

*** Applied Machine Learning Fundamentals ***

Probability Density Estimation (PDE)

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SAP SE

November 29, 2019



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Lecture Overview

Unit I	Machine Learning Introduction
Unit II	Mathematical Foundations
Unit III	Bayesian Decision Theory
Unit IV	Probability Density Estimation
Unit V	Regression
Unit VI	Classification I
Unit VII	Evaluation
Unit VIII	Classification II
Unit IX	Clustering
Unit X	Dimensionality Reduction

Agenda November 29, 2019

① Introduction

What about continuous Data?
Methods for PDE

② Parametric Models

General Idea
Parameter Learning and Assumptions
Maximum Likelihood Estimation (MLE)

③ Non-parametric Models

④ Mixture Models

General Idea
Mixture of Gaussians
Expectation Maximization for MoG

⑤ Wrap-Up

Summary
Self-Test Questions
Lecture Outlook
Recommended Literature and further Reading
Meme of the Day

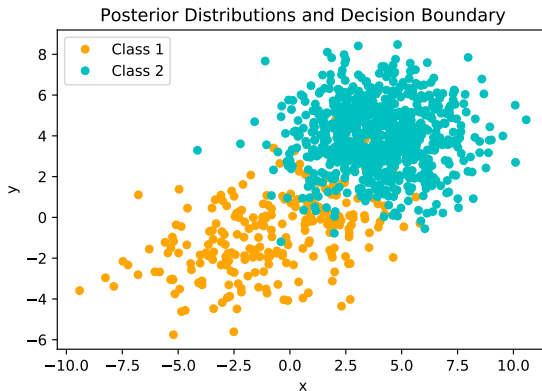
Section:
Introduction



Probability Density Estimation (PDE)

- We have learned about Bayes' optimal classifiers which classify data based on the probability distribution $p(\mathbf{x}|\mathcal{C}_k) \cdot p(\mathcal{C}_k)$
- Naïve Bayes is an instance of PDE for **discrete data**
- **How to get these probabilities in the continuous case?**
 - The prior $p(\mathcal{C}_k)$ is still easy to compute
 - The estimation of class conditional probabilities $p(\mathbf{x}|\mathcal{C}_k)$ is more complicated
 - Assume labeled data; estimate the density separately for each class \mathcal{C}_k
- NB: For ease of notation: $p(\mathbf{x}) \equiv p(\mathbf{x}|\mathcal{C}_k)$

Training Data Example



Overview of the Methods for PDE

① Parametric models (maximum likelihood estimation)

- Assume a fixed parametric form (e. g. a Gaussian distribution)
- Estimate the parameters such that the model fits the data best

② Non-parametric models

- Often we do not know the functional form of the density
- Estimate probability directly from the data without an explicit model

③ Mixture models

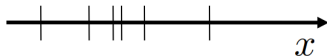
- Combination of ① and ②
- EM algorithm

Section:
Parametric Models



General Approach

- Given some (continuous) training data $\mathbf{X} = \{x^{(i)}\}_{i=1}^n$ (where all $x^{(i)}$ belong to the same class):



- Estimate $p(x)$ using a fixed parametric form:



Example: Gaussian Distribution

- One common case is the **Gaussian distribution**:

$$p(x|\mu, \sigma^2) = \mathcal{N}(x|\mu, \sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{(x - \mu)^2}{2\sigma^2} \right\} \quad (1)$$

- Notation for parametric models:
 - $p(x|\theta)$
 - In the case of a Gaussian: $\theta = \{\mu, \sigma^2\}$

$\mu \hat{=}$ mean
 $\sigma^2 \hat{=}$ variance

Learning the Parameters

- Learning = Estimation of the parameters θ given the data \mathbf{X}
- **Likelihood** of the parameters θ :
 - Is defined as the probability that \mathbf{X} was generated by a probability density function (pdf) with parameters θ

$$\mathcal{L}(\theta) = p(\mathbf{X}|\theta) \quad (2)$$

- We want to **maximize** the likelihood

⇒ **Maximum likelihood estimation (MLE)**

A fundamental Assumption

- How to compute $\mathcal{L}(\boldsymbol{\theta})$?
- The data is assumed to be **i.i.d.** (independent and identically distributed):
 - Two random variables x_1 and x_2 are independent if

$$P(x_1 \leq \alpha, x_2 \leq \beta) = P(x_1 \leq \alpha) \cdot P(x_2 \leq \beta) \quad \forall \alpha, \beta \in \mathbb{R} \quad (3)$$

- Two random variables x_1 and x_2 are identically distributed if

$$P(x_1 \leq \alpha) = P(x_2 \leq \alpha) \quad \forall \alpha \in \mathbb{R} \quad (4)$$

Computation of the Likelihood

$$\begin{aligned}\mathcal{L}(\boldsymbol{\theta}) &= p(\mathbf{X}|\boldsymbol{\theta}) \\ &= p(x^{(1)}, x^{(2)}, \dots, x^{(n)}|\boldsymbol{\theta})\end{aligned}$$

data is independent:

$$= p(x^{(1)}|\boldsymbol{\theta}) \cdot p(x^{(2)}|\boldsymbol{\theta}) \cdot \dots \cdot p(x^{(n)}|\boldsymbol{\theta})$$

data is identically distributed:

$$= \prod_{i=1}^n p(x^{(i)}|\boldsymbol{\theta})$$

What is the problem here?

(5)

Computation of the Likelihood (Ctd.)

- **Problem:** Large n might cause arithmetic underflows! (why?)
- Transform the likelihood using the logarithm \Rightarrow **log-likelihood**

$$\mathcal{LL}(\boldsymbol{\theta}) = \log \mathcal{L}(\boldsymbol{\theta})$$

Why is this an allowed transformation?

$$= \log \prod_{i=1}^n p(x^{(i)} | \boldsymbol{\theta})$$

$$\log \Pi = \Sigma \log$$

$$= \sum_{i=1}^n \log p(x^{(i)} | \boldsymbol{\theta}) \quad (6)$$

Maximum Likelihood of a Gaussian

- $\theta = \{\mu, \sigma^2\}$

$$\mathcal{LL}(\{\mu, \sigma^2\}) = \sum_{i=1}^n \log \mathcal{N}(x^{(i)} | \mu, \sigma^2) \quad (7)$$

$$= \sum_{i=1}^n \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{(x^{(i)} - \mu)^2}{2\sigma^2} \right\} \quad (8)$$

- Find μ_{ml} and σ_{ml}^2 which maximize the log-likelihood:

$$\mu_{ml}, \sigma_{ml}^2 = \arg \max_{\mu, \sigma^2} \mathcal{LL}(\theta)$$

Maximum Likelihood of a Gaussian (Ctd.)

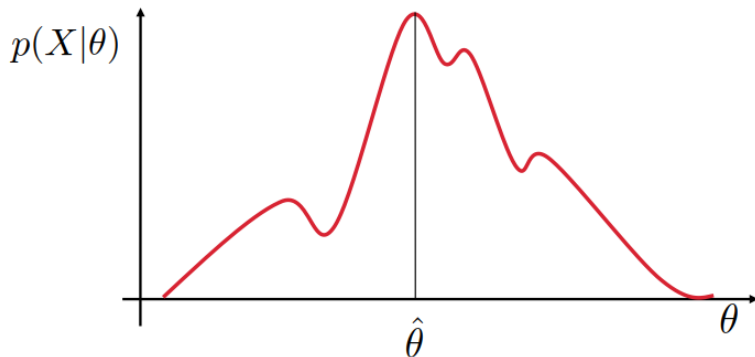
- Compute the partial derivatives with respect to the parameters θ
- Derivative w. r. t. μ :

$$\nabla_{\mu} \mathcal{L}(\theta) = \nabla_{\mu} \sum_{i=1}^n \log \frac{1}{\sqrt{2\pi\sigma^2}} \exp \left\{ -\frac{(x^{(i)} - \mu)^2}{2\sigma^2} \right\} = \sum_{i=1}^n \frac{x^{(i)} - \mu}{\sigma^2}$$

- Set derivative to zero and solve:

$$\sum_{i=1}^n x^{(i)} - \mu \stackrel{!}{=} 0 \Leftrightarrow n \cdot \mu = \sum_{i=1}^n x^{(i)} \Leftrightarrow \mu = \frac{1}{n} \sum_{i=1}^n x^{(i)}$$

Maximization of the Likelihood



We can classify!

Looks familiar?

- Maximum likelihood parameters:

$$\mu_{ml} = \frac{1}{n} \sum_{i=1}^n x^{(i)}$$

$$\sigma_{ml}^2 = \frac{1}{n} \sum_{i=1}^n (x^{(i)} - \mu_{ml})^2$$

- Now we can use Bayes' rule to predict class labels
 - We have the priors...
 - ...and the class conditionals
- Also, the **decision boundary** can be computed

Multivariate Case

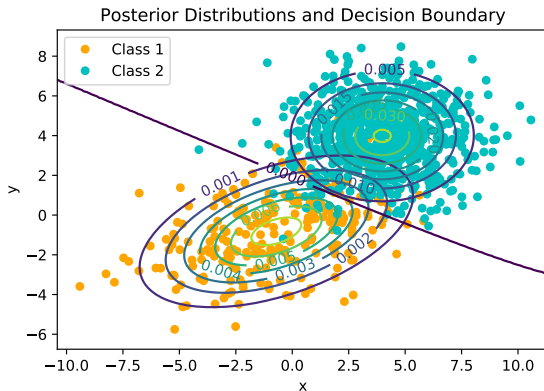
- The solution above is for 1-D data, what if we have more dimensions?
- **Multivariate Gaussian distribution:**

$$\mathcal{N}_D(\mathbf{x}|\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{\sqrt{(2\pi)^D |\boldsymbol{\Sigma}|}} \exp \left\{ -\frac{1}{2} (\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{x} - \boldsymbol{\mu}) \right\} \quad (9)$$

- Luckily, the derivations don't change:

$$\boldsymbol{\mu}_{ml} = \frac{1}{n} \sum_{i=1}^n \mathbf{x}^{(i)} \quad \boldsymbol{\Sigma}_{ml} = \frac{1}{n} \sum_{i=1}^n (\mathbf{x}^{(i)} - \boldsymbol{\mu}_{ml})(\mathbf{x}^{(i)} - \boldsymbol{\mu}_{ml})^\top \quad (10)$$

MLE for the Example Data Set



Section:
Non-parametric Models



Disadvantages of parametric Models

- Until now we used a fixed parametric form (e.g. a Gaussian) which is governed by a small amount of parameters
- **This assumption may be wrong:**
 - Another distribution (exponential, gamma, ...) may fit better
 - A suitable 'text-book distribution' may not exist

We don't want to make any assumptions about the underlying distribution!

Non-parametric Approaches

- ① Histograms (Binning)
- ② Kernel density estimation (KDE)
- ③ Nearest neighbors (kNN)

Histograms

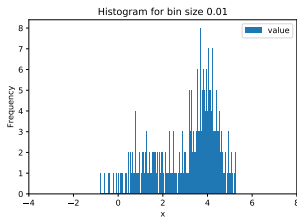
- Histograms partition the data $\mathbf{X} = \{\mathbf{x}^{(i)}\}_{i=1}^n$ into distinct **bins** of volume v_j ...
- ...and subsequently count the number of instances k_j falling into the j -th bin
- Approximate the probability $p(\mathbf{x})$ by:

$$p(\mathbf{x}) \approx \frac{k_j}{n \cdot v_j} \quad \text{for } \mathbf{x} \text{ in bin } j \quad (11)$$

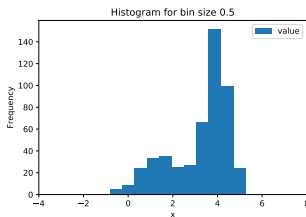
- The sum of all probabilities equals 1: $\sum_{j=1}^m \frac{k_j}{n \cdot v_j} = 1$
- v_j is a **hyper-parameter** (usually, all bins have equal size)

Histograms (Ctd.)

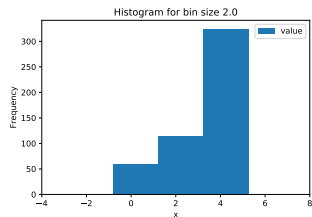
Too narrow



About right

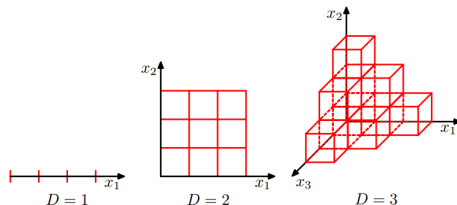


Too wide



Drawbacks of Histograms

- Histograms are mostly unsuited for many applications
- **Drawbacks:**
 - ① **Discontinuities** due to bin edges
 - ② Number of bins **explodes** with growing number of dimensions D



The latter issue is known as the curse of dimensionality

An alternative Approach

- Don't use a fixed number of pre-determined bins
- Instead, employ a **sliding window** approach by centering a region \mathcal{R} (bin) around the data point of interest \mathbf{x}

$$p(\mathbf{x}) \approx \frac{k}{n \cdot v} \quad (12)$$

- This gives rise to two different techniques:
 - ① **Kernel density estimation** (Fix v and determine k)
 - ② **k-nearest neighbors** (Fix k and determine v)

Kernel Density Estimation: Parzen Window

- \mathcal{R} is a D -dimensional **hyper-cube** of edge length h centered on \mathbf{x}
- Determine if a data point falls into region \mathcal{R} :

$$H(\mathbf{u}) = \begin{cases} 1 & \text{if } |u_d| \leq h/2, d = 1, 2, \dots, D \\ 0 & \text{otherwise} \end{cases} \quad (13)$$

- The total number of data points falling into region \mathcal{R} is given by:

$$k(\mathbf{x}) = \sum_{i=1}^n H(\mathbf{x} - \mathbf{x}^{(i)}) \quad (14)$$

Kernel Density Estimation: Parzen Window (Ctd.)

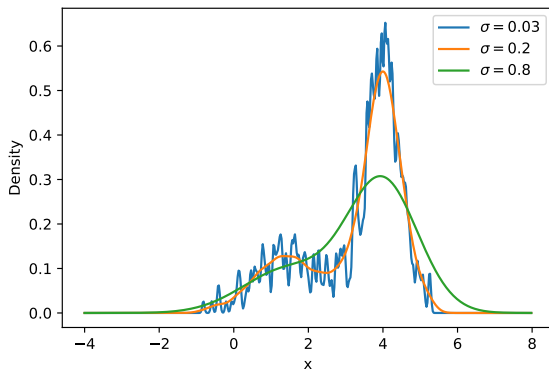
- The volume v is simple to compute:

$$v = \int H(\mathbf{u}) \, d\mathbf{u} = h^D \quad (15)$$

- Putting it all together we get:

$$p(\mathbf{x}) \approx \frac{k(\mathbf{x})}{n \cdot v} = \frac{1}{n \cdot h^D} \sum_{i=1}^n H(\mathbf{x} - \mathbf{x}^{(i)}) \quad (16)$$

Kernel Density Estimation: Parzen Window (Ctd.)



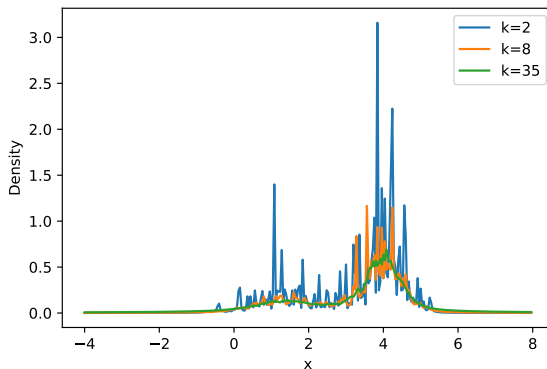
k-Nearest Neighbors

- Increase the size of a sphere until k data points fall into this sphere, keep the number K of data points fixed

$$p(\mathbf{x}) \approx \frac{k}{n \cdot v(\mathbf{x})} \quad (17)$$

- We will also look at k-Nearest Neighbors as a classification model later
→ you can use a majority vote among the k closest training data points to classify a new data point \mathbf{x}

k-Nearest Neighbors (Ctd.)



Section:
Mixture Models



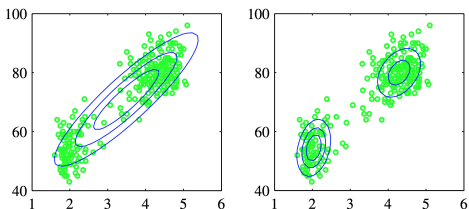
Why do we need mixture models?

- Parametric models have low memory footprint, are quick at runtime and often have nice analytic properties
- Non-parametric models make fewer assumptions about the data, but are slower and have a high memory footprint
- We can combine different models in a mixture model!

$$p(\mathbf{x}) = \sum_{j=1}^M p(\mathbf{x}|j)p(j) \quad (18)$$

Why do we need mixture models?

- A single parametric model might not be able to capture the structure of the dataset at hand
- Mixture distributions (e. g. a linear combination of Gaussians) can approximate almost any continuous density to arbitrary accuracy (given a sufficient number of Gaussians is used)

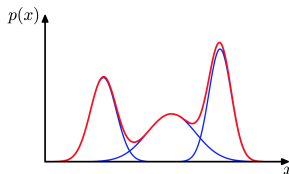


Mixture of Gaussians Definition

- Sum up the M individual Gaussian distributions

$$p(x) = \sum_{j=1}^M p(x|z_j) \quad (19)$$

$$p(x|z_j) = \mathcal{N}(x|\mu_j, \sigma_j) = \frac{1}{\sqrt{2\pi}\sigma_j} \exp\left(-\frac{(x - \mu_j)^2}{2\sigma_j^2}\right) \quad (20)$$



Mixture of Gaussians Definition

- Sum up the M individual Gaussian distributions

$$p(x) = \sum_{j=1}^M p(x|z_j) \quad (21)$$

$$p(x|z_j) = \mathcal{N}(x|\mu_j, \sigma_j) = \frac{1}{\sqrt{2\pi}\sigma_j} \exp\left(-\frac{(x - \mu_j)^2}{2\sigma_j^2}\right) \quad (22)$$

- z_j for $j \in \{1, \dots, M\}$ are the *mixing coefficients*
- $z_j \sim \text{Multinomial}(\phi)$, where $\phi_j \geq 0$ and $\sum_{j=1}^M \phi_j = 1$

Mixture of Gaussians Definition

- Sum up the M individual Gaussian distributions

$$p(x) = \sum_{j=1}^M p(x|z_j) \quad (23)$$

$$p(x|z_j) = \mathcal{N}(x|\mu_j, \sigma_j) = \frac{1}{\sqrt{2\pi}\sigma_j} \exp\left(-\frac{(x - \mu_j)^2}{2\sigma_j^2}\right) \quad (24)$$

- The overall mixture density $p(x)$ integrates to 1
- The parameters: $\theta = \{\mu_1, \sigma_1, \pi_1, \dots, \mu_M, \sigma_M, \pi_M\}$

Which function do we need to maximize to estimate the parameters?

Maximum Likelihood Estimation for MoG

- We have defined our Gaussian mixture model: $p(x) = \sum_{j=1}^M p(x|z_j)$
- Maximize the log-likelihood to estimate the parameters θ :

$$\mathcal{L} = \log L(\theta) = \sum_{i=1}^N \log p(x_i|\theta) \quad (25)$$

$$\frac{\partial \mathcal{L}}{\partial \mu_j} = 0 \quad (26)$$

$$\mu_j = \frac{\sum_{i=1}^N p(z_j|x_i) x_i}{\sum_{i=1}^N p(z_j|x_i)} \quad \Leftarrow \text{Do you see the issue?} \quad (27)$$

Maximum Likelihood Estimation for MoG

- Maximize the log-likelihood to estimate the parameters θ :

$$\mathcal{L} = \log L(\theta) = \sum_{i=1}^N \log p(x_i|\theta) \quad (28)$$

$$\frac{\partial \mathcal{L}}{\partial \mu_j} = 0 \quad (29)$$

$$\mu_j = \frac{\sum_{i=1}^N p(z_j|x_i) x_i}{\sum_{i=1}^N p(z_j|x_i)} \quad \Leftarrow \text{Do you see the issue?} \quad (30)$$

- There is circular dependency, μ_j depends on $z_j \Rightarrow$ there is no analytical solution!

Maximum Likelihood Estimation for MoG

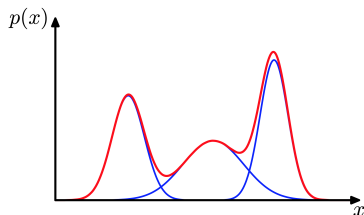
- There is circular dependency, μ_j depends on $z_j \Rightarrow$ there is no analytical solution!
- How about using gradient descent?
 - Complex gradient (circular dependency, non-linearity)
 - Optimization of each Gaussian mixture component depends on all of the other mixture components!
- We have to take a different approach: *Expectation Maximization*

Expectation Maximization

- We have observed data (without labels) x_i and unobserved / hidden / latent variables z_j
- If we knew which Gaussian has generated which data point x_i , then we could do a maximum likelihood estimation for each mixture component j
- If we knew the Gaussians already, i. e. their means and (co)variances, then we could assign each data point x_i to mixture component η if
$$p(z_\eta = 1|x_i) > p(z_j = 2|x_i) \quad \forall j$$

Expectation Maximization

- This is a chicken-and-egg problem:
 - We don't know which mixture component generated which x_i
 - We don't know the distributions
- How can we estimate the parameters of the distribution and the mixture coefficients z_j ?



Expectation Maximization

- How can we estimate the parameters of the distribution and the mixture coefficients z_j ?
- Idea:
 - Assign a *mixture label* to each data point x_i , cluster the data points and assign them to the mixture components
 - The mixture coefficients z_j are *soft assignments* of data points to mixture components
 - Iterate between updating the estimate of the mixture coefficients z_j and the distribution parameters μ_j, σ_j

Expectation Maximization

EM for Gaussian Mixture Models:

- ① Initialize μ_j, σ_j, z_j and evaluate the initial log-likelihood
- ② **E step:** Compute the posterior distribution (*responsibilities* γ_{ij}) for each mixture coefficient z_j and all data points x_i using the current parameter values

$$\gamma_{ij} = p(z_j|x_i) = \frac{z_j \mathcal{N}(x_i|\mu_j, \sigma_j)}{\sum_{m=1}^M z_m \mathcal{N}(x_i|\mu_m, \sigma_m)} \quad (31)$$

Expectation Maximization

EM for Gaussian Mixture Models:

③ **M step:** Compute the new parameters using weighted estimates

$$N_j = \sum_{i=1}^N \gamma_{ij} \quad z_j^{\text{new}} = \frac{N_j}{N} \quad (32)$$

$$\mu_j^{\text{new}} = \frac{1}{N_j} \sum_{i=1}^N \gamma_{ij} x_i \quad (\sigma_j^{\text{new}})^2 = \frac{1}{N_j} \sum_{i=1}^N \gamma_{ij} (x_i - \mu_j^{\text{new}})^2 \quad (33)$$

Expectation Maximization

EM for Gaussian Mixture Models:

- 4 Iterate **E Step** and **M Step** until convergence, i. e. evaluate the log-likelihood after each run and check if the parameters converge:

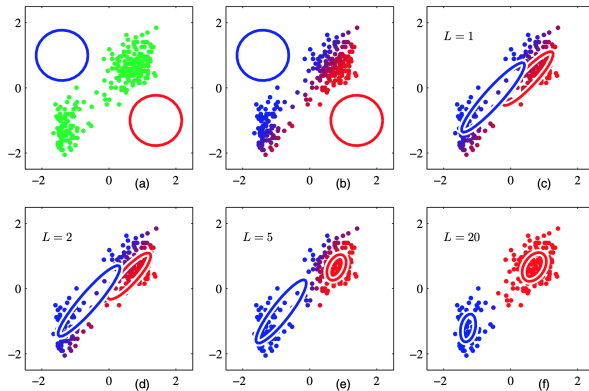
$$\mathcal{L} = \sum_{i=1}^N \log p(x_i | \boldsymbol{\theta}) = \sum_{i=1}^N \log \left(\sum_{j=1}^M z_j \mathcal{N}(x_i | \mu_j, \sigma_j) \right) \quad (34)$$

Expectation Maximization

- We can use EM for performing maximum likelihood estimation even when the data is incomplete (missing features)
- The incomplete log-likelihood \mathcal{L} is maximized locally
- The log-likelihood is guaranteed to improve or stay the same in every EM iteration
- There are great visualizations of EM for Gaussian mixture models:
 - EM density estimation animation
 - 2-dimensional EM animation

Expectation Maximization

Let's look at some iterations of expectation maximization on a simple dataset:



Expectation Maximization

- How do we initialize the parameters for EM?
 - EM depends on a good initialization of the parameters, a poor initialization can lead to getting stuck in bad local optima
 - We can use *k-means* to get an initial clustering
- How many mixture components do we need?
 - There are different criteria, we can for instance maximize the *Bayesian Information Criterion* with respect to K :

$$\log p(X|\theta_{ML}) - \frac{1}{2}K \log N \quad (35)$$

- K : Number of parameters
- N : Number of data points

Section:
Wrap-Up



Summary

- We can use parametric, non-parametric and mixture models for density estimation
- This allows us to estimate the probabilities needed for e. g. a Naïve Bayes model to work with continuous features
- Parametric models assume a certain form of the density, governed by parameters like mean and variance for a Gaussian
- Maximum likelihood estimation allows us to determine the parameters based on our dataset
- Non-parametric models directly use the (training) data points themselves
- We can use expectation maximization to optimize the parameters of mixture models



Self-Test Questions

- What is maximum likelihood estimation? How can you get the maximum likelihood estimate for a Gaussian distribution?
- What does *non-parametric* mean in *non-parametric models*?
- What is different between kernel density estimation and k-nearest neighbors?
- Why can't we use a simple maximum likelihood estimate for a mixture model as with a single Gaussian distribution?
- What happens in the E and M steps in expectation maximization?

What's next...?

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Recommended Literature and further Reading



[1] Pattern Recognition and Machine Learning - Chapter 1.2.4 *The Gaussian Distribution*

Bishop. 2006.

→ [Link](#)



[1] Pattern Recognition and Machine Learning - Chapter 2.5 *Non-Parametric Methods*

Bishop. 2006.

→ [Link](#)



[1] Pattern Recognition and Machine Learning - Chapter 9.2 *Mixtures of Gaussians*

Bishop. 2006.

→ [Link](#)

Meme of the Day



Thank you very much for the attention!

Topic: *** Applied Machine Learning Fundamentals *** Probability Density Estimation (PDE)

Term: November 29, 2019

Contact:

Clemens Biehl

Moodle Forum

Do you have any questions?