

X : random variable

$E[X^k]$ is called the k^{th} moment of X .

$$\hookrightarrow = \begin{cases} \sum_{j=0}^{\infty} x^j P(X=j) \\ \int x^k f(x) dx \end{cases}$$

$\phi_X(t) = E[e^{tX}]$ is called the moment generating function of X .

Thm: Suppose $\phi_X(t) < \infty$ for $t \in (-t_0, t_0)$ for some $t_0 > 0$.

Then $\phi^{(n)}(0) = E[X^n]$.
 n^{th} derivative

Ex: $X \sim \text{Exp}(\lambda)$ density $\lambda e^{-\lambda x}$.

$$E[e^{tX}] = \int_0^{\infty} e^{tx} \lambda e^{-\lambda x} dx$$

$$\begin{aligned}
 &= \lambda \int_0^{\infty} e^{-(\lambda-t)x} dx \\
 &= -\frac{\lambda}{\lambda-t} e^{-(\lambda-t)x} \Big|_0^{\infty} \\
 \text{when } t < \lambda & \\
 &= 0 + \frac{\lambda}{\lambda-t} e^{-0} \\
 &= \frac{\lambda}{\lambda-t} \leftarrow \text{mgf of } X.
 \end{aligned}$$

Ex: More generally, if $X \sim \text{Gamma}(n, \lambda)$,

$$\begin{aligned}
 E[e^{tx}] &= \int_0^{\infty} e^{tx} \frac{\lambda^n x^{n-1}}{(n-1)!} e^{-\lambda x} dx \\
 &= \frac{\lambda^n}{(\lambda-t)^n} \underbrace{\int_0^{\infty} \frac{(\lambda-t)^n x^{n-1}}{(n-1)!} e^{-(\lambda-t)x} dx}_{1 \quad \text{Gamma}(n, \lambda-t)}
 \end{aligned}$$

$$\phi_X(t) = \frac{\lambda^n}{(\lambda - t)^n} \quad \text{mgf of Gamma}(n, \lambda).$$

Moments of Gamma(n, λ) dist.

$$\phi(t) = \lambda^n (\lambda - t)^{-n}$$

$$\phi'(t) = n \lambda^n (\lambda - t)^{-(n+1)}$$

$$\phi''(t) = n(n+1) \lambda^n (\lambda - t)^{-(n+2)}$$

\vdots

$$\phi^{(k)}(t) = n(n+1) \dots (n+k-1) \lambda^n (\lambda - t)^{-(n+k)}$$

So $E[X^k] = \phi^{(k)}(0) =$

$$\boxed{\frac{n(n+1) \dots (n+k-1)}{\lambda^k}}$$

In particular, if $X \sim \text{Exp}(\lambda)$

$$E[X^k] = \frac{k!}{\lambda^k}$$

Ex: Mgf of Poisson dist. $\text{Poisson}(\lambda)$

$$P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}, \quad k=0, 1, 2, \dots$$

$$E[e^{tX}] = \sum_{k=0}^{\infty} e^{tk} e^{-\lambda} \frac{\lambda^k}{k!}$$

$$= \sum_{k=0}^{\infty} e^{-\lambda} \frac{(\lambda e^t)^k}{k!}$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda e^t)^k}{k!}$$

$$= e^{-\lambda} e^{\lambda e^t}$$

$$= e^{\lambda e^t - \lambda}$$

$$E[X] = \phi'(0), \quad E[X^2] = \phi''(0)$$

$$\phi' = e^{\lambda e^t - \lambda} (\lambda e^t)$$

$$\phi'(t) = \lambda e^{\lambda e^t - \lambda + t}$$

$$\phi''(t) = \lambda e^{\lambda e^t - \lambda + t} (\lambda e^t + 1)$$

$$E[X] = \phi'(0) = \lambda e^{\lambda - \lambda + 0} = \lambda$$

$$\begin{aligned} E[X^2] &= \phi''(0) = \lambda e^{\lambda - \lambda + 0} (\lambda e^0 + 1) \\ &= \lambda(\lambda + 1) = \lambda^2 + \lambda \end{aligned}$$

$$\text{Var}(X) = E[X^2] - (E[X])^2$$

$$= \lambda^2 + \lambda - \lambda^2 = \lambda$$

When X is a non-negative integer-valued rv, the generating function of X is

$$\gamma_X(z) = E[z^X] = \sum_{k=0}^{\infty} z^k P(X=k)$$

(Use convention $0^0 = 1$).

Hint: If we set $z = e^t$, get

$$E[e^{tX}] = \text{mgf.}$$

Ex: Poisson dist. (λ)

$$P(X=k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

$$\gamma(z) = E[z^X] = \sum_{k=0}^{\infty} e^{-\lambda} \frac{\lambda^k}{k!} z^k$$

$$= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda z)^k}{k!}$$

$$= e^{-\lambda} e^{\lambda z} = e^{\lambda(z-1)}$$

Thm: let X be a non neg. integer-valued r.v.

Then
$$\gamma_X^{(n)}(1) = E[X(X-1)\dots(X-n+1)]$$

In particular, $\gamma_X'(1) = E[X]$

and $\gamma_X''(1) = E[X^2 - X] = E[X^2] - E[X].$

Ex: Poisson(λ)

$$\gamma(z) = e^{\lambda(z-1)}$$

$$\gamma'(z) = \lambda e^{\lambda(z-1)} \Rightarrow E[X] = \gamma'(1) = \lambda$$

$$\gamma''(z) = \lambda^2 e^{\lambda(z-1)} \Rightarrow E[X^2 - X] = \lambda^2$$

\vdots

$$\gamma^{(n)}(z) = \lambda^n e^{\lambda(z-1)}$$

$$\Rightarrow E[X(X-1)\dots(X-n+1)] = \lambda^n$$

$$\text{var}(X) = E[X^2] - (EX)^2$$

$$E(X^2) - E(X) = \lambda^2 \Rightarrow E[X^2] = \lambda^2 + \lambda$$

$$\Rightarrow \text{var}(X) = \lambda^2 + \lambda - \lambda^2 = \lambda$$

Ex: $X \sim \text{Geom}(p)$

$$P(X=k) = (1-p)^{k-1} p$$

Calculate generating fn and find EX .

$$\gamma(z) = E[z^X] = \sum_{k=1}^{\infty} z^k (1-p)^{k-1} p$$

$$= \frac{p}{1-p} \sum_{k=1}^{\infty} [z(1-p)]^k$$

$$= \frac{p}{1-p} \cdot \frac{z(1-p)}{1-z(1-p)}$$

$$\gamma(z) = \frac{pz}{1-z(1-p)}$$

$$\gamma'(z) = p \left[\frac{1-z(1-p) - z[-(1-p)]}{(1-z(1-p))^2} \right]$$

$$\gamma'(1) = p \left[\frac{1-(1-p) - (-1+p)}{[1-(1-p)]^2} \right]$$

$$= p \left[\frac{p - (-1+p)}{p^2} \right]$$

$$= \frac{p}{p^2} = \left(\frac{1}{p} \right)$$