Lecture 8:

Bayesian Estimation: Direct Sampling

Big Data and Machine Learning for Applied Economics Econ 4676

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Agenda

- 1 Bayesian Estimation
- 2 Simulation-based methods for Bayesian analysis
 - Direct Sampling
- 3 Recap
- 4 Further Readings

Bayesian Estimation

- ► The Bayesian approach to stats is fundamentally different from the classical approach we have been taking
- ► In the classical approach, the parameter β is thought to be an unknown, but fixed quantity, e.g., $X_i \sim f(β)$
- ▶ In the Bayesian approach β is considered to be a quantity whose variation can be described by a probability distribution (*prior distribution*)
- ► Then a sample is taken from a population indexed by β and the prior is updated with this sample
- ▶ The resulting updated prior is the *posterior distribution*

Bayes Approach

Bayes Theorem

$$\pi(\beta|X) = \frac{f(X|\beta)p(\beta)}{m(X)} \tag{1}$$

with m(X) is the marginal distribution of X, i.e.

$$m(X) = \int f(X|\beta)p(\beta)d\beta \tag{2}$$

It is important to note that Bayes' theorem does not tell us what our beliefs should be, it tells us how they should change after seeing new information.

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Frequentist Approach

 \blacktriangleright The interest is on β , frequentist estimation procedures give us that, for example

$$\hat{\beta}_{MLE} = (X'X)^{-1}X'y \tag{3}$$

- Now in Bayes world, I have the full distribution. Which β I use?
- ▶ I can use any moment, but usually the interest lies on

$$E(\beta) = \int \beta \pi(\beta|X) \tag{4}$$

▶ Why? note that if you use MSE as loss function, the Bayes estimate of the unknown parameter is the mean of the posterior distribution

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Bayesian Estimation

- ▶ We are going to have an overview simulation-based methods for Bayesian analysis.
 - 1 Direct sampling algorithm
 - 2 Gibbs sampling algorithm
- ► As a running example, we use the linear regression framework

$$y_i = \beta x_i + u_i , \quad u_i \sim N(0, \sigma^2)$$
 (5)

- with σ^2 known, and
- ▶ with prior distribution $\beta \sim N(\beta_0, \tau^2)$



- ▶ Using the knowledge of conjugate priors + the trick for exponentials
- \blacktriangleright The posterior distribution β follows the normal distribution:

$$\beta \mid Y, X \sim N \left(\frac{\frac{1}{\sigma^2} \sum_{i=1}^{N} y_i x_i + \frac{1}{\tau^2} \beta_0}{\frac{1}{\sigma^2} \sum_{i=1}^{N} x_i^2 + \frac{1}{\tau^2}}, \frac{1}{\frac{1}{\sigma^2} \sum_{i=1}^{N} x_i^2 + \frac{1}{\tau^2}} \right)$$
(6)

▶ We where able to characterize the full posterior distribution for the unknown object.

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- Suppose now, that our object of interest is not β per se, but some nonlinear function of unknown parameter β , e.g. $h(\beta)$.
- ► For example:

$$\blacktriangleright h(\beta) = \beta$$

$$\blacktriangleright h(\beta) = |\beta|$$

►
$$h(β) = α\%$$
 quantile of β

$$h(\beta) = \beta^3$$

$$h(\beta) = \beta_1 \beta_2$$

▶ The goal is to obtain posterior moments of $h(\beta)$.

Side note: Frequentist's approach

- Frequentist obtain the sampling distribution of $h(\beta)$ using the delta method:
- ▶ If we have

$$\sqrt{N} \left(\widehat{\beta} - \beta_0 \right) \to_d N \left(0, V_{\text{asy}} \right)$$
 (7)

► Then we

$$\sqrt{N} \left(h\left(\widehat{\beta}\right) - h(\beta_0) \to_d N\left(0, V_{\text{asy}}\left[h'(\beta_0)\right]^2\right]$$
 (8)

As $N \to \infty$ where n is the number of observations.



- ► Idea: Monte Carlo integration.
- ► Requirement
 - ► Know how to generate i.i.d. samples from the posterior distribution of β , $\pi(\beta|Y)$
- ► The requirement is satisfied for our linear regression example:
 - ightharpoonup The posterior distribution of β follows the normal distribution.
 - Most modern statistical program languages provide random number generators for many parametric distributions including the normal distribution.

▶ Direct sampling approach simply approximates the posterior expectations of a function $h(\beta)$ by

$$E_{Y}^{\beta}[h(\beta)] = \int h(\beta) \pi(\beta|Y) d\beta$$

$$\approx \frac{1}{S} \sum_{i=1}^{S} h(\beta^{i})$$
(9)

- ▶ Where β^i is i.i.d. samples from $\pi(\beta|Y)$
- ➤ S is "number of random samples from the posterior" or "number of generated draws" NOT the number of observations.



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- ▶ Provided that $E_Y^{\beta} \left[h \left(\beta \right)^2 \right] < \infty$,
- ▶ we can use the Strong Law of Large Numbers (SLLN)

$$\frac{1}{S} \sum_{i=1}^{S} h(\beta^{i}) \to a.s. \int h(\beta) p(\beta|Y) d\beta$$

and the Central Limit Theorem (CLT)

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h(\beta^{i}) - \int h(\beta) \pi(\beta|Y) d\beta \right) \rightarrow_{d} N(0, V_{\pi})$$

Where

$$V_{\pi} = \operatorname{Var}_{Y}^{\beta}(h(\beta)) = \int \left(h(\beta) - E_{Y}^{\beta}[h(\beta)]\right)^{2} \pi(\beta|Y) d\beta$$

► *S* is the number of simulated draws from the posterior distribution

▶ Note that we turned a complicated integration into a simple average

$$\frac{1}{S} \sum_{i=1}^{S} h(\beta^{i}) \to a.s. \int h(\beta) p(\beta|Y) d\beta$$
 (10)

- ▶ As the number of simulated draws increases, this simple average converges to the object of interest.
- ► Numerical accuracy?



- ▶ The CLT result provides a way to measure the numerical accuracy of this
- ► Monte Carlo approximation:

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h(\beta^{i}) - \int h(\beta) p(\beta|Y) d\beta \right) \rightarrow_{d} N(0, V_{\pi})$$
(11)

► That is,

$$\frac{1}{S} \sum_{i=1}^{S} h(\beta^{i}) \approx_{d} N\left(E_{Y}^{\beta}\left[h\left(\beta\right)\right], \frac{V_{\pi}}{S}\right)$$
(12)

- ▶ Where $V_{\pi} = \operatorname{Var}_{Y}^{\beta}(h(\beta))$. Posterior variance of $h(\beta)$ scaled by 1/S determines the numerical accuracy. As $S \to \infty$, numerical approximation goes to zero
- ► Trade-off
 - Large N: high computational cost (time) but more accurate approximation
 - Small N: low computation cost (time) but less accurate approximation

Example: Linear regression

► Consider the following linear regression model

$$y_i = \beta x_i + u_i, \quad \mathbf{u}_i \sim N(0, \sigma^2) \tag{13}$$

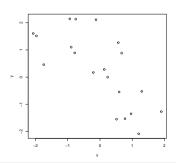
- with prior distribution $\beta \sim N(\beta_0, \tau^2)$, and suppose σ^2 is known.
- ▶ Then, we now all know that the posterior distribution β follows the normal distribution:

$$\beta|Y,X \sim N\left(\frac{\frac{1}{\sigma^2}\sum_{i=1}^{N}y_ix_i + \frac{1}{\tau^2}\beta_0}{\frac{1}{\sigma^2}\sum_{i=1}^{N}x_i^2 + \frac{1}{\tau^2}}, \frac{1}{\frac{1}{\sigma^2}\sum_{i=1}^{N}x_i^2 + \frac{1}{\tau^2}}\right)$$
(14)

- ▶ Goal: posterior mean and equal-tail-probability credible set for $|\beta|$
- ▶ I generate data y_i x_i with

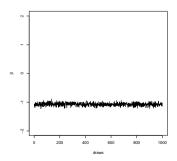
$$y_i = \beta x_i + u_i, \quad u_i \sim N(0, \sigma^2)$$

- N = 20
- $ightharpoonup eta_{\text{true}} = -1 \text{ and } \sigma^2 = 1$
- $\beta_0 = 0$ and i = 100



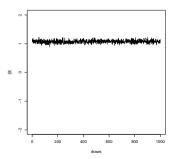
- ▶ Step 1: we generate N draws from the N(m, V), $\{\beta^i\}$ 1,...,S
- M = -1.07
- V = 0.0510

Figure 1: Example of draws $(\{\beta^i\}_{1,\dots,N})$, S = 1,000



- ▶ Step 2: we are interested in posterior moments of $|\beta|$.
- ► Turn draws into $\{|\beta|\}_{1,...,S}$

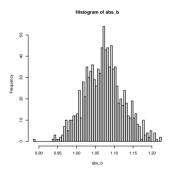
Figure 2: Example of draws $\left(\left\{|\beta^i|\right\}_{1,\dots,S}\right)$



Example: Linear regression

▶ Histogram approximation to $\pi(|\beta| | Y)$ using $\{|\beta^i|\}_{1,\dots,S}$

Figure 3: Example of draws $\left(\left\{|\beta^i|\right\}_{1,\dots,S}\right)$



Example: Linear regression

▶ The posterior mean of $|\beta|$ is approximated by

$$E_Y^{\beta}[|\beta|] \approx \frac{1}{S} \sum_{i=1}^{S} |\beta^i| = 1.0719$$
 (15)

▶ The 90% equal-tail-probability interval is approximated by

$$C_Y = [q_1, q_u] = [0.719, 1.441]$$
 (16)

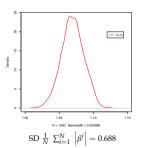
 \blacktriangleright Where q_1 and q_u such that

$$5\% = \frac{1}{S} \sum_{i=1}^{S} 1\{|\beta| < q_u\}$$
 (17)

- ▶ Numerical accuracy of $\frac{1}{N} \sum_{i=1}^{N} 1 \left| \beta^{i} \right|$
- We know that if we generate enough number of β^i , we get an accurate approximation to the posterior moments
- How many draws are enough?
- ▶ In other words, "Will I get different answer if I construct the same quantity using different set of draws $\{\beta^i\}$?
- ▶ Is S = 10 enough? Or, is S = 10,000 enough?

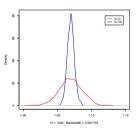
- ► To see the numerical error I generate 1,000 sets of $\{\beta^i\}_{i=1,\dots,S}$
- ► Compute 1,000 of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ to see how variable this Monte Carlo approximation with different *S*

Figure 4: Distribution of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ over $\{\beta^{i}\}_{i=1,\dots,S}$



- ► To see the numerical error I generate 1,000 sets of $\{\beta^i\}_{i=1,...,S}$
- ► Compute 1,000 of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ to see how variable this Monte Carlo approximation with different S

Figure 5: Distribution of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ over $\{\beta^{i}\}_{i=1,...,S}$

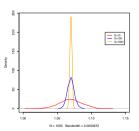


SD
$$\frac{1}{N} \sum_{i=1}^{N} |\beta^{i}| = 0.022$$



- ▶ To see the numerical error I generate 1,000 sets of $\{\beta^i\}_{i=1,\dots,S}$
- ► Compute 1,000 of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ to see how variable this Monte Carlo approximation with different S

Figure 6: Distribution of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ over $\{\beta^{i}\}_{i=1,...,S}$

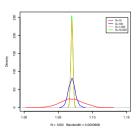


SD
$$\frac{1}{N} \sum_{i=1}^{N} |\beta^{i}| = 0.0074$$



- ► To see the numerical error I generate 1,000 sets of $\{\beta^i\}_{i=1,...,S}$
- ► Compute 1,000 of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ to see how variable this Monte Carlo approximation with different S

Figure 7: Distribution of $\frac{1}{S} \sum_{i=1}^{S} |\beta^{i}|$ over $\{\beta^{i}\}_{i=1,...,S}$



SD
$$\frac{1}{N} \sum_{i=1}^{N} \left| \beta^i \right| = 0.0023$$



- ▶ What do we try to capture in this exercise?
- We try to mimic the distribution of Monte Carlo approximation offered by

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h\left(\beta^{i}\right) - \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \right) \rightarrow_{d} N(0, V_{\pi})$$
 (18)

- ▶ All variation in this Monte Carlo approximation is due to numerical "simulation".
- ▶ Throughout this example, we fix $Y_{1:N}$, $X_{1:N}$ (not a sampling variation).

Recap

- ▶ If you know how to generate i.i.d draws from the posterior distribution of β ,
- \blacktriangleright You also can posterior moments of $h(\beta)$ by simple average:

$$\int h(\beta) p(\beta|Y) d\beta \approx \frac{1}{N} \sum_{i=1}^{N} h(\beta^{i})$$
(19)

► SLLN guarantees this Monte Carlo average to the right limit:

$$\frac{1}{N} \sum_{i=1}^{N} h\left(\beta^{i}\right) \to_{\text{a.s.}} \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \tag{20}$$

► CLT tells you that the Monte Carlo average always has a numerical error:

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h\left(\beta^{i}\right) - \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \right) \to_{d} N\left(0, V_{\pi}\right)$$
 (21)

▶ It is important to check how good is your numerical approximation

Review & Next Steps

- ▶ Direct Sampler
- ► Next Class: Gibbs Sampler

Further Readings

- ► Casella, G., & Berger, R. L. (2002). Statistical inference (Vol. 2, pp. 337-472). Pacific Grove, CA: Duxbury. Chapter 7
- ► Hoff, P. D. (2009). A first course in Bayesian statistical methods (Vol. 580). New York: Springer.