \sigma^2)$$

$$\{x_i\}_{i=1}^n \perp \{y_i\}_{i=1}^n$$

$$\text{then } \frac{S_x^2}{S_y^2} \sim F_{n-1, n-1}$$

CLT

$$X_1, \dots, X_n \text{ iid } \sim (\mu, \sigma^2)$$

not necessarily normal

$$\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma} \rightarrow N(0, 1) \text{ in distribution as } n \rightarrow \infty$$

$$X_1, \dots, X_n \text{ iid Bern}(p)$$

$$\hat{p} = \bar{X} \quad p(\text{of interest})$$

$$\frac{\sqrt{n}(\hat{p} - p)}{\sqrt{p(1-p)}} \sim N(0, 1)$$

$$\frac{\sqrt{n}(\hat{p} - p)}{\hat{p}(1-\hat{p})} \sim N(0, 1)$$

$1-\alpha$ is confidence level

$(100(1-\alpha)\% \text{ CI})$

$$X_1, \dots, X_n \text{ iid } N(\mu, \sigma_0^2)$$

σ_0^2 is known, μ is unknown

Find $100(1-\alpha)\% \text{ CI for } \mu$

$$\frac{\sqrt{n}(\bar{X} - \mu)}{\sigma_0} \sim N(0, 1) \quad \text{pivot}$$

$$P(-z_{\alpha/2} \leq \frac{\sqrt{n}(\bar{X} - \mu)}{\sigma_0} \leq z_{\alpha/2}) = 1-\alpha$$

$$\bar{X} + z_{\alpha/2} \frac{\sigma_0}{\sqrt{n}} \geq \mu$$

$$\bar{X} - z_{\alpha/2} \frac{\sigma_0}{\sqrt{n}} \leq \mu$$

$$\bar{X} - z_{\alpha/2} \cdot \frac{\sigma_0}{\sqrt{n}} \leq \mu \leq \bar{X} + z_{\alpha/2} \cdot \frac{\sigma_0}{\sqrt{n}}$$

multiplier
standard error

$$t\text{-interval } \mu \quad X_1, \dots, X_n \sim \text{iid } N(\mu, \sigma^2)$$

μ, σ^2 unknown

$$(\bar{X} - t_{n-1, \alpha/2} \cdot \frac{s}{\sqrt{n}}, \bar{X} + t_{n-1, \alpha/2} \cdot \frac{s}{\sqrt{n}})$$

$$X_1, \dots, X_n \text{ Bernoulli}(p) \quad \text{CI for } p$$

$$\hat{p} - z_{\alpha/2} \cdot \frac{\sqrt{\hat{p}(1-\hat{p})}}{\sqrt{n}}, \hat{p} + z_{\alpha/2} \cdot \frac{\sqrt{\hat{p}(1-\hat{p})}}{\sqrt{n}}$$

non symmetric example

X_1, \dots, X_n iid $N(\mu, \sigma^2)$ find 100(1- α)% CI for σ^2

$$\frac{(n-1)s^2}{\sigma^2} \sim \chi^2_{n-1}$$

$$P\left(\chi^2_{n-1, 1-\frac{\alpha}{2}} \leq \frac{(n-1)s^2}{\sigma^2} \leq \chi^2_{n-1, \frac{\alpha}{2}}\right) = 1-\alpha$$

$$\left[\frac{(n-1)s^2}{\chi^2_{n-1, \frac{\alpha}{2}}}, \frac{(n-1)s^2}{\chi^2_{n-1, 1-\frac{\alpha}{2}}} \right]$$

model (population)

$\{F_\theta : \theta \in \Theta\}$ Θ is the parameter space, $\theta_0 \in \Theta$

$$X_1, \dots, X_n \sim F_\theta$$

hypothesis, for example

$$\theta_0 \in \Theta, \theta_1 \in \Theta$$

null hypothesis H_0 alternative hypothesis H_a

complements of each other

For example

$$H_0 : \theta = \theta_0$$

$$H_a : \theta \neq \theta_0$$

$$H_0 : \theta \leq \theta_0 \text{ versus } H_a : \theta > \theta_0$$

$$H_0 : \theta \geq \theta_0 \text{ versus } H_a : \theta < \theta_0$$

$$H_0 : \bar{x} = 0 \text{ vs } H_a : \bar{x} \neq 0 \quad \leftarrow \text{this is wrong, no hypothesis, just fact}$$

Hypothesis tests are based on samples: so there is uncertainty involved
never say "H₀ is proved true", only say "H₀ is not rejected at significance level α "

$$H_0 : \theta = 0 \text{ vs } H_a : \theta = 1 \quad (1) \quad \text{reject } H_0 \text{ in } (1)$$

$$H_0 : \theta = 1 \text{ vs } H_a : \theta = 0 \quad (2) \quad \text{accept } H_0 \text{ in } (2)$$

not put sum

confidence in these two

H_0 & H_a not treated equally

conclusions are not the same

H_0 is favored unless evidence against it is very strong

H_0 is not rejected