

Machine Learning

4771

Instructor: Itsik Pe'er

Reminder: Ensembles

Many weak classifiers \rightarrow a powerful one



(Classification) models

Parametric

Estimate parameters of the distribution of the data

Non-parametric

Reason about data assuming unknown distribution

(Classification) models

Parametric

Estimate parameters of the distribution of the data

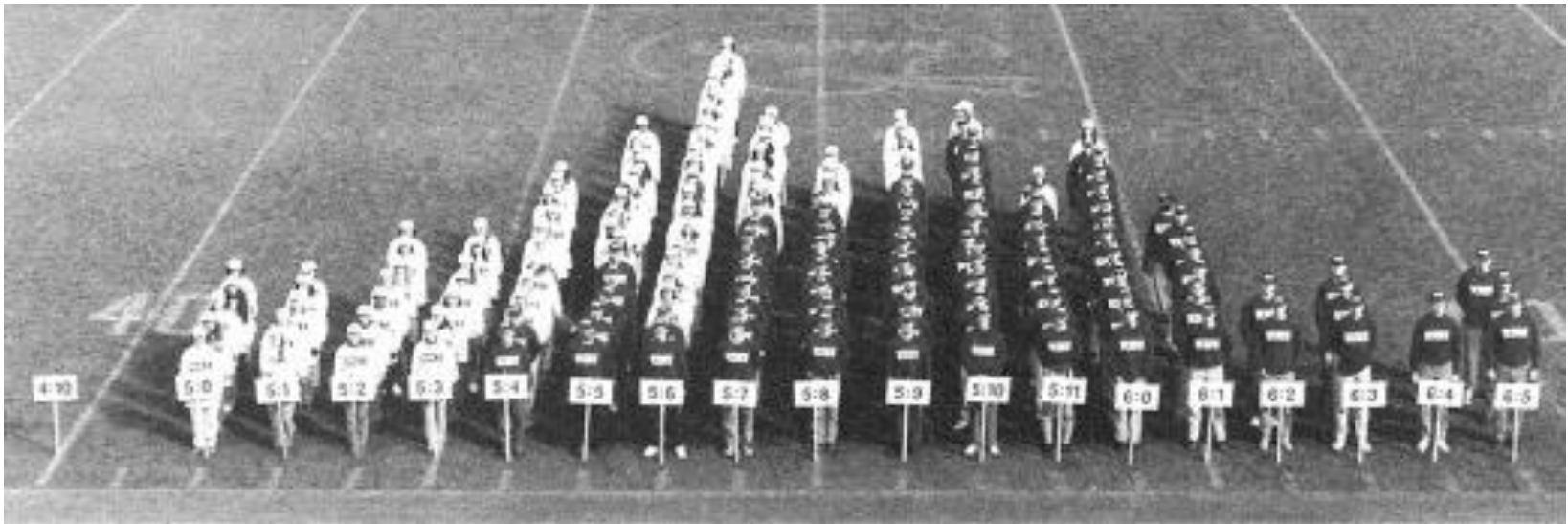
- ◆ Logistic regression
- ◆ Least squares regression

Non-parametric

Reason about data assuming unknown distribution

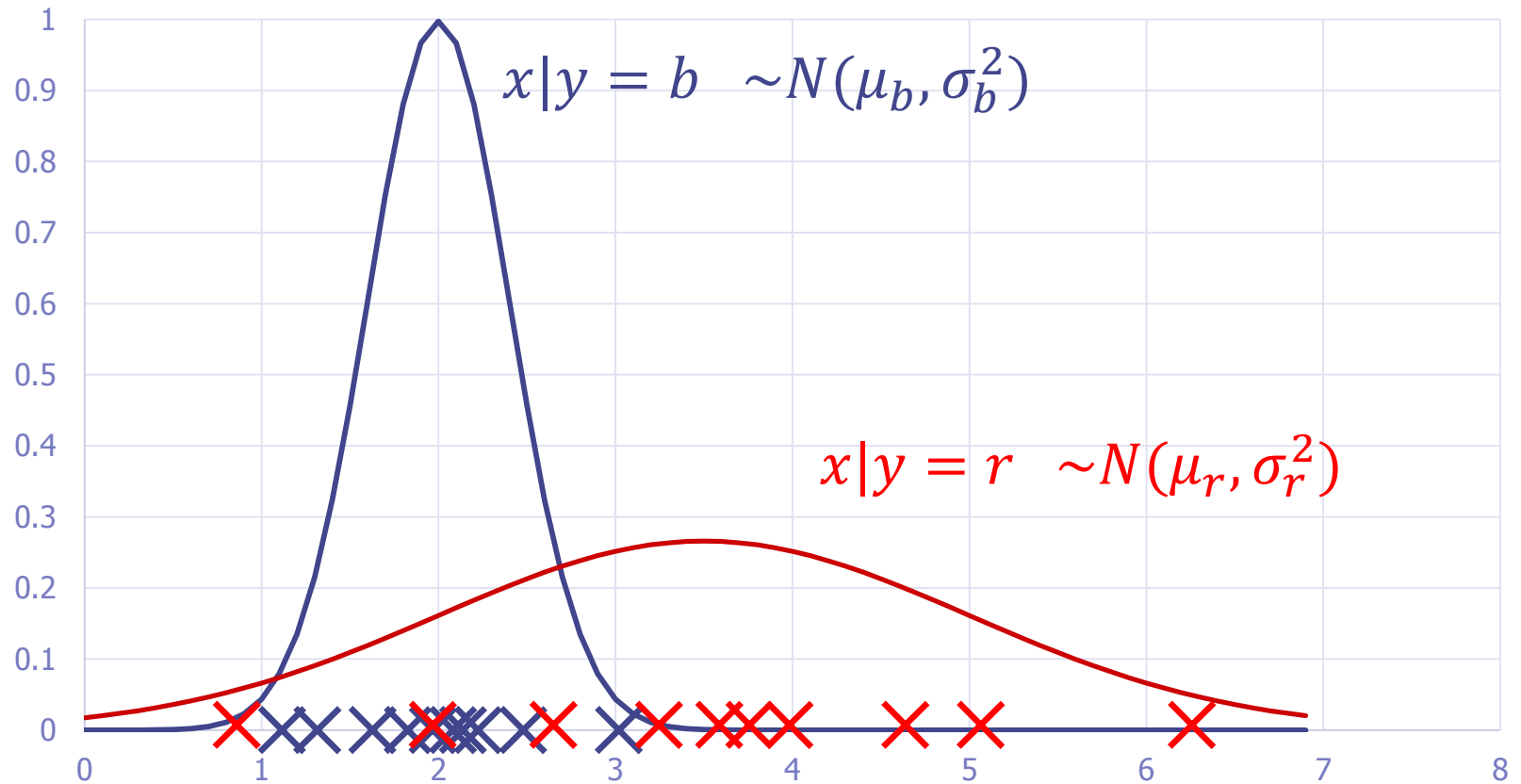
- ◆ Nearest neighbors
- ◆ Decision trees
- ◆ SVM
- ◆ RBF regression

Bayesian Classification

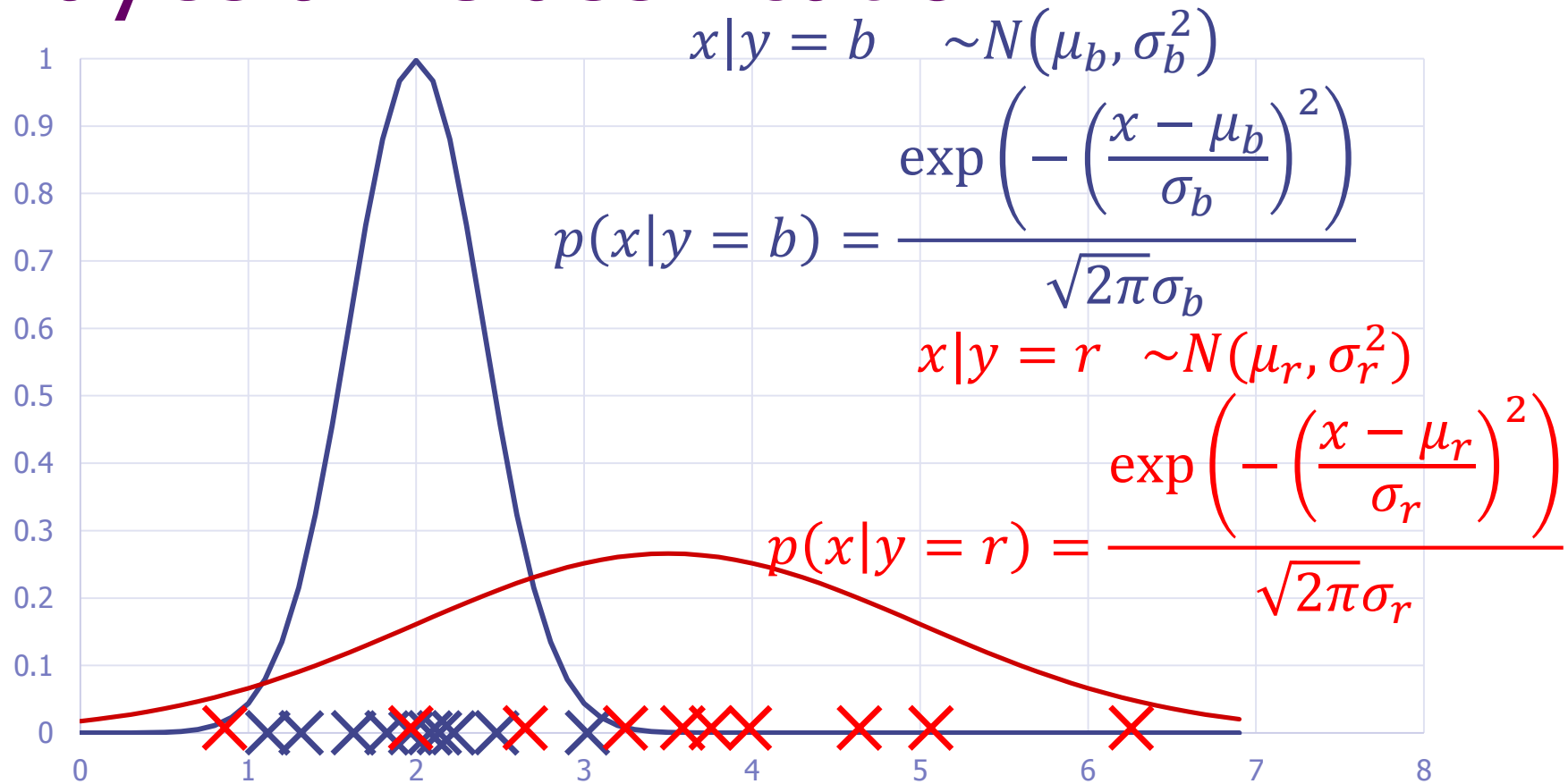


If distribution of each cluster known/inferable:
Can be used for classification

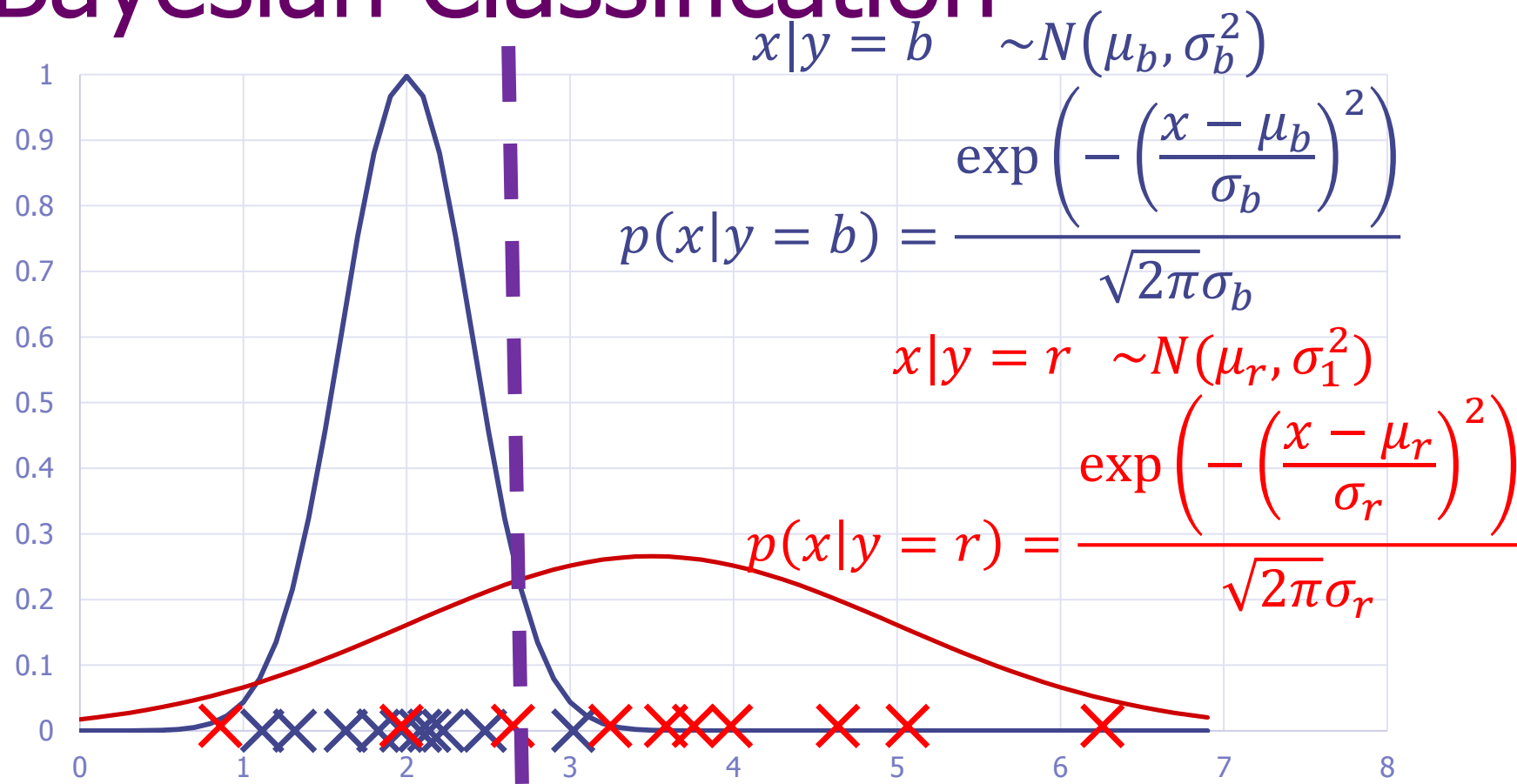
Bayesian Classification



Bayesian Classification



Bayesian Classification

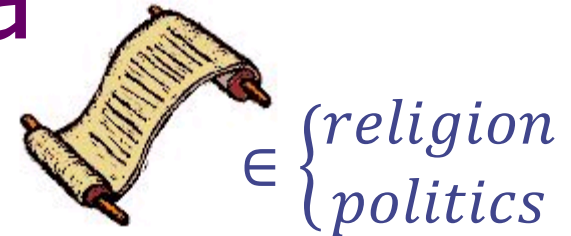


$$p(y_i = b|x_i) = \frac{p(x_i|y_i = b)p(y_i = b)}{\sum_{y=\{b,r\}} p(x_i|y_i = y)p(y_i = y)}$$

If uniform prior: $p(y_i = b|x_i) = C p(x_i|y_i = b)$ so max likelihood

High Dimensional Data

- Text classification: simplest model
- 10^5 - 10^6 words in English
- Each document is $D=10^5$ dimensional binary vector \vec{x}_i
- Each dimension is a word, set to 1 if word in the document



Dim1: "we" = 1

Dim2: "hello" = 0

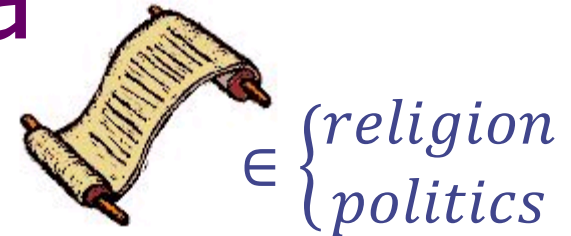
Dim3: "people" = 1

Dim4: "justice" = 1

...

$$p(\vec{x}) = p(\vec{x}(1), \vec{x}(2), \dots, \vec{x}(D))$$

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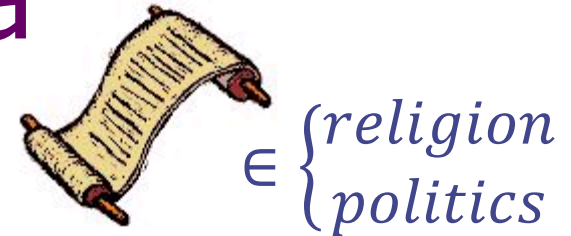
Dim4: "justice" = 1

...

$$p(\vec{x}) = p(\vec{x}(1), \vec{x}(2), \dots, \vec{x}(D))$$

- Each 1 dimensional $\vec{x}(d)$ is a Bernoulli variable
- \vec{x} is multivariate Bernoulli

High Dimensional Data



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Dim1: "we" = 1

Dim2: "hello" = 0

Dim3: "people" = 1

Dim4: "justice" = 1

...

- Naïve Bayes: assumes each word is independent

$$p(\vec{x}) = p(\vec{x}(1), \vec{x}(2), \dots, \vec{x}(D)) = \prod_{d=1}^D p(\vec{x}(d))$$

$$= \prod_{d=1}^D \vec{\theta}(d)^{\vec{x}(d)} (1 - \vec{\theta}(d))^{1-\vec{x}(d)}$$

Text: Naïve Bayes

- Maximum likelihood: assume we have several IID vectors
- Have N documents, each a 100,000 dimension binary vector
- Each dimension is a word, set to 1 if word in the document

		\vec{x}_1	\vec{x}_2	\vec{x}_3	\vec{x}_4
Dim1:	"the"	=	1	0	1
Dim2:	"hello"	=	0	1	0
Dim3:	"and"	=	1	1	0
Dim4:	"happy"	=	1	0	0

- Likelihood=

Text: Naïve Bayes

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Dim4:	"happy"	=	1	0	0

- Likelihood = $\prod_{i=1}^N p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^N \prod_{d=1}^D \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$
- Max likelihood solution:

Text: Naïve Bayes

- Likelihood = $\prod_{i=1}^N p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^N \prod_{d=1}^D \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$
 - Max likelihood solution: for each word d count $\vec{\theta}(d) = \frac{N_d}{N}$
 - Assuming beta-prior: $p(\vec{\theta}(d)) \sim \text{Beta}(1, 1)$
- posterior: $p(\vec{\theta}(d) | \text{data}) \sim \text{Beta}(N_d + 1, (N - N_d) + 1)$
- EAP($\vec{\theta}(d)$) = $\frac{N_d + 1}{N + 2}$

Text: Naïve Bayes

- Likelihood = $\prod_{i=1}^N p(\vec{x}_i | \vec{\theta}) = \prod_{i=1}^N \prod_{d=1}^D \vec{\theta}(d)^{\vec{x}_i(d)} (1 - \vec{\theta}(d))^{(1 - \vec{x}_i(d))}$
- Max likelihood solution: for each word d count $\vec{\theta}(d) = \frac{N_d}{N}$
- Assuming (conjugate) beta-prior: $p(\vec{\theta}(d)) \sim \text{Beta}(\alpha, \beta)$
posterior: $p(\vec{\theta}(d) | \text{data}) \sim \text{Beta}(N_d + \alpha, (N - N_d) + \beta)$
- EAP($\vec{\theta}(d)$) = $\frac{N_d + \alpha}{N + \alpha + \beta}$
- To classify new document \vec{x}_{new} , build two models $\vec{\theta}_{religion}, \vec{\theta}_{politics}$
Compare: $\text{prediction} = \operatorname{argmax}_{y \in \{\text{religion}, \text{politics}\}} p(\vec{x}_{new} | \vec{\theta}_y)$

Text: Naïve Bayes

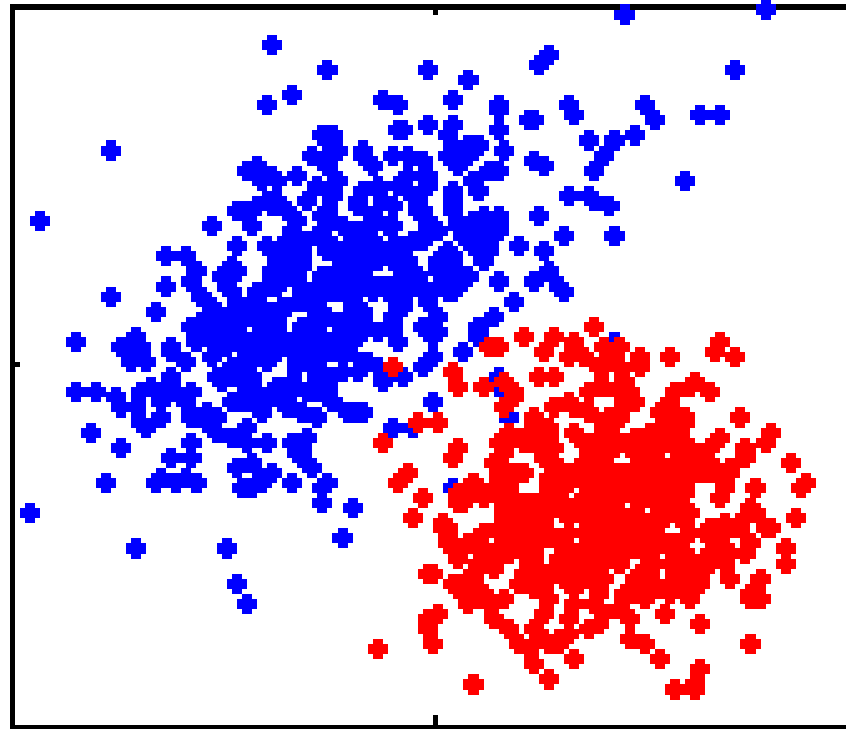
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Compare: $\text{prediction} = \underset{y \in \{religion, politics\}}{\operatorname{argmax}} \log p(\vec{x}_{new} | \vec{\theta}_y) =$

$$\underset{y}{\operatorname{argmax}} \sum_{d=1}^D \left(\vec{x}_{new}(d) \log \vec{\theta}_y(d) + (1 - \vec{x}_{new}(d)) \log (1 - \vec{\theta}_y(d)) \right)$$

$$= \underset{y}{\operatorname{argmax}} \sum_{d=1}^D \vec{x}_{new}(d) \log \frac{\vec{\theta}_y(d)}{1 - \vec{\theta}_y(d)}$$

Handling Dependencies: Two 2D Gaussians

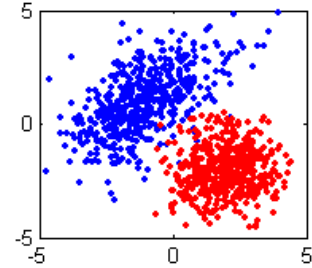
Height



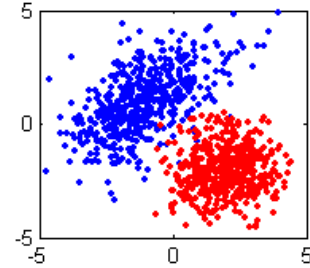
pitch

Classification with Gaussians

- Have two classes, each with their own Gaussian:
 $\{(x_1, y_1), \dots, (x_N, y_N)\} \quad x \in \mathbf{R}^D, y \in \{0, 1\}$

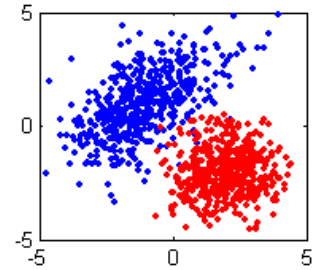


Classification with Gaussians



- Have two classes, each with their own Gaussian:
$$\{(x_1, y_1), \dots, (x_N, y_N)\} \quad x \in \mathbf{R}^D, y \in \{0, 1\}$$
- Given parameters $\theta = \{\alpha, \mu_0, \Sigma_0, \mu_1, \Sigma_1\}$ we can generate iid data from $p(x, y|\theta) = p(y|\theta)p(x|y, \theta)$ by:
 - 1) flipping a coin to get y via Bernoulli $p(y|\theta) = \alpha^y(1 - \alpha)^{1-y}$
 - 2) sampling an x from y 'th Gaussian $p(x|y, \theta) = N(\mu_y, \Sigma_y)$
- Recover parameters from data using maximum likelihood $l(\theta)$

Classification with Gaussians



- Have two classes, each with their own Gaussian:

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2) sampling an x from y 'th Gaussian $p(x|y, \theta) = N(\mu_y, \Sigma_y)$

- Recover parameters from data using maximum likelihood $l(\theta)$

$$\begin{aligned} \log p(\text{data}|\theta) &= \sum_{i=1}^N \log p(x_i, y_i|\theta) = \sum_{i=1}^N \log p(y_i|\theta) + \sum_{i=1}^N \log p(x_i|y_i, \theta) \\ &= \sum_{i=1}^N \log p(y_i|\alpha) + \sum_{y_i \in 0}^N \log p(x_i|\mu_0, \Sigma_0) + \sum_{y_i \in 1}^N \log p(x_i|\mu_1, \Sigma_1) \end{aligned}$$

Classification with Gaussians

- Max Likelihood can be done separately for the 3 terms

$$l(\theta) = \sum_{i=1}^N \log p(y_i|\alpha) + \sum_{y_i \in 0}^N p(x_i|\mu_0, \Sigma_0) + \sum_{y_i \in 1}^N p(x_i|\mu_1, \Sigma_1)$$

Classification with Gaussians

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- Count # of pos & neg examples (class prior): $\alpha = \frac{N_1}{N_0 + N_1}$
- Get mean & cov of negatives and mean & cov of positives:

$$\mu_0 = \bar{x} \Big|_{y_i=0} = \frac{1}{N_0} \sum_{y_i=0} x_i \quad \Sigma_0 = \frac{1}{N_0} \sum_{y_i=0} (x_i - \mu_0)(x_i - \mu_0)^T$$

$$\mu_1 = \bar{x} \Big|_{y_i=1} = \frac{1}{N_1} \sum_{y_i=1} x_i \quad \Sigma_1 = \frac{1}{N_1} \sum_{y_i=1} (x_i - \mu_1)(x_i - \mu_1)^T$$

Classification with Gaussians

- Max Likelihood can be done separately for the 3 terms

$$l(\theta) = \sum_{i=1}^N \log p(y_i|\alpha) + \sum_{y_i \in 0}^N p(x_i|\mu_0, \Sigma_0) + \sum_{y_i \in 1}^N p(x_i|\mu_1, \Sigma_1)$$

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$$\mu_1 = \bar{x}_1 = \bar{x} \Big|_{y_i=1} = \frac{1}{N_1} \sum_{y_i=1} x_i \quad \Sigma_1 = \frac{1}{N_1} \sum_{y_i=1} (x_i - \mu_1)(x_i - \mu_1)^T$$

Posterior μ_y if (conjugate) prior is $N(\mu_p, \Sigma_p)$ and known Σ_y :

$$\mu_y \sim N(\mu_{post}, \Sigma_{post})$$

$$\text{where : } \mu_{post} = \Sigma_{post}(\Sigma_p^{-1} + N\Sigma_y\bar{x}_y), \Sigma_{post} = (\Sigma_p^{-1} + N\Sigma_y^{-1})^{-1}$$

Classification with Gaussians

- Max Likelihood can be done separately for the 3 terms

$$l(\theta) = \sum_{i=1}^N \log p(y_i | \alpha) + \sum_{y_i \in 0}^N p(x_i | \mu_0, \Sigma_0) + \sum_{y_i \in 1}^N p(x_i | \mu_1, \Sigma_1)$$

- Given (x, y) pair, can now compute likelihood $p(x, y)$

- Bayesian classification:

- Without x , can compute prior guess for y : $p(y)$

- Give me x , want y , I need posterior $p(y|x)$

- Bayes optimal decision: $\hat{y} = \operatorname{argmax}_{y \in \{0,1\}} p(y|x)$

- Optimal if we have true probability

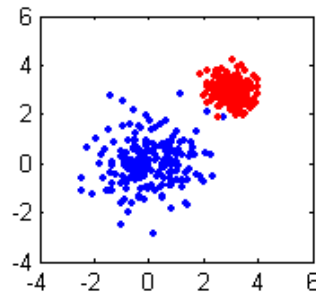
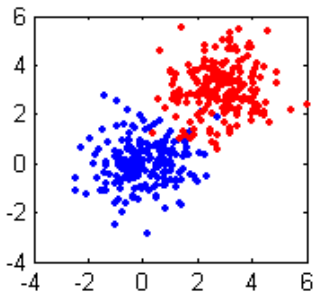
Deciding between Gaussians

$$\begin{aligned} p(y = 1|x) &= \frac{p(x, y = 1)}{p(x, y = 0) + p(x, y = 1)} \\ &= \frac{\alpha N(x|\mu_1, \Sigma_1)}{(1 - \alpha)N(x|\mu_0, \Sigma_0) + \alpha N(x|\mu_1, \Sigma_1)} \end{aligned}$$

Mahalanobis Distance

$$\begin{aligned}
 p(y = 1|x) &= \frac{p(x, y = 1)}{p(x, y = 0) + p(x, y = 1)} \\
 &= \frac{\alpha N(x|\mu_1, \Sigma_1)}{(1 - \alpha)N(x|\mu_0, \Sigma_0) + \alpha N(x|\mu_1, \Sigma_1)}
 \end{aligned}$$

$$\begin{aligned}
 \log p(y_{new} = y|x_{new}) &= C - \underbrace{(x_{new} - \mu_y)^T \Sigma_y^{-1} (x_{new} - \mu_y)}_{\text{Mahalanobis Distance}(x_{new}, \mu_y)} \\
 C &= C_{\alpha, x_{new}} + C_{\alpha, \Sigma_y}
 \end{aligned}$$



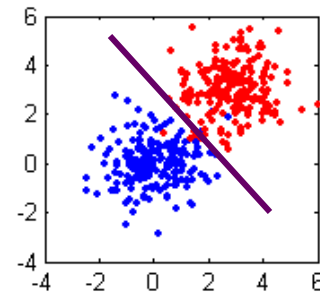
Linear or Quadratic Decisions

- Example cases, plotting decision boundary when $\alpha = 0.5$

$$\begin{aligned}
 p(y = 1|x) &= \frac{p(x, y = 1)}{p(x, y = 0) + p(x, y = 1)} \\
 &= \frac{\alpha N(x|\mu_1, \Sigma_1)}{(1 - \alpha)N(x|\mu_0, \Sigma_0) + \alpha N(x|\mu_1, \Sigma_1)}
 \end{aligned}$$

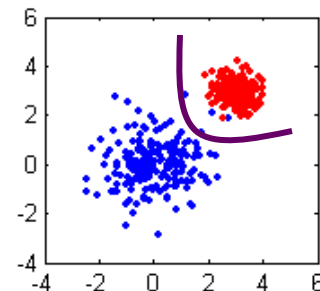
- If covariances are equal:

linear decision



- If covariances are different:

quadratic decision



Summary

◆ Naïve Bayes:

- Assuming independence of features

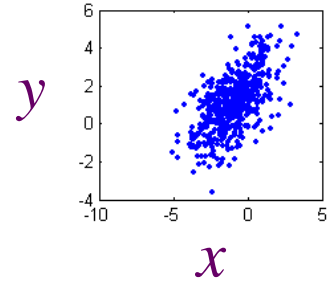
◆ Classifying Gaussians:

- Bayesian
- Mahalanobis Distance

More Fun with Gaussians

- Have input and output, each Gaussian:

$$\{(x_1, y_1), \dots, (x_N, y_N)\} \quad x \in \mathbf{R}^{D_x}, y \in \mathbf{R}^{D_y}, D = D_x + D_y$$



Multiplying Gaussians

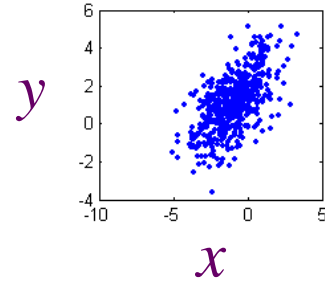
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$$\{(x_1, y_1), \dots, (x_N, y_N)\} \quad x \in \mathbf{R}^{D_x}, y \in \mathbf{R}^{D_y}, D = D_x + D_y$$

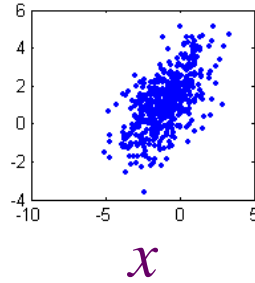
$$\text{concatenate } z_i = \begin{bmatrix} x_i \\ y_i \end{bmatrix}$$

$$p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} \sqrt{|\Sigma|}} \exp \left(-\frac{1}{2} (z - \mu)^T \Sigma^{-1} (z - \mu) \right)$$

- Maximum Likelihood



Regression with Gaussians



- Have input and output, each Gaussian:

$$\{(x_1, y_1), \dots, (x_N, y_N)\} \quad x \in \mathbf{R}^{D_x}, y \in \mathbf{R}^{D_y}, D = D_x + D_y$$

$$\text{concatenate } z_i = \begin{bmatrix} x_i \\ y_i \end{bmatrix}$$

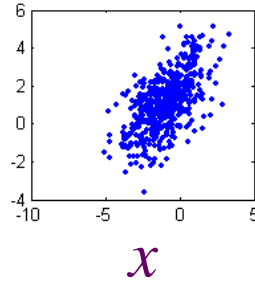
$$p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} \sqrt{|\Sigma|}} \exp \left(-\frac{1}{2} (z - \mu)^T \Sigma^{-1} (z - \mu) \right)$$

- Maximum Likelihood is as usual for a multivariate Gaussian

$$\mu = \frac{1}{N} \sum_i^N z_i \quad \Sigma = \frac{1}{N} \sum_i^N (z_i - \mu)^T (z_i - \mu)$$

- Bayes optimal decision:

Regression with Gaussians



- Have input and output, each Gaussian:

$$\{(x_1, y_1), \dots, (x_N, y_N)\} \quad x \in \mathbf{R}^{D_x}, y \in \mathbf{R}^{D_y}, D = D_x + D_y$$

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- Maximum Likelihood is as usual for a multivariate Gaussian

$$\mu = \frac{1}{N} \sum_i z_i \quad \Sigma = \frac{1}{N} \sum_i (z_i - \mu)^T (z_i - \mu)$$

- Bayes optimal decision: $\hat{y} = \operatorname{argmax}_{y \in \mathbf{R}} p(y|x)$

- Or we can use: $\hat{y} = E_{p(y|x)}\{y\}$

- Have joint, need conditional: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_y p(x,y)}$

Gaussian Marginals/Conditionals

- Conditional & marginal from joint: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_y p(x,y)}$
- Gaussian: $p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} \sqrt{|\Sigma|}} \exp\left(-\frac{1}{2}(z - \mu)^T \Sigma^{-1}(z - \mu)\right)$
 $p(x, y) =$

Gaussian Marginals/Conditionals

$$p(x, y) = \frac{\exp\left(-\frac{1}{2}\left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)\right)}{(2\pi)^{D/2} \sqrt{|\Sigma|}} =$$

Gaussian Marginals/Conditionals

$$p(x, y) = \frac{\exp\left(-\frac{1}{2}\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)}{(2\pi)^{D/2} \sqrt{|\Sigma|}} = F_{xx} F_{xy} F_{yy}$$

$$\begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Sigma_{xx}^{-1} & M \\ M & \Sigma_{yy}^{-1} \end{bmatrix}$$

$$F_{xx} = \frac{\exp\left(-\frac{(x - \mu_x)^T \Sigma_{xx}^{-1} (x - \mu_x)}{2}\right)}{(2\pi)^{D_x/2} \sqrt{|\Sigma_{xx}|}}$$

$$F_{yy} = \frac{\exp\left(-\frac{(y - \mu_y)^T \Sigma_{yy}^{-1} (y - \mu_y)}{2}\right)}{(2\pi)^{D_y/2} \sqrt{|\Sigma_{yy}|}}$$

Gaussian Marginals/Conditionals

$$p(x, y) = \frac{\exp\left(-\frac{1}{2}\left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)\right)}{(2\pi)^{D/2} \sqrt{|\Sigma|}} = F_{xx} F_{xy} F_{yy} ;$$

$$\begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} = \begin{bmatrix} \Sigma_{xx}^{-1} & M \\ M & \Sigma_{yy}^{-1} \end{bmatrix} \quad M = -\Sigma_{xx}^{-1} \Sigma_{xy} \Sigma_{yy}^{-1}$$

$$F_{xx} = \frac{\exp\left(-\frac{(x-\mu_x)^T \Sigma_{xx}^{-1} (x-\mu_x)}{2}\right)}{(2\pi)^{D_x/2} \sqrt{|\Sigma_{xx}|}} ; \quad F_{yy} = \frac{\exp\left(-\frac{(y-\mu_y)^T \Sigma_{yy}^{-1} (y-\mu_y)}{2}\right)}{(2\pi)^{D_y/2} \sqrt{|\Sigma_{yy}|}}$$

$$F_{xy} = \sqrt{\frac{|\Sigma_{xx} \Sigma_{yy}|}{|\Sigma|}} \exp\left(-(x - \mu_x)^T M (y - \mu_y)\right)$$

Gaussian Marginals/Conditionals

- Conditional & marginal from joint: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_y p(x,y)}$
 - Gaussian: $p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} \sqrt{|\Sigma|}} \exp\left(-\frac{1}{2} (z - \mu)^T \Sigma^{-1} (z - \mu)\right)$
- $$p(x, y) = \frac{\exp\left(-\frac{1}{2} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)\right)}{(2\pi)^{D/2} \sqrt{|\Sigma|}}$$
- $$p(x) = \frac{1}{(2\pi)^{D_x/2} \sqrt{|\Sigma_{xx}|}} \exp\left(-\frac{1}{2} (x - \mu_x)^T \Sigma_{xx}^{-1} (x - \mu_x)\right) = N(\mu_x, \Sigma_{xx})$$

Gaussian Marginals/Conditionals

- Conditional & marginal from joint: $p(y|x) = \frac{p(x,y)}{p(x)} = \frac{p(x,y)}{\int_y p(x,y)}$
 - Gaussian: $p(z|\mu, \Sigma) = \frac{1}{(2\pi)^{D/2} \sqrt{|\Sigma|}} \exp\left(-\frac{1}{2} (z - \mu)^T \Sigma^{-1} (z - \mu)\right)$
- $$p(x, y) = \frac{\exp\left(-\frac{1}{2} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)^T \begin{bmatrix} \Sigma_{xx} & \Sigma_{xy} \\ \Sigma_{yx} & \Sigma_{yy} \end{bmatrix}^{-1} \left(\begin{bmatrix} x \\ y \end{bmatrix} - \begin{bmatrix} \mu_x \\ \mu_y \end{bmatrix}\right)\right)}{(2\pi)^{D/2} \sqrt{|\Sigma|}}$$
- $$p(x) = \frac{1}{(2\pi)^{D_x/2} \sqrt{|\Sigma_{xx}|}} \exp\left(-\frac{1}{2} (x - \mu_x)^T \Sigma_{xx}^{-1} (x - \mu_x)\right) = N(\mu_x, \Sigma_{xx})$$
- $$p(y|x) = N(\mu_y + \Sigma_{yx} \Sigma_{xx}^{-1} (x - \mu_x), \Sigma_{yy} - \Sigma_{yx} \Sigma_{xx}^{-1} \Sigma_{xy})$$
- Here argmax is conditional expectation:

$$\hat{y} = \mu_y + \Sigma_{yx} \Sigma_{xx}^{-1} (x - \mu_x)$$

