Lecture 9:

Bayesian Estimation: Gibbs Sampling

Big Data and Machine Learning for Applied Economics Econ 4676

Ignacio Sarmiento-Barbieri

Universidad de los Andes

September 7, 2021

Agenda

- 1 Recap
 - Bayesian Estimation
 - Direct Sampling
- 2 Gibbs sampling
- 3 Review & Next Steps
- 4 Further Readings

Bayesian Estimation

► Bayes Theorem

$$\pi(\underline{\beta}|\underline{X}) = \frac{f(X|\beta)p(\underline{\beta})}{[m(X)]} \pi(\underline{\beta}|\underline{X}) \propto f(X|\underline{\beta}) P(\underline{\beta})$$
(1)

Interes B

 \blacktriangleright with m(X) is the marginal distribution of X, i.e.

$$m(X) = \int f(X|\beta)p(\beta)d\beta \tag{2}$$

▶ Bayes' theorem does not tell us what our beliefs should be, it tells us how they should change after seeing new information.

Frequentist Approach

► Suppose the model is

$$y_i = \widehat{\beta x_i} + u_i$$
(3)

$$u_i \sim N(0, \underline{\sigma^2}) \tag{4}$$

$$\sigma^2$$
 is known (5)

- ▶ Interest is on some form of $h(\beta)$ e.g. $|\beta|$
- Frequentists

$$\hat{\beta}_{MLE} = (X'X)^{-1}X'y$$

- then use the Delta method (or bootsrap?)
- Bayesians
 - ▶ Using simulation based methods → direct sampling algorithm

Simulation methods

- M Posteror

 E(Prosteror)
- ▶ Bayesians specify a prior distribution $\beta \sim N(\beta_0, \tau_+^2)$
- Use conjugate priors and get the posterior

$$\beta \mid Y, X \sim N \left(\frac{\frac{1}{\sigma^2} \sum_{i=1}^{N} y_i x_i + \frac{1}{\tau^2} \beta_0}{\frac{1}{\sigma^2} \sum_{i=1}^{N} x_i^2 + \frac{1}{\tau^2}} \right), \frac{1}{\frac{1}{\sigma^2} \sum_{i=1}^{N} x_i^2 + \frac{1}{\tau^2}} \right)$$
(6)

- \triangleright generate i.i.d. samples from the posterior distribution of β , $\pi(\beta|Y)$
- ightharpoonup get $h(\beta)$ e.g. $|\beta|$



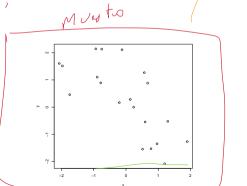
4/28

Example: Linear regression

- ▶ Goal: posterior mean for $|\beta|$ (quantile t = 50)
- ▶ I generate data $y_i x_i$ with

$$y_i = \beta x_i + u_i, \quad u_i \sim N(0, \sigma^2)$$

- N = 20
- $ightharpoonup eta_{
 m true} = -1$ and $\sigma^2 = 1$
- $\beta_0 = 0$ and $\tau = 100$



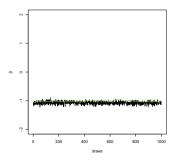
Par 6 (B) P(B) ~ N(0,1) pend st N(0,1)

LIK N -s post N(M,V) 20

Example: Linear regression

▶ Step 1: we generate S draws from the N(m,V), $\{\beta^s\}1,...,S$

Figure 1: Example of draws
$$\{\{\beta^s\}_{1,...,s}\}$$
, $S=1,000$



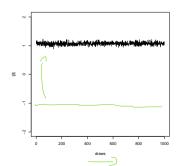


6/28

Example: Linear regression

- ▶ Step 2: we are interested in posterior moments of $|\beta|$.
- ► Turn draws into $\{|\beta|\}_{1,...,S}$

Figure 2: Example of draws $(\{|\beta^s|\}_{1,...,S})$



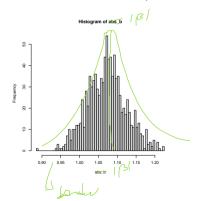
$$\beta^{2}$$

$$h(\beta) = \omega(\beta)$$

Example: Linear regression

► Histogram approximation to $\pi(|\beta| | Y)$ using $\{|\beta^s|\}_{1,...,S}$

Figure 3: Example of draws $(\{|\beta^s|\}_{1,...,S})$



Example: Linear regression

▶ The posterior mean of $|\beta|$ is approximated by

$$E_Y^{\beta}[|\beta|] \approx \frac{1}{S} \sum_{s=1}^{S} |\beta^s| = 1.0719$$

Numerical accuracy use CLT

► Consider now the linear regression model

$$y_i = \beta x_i + u_i, \quad \underline{u}_i \sim \underline{N}(0) \underline{\sigma}^2$$

- ightharpoonup but assume σ^2 is unknown
- ▶ The prior on β is the same as before

$$\beta \sim N(\beta_0, \tau^2)$$

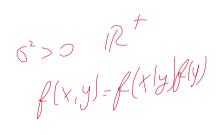
• We add now a prior on σ^2 is the Inverse-Gamma

$$\sigma^2 \sim IG(\underline{a}, \underline{b})$$

We want to know the joint posterior distribution

$$\pi(\beta, \sigma^2 \mid Y, X)$$

• Or we want to know the marginal distribution of β and $\underline{\sigma}^2$



 \blacktriangleright We know the conditional distribution of β

$$\beta | Y, X, \sigma^2 \sim N(m, V)$$

► We can show that that

$$\sigma^2 | Y, \beta \sim \text{IG} \left(\underline{a} + \underbrace{\frac{N}{2}}, b + \frac{1}{2} \sum_{i=1}^{N} (\underline{y_i - x_i \beta})^2 \right)$$

► That is, IG is conjugate prior for σ^2



Derivation of the conditional posterior distribution

$$p(\sigma^{2} | Y, X, \beta) \propto p\left(Y | X, \sigma^{2}, \beta\right) p\left(\sigma^{2}\right)$$

$$\propto \left(\frac{1}{\sqrt{2\pi\sigma^{2}}}\right)^{N} \exp\left(-\frac{1}{2}\sum_{i=1}^{N} \frac{(y_{i} - x_{i}\beta)^{2}}{\sigma^{2}}\right) \times \frac{b^{a}}{\Gamma(a)} \left(\sigma^{2}\right)^{-a-1} \exp\left(-\frac{b}{\sigma^{2}}\right)$$

$$\propto \left(\sigma^{2}\right)^{\left(-a-1-\frac{N}{2}\right)} \exp\left(-\frac{b}{\sigma^{2}} - \frac{1}{2}\sum_{i=1}^{N} \frac{(y_{i} - x_{i}\beta)^{2}}{\sigma^{2}}\right)$$

(keep terms that are related to
$$\sigma^2$$
)



- Derivation of the conditional posterior distribution
- We can see that the conditional posterior has the IG density form and we can easily deduce that

$$\sigma^2 | \Upsilon, \beta \sim IG(\bar{a} \, \bar{b})$$

▶ Where

$$\bar{a} = \underbrace{a + \frac{N}{2}}_{\bar{b} = b + \frac{1}{2}} \sum_{i=1}^{N} (y_i - x_i \beta)^2$$

$$(9)$$

▶ We know conditional posteriors. Can we use these to recover joint distribution?

- ▶ We know conditional posteriors. Can we use these to recover joint distribution?
- ► The answer turns out to be yes
- ▶ This is very cool. Joint distribution may be nasty and high-dimensional.
- ▶ But, Gibbs sampling allows us to break the nasty joint distribution piece by piece

Example: Linear regression

- ▶ In the context of linear regression example, the Gibbs algorithm works as below

Enter the following iteration with
$$\beta^0$$
 and $s=1$

$$1 \quad (\sigma^2)^s \sim p(\sigma^2|Y, \beta^{(s-1)})$$

$$2 \quad \beta^s) \sim p(\beta|Y, \sigma^{(s-1)})$$

- 3 Go to step 1 with i = i + 1 if i < S. Otherwise, exit loop
- At the end of the algorithm, you get draws $\{\beta^{(s)}, (\sigma^2)^{(s)}\}_{s=1,\ldots,s}$.

Example: Linear regression

► Under regular conditions,

$$\frac{1}{S - S_0} \sum_{s=S_{0+1}}^{S} h\left(\beta^{(s)}, \left(\sigma^2\right)^{(s)}\right) \rightarrow_{\text{a.s.}} \int h\left(\beta, \sigma^2 \middle| Y\right) d\beta d\sigma^2$$

- \blacktriangleright Where the first S_0 draws are discarded
- ► <u>CLT</u> also holds so that we can evaluate the <u>quality of the numerical approximation</u>

Example: Linear regression

ightharpoonup Like before, I generate data y_i and x_i with

$$y_i = \beta x_i + u_i, \quad u_i \sim N(0, \sigma^2)$$

with

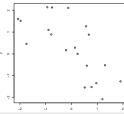
$$N = 100$$

$$ightharpoonup \beta_{\text{true}} = -1 \text{ and } \sigma^2 = 1$$

Prior for
$$\beta$$
: $\beta_0 = 0$ and $\tau = 1$

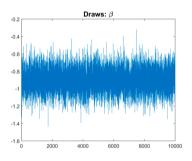
• Prior for
$$\sigma^2 : a = 10, b = 20$$

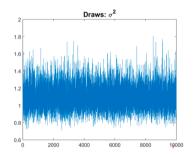
Start the algorithm with $\beta^{(0)} = 5$



Example: Linear regression

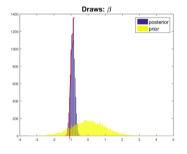
► Gibbs sampling: Posterior draws

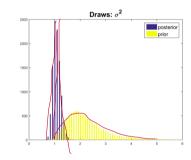




Example: Linear regression

► Gibbs sampling: Posterior draws





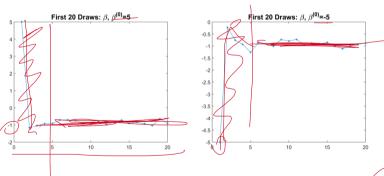
- ▶ Once we observe data, we update our belief accordingly.
- ▶ Distribution shrinks. Center of the distribution moved toward where the data generated from:

$$\beta_{\text{true}} = -1 \tag{10}$$

$$\sigma^2 = 1 \tag{11}$$

Initial value effect

- ► Initial value effect
- We start the algorithm with some arbitrary number β⁰. Theory tells us that this initial value should not matter in the long run.



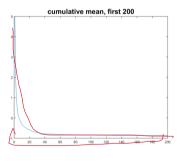
► To get rid of the initial value effect, we usually take out first *x* draws (say, 1,000 draws)

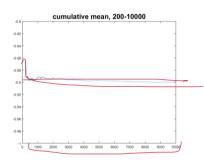
Checking convergence

- Checking convergence
- We start from arbitrary $\beta^{(0)}$
- We can view $\beta^{(0)}$ as a draw from the arbitrary distribution (does not have to be the distribution we are interested in).
- ► Theory tells us that the sequence from Gibbs algorithm converges to the joint posterior distribution.
 - After some iteration, $\beta^{(i)}$, $(\sigma^2)^i$ is a draw from $\pi(\beta, \sigma^2|Y)$
- ► How can we check the convergence? There are many ways.

Checking convergence: Running Mean Plots

Cumulative mean over draws



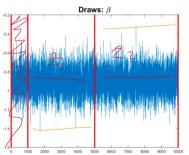


► This sequence behaves well in the sense that the Monte Carlo average converges to some number (posterior moment of interest) as the number of draws increases.

Checking convergence: Geweke's diagnostic check

- ightharpoonup Take out the first x draws
- ► Split draws into two non-overlapping parts (e.g. first 10% vs last 50%)

► Compare Z-score

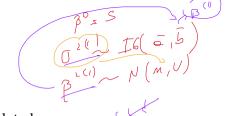


- ► Standardized mean (Z-score) of the left sample: -6.9363
- ► Standardized mean (Z-score) of the right sample: -6.9401
- (you can formally test via hypothesis testing).

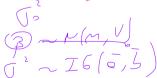
vs last 50%) $\frac{N}{\sqrt{2}}$ $\frac{N}{\sqrt{2}}$ $\frac{N}{\sqrt{2}}$



Gibbs sampling versus Direct sampling



- ► A sequence from the Gibbs sampling is serially correlated
 - Previous draw affects the current draw.
 - ▶ Gibbs sampler creates a Markov chain, For this reason, it belongs to the class of *Markov chain Monte Carlo (MCMC)* procedures.
- ▶ Recall Direct sampling procedure generates i.i.d. draws.



Gibbs sampling versus Direct sampling

- Both "Direct sampling" and "Gibbs sampling" generate draws that satisfy CLT.
- For draws from direct sampling:

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h\left(\beta^{i}\right) - \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \right) \rightarrow_{d} N\left(0 \middle|V_{\pi}\right)$$
where

For draws from Gibbs sampling (after discarding first few draws):

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h\left(\beta^{i}\right) - \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \right) \rightarrow_{d} N\left(0, V_{G}\right)$$

- $V_{\pi} < V_{G}$ To achieve the same level of approximation error, we need more draws X for Gibbs sampler.

Review

General Gibbs sampler

- Gibbs sampler works for more than two parameters case. Let be θ unknown parameter with dim $\theta > 1$
- ► Requirements:
 - Parameter vector θ can be partitioned into $\theta = (\theta_1, \theta_2, \dots, \theta_m)$
 - \triangleright For each s it is possible to generate draws of θ_s from the conditional distribution, $p(\theta_s|\theta_{s-1}, Y)$ where θ_{-s} denotes the vector θ without the partition θ_s
- ► Gibbs sampler: For s = 1, ..., S: 9_{L} ~ volor (NIC
 - ► Draw $\theta_1^{(s+1)}$ from the density $p(\theta_1|\theta_2^{(s)}, \theta_3^{(s)}, \dots, \theta_m^{(s)}, Y)$ ► Draw $\theta_2^{(s+1)}$ from the density $p(\theta_2|\theta_1^{(s+1)}, \theta_3^{(s)}, \dots, \theta_m^{(s)}, Y)$

 - Draw $\theta_m^{(s+1)}$ from the density $p\left(\theta_m | \theta_1^{(s)}, \theta_2^{(s+1)}, \theta_3^{(s+1)}, \dots, \theta_{m-1}^{(s\mp 1)}, Y\right)$

Review & Next Steps

- ► Next Class: Empirical Bayes
- Next Week: PS 2

Defer apa coler

Further Readings

- ► Casella, G., & Berger, R. L. (2002). Statistical inference (Vol. 2, pp. 337-472). Pacific Grove, CA: Duxbury. Chapter 7
- ► Hoff, P. D. (2009). A first course in Bayesian statistical methods (Vol. 580). New York: Springer.
- ▶ Roberts, G.O. and J.S. Rosenthal (2004): General State Space Markov Chains and MCMC algorithms, Probability Surveys, 1, 20–71.
- Geweke, J. (2005): Contemporary Bayesian Econometrics and Statistics, John Wiley & Sons.