Machine Learning 4771

Instructor: Itsik Pe'er

Reminder: Ensembles

Many weak classifiers → a powerful one





(Classification) models

Parametric

Estimate parameters of the distribution of the data

Non-parametric

Reason about data assuming unknown distribution

(Classification) models

Parametric

Estimate parameters of the distribution of the data

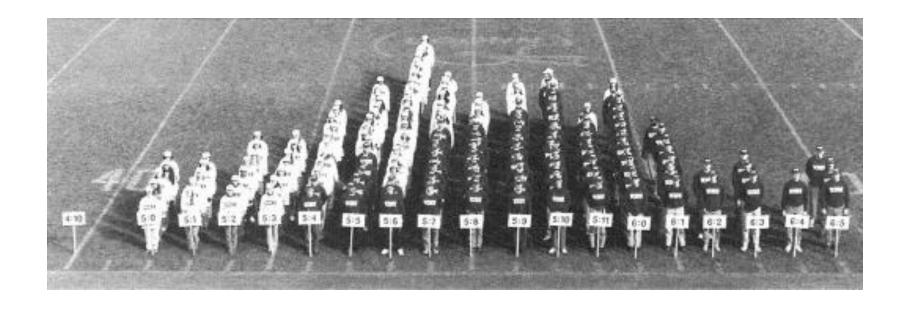
- Logistic regression
- Least squares regression

Non-parametric

Reason about data assuming unknown distribution

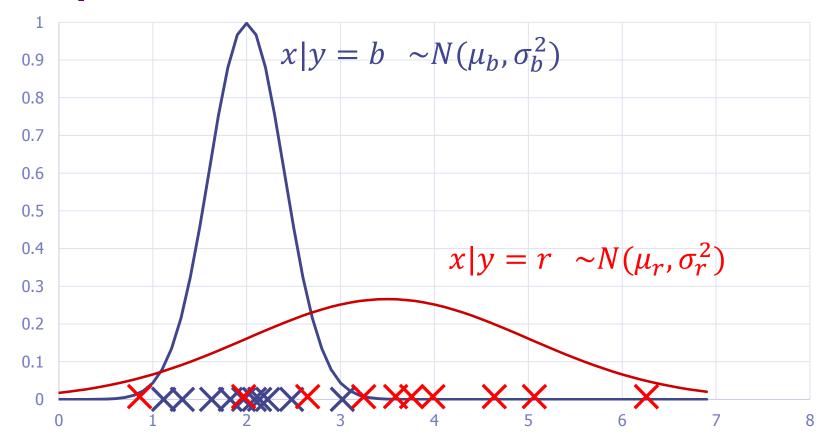
- Nearest neighbors
- Decision trees
- SVM
- RBF regression

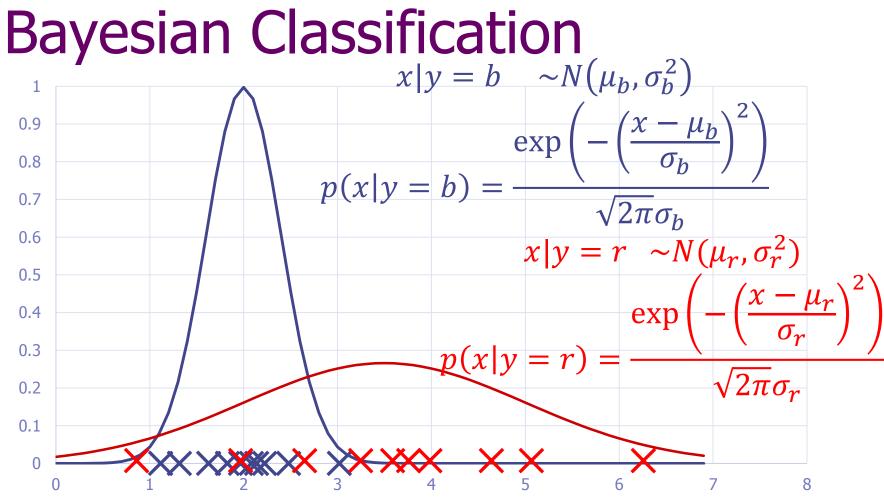
Bayesian Classification



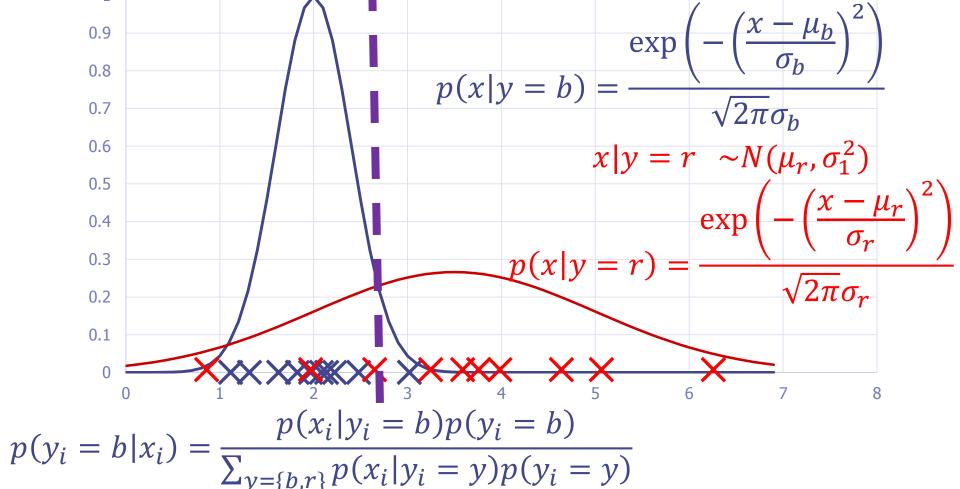
If distribution of each cluster known/inferable: Can be used for classification

Bayesian Classification





Bayesian Classification $x|y=b \sim N(\mu_b, \sigma_b^2)$



If uniform prior: $p(y_i = b|x_i) = Cp(x_i|y_i = b)$ so max likelihood

High Dimensional Data

- Text classification: simplest model
- •10⁵-10⁶ words in English
- •Each document is $D=10^5$ dimensional binary vector \vec{x}_i
- •Each dimension is a word, set to 1 if word in the document

```
Dim1: "we" = 1
Dim2: "hello" = 0
Dim3: "people" = 1
Dim4: "justice" = 1
```

$$p(\vec{x}) = p(\vec{x}(1), \vec{x}(2), ..., \vec{x}(D))$$

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...
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```

- •Each 1 dimensional $\vec{x}(d)$ is a Bernoulli variable
- \vec{x} is multivariate Bernoulli

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Dim1: "we" = 1
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•

•Naïve Bayes: assumes each word is independent

$$p(\vec{x}) = p(\vec{x}(1), \vec{x}(2), ..., \vec{x}(D)) = \prod_{d=1}^{D} p(\vec{x}(d))$$
$$= \prod_{d=1}^{D} \vec{\theta}(d)^{\vec{x}(d)} (1 - \vec{\theta}(d))^{1 - \vec{x}(d)}$$

- Maximum likelihood: assume we have several IID vectors
- •Have N documents, each a 100,000 dimension binary vector
- •Each dimension is a word, set to 1 if word in the document

•Likelihood=

- Maximum likelihood: assume we have several IID vectors
- •Have N documents, each a 100,000 dimension binary vector
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•Likelihood=
$$\prod_{i=1}^{N} p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^{N} \prod_{d=1}^{D} \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$$

•Max likelihood solution:

•Likelihood=
$$\prod_{i=1}^{N} p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^{N} \prod_{d=1}^{D} \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$$

- •Max likelihood solution: for each word d count $\vec{\theta}(d) = \frac{N_d}{N}$
- •Assuming beta-prior: $p\left(\vec{\theta}(d)\right) \sim Beta(1,1)$
- posterior: $p(\vec{\theta}(d)|data) \sim Beta(N_d + 1, (N N_d) + 1)$
- •EAP($\vec{\theta}(d)$) = $\frac{N_d+1}{N+2}$

•Likelihood=
$$\prod_{i=1}^{N} p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^{N} \prod_{d=1}^{D} \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$$

- •Max likelihood solution: for each word d count $\vec{\theta}(d) = \frac{N_d}{N}$
- •Assuming (conjugate) beta-prior: $p(\vec{\theta}(d)) \sim Beta(\alpha, \beta)$
- posterior: $p(\vec{\theta}(d)|data) \sim Beta(N_d + \alpha, (N N_d) + \beta)$
- •EAP($\vec{\theta}(d)$) = $\frac{N_d + \alpha}{N + \alpha + \beta}$
- •To classify new document \vec{x}_{new} , build two models $\vec{\theta}_{religion}$, $\vec{\theta}_{politics}$ Compare: $prediction = argmax_{y \in \{religion, politics\}} p(\vec{x}_{new} | \vec{\theta}_y)$

•Likelihood=
$$\prod_{i=1}^{N} p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^{N} \prod_{d=1}^{D} \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$$

- •Max likelihood solution: for each word d count $\vec{\theta}(d) = \frac{N_d}{N_d}$
- •Assuming (conjugate) beta-prior: $p(\vec{\theta}(d)) \sim Beta(\alpha, \beta)$

posterior:
$$p(\vec{\theta}(d)|data) \sim Beta(N_d + \alpha, (N - N_d) + \beta)$$

- •EAP($\vec{\theta}(d)$) = $\frac{N_d + \alpha}{N + \alpha + \beta}$
- •To classify new document \vec{x}_{new} , build two models $\vec{\theta}_{religion}$, $\vec{\theta}_{politics}$

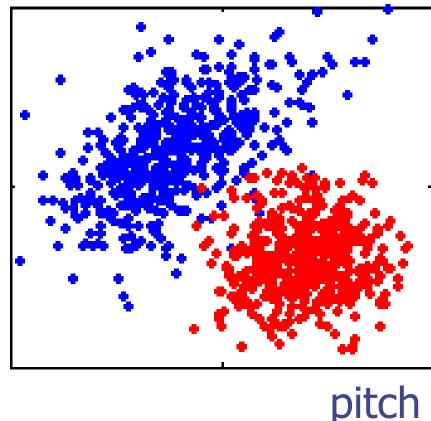
Compare:
$$prediction = argmax_{y \in \{religion, politics\}} \log p(\vec{x}_{new} | \vec{\theta}_y) = argmax_y \sum_{d=1}^{D} (\vec{x}_{new}(d) \log \vec{\theta}_y(d) + (1 - \vec{x}_{new}(d)) \log (1 - \vec{\theta}_y(d)))$$

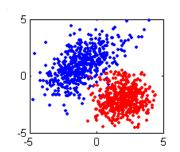
$$\begin{aligned}
gmax_y & \sum_{d=1}^{D} \left(\vec{x}_{new}(d) \log \theta_y(d) + \left(1 - \vec{x}_{new}(d) \right) \right) \\
&= argmax_y \sum_{d=1}^{D} \vec{x}_{new}(d) \log \frac{\vec{\theta}_y(d)}{1 - \vec{\theta}_y(d)}
\end{aligned}$$

Itsik Pe'er, Columbia University

Handling Dependencies: Two 2D Gaussians

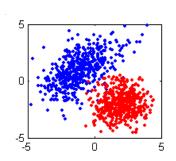
Height





Have two classes, each with their own Gaussian:

$$\{(x_1,y_1),\dots,(x_N,y_N)\}\quad x\in I\!\!R^D,y\in\{0,1\}$$



• Have two classes, each with their own Gaussian:

$$\{(x_1, y_1), \dots, (x_N, y_N)\}\ x \in \mathbb{R}^D, y \in \{0, 1\}$$

- •Given parameters $\theta = \{\alpha, \mu_0, \Sigma_0, \mu_1, \Sigma_1\}$ we can generate iid data from $p(x, y|\theta) = p(y|\theta)p(x|y, \theta)$ by:
 - 1) flipping a coin to get y via Bernoulli $p(y|\theta) = \alpha^y (1-\alpha)^{1-y}$
 - 2) sampling an x from y'th Gaussian $p(x|y,\theta) = N(\mu_y, \Sigma_y)$
- •Recover parameters from data using maximum likelihood $l(\theta)$

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• Have two classes, each with their own Gaussian:

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 - 1) flipping a coin to get y via Bernoulli $p(y|\theta) = \alpha^y (1-\alpha)^{1-y}$
 - 2) sampling an x from y'th Gaussian $p(x|y,\theta) = N(\mu_y, \Sigma_y)$
- •Recover parameters from data using maximum likelihood $l(\theta)$

$$\log p(data|\theta) = \sum_{i=1}^{N} \log p(x_i, y_i|\theta) = \sum_{i=1}^{N} \log p(y_i|\theta) + \sum_{i=1}^{N} p(x_i|y_i, \theta)$$

$$= \sum_{i=1}^{N} \log p(y_i|\alpha) + \sum_{y_i \in 0}^{N} p(x_i|\mu_0, \Sigma_0) + \sum_{y_i \in 1}^{N} p(x_i|\mu_1, \Sigma_1)$$

Max Likelihood can be done separately for the 3 terms

$$l(\theta) = \sum_{i=1}^{N} \log p(y_i | \alpha) + \sum_{y_i \in 0}^{N} p(x_i | \mu_0, \Sigma_0) + \sum_{y_i \in 1}^{N} p(x_i | \mu_1, \Sigma_1)$$

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- •Count # of pos & neg examples (class prior): $\alpha = \frac{N_1}{N_0 + N_1}$
- •Get mean & cov of negatives and mean & cov of positives:

$$\mu_0 = \bar{x} \Big|_{y_i=0} = \frac{1}{N_0} \sum_{y_i=0} x_i \quad \Sigma_0 = \frac{1}{N_0} \sum_{y_i=0} (x_i - \mu_0)(x_i - \mu_0)^T$$

$$\mu_1 = \bar{x} \Big|_{y_i=1} = \frac{1}{N_1} \sum_{y_i=1} x_i \quad \Sigma_1 = \frac{1}{N_1} \sum_{y_i=1} (x_i - \mu_1)(x_i - \mu_1)^T$$

Max Likelihood can be done separately for the 3 terms

$$l(\theta) = \sum_{i=1}^{N} \log p(y_i | \alpha) + \sum_{y_i \in 0}^{N} p(x_i | \mu_0, \Sigma_0) + \sum_{y_i \in 1}^{N} p(x_i | \mu_1, \Sigma_1)$$

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$$\mu_1 = \bar{x}_1 = \bar{x} \Big|_{y_i = 1} = \frac{1}{N_1} \sum_{y_i = 1} x_i \quad \Sigma_1 = \frac{1}{N_1} \sum_{y_i = 1} (x_i - \mu_1) (x_i - \mu_1)^T$$

Posterior μ_y if (conjugate) prior is $N(\mu_p, \Sigma_p)$ and known Σ_y :

$$\mu_y \sim N(\mu_{post}, \Sigma_{post})$$

where
$$:\mu_{post} = \Sigma_{post} (\Sigma_p^{-1} + N\Sigma_y \bar{x}_y), \Sigma_{post} = (\Sigma_p^{-1} + N\Sigma_y^{-1})^{-1}$$

•Max Likelihood can be done separately for the 3 terms

$$l(\theta) = \sum_{i=1}^{N} \log p(y_i | \alpha) + \sum_{y_i \in 0}^{N} p(x_i | \mu_0, \Sigma_0) + \sum_{y_i \in 1}^{N} p(x_i | \mu_1, \Sigma_1)$$

- •Given (x,y) pair, can now compute likelihood p(x,y)
- Bayesian classification:
 - •Without x, can compute prior guess for y: p(y)
 - •Give me x, want y, I need posterior p(y|x)
 - •Bayes optimal decision: $\hat{y} = argmax_{y \in \{0,1\}} p(y|x)$
 - Optimal if we have true probability

Deciding between Gaussians

$$p(y = 1|x) = \frac{p(x, y = 1)}{p(x, y = 0) + p(x, y = 1)}$$
$$= \frac{\alpha N(x|\mu_1, \Sigma_1)}{(1 - \alpha)N(x|\mu_0, \Sigma_0) + \alpha N(x|\mu_1, \Sigma_1)}$$

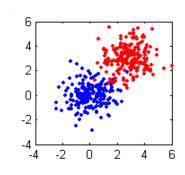
Mahalanobis Distance

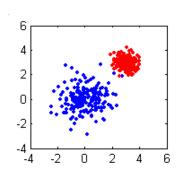
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$$= \frac{\alpha N(x|\mu_1, \Sigma_1)}{(1 - \alpha)N(x|\mu_0, \Sigma_0) + \alpha N(x|\mu_1, \Sigma_1)}$$

$$\log p(y_{new} = y | x_{new}) = C - (x_{new} - \mu_y)^T \Sigma_y^{-1} (x_{new} - \mu_y)$$

$$C = C_{\alpha, x_{new}} + C_{\alpha, \Sigma_y}$$

Mahalanobis Distance (x_{new}, μ_y)





Linear or Quadratic Decisions

•Example cases, plotting decision boundary when = 0.5

$$p(y = 1|x) = \frac{p(x, y = 1)}{p(x, y = 0) + p(x, y = 1)}$$

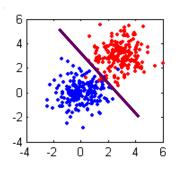
$$= \frac{\alpha N(x|\mu_1, \Sigma_1)}{(1 - \alpha)N(x|\mu_0, \Sigma_0) + \alpha N(x|\mu_1, \Sigma_1)}$$

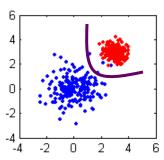
•If covariances are equal:

linear decision

•If covariances are different:

quadratic decision





Summary

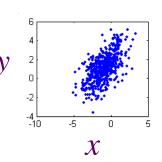
- Naïve Bayes:
 - Assuming independence of features

- Classifying Gaussians:
 - Bayesian
 - Mahalanobis Distance

More Fun with Gaussians

Have input and output, each Gaussian:

$$\{(x_1,y_1),\dots,(x_N,y_N)\}\ x\in {\pmb R}^{D_{\mathcal X}},y\in {\pmb R}^{D_{\mathcal Y}},D=D_{\mathcal X}+D_{\mathcal Y}$$



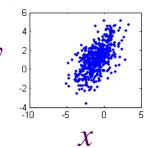
Multiplying Gaussians

•Have input and output, each Gaussian:

Frave input and output, each Gaussian:
$$\{(x_1,y_1),...,(x_N,y_N)\} \ x \in \mathbf{R}^{D_x}, y \in \mathbf{R}^{D_y}, D = D_x + D_y$$
 concatenate
$$z_i = \begin{bmatrix} x_i \\ y_i \end{bmatrix}$$

$$p(z|\mu,\Sigma) = \frac{1}{(2\pi)^{D/2}\sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z-\mu)^T\Sigma^{-1}(z-\mu)\right)$$

Maximum Likelihood



Regression with Gaussians

Have input and output, each Gaussian:

$$\{(x_1, y_1), \dots, (x_N, y_N)\} \ x \in \mathbf{R}^{D_x}, y \in \mathbf{R}^{D_y}, D = D_x + D_y$$

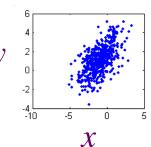
$$\text{concatenate } z_i = \begin{bmatrix} x_i \\ y_i \end{bmatrix}$$

$$p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} \sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z-\mu)^T \Sigma^{-1}(z-\mu)\right)$$

Maximum Likelihood is as usual for a multivariate Gaussian

$$\mu = \frac{1}{N} \sum_{i}^{N} z_{i} \quad \Sigma = \frac{1}{N} \sum_{i}^{N} (z_{i} - \mu)^{T} (z_{i} - \mu)$$

Bayes optimal decision:



 \mathcal{X}

Regression with Gaussians

•Have input and output, each Gaussian:

$$\{(x_1, y_1), \dots, (x_N, y_N)\}\ x \in \mathbf{R}^{D_{\mathcal{X}}}, y \in \mathbf{R}^{D_{\mathcal{Y}}}, D = D_{\mathcal{X}} + D_{\mathcal{Y}}$$

concatenate $z_i = \begin{bmatrix} x_i \\ y_i \end{bmatrix}$

$$p(z|\mu,\Sigma) = \frac{1}{(2\pi)^{D/2}\sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z-\mu)^T \Sigma^{-1}(z-\mu)\right)$$

Maximum Likelihood is as usual for a multivariate Gaussian

$$\mu = \frac{1}{N} \sum_{i}^{N} z_{i} \quad \Sigma = \frac{1}{N} \sum_{i}^{N} (z_{i} - \mu)^{T} (z_{i} - \mu)$$

- •Bayes optimal decision: $\hat{y} = argmax_{y \in R} p(y|x)$
- •Or we can use: $\hat{y} = E_{p(y|x)}\{y\}$

•Have joint, need conditional:
$$p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_{y} p(x,y)}$$

- •Conditional & marginal from joint: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_V p(x,y)}$
- •Gaussian: $p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2}\sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z-\mu)^T \Sigma^{-1}(z-\mu)\right)$ p(x,y) =

$$p(x,y) = \frac{\exp\left(-\frac{1}{2}\left(\begin{bmatrix}x\\y\end{bmatrix} - \begin{bmatrix}\mu_x\\\mu_y\end{bmatrix}\right)^T \begin{bmatrix}\Sigma_{xx} & \Sigma_{xy}\\\Sigma_{yx} & \Sigma_{yy}\end{bmatrix}^{-1}\left(\begin{bmatrix}x\\y\end{bmatrix} - \begin{bmatrix}\mu_x\\\mu_y\end{bmatrix}\right)\right)}{(2\pi)^{D/2}\sqrt{|\Sigma|}} = \frac{1}{2}\left(\frac{1}{2}\left(\frac{x}{y}\right) - \frac{1}{2}\left(\frac{x}{y}\right) - \frac{1}{2}\left(\frac{x}{y}$$

$$p(x,y) = \frac{\exp\left(-\frac{1}{2}(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_{x} \\ \mu_{y} \end{bmatrix})^{T} \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} (\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_{x} \\ \mu_{y} \end{bmatrix}) \right)}{(2\pi)^{D/2}\sqrt{|\Sigma|}} = F_{xx}F_{xy}F_{yy}$$

$$\begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Sigma_{xx}^{-1} & M \\ M^{T} & \Sigma_{yy}^{-1} \end{bmatrix}$$

$$F_{xx} = \frac{\exp\left(-\frac{(x - \mu_{x})^{T}\Sigma_{xx}^{-1}(x - \mu_{x})}{2}\right)}{(2\pi)^{D_{x}/2}\sqrt{|\Sigma_{xx}|}}$$

$$F_{yy} = \frac{\exp\left(-\frac{(y - \mu_{y})^{T}\Sigma_{yy}^{-1}(y - \mu_{y})}{2}\right)}{(2\pi)^{D_{y}/2}\sqrt{|\Sigma_{yy}|}}$$

$$p(x,y) = \frac{\exp\left(-\frac{1}{2}(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_{x} \\ \mu_{y} \end{bmatrix})^{T} \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} (\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_{x} \\ \mu_{y} \end{bmatrix}) \right)}{(2\pi)^{D/2}\sqrt{|\Sigma|}} = F_{xx}F_{xy}F_{yy};$$

$$\begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Sigma_{xx}^{-1} & M \\ M^{T} & \Sigma_{yy}^{-1} \end{bmatrix} M = -\Sigma_{xx}^{-1}\Sigma_{xy}\Sigma_{yy}^{-1}$$

$$F_{xx} = \frac{\exp\left(-\frac{(x-\mu_{x})^{T}\Sigma_{xx}^{-1}(x-\mu_{x})}{2}\right)}{(2\pi)^{Dx/2}\sqrt{|\Sigma_{xx}|}}; F_{yy} = \frac{\exp\left(-\frac{(y-\mu_{y})^{T}\Sigma_{yy}^{-1}(y-\mu_{y})}{2}\right)}{(2\pi)^{Dy/2}\sqrt{|\Sigma_{yy}|}}$$

$$F_{xy} = \sqrt{\frac{|\Sigma_{xx}||\Sigma_{yy}|}{|\Sigma|}} \exp\left(-(x - \mu_x)^T M(y - \mu_y)\right)$$

•Conditional & marginal from joint: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_{\mathbb{R}^n} p(x,y)}$

•Gaussian:
$$p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2}\sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z-\mu)^T \Sigma^{-1}(z-\mu)\right)$$

$$exp\left(-\frac{1}{2}\left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)\right)$$

$$p(x,y) = \frac{1}{(2\pi)^{D_x/2}\sqrt{|\Sigma_{xx}|}} exp\left(-\frac{1}{2}(x - \mu_x)^T \Sigma_{xx}^{-1}(x - \mu_x)\right) = N(\mu_x, \Sigma_{xx})$$

$$p(x) = \frac{1}{(2\pi)^{D_{X}/2} \sqrt{|\Sigma_{XX}|}} \exp\left(-\frac{1}{2}(x - \mu_{X})^{T} \Sigma_{XX}^{-1}(x - \mu_{X})\right) = N(\mu_{X}, \Sigma_{XX})$$

•Conditional & marginal from joint: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_{\mathbb{R}^n} p(x,y)}$

•Gaussian:
$$p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2}\sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z-\mu)^T \Sigma^{-1}(z-\mu)\right)$$

$$p(x,y) = \frac{\exp\left(-\frac{1}{2}\left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)\right)}{(2\pi)^{D/2}\sqrt{|\Sigma|}}$$
$$p(x) = \frac{1}{(2\pi)^{Dx/2}\sqrt{|\Sigma_{xx}|}} \exp\left(-\frac{1}{2}(x - \mu_x)^T \Sigma_{xx}^{-1}(x - \mu_x)\right) = N(\mu_x, \Sigma_{xx})$$

$$p(x) = \frac{1}{(2\pi)^{D_{X}/2} \sqrt{|\Sigma_{XX}|}} \exp\left(-\frac{1}{2}(x - \mu_{X})^{T} \Sigma_{XX}^{-1}(x - \mu_{X})\right) = N(\mu_{X}, \Sigma_{XX})$$

$$p(y|x) = N(\mu_y + \Sigma_{yx}\Sigma_{xx}^{-1}(x - \mu_x), \Sigma_{yy} - \Sigma_{yx}\Sigma_{xx}^{-1}\Sigma_{xy})$$

•Here argmax is conditional expectation:

$$\hat{y} = \mu_{v} + \Sigma_{vx} \Sigma_{xx}^{-1} (x - \mu_{x})$$

