Lecture 9:

Bayesian Estimation: Direct Sampling

Big Data and Machine Learning for Applied Economics Econ 4676

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Agenda

- 1 Recap
 - Bayesian Estimation
 - Direct Sampling
- 2 Gibbs sampling
- 3 Review & Next Steps
- 4 Further Readings

Bayesian Estimation

Bayes Theorem

$$\pi(\beta|X) = \frac{f(X|\beta)p(\beta)}{m(X)} \tag{1}$$

with m(X) is the marginal distribution of X, i.e.

$$m(X) = \int f(X|\beta)p(\beta)d\beta \tag{2}$$

▶ Bayes' theorem does not tell us what our beliefs should be, it tells us how they should change after seeing new information.

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Frequentist Approach

► Suppose the model is

$$y_i = \beta x_i + u_i \tag{3}$$

$$u_i \sim N(0, \sigma^2) \tag{4}$$

$$\sigma^2$$
 is known (5)

- ▶ Interest is on some form of $h(\beta)$ e.g. $|\beta|$
- Frequentists
 - $\hat{\beta}_{MLE} = (X'X)^{-1}X'y$
 - ▶ then use the Delta method (or bootsrap?)
- Bayesians
 - ▶ Using simulation based methods → direct sampling algorithm

Simulation methods

- ▶ Bayesians specify a prior distribution $\beta \sim N(\beta_0, \tau^2)$
- Use conjugate priors and get the posterior

$$\beta|Y,X \sim N\left(\frac{\frac{1}{\sigma^2}\sum_{i=1}^{N}y_ix_i + \frac{1}{\tau^2}\beta_0}{\frac{1}{\sigma^2}\sum_{i=1}^{N}x_i^2 + \frac{1}{\tau^2}}, \frac{1}{\frac{1}{\sigma^2}\sum_{i=1}^{N}x_i^2 + \frac{1}{\tau^2}}\right)$$
(6)

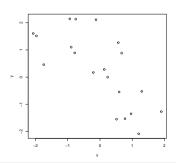
- generate i.i.d. samples from the posterior distribution of β , $\pi(\beta|Y)$
- ightharpoonup get $h(\beta)$ e.g. $|\beta|$

Example: Linear regression

- ▶ Goal: posterior mean for $|\beta|$ (quantile t = 50)
- ▶ I generate data y_i x_i with

$$y_i = \beta x_i + u_i, \quad u_i \sim N(0, \sigma^2)$$

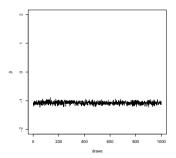
- N = 20
- $ightharpoonup eta_{
 m true} = -1$ and $\sigma^2 = 1$
- $\beta_0 = 0 \ and \ \tau = 100$



Example: Linear regression

▶ Step 1: we generate S draws from the N(m, V), $\{\beta^s\}1,...,S$

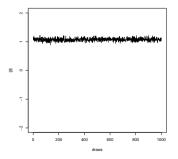
Figure 1: Example of draws $\left(\left\{\beta^{s}\right\}_{1,\dots,S}\right)$, S=1,000



Example: Linear regression

- ▶ Step 2: we are interested in posterior moments of $|\beta|$.
- ► Turn draws into $\{|\beta|\}_{1,...,S}$

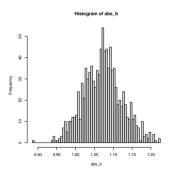
Figure 2: Example of draws $(\{|\beta^s|\}_{1,...,S})$



Example: Linear regression

► Histogram approximation to $\pi(|\beta| | Y)$ using $\{|\beta^s|\}_{1,\dots,S}$

Figure 3: Example of draws $(\{|\beta^s|\}_{1,...,S})$



Example: Linear regression

▶ The posterior mean of $|\beta|$ is approximated by

$$E_Y^{\beta}[|\beta|] \approx \frac{1}{S} \sum_{s=1}^{S} |\beta^s| = 1.0719$$
 (7)

Numerical accuracy use CLT

► Consider now the linear regression model

$$y_i = \beta x_i + u_i, \quad u_i \sim N(0, \sigma^2)$$

- ightharpoonup but assume σ^2 is unknown
- ▶ The prior on β is the same as before

$$\beta \sim N(\beta_0, \tau^2)$$

• We add now a prior on σ^2 is the Inverse-Gamma

$$\sigma^2 \sim IG(a,b)$$

▶ We want to know the joint posterior distribution

$$\pi(\beta, \sigma^2 | Y, X)$$

• Or we want to know the marginal distribution of β and σ^2

 \blacktriangleright We know the conditional distribution of β

$$\beta | Y, X, \sigma^2 \sim N(m, V)$$

▶ We can show that that

$$\sigma^{2}|Y,\beta \sim IG\left(a+\frac{N}{2}, b+\frac{1}{2}\sum_{i=1}^{N}(y_{i}-x_{i}\beta)^{2}\right)$$

► That is, IG is conjugate prior for σ^2



Derivation of the conditional posterior distribution

$$p(\sigma^{2} | Y, X, \beta) \propto p\left(Y | X, \sigma^{2}, \beta\right) p\left(\sigma^{2}\right)$$

$$\propto \left(\frac{1}{\sqrt{2\pi\sigma^{2}}}\right)^{N} \exp\left(-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_{i} - x_{i}\beta)^{2}}{\sigma^{2}}\right) \times \frac{b^{a}}{\Gamma(a)} \left(\sigma^{2}\right)^{-a-1} \exp\left(-\frac{b}{\sigma^{2}}\right)$$

$$\propto \left(\sigma^{2}\right)^{\left(-a - 1 - \frac{N}{2}\right)} \exp\left(-\frac{b}{\sigma^{2}} - \frac{1}{2} \sum_{i=1}^{N} \frac{(y_{i} - x_{i}\beta)^{2}}{\sigma^{2}}\right)$$

$$(8)$$

(keep terms that are related to σ^2)



- Derivation of the conditional posterior distribution
- We can see that the conditional posterior has the IG density form and we can easily deduce that

$$\sigma^2 | \Upsilon, \beta \sim IG(\bar{a} \, \bar{b})$$

▶ Where

$$\bar{a} = a + \frac{N}{2}$$

$$\bar{b} = b + \frac{1}{2} \sum_{i=1}^{N} (y_i - x_i \beta)^2$$
(9)

▶ We know conditional posteriors. Can we use these to recover joint distribution?

- ▶ We know conditional posteriors. Can we use these to recover joint distribution?
- ► The answer turns out to be yes
- ► This is very cool. Joint distribution may be nasty and high-dimensional.
- ▶ But, Gibbs sampling allows us to break the nasty joint distribution piece by piece

Example: Linear regression

- ▶ In the context of linear regression example, the Gibbs algorithm works as below
- ▶ Enter the following iteration with β^0 and s = 1

 - 3 Go to step 1 with i = i + 1 if i < S. Otherwise, exit loop
- ► At the end of the algorithm, you get draws $\{\beta^{(s)}, (\sigma^2)^s\}_{s=1,...,S}$.

Example: Linear regression

► Under regular conditions,

$$\frac{1}{S-S_0} \sum_{s=S_{0+1}}^{S} h\left(\beta^{(s)}, \left(\sigma^2\right)^{(s)}\right) \rightarrow_{\text{a.s.}} \int h\left(\beta, \sigma^2 \middle| Y\right) d\beta d\sigma^2$$

- \blacktriangleright Where the first S_0 draws are discarded
- ▶ CLT also holds so that we can evaluate the quality of the numerical approximation

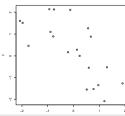
Example: Linear regression

ightharpoonup Like before, I generate data y_i and x_i with

$$y_i = \beta x_i + u_i, \quad u_i \sim N(0, \sigma^2)$$

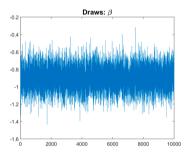
with

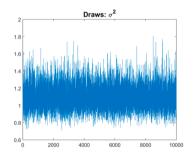
- N = 100
- $ightharpoonup \beta_{\text{true}} = -1 \text{ and } \sigma^2 = 1$
- Prior for β : $\beta_0 = 0$ and $\tau = 1$
- Prior for σ^2 : a = 10, b = 20
- Start the algorithm with $\beta^{(0)} = 5$



Example: Linear regression

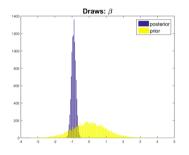
► Gibbs sampling: Posterior draws

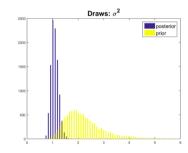




Example: Linear regression

► Gibbs sampling: Posterior draws





- Once we observe data, we update our belief accordingly.
- ▶ Distribution shrinks. Center of the distribution moved toward where the data generated from:

$$\beta_{\text{true}} = -1 \tag{10}$$

$$\sigma^2 = 1 \tag{11}$$

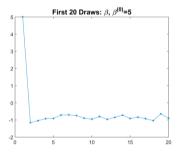
Sarmiento-Barbieri (Uniandes)

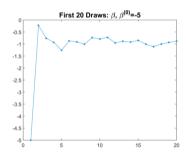
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Initial value effect

- ► Initial value effect
- We start the algorithm with some arbitrary number β⁰. Theory tells us that this initial value should not matter in the long run.





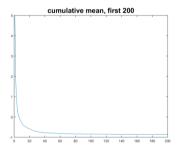
► To get rid of the initial value effect, we usually take out first *x* draws (say, 1,000 draws)

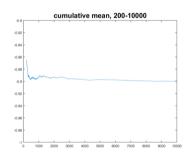
Checking convergence

- Checking convergence
- We start from arbitrary $\beta^{(0)}$
- We can view $\beta^{(0)}$ as a draw from the arbitrary distribution (does not have to be the distribution we are interested in).
- ► Theory tells us that the sequence from Gibbs algorithm converges to the joint posterior distribution.
 - After some iteration, $\beta^{(i)}$, $(\sigma^2)^i$ is a draw from $\pi(\beta, \sigma^2|Y)$
- ► How can we check the convergence? There are many ways.

Checking convergence: Running Mean Plots

Cumulative mean over draws

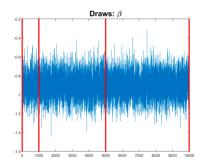




► This sequence behaves well in the sense that the Monte Carlo average converges to some number (posterior moment of interest) as the number of draws increases.

Checking convergence: Geweke's diagnostic check

- ► Take out the first *x* draws
- ▶ Split draws into two non-overlapping parts (e.g. first 10% vs last 50%)
- ► Compare Z-score



- ► Standardized mean (Z-score) of the left sample: -6.9363
- ► Standardized mean (Z-score) of the right sample: -6.9401
- (you can formally test via hypothesis testing).



Gibbs sampling versus Direct sampling

- ▶ A sequence from the Gibbs sampling is serially correlated
 - Previous draw affects the current draw.
 - ► Gibbs sampler creates a Markov chain. For this reason, it belongs to the class of *Markov chain Monte Carlo (MCMC)* procedures.
- ▶ Recall Direct sampling procedure generates i.i.d. draws.

Gibbs sampling versus Direct sampling

- ▶ Both "Direct sampling" and "Gibbs sampling" generate draws that satisfy CLT.
- ► For draws from direct sampling:

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h\left(\beta^{i}\right) - \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \right) \to_{d} N\left(0, V_{\pi}\right)$$

► For draws from Gibbs sampling (after discarding first few draws):

$$\sqrt{S} \left(\frac{1}{S} \sum_{i=1}^{N} h\left(\beta^{i}\right) - \int h\left(\beta\right) \pi\left(\beta|Y\right) d\beta \right) \to_{d} N\left(0, V_{G}\right)$$

- $ightharpoonup V_{\pi} < V_{G}$
- ► To achieve the same level of approximation error, we need more draws *X* for Gibbs sampler.

Review

General Gibbs sampler

- ▶ Gibbs sampler works for more than two parameters case. Let be θ unknown parameter with dim $\theta > 1$
- ► Requirements:
 - Parameter vector θ can be partitioned into $\theta = (\theta_1, \theta_2, ..., \theta_m)$
 - For each s it is possible to generate draws of θ_s from the conditional distribution, $p(\theta_s|\theta_{s-1}, Y)$ where θ_{-s} denotes the vector θ without the partition θ_s
- ▶ Gibbs sampler: For s = 1, ..., S:
 - ▶ Draw $\theta_1^{(s+1)}$ from the density $p(\theta_1|\theta_2^{(s)},\theta_3^{(s)},\ldots,\theta_m^{(s)},Y)$
 - ▶ Draw $\theta_2^{(s+1)}$ from the density $p(\theta_2|\theta_1^{(s+1)},\theta_3^{(s)},\ldots,\theta_m^{(s)},Y)$

 - Draw $\theta_m^{(s+1)}$ from the density $p(\theta_m | \theta_1^{(s)}, \theta_2^{(s+1)}, \theta_3^{(s+1)}, \dots, \theta_{m-1}^{(s+1)}, Y)$

Review & Next Steps

► **Next Class:** Empirical Bayes

► Next Week: PS 2

Further Readings

- ► Casella, G., & Berger, R. L. (2002). Statistical inference (Vol. 2, pp. 337-472). Pacific Grove, CA: Duxbury. Chapter 7
- ► Hoff, P. D. (2009). A first course in Bayesian statistical methods (Vol. 580). New York: Springer.
- ▶ Roberts, G.O. and J.S. Rosenthal (2004): General State Space Markov Chains and MCMC algorithms, Probability Surveys, 1, 20–71.
- Geweke, J. (2005): Contemporary Bayesian Econometrics and Statistics, John Wiley & Sons.