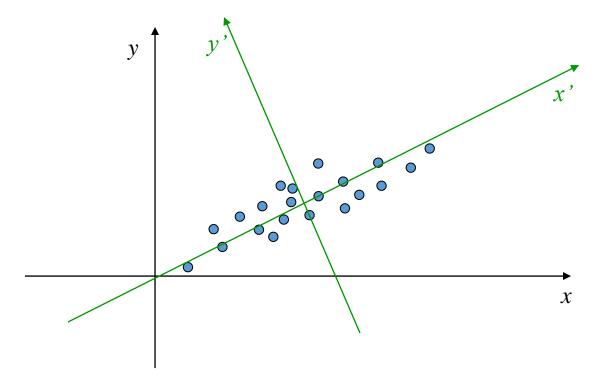
Dimensionality Reduction

Yu-Shuen Wang, CS, NCTU

Principal component analysis

PCA – the general idea

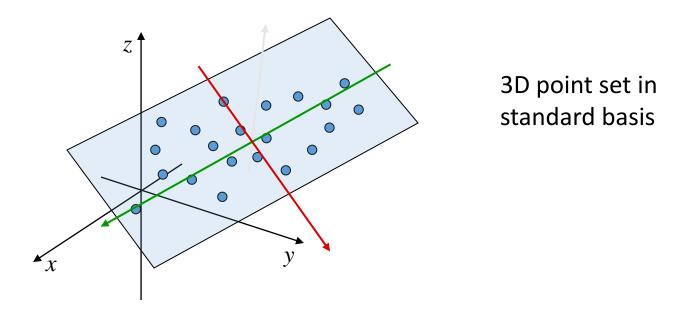
• PCA finds an orthogonal basis that best represents given data set.



• The sum of distances² from the x' axis is minimized.

PCA – the general idea

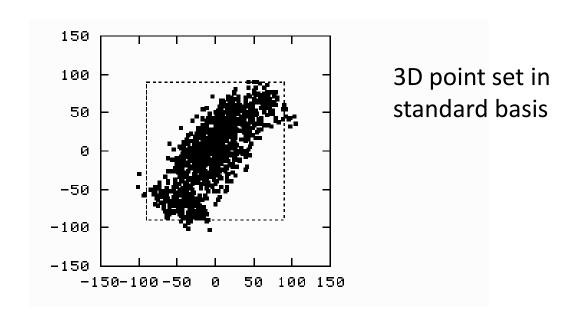
• PCA finds an orthogonal basis that best represents given data set.



• PCA finds a best approximating plane (again, in terms of $\Sigma distances^2$)

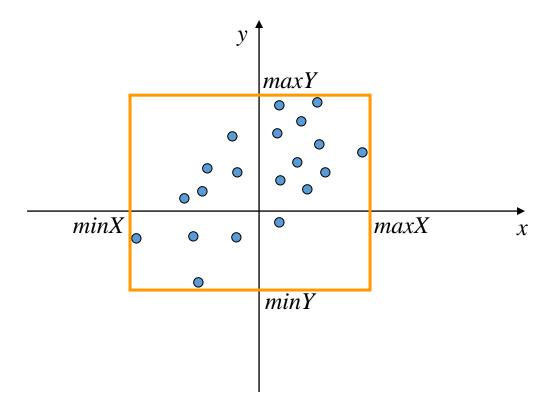
PCA – the general idea

• PCA finds an orthogonal basis that best represents given data set.

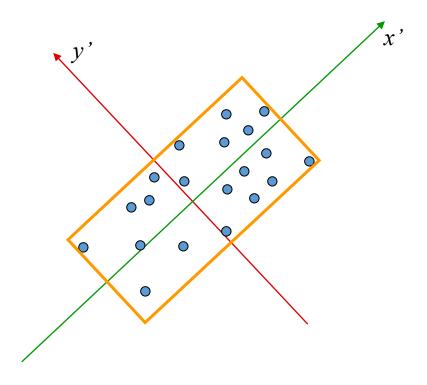


• PCA finds a best approximating plane (again, in terms of $\Sigma distances^2$)

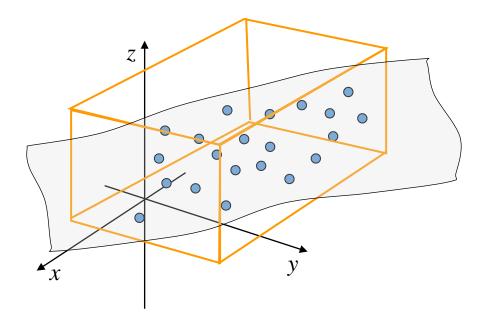
• An axis-aligned bounding box: agrees with the axes



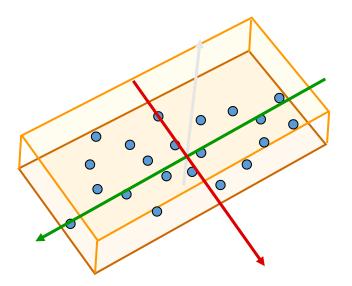
Oriented bounding box: we find better axes!



This is not the optimal bounding box



Oriented bounding box: we find better axes!



Notations

• Denote our data points by \mathbf{x}_1 , \mathbf{x}_2 , ..., $\mathbf{x}_n \in R^d$

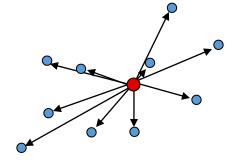
$$\mathbf{x_1} = \begin{pmatrix} x_1^1 \\ x_1^2 \\ \vdots \\ x_1^d \end{pmatrix}, \quad \mathbf{x_2} = \begin{pmatrix} x_2^1 \\ x_2^2 \\ \vdots \\ x_2^d \end{pmatrix}, \quad \dots, \quad \mathbf{x_n} = \begin{pmatrix} x_n^1 \\ x_n^2 \\ \vdots \\ x_n^d \end{pmatrix}$$

The origin of the new axes

- The origin is zero-order approximation of our data set (a point)
- It will be the center of mass:

$$\mathbf{m} = \frac{1}{n} \sum_{i=1}^{n} \mathbf{x}_{i}$$

• It can be shown that:



$$\mathbf{m} = \underset{\mathbf{x}}{\operatorname{argmin}} \sum_{i=1}^{n} \|\mathbf{x}_{i} - \mathbf{x}\|^{2}$$

Scatter matrix

• Denote $\mathbf{y}_i = \mathbf{x}_i - \mathbf{m}$, i = 1, 2, ..., n $S = YY^T$

where Y is $d \times n$ matrix with \mathbf{y}_k as columns (k = 1, 2, ..., n)

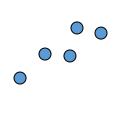
$$S = \begin{pmatrix} y_1^1 & y_2^1 & \cdots & y_n^1 \\ y_1^2 & y_2^2 & & y_n^2 \\ \vdots & \vdots & & \vdots \\ y_1^d & y_2^d & \cdots & y_n^d \end{pmatrix} \begin{pmatrix} y_1^1 & y_1^2 & \cdots & y_1^d \\ y_1^1 & y_2^2 & \cdots & y_2^d \\ \vdots & & \vdots & & \vdots \\ y_n^1 & y_n^2 & \cdots & y_n^d \end{pmatrix}$$

$$\mathbf{V} \qquad \mathbf{V}^T$$

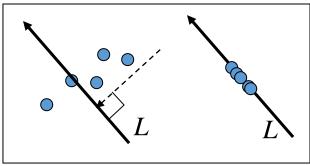
Variance of projected points

- In a way, S measures variance (= scatterness) of the data in different directions.
- Let's look at a line L through the center of mass \mathbf{m} , and project our points \mathbf{x}_i onto it. The variance of the projected points \mathbf{x}_i is:

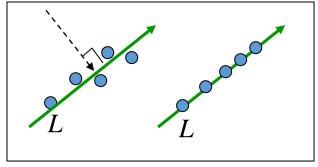
$$\operatorname{var}(L) = \frac{1}{n} \sum_{i=1}^{n} ||\mathbf{x}_{i}' - \mathbf{m}||^{2}$$



Original set



Small variance

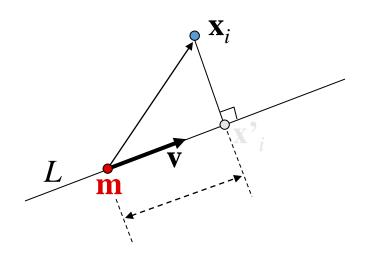


Large variance

Variance of projected points

• Given a direction \mathbf{v} , $||\mathbf{v}|| = 1$, the projection of \mathbf{x}_i onto $L = \mathbf{m} + \mathbf{v}t$ is:

$$\left| || \mathbf{x}_{i}' - \mathbf{m} || = \langle \mathbf{v}, \mathbf{x}_{i} - \mathbf{m} \rangle / || \mathbf{v} || = \langle \mathbf{v}, \mathbf{y}_{i} \rangle = \mathbf{v}^{T} \mathbf{y}_{i} \right|$$



Variance of projected points

So,

$$var(L) = \frac{1}{n} \sum_{i=1}^{n} || \mathbf{x}_{i}' - \mathbf{m} ||^{2} = \frac{1}{n} \sum_{i=1}^{n} (\mathbf{v}^{T} \mathbf{y}_{i})^{2} = \frac{1}{n} || \mathbf{v}^{T} Y ||^{2} =$$

$$= \frac{1}{n} || Y^{T} \mathbf{v} ||^{2} = \frac{1}{n} \langle Y^{T} \mathbf{v}, Y^{T} \mathbf{v} \rangle = \frac{1}{n} \mathbf{v}^{T} Y Y^{T} \mathbf{v} = \frac{1}{n} \mathbf{v}^{T} S \mathbf{v} = \frac{1}{n} \langle S \mathbf{v}, \mathbf{v} \rangle$$

$$\sum_{i=1}^{n} (\mathbf{v}^{T} \mathbf{y_{i}})^{2} = \sum_{i=1}^{n} \left((v^{1} \quad v^{2} \quad \cdots \quad v^{d}) \begin{pmatrix} y_{i}^{1} \\ y_{i}^{2} \\ \vdots \\ y_{i}^{d} \end{pmatrix} \right)^{2} = \left\| (v^{1} \quad v^{2} \quad \cdots \quad v^{d}) \begin{pmatrix} y_{1}^{1} \quad y_{2}^{1} \quad \cdots \quad y_{n}^{1} \\ y_{1}^{2} \quad y_{2}^{2} \quad \cdots \quad y_{n}^{2} \\ \vdots \quad \vdots \quad & \vdots \\ y_{1}^{d} \quad y_{2}^{d} \quad \cdots \quad y_{n}^{d} \end{pmatrix} \right\|^{2} = \left\| (\mathbf{v}^{T} \mathbf{y}^{T}) \right\|^{2}$$

Directions of maximal variance

- So, we have: $var(L) = \langle Sv, v \rangle$
- Theorem:

```
Let f: \{\mathbf{v} \in R^d \ / \ //\mathbf{v}// = 1\} \to R, f(\mathbf{v}) = \langle S\mathbf{v}, \mathbf{v} \rangle (and S is a symmetric matrix).
```

Then, the extrema of f are attained at the eigenvectors of S.

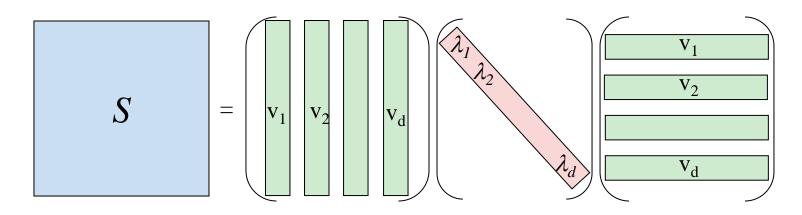
So, eigenvectors of S are directions of maximal/minimal variance!

Summary so far

- We take the centered data vectors \mathbf{y}_1 , \mathbf{y}_2 , \mathbf{y}_2 , \mathbf{y}_2 , \mathbf{y}_2 , \mathbf{y}_2 , \mathbf{y}_3
- Construct the scatter matrix
- S measures the variance of the data points
- ullet Eigenvectors of S are directions of maximal variance.

Scatter matrix - eigendecomposition

- S is symmetric
- \Rightarrow S has eigendecomposition: $S = VAV^{T}$



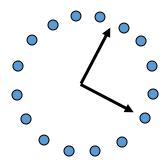
The eigenvectors form orthogonal basis

Principal components

- Eigenvectors that correspond to big eigenvalues are the directions in which the data has strong components (= large variance).
- If the eigenvalues are more or less the same there is no preferable direction.

Note: the eigenvalues are always non-negative. Think why...

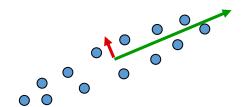
Principal components



- There's no preferable direction
- *S* looks like this:

$$Vegin{pmatrix} \lambda & & \ & \lambda \end{pmatrix} V^T$$

• Any vector is an eigenvector



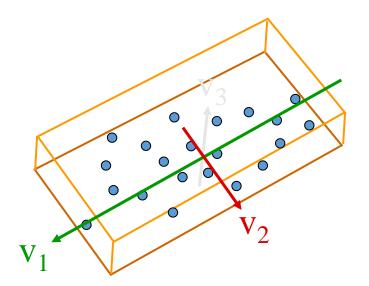
- There is a clear preferable direction
- S looks like this:

$$V \begin{pmatrix} \lambda & \\ & \mu \end{pmatrix} V^T$$

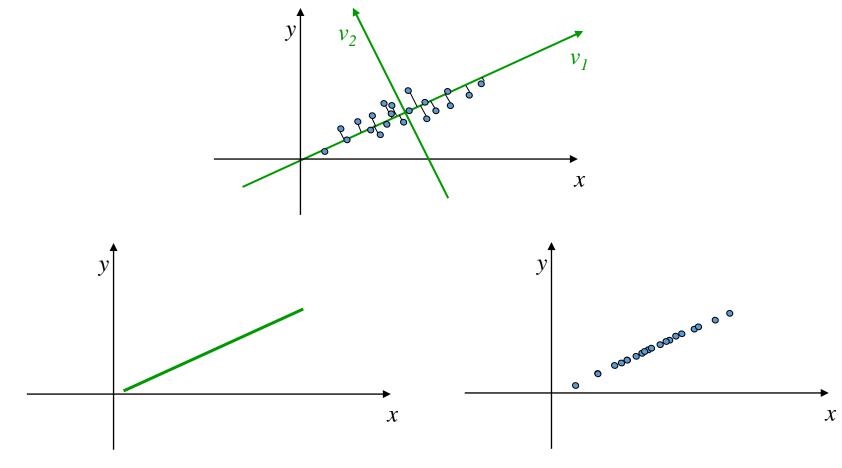
 μ is close to zero, much smaller than λ .

How to use what we got

• For finding oriented bounding box – we simply compute the bounding box with respect to the axes defined by the eigenvectors. The origin is at the mean point \mathbf{m} .



For approximation



This line segment approximates the original data set

The projected data set approximates the original data set

For approximation

• In general dimension d, the eigenvalues are sorted in descending order:

$$\lambda_1 \geq \lambda_2 \geq \dots \geq \lambda_d$$

- The eigenvectors are sorted accordingly.
- To get an approximation of dimension d' < d, we take the d' first eigenvectors and look at the subspace they span (d' = 1 is a line, d' = 2 is a plane...)

For approximation

• To get an approximating set, we project the original data points onto the chosen subspace:

$$\mathbf{x}_i = \mathbf{m} + \alpha_1 \mathbf{v}_1 + \alpha_2 \mathbf{v}_2 + \dots + \alpha_d \mathbf{v}_d + \dots + \alpha_d \mathbf{v}_d$$

Projection:

$$\mathbf{x}_{i}' = \mathbf{m} + \alpha_{1}\mathbf{v}_{1} + \alpha_{2}\mathbf{v}_{2} + \dots + \alpha_{d}\mathbf{v}_{d} + \mathbf{0}\mathbf{v}_{d} + \dots + \mathbf{0}\mathbf{v}_{d} + \dots + \mathbf{0}\mathbf{v}_{d}$$

Technical remarks:

- $\lambda_i \ge 0$, i = 1,...,d (such matrices are called positive semi-definite). So we can indeed sort by the magnitude of λ_i
- Theorem: $\lambda_i \ge 0 \iff \langle S\mathbf{v}, \, \mathbf{v} \rangle \ge 0 \quad \forall \, \mathbf{v}$ Proof:

$$S = V\Lambda V^{T} \Rightarrow \langle S\boldsymbol{v}, \boldsymbol{v} \rangle = \boldsymbol{v}^{T} S\boldsymbol{v} = \boldsymbol{v}^{T} V\Lambda V^{T} \boldsymbol{v}$$

= $(V^{T}\boldsymbol{v})^{T} \Lambda(\boldsymbol{v}^{T} V) = \boldsymbol{u}^{T} \Lambda \boldsymbol{u} = \langle \Lambda \boldsymbol{u}, \boldsymbol{u} \rangle$

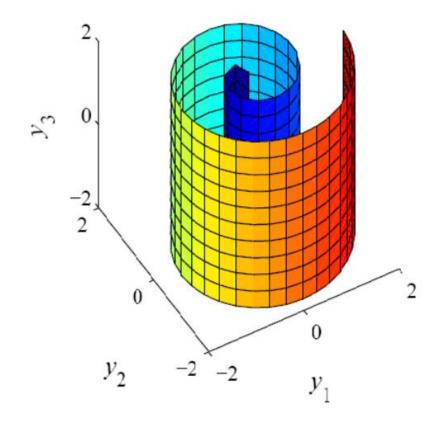
$$< S\mathbf{v}, \mathbf{v} > = \lambda_1 \mathbf{u_1}^2 + \lambda_2 \mathbf{u_2}^2 + \dots + \lambda_d \mathbf{u_d}^2$$

Therefore, $\lambda_i \ge 0 \iff \langle S\mathbf{v}, \mathbf{v} \rangle \ge 0 \quad \forall \mathbf{v}$

Visualizing Data using t-SNE

Dimensionality Reduction is a helpful tool for visualization

- Dimensionality reduction algorithms
 - Map high-dimensional data to a lower dimension
 - While preserving structure
- They are used for
 - Visualization
 - Performance
 - Curse of dimensionality
- A ton of algorithms exist
 - t-SNE is specialised for visualization and has gained a lot of popularity



Dimensionality Reduction techniques solve optimization problems

$$X = \{x_1, x_2, \dots, x_n \in \mathbb{R}^h\} \Rightarrow Y = \{y_1, y_2, \dots, y_n \in \mathbb{R}^l\}$$

$$\min_{\mathcal{Y}} C(X, Y)$$

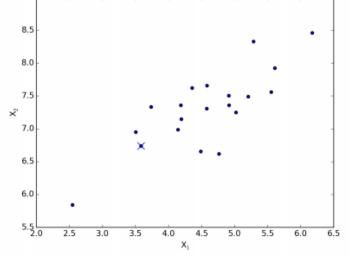
- Three approaches for Dimensionality Reduction
 - Distance preserving
 - Topology preserving
 - Information preserving
- t-SNE is distance-based but tends to preserve topology

SNE computes pair-wise similarities

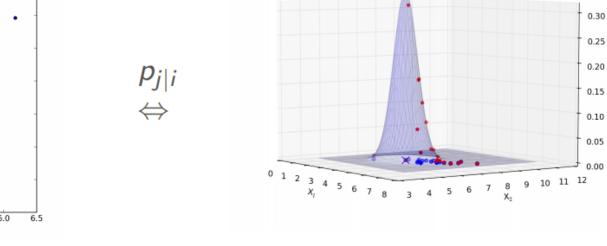
$$p_{j|i} = \frac{exp(-\|x_i - x_j\|^2 / 2\sigma_i^2)}{\sum_{k!=i} exp(-\|x_i - x_k\|^2 / 2\sigma_i^2)}, \quad q_{j|i} = \frac{exp(-\|y_i - y_j\|^2)}{\sum_{k!=i} exp(-\|y_i - y_k\|^2)}$$

$$p_{i|i}=0, \quad q_{i|i}=0$$

 SNE converts euclidean distances to similarities, that can be interpreted as probabilities

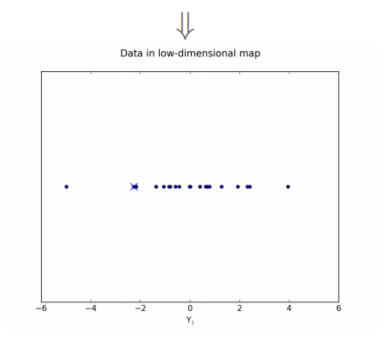


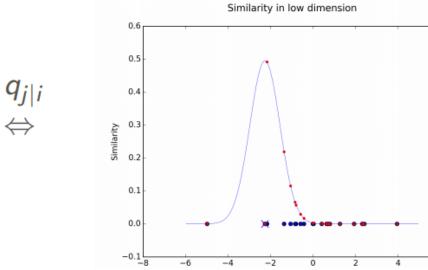
Data in high-dimensional space



Similarity in high dimension







Kullback-Leiber Divergence measures the faithfulness with wich $q_{j|i}$ models $p_{j|i}$

- $P_i=\{p_{1|i},p_{2|i},\dots,p_{n|i}\}$ and $Q_i=\{q_{1|i},q_{2|i},\dots,q_{n|i}\}$ are the distribution on the neighbors of point i.
- Kullback-Leiber Divergence (KL) compares two distributions.

$$C = \sum_{i} KL(P_i||Q_i) = \sum_{i} \sum_{j} p_{j|i} log \frac{p_{j|i}}{q_{j|i}}$$

- KL divergence is asymmetric
- KL divergence is always positive.
- We have our minimization problem: $\min_{v} C(X, Y)$

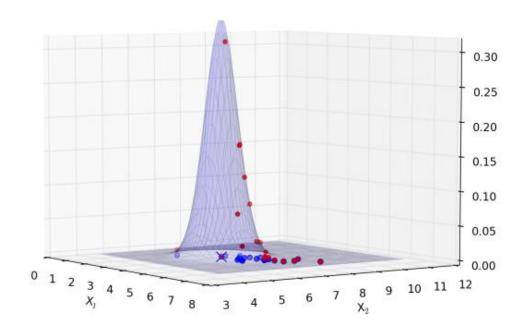
Why radial basis function (exponential)?

$$p_{j|i} = \frac{exp(-\|x_i - x_j\|^2 / 2\sigma_i^2)}{\sum_{k!=i} exp(-\|x_i - x_k\|^2 / 2\sigma_i^2)}, \quad q_{j|i} = \frac{exp(-\|y_i - y_j\|^2)}{\sum_{k!=i} exp(-\|y_i - y_k\|^2)}$$

Similarity in high dimension

Focus on local geometry.

This is why t-SNE can be interpreted as topology-based

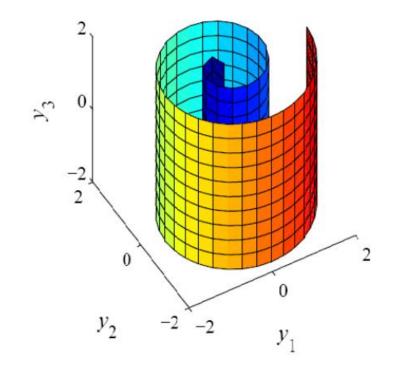


Why probabilities?

$$p_{j|i} = \frac{exp(-\|x_i - x_j\|^2 / 2\sigma_i^2)}{\sum_{k!=i} exp(-\|x_i - x_k\|^2 / 2\sigma_i^2)}, \quad q_{j|i} = \frac{exp(-\|y_i - y_j\|^2)}{\sum_{k!=i} exp(-\|y_i - y_k\|^2)}$$

Small distance does not mean proximity on manifold.

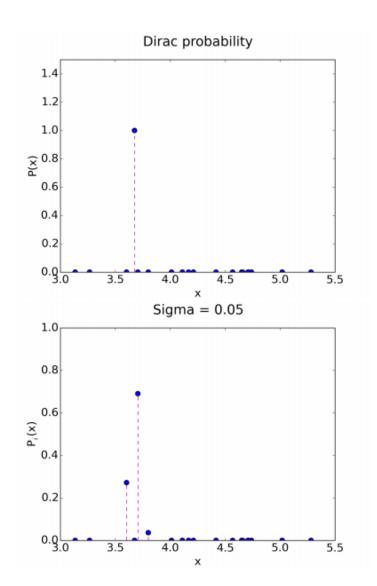
Probabilities are appropriate to model this uncertainty.

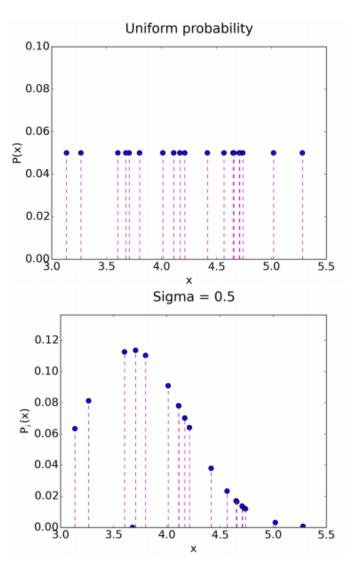


How do you choose σ_i ?

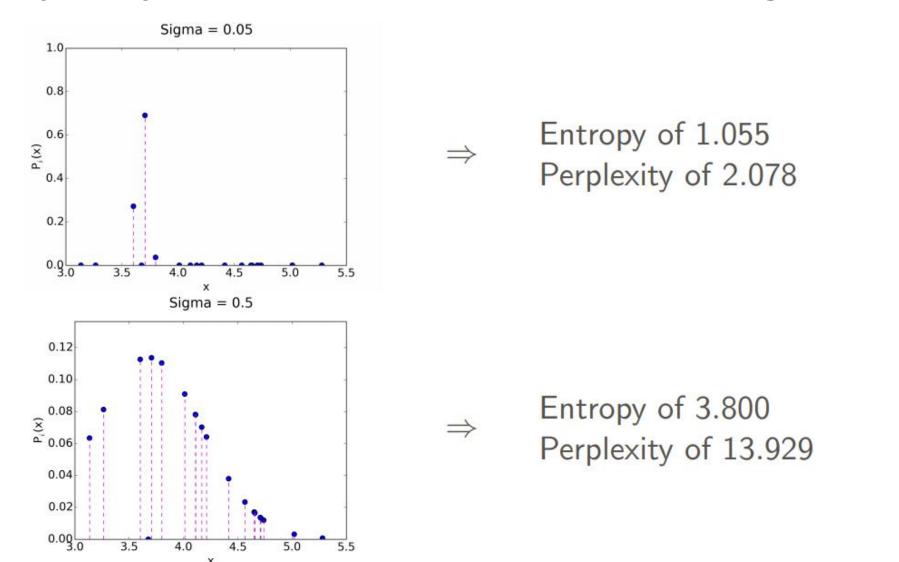
$$H(P) = -\sum_{i} p_i \log p_i$$

The entropy of P_i increases with σ_i





Perplexity, a smooth measure of the # of neighbors.



From SNE to t-SNE

SNE

Modelisation

$$p_{j|i} = \frac{exp(-\|x_i - x_j\|^2 / 2\sigma_i^2)}{\sum_{k!=i} exp(-\|x_i - x_k\|^2 / 2\sigma_i^2)},$$

$$q_{j|i} = \frac{exp(-\|y_i - y_j\|^2)}{\sum_{k!=i} exp(-\|y_i - y_k\|^2)}$$

Cost Function

$$C = \sum_{i} KL(P_i||Q_i)$$

Derivatives

$$\frac{dC}{dy_i} = 2\sum_{j} (p_{j|i} - q_{j|i} + p_{i|j} - q_{i|j})(y_i - y_j) \qquad \frac{dC}{dy_i} = 4\sum_{j} (p_{ij} - q_{ij})(y_i - y_j)$$

Symmetric SNE

Modelisation

$$p_{ij} = \frac{p_{j|i} + p_{i|j}}{2n},$$

$$q_{ij} = \frac{exp(-\|y_i - y_j\|^2)}{\sum_{k!=i} exp(-\|y_i - y_k\|^2)}$$

Cost Function

$$C = \sum KL(P||Q)$$

Derivatives

$$\frac{dC}{dy_i} = 4\sum_j (p_{ij} - q_{ij})(y_i - y_j)$$

t-SNE

Modelisation

$$p_{ij} = \frac{p_{j|i} + p_{i|j}}{2n},$$

$$q_{ij} = \frac{\left(1 + \|y_i - y_j\|^2\right)^{-1}}{\sum_{k!=i} (1 + \|y_i - y_k\|^2)^{-1}}$$

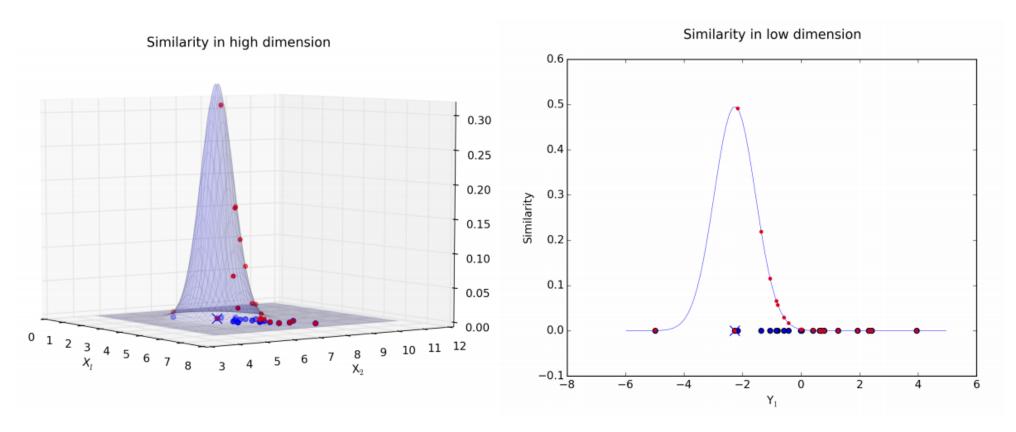
Cost Function

$$C = \sum KL(P||Q)$$

Derivatives

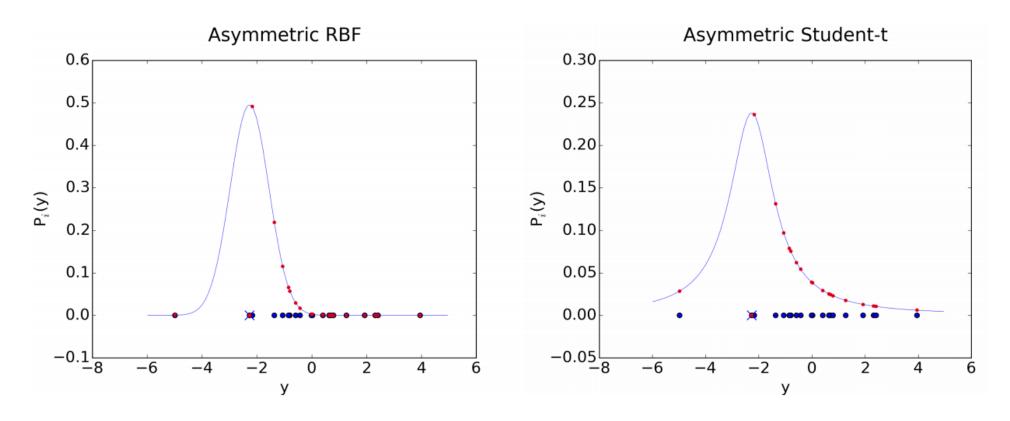
$$\frac{dC}{dy_i} = 4\sum_{j} (p_{ij} - q_{ij})(y_i - y_j) (1 + ||y_i - y_j||^2)^{-1}$$

The "Crowding problem"



• There is much more space in high dimensions.

Mismatched Tails can Compensate for Mismatched Dimensionalities



• Student-t distribution has heavier tails.

Last but not least: Optimization

$$\min_{\mathcal{Y}} C(X,Y)$$

$$C = \sum_{i} KL(P_i||Q_i) = \sum_{i} \sum_{j} p_{j|i} log \frac{p_{j|i}}{q_{j|i}}$$

- Non-convex
- Gradient descent + Adaptive learning rate + Momentum

•
$$y^t = y^{t-1} + \eta \frac{\partial C}{\partial Y} + \alpha (y^{t-1} - y^{t-2})$$

