

Comp Sci, Nov 9, 2022

$$\begin{array}{l|l} X \in \mathbb{R}^{m \times p} & U U^T = U^T U = \mathbb{1} \\ X^T X \in \mathbb{R}^{p \times p} & V V^T = V^T V = \mathbb{1} \\ X = U \Sigma V^T & U \in \mathbb{R}^{m \times m} \\ & V \in \mathbb{R}^{p \times p} \end{array}$$

$$\Sigma \in \mathbb{R}^{m \times p}$$

$$\Sigma = \begin{bmatrix} \sigma_0 & & & 0 \\ & \sigma_1 & & \\ & 0 & \ddots & \\ & & & \sigma_{p-1} \\ & & & & 0 \end{bmatrix}$$
$$\sigma_0 > \sigma_1 > \sigma_2 \dots > \sigma_{p-1} > 0$$

$$U = [u_0 \ u_1 \ \dots \ u_{m-1}]$$

$$V = [v_0 \ v_1 \ \dots \ v_{p-1}]$$

$$\begin{aligned}
X^T X &= (U \Sigma V^T)^T U \Sigma V^T \\
&= V \Sigma^T \underbrace{U^T U}_{I} \Sigma V^T \\
&= V \Sigma^T \Sigma V^T
\end{aligned}$$

$$\Sigma \in \mathbb{R}^{n \times p}$$

$$\Sigma^T \Sigma \in \mathbb{R}^{p \times p}$$

$$\Sigma^T \Sigma = \begin{bmatrix} \sigma_0^2 & & 0 \\ & \ddots & \\ 0 & & \sigma_{p-1}^2 \end{bmatrix}$$

$$\Sigma \Sigma^T \in \mathbb{R}^{n \times n}$$

$$X^T X = V \Sigma^T \Sigma V^T$$

intermediate

multiply from the right with
 V

$$(X^T X) V = V \Sigma^T \Sigma$$

$$(X^T X) v_i = v_i \sigma_i^2$$

Eigenvalues of $X^T X$ are $\sigma_i^2 > 0$

$X^T X$ is positive definite.

$$\frac{\partial^2 C^{OLS}}{\partial \beta \partial \beta^T} = \frac{2}{n} X^T X \Rightarrow$$

convex problem with a global minimum.

in ridge regression

$X^T X + \lambda I$, also a convex problem.

intermezzo 2

$$X^T X v_i = \sigma_i^2 v_i$$

$$\text{cov}(x_i, x_j) = \int_D dx_i dx_j' \frac{1}{n} (x_i - \mu_i)(x_j' - \mu_j)$$

$$\times p(x_i, x_j)$$

if x_i, x_j are i.i.d. ($p(x_i)p(x_j)$
 $= p(x_i, x_j)$)
 then $\text{cov}(x_i, x_j) = 0$

$$X = [x_0 \ x_1 \ x_2 \ \dots \ x_{p-1}]$$

$$\text{cov}(x_i, x_j) = \frac{1}{n} \sum_{k=0}^{n-1} (x_{ki} - \mu_i)(x_{kj} - \mu_j)$$

$$X = \begin{bmatrix} (x_{00}) & x_{01} & \dots & x_{0p-1} \\ x_{10} & & & \\ \vdots & & & \\ x_{n-1,0} & x_{n-1,1} & \dots & x_{n-1,p-1} \\ x_0 & x_1 & \dots & x_{p-1} \end{bmatrix}$$

$$\text{cov}(x_i, x_j) \rightarrow \text{cov}[X]$$

$$= \frac{1}{n} X^T X$$

The eigen values
 are proportional
 with Σ

$$\text{cov}[\bar{x}] = \begin{bmatrix} \sigma_0^2 & \text{cov}[\bar{x}_0, \bar{x}_1] & \dots & \text{cov}[\bar{x}_0, \bar{x}_{p-1}] \\ \times & \sigma_1^2 & \times & \times \\ \times & \times & \times & \times \\ \times & \times & \times & \times & \sigma_{p-1}^2 \end{bmatrix}$$

$$S \text{cov}[\bar{x}] S^T = D$$

$$= \begin{bmatrix} \sigma_0^2 & & & 0 \\ & \sigma_1^2 & & \\ & & \ddots & \\ 0 & & & \end{bmatrix}$$

$$\text{cov}[\bar{y}] = S \text{cov}[\bar{x}] S^T$$

$$Ax = \lambda x$$

$$SAS^T = D = \begin{bmatrix} \lambda_0 & & \\ & \ddots & \\ & & \lambda_{n-1} \end{bmatrix}$$

$$SA \uparrow x = S \lambda x = \lambda Sx$$

$$SS^T = \underline{1} = S^T S$$

$$\underbrace{SA^T}_{B} \underbrace{Sx}_y = \lambda \underbrace{Sx}_y$$

$$B \cdot y = \lambda y$$

Back to OLS and Ridge with SVD

$$X^T X = V \Sigma^T \Sigma V^T$$

$$\hat{\beta}_{OLS} = (X^T X)^{-1} X^T y$$

$$\hat{y} = X \hat{\beta}_{OLS}$$

$$= X (X^T X)^{-1} X^T y$$

$$= U \Sigma V^T (V \Sigma^T \Sigma V^T)^{-1} V \Sigma^T U^T y$$

$$(AB)^{-1} = B^{-1} A^{-1}$$

A & B are square matrices
and both are invertible

$$V V^T = \mathbb{1} \quad V = (V^T)^{-1}$$

$$\begin{aligned} (V \Sigma^T \Sigma V^T)^{-1} &= (V^T)^{-1} (\Sigma^T \Sigma)^{-1} V^{-1} \\ &= V (\Sigma^T \Sigma)^{-1} V^T \end{aligned}$$

$$\begin{aligned} \tilde{y} &= X \hat{\beta}_{OLS} = U U^T y \\ &= \left(\sum_{i=0}^{p-1} u_i u_i^T \right) y \end{aligned}$$

Ridge

$$\hat{\beta}_{Ridge} = (X^T X + \lambda \mathbb{I})^{-1} X^T y$$

SVD :

$$\begin{aligned} \tilde{y}_{Ridge} &= X \hat{\beta}_{Ridge} \\ &= \sum_{j=0}^{p-1} u_j u_j^T \frac{\sigma_j^2}{\sigma_j^2 + \lambda} \end{aligned}$$

$$\sigma_0^2 > \sigma_1^2 > \dots > \sigma_{p-1}^2 > 0$$

Simple example

$$\begin{array}{l} X = \underline{1} \\ \tilde{y}_{OLS} = \hat{\beta}_{OLS} \end{array} \quad \left| \quad n = p \right.$$

$$C(\beta)_{OLS} = \frac{1}{n} \sum_{i=0}^{p-1} (y_i - \beta_i)^2$$

$$\tilde{y}_i = \sum_{j=0}^{p-1} x_{ij} \beta_j = \beta_i$$

$$x_{ij} = 1 \text{ if } i = j$$

minimize with respect to

$$\beta_j \Rightarrow \hat{\beta}_i = y_i$$

$$\frac{\partial C_{OLS}}{\partial \beta_j} = - \frac{2}{n} (y_j - \beta_j) = 0$$

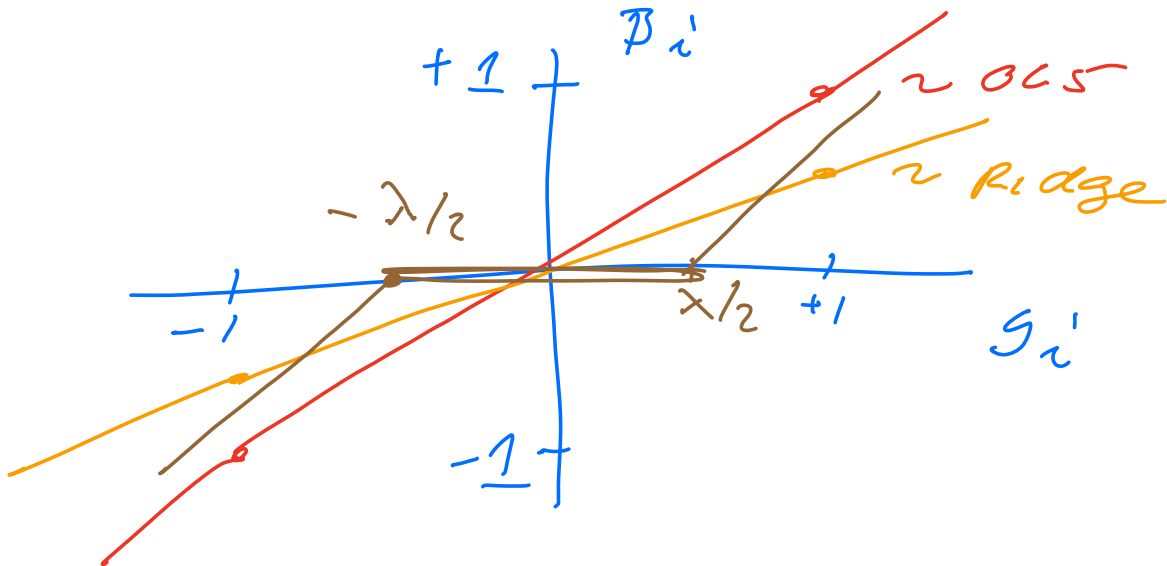
Ridge regression

$$C(\beta)_{\text{ridge}} = \frac{1}{n} \sum_{i=0}^{p-1} (y_i - \beta_i)^2 + \lambda \sum_{i=0}^{p-1} \beta_i^2$$

$$\frac{\partial C}{\partial \beta_i} = 0 = -\frac{2}{n} (y_i - \beta_i) + 2\lambda \beta_i = 0$$

\Rightarrow

$$\beta_i^{\text{ridge}} = \frac{y_i}{1 + \lambda}$$



Lasso Regression

$$C(\beta)_{\text{Lasso}} = \frac{1}{n} \sum_{j=0}^{p-1} (y_j - \beta_j)^2 + \lambda \sum_{j=0}^{p-1} |\beta_j|$$

$$\frac{\partial C_{\text{Lasso}}}{\partial \beta_i} = -\frac{2}{n} (y_i - \beta_i) + \lambda \frac{\beta_i}{|\beta_i|}$$

$$\frac{d|x|}{dx} = \begin{cases} +1 & \text{if } x > 0 \\ -1 & \text{if } x < 0 \end{cases}$$

$$\hat{\beta}_i^{\text{Lasso}} = \begin{cases} y_i - \lambda/2 & \text{if } y_i > \lambda/2 \\ y_i + \lambda/2 & \text{if } y_i < -\lambda/2 \\ 0 & \text{if } |y_i| \leq \lambda/2 \end{cases}$$

