Repeated Measures ANOVA in R

The **repeated-measures ANOVA** is used for analyzing data where same subjects are measured more than once. This test is also referred to as a *within-subjects ANOVA* or *ANOVA with repeated measures*. The "within-subjects" term means that the same individuals are measured on the same outcome variable under different time points or conditions.

For example, you might have measured 10 individuals' self-esteem score (the outcome or dependent variable) on three time points during a specific diet to determine whether their self-esteem improved.

This chapter describes the different types of repeated measures ANOVA, including:

- **One-way repeated measures ANOVA**, an extension of the paired-samples t-test for comparing the means of three or more levels of a *within-subjects* variable.
- **two-way repeated measures ANOVA** used to evaluate simultaneously the effect of two within-subject factors on a continuous outcome variable.
- **three-way repeated measures ANOVA** used to evaluate simultaneously the effect of three within-subject factors on a continuous outcome variable.

The main goal of two-way and three-way repeated measures ANOVA is, respectively, to evaluate if there is a statistically significant interaction effect between two and three within-subjects factors in explaining a continuous outcome variable.

You will learn how to:

- Compute and interpret the different repeated measures ANOVA in R.
- Check repeated measures ANOVA test assumptions
- **Perform post-hoc tests**, multiple pairwise comparisons between groups to identify which groups are different
- **Visualize the data** using box plots, add ANOVA and pairwise comparisons p-values to the plot

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Related Book

Practical Statistics in R II - Comparing Groups: Numerical Variables

Assumptions

The repeated measures ANOVA makes the following assumptions about the data:

- **No significant outliers** in any cell of the design. This can be checked by visualizing the data using box plot methods and by using the function <code>identify_outliers()</code> [rstatix package].
- Normality: the outcome (or dependent) variable should be approximately normally distributed in each cell of the design. This can be checked using the Shapiro-Wilk normality test (shapiro_test() [rstatix]) or by visual inspection using QQ plot (ggqqplot() [ggpubr package]).
- **Assumption of sphericity**: the variance of the differences between groups should be equal. This can be checked using the **Mauchly's test of sphericity**, which is automatically reported when using the R function anova_test() [rstatix package]. Read more in Chapter @ref(mauchly-s-test-of-sphericity-in-r).

Before computing repeated measures ANOVA test, you need to perform some preliminary tests to check if the assumptions are met.

Note that, if the above assumptions are not met there are a non-parametric alternative (*Friedman test*) to the one-way repeated measures ANOVA.

Unfortunately, there are no non-parametric alternatives to the two-way and the three-way repeated measures ANOVA. Thus, in the situation where the assumptions are not met, you could consider running the two-way/three-way repeated measures ANOVA on the transformed and non-transformed data to see if there are any meaningful differences.

If both tests lead you to the same conclusions, you might not choose to transform the outcome variable and carry on with the two-way/three-way repeated measures ANOVA on the original data.

It's also possible to perform robust ANOVA test using the **WRS2** R package.

No matter your choice, you should report what you did in your results.

Prerequisites

Make sure that you have installed the following R packages:

- tidyverse for data manipulation and visualization
- ggpubr for creating easily publication ready plots
- rstatix provides pipe-friendly R functions for easy statistical analyses
- datarium: contains required data sets for this chapter

Start by loading the following R packages:

```
1 library(tidyverse)
2 library(ggpubr)
3 library(rstatix)
```

Key R functions:

```
• 1 anova_test()
```

[rstatix package], a wrapper around

```
1 | car::Anova()
```

for making easy the computation of repeated measures ANOVA. Key arguments for performing repeated measures ANOVA:

- o data: data frame
- o dv: (numeric) the dependent (or outcome) variable name.
- wid: variable name specifying the case/sample identifier.
- within: within-subjects factor or grouping variable
- get_anova_table() [rstatix package]. Extracts the ANOVA table from the output of anova_test(). It returns ANOVA table that is automatically corrected for eventual deviation from the sphericity assumption. The default is to apply automatically the Greenhouse-Geisser sphericity correction to only within-subject factors violating the sphericity assumption (i.e., Mauchly's test p-value is significant, p <= 0.05). Read more in Chapter @ref(mauchly-s-test-of-sphericity-in-r).</p>

One-way repeated measures ANOVA

Data preparation

We'll use the self-esteem score dataset measured over three time points. The data is available in the datarium package.

```
# Data preparation
# Wide format
data("selfesteem", package = "datarium")
head(selfesteem, 3)
```

```
5 ## # A tibble: 3 x 4
 6
   ##
          id t1 t2 t3
 7
   ##
        <int> <dbl> <dbl> <dbl>
   ## 1
          1 4.01 5.18 7.11
8
   ## 2
           2 2.56 6.91 6.31
9
10 ## 3
          3 3.24 4.44 9.78
11 # Gather columns t1, t2 and t3 into long format
12 # Convert id and time into factor variables
   selfesteem <- selfesteem %>%
13
14
    gather(key = "time", value = "score", t1, t2, t3) %>%
    convert_as_factor(id, time)
15
16 head(selfesteem, 3)
   ## # A tibble: 3 x 3
17
18 ## id time score
19 ## <fct> <fct> <dbl>
20 ## 1 1 t1 4.01
21 ## 2 2 t1
                   2.56
22 ## 3 3 t1
                  3.24
```

The one-way repeated measures ANOVA can be used to determine whether the means selfesteem scores are significantly different between the three time points.

Summary statistics

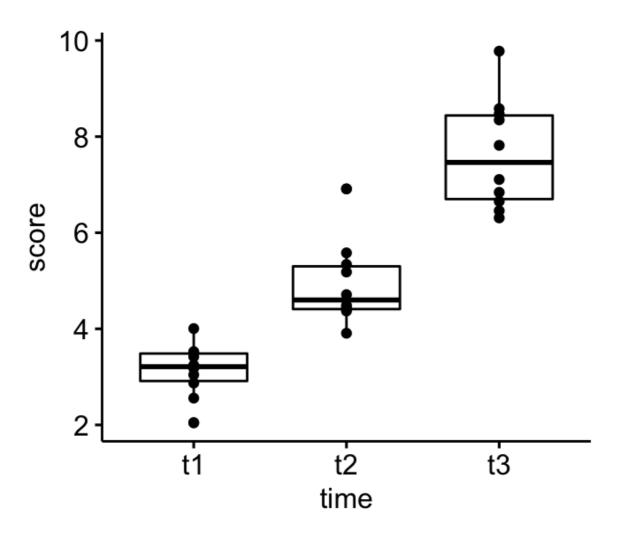
Compute some summary statistics of the self-esteem score by groups (time): mean and sd (standard deviation)

```
1 selfesteem %>%
2
   group_by(time) %>%
3
    get_summary_stats(score, type = "mean_sd")
  ## # A tibble: 3 x 5
4
                      n mean
5
  ## time variable
  ## <fct> <chr> <dbl> <dbl> <dbl> <
6
  ## 1 t1 score 10 3.14 0.552
7
  ## 2 t2 score
                     10 4.93 0.863
8
  ## 3 t3 score
                     10 7.64 1.14
```

Visualization

Create a box plot and add points corresponding to individual values:

```
bxp <- ggboxplot(selfesteem, x = "time", y = "score", add = "point")
bxp</pre>
```



Check assumptions

Outliers

Outliers can be easily identified using box plot methods, implemented in the R function identify_outliers() [rstatix package].

```
selfesteem %>%
2
    group_by(time) %>%
3
     identify_outliers(score)
  ## # A tibble: 2 x 5
5
        time id score is.outlier is.extreme
       <fct> <fct> <dbl> <lgl>
6
                                  <1g1>
                    2.05 TRUE
7
   ## 1 t1
             6
                                    FALSE
   ## 2 t2
                    6.91 TRUE
                                    FALSE
```

There were no extreme outliers.

Note that, in the situation where you have extreme outliers, this can be due to: 1) data entry errors, measurement errors or unusual values.

You can include the outlier in the analysis anyway if you do not believe the result will be substantially affected. This can be evaluated by comparing the result of the ANOVA with and without the outlier.

It's also possible to keep the outliers in the data and perform robust ANOVA test using the WRS2 package.

Normality assumption

The normality assumption can be checked by computing Shapiro-Wilk test for each time point. If the data is normally distributed, the p-value should be greater than 0.05.

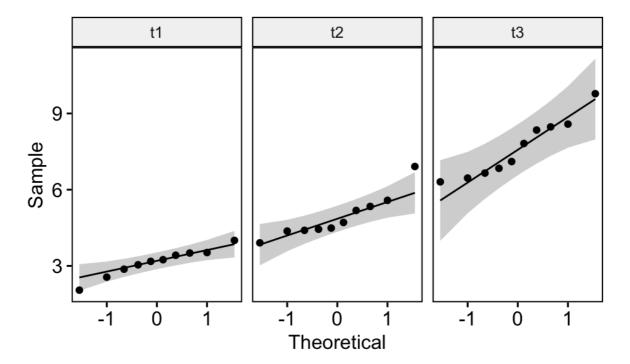
```
selfesteem %>%
1
2
    group_by(time) %>%
3
    shapiro_test(score)
  ## # A tibble: 3 x 4
4
5
       time variable statistic
       <fct> <chr>
6
  ##
                         <dbl> <dbl>
7
  ## 1 t1 score
                         0.967 0.859
8
  ## 2 t2 score
                         0.876 0.117
  ## 3 t3 score
                         0.923 0.380
```

The self-esteem score was normally distributed at each time point, as assessed by Shapiro-Wilk's test (p > 0.05).

Note that, if your sample size is greater than 50, the normal QQ plot is preferred because at larger sample sizes the Shapiro-Wilk test becomes very sensitive even to a minor deviation from normality.

QQ plot draws the correlation between a given data and the normal distribution. Create QQ plots for each time point:

```
1 | ggqqplot(selfesteem, "score", facet.by = "time")
```



From the plot above, as all the points fall approximately along the reference line, we can assume normality.

Assumption of sphericity

As mentioned in previous sections, the assumption of sphericity will be automatically checked during the computation of the ANOVA test using the R function [anova_test()] [rstatix package]. The Mauchly's test is internally used to assess the sphericity assumption.

By using the function <code>get_anova_table()</code> [rstatix] to extract the ANOVA table, the Greenhouse-Geisser sphericity correction is automatically applied to factors violating the sphericity assumption.

Computation

```
res.aov <- anova_test(data = selfesteem, dv = score, wid = id, within = time)
get_anova_table(res.aov)
## ANOVA Table (type III tests)
##
##
##
##
## Effect DFn DFd F p p<.05 ges
## 1 time 2 18 55.5 2.01e-08 * 0.829
```

The self-esteem score was statistically significantly different at the different time points during the diet, F(2, 18) = 55.5, p < 0.0001, eta2[g] = 0.83.

where,

- F Indicates that we are comparing to an F-distribution (F-test); (2, 18) indicates the degrees of freedom in the numerator (DFn) and the denominator (DFd), respectively; 55.5 indicates the obtained F-statistic value
- p specifies the p-value
- ges is the generalized effect size (amount of variability due to the within-subjects factor)

Post-hoc tests

You can perform multiple **pairwise paired t-tests** between the levels of the within-subjects factor (here time). P-values are adjusted using the Bonferroni multiple testing correction method.

```
1 # pairwise comparisons
pwc <- selfesteem %>%
   pairwise_t_test(
    score ~ time, paired = TRUE,
     p.adjust.method = "bonferroni"
5
6
7 pwc
8 | ## # A tibble: 3 x 10
9 | ## .y. group1 group2 n1 n2 statistic df
                                                   p p.adj
   p.adj.signif
10 | ## * <chr> <chr> <int> <int> <dbl> <dbl> <dbl> <dbl>
   <chr>>
11 ## 1 score t1 t2
                       10
                              10 -4.97 9 0.000772 0.002 **
12 ## 2 score t1 t3 10
                              10 -13.2 9 0.000000334 0.000001
   ****
  ## 3 score t2 t3
                         10
                              10 -4.87
                                            9 0.000886
                                                       0.003
                                                              **
```

All the pairwise differences are statistically significant.

Report

We could report the result as follow:

The self-esteem score was statistically significantly different at the different time points, F(2, 18) = 55.5, p < 0.0001, generalized eta squared = 0.82.

Post-hoc analyses with a Bonferroni adjustment revealed that all the pairwise differences, between time points, were statistically significantly different ($p \le 0.05$).

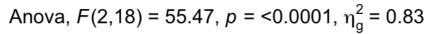
```
# visualization: box plots with p-values
pwc <- pwc %>% add_xy_position(x = "time")

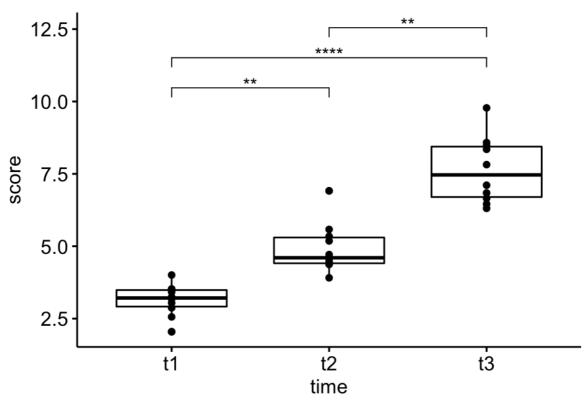
bxp +

stat_pvalue_manual(pwc) +

labs(
subtitle = get_test_label(res.aov, detailed = TRUE),
caption = get_pwc_label(pwc)

)
```





pwc: T test; p.adjust: Bonferroni

Two-way repeated measures ANOVA

Data preparation

We'll use the selfesteem2 dataset [in datarium package] containing the self-esteem score measures of 12 individuals enrolled in 2 successive short-term trials (4 weeks): control (placebo) and special diet trials.

Each participant performed all two trials. The order of the trials was counterbalanced and sufficient time was allowed between trials to allow any effects of previous trials to have dissipated.

The self-esteem score was recorded at three time points: at the beginning (t1), midway (t2) and at the end (t3) of the trials.

The question is to investigate if this short-term diet treatment can induce a significant increase of self-esteem score over time. In other terms, we wish to know if there is significant interaction between diet and time on the self-esteem score.

The two-way repeated measures ANOVA can be performed in order to determine whether there is a significant interaction between diet and time on the self-esteem score.

Load and show one random row by treatment group:

```
1 # Wide format
 2 set.seed(123)
 data("selfesteem2", package = "datarium")
 4 | selfesteem2 %>% sample_n_by(treatment, size = 1)
 5 ## # A tibble: 2 x 5
 6 ## id treatment t1 t2 t3
 8 ## 1 4 ctr 92 92 89
9 ## 2 10 Diet 90 93 95
10 | # Gather the columns t1, t2 and t3 into long format.
11 # Convert id and time into factor variables
12
    selfesteem2 <- selfesteem2 %>%
    gather(key = "time", value = "score", t1, t2, t3) %>%
13
14
     convert_as_factor(id, time)
15 # Inspect some random rows of the data by groups
16 set.seed(123)
| selfesteem2 %>% sample_n_by(treatment, time, size = 1)
18 ## # A tibble: 6 x 4
19 ## id treatment time score
20 | ## <fct> <fct> <fct> <dbl>
                      t1 92
21 ## 1 4 ctr
22 ## 1 4 Ctr t1
22 ## 2 10 ctr t2
23 ## 3 5 ctr t3
24 ## 4 11 Diet t1
                             84
                              68
                               93
25 ## 5 12 Diet t2
26 ## 6 1 Diet t3
                                80
                                88
```

In this example, the effect of "time" on self-esteem score is our **focal variable**, that is our primary concern.

However, it is thought that the effect "time" will be different if treatment is performed or not. In this setting, the "treatment" variable is considered as **moderator variable**.

Summary statistics

Group the data by treatment and time, and then compute some summary statistics of the score variable: mean and sd (standard deviation).

```
selfesteem2 %>%
     group_by(treatment, time) %>%
3
     get_summary_stats(score, type = "mean_sd")
   ## # A tibble: 6 x 6
      treatment time variable
                                n mean
       <fct>
                <fct> <chr>
                             <dbl> <dbl> <dbl> <dbl>
   ## 1 ctr
               t1 score
                              12 88
                                         8.08
                t2
   ## 2 ctr
                               12 83.8 10.2
                     score
                               12 78.7 10.5
9
   ## 3 ctr
               t3 score
                t1 score
10
   ## 4 Diet
                              12 87.6 7.62
   ## 5 Diet
              t2 score
                              12 87.8 7.42
11
            t3
   ## 6 Diet
                                12 87.7 8.14
                     score
```

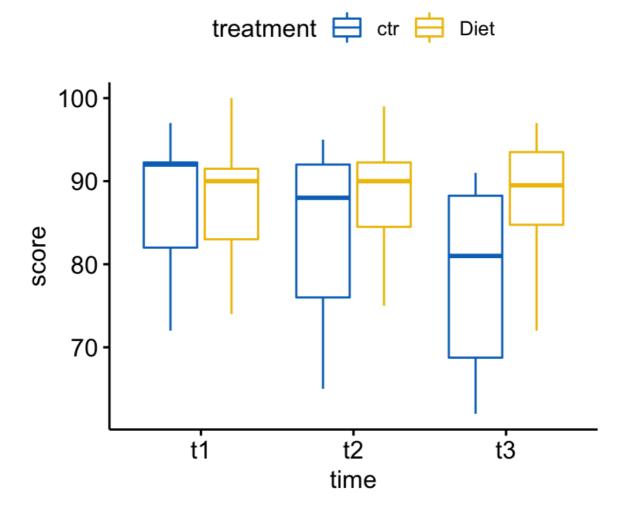
Visualization

Create box plots of the score colored by treatment groups:

```
bxp <- ggboxplot(
selfesteem2, x = "time", y = "score",
color = "treatment", palette = "jco"

bxp

bxp</pre>
```



Chek assumptions

Outliers

```
selfesteem2 %>%
group_by(treatment, time) %>%
identify_outliers(score)
## [1] treatment time id score is.outlier is.extreme
## <0 rows> (or 0-length row.names)
```

There were no extreme outliers.

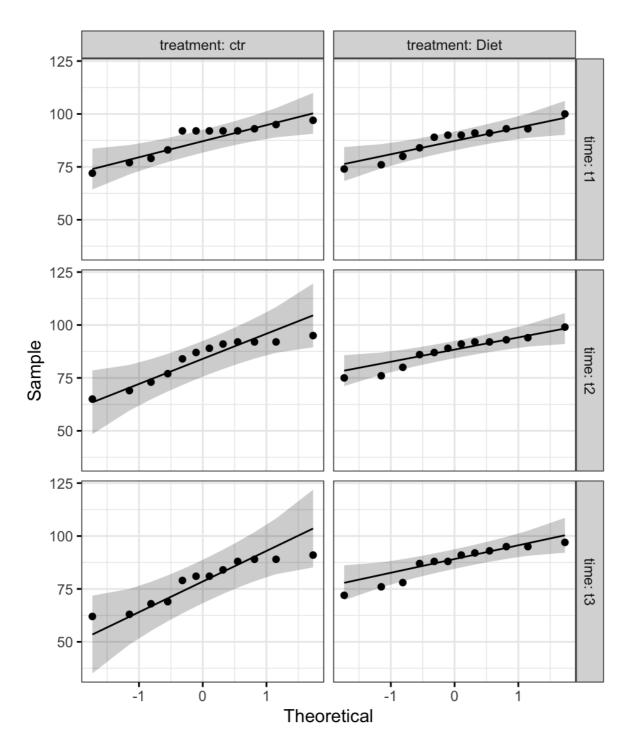
Normality assumption

Compute Shapiro-Wilk test for each combinations of factor levels:

The self-esteem score was normally distributed at each time point (p > 0.05), except for ctr treatment at t1, as assessed by Shapiro-Wilk's test.

Create QQ plot for each cell of design:

```
ggqqplot(selfesteem2, "score", ggtheme = theme_bw()) +
facet_grid(time ~ treatment, labeller = "label_both")
```



From the plot above, as all the points fall approximately along the reference line, we can assume normality.

Computation

```
1 res.aov <- anova_test(</pre>
    data = selfesteem2, dv = score, wid = id,
3
    within = c(treatment, time)
4
5 | get_anova_table(res.aov)
6 ## ANOVA Table (type III tests)
7 ##
8 ##
             Effect DFn DFd F pp<.05 ges
9 ## 1
           treatment 1.00 11.0 15.5 2.00e-03 * 0.059
10 ## 2
            time 1.31 14.4 27.4 5.03e-05
                                               * 0.049
   ## 3 treatment:time 2.00 22.0 30.4 4.63e-07 * 0.050
```

There is a statistically significant two-way interactions between treatment and time, F(2, 22) = 30.4, p < 0.0001.

Post-hoc tests

A **significant two-way interaction** indicates that the impact that one factor (e.g., treatment) has on the outcome variable (e.g., self-esteem score) depends on the level of the other factor (e.g., time) (and vice versa). So, you can decompose a significant two-way interaction into:

- **Simple main effect**: run one-way model of the first variable (factor A) at each level of the second variable (factor B),
- **Simple pairwise comparisons**: if the simple main effect is significant, run multiple pairwise comparisons to determine which groups are different.

For a **non-significant two-way interaction**, you need to determine whether you have any statistically significant **main effects** from the ANOVA output.

Procedure for a significant two-way interaction

Effect of treatment. In our example, we'll analyze the effect of treatment on self-esteem score at every time point.

Note that, the treatment factor variable has only two levels ("ctr" and "Diet"); thus, ANOVA test and paired t-test will give the same p-values.

```
1 # Effect of treatment at each time point
2 one.way <- selfesteem2 %>%
   group_by(time) %>%
   anova_test(dv = score, wid = id, within = treatment) %>%
   get_anova_table() %>%
5
   adjust_pvalue(method = "bonferroni")
7 one.way
8 ## # A tibble: 3 x 9
   ## time Effect DFn
                                       p `p<.05` ges
9
                          DFd F
                                                        p.adj
<dbl> <dbl>
   ## 1 t1 treatment 1 11 0.376 0.552 ""
                                              0.000767 1
11
   ## 2 t2 treatment
                      1 11 9.03 0.012 *
                                              0.052 0.036
12
   ## 3 t3 treatment 1 11 30.9 0.00017 *
                                              0.199 0.00051
13
14 # Pairwise comparisons between treatment groups
   pwc <- selfesteem2 %>%
15
16
   group_by(time) %>%
17
    pairwise_t_test(
   score ~ treatment, paired = TRUE,
18
```

```
19
      p.adjust.method = "bonferroni"
20
      )
21 pwc
22 | ## # A tibble: 3 x 11
23 ## time .y. group1 group2 n1 n2 statistic df
                                                         p.adj
   p.adj.signif
   ## * <fct> <chr> <chr> <int> <int> <dbl> <dbl> <dbl> <dbl>
24
   <chr>>
                                  12 0.613 11 0.552 0.552
   ## 1 t1 score ctr Diet
25
                            12
   ns
                             12
                                  12 -3.00 11 0.012 0.012 *
26 ## 2 t2 score ctr Diet
27 ## 3 t3 score ctr
                             12
                                  12 -5.56 11 0.00017 0.00017
                     Diet
```

Considering the Bonferroni adjusted p-value (p.adj), it can be seen that the simple main effect of treatment was not significant at the time point t1 (p = 1). It becomes significant at t2 (p = 0.036) and t3 (p = 0.00051).

Pairwise comparisons show that the mean self-esteem score was significantly different between ctr and Diet group at t2 (p = 0.12) and t3 (p = 0.00017) but not at t1 (p = 0.55).

Effect of time. Note that, it's also possible to perform the same analysis for the time variable at each level of treatment. You don't necessarily need to do this analysis.

The R code:

```
# Effect of time at each level of treatment
 2 one.way2 <- selfesteem2 %>%
 3
    group_by(treatment) %>%
4
    anova_test(dv = score, wid = id, within = time) %>%
     get_anova_table() %>%
 6
     adjust_pvalue(method = "bonferroni")
 7
    # Pairwise comparisons between time points
8 pwc2 <- selfesteem2 %>%
9
     group_by(treatment) %>%
10
    pairwise_t_test(
       score ~ time, paired = TRUE,
11
       p.adjust.method = "bonferroni"
12
13
14
    pwc2
```

After executing the R code above, you can see that the effect of time is significant only for the control trial, F(2, 22) = 39.7, p < 0.0001. Pairwise comparisons show that all comparisons between time points were statistically significant for control trial.

Procedure for non-significant two-way interaction

If the interaction is not significant, you need to interpret the main effects for each of the two variables: treatment and time. A significant main effect can be followed up with pairwise comparisons.

In our example (see ANOVA table in res.aov), there was a statistically significant main effects of treatment (F(1, 11) = 15.5, p = 0.002) and time (F(2, 22) = 27.4, p < 0.0001) on the self-esteem score.

Pairwise paired t-test comparisons:

```
1 # comparisons for treatment variable
2
   selfesteem2 %>%
 3
    pairwise_t_test(
 4
        score ~ treatment, paired = TRUE,
 5
        p.adjust.method = "bonferroni"
 6
        )
 7 # comparisons for time variable
8
    selfesteem2 %>%
9
      pairwise_t_test(
10
        score ~ time, paired = TRUE,
        p.adjust.method = "bonferroni"
11
12
```

All pairwise comparisons are significant.

Report

We could report the result as follow:

A two-way repeated measures ANOVA was performed to evaluate the effect of different diet treatments over time on self-esteem score.

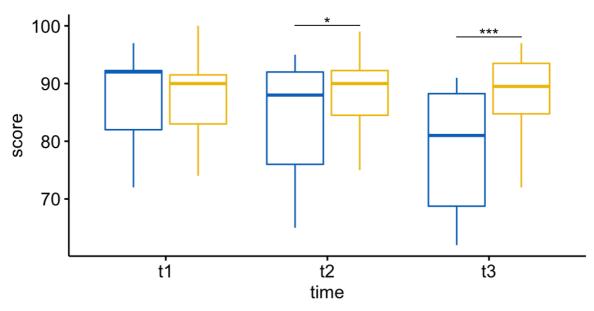
There was a statistically significant interaction between treatment and time on self-esteem score, F(2, 22) = 30.4, p < 0.0001. Therefore, the effect of treatment variable was analyzed at each time point. P-values were adjusted using the Bonferroni multiple testing correction method. The effect of treatment was significant at t2 (p = 0.036) and t3 (p = 0.00051) but not at the time point t1 (p = 1).

Pairwise comparisons, using paired t-test, show that the mean self-esteem score was significantly different between ctr and Diet trial at time points t2 (p = 0.012) and t3 (p = 0.00017) but not at t1 (p = 0.55).

```
# Visualization: box plots with p-values
pwc <- pwc %>% add_xy_position(x = "time")
bxp +
stat_pvalue_manual(pwc, tip.length = 0, hide.ns = TRUE) +
labs(
subtitle = get_test_label(res.aov, detailed = TRUE),
caption = get_pwc_label(pwc)
)
```

Anova,
$$F(2,22) = 30.42$$
, $p = <0.0001$, $\eta_q^2 = 0.05$

treatment 🖨 ctr 📋 Diet



pwc: T test; p.adjust: Bonferroni

Three-way repeated measures ANOVA

Data preparation

We'll use the weightloss dataset [datarium package]. In this study, a researcher wanted to assess the effects of Diet and Exercises on weight loss in 10 sedentary individuals.

The participants were enrolled in four trials: (1) no diet and no exercises; (2) diet only; (3) exercises only; and (4) diet and exercises combined.

Each participant performed all four trials. The order of the trials was counterbalanced and sufficient time was allowed between trials to allow any effects of previous trials to have dissipated.

Each trial lasted nine weeks and the weight loss score was measured at the beginning (t1), at the midpoint (t2) and at the end (t3) of each trial.

Three-way repeated measures ANOVA can be performed in order to determine whether there is a significant interaction between diet, exercises and time on the weight loss score.

Load the data and show some random rows by groups:

```
# Wide format
1
    set.seed(123)
    data("weightloss", package = "datarium")
   weightloss %>% sample_n_by(diet, exercises, size = 1)
5
    ## # A tibble: 4 x 6
         id
               diet exercises
                                  t1
                                        t2
                                              t3
         <fct> <fct> <fct>
                               <dbl> <dbl> <dbl>>
    ## 1 4
                                11.1
                                       9.5 11.1
               no
                     no
    ## 2 10
9
                                10.2 11.8 17.4
               no
                     yes
10
    ## 3 5
               yes
                                11.6 13.4 13.9
```

```
11 ## 4 11 yes yes 12.7 12.7 15.1
12
   # Gather the columns t1, t2 and t3 into long format.
13 # Convert id and time into factor variables
14 | weightloss <- weightloss %>%
   gather(key = "time", value = "score", t1, t2, t3) %>%
15
16
    convert_as_factor(id, time)
   # Inspect some random rows of the data by groups
17
18 set.seed(123)
   weightloss %>% sample_n_by(diet, exercises, time, size = 1)
19
20 ## # A tibble: 12 x 5
      id diet exercises time score
21 ##
## 1 4 no no
                       t1
                             11.1
23
24 ## 2 10 no no
                       t2
                             10.7
25 ## 3 5 no no
                             12.3
                       t3
26 ## 4 11 no yes
                             10.2
                       t1
27 ## 5 12 no yes
                             13.2
                       t2
28 ## 6 1 no yes
                     t3
                             15.8
29 | ## # ... with 6 more rows
```

In this example, the effect of the "time" is our **focal variable**, that is our primary concern.

It is thought that the effect of "time" on the weight loss score will depend on two other factors, "diet" and "exercises", which are called **moderator variables**.

Summary statistics

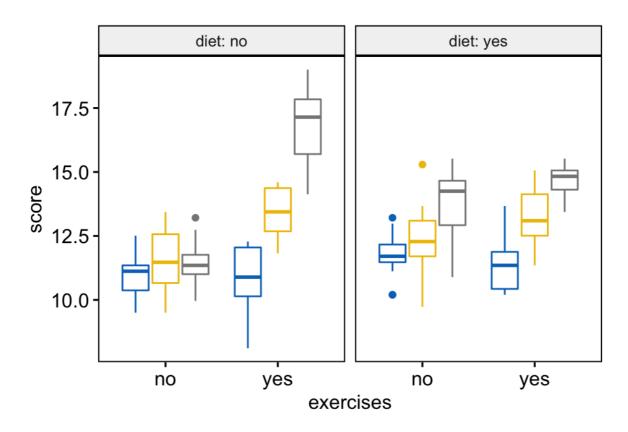
Group the data by diet, exercises and time, and then compute some summary statistics of the score variable: mean and sd (standard deviation)

```
1 weightloss %>%
    group_by(diet, exercises, time) %>%
2
    get_summary_stats(score, type = "mean_sd")
4
  ## # A tibble: 12 x 7
5 ## diet exercises time variable
                                  n mean
  ## <fct> <fct> <fct> <chr> <dbl> <dbl> <dbl>
6
                                12 10.9 0.868
7
   ## 1 no no
                   t1 score
   ## 2 no no
                  t2 score
8
                                  12 11.6 1.30
  ## 3 no no
9
                  t3 score
                                 12 11.4 0.935
                                12 11.4 0.93
   ## 4 no yes
10
                   t1 score
11 ## 5 no yes
                   t2 score
                                 12 13.4 1.01
                                12 16.8 1.53
12 ## 6 no yes
                    t3 score
13 | ## # ... with 6 more rows
```

Visualization

Create box plots:

```
bxp <- ggboxplot(
weightloss, x = "exercises", y = "score",
color = "time", palette = "jco",
facet.by = "diet", short.panel.labs = FALSE
)
bxp</pre>
```



Check assumptions

Outliers

```
weightloss %>%
2
      group_by(diet, exercises, time) %>%
      identify_outliers(score)
    ## # A tibble: 5 x 7
 5
         diet exercises time id
                                      score is.outlier is.extreme
6
         <fct> <fct>
                         <fct> <fct> <dbl> <lgl>
                                                       <1g1>
7
    ## 1 no
                          t3
                                2
                                       13.2 TRUE
                                                       FALSE
                                       10.2 TRUE
8
    ## 2 yes
               no
                         t1
                                1
                                                       FALSE
9
    ## 3 yes
                         t1
                                3
                                       13.2 TRUE
                                                       FALSE
               no
10
                          t1
                                4
                                       10.2 TRUE
    ## 4 yes
               no
                                                       FALSE
                          t2
                                10
                                       15.3 TRUE
11
    ## 5 yes
               no
                                                       FALSE
```

There were no extreme outliers.

Normality assumption

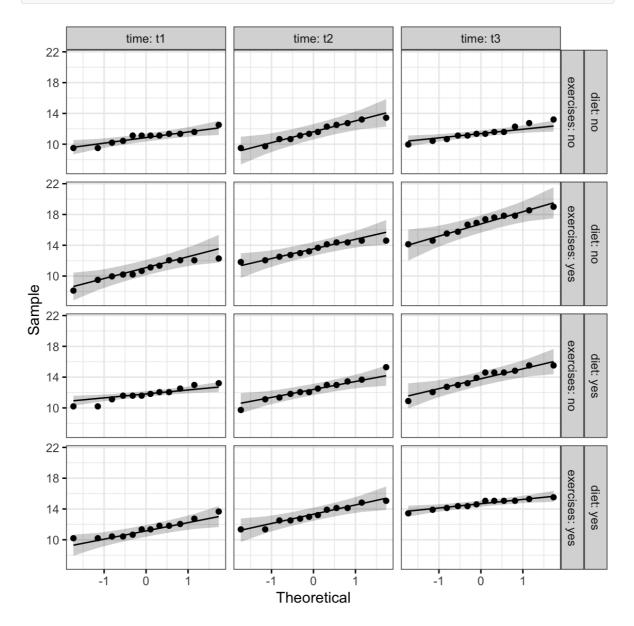
Compute Shapiro-Wilk test for each combinations of factor levels:

```
weightloss %>%
 2
      group_by(diet, exercises, time) %>%
 3
      shapiro_test(score)
    ## # A tibble: 12 x 6
 4
         diet exercises time variable statistic
 6
         <fct> <fct>
                         <fct> <chr>
                                             <dbl> <dbl>
 7
    ## 1 no
                                             0.917 0.264
                         t1
                                score
    ## 2 no
                         t2
                                             0.957 0.743
                                score
 9
    ## 3 no
                         t3
                               score
                                             0.965 0.851
               no
10
    ## 4 no
               yes
                         t1
                                score
                                             0.922 0.306
                                             0.912 0.229
11
    ## 5 no
               yes
                         t2
                               score
                                             0.953 0.674
12
    ## 6 no
               yes
                         t3
                                score
13
    ## # ... with 6 more rows
```

The weight loss score was normally distributed, as assessed by Shapiro-Wilk's test of normality (p > .05).

Create QQ plot for each cell of design:

```
ggqqplot(weightloss, "score", ggtheme = theme_bw()) +
facet_grid(diet + exercises ~ time, labeller = "label_both")
```



From the plot above, as all the points fall approximately along the reference line, we can assume normality.

Computation

```
1 res.aov <- anova_test(</pre>
    data = weightloss, dv = score, wid = id,
    within = c(diet, exercises, time)
5 | get_anova_table(res.aov)
6 | ## ANOVA Table (type III tests)
7 ##
8 ##
                 Effect DFn DFd
                                   F p p<.05 ges
9 ## 1
              diet 1.00 11.0 6.021 3.20e-02 * 0.028
10 ## 2 exercises 1.00 11.0 58.928 9.65e-06
                                                  * 0.284
                time 2.00 22.0 110.942 3.22e-12
11 | ## 3
                                                  * 0.541
12 ## 4 diet:exercises 1.00 11.0 75.356 2.98e-06
                                                  * 0.157
           diet:time 1.38 15.2 0.603 5.01e-01
13 ## 5
                                                    0.013
14 ## 6 exercises:time 2.00 22.0 20.826 8.41e-06 * 0.274
15 | ## 7 diet:exercises:time 2.00 22.0 14.246 1.07e-04
                                                  * 0.147
```

From the output above, it can be seen that there is a statistically significant three-way interactions between diet, exercises and time, F(2, 22) = 14.24, p = 0.00011.

Note that, if the three-way interaction is not statistically significant, you need to consult the two-way interactions in the output.

In our example, there was a statistically significant two-way diet:exercises interaction (p < 0.0001), and two-way exercises:time (p < 0.0001). The two-way diet:time interaction was not statistically significant (p = 0.5).

Post-hoc tests

If there is a significant three-way interaction effect, you can decompose it into:

- Simple two-way interaction: run two-way interaction at each level of third variable,
- Simple simple main effect: run one-way model at each level of second variable, and
- simple simple pairwise comparisons: run pairwise or other post-hoc comparisons if necessary.

Compute simple two-way interaction

You are free to decide which two variables will form the simple two-way interactions and which variable will act as the third (moderator) variable. In the following R code, we have considered the simple two-way interaction of exercises*time at each level of diet.

Group the data by diet and analyze the simple two-way interaction between exercises and time:

```
# Two-way ANOVA at each diet level
two.way <- weightloss %>%
group_by(diet) %>%
anova_test(dv = score, wid = id, within = c(exercises, time))
two.way
## # A tibble: 2 x 2
```

```
7 ## diet anova
 8 ## <fct> 1ist>
 9 ## 1 no <anov_tst>
10 | ## 2 yes <anov_tst>
11 # Extract anova table
12 get_anova_table(two.way)
13 ## # A tibble: 6 x 8
                                          DFn DFd F
14 ## diet Effect
                                                                           p `p<.05` ges

      18
      ## 3 no
      exercises:time
      2
      22 28.9 6.92e- 7 *

      19
      ## 4 yes
      exercises
      1
      11 13.4 4.00e- 3 *

      20
      ## 5 yes
      time
      2
      22 20.5 9.30e- 6 *

      21
      ## 6 yes
      exercises:time
      2
      22 2.57 9.90e- 2 ""

                                                                                         0.504
                                                                                         0.038
                                                                                         0.49
                                                                                          0.06
```

There was a statistically significant simple two-way interaction between exercises and time for "diet no" trial, F(2, 22) = 28.9, p < 0.0001, but not for "diet yes" trial, F(2, 22) = 2.6, p = 0.099.

Note that, it's recommended to adjust the p-value for multiple testing. One common approach is to apply a Bonferroni adjustment to downgrade the level at which you declare statistical significance.

This can be done by dividing the current level you declare statistical significance at (i.e., p < 0.05) by the number of simple two-way interaction you are computing (i.e., 2).

Thus, you only declare a two-way interaction as statistically significant when p < 0.025 (i.e., p < 0.05/2). Applying this to our current example, we would still make the same conclusions.

Compute simple simple main effect

A statistically significant simple two-way interaction can be followed up with **simple simple main effects**.

In our example, you could therefore investigate the effect of time on the weight loss score at every level of exercises or investigate the effect of exercises at every level of time.

You will only need to consider the result of the simple simple main effect analyses for the "diet no" trial as this was the only simple two-way interaction that was statistically significant (see previous section).

Group the data by diet and exercises, and analyze the simple main effect of time. The Bonferroni adjustment will be considered leading to statistical significance being accepted at the p < 0.025 level (that is 0.05 divided by the number of tests (here 2) considered for "diet:no" trial.

```
# Effect of time at each diet X exercises cells
time.effect <- weightloss %>%
group_by(diet, exercises) %>%
anova_test(dv = score, wid = id, within = time)
time.effect
## # A tibble: 4 x 3
## diet exercises anova
## <fct> <fct> <
```

```
12 | ## 4 yes yes <anov_tst>
13
   # Extract anova table
   get_anova_table(time.effect) %>%
14
    filter(diet == "no")
15
16 | ## # A tibble: 2 x 9
17
   ##
      diet exercises Effect DFn DFd
                                         F
                                                  p `p<.05`
                                                             ges
        <fct> <fct> <chr> <dbl> <dbl> <dbl> <dbl>
18 ##
                                               <dbl> <chr>
                                                           <db1>
                             2
                     time
                                    22 1.32 2.86e- 1 ""
   ## 1 no no
                                                           0.075
19
                               2
                                    22 78.8 9.30e-11 *
                                                           0.801
   ## 2 no
             yes
                      time
20
```

There was a statistically significant simple simple main effect of time on weight loss score for "diet:no,exercises:yes" group (p < 0.0001), but not for when neither diet nor exercises was performed (p = 0.286).

Compute simple simple comparisons

A statistically significant simple simple main effect can be followed up by **multiple pairwise comparisons** to determine which group means are different.

Group the data by diet and exercises, and perform pairwise comparisons between time points with Bonferroni adjustment:

```
# Pairwise comparisons
 2
   pwc <- weightloss %>%
 3
    group_by(diet, exercises) %>%
     pairwise_t_test(score ~ time, paired = TRUE, p.adjust.method =
   "bonferroni") %>%
 5
     select(-df, -statistic) # Remove details
   # Show comparison results for "diet:no,exercises:yes" groups
 6
 7
    pwc %>% filter(diet == "no", exercises == "yes") %>%
    select(-p) # remove p columns
8
 9
   ## # A tibble: 3 x 9
10
       diet exercises .y. group1 group2
                                            n1
                                                  n2
                                                           p.adj
    p.adj.signif
   ## <fct> <fct> <chr> <chr> <chr> <int> <int>
                                                           <db1> <chr>
11
12
    ## 1 no
             yes
                       score t1
                                   t2
                                            12
                                                 12 0.000741
                                            13
    ## 2 no
                       score t1
                                   t3
             yes
                                                                ***
14
   ## 3 no
             yes
                       score t2
                                   t3
                                            12
                                                  12 0.000257
```

In the pairwise comparisons table above, we are interested only in the simple simple comparisons for "diet:no,exercises:yes" groups. In our example, there are three possible combinations of group differences. We could report the pairwise comparison results as follow.

All simple simple pairwise comparisons were run between the different time points for "diet:no,exercises:yes" trial. The Bonferroni adjustment was applied. The mean weight loss score was significantly different in all time point comparisons when exercises are performed (p < 0.05).

Report

A three-way repeated measures ANOVA was performed to evaluate the effects of diet, exercises and time on weight loss. There was a statistically significant three-way interaction between diet, exercises and time, F(2, 22) = 14.2, p = 0.00011.

For the simple two-way interactions and simple simple main effects analyses, a Bonferroni adjustment was applied leading to statistical significance being accepted at the p < 0.025 level.

There was a statistically significant simple two-way interaction between exercises and time for "diet no" trial, F(2, 22) = 28.9, p < 0.0001, but not for "diet yes"" trial, F(2, 22) = 2.6, p = 0.099.

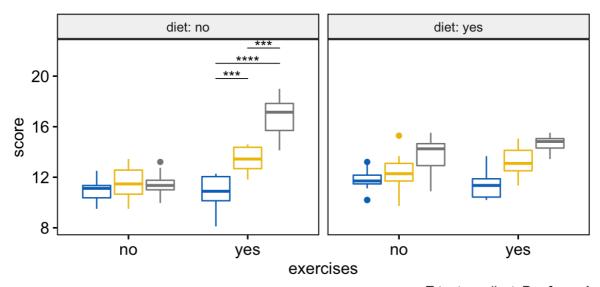
There was a statistically significant simple simple main effect of time on weight loss score for "diet:no,exercises:yes" trial (p < 0.0001), but not for when neither diet nor exercises was performed (p = 0.286).

All simple simple pairwise comparisons were run between the different time points for "diet:no,exercises:yes" trial with a Bonferroni adjustment applied. The mean weight loss score was significantly different in all time point comparisons when exercises are performed (p < 0.05).

```
# Visualization: box plots with p-values
2
    pwc <- pwc %>% add_xy_position(x = "exercises")
    pwc.filtered <- pwc %>%
3
4
      filter(diet == "no", exercises == "yes")
5
    bxp +
      stat_pvalue_manual(pwc.filtered, tip.length = 0, hide.ns = TRUE) +
6
7
8
        subtitle = get_test_label(res.aov, detailed = TRUE),
9
        caption = get_pwc_label(pwc)
10
      )
```

Anova,
$$F(2,22) = 14.25$$
, $p = 0.00011$, $\eta_q^2 = 0.15$





pwc: T test; p.adjust: Bonferroni

Summary

This chapter describes how to compute, interpret and report repeated measures ANOVA in R. We also explain the assumptions made by repeated measures ANOVA tests and provide practical examples of R codes to check whether the test assumptions are met.