EC 320: Introduction to Econometrics

Kyle Raze Fall 2019

Prologue

Housekeeping

Thanksgiving

- No lab this week.
- Still have lecture on Wednesday.

Problem Set 5

- Last problem set!
- Due Saturday, December 7 by 11:59pm.

Midterm 2

• Grades posted.

Goal: Make quantitative statements about qualitative information.

• e.g., race, gender, being employed, living in Oregon, etc.

Approach: Construct binary variables.

- a.k.a. dummy variables or indicator variables.
- Value equals 1 if observation is in the category or 0 if otherwise.

Regression implications

- 1. Binary variables change the interpretation of the intercept.
- 2. Coefficients on binary variables have different interpretations than those on continuous variables.

Consider the relationship

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + u_i$$

where

- ullet \mathbf{Pay}_i is a continuous variable measuring an individual's pay
- $School_i$ is a continuous variable that measures years of education

Interpretation

- β_0 : y-intercept, i.e., Pay when School = 0
- β_1 : expected increase in Pay for a one-unit increase in School

Consider the relationship

$$\mathrm{Pay}_i = \beta_0 + \beta_1 \mathrm{School}_i + u_i$$

Derive the slope's interpretation:

$$egin{aligned} \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell+1] - \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell] \ &= \mathbb{E}[eta_0 + eta_1(\ell+1) + u] - \mathbb{E}[eta_0 + eta_1\ell + u] \ &= [eta_0 + eta_1(\ell+1)] - [eta_0 + eta_1\ell] \ &= eta_0 - eta_0 + eta_1\ell - eta_1\ell + eta_1 \ &= eta_1. \end{aligned}$$

The slope gives the expected increase in pay for an additional year of schooling.

Consider the relationship

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + u_i$$

Alternative derivation

Differentiate the model with respect to schooling:

$$\frac{d\text{Pay}}{d\text{School}} = \beta_1$$

The slope gives the expected increase in pay for an additional year of schooling.

If we have multiple explanatory variables, e.g.,

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + \beta_2 \text{Ability}_i + u_i$$

then the interpretation changes slightly.

$$egin{aligned} \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell + 1 \wedge \operatorname{Ability} = lpha] - \mathbb{E}[\operatorname{Pay}|\operatorname{School} &= \ell \wedge \operatorname{Ability} = lpha] \\ &= \mathbb{E}[eta_0 + eta_1(\ell+1) + eta_2lpha + u] - \mathbb{E}[eta_0 + eta_1\ell + eta_2lpha + u] \\ &= [eta_0 + eta_1(\ell+1) + eta_2lpha] - [eta_0 + eta_1\ell + eta_2lpha] \\ &= eta_0 - eta_0 + eta_1\ell - eta_1\ell + eta_1 + eta_2lpha - eta_2lpha \\ &= eta_1 \end{aligned}$$

The slope gives the expected increase in pay for an additional year of schooling, **holding ability constant**.

If we have multiple explanatory variables, e.g.,

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + \beta_2 \text{Ability}_i + u_i$$

then the interpretation changes slightly.

Alternative derivation

Differentiate the model with respect to schooling:

$$\frac{\partial \text{Pay}}{\partial \text{School}} = \beta_1$$

The slope gives the expected increase in pay for an additional year of schooling, **holding ability constant**.

Consider the relationship

$$\mathrm{Pay}_i = eta_0 + eta_1 \mathrm{Female}_i + u_i$$

where Pay_i is a continuous variable measuring an individual's pay and $Female_i$ is a binary variable equal to 1 when i is female.

Interpretation

 β_0 is the expected Pay for males (*i.e.*, when Female = 0):

$$egin{aligned} \mathbb{E}[\operatorname{Pay}|\operatorname{Male}] \ &= \mathbb{E}[eta_0 + eta_1 imes 0 + u_i] \ &= \mathbb{E}[eta_0 + 0 + u_i] \ &= eta_0 \end{aligned}$$

Consider the relationship

$$\text{Pay}_i = \beta_0 + \beta_1 \text{Female}_i + u_i$$

where Pay_i is a continuous variable measuring an individual's pay and $Female_i$ is a binary variable equal to 1 when i is female.

Interpretation

 β_1 is the expected difference in Pay between females and males:

$$egin{aligned} \mathbb{E}[ext{Pay}| ext{Female}] &- \mathbb{E}[ext{Pay}| ext{Male}] \ &= \mathbb{E}[eta_0 + eta_1 imes 1 + u_i] - \mathbb{E}[eta_0 + eta_1 imes 0 + u_i] \ &= \mathbb{E}[eta_0 + eta_1 + u_i] - \mathbb{E}[eta_0 + 0 + u_i] \ &= eta_0 + eta_1 - eta_0 \ &= eta_1 \end{aligned}$$

Consider the relationship

$$\mathrm{Pay}_i = eta_0 + eta_1 \mathrm{Female}_i + u_i$$

where Pay_i is a continuous variable measuring an individual's pay and $Female_i$ is a binary variable equal to 1 when i is female.

Interpretation

 $\beta_0 + \beta_1$: is the expected Pay for females:

$egin{aligned} \mathbb{E}[ext{Pay}| ext{Female}] \ &= \mathbb{E}[eta_0 + eta_1 imes 1 + u_i] \ &= \mathbb{E}[eta_0 + eta_1 + u_i] \ &= eta_0 + eta_1 \end{aligned}$

Consider the relationship

$$\mathrm{Pay}_i = \beta_0 + \beta_1 \mathrm{Female}_i + u_i$$

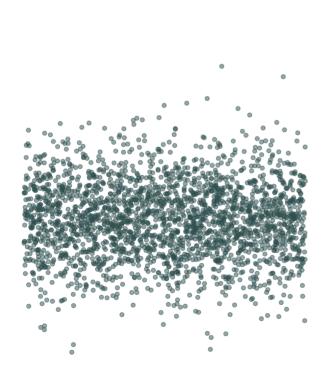
Interpretation

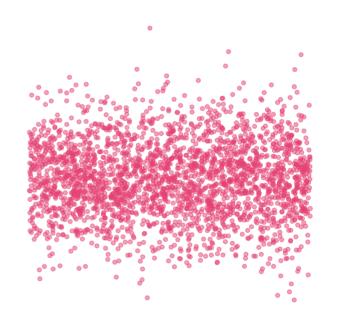
- β_0 : expected Pay for males (*i.e.*, when Female = 0)
- β_1 : expected difference in Pay between females and males
- $\beta_0 + \beta_1$: expected Pay for females
- Males are the reference group

Note: If there are no other variables to condition on, then $\hat{\beta}_1$ equals the difference in group means, e.g., $\bar{X}_{\text{Female}} - \bar{X}_{\text{Male}}$.

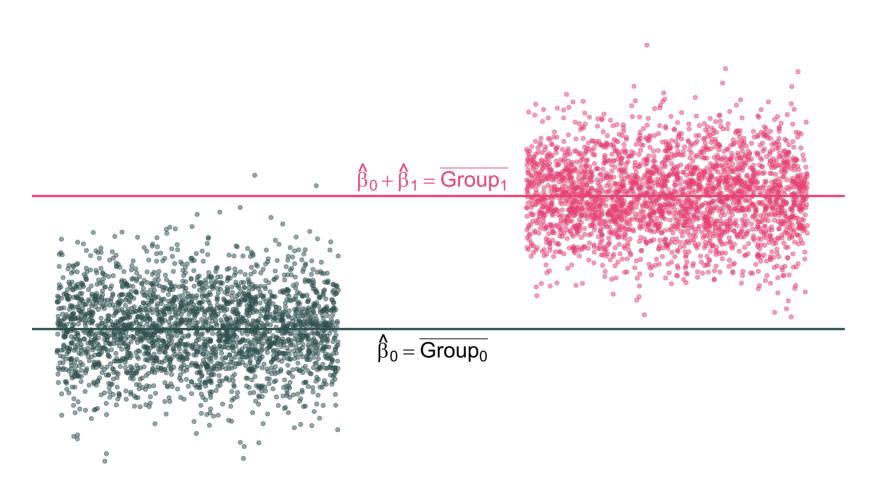
Note₂: The *holding all other variables constant* interpretation also applies for categorical variables in multiple regression settings.

$$Y_i = \beta_0 + \beta_1 X_i + u_i$$
 for binary variable $X_i = \{0, 1\}$



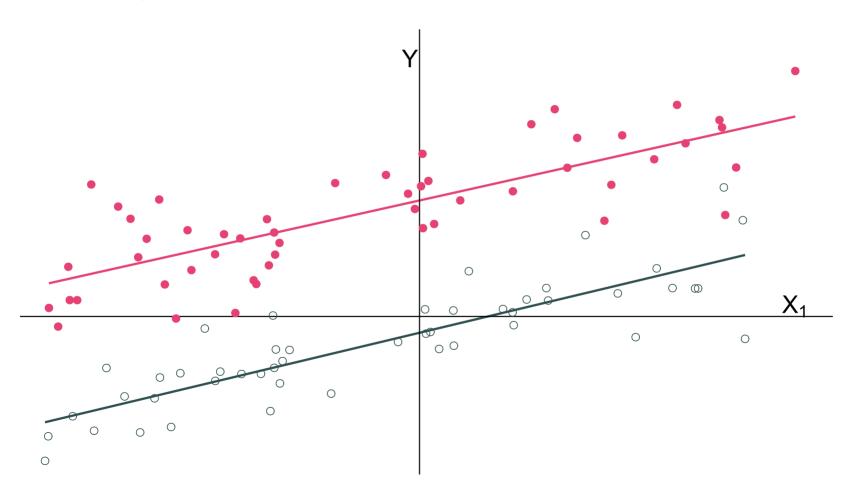


$$Y_i = \beta_0 + \beta_1 X_i + u_i$$
 for binary variable $X_i = \{0, 1\}$



Multiple Regression

Another way to think about it:



Question: Why not estimate $Pay_i = \beta_0 + \beta_1 Female_i + \beta_2 Male_i + u_i$?

Answer: The intercept is a perfect linear combination of $Male_i$ and $Female_i$.

- Violates no perfect collinearity assumption.
- OLS can't estimate all three parameters simultaneously.
- Known as dummy variable trap.

Practical solution: Select a reference category and drop its indicator.

Dummy Variable Trap?

Don't worry, R will bail you out if you include perfectly collinear indicators.

Example

Thanks, R.

Omitted variable bias (OVB) arises when we omit a variable that

- 1. Affects the outcome variable Y
- 2. Correlates with an explanatory variable X_i

Biases OLS estimator of β_j .

Example

Let's imagine a simple population model for the amount individual i gets paid

$$\text{Pay}_i = \beta_0 + \beta_1 \text{School}_i + \beta_2 \text{Male}_i + u_i$$

where $School_i$ gives i's years of schooling and $Male_i$ denotes an indicator variable for whether individual i is male.

Interpretation

- β_1 : returns to an additional year of schooling (ceteris paribus)
- β_2 : premium for being male (*ceteris paribus*) If $\beta_2 > 0$, then there is discrimination against women.

Example, continued

From the population model

$$\mathrm{Pay}_i = eta_0 + eta_1 \mathrm{School}_i + eta_2 \mathrm{Male}_i + u_i$$

An analyst focuses on the relationship between pay and schooling, i.e.,

$$egin{aligned} ext{Pay}_i &= eta_0 + eta_1 ext{School}_i + (eta_2 ext{Male}_i + u_i) \ & ext{Pay}_i &= eta_0 + eta_1 ext{School}_i + arepsilon_i \end{aligned}$$

where $arepsilon_i = eta_2 \mathrm{Male}_i + u_i$.

We assumed exogeneity to show that OLS is unbiasedness. But even if $\mathbb{E}[u|X]=0$, it is not necessarily true that $\mathbb{E}[\varepsilon|X]=0$ (false if $\beta_2\neq 0$).

Specifically, $\mathbb{E}[\varepsilon|\mathrm{Male}=1]=\beta_2+\mathbb{E}[u|\mathrm{Male}=1]\neq 0$. Now OLS is biased.

Let's try to see this result graphically.

The true population model:

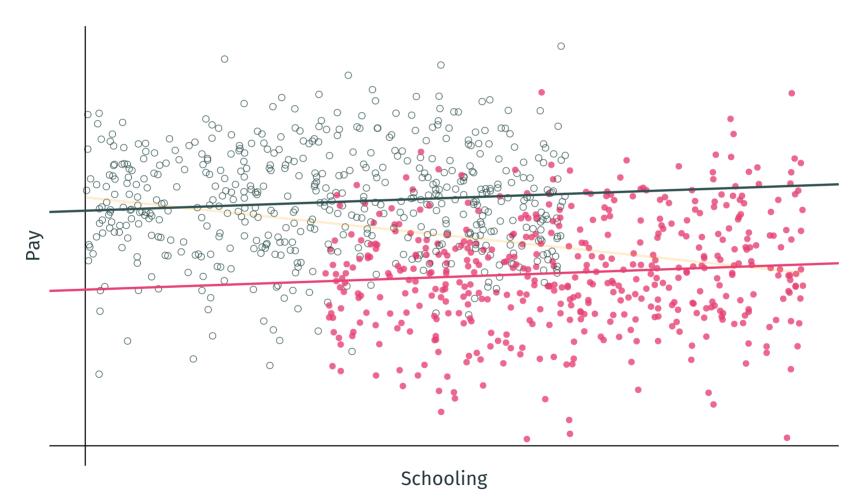
$$\mathrm{Pay}_i = 20 + 0.5 imes \mathrm{School}_i + 10 imes \mathrm{Male}_i + u_i$$

The regression model that suffers from omitted-variable bias:

$$\text{Pay}_i = \hat{\beta}_0 + \hat{\beta}_1 \times \text{School}_i + e_i$$

Finally, imagine that women, on average, receive more schooling than men.

Unbiased regression: $\widehat{\mathrm{Pay}}_i = 20.3 + 0.4 imes \mathrm{School}_i + 10.2 imes \mathrm{Male}_i$



Example: Weekly Wages

Q₁: What is the reference category?

Q₂: Interpret the coefficients.

Q₃: Suppose you ran lm(wage ~ nonsouth, data = wage_data) instead. What is the coefficient estimate on nonsouth? What is the intercept estimate?

Example: Weekly Wages

Q₁: What is the reference category?

Q₂: Interpret the coefficients.

Q₃: Suppose you ran lm(wage ~ south + nonblack, data = wage_data) instead. What is the coefficient estimate on nonblack? What is the coefficient estimate on south? What is the intercept estimate?

Example: Weekly Wages

Answer to Q_3 :