

Probability and Mathematical Statistics: Homework #10

Due on September 18, 2022 at 11:59am

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Problem 1 (mention the source of question, e.g., BH CH0 #1)

$$(a) \ E(X|X \leq 1) = \sum k(P(X = k|X \geq 1)) = \sum \frac{kP(X=k, X \geq 1)}{P(X \geq 1)} = \sum \frac{ke^{-\lambda}(\lambda)^k}{k!(1-e^{-\lambda})} = \frac{\lambda}{1-e^{-\lambda}}$$

(b)

$$\begin{aligned} Var(X|X \geq 1) &= E(X^2|X \geq 1) - (E(X|X \geq 1))^2 \\ &= \sum_{k=0}^{\infty} \frac{k^2 e^{-\lambda} \lambda^k}{k!(1-e^{-\lambda})} - \left(\sum_{k=0}^{\infty} \frac{k e^{-\lambda} \lambda^k}{k!(1-e^{-\lambda})} \right)^2 \\ &= \frac{\lambda^2 + \lambda}{1-e^{-\lambda}} - \frac{\lambda^2}{(1-e^{-\lambda})^2} \end{aligned}$$

Problem 2 (BH CH0 #2)

- (a) let $Y = X^2$ and $X \sim \text{unif}(0, 1)$ then $E(\frac{X}{X+Y}) = E(\frac{1}{X+1}) = \int_0^1 \frac{x}{x+1} dx = 1 - \ln 2$, $\frac{E(X)}{E(X+Y)} = \frac{\frac{1}{2}}{\frac{1}{2} + \frac{1}{3}} = \frac{3}{5}$
- (b) it is true, since they are iid we can know $E(\frac{X}{X+Y}) = E(\frac{Y}{X+Y})$ (by symmetry), since add them up you can get $E(\frac{X+Y}{X+Y}) = 1$, so $E(\frac{X}{X+Y}) = \frac{1}{2}$, then the rhs $\frac{E(X)}{E(X+Y)} = \frac{E(X)}{E(X)+E(Y)}$ since X and Y are iid, so it is equal to $\frac{1}{2}$ lhs equal to rhs so it is true.
- (c) let T be $(X+Y)^c$, let W be $\frac{X}{X+Y}$. X and Y are independent, and $\sim \text{Gamma}(a, \lambda)$ and $\sim \text{Gamma}(b, \lambda)$, T and W are independent, so does T^c and W^c , so we have $E(T^c W^c) = E(T^c)E(W^c)$, so $E(X^c) = E((X+Y)^c)E((\frac{X}{X+Y})^c)$, so $\frac{E(X^c)}{E(X+Y)^c} = E((\frac{X}{X+Y})^c)$

Problem 3 (BH CH0 #3)

- (a) let X be $\frac{T_1}{T_2}$ Y be $T_1 + T_2$ the Jacobian of $|\frac{\partial(t_1, t_2)}{\partial(x, y)}| = \frac{y}{(x+1)^2}$
 $f_{X,Y}(x, y) = (\lambda_1)^2 e^{-\lambda_1 y} y \frac{1}{(x+1)^2}$ it can be divided into $f(x) \cdot g(y)$ so it is independent
- (b) using lotp, this is to compute $\int_0^\infty P(T_1 < T_2 | T_2 = t_2) f_{T_2}(t_2) dt_2 = \int_0^\infty (1 - e^{-\lambda_1 t_2}) \lambda_2 e^{-\lambda_2 t_2} dt_2 = \int_0^\infty \lambda_2 e^{-\lambda_2 t_2} - \lambda_2 e^{-(\lambda_1 + \lambda_2) t_2} dt_2 = \frac{\lambda_1}{\lambda_1 + \lambda_2}$, especially when $\lambda_1 = \lambda_2$ it is equal to $\frac{1}{2}$
- (c) it is equal to find the minum expectation of wait and serve, to calculate the wait time $E(\min(T_1, T_2))$ $\min(T_1, T_2)$ satisfy $\sim \text{Expo}(\lambda_1 + \lambda_2)$ so the expectation is $\frac{1}{\lambda_1 + \lambda_2}$ then the wait time, suppose the random variable of waiting time is W $f_W(w) = f_W(w | T_1 < T_2) P(T_1 < T_2) + f_W(w | T_1 \geq T_2) P(T_1 \geq T_2)$ conditional on $T_1 < T_2$ it is $\text{Expo}(\lambda_1)$ conditional on $T_1 \geq T_2$ the W is $\text{Expo}(\lambda_2)$ so the expectation of serving time is $\frac{2}{\lambda_1 + \lambda_2}$, add them up, they are totally $\frac{3}{\lambda_1 + \lambda_2}$

Problem 4 (BH CH0 #4)

- (a) let I_j is the position of j_{th} can form a sequence of "CATCAT", the total sequence $E(\sum I_j) = E(I_1 + I_2 + I_3 \cdots + I_n) = 110(p_1 p_2 p_3)^2, E(I_j) = (p_1 p_2 p_3)^2$
- (b) considering when the first A appears before C, it is similar as FS distribution each trial one by one, if it isn't A and C then continue which has probability $1 - p_1 - p_2$ when first A happen it finish so the total probability is $\sum_{t=0}^{\infty} (1 - p_1 - p_2)^t p_1 = \frac{p_1}{p_1 + p_2}$
- (c) the prior distribution of p is $Beta(1, 1)$ after observing the data, in the observed data the number of success is updated by 1, and the number of failure is updated by 2, using the Beta-Binomial conjugacy we can know the posterior is $Beta(2, 3)$ let X be the number of C in the sequence we've observed, to calculate it first we get $f(p|X=1) = \frac{p(1-p)^2}{\beta(2,3)}$, then $\int_0^1 \frac{p^2(1-p)^2}{\beta(2,3)} dp = 0.4$

Problem 5 (BH CH0 #5)

- (a) let the event of HT happen be divided into two independent event W_1 be the number of tosses until first H happens which satisfy $FS(p)$, and W_2 be the number of tosses after the first H happen first T happen which satisfy $FS(1-p)$, so $E(W_1 + W_2) = E(W_1) + E(W_2) = \frac{1}{(1-p)} + \frac{1}{p}$
- (b) using the conditional expectation let W_{HH} be the number of toss when first HT happens $E(W_{HH}) = E(W_{HH}|\text{first toss H})P(\text{first toss H}) + E(W_{HH}|\text{first toss T})P(\text{first toss T})$,
 $E(W_{HH}|\text{first toss H}) = E(W_{HH}|\text{first toss H, second toss H})P(\text{second H}) + E(W_{HH}|\text{first toss H, second toss T})P(\text{second toss T}) = 2p + (E(W_{HH}) + 2)(1-p)$ so the total E is $\frac{1}{p} + \frac{1}{p^2}$
- (c) for (a) the total expectation is $\int_0^1 E(W_{HH}|p)f(p)dp = \int_0^1 (\frac{1}{p} + \frac{1}{1-p}) (\frac{p^{a-1}(1-p)^{b-1}}{\beta(a,b)})dp = \frac{\beta(a-1,b) + \beta(a,b-1)}{\beta(a,b)} = \frac{\Gamma(a-1)\Gamma(a+b)}{\Gamma(a+b-1)\Gamma(a)} + \frac{\Gamma(b-1)\Gamma(a+b)}{\Gamma(a+b-1)\Gamma(a)} = \frac{a+b-1}{a-1} + \frac{a+b-1}{b-1}$ for (b) the total expectation is $\int_0^1 (\frac{1}{p} + \frac{1}{p^2}) \frac{p^{a-1}(1-p)^{b-1}}{\beta(a,b)} dp = \frac{\beta(a-1,b)}{\beta(a,b)} + \frac{\beta(a-2,b)}{\beta(a,b)} = \frac{a+b-1}{a-1} + \frac{(a+b-1)(a+b-2)}{(a-1)(a-2)}$