

# Nested sampling

Common implementations:

- MultiNest
- PolyChord

- o Original method due to Skilling (2004)
- o Actually a method for computing the Bayesian evidence,  $Z$
- o Useful by-product: We get  $\bar{\theta}$  samples distributed according to  $p(\bar{\theta} | D)$

o Want to compute  $Z = \int \mathcal{L}(\bar{\theta}) \pi(\bar{\theta}) d\bar{\theta}$  (\*)

o High-dim integrals are hard! one-dim. are easy!

o (can we turn (\*) into a one-dim. integral?)

o Introduce variable: "prior mass"

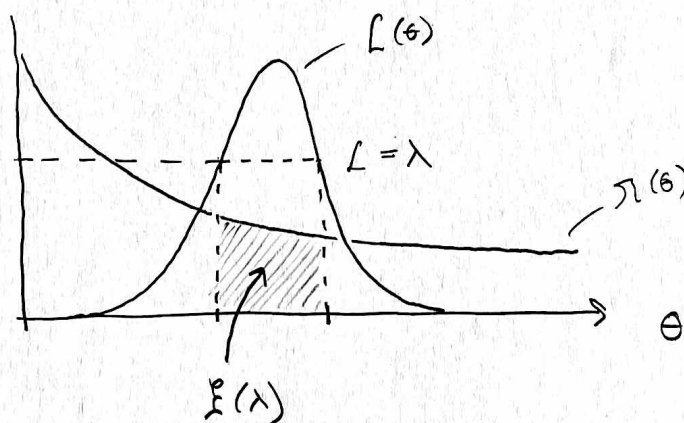
Terminology:

$$\text{prob. density} = \frac{d(\text{prob. mass})}{d(\text{volume})}$$

$$\tilde{Z}(\lambda) = \int_{\mathcal{L}(\theta) > \lambda} \pi(\theta) d\theta$$

$$d\tilde{Z} = \pi(\theta) d\theta$$

Transf. of random variable:  
 $p_x(x) dx = p_y(y) dy$   
 Here:  $\pi_{\tilde{Z}}(\tilde{Z}) d\tilde{Z} = \pi_{\theta}(\theta) d\theta$   
 with  $\pi_{\tilde{Z}}(\tilde{Z}) = 1$   
 and  $\tilde{Z} \in [0, 1]$



$\tilde{Z}(\lambda)$ : The amount of prior probability contained within the regions of parameter space where the likelihood  $\mathcal{L}(\theta)$  is greater than some value  $\lambda$

$d\tilde{Z}$ : The small additional prior mass included by lowering the likelihood threshold by  $dL$

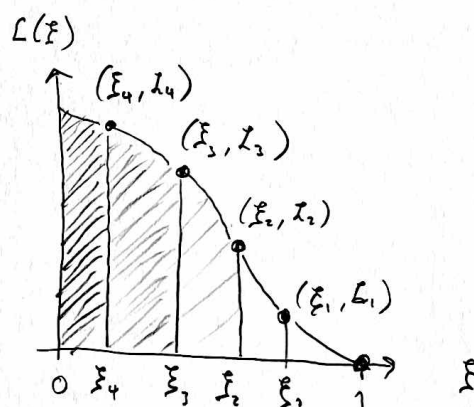
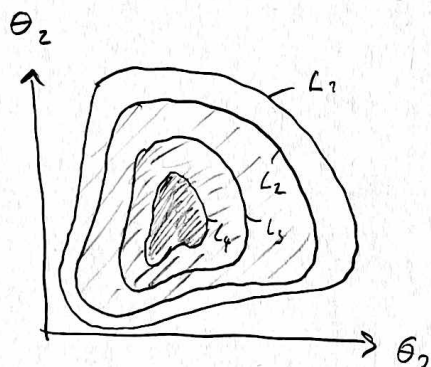
Examples:  $\tilde{Z}(0) = 1$  ,  $\tilde{Z}(\lambda = L_{\max}) = 0$

- Note that  $\xi(\lambda)$  is a one-dim, decreasing function of  $\lambda$
- Inverse function, denoted as  $L(\xi)$  is simply

$L(\xi(\lambda)) \equiv \lambda$  = the value for the likelihood contour that contains a given prior mass  $\xi$

- Can now express  $Z$  as one-dim, integral over  $\xi$  :

$$Z = \int_0^1 L(\xi) d\xi$$



- If we can get a set of ordered pairs of values  $(\xi_i, L_i)$  we can evaluate  $Z$  integral using standard methods, (e.g. trapezoidal rule)
- Nested sampling is algorithm to get these samples

## Algorithm [show slides]

- 1) Draw  $N$  "live points"  $\bar{\theta}$  according to prior  $\pi(\bar{\theta})$
- 2) Evaluate  $L(\bar{\theta})$  at each live point
- 3) Discard (but record) point with lowest likelihood
- 4) Draw a new point from prior, but with additional req. that  $L(\bar{\theta}_{\text{new}}) > L(\bar{\theta}_{\text{disc.}})$
- 5) Repeat from step 3

← Main challenge for algo. efficiency!

• The discarded points form ordered set of likelihood samples

$$0 < L_1 < L_2 < \dots$$

• For each likelihood sample, run estimate corresponding prior mass  $\xi_i$  to obtain (will show this later)

$$1 > \xi_1 > \xi_2 > \dots$$

• Result:

Evidence estimate:

$$Z \approx \sum_{i=1}^M L_i w_i = \sum_{i=1}^M L_i \underbrace{\frac{1}{2} [\xi_{i-1} - \xi_{i+1}]}_{w_i}$$

Posterior samples:

Assign each discarded parameter sample  $\bar{\theta}_i$  its share of the posterior prob.

$$p_i = \frac{L_i w_i}{Z}$$

Main challenge: How to efficiently draw replacement samples from the "likelihood-constrained prior"?

Multilist + friends solve this!

$w_i$ : The slice of prior mass associated with the likelihood value  $L_i$ .  
(Here chosen according to trapezoidal rule)

- How can we estimate the prior mass  $\xi_i$  corresponding to a likelihood value  $L_i$ ?

$$0 < L_1 < L_2 < \dots$$

$$1 > \xi_1 > \xi_2 > \dots$$

- From  $d\xi = \pi(\bar{\theta}) d\bar{\theta}$ , we know that sampling  $\bar{\theta}$  according to  $\pi(\bar{\theta})$  corresponds to sampling  $\xi$  from uniform distribution  $U(0,1)$

$$\left[ \begin{array}{l} \text{Recall relation:} \\ \bar{\theta} \rightarrow L(\bar{\theta}) \rightarrow \lambda \rightarrow \begin{array}{l} \text{Integration} \\ \text{contour for} \\ \int_{L(\bar{\theta}) > \lambda} \pi(\bar{\theta}) d\bar{\theta} \end{array} \rightarrow \xi \end{array} \right]$$

- Sampling constraint  $L(\bar{\theta}_{\text{new}}) > L(\bar{\theta}_{\text{disc}})$  ensures that the prior mass associated with  $\bar{\theta}_{\text{new}}$  is smaller than that for  $\bar{\theta}_{\text{disc}}$ .

$$\xi_{\text{new}} < \xi_{\text{disc}}$$

- We use  $N$  live points. At start of iteration  $i$  we have  $N$   $\bar{\theta}$ -samples that should correspond to  $N$   $\xi$ -samples from uniform distr. on  $(0, \xi_{i-1})$
- The prior mass  $\xi_i$  of next point to be discarded is an unknow/random variable

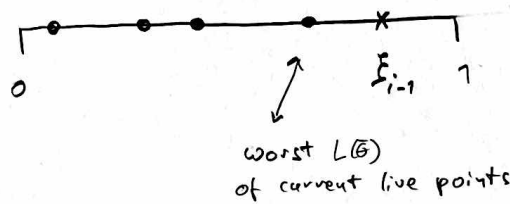
$$\xi_i = t_i \xi_{i-1}$$

where the shrinkage factor  $t_i = \frac{\xi_i}{\xi_{i-1}}$  has a pdf

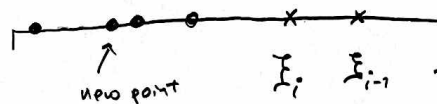
$$p(t) = N t^{N-1}$$

$$\left[ \begin{array}{l} \text{pdf for the largest value } t \text{ of} \\ N \text{ samples drawn from } U(0,1) \end{array} \right]$$

Iteration  $i$  ( $N=4$ )

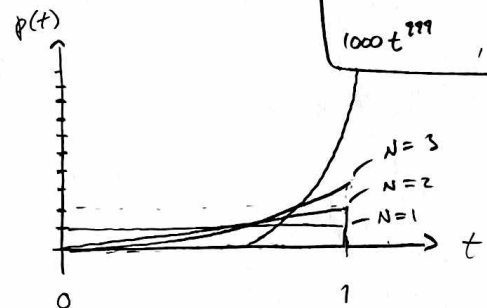


Discard and sample a new point under constraint



(We don't try to compute the exact  $\xi_i$  for the iteration  $i$ , we just estimate it)

$$p(t) = Nt^{N-1} = \begin{cases} 1 & , N=1 \\ 2t & , N=2 \\ 3t^2 & , N=3 \\ 1000t^{999} & , N=1000 \end{cases}$$



- Since we start from  $\xi_0 = 1$ , we can express  $\xi_i$  as the random variable

$$\xi_i = t_i t_{i-1} \dots t_1 \quad \left( \text{since } \xi_i = t_i \xi_{i-1} = t_i t_{i-1} \xi_{i-2} = \dots \right)$$

$$\text{or } \ln \xi_i = \ln t_i + \ln t_{i-1} + \dots$$

- All the  $t_i$  have pdf  $p(t) = Nt^{N-1}$  which give

$$E[\ln t] = -\frac{1}{N}, \quad \text{Var}[\ln t] = \frac{1}{N^2}$$

- This means that the sum  $\ln \xi_i = \ln t_i + \ln t_{i-1} + \dots = \sum$  has expectation and variance

$$\begin{aligned} E[\ln \xi_i] &= E[\ln t_i] + E[\ln t_{i-1}] + \dots = \frac{1}{N} \\ &= -\frac{1}{N} - \frac{1}{N} - \dots \\ &= -\frac{i}{N} \end{aligned}$$

$$\begin{aligned} \text{Var}[\ln \xi_i] &= \text{Var}[\ln t_i] + \text{Var}[\ln t_{i-1}] + \dots \quad \left( \text{Since the } t_i \text{ are uncorrelated} \right) \\ &= \frac{i}{N^2} \end{aligned}$$

- In short:  $\ln \xi_i \approx -\frac{i}{N} \pm \frac{\sqrt{i}}{N}$

- So we approximate the prior mass  $\xi_i$  associated with the likelihood value  $L_i$  of the discarded point  $\bar{\theta}$  at iteration  $i$  as

$$\boxed{\xi_i \approx e^{-\frac{i}{N}}}$$

- Then we have what we need to compute a new approx. for  $Z$  after each new iteration  $i$

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- The sampling stops when the largest possible contribution  $\Delta Z$  from the current live points is much smaller than the current estimate for  $Z$   
(But this will fail if the sampling has missed some region of high likelihood)
- Uncertainty on final evidence estimate is dominated by uncertainty in  $\xi_i$  estimates  
(assuming the sampling has found all relevant parameter regions.)



• Efficiency challenge :

— Naively sampling  $\bar{\theta}$  points from entire  $\pi(\bar{\theta})$  at every iteration will lead to ever decreasing efficiency, due to constraint  $L(\bar{\theta}_{\text{new}}) > L(\bar{\theta}_{\text{disc.}})$

— One appr. used to alleviate problem :

— Draw samples from ellipsoids containing current live points

— Use clustering algo. to assign sep. ellipsoids to sep. clusters of live points

• Much used packages : MultiNest , PolyChord  
(`py-multinest`)