# Clustering: Models and Algorithms

Shikui Tu
Shanghai Jiao Tong University
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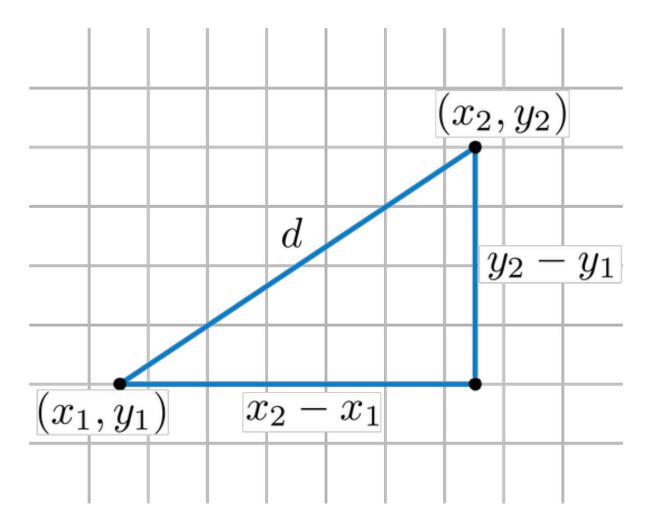
#### Outline

Gaussian Mixture Models (GMM)

Expectation-Maximization (EM) for maximum likelihood

- Gaussian Mixture Models (GMM)
  - From generation process perspective

### **Euclidean Distance**



# Euclidian distance may not be a good measure for some data

# Euclidean distance Equal distance line In general Mahalanobis distance

Distances at different directions could be different!

$$d(\vec{x}, \vec{y}) = \sqrt{(\vec{x} - \vec{y})^T \Sigma^{-1} (\vec{x} - \vec{y})}.$$

 $\Sigma$  is the covariance matrix

### More Distance Measures

| Table 1 Gene expression similarity measures               |  |
|---|--|
| Manhattan distance (city-block distance, L1 norm)         | $d_{fg} = \sum_{c} \left  e_{fc} - e_{gc} \right $   |
| Euclidean distance<br>(L2 norm)                           | $d_{fg} = \sqrt{\sum_{c} \left(e_{fc} - e_{gc}\right)^2}$  |
| Mahalanobis distance                                      | $d_{fg} = (e_f - e_g)^{\text{I}} \Sigma^{-1} (e_f - e_g)$ , where $\Sigma$ is the (full or within-cluster) covariance matrix of the data                                 |
| Pearson correlation (centered correlation)                | $d_{fg} = 1 - r_{fg}$ , with $r_{fg} = \frac{\sum_{c} (e_{fc} - \bar{e}_f)(e_{gc} - \bar{e}_g)}{\sqrt{\sum_{c} (e_{fc} - \bar{e}_f)^2 \sum_{c} (e_{gc} - \bar{e}_g)^2}}$ |
| Uncentered correlation (angular separation, cosine angle) | $d_{fg} = 1 - r_{fg}$ , with $r_{fg} = \frac{\sum_{c} e_{fc} e_{gc}}{\sqrt{\sum_{c} e_{fc}^2 \sum_{c} e_{gc}^2}}$  |
| Spellman rank correlation                                 | As Pearson correlation, but replace $e_{gc}$ with the rank of $e_{gc}$ within the expression values of gene $g$ across all conditions $c=1C$                             |
| Absolute or squared correlation                           | $d_{fg} = 1 -  r_{fg}  \text{ or } d_{fg} = 1 - r_{fg}^2$  |
| $d_{fg}$ distance between expression patterns for genes   | $f$ and $g$ . $e_{gc}$ , expression level of gene $g$ under condition $c$ .  |

### From distance to probability

distance

likely

$$\|x - \mu\|^2$$

$$\longrightarrow \exp\{-\lambda \|x - \mu\|^2\}$$

"The closer, the more likely."

Sum or integral to be one

**Probability** 

It is more powerful to consider everything in probability framework!

$$\frac{1}{\sqrt{(2\pi)^k |\mathbf{\Sigma}|}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^\mathrm{T} \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu})\right)$$

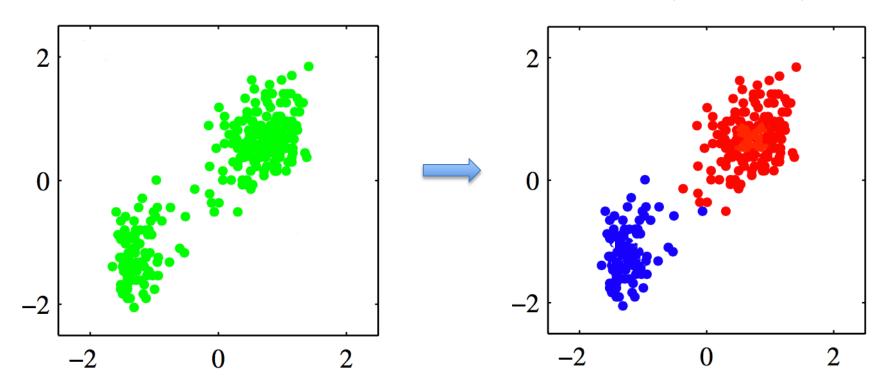
Gaussian distribution with the Mahalanobis distance

$$D_M(x) = \sqrt{(x-\mu)^T \Sigma^{-1}(x-\mu)}.$$

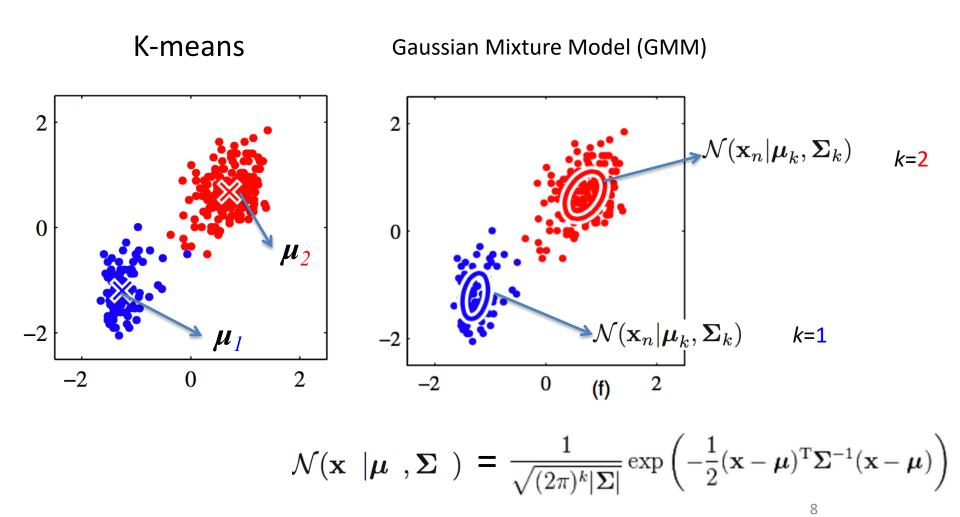
### Review the clustering problem again

We have the following data:

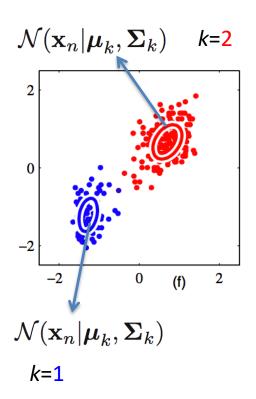
We want to cluster the data into two clusters (red and blue)



# Instead if using $\{\mu_{I_1}, \mu_2\}$ , each cluster is represented as a Gaussian distribution



## Gaussian Mixture Model (GMM)



We use  $z_k = 1$  to indicate a point **x** belongs to cluster k

$$\mathbf{z} = (z_1, ..., z_K)$$
  $z_k \in \{0, 1\}$   $\sum_k z_k = 1$ 

Assume the points in the same cluster follow a **Gaussian distribution** 

$$p(\mathbf{x}|z_k=1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

A mixing weight for each cluster:

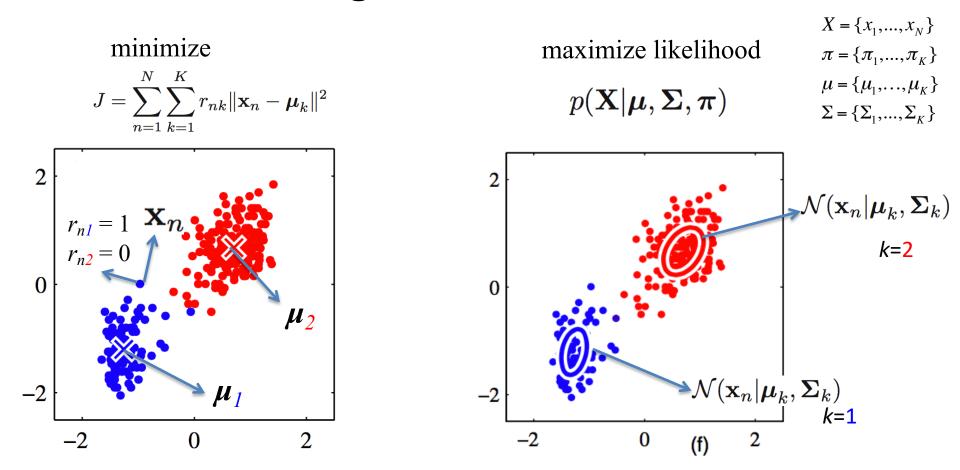
$$p(z_k=1)=\pi_k$$
  $0\leqslant\pi_k\leqslant 1$   $\sum_{k=1}^K\pi_k=1$ 

prior probability of point belonging to a cluster

So, we get a distribution for the data point  $\mathbf{x}$ :

$$p(\mathbf{x}) = \sum_{\mathbf{z}} p(\mathbf{z}) p(\mathbf{x}|\mathbf{z}) = \sum_{k=1}^K \pi_k \mathcal{N}(\mathbf{x}|oldsymbol{\mu}_k, oldsymbol{\Sigma}_k)$$

# From minimizing sum of square distances to finding maximum likelihood



Remember: The closer the distance, the more likely the probability.

#### Maximum likelihood

Given a data set  $\mathbf{X} = (\mathbf{x}_1, \dots, \mathbf{x}_N)^T$  in which the observations  $\{\mathbf{x}_n\}$  are assumed to be drawn independently from a multivariate Gaussian distribution, we can estimate the parameters of the distribution by maximum likelihood. The log likelihood function is given by

$$\ln p(\mathbf{X}|\boldsymbol{\mu},\boldsymbol{\Sigma}) = -\frac{ND}{2}\ln(2\pi) - \frac{N}{2}\ln|\boldsymbol{\Sigma}| - \frac{1}{2}\sum_{n=1}^{N}(\mathbf{x}_n - \boldsymbol{\mu})^{\mathrm{T}}\boldsymbol{\Sigma}^{-1}(\mathbf{x}_n - \boldsymbol{\mu})$$

Maximizing the log-likelihood function:

$$\frac{\partial}{\partial \boldsymbol{\mu}} \ln p(\mathbf{X}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \boldsymbol{\Sigma}^{-1} (\mathbf{x}_n - \boldsymbol{\mu}) = 0 \qquad \longrightarrow \qquad \boldsymbol{\mu}_{\mathrm{ML}} = \frac{1}{N} \sum_{n=1}^{N} \mathbf{x}_n$$

Similarly we get 
$$\mathbf{\Sigma}_{\mathrm{ML}} = rac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - oldsymbol{\mu}_{\mathrm{ML}}) (\mathbf{x}_n - oldsymbol{\mu}_{\mathrm{ML}})^{\mathrm{T}}$$

 $m{\mu}_{
m ML}$  and  $m{\Sigma}_{
m ML}$  are the maximum likelihood estimates of the mean and the co-variance matrix.

### Matrix-cook-book

```
(A is a constant)
                             \partial \mathbf{A} = 0
                   \partial(\alpha \mathbf{X}) = \alpha \partial \mathbf{X}
          \partial(\mathbf{X} + \mathbf{Y}) = \partial\mathbf{X} + \partial\mathbf{Y}
            \partial(\operatorname{Tr}(\mathbf{X})) = \operatorname{Tr}(\partial \mathbf{X})
                  \partial(\mathbf{XY}) = (\partial\mathbf{X})\mathbf{Y} + \mathbf{X}(\partial\mathbf{Y})
            \partial(\mathbf{X} \circ \mathbf{Y}) = (\partial \mathbf{X}) \circ \mathbf{Y} + \mathbf{X} \circ (\partial \mathbf{Y})
          \partial(\mathbf{X}\otimes\mathbf{Y}) = (\partial\mathbf{X})\otimes\mathbf{Y} + \mathbf{X}\otimes(\partial\mathbf{Y})
                 \partial(\mathbf{X}^{-1}) = -\mathbf{X}^{-1}(\partial\mathbf{X})\mathbf{X}^{-1}
          \partial(\det(\mathbf{X})) = \det(\mathbf{X})\operatorname{Tr}(\mathbf{X}^{-1}\partial\mathbf{X})
\partial(\ln(\det(\mathbf{X}))) = \operatorname{Tr}(\mathbf{X}^{-1}\partial\mathbf{X})
                        \partial \mathbf{X}^T = (\partial \mathbf{X})^T
                        \partial \mathbf{X}^H = (\partial \mathbf{X})^H
```

#### Outline

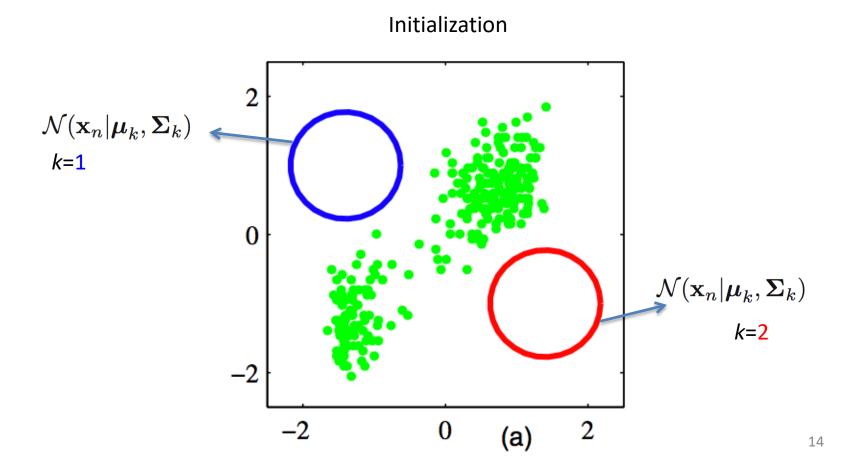
Gaussian Mixture Models (GMM)

 Expectation-Maximization (EM) for maximum likelihood

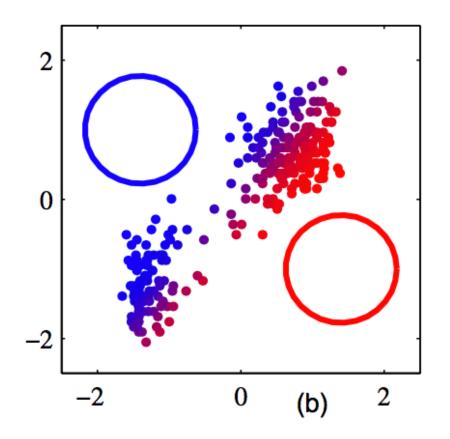
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# Expectation-Maximization (EM) algorithm for maximum likelihood

$$\ln p(\mathbf{X}|\boldsymbol{\pi}, \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \sum_{n=1}^{N} \ln \left\{ \sum_{k=1}^{K} \pi_k \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) \right\}$$



### E Step



When the parameters are given, the assignments of the points can be calculated by the posterior probability, i.e., the probability of a data point belonging to a cluster once we have observed the data point.

$$\gamma(z_{nk}) = rac{\pi_k \mathcal{N}(\mathbf{x}_n | oldsymbol{\mu}_k, oldsymbol{\Sigma}_k)}{\displaystyle\sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}_n | oldsymbol{\mu}_j, oldsymbol{\Sigma}_j)}$$

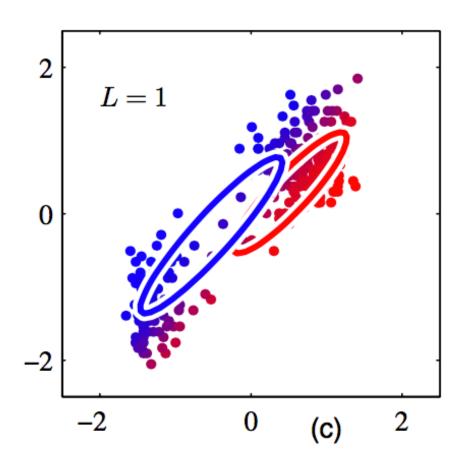
Soft assignment:

A point fractionally belongs to two clusters.

For example,

- 0.2 belong to cluster 1
- 0.8 belong to cluster 2

### M Step



When the assignments  $\gamma(z_{nk})$  of the points to the clusters are known, parameters could be calculated for each cluster (Gaussian) separately.

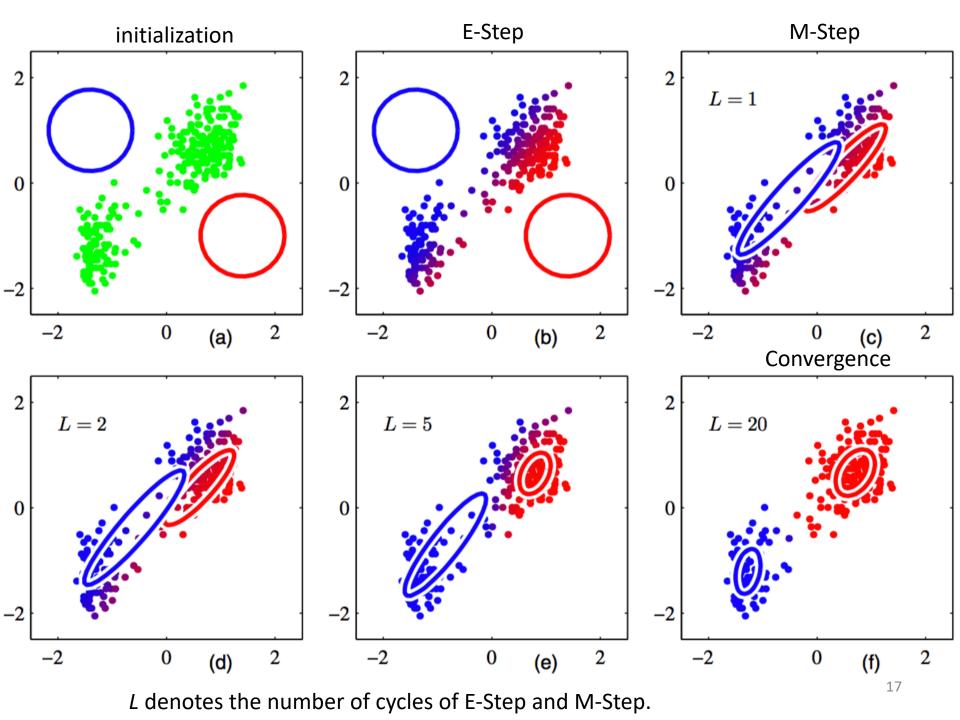
Mixing weight  $\pi_k$ : the proportion of number of points in cluster k within all data points

$$\pi_{m{k}} = rac{N_{m{k}}}{N} \quad ext{;} \quad N_{m{k}} = \sum_{n=1}^{N} \gamma(z_{nm{k}}).$$

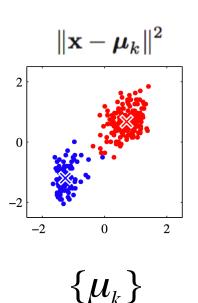
 $\mu_{k_i} \Sigma_k$ : the mean and the covariance matrix are calculated for each cluster  $_N$ 

$$oldsymbol{\mu}_k = rac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$$

$$\mathbf{\Sigma}_k = rac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \boldsymbol{\mu}_k) (\mathbf{x}_n - \boldsymbol{\mu}_k)^{\mathrm{T}}$$

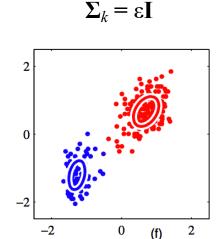


#### Relation to K-means



One-in-K assignment

$$r_{nk} = egin{cases} 1 & ext{if } k = rg \min_j \|\mathbf{x}_n - oldsymbol{\mu}_j\|^2 \ 0 & ext{otherwise}. \end{cases}$$



GMM considers covariance and mixing weights.

$$p(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k) = \frac{1}{(2\pi\epsilon)^{1/2}} \exp\left\{-\frac{1}{2\epsilon}\|\mathbf{x} - \boldsymbol{\mu}_k\|^2\right\}$$

Soft assignment

$$\gamma(z_{nk}) = \frac{\pi_k \exp\left\{-\|\mathbf{x}_n - \boldsymbol{\mu}_k\|^2 / 2\epsilon\right\}}{\sum_j \pi_j \exp\left\{-\|\mathbf{x}_n - \boldsymbol{\mu}_j\|^2 / 2\epsilon\right\}}$$

### Summary for the EM algorithm for GMM

- Does it find the global optimum?
  - No, like K-means, EM only finds the nearest local optimum and the optimum depends on the initialization

 GMM is more general then K-means by considering mixing weights, covariance matrices, and soft assignments.

Like K-means, it does not tell you the best K.

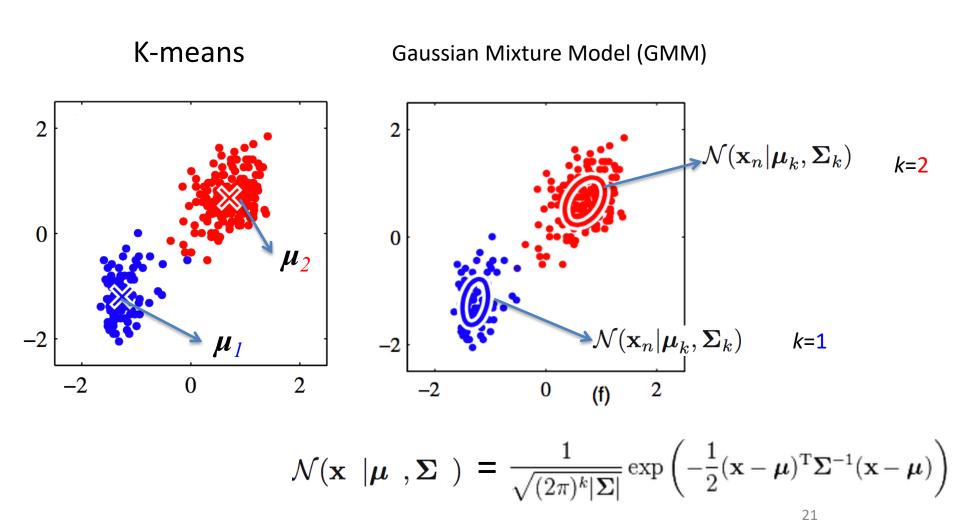
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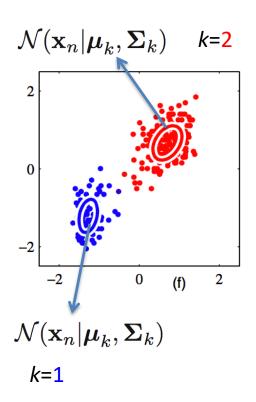
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## Gaussian Mixture Model (GMM)



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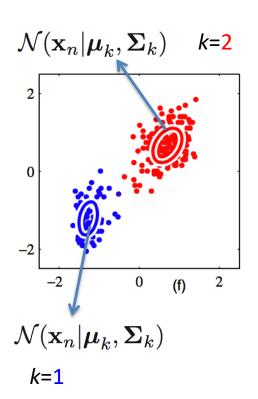
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prior probability of point belonging to a cluster

So, we get a distribution for the data point  $\mathbf{x}$ :

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### Introduce a latent variable



We use  $z_k = 1$  to indicate a point **x** belongs to cluster k

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A mixing weight for each cluster:

$$p(z_k=1)=\pi_k$$
  $0\leqslant\pi_k\leqslant 1$   $\sum_{k=1}^K\pi_k=1$ 

prior probability of point belonging to a cluster

$$p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$$

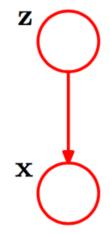
Assume the points in the same cluster follow a **Gaussian distribution** 

$$p(\mathbf{x}|z_k=1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$$

### Gaussian Mixture Model (GMM)

#### **Generative process**

- Randomly sample a **z** from a categorical distribution  $[\pi_1, ..., \pi_K]$ ;
- Generate  $\mathbf{x}$  according to Gaussian distribution  $p(\mathbf{x}|z_k=1) = \mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$

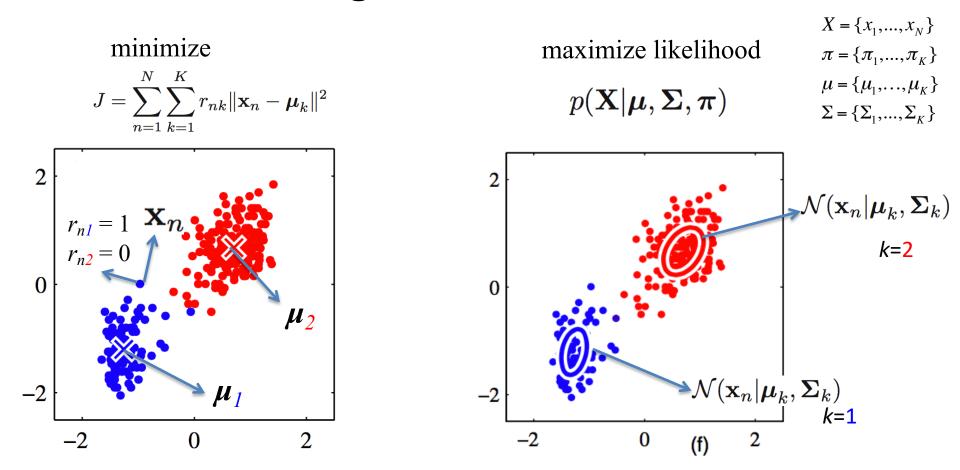


Graphical representation of p(x, z) = p(z)p(x|z)

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# From minimizing sum of square distances to finding maximum likelihood



Remember: The closer the distance, the more likely the probability.

## Thank you!