

# Quiz 2

(MA6.102) Probability and Random Processes, Monsoon 2023

Question cum Answer Booklet

19 October, 2023

Max. Duration: 45 Minutes

Max. Marks: 20

Roll Number: \_\_\_\_\_

Programme: \_\_\_\_\_

Special Instruction: - Please don't do rough work anywhere in this booklet except on the last 3 pages.

Marks Table

Q1	Q2	Q3

**Question 1.** Let  $X_1, X_2, \dots, X_n, \dots$  be a sequence of independent and identically distributed continuous random variables with the common PDF  $f_X(x)$ . Note that  $P(X = \alpha) = 0$ , for all  $\alpha \in \mathbb{R}$ , and that  $P(X_i = X_j) = 0$ , for all  $i \neq j$ . For  $n \geq 2$ , define  $X_n$  as a record-to-date of the sequence if  $X_n > X_i$ , for all  $i < n$ .

- (a) [5 Marks] Find an expression for the expected number of records-to-date that occur over the first  $m$  trials for any given integer  $m$ .

Hint: Use indicator random variables.

- (b) [5 Marks] Let  $A_n$  denote the event that  $X_n$  is a record-to-date, for  $n \geq 2$ . Is  $A_n$  independent of  $A_{n+1}$ , i.e., is  $P(A_n \cap A_{n+1}) = P(A_n)P(A_{n+1})$ ?

(a)

Let  $I_i$  denote the indicator R.V. for the event that  $X_i$  is a record-to-date, for  $i \geq 1$

Let  $N$  denote the no. of records-to-date that occur over first  $m$  trials

$$\text{We then have } N = \sum_{i=1}^m I_i$$

$$\Rightarrow \mathbb{E}[N] = \mathbb{E}\left[\sum_{i=1}^m I_i\right] = \sum_{i=1}^m \mathbb{E}[I_i] \quad [\text{By Linearity of Expectations}]$$

$$\text{Now, } \mathbb{E}[I_i] = 1 \cdot P(I_i = 1) + 0 \cdot P(I_i = 0)$$

$$= P(X_i \text{ is record-to-date})$$

$$= P(X_i > \max\{X_1, \dots, X_{i-1}\})$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{x_i} \dots \int_{-\infty}^{x_i} f_{X_1, \dots, X_i}(x_1, x_2, \dots, x_i) dx_1 dx_2 \dots dx_i$$

$X_i$ 's are independent

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{x_i} \dots \int_{-\infty}^{x_i} f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_i}(x_i) dx_1 dx_2 \dots dx_i$$

$X_i$ 's are identically distr.

$$= \int_{-\infty}^{\infty} \int_{-\infty}^{x_i} \dots \int_{-\infty}^{x_i} f_X(x_1) f_X(x_2) \dots f_X(x_i) dx_1 dx_2 \dots dx_i$$

$$= \int_{-\infty}^{\infty} f_x(x_i) \left( \int_{-\infty}^{x_i} f_x(x_{i-1}) dx_{i-1} \right) \dots \left( \int_{-\infty}^{x_1} f_x(x_1) dx_1 \right) dx_i$$

By defn. of CDF

$$= \int_{-\infty}^{\infty} f_x(x_i) F_x(x_i) \dots F_x(x_1) dx_i$$

$$= \int_{-\infty}^{\infty} f_{x_i}(x_i) (F_x(x_i))^{i-1} dx_i$$

$$\mathbb{E}[I_i] = \int_0^1 t^{i-1} dt = \left[ \frac{t^i}{i} \right]_0^1 = \frac{1}{i}$$

$$\Rightarrow \mathbb{E}[N] = \mathbb{E}[I_1] + \sum_{i=2}^m \mathbb{E}[I_i] = \sum_{i=2}^m \frac{1}{i} = \frac{1}{2} + \frac{1}{3} + \dots + \frac{1}{m} \quad \underline{\text{Ans}}$$

let  $F_x(x_i) = t$   
 $\Rightarrow f_x(x_i) dx = dt$   
 $x_i = -\infty \Rightarrow F_x(x_i) = 0 \Rightarrow t = 0$   
 $x_i = \infty \Rightarrow F_x(x_i) = 1 \Rightarrow t = 1$

[Note  $\mathbb{E}[I_1] = 0$ ,  $\therefore P(X_1 \text{ is record-to-date}) = 0$  as  $X_1$  being r-t-d is an empty event]

(b)  $P(A_n) = P(X_n \text{ is record-to-date}) = \frac{1}{n}$  [from part (a), as our choice of  $i$  was arbitrary, the result in part (a) holds  $\forall i \geq 2$ ]

Similarly,  $P(A_{n+1}) = \frac{1}{n+1}$

$$\begin{aligned} \text{Now, } P(A_n \cap A_{n+1}) &= P(X_n \text{ is record-to-date} \cap X_{n+1} \text{ is record-to-date}) \\ &= P(X_n > \max\{X_1, \dots, X_{n-1}\} \cap X_{n+1} > \max\{X_1, \dots, X_n\}) \\ &= P(X_{n+1} > X_n > \max\{X_1, \dots, X_{n-1}\}) \quad \text{again from i.i.d.} \\ &= \int_{-\infty}^{\infty} \int_{-\infty}^{x_{n+1}} \int_{-\infty}^{x_n} \dots \int_{-\infty}^{x_1} f_x(x_1) f_x(x_2) \dots f_x(x_n) f_x(x_{n+1}) dx_1 \dots dx_{n+1} \\ &= \int_{-\infty}^{\infty} f_x(x_{n+1}) \left[ \int_{-\infty}^{x_{n+1}} f_x(x_n) \left[ \int_{-\infty}^{x_n} f_x(x_{n-1}) dx_{n-1} \dots \left[ \int_{-\infty}^{x_1} f_x(x_1) dx_1 \right] dx_n \right] dx_{n+1} \right] \\ &= \int_{-\infty}^{\infty} f_x(x_{n+1}) \int_{-\infty}^{x_{n+1}} f_x(x_n) (F_x(x_n))^{n-1} dx_n dx_{n+1} = \int_{-\infty}^{\infty} f_x(x_{n+1}) \left[ \frac{F_x(x_n)}{n} \right]_{-\infty}^{x_{n+1}} \end{aligned}$$

$$= \int_{-\infty}^{\infty} f_x(x_{n+1}) F_x(x_{n+1})^n dx_{n+1} = \frac{1}{n} \int_{-\infty}^{\infty} f_x(x_{n+1}) F_x(x_{n+1})^n dx_{n+1} = \frac{1}{n} \left[ \frac{F_x(x_{n+1})^{n+1}}{n+1} \right]_{-\infty}^{\infty} = \frac{1}{n} \frac{1}{n+1}$$

**Question 2** (5 Marks). Consider a discrete integer-valued random variable  $Y$  with CDF

$$F_Y(y) = 1 - \frac{2}{(y+1)(y+2)}, \text{ for integer values } y \geq 0. \quad = P(A_n) P(A_{n+1})$$

Let  $Z$  be another integer valued random variable with the conditional PMF

$$P_{Z|Y}(z|y) = \frac{1}{y^2}, \text{ for } 1 \leq z \leq y^2.$$

Find  $\mathbb{E}[Z]$ .

$$P_Y(y) = F_Y(y) - F_Y(y-1), \text{ for } y \geq 1 \quad [P_Y(0) = 0, \because F_Y(0) = 0]$$

$$= 1 - \frac{2}{(y+1)(y+2)} - 1 + \frac{2}{y(y+1)}$$

$$= \frac{2}{y+1} \left[ \frac{1}{y} - \frac{1}{y+2} \right] = \frac{2}{y+1} \cdot \frac{2}{y(y+2)}$$

$$\mathbb{E}[Z|Y=y] = \sum_{z=1}^{y^2} z P_{Z|Y}(z|y) = \sum_{z=1}^{y^2} z \frac{1}{y^2} = \frac{1}{y^2} \cdot \frac{y^2(y^2+1)}{2}$$

$$\Rightarrow \mathbb{E}[Z|Y] = \frac{y^2+1}{2}$$

$$\mathbb{E}[Z] = \mathbb{E}[\mathbb{E}[Z|Y]] = \mathbb{E}\left[\frac{y^2+1}{2}\right] = \frac{1}{2} \mathbb{E}[Y^2] + \frac{1}{2}$$

By Law of Iterated Expectations

$$\begin{aligned} \text{Now, } \mathbb{E}[Y^2] &= \sum_{y=0}^{\infty} y^2 P_Y(y) = \sum_{y=1}^{\infty} y^2 P_Y(y) \\ &= \sum_{y=1}^{\infty} y^2 \frac{4}{y(y+1)(y+2)} \\ &= 4 \sum_{y=1}^{\infty} \frac{y}{(y+1)(y+2)} \end{aligned}$$

$$= 4 \sum_{y=1}^{\infty} \frac{2(y+1) - (y+2)}{(y+1)(y+2)} = 4 \sum_{y=1}^{\infty} \left( \frac{2}{y+2} - \frac{1}{y+1} \right)$$

$$= 4 \lim_{N \rightarrow \infty} \left[ \sum_{y=1}^N \frac{2}{y+2} - \frac{1}{y+1} \right]$$

$$= 4 \lim_{N \rightarrow \infty} \left[ \sum_{y=1}^N \frac{1}{y+2} + \sum_{y=1}^N \left( \frac{1}{y+2} - \frac{1}{y+1} \right) \right]$$

In general,  
 $\lim_{N \rightarrow \infty} \sum_{n=1}^N A(n) + B(n)$

$= \lim_{N \rightarrow \infty} \sum_{n=1}^N A(n) + \lim_{N \rightarrow \infty} \sum_{n=1}^N B(n)$   
 if either one of the limits  
 is finite

$$= 4 \lim_{N \rightarrow \infty} \left( \sum_{y=1}^N \frac{1}{y+2} \right) + 4 \lim_{N \rightarrow \infty} \left( \frac{1}{N+2} - \frac{1}{2} \right)$$

$$= 4 \underbrace{\lim_{N \rightarrow \infty} \left( \sum_{y=1}^N \frac{1}{y+2} \right)}_{-2} - 2$$

$= \infty$  (since, this is a divergent series)

Hence,  $E[Z] = \infty$

This splitting of sum  
 is allowed since  
 the summation on  
 right is a finite  
 quantity

**Question 3** (5 Marks). Let  $X$  be a continuous random variable with PDF  $f_X$  such that  $f_X(x) > 0$ , for all  $x \in \mathbb{R}$ . Suppose  $g : \mathbb{R} \rightarrow \mathbb{R}$  is a strictly monotonic differentiable function. Find the PDF of  $Y \triangleq g(X)$  and use it to show that

$$\text{Let } g \text{ be monotonically increasing (strictly)} \quad \mathbb{E}[Y] = \int_{-\infty}^{\infty} g(x)f_X(x) dx.$$

$\downarrow$  Ineq. sign remains unchanged  
 $\because g$  is incr.

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \leq g^{-1}(y))$$

$$= F_X(g^{-1}(y)) \quad (\because g \text{ is increasing, } g^{-1} \text{ is well-defined})$$

$$\Rightarrow f_Y(y) = \frac{dF_X(g^{-1}(y))}{d(g^{-1}(y))} \frac{dg^{-1}(y)}{dy} = f_X(g^{-1}(y)) (g^{-1})'(y)$$

$$\text{Now, } \mathbb{E}[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy = \int_{-\infty}^{\infty} y f_X(g^{-1}(y)) (g^{-1})'(y) dy$$

$$\text{let } g^{-1}(y) = t \Rightarrow (g^{-1})'(y) dy = dt$$

$$\Rightarrow \mathbb{E}[Y] = \int_{-\infty}^{\infty} g(t) f_X(t) dt = \int_{-\infty}^{\infty} g(x) f_X(x) dx$$

$\xleftarrow{\text{change of variables}}$

Let  $g$  be monotonically decreasing (strictly)

$$F_Y(y) = P(Y \leq y) = P(g(X) \leq y) = P(X \geq g^{-1}(y))$$

[ Inequality gets reversed,  $\because g$  is decr. ]

$$= P(X > g^{-1}(y))$$

$$= 1 - P(X \leq g^{-1}(y)) = 1 - F_X(g^{-1}(y))$$

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$$\Rightarrow f_Y(y) = -f_X(g^{-1}(y)) (g^{-1})'(y)$$

$$\Rightarrow E[Y] = \int_{-\infty}^{\infty} y f_Y(y) dy = - \int_{-\infty}^{\infty} y f_X(g^{-1}(y)) (g^{-1})'(y) dy$$

$$\text{let } g^{-1}(y) = t \Rightarrow (g^{-1})'(y) dy = dt$$

$$\Rightarrow E[Y] = - \int_{\infty}^{-\infty} g(t) f_X(t) dt \quad \begin{bmatrix} y = -\infty \Rightarrow g^{-1}(y) = \infty \Rightarrow t = \infty \\ y = \infty \Rightarrow g^{-1}(y) = -\infty \Rightarrow t = -\infty \end{bmatrix}$$

$$= - \left[ - \int_{-\infty}^{\infty} g(t) f_X(t) dt \right] = \int_{-\infty}^{\infty} g(t) f_X(t) dt$$

using property of definite integrals

$$= \int_{-\infty}^{\infty} g(x) f_X(x) dx$$



**ROUGH SHEET-1**

**ROUGH SHEET-2**

**ROUGH SHEET-3**