



Question 1.

证明: $\because P(w_i|x) \geq P(w_j|x)$ 所以在 w_i 有最大后验概率.

(1) ① 若在特征集 x , 则判为非 i 类的所有损失为: $\lambda_s \sum_{j \neq i} P(w_j|x) = \lambda_s [1 - P(w_i|x)]$

若选为非 i 类的 k 类, 风险为 $\lambda_s \sum_{j \neq i} P(w_j|x) = \lambda_s [1 - P(w_i|x)] \geq \lambda_s [1 - P(w_i|x)]$

\therefore 在非拒绝情况下, 判为 w_i 风险损失最小.

② 若 $\lambda_r \leq \lambda_s [1 - P(w_i|x)]$ 即 $P(w_i|x) \leq 1 - \frac{\lambda_r}{\lambda_s}$ 则拒绝风险最小.

综上: 在 $P(w_i|x) \geq P(w_j|x)$ 且 $P(w_i|x) \geq 1 - \lambda_r/\lambda_s$ 判为 w_i , 否则拒绝.

(2) 若 $\lambda_r = 0$ 则拒绝情况下的风险也为 0, 即对任意 j , 判 w_i 或拒绝均可.

(3) 若 $\lambda_r > \lambda_s$ 则恒有 $P(w_i|x) > 1 - \frac{\lambda_r}{\lambda_s}$, 则判为 w_i 风险最小, 不会拒绝.

Question 2

$$(a) P_e = P(|x - \mu_1| > |x - \mu_2| | w_1) P(w_1) + P(|x - \mu_1| < |x - \mu_2| | w_2) P(w_2)$$

$$= P(x > \frac{\mu_1 + \mu_2}{2} | w_1) P(w_1) + P(x < \frac{\mu_1 + \mu_2}{2} | w_2) P(w_2)$$

$$= P(x > \frac{\mu_1 + \mu_2}{2} | w_1)$$

$$= \frac{1}{\sqrt{2\pi}\sigma} \int_{\frac{\mu_1 + \mu_2}{2}}^{\infty} e^{-\frac{1}{2}(\frac{x - \mu_1}{\sigma})^2} dx \quad (\text{令 } u = \frac{x - \mu_1}{\sigma})$$

$$= \frac{1}{\sqrt{2\pi}} \int_{\frac{\mu_2 - \mu_1}{2\sigma}}^{\infty} e^{-\frac{u^2}{2}} \cdot u' du \quad (u' = \frac{1}{\sigma})$$

$$= \frac{1}{\sqrt{2\pi}} \int_a^{\infty} e^{-\frac{u^2}{2}} du \quad (a = \frac{\mu_2 - \mu_1}{2\sigma})$$

$$(b) \text{ 当 } \frac{|\mu_1 - \mu_2|}{\sigma} \rightarrow \infty \text{ 则 } a \rightarrow \infty, \frac{1}{\sqrt{2\pi}a} e^{-\frac{a^2}{2}} \rightarrow 0$$

$$\therefore P_e = \frac{1}{\sqrt{2\pi}a} \int_a^{\infty} e^{-\frac{u^2}{2}} du \leq \frac{1}{\sqrt{2\pi}} \int_a^{\infty} \frac{u}{a} e^{-\frac{u^2}{2}} du = \frac{1}{\sqrt{2\pi}a} e^{-\frac{a^2}{2}} \therefore P_e \rightarrow 0$$

Question 3

$$(a) p(x) = \frac{1}{(2\pi)^{d/2} |\Sigma|^{1/2}} \exp \left[-\frac{1}{2} (x - \mu)^t \Sigma^{-1} (x - \mu) \right]$$

$$\mu = E[x] = \int x p(x) dx$$

$$\Sigma = E[(x - \mu)(x - \mu)^t] = \int (x - \mu)(x - \mu)^t p(x) dx$$

(b) (a) $\Sigma_i = \frac{1}{n} \sum_{i=1}^n \Sigma_i$

$$g_i(x) = x^T W_i x + w_i^T x + w_{i0}$$

$$W_i = -\frac{1}{2} \Sigma_i^{-1} \quad w_i = \Sigma_i^{-1} \mu_i$$

$$w_{i0} = -\frac{1}{2} \mu_i^T \Sigma_i^{-1} \mu_i - \frac{1}{2} \ln |\Sigma_i| + \ln P(w_i)$$

(b) $\Sigma_i = \Sigma$

$$g_i(x) = -\frac{1}{2} (x - \mu_i)^T \Sigma^{-1} (x - \mu_i) + \ln P(w_i)$$

$$g_i(x) = w_i^T x + w_{i0}$$

$$w_i^T = \Sigma^{-1} \mu_i$$

$$w_{i0} = -\frac{1}{2} \mu_i^T \Sigma^{-1} \mu_i + \ln P(w_i)$$

(c) 求矩阵的伪逆、特征降维、扩大非相关样本等。

Question 4

判别面 $x_0 = \frac{1}{2} (\mu_1 + \mu_2) - \frac{\ln [P(w_1) / P(w_2)]}{(\mu_1 - \mu_2)^T \Sigma^{-1} (\mu_1 - \mu_2)} (\mu_1 - \mu_2)$ 假设 $\mu_1 < \mu_2$

当 $x_0 < \mu_1$ 时: $\ln [P(w_1) / P(w_2)] > -\frac{1}{2} (\mu_1 - \mu_2)^T \Sigma^{-1} (\mu_1 - \mu_2)$

当 $x_0 > \mu_2$ 时: $\ln [P(w_1) / P(w_2)] < \frac{1}{2} (\mu_2 - \mu_2)^T \Sigma^{-1} (\mu_2 - \mu_2)$

Question 5

依题意 $z_{ik} = \begin{cases} 1 & \text{第 } k \text{ 个样本的取态状态为 } w_i \\ 0 & \text{其他} \end{cases}$

(a) $Pr [z_{ik}=1 | P(w_i)] = P(w_i)$ 统一为 $P[z_{ik} | P(w_i)] = [P(w_i)]^{z_{ik}} [1 - P(w_i)]^{1-z_{ik}}$

$Pr [z_{ik}=0 | P(w_i)] = 1 - P(w_i)$

连续独立抽取 n 个 z :

$$P[z_{i1}, z_{i2}, \dots, z_{in} | P(w_i)] = \prod_{k=1}^n P[z_{ik} | P(w_i)] = \prod_{k=1}^n [P(w_i)]^{z_{ik}} [1 - P(w_i)]^{1-z_{ik}}$$

(b) $P(w_i)$ 的对数似然度为

$$L(P(w_i)) = \ln \{P[z_{i1}, \dots, z_{in} | P(w_i)]\} = \ln \left\{ \prod_{k=1}^n [P(w_i)]^{z_{ik}} [1 - P(w_i)]^{1-z_{ik}} \right\}$$

$$= \sum_{k=1}^n \{z_{ik} \ln [P(w_i)] + (1 - z_{ik}) \ln [1 - P(w_i)]\}$$

求最大似然估计则 $\nabla_{P(w_i)} L(P(w_i)) = \frac{1}{P(w_i)} \sum_{k=1}^n z_{ik} - \frac{1}{1 - P(w_i)} \sum_{k=1}^n (1 - z_{ik}) = 0$



中国科学院大学

University of Chinese Academy of Sciences

2020 E801782032

薛亮

$$\begin{aligned} \text{即 } \frac{1}{\hat{p}(w_i)} \sum_{k=1}^n z_{ik} &= \frac{1}{1 - \hat{p}(w_i)} \sum_{k=1}^n (1 - z_{ik}) \\ (1 - \hat{p}(w_i)) \sum_{k=1}^n z_{ik} &= n \hat{p}(w_i) - \hat{p}(w_i) \sum_{k=1}^n z_{ik} \\ \hat{p}(w_i) &= \frac{1}{n} \sum_{k=1}^n z_{ik} \end{aligned}$$

Question 6.

$$\begin{aligned} (a) \quad \hat{\mu}_{n+1} &= \frac{1}{n+1} \sum_{k=1}^{n+1} x_k = \frac{1}{n+1} \left(\sum_{k=1}^n x_k + x_{n+1} \right) = \frac{n}{n+1} \left(\frac{1}{n} \sum_{k=1}^n x_k \right) + \frac{1}{n+1} x_{n+1} \\ &= \frac{n+1-1}{n+1} \hat{\mu}_n + \frac{1}{n+1} x_{n+1} = \hat{\mu}_n + \frac{1}{n+1} (x_{n+1} - \hat{\mu}_n) \\ C_{n+1} &= \frac{1}{n} \sum_{k=1}^{n+1} (x_k - \hat{\mu}_{n+1})(x_k - \hat{\mu}_{n+1})^T \\ &= \frac{1}{n} \sum_{k=1}^n (x_k - \hat{\mu}_{n+1})(x_k - \hat{\mu}_{n+1})^T + \frac{1}{n} [x_{n+1} - \hat{\mu}_n - \frac{1}{n+1}(x_{n+1} - \hat{\mu}_n)] [x_{n+1} - \hat{\mu}_n - \frac{1}{n+1}(x_{n+1} - \hat{\mu}_n)]^T \\ &= \frac{1}{n} \sum_{k=1}^n [x_k - \hat{\mu}_n - \frac{1}{n+1}(x_{n+1} - \hat{\mu}_n)] [x_k - \hat{\mu}_n - \frac{1}{n+1}(x_{n+1} - \hat{\mu}_n)]^T \\ &\quad + \frac{n}{(n+1)^2} (x_{n+1} - \hat{\mu}_n)(x_{n+1} - \hat{\mu}_n)^T \quad \begin{matrix} 0 \\ // \end{matrix} \\ &= \frac{1}{n} \left[\sum_{k=1}^n (x_k - \hat{\mu}_n)(x_k - \hat{\mu}_n)^T - \frac{1}{n+1} \sum_{k=1}^n (x_k - \hat{\mu}_n)(x_{n+1} - \hat{\mu}_n)^T - \frac{1}{n+1} (x_{n+1} - \hat{\mu}_n) \sum_{k=1}^n (x_k - \hat{\mu}_n)^T \right. \\ &\quad \left. + \frac{1}{(n+1)^2} \sum_{k=1}^n (x_{n+1} - \hat{\mu}_n)(x_{n+1} - \hat{\mu}_n)^T \right] + \frac{n}{(n+1)^2} (x_{n+1} - \hat{\mu}_n)(x_{n+1} - \hat{\mu}_n)^T \\ &= \frac{n-1}{n} \cdot \frac{1}{n-1} \sum_{k=1}^n (x_k - \hat{\mu}_n)(x_k - \hat{\mu}_n)^T + \left[\frac{1}{n} \cdot n \cdot \frac{1}{(n+1)^2} + \frac{n}{(n+1)^2} \right] (x_{n+1} - \hat{\mu}_n)(x_{n+1} - \hat{\mu}_n)^T \\ &= \frac{n-1}{n} C_n + \frac{1}{n+1} (x_{n+1} - \hat{\mu}_n)(x_{n+1} - \hat{\mu}_n)^T \end{aligned}$$

(b) 对于 $\hat{\mu}_n$ 复杂度 $O(dn)$; 对 C_n 复杂度 $O(dn^2)$

年 月 日