

## Binomial Distribution

- An experiment with repeated trials.
- Trials are independent.
- Each trial results in SOME EVENT either happening or not happening.
- Occurrence of that event is called SUCCESS and non-occurrence called FAILURE.
- There is the same probability  $p$ , of success in each trail.
- Correspondingly, each trial has the probability of failure,  $q = 1 - p$ .

## Binomial distribution

- $n$  independent trials
- probability  $p$  of success in each trial
- $0 \leq p \leq 1$ ,  $n$  is a positive integer.

**Random variable  $X$ :** number of successes in  $n$  independent trials.

**Probability Distribution of  $X$ :**

$X$  follows *Binomial Distribution*

$$X \sim \text{Bin}(n, p)$$

$P(k \text{ success in } n \text{ trials}) = P(X = k)$ , where

$$P(X = k) = \binom{n}{k} p^k (1 - p)^{n-k}; \quad k = 0, 1, \dots, n$$

**Range of  $X$ :**  $\{0, 1, 2, \dots, n\}$

The binomial probabilities  $P(X = k)$  are the terms in the binomial expansion:

$$(p + q)^n = \sum_{k=0}^n \binom{n}{k} p^k q^{n-k},$$

where,  $q = 1 - p$ .

Exercise:

Check that the binomial distribution is

actually, a distribution by showing that

- $P(X = k) \geq 0$  for all  $k$ .
- $\sum_{k=0}^n P(X = k) = 1$

**Why  $P(k \text{ success in } n \text{ trials}) = \binom{n}{k} p^k (1 - p)^{n-k}$  ???**

Exercise:

See Topic 1: Sampling with replacement example.

See Test 1: Q5

Compare the probabilities and events with binomial distribution.

## **Examples:**

Example 1: Draw 5 balls from an urn, with 6 red balls and 4 green balls. Find the probability distribution of number of red balls. what is the probability of drawing 2 red balls?

Example 2: The probability that a patient recovers from a rare blood disease is 0.4. If 15 people are known to have contracted that diseases, what is the probability that

- at least 3 survive.
- from 3 to 8 survive

Example 3: Text book section 2.1 page 83

What is the probability that among five families, each with six children, at least three of the families have four or more girls?

## Most Likely Number of Successes/Mode of Binomial Distribution

For  $X \sim \text{Bin}(n, p)$ ,  $0 < p < 1$ , the **Mode** or the **most likely number of successes** in  $n$  independent trials with probability  $p$  of success on each trial is  $m$ :

$$m = \lfloor (n + 1)p \rfloor \text{ i.e., the greatest integer } \leq (n + 1)p.$$

- There may be more than one mode in case of ties.
- The Binomial Distribution is *unimodal*. The probabilities are strictly increasing before they reach the maximum and strictly decreasing after the maximum.
- If  $(n + 1)p$  is an integer, then there 2 modes/ most likely numbers for the distribution  $m$  and  $m - 1$ .

Exercise: Book page 87, For the given  $p$  and  $n$ , check the mode formula in a few of these cases to see how it works.

## Normal Distribution

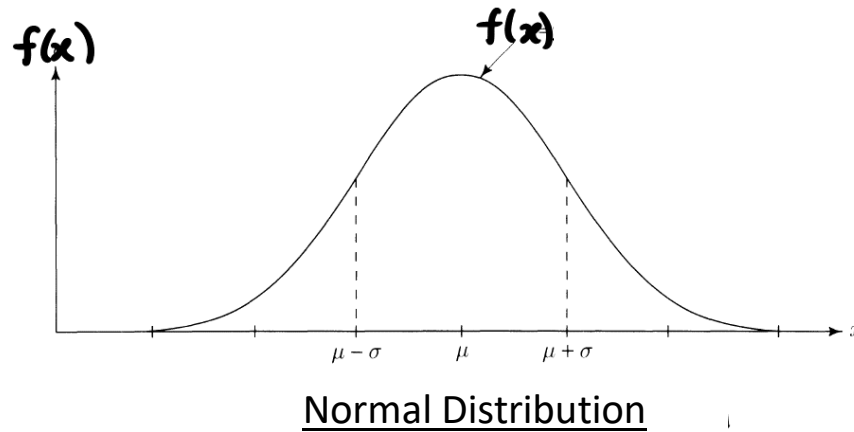
Random variable  $X$  has a **Normal or Gaussian distribution**, with parameters  $\mu$  and  $\sigma^2$ :

$$X \sim N(\mu, \sigma^2)$$

$$\text{Density function: } f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

$\sigma > 0$  and  $\mu$  is arbitrary.

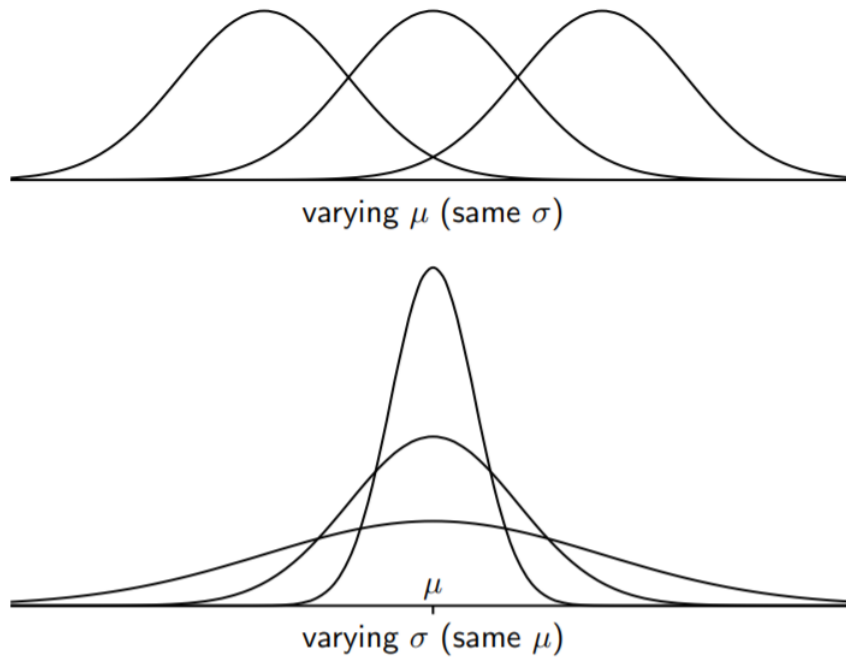
- $\mu$  is mean of the distribution.
- $\sigma^2$  is the variance.
- $\sigma$  is the standard deviation.
- **Range of  $X$ :**  $-\infty < x < \infty$
- $N(\mu, \sigma^2)$  is unimodal with **mode** at  $\mu$ .



### Normal Densities:

$$\text{Density function: } f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

- $\mu$  is a location parameter (changing  $\mu$  just shifts the distribution without changing its shape)
- $\sigma$  is a scale parameter (the distribution is concentrated around  $\mu$  when  $\sigma$  is small, and is spread out when  $\sigma$  is big.)



- Normal densities satisfy the conditions for a probability density function:

1.  $f(x) \geq 0$ , for all  $x$

2.  $\int_{-\infty}^{\infty} \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx = 1$



## Standard Normal Distribution:

Random variable  $Z$  is called **Standard Normal distribution** if parameters  $\mu = 0$  and  $\sigma = 1$

$$Z \sim N(0, 1)$$

$$\text{Density function: } \phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

## CDF of standard normal $Z \sim N(0, 1)$

Let  $\Phi(z)$  be the cdf of  $Z \sim N(0, 1)$

$$\Phi(z) = \int_{-\infty}^z \phi(t) dt$$

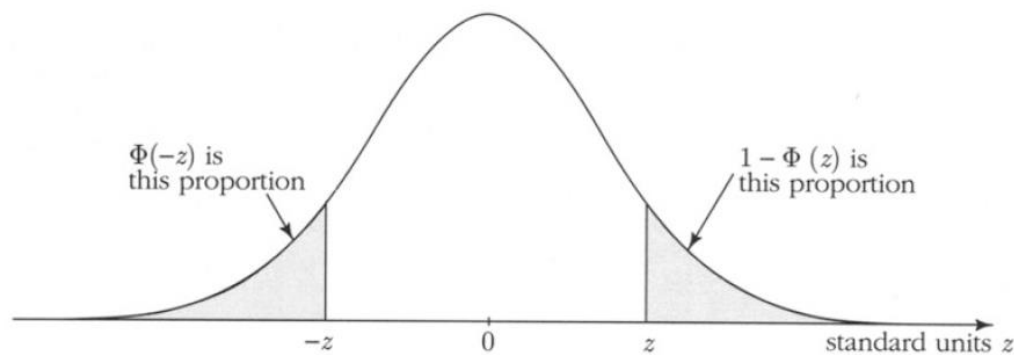
- No closed-form expression for  $\Phi$ , use table of values instead (Appendix 5, for  $z \geq 0$ ).

- $\Phi'(z) = \phi(z) = \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$
- $\Phi(z)$  is symmetric about mean  $\mu = 0$

$$\Phi(z) = P(Z \leq -z) = P(Z \geq z) = 1 - \Phi(z)$$

- $\Phi(a, b) = \Phi(b) - \Phi(a)$
- $\Phi(-z, z) = 2\Phi(z) - 1$

FIGURE 3. Symmetry of the normal curve.



## Cumulative Distribution Function of Normal $X \sim N(\mu, \sigma^2)$

The cdf for normal random variable  $N(\mu, \sigma^2)$  is:

$$F(x) = \Phi\left(\frac{x - \mu}{\sigma}\right)$$

- Transformation of Standard Normal Distribution:

Lemma: If  $X = \mu + \sigma Z$  then  $X \sim N(\mu, \sigma^2) \Leftrightarrow Z \sim N(0, 1)$ .

Exercise:

Prove the converse. If  $X \sim N(\mu, \sigma^2)$  and  $X = \mu + \sigma Z$ . Show that  $Z \sim N(0, 1)$

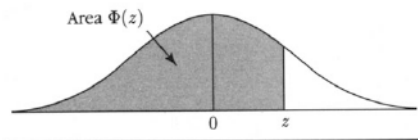
## How to find Normal Probabilities?

Calculate probabilities for Normal  $X \sim N(\mu, \sigma^2)$  using

1. the general cdf formulae obtained using transformations.
2. Standard normal cdf  $\Phi$  (table)
3. continuity (  $\Phi(z-) = \Phi(z)$  );
4. symmetry (  $\Phi(z) = 1 - \Phi(-z)$  )

**Example:** Let  $X \sim N(1, 4)$ . Find  $P(0.5 \leq X \leq 3.46)$ .

Solve from notes IV



## Appendix 5

## Normal Table

Table shows values of  $\Phi(z)$  for  $z$  from 0 to 3.59 by steps of .01. Example: to find  $\Phi(1.23)$ , look in row 1.2 and column .03 to find  $\Phi(1.2 + .03) = \Phi(1.23) = .8907$ . Use  $\Phi(z) = 1 - \Phi(-z)$  for negative  $z$ .

[illegible]

What if the  $z$  you are looking for is not in the  $\Phi$  table?

- software (The best choice)
  1. *NORMSDIST* function in Excel
  2. *pnorm* function in R
  3. *normcdf* function in Matlab
  4. etc
- crude approximation (round  $z$  to 2 decimals and use the corresponding value from the table)
- Linear interpolation for  $l \leq z \leq r$   
 $z = l + \lambda(r - l)$  then

$$\Phi(z) \approx \Phi(l) + \lambda(\Phi(r) - \Phi(l)),$$

$$\Phi(z) \approx \Phi(l) + \frac{z-l}{r-l} (\Phi(r) - \Phi(l))$$

Note: For  $x = r$  or  $x = l$ , the formula is exact.

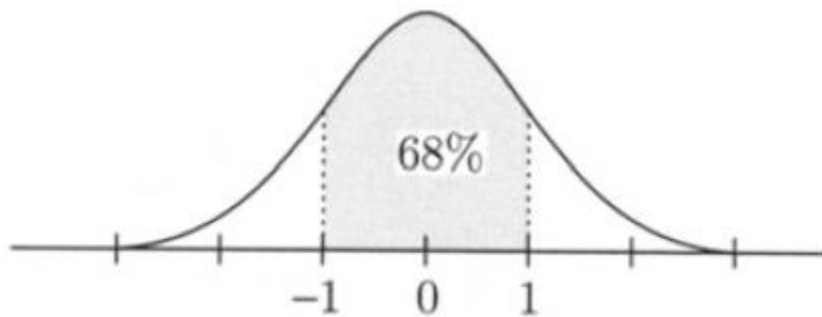
- Program an approximation formula for  $\Phi(z)$  [eg. on pg. 95 of the textbook]

Example:  $X \sim N(2, 3)$ . Find  $P(X \leq 4)$

**Rule of Thumb regarding Normal densities:**

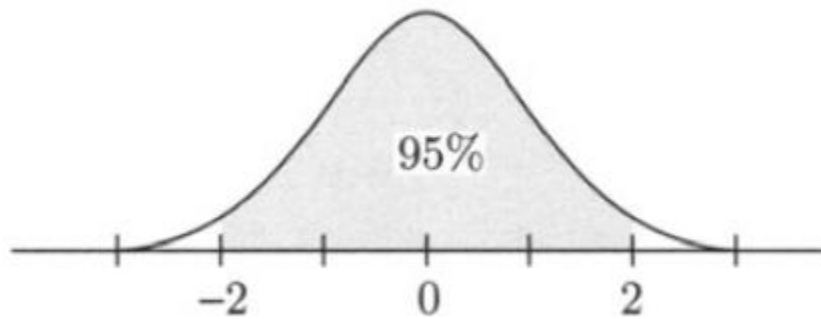
- 70% of the mass lies within 1 standard deviation of the mean.

$$P(\mu - \sigma \leq X \leq \mu + \sigma) \approx 0.70$$



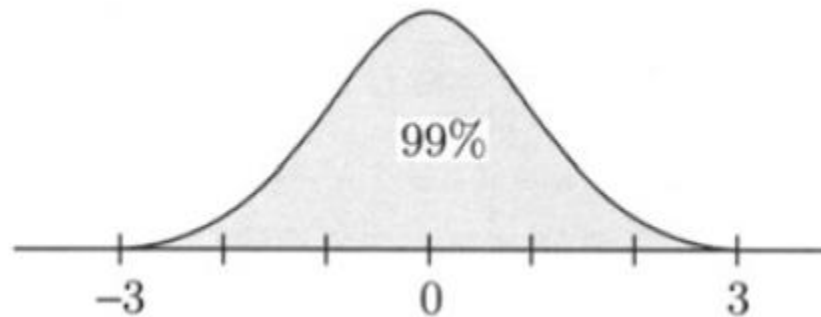
- 95% of the mass lies within 2 standard deviations of the mean.

$$P(\mu - 2\sigma \leq X \leq \mu + 2\sigma) \approx 0.95$$



- 99% of the mass lies within 3 standard deviations of the mean.

$$P(\mu - 3\sigma \leq X \leq \mu + 3\sigma) \approx 0.99$$



## Normal Approximation to Binomial Distribution:

- $Bin(n, p)$  prob's can be worked out exactly, when  $n$  is small.
- For **large  $n$** , we approximate binomial distribution by normal distribution.

Let  $n$  be large and  $X \sim Bin(n, p)$ . Then,

$X \approx Y$ , where  $Y \sim N(np, np(1 - p))$ .

where,  $\mu = np$  and  $\sigma^2 = np(1 - p)$ .

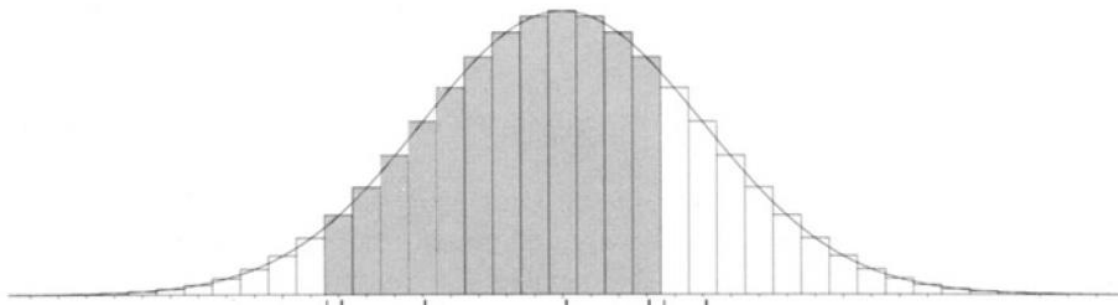
- (Crude) approximation formula for Binomial probabilities:

$X \sim Bin(n, p)$  and  $n$  large

$\Rightarrow P(X \leq x) \approx P(Y \leq x)$

where  $Y \sim N(np, np(1 - p))$ .





Example:  $X \sim \text{Bin}(1000, 0.5)$ . Find  $P(X \leq 495)$ .

Note: Under this crude approximation,  $P(X = 495) \approx P(Y = 495) = 0$  since  $Y$  has a continuous distribution.

For discrete Binomial distribution It may be true that  $P(X = 495)$  is small. But how small?

### Continuity Correction

Need to correct for approximating a discrete distribution by a continuous one.

For a general discrete r.v.  $X$ :

- Take possible values  $x_1, \dots, x_n$ .
- Let  $\delta_i = x_{i+1} - x_i$  be the distance between neighbouring values. (Note: take  $\delta_0 = -\infty, \delta_n = +\infty$ ).

- Split the difference between neighbouring values, we have that  $x_i$  is the only possible value for  $X$  in the interval  $[x_i - \frac{\delta_{i-1}}{2}, x_i + \frac{\delta_i}{2}]$ .
- $$P(X \leq x_i) = P(X \leq x_i + \frac{\delta_i}{2}), P(X \geq x_i) = P(X \geq x_i - \frac{\delta_{i-1}}{2})$$
- $$P(X = x_i) = P\left(x_i - \frac{\delta_{i-1}}{2} \leq X \leq x_i + \frac{\delta_i}{2}\right).$$
- Approximating  $X$  by a r.v. with a continuous distribution, apply the approximation to these expanded events to get more accurate results.
- In the binomial case  $\delta_i = 1$ , for all  $i$ .

Example:  $X \sim \text{Bin}(1000, 0.5)$ . Find

- $P(X = 495)$
- $P(X \leq 495)$

To summarize:

- Can do normal approximation with or without a continuity correction.
- Including the correction gives greater accuracy when approximating binomials
- The normal probabilities can be found using software, crude rounding, or linear interpolation.

Example: Batting averages (§2.2 Problem 11a)

If a player's true batting average is .300, what is the probability of hitting .310 or better over the next 100 at bats?