Statistics: Homework 4

1. (a) Given p_i and q_i denote the probability of choosing box 1 and 2 respectively if the ball color chosen is i where $i = \{B, W, G\}$, denoting the three different colors. With the given information of the number of different color balls in the different boxes,

$$\begin{array}{lll} \mathbb{P}(B|1) = 4/10 & \mathbb{P}(B|2) = 3/10 & \mathbb{P}(B|3) = 2/10 \\ \mathbb{P}(W|1) = 2/10 & \mathbb{P}(W|2) = 6/10 & \mathbb{P}(W|3) = 0 \\ \mathbb{P}(G|1) = 4/10 & \mathbb{P}(G|2) = 1/10 & \mathbb{P}(G|3) = 8/10 \end{array}$$

The risk function is represented by,

$$R(\theta, \hat{\theta}_{p,q}) = \mathbb{E}_{\theta}(|\theta^2 - \hat{\theta}_{p,q}|^2)$$
$$= \mathbb{E}_{\theta}\left(\sum_{i \in \{B,W,G\}} L(\theta, \hat{\theta}_{p,q}(i)) \mathbb{P}(i|\theta)\right)$$

where $L(\theta, \hat{\theta}_{p,q}(i)) = L(\theta, 1)p_i + L(\theta, 2)q_i + L(\theta, 3)(1 - p_i - q_i)$. Therefore,

$$\begin{split} R(1,\hat{\theta}_{p,q}) &= [q_B + 4(1-p_B-q_B)]\frac{4}{10} + [q_W + 4(1-p_W-q_W)]\frac{2}{10} + [q_G + 4(1-p_G-q_G)]\frac{4}{10} \\ R(2,\hat{\theta}_{p,q}) &= [9p_B + 4q_B + 49(1-p_B-q_B)]\frac{3}{10} + [9p_W + 4q_W + 49(1-p_W-q_W)]\frac{6}{10} \\ &+ [9p_G + 4q_G + 49(1-p_G-q_G)]\frac{1}{10} \end{split}$$

(b) Bayes risk is given by

$$r(f,\theta) = \int R(\theta, \hat{\theta}_{p,q}) f(\theta) d\theta$$

but since our scenario is discrete, we instead have

$$\begin{split} r(f,\theta) &= \sum_{\theta=1,2} R(\theta,\hat{\theta}_{p,q}) \mathbb{P}(\theta) \\ &= \lambda R(1,\hat{\theta}_{p,q}) + (1-\lambda) R(2,\hat{\theta}_{p,q}) \end{split}$$

where $R(1, \hat{\theta}_{p,q})$ and $R(2, \hat{\theta}_{p,q})$ are the values are from (a).

(c) Given $\lambda = 1/2$, we have

$$r(f,\theta) = \frac{1}{2} \left(R(1,\hat{\theta}_{p,q}) + R(2,\hat{\theta}_{p,q}) \right) = \frac{1}{2} \left(16.3 - 13.6p_B - 14.7q_B - 24.8p_W - 27.6q_W - 5.6p_G - 5.6q_G \right)$$

thus to the infimum of Bayes risk is when $q_B = q_W = q_G = 1$.

library(leaps)

Read csv file into dataframe car
car <- read.csv('carmpgdat.csv')</pre>

3.
library(dplyr)
library(magrittr)

Reading data with separator tab
raw_riasec <- read.csv('RIASEC.csv',sep = '\t')</pre>

```
# (a) CLEANING UP
   # List of realistic traits
   realistic_trait <- c('R1','R2','R3','R4','R5','R6','R7','R8')</pre>
   # Extracting out the realistic traits
   raw_realistic <- raw_riasec[,realistic_trait]</pre>
   # Removing rows with -1 from the dataframe
   realistic <- raw_realistic %>%
    filter(R1* R2 * R3 * R4 * R5 * R6 * R7 * R8 > 0)
   # (b) MODEL SELECTION
   # Computing the score for the R trait
  realistic <- realistic %>%
    rowwise() %>%
    mutate(Rscore = mean(c(R1, R2, R3, R4, R5, R6, R7, R8)))
   # Generating the training and validation set
   tr_realistic <- realistic[1:6500,]</pre>
   val_realistic <- realistic[-(1:6500),]</pre>
   # Building the linear model
  lm_riasec <- lm(Rscore~R1, data = tr_realistic)</pre>
   avg_RSS_tr = mean(lm_riasec$residuals^2)
   # estimated regression function and residual sum of squares
  print(lm$coefficients)
   # R1 = 1.0011862 + 0.4282061 * R1
   # (c)VALIDATION
   reg.fn <- function(x) 1.0011862 + 0.4282061 * x
   val_realistic %<>%
    mutate(pred_Rscore = reg.fn(R1),
           residuals = reg.fn(R1) - Rscore )
   avg_RSS_val = mean(val_realistic$residuals^2)
   print (avg_RSS_tr) # 0.4540176
  print (avg_RSS_val) # 0.5376852
   # The residual sum of squares for the validation set using the regression function
   # is larger than the residual sum of square for the training set but are of the
   # same order. Thus the model generalizes well.
4. (a) -
      # (a) Using Monte Carlo method with N samples
      N <- 100000
      I \leftarrow function(u) \{X = rgamma(1.5, 1/2.3, n = N); return(sum(1 < X & X < 2)/N)\}
      I_hat <- I
       # I_hat = 0.20146
       # Estimated standard error by resampling Monte Carlo 10000 times.
      se_mc <- sqrt(var(sapply(1:N,I)))</pre>
       \# se_mc = 0.001272072
      # Estimated standard error using 10000 bootstrap samples.
      B <- 10000
```

```
 base\_sample \leftarrow rgamma(1.5,1/2.3,n = N) \\ I\_bootstrap \leftarrow function(X)\{X = sample(base\_sample, N, replace = TRUE); return(sum(1<X & X<2)/N)\} \\ se\_bootstrap \leftarrow sqrt(var(sapply(1:B,I\_bootstrap)))
```

se_bootstrap = 0.001273952

(b) We see that we can rewrite I as the following,

$$I = \int_{1}^{2} f(x|1.5, 2.3) dx = \int_{0}^{\infty} h(x) f(x|1.5, 2.3) dx$$

where $h(x) = \mathbb{1}(1 \le x \le 2)$, an indicator function that is 1 when $1 \le x \le 2$ and 0 otherwise. Thus we can use the following estimator \hat{I} for I where $X_i \sim \text{Gamma}(1.5, 2.3)$,

$$\hat{I} = \frac{1}{N} \sum_{i=1}^{N} h(X_i)$$

Hence by viewing $h(X) \sim \text{Bernoulli}(p)$, where $p = \mathbb{P}(1 \leq X \leq 2)$ for $X \sim \text{Gamma}(1.5, 2.3)$ the standard error is given by

$$Var(\hat{I}) = \frac{p(1-p)}{N}$$

where we sourced for the value of p from the website thus getting

$$\mathbb{P}(1 \le X \le 2) = \int_0^2 f(x|1.5, 2.3) \, dx - \int_0^1 f(x|1.5, 2.3) \, dx = 0.37173 - 0.16723$$

Evaluating the value of the standard error we get se = 0.00127545972