

# Probability and Mathematical Statistics: Homework #5

Due on September 18, 2022 at 11:59am

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**Problem 1** (mention the source of question, e.g., BH CH0 #1)**Solution**

$P_k = P(N = K); P_0 = 0, P_1 = 0, P_2 = 0, P_3 = p^3$ , then to use first step method, I can define  $S_i$ : the result of  $i_{th}$  trial,  $S_i = 1(H), S_i = 0(T), P(S_i = 1) = p = \frac{1}{2}, P(S_i = 0) = q = \frac{1}{2}$

$$P(N = k) = P(N = k, |S_1 = 0)P(S_1 = 0) + P(N = K | S_1 = 1, S_2 = 0)P(S_1 = 1, S_2 = 0) + P(N = K | S_1 = 1, S_2 = 1, S_3 = 0) \cdot P(S_1 = 1, S_2 = 1, S_3 = 0) = P(N = K - 1)q + P(N = k - 2)pq + P(N = k - 3)p^2q \quad (1)$$

then to use PGF  $g(t) = E(t^N) = \sum_{k=0}^{\infty} P_k t^k = \sum_{k=3}^{\infty} P_k t^k = p^2 q t^3 + \sum_{k=4}^{\infty} P_k t^k$

on the other hand

$$\sum_{k=4}^{\infty} P_k t^k = q \sum_{k=4}^{\infty} P_{k-1} t^k + pq \sum_{k=4}^{\infty} P_{k-2} t^k + p^2 q \sum_{k=4}^{\infty} P_{k-3} t^k$$

$$g(t) - p^2 q t^3 = qt \cdot g(t) + pqt^2 \cdot g(t) + p^2 q t^3 g(t) \text{ then I can get } g(t) = \frac{-p^2 q t^3}{qt + pqt^2 + p^2 q t^3 - 1}$$

$$E(N) = g'(t)|_{t=1}, \text{ since } p=q=\frac{1}{2}, \text{ so I can get that } g'(t) = \frac{24t^2 - 8t^3 - 2t^4}{(4t + 2t^2 + t^3 - 8)^2}, g'(1) = 14$$

$$\text{Var}(N) = E(N^2) - (E(N))^2 = E(N(N-1)) + E(N) - (E(N))^2 = g''(1) + E(N) - (E(N))^2 = 142$$

**Problem 2 (BH CH0 #2)**

- (a)  $E(T) = E(e^{-3X}) = \sum_{k=0}^{\infty} \frac{e^{-3k} \cdot e^{-\lambda} \lambda^k}{k!} = \sum_{k=0}^{\infty} \frac{e^{-\lambda} (\frac{\lambda}{e^3})^k}{k!} = e^{(\frac{1}{e^3} - 1)\lambda}$ , since it is not equal to  $e^{-3\lambda}$ , so it isn't unbiased.
- (b) To show is to calculate  $E(g(X)) = \sum_{k=0}^{\infty} \frac{(-2)^k e^{-\lambda} \lambda^k}{k!} = e^{-\lambda} \sum_{k=0}^{\infty} \frac{(-2\lambda)^k}{k!} = e^{-3\lambda} = \theta$ , so  $E(g(X)) - \theta = 0$ , so it turns out to be unbiased
- (c) The estimator  $g(X)$  could be negative and will be very large in terms of absolute value if  $X$  is large, while  $\theta$  is a positive small number, so we can change it to a function  $h(X)$

$$h(X) = \begin{cases} 1 & X = 2k (k \in N) \\ 0 & X = 2k + 1 (k \in N) \end{cases}$$

since  $\theta \leq 1$ , it is strictly greater.

**Problem 3 (BH CH0 #3)**

- (a) To find the expected number, we first define  $X$  as the number of times all sixes have achieved all sixes and we can use indicator r.v.  $I_j$  to represent if the  $j_{th}$  roll satisfy all sixes, the trial satisfy a possibility  $p = (\frac{1}{6})^n$ , it satisfy  $X \sim Binom(4 \cdot 6^{n-1}, (\frac{1}{6})^n)$ , so  $E(X) = 4 \cdot 6^{n-1} \times (\frac{1}{6})^n = \frac{2}{3}$
- (b) when  $n$  is a very large number, the  $(\frac{1}{6})^n$  will turn to be very small, but among the  $4 \cdot 6^{n-1}$  trials, the total probability  $\frac{2}{3}$ ,  $X \sim Binom(4 \cdot 6^{n-1}, (\frac{1}{6})^n)$ , so it can use  $X \sim Pois(\frac{2}{3})$ ,  $P(X \geq 1) = 1 - P(X=0) = 1 - e^{-\lambda} = 1 - e^{-\frac{2}{3}}$
- (c) It will not change, since the linearity of expectation doesn't change, we can easily conclude that  $E(X) = E(I_1 + I_2 + \dots)$  can be divided into  $E(I_1) + E(I_2) + \dots + E(I_n)$  though the variable is dependent, the expectation doesn't change, so the answer is same.

**Problem 4 (BH CH0 #4)**

- (a) considering when  $X=k$ , it is same as randomly choose  $m+k$  without replacement and the last time we get the last element as the labeled, the last time the probability we get the labeled is  $\frac{n-m+1}{N-m-k+1}$ , the probability to get  $(m-1)$  labeled in choosing  $m+k$  elements, is  $\frac{\binom{n}{m-1} \cdot \binom{N-n}{k}}{\binom{N}{m+k-1}}$  so the total probability is
- $$P(X = k) = \frac{\binom{n}{m-1} \cdot \binom{N-n}{k}}{\binom{N}{m+k-1}} \cdot \frac{n-m+1}{N-m-k+1}$$
- we can easily get that  $Y=X+m$  so  $P(Y = k) = P(X = k - m) = \frac{\binom{n}{m-1} \cdot \binom{N-n}{k-m}}{\binom{N}{k-1}} \cdot \frac{n-m+1}{N-k+1}$
- (b) there are  $n$  labeled elks, so we can use the labeled elks as indicator variables in the  $N$  trials, just as  $I_j$  : whether this capture happens between the elks labeled  $j$  and  $j-1$ , and the probability I can get is  $\frac{1}{n+1}$ , then I can define that  $I_{ij}$  : which is the indicator variable about whether the  $j_{th}$  unlabeled elk captured between  $i$  and  $i-1$  labeled elk, then the probability of  $I_{ij} = \frac{1}{n+1}$ , then  $E(X) = E(X_1 + X_2 + \dots + X_n) = E(X_1) + E(X_2) + \dots = E(I_{11} + I_{12} + I_{13} + \dots) + E(I_{21} + I_{22} + \dots) + \dots = \sum_{i=1}^m \sum_{j=1}^{N-n} \frac{1}{n+1} = \frac{m(N-n)}{n+1}$
- $$E(Y) = E(X + m) = \frac{mN+m}{n+1}$$
- (c) since the sample is fixed, I can represent  $E(Y)=s$ , supposing  $Z$  as the number of labeled elks in the sampling  $Z \sim \text{Geom}(n, N-n, s)$ , then first I want to calculate  $E(Z)$
- $$E(Z) = \sum \frac{k \binom{n}{k} \binom{N-n}{s-k}}{\binom{N}{s}} = \sum \frac{n \binom{n-1}{k-1} \binom{N-n}{s-k}}{\binom{N}{s}} = n \frac{\binom{N-1}{s-1}}{\binom{N}{s}} = \frac{nE(Y)}{N} = \frac{(nN+n)m}{nN+N}$$
- it is less than  $m$ , (specially when  $N$  is very large, these two are very close)

**Problem 5 (BH CH0 #6)**

- (a) I can separate it into two parts during the exploration phase ,you randomly eat  $k$  different dishes,so on average ,we try the dishes rank= $\frac{n+1}{2}$  the total number of ranks in this phase is  $\frac{n+1}{2}k$  during the exploitation phase,you will always choose  $X$  so using LOTP ,we can get  $\sum P(X = k)X(m-k)=E(X)(m-k)$ ,so the total is  $\frac{n+1}{2}k + E(X)(m - k)$
- (b) the corresponding sequence when  $X=j$  should be has the largest rank  $j$ ,and the other  $k - 1$  dishes rank smaller than  $j$ ,which means pick from  $1 - (j - 1)$ so the PMF  $P(X=k)=\frac{\binom{j-1}{k-1}}{\binom{n}{k}}$
- (c) first,I want to use two equation

i  $j\binom{j-1}{k-1}=k\binom{j}{k}$  using choosing captain or choosing monitor story telling to prove

ii  $\sum_{j=k}^{j=n} \binom{j}{k} = \binom{n+1}{k+1}$  considering a queue with  $n+1$  people with different ages,choosing  $k+1$  from  $n+1$  is equal to considering choosing  $j+1$  oldest person first,than choose the  $k$  people smaller than him.

using following equation, $E(X)=\sum_{j=k}^{j=n} \frac{\binom{j-1}{k-1}j}{\binom{n}{k}} = \sum_{j=k}^{j=n} \frac{\binom{j}{k}k}{\binom{n}{k}} = \frac{k(n+1)}{k+1}$

- (d) we can represent the whole function as  $f(k) = \frac{n+1}{2}k + \frac{k(n+1)(m-k)}{k+1}$  then to solve the problem  $f'(k) = 0$  I can get that  $(k+1)^2 = 2(m+1)$ ,since  $k$  is larger than zero,so  $k = \sqrt{2m+2} - 1$