

## 0.1 Wiener's integral

Consider an *unconstrained* Brownian particle, moving on the real line, starting in  $x_0$  at  $t_0$ . By solving the diffusion equation we found that the probability of finding the particle in  $[x, x + dx]$  at time  $t > t_0$  is given by the **propagator**:

$$\begin{aligned}\mathbb{P}\{x(t) \in [x, x + dx] | x(t_0) = x_0\} &= W(x, t | x_0, t_0) dx = \\ &= \frac{1}{\sqrt{4\pi D(t - t_0)}} \exp\left(-\frac{(x - x_0)^2}{4D(t - t_0)}\right) dx \quad (1)\end{aligned}$$

By integrating (1) we can then find the probability of finding the particle inside an interval  $[A, B]$  at time  $t$ :

$$\mathbb{P}\{x(t) \in [A, B] | x(t_0) = x_0\} = \int_A^B dx W(x, t | x_0, t_0) \quad t > t_0$$

We are now interested in computing the expected value  $\langle f \rangle$  of **functionals**  $f$  of the trajectory, i.e. of quantities depending on several (or all) points of the trajectory  $x(\tau)$  of a Brownian particle.

- The simplest example is the **correlation function**, which is defined as the product of the particle's position at two different times  $t_1 < t_2$ :

$$f(\{x(t_1), x(t_2)\}) = x(t_1)x(t_2) \quad t_1 < t_2$$

- A more general (and difficult) case is given by a function of the *entire* trajectory, such as:

$$f(\{x(\tau) : 0 < \tau \leq \tau\}) = g\left(\int_0^t x(\tau)a(\tau) d\tau\right) \quad a, g: \mathbb{R} \rightarrow \mathbb{R}$$

### 0.1.1 Functions of a discrete number of points

Let's start from the simplest case, and consider the **correlation function**:

$$f(\{x(t_1), x(t_2)\}) = x(t_1)x(t_2) \quad t_1 < t_2$$

To compute  $\langle f \rangle$  we will need the *joint probability distribution*  $g(x_1, x_2)$  that gives the probability of  $x(t_1)$  being “close to”  $x_1$  **and**  $x(t_2)$  “close to”  $x_2$  *for the same trajectory*. Let us denote the three events of interest:

$A$ : Particle starts in  $x_0$  at  $t_0$

$B$ : Particle is close to  $x_1$  at  $t_1$  ( $x(t_1) \in [x_1, x_1 + dx_1]$ )

$C$ : Particle is close to  $x_2$  at  $t_2$  ( $x(t_2) \in [x_2, x_2 + dx_2]$ )

We are interested in the joint probability  $\mathbb{P}(C, B|A)$  (the order is defined by  $t_2 > t_1 > t_0$ ). From probability theory:

$$\mathbb{P}(C, B|A) = \mathbb{P}(C|B, A)\mathbb{P}(B|A)$$

We already know how to compute probabilities like  $\mathbb{P}(B|A)$ , but not like  $\mathbb{P}(C|B, A)$ . Fortunately, that is not needed.

Recall, in fact, that Brownian motion is a *Markovian process*, meaning that the *future* depends only on the *present state*, i.e. the particle *has no memory*. So, **subsequent displacements are independent**: the probability of the particle going from  $x_1$  to  $x_2$  is the same whether it has started at  $x_0$  or at any other point  $\tilde{x}_0$ . In other words, if we take the *present state* as the particle being in  $x_1$  at  $t_1$ , the future (position at  $t_2 > t_1$ ) depends only on that, and not on the past (position at  $t_0$ ). So:

$$\mathbb{P}(C|B, A) = \mathbb{P}(C|B)$$

leading to:

$$\mathbb{P}(C, B|A) = \mathbb{P}(C|B)\mathbb{P}(B|A)$$

Inserting the *propagators* (1):

$$d\mathbb{P}_{t_1, t_2}(x_1, x_2) \equiv W(x_2, t_2|x_1, t_1)W(x_1, t_1|x_0, t_0) dx_1 dx_2$$

This is the joint probability we need to compute  $\langle f \rangle$ . Of course, nothing stops us from considering  $N$  “jumps” instead of only 2:

$$\begin{aligned} d\mathbb{P}_{t_1, \dots, t_n}(x_1, \dots, x_n|x_0, t_0) &\equiv W(x_n, t_n|x_{n-1}, t_{n-1}) \dots W(x_1, t_1|x_0, t_0) dx_1 dx_2 \dots dx_n = \\ &= \exp\left(-\sum_{i=1}^n \frac{(x_i - x_{i-1})^2}{4\pi D \Delta t_i}\right) \prod_{i=1}^n \frac{dx_i}{\sqrt{4\pi D \Delta t_i}} \end{aligned} \quad (2)$$

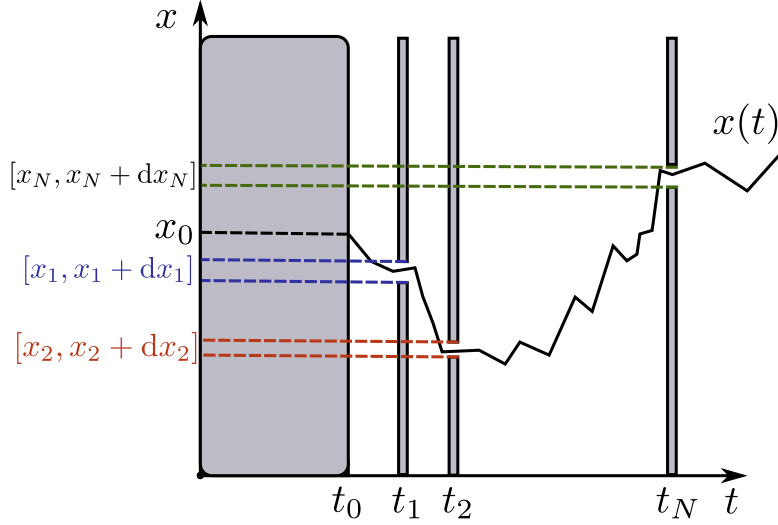
Then, the average of a generic function  $f(x(t_1), \dots, x(t_n))$  of the positions of the particle at times  $t_1 < t_2 < \dots < t_n$  is defined as:

$$\langle f(x(t_1), \dots, x(t_n)) \rangle_w = \int_{\mathbb{R}^n} f(x_1, \dots, x_n) d\mathbb{P}_{t_1, \dots, t_n}(x_1, \dots, x_n|x_0, t_0)$$

### 0.1.2 Functionals of the whole trajectory

The quantity in (2) can be interpreted as the *infinitesimal volume element* spanned by all the *trajectories* passing through a set of *tiny gates*, as represented in figure 1.

The underlying idea is that *probabilities* satisfy the axioms of *measures*, that is *functions that assign a measure*, i.e. a generalization of “size”, *to all sets included in a specific collection*.



**Figure (1)** – All trajectories that pass through the set of *gates*  $[x_i, x_i + dx_i]$  at times  $t_i$  (such as the  $x(t)$  here represented) contribute to the *volume*  $d\mathbb{P}_{t_1, \dots, t_n}(x_1, \dots, x_N)$

We now try to formalize this idea in order to extend the results of the previous section to the case of functions depending on a *infinite* number of trajectory points.

1. **Space definition.** Let  $T \subset \mathbb{R}$  (**index set**), denote with  $\mathbb{R}^T$  the set of all functions (**stochastic processes**)  $k: T \rightarrow \mathbb{R}$ . The idea is that an element of  $\mathbb{R}^T$  is a collection of *random variables* indexed by  $T$ .

In our case  $T$  is a collection of time instants (e.g.  $T = [0, +\infty)$ ) and a generic element of  $\mathbb{R}^T$  is made of *all the traversed points of a trajectory at times  $T$* :

$$\{x(t): t \in T\} \in \mathbb{R}^T$$

2. **Probability measure on finite points.** The expression in (2), as observed, allows us to *measure* the volumes spanned by trajectories traversing a set of gates. Let's formalize this idea. Consider a *finite* set of times  $T = \{t_i\}_{i=1, \dots, n}$  with  $n \in \mathbb{N}, t_i \in \mathbb{R}$  and  $t_1 < t_2 < \dots < t_n$ , each associated to a *gate*  $H_i = [a_i, b_i]$ , with  $a_i, b_i \in \mathbb{R}$  and  $a_i < b_i$ . All the trajectories  $\mathbb{R}^T$  traversing each  $H_i$  at a time  $t_i \in T$  span a **cylindrical set**  $A$  of the form:

$$A = \{x(t): x(t_1) \in H_1, \dots, x(t_n) \in H_n\} \subset \mathbb{R}^T$$

Using (2) and integrating over the *gates* we can define the measure of  $A$ -like sets as:

$$P_w(A) \equiv \int_{\mathbb{R}^n} d\mathbb{P}_{t_1, \dots, t_n}(x_1, \dots, x_n | x_0, t_0) \mathbb{I}_{H_1}(x_1) \dots \mathbb{I}_{H_n}(x_n)$$

where  $\mathbb{I}_{H_i}(x_i)$  are **characteristic functions** of the gates:

$$\mathbb{I}_{H_i}(x) = \begin{cases} 1 & x \in H_i \\ 0 & \text{otherwise} \end{cases}$$

In our case  $H_i$  are just intervals, and so:

$$P_w(A) \equiv \mathbb{P}_{t_1, \dots, t_n}(A) \stackrel{(2)}{=} \int_{H_1} dx_1 \int_{H_2} dx_2 \cdots \int_{H_n} dx_n \prod_{i=1}^n \frac{1}{\sqrt{4\pi D \Delta t_i}} \exp\left(-\frac{(x_i - x_{i-1})^2}{4D \Delta t_i}\right) \quad (3)$$

with  $\Delta t_i = t_i - t_{i-1}$ .

3. **Generalization on infinite points.** Note that (3) holds for any  $n$ . So, using  $A$ -like sets, we can construct a  $\sigma$ -algebra<sup>1</sup>  $\mathcal{F}$  of  $\mathbb{R}^T$ . Then, by applying **Kolmogorov extension theorem** we can extend the measure  $P_w$  we just found to the entire  $\mathcal{F}$ .
4. **Probability space.** We now have a set of *all possible outcomes*  $\mathbb{R}^T$  (in our case, all the possible trajectories that can be produced by a Brownian motion). We also have the collection of all *events*  $\mathcal{F}$ , that is subsets of  $\mathbb{R}^T$  for which is meaningful to assign a *probability measure*  $P_w: \mathcal{F} \rightarrow [0, 1]$ . The triad  $(\mathbb{R}^T, \mathcal{F}, P_w)$  forms a **probability space**, that gives a rigorous meaning to the concept of “computing the probability of a trajectory”.

The **measure** so obtained is called **Wiener measure**, and denoted as the following:

$$P_w(A) \equiv \int_A d_w x(\tau)$$

Then we can compute expected values. For example, if  $f(\{x(\tau): \tau \in T\})$  is a function depending on the points traversed at times in a set  $T$ , then:

$$\langle f \rangle_w \equiv \int_{\mathbb{R}^T} f(x(\tau)) d_w x(\tau) \quad T = [0, \infty)$$

Note that the Wiener measure *exists* and it's well defined (Kolmogorov's theorem), but we know it explicitly only in specific *finite cases*. So, to compute the expected value of functionals  $F(\{x(t)\})$  over *continuous trajectories* we first *discretize* the trajectory, and then take a *continuum limit*.

1. Suppose we have a functional  $F(\{x(\tau): 0 < \tau < t\})$ , and we want to compute  $\langle F \rangle$ .
2. We *discretize* the problem by *arbitrarily* subdividing the time interval  $[0, t]$  in  $n$  parts  $0 = t_0 < t_1 < t_2 < \cdots < t_n = t$ . Then we consider an *approximated* functional  $F_N(\{x(t_0), \dots, x(t_n)\})$  (for example approximating the path  $x(\tau)$  with a piecewise linear function, depending only on  $x(t_0), \dots, x(t_n)$ ), so that:

$$F = \lim_{N \rightarrow \infty} F_N$$

where  $N \rightarrow \infty$  means that  $\max \Delta t_i \rightarrow 0$ , with  $\Delta t_i = t_i - t_{i-1}$ . This limit needs to be properly defined (by using the Wiener measure to define a norm in a space of *integrable* functionals, etc.), but we will not do that here.

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<sup>1</sup> A  $\sigma$ -algebra on a set  $X$  is a collection  $\Sigma$  of subsets of  $X$  that includes  $X$  itself, is closed under complement, and is closed under countable unions

3. Then the **Weiner path integral** is defined as:

$$\begin{aligned}\langle F \rangle_w &= \int d_w x(\tau) F(\{x(\tau) : 0 < \tau < t\}) \equiv \lim_{N \rightarrow \infty} \langle F_N \rangle_w = \\ &= \lim_{N \rightarrow \infty} \int d\mathbb{P}_{t_1, \dots, t_n} (x_1, \dots, x_n | x_0, t_0) F_N(x(t_0), \dots, x(t_n))\end{aligned}$$

Geometrically, we are evaluating  $F$  for every possible Brownian path  $x(\tau)$ , and then averaging all these results, each weighted by the *probability* of the corresponding path.

**Example 1** (Correlation function and ESCK property):

As expected, the more general definition of the Wiener measure - involving the continuum limit  $N \rightarrow \infty$  - reduces to (2) when evaluated for a function depending only on a finite set of particle's positions.

For example, consider the expected value of the *correlation function* (assume the particle starting in 0 at time 0 for simplicity):

$$\begin{aligned}\langle x(t'_1)x(t'_2) \rangle &= \int d_w x x_1(t'_1)x_2(t'_2) = \\ &= \lim_{N \rightarrow \infty} \int d\mathbb{P}_{t_1, \dots, t_N} (x_1, \dots, x_N | 0, 0) x(t_k)x(t_n) =\end{aligned}$$

where we chose the discretization so that  $t_k = t'_1$  and  $t_n = t'_2$ . Then, by expanding the measure and applying the ESCK property we get (the limit is omitted):

$$\begin{aligned}&= \int dx_1 \dots dx_N W(x_N, t_N | x_{N-1}, t_{N-1}) \dots W(x_1, t_1 | 0, 0) x_k x_n = \\ &\stackrel{(a)}{=} \int dx_k dx_n W(x_n, t_n | x_k, t_k) W(x_k, t_k | 0, 0) x_k x_n\end{aligned}$$

which is the same result we could have obtained directly from (2):

$$\langle x(t'_1)x(t'_2) \rangle = \int dx'_1 dx'_2 W(x'_2, t'_2 | x'_1, t'_1) W(x'_1, t'_1 | 0, 0) x'_1 x'_2$$

## 0.2 Change of random variables

Consider a random variable  $X \sim q(x)$ , with  $q(x)$  being a generic distribution (e.g.  $q(x) = \mu e^{-\mu x}$ ). Now consider a function  $y(x)$ , e.g.  $y(x) = x^2$ .  $Y$  is then a new random variable, with a certain distribution  $p(y)$ . We now want to compute  $p(y)$  starting from  $q(x)$  and  $y(x)$ .

Suppose that  $y(x)$  is invertible. Then, if we extract a value from  $X$ , it will be inside  $[x, x + dx]$  with a probability  $q(x) dx$ . Knowing  $X$ , we can use the relation  $y(x)$  to uniquely determine  $Y$ , that will be in  $[y, y + dy]$  with the same probability. So, the following holds:

$$q(x) dx = p(y) dy \tag{4}$$

We can compute  $dy$  by *nudging*  $y(x)$ , and expanding in Taylor series:

$$y(x + dx) \equiv y + dy + O(dy^2) = y(x) + \underbrace{dx y'(x)}_{dy} + O(dx^2)$$

and so  $dy = dx y'(x)$ . Substituting in (4) we get:

$$q(x) dx = p(y) dy = p(y(x)) y'(x) dx \Rightarrow p(y) = q(x(y)) \frac{dx}{dy}(x(y)) \quad (5)$$

For a more general  $y(x)$ ,  $x \sim q(x)$  and  $y = y(x)$ , the expected value of a function  $f$  in terms of  $q(x)$ :

$$\begin{aligned} \langle f(y) \rangle &= \int_{\mathbb{R}} dx f(y(x)) q(x) = \\ &= \int_{\mathbb{R}} dx f(y(x)) q(x) \underbrace{\int_{\mathbb{R}} dz \delta(z - y(x))}_{=1} = \\ &\stackrel{(a)}{=} \int_{\mathbb{R}} dz f(z) \underbrace{\int_{\mathbb{R}} dx q(x) \delta(z - y(x))}_{\langle \delta(z - y(x)) \rangle_{q(x)}} \end{aligned} \quad (6)$$

where in (a) we used the fact that  $\delta(z - y(x)) = 1$  only when  $z = y(x)$ , and it's 0 otherwise, and so:

$$f(y(x)) = \int_{\mathbb{R}} dz f(z) \delta(z - y(x))$$

Of course we can rewrite  $\langle f \rangle$  directly in terms of  $p(y)$ :

$$\langle f(y) \rangle = \int_{\mathbb{R}} dy f(y) p(y) \quad (7)$$

Comparing (8) with (7) and renaming  $y \rightarrow z$  leads to:

$$p(z) = \int_{\mathbb{R}} dx q(x) \delta(z - y(x)) = \langle \delta(z - y(x)) \rangle_{q(x)} \quad (8)$$

which, in general, is not the same as the previously obtained result:

$$p(z) \neq q(x(z)) \frac{dx(z)}{dz}$$

To retrieve this special case we must assume  $y(x)$  to be invertible, with inverse  $x(y)$ . This means that  $\text{sgn } y'(x) = A$ , with  $A \in \mathbb{R} \setminus \{0\}$  constant.

We want now to compute  $\delta(z - y(x))$  in this case. Recall that  $\delta \circ g$ , if  $g$  is a continuously differentiable function with  $g(x_0) = 0$  and  $g'(x) \neq 0 \forall x$  is:

$$\delta(g(x)) = \frac{\delta(x - x_0)}{|g'(x_0)|}$$

So, if we let  $g(x) = z - y(x)$ , the only zero is at  $x = x(z)$ , as then  $y(x(z)) = z$ . So:

$$\delta(z - y(x)) = \frac{\delta(x - x(z))}{|y'(x(z))|}$$

Substituting back in (8):

$$\begin{aligned} p(z) &= \left\langle \frac{\delta(x - x(z))}{|y'(x(z))|} \right\rangle_{q(x)} = \int_{\mathbb{R}} dx q(x) \frac{\delta(x - x(z))}{|y'(x(z))|} = q(x(z)) |y'(x(z))|^{-1} = \\ &= q(x(z)) \frac{dx(y)}{dy} \Big|_{y=x(z)} \end{aligned} \quad (9)$$

which is the same rule found in (5).

### 0.3 The 1st integral

Thanks to the Wiener measure we have a way to assign probabilities to *paths*  $x(\tau)$ . We can recover from this the *transition probabilities* we started from, by considering the functional that *evaluates* a path at an instant  $t$ :  $x(\tau) \mapsto x(t) \equiv x_t$ . Then, by applying (8) we can compute the distribution followed by  $x_t$ :

$$p(x_t) = W(x_t, t|0, 0) = \langle \delta(x_t - x(\tau)) \rangle_w = \int d_W x(\tau) \delta(x_t - x(\tau)) \quad (10)$$

(The starting condition  $x(t_0) = x_0$  is *contained* in the definition of the measure  $d_W x(\tau)$ ).

So we can now write:

$$\begin{aligned} W(x, t|0, 0) &= \int d_W x \delta(x(t) - x) = \\ &= \text{“} \lim_{N \rightarrow \infty} \text{”} \int \prod_{i=1}^{N+1} \frac{dx_i}{\sqrt{4\pi D \Delta t_i}} \exp \left( - \sum_{i=1}^N \frac{(x_i - x_{i-1})^2}{4D \Delta t_i} \right) \delta(x_{N+1} - x) \end{aligned}$$

where  $t_n = t$ ,  $x(t_n) = x_{N+1}$ .

Recall that:

$$W(x_t, t|0, 0) = \frac{1}{\sqrt{4\pi D t}} \exp \left( - \frac{x_t^2}{4D t} \right)$$

If we set  $x_t = 0$  (for simplicity), we get:

$$W(0, t|0, 0) = \frac{1}{\sqrt{4\pi D t}} \quad (11)$$

As an exercise to get some familiarity with Wiener integrals, we want now to *re-derive* this result, by evaluating the Wiener path integral in (10), with  $x_t = 0$ :

$$W(0, t|0, 0) = \langle \delta(0 - x(\tau)) \rangle_w = \int d_W x(\tau) \delta(x(\tau)) \equiv I_1 \quad (12)$$

First, it is convenient to establish some additional notation. Let  $T = [0, \infty)$ . We denote with  $\mathcal{C}\{0, 0; t'\}$  the subset of trajectories in  $\mathbb{R}^T$  starting from  $x = 0$  at  $t = 0$ , and lasting a time span  $t'$ . Then,  $\mathcal{C}\{0, 0; x', t'\}$  is the subset of  $\mathcal{C}\{0, 0; t'\}$  when even the end-point is fixed to be  $x(t') = x'$ . The following *normalization* property holds:

$$\langle 1 \rangle_w = \int_{\mathcal{C}\{0, 0; t\}} 1 \cdot d_w x(\tau) = 1$$

We can then rewrite (12) as:

$$I_1 = \int_{\mathcal{C}\{0, 0; 0, t\}} d_W x(\tau)$$

Geometrically,  $W(0, t|0, 0) = I_1$  is the probability that a Brownian particle starting at the origin *returns* in it after a finite amount of time  $t$ .

The standard way to compute a Wiener integral is to *discretize* it, and then take a *continuum* limit. So, consider for simplicity a **uniform** time discretization  $\{t_i\}_{i=1, \dots, N+1}$ , with instants  $\epsilon$ -apart from each other, so that:

$$t_i - t_{i-1} \equiv \epsilon = \frac{t}{N+1} \quad \forall i = 1, \dots, N+1$$

Note that the end-points are  $x_0 = x_{N+1} = 0$ .

Then:

$$I_1 \equiv \lim_{\substack{\epsilon \rightarrow 0 \\ N \rightarrow \infty}} I_1^{(N)} \tag{13}$$

$$I_1^{(N)} \equiv \frac{1}{(\sqrt{4\pi D\epsilon})^{N+1}} \int_{-\infty}^{+\infty} dx_1 \int_{-\infty}^{+\infty} dx_2 \cdots \int_{-\infty}^{+\infty} dx_N \exp\left(-\frac{1}{4D\epsilon} \sum_{i=0}^N (x_{i+1} - x_i)^2\right) \tag{14}$$

Let's focus on the summation in the exponential:

$$\begin{aligned} \sum_{i=0}^N (x_{i+1} - x_i)^2 &= x_1^2 + \cancel{x_0^2} - 2x_0x_1 + x_2^2 + x_1^2 - 2x_1x_2 + \cdots + \cancel{x_{N+1}^2} + x_N^2 - 2x_Nx_{N+1} = \\ &= 2(x_1^2 + \cdots + x_N^2) - 2(x_1x_2 + x_2x_3 + \cdots + x_{N-1}x_N) = \\ &= 2\left(\sum_{i=1}^N x_i^2\right) - 2\left(\sum_{i=1}^{N-1} x_i x_{i+1}\right) \end{aligned}$$

This is a *quadratic form*, i.e. a polynomial with all terms of order 2. So, it can be written in *matrix form*:

$$= \sum_{k,l=1}^N x_k A_{kl} x_l = \mathbf{x}^T A_N \mathbf{x}$$



for an appropriate choice of entries  $A_{kl}$  of the  $N \times N$  matrix  $A_N$ :

$$A_{kk} = 2; A_{kl} = -(\delta_{kl} + 1 + \delta_{k+1,l}) \Rightarrow A_N = \begin{pmatrix} 2 & -1 & 0 & \cdots & 0 \\ -1 & 2 & -1 & \ddots & \vdots \\ 0 & \ddots & \ddots & \ddots & 0 \\ \vdots & \ddots & -1 & 2 & -1 \\ 0 & \cdots & 0 & -1 & 2 \end{pmatrix}$$

Substituting back in (14):

$$I_1^{(N)} = \frac{1}{(\sqrt{4\pi D\epsilon})^{N+1}} \int_{-\infty}^{+\infty} dx_1 \cdots \int_{-\infty}^{+\infty} dx_N \exp\left(-\frac{\mathbf{x}^T A_N \mathbf{x}}{4D\epsilon}\right)$$

Recall the multivariate Gaussian integral:

$$\int_{-\infty}^{+\infty} dx_1 \cdots dx_N \exp\left(-\sum_{ij} B_{ij} x_i x_j\right) = \frac{(\sqrt{\pi})^N}{\sqrt{\det B}}$$

leading to:

$$I_1^{(N)} = \frac{1}{(\sqrt{4\pi D\epsilon})^{N+1}} \sqrt{\frac{\pi^N}{\det(A_N [\frac{1}{4D\epsilon}]^N)}} \stackrel{(a)}{=} \frac{1}{(\sqrt{4\pi D\epsilon})^{N+1}} \frac{\sqrt{4\pi D\epsilon}^N}{\sqrt{\det A_N}} = \frac{1}{\sqrt{4\pi D\epsilon}} \frac{1}{\sqrt{\det A_N}} \quad (15)$$

where in (a) we used the property of the determinant  $\det(cA) = c^n \det(A) \forall c \in \mathbb{R}$ . Now, all that's left is to compute the determinant of  $A_N$ . Fortunately, as  $A_N$  is a *tri-diagonal* matrix, there is a recurrence relation in terms of the leading principal minors of  $A_N$ , which turns out to be multiples of the determinants of  $A_{N-1}$  and  $A_{N-2}$ .

Explicitly, consider  $A_N$ :

$$\det A_N \equiv \begin{vmatrix} 2 & -1 & 0 & \cdots & 0 \\ -1 & 2 & -1 & \ddots & \vdots \\ 0 & \ddots & \ddots & \ddots & 0 \\ \vdots & \ddots & -1 & 2 & -1 \\ 0 & \cdots & 0 & -1 & 2 \end{vmatrix}_{N \times N}$$

and start computing the determinant following the last column. The only non-zero

contributions are:

$$\begin{aligned}
\det A_N &= \underbrace{(-1)^{(N-1)+N}}_{+1} (-1) \begin{vmatrix} 2 & -1 & 0 & \cdots & 0 \\ -1 & 2 & -1 & \ddots & \vdots \\ 0 & \ddots & \ddots & \ddots & 0 \\ \vdots & \ddots & -1 & 2 & -1 \\ 0 & \cdots & 0 & 0 & \textcolor{violet}{-1} \end{vmatrix}_{(N-1) \times (N-1)} + (-1)^{2N} \textcolor{red}{(2)} \det A_{N-1} = \\
&= (-1)^{2(N-1)} \textcolor{violet}{(-1)} \det A_{N-2} + 2 \det A_{N-1} = 2 \det A_{N-1} - \det A_{N-2} \quad (16)
\end{aligned}$$

where the terms in blue are just the alternating signs from the determinant expansion, and the other colours identify the matrix entries that are being used. Then, it is just a matter of computing the first two terms of the succession ( $|A_N| = \det A_N$  for brevity):

$$|A_1| = 2 \quad |A_2| = \begin{vmatrix} 2 & -1 \\ -1 & 2 \end{vmatrix} = 4 - 1 = 3$$

And now we can use (16) to iteratively compute all  $|A_N|$ , e.g.  $|A_3| = 2 \cdot 3 - 2 = 4$ . To find  $|A_N|$  for a *generic*  $N$ , we need to make an hypothesis, and then verify that it is compatible with (16). In this case, note that  $|A_N| = N + 1$  (\*) for all the examples we explicitly computed. Then, by induction:

$$|A_{N+1}| \stackrel{(16)}{=} 2 \cdot |A_N| - |A_{N-1}| \stackrel{(*)}{=} 2 \cdot (N + 1) - (N - 1 + 1) = 2N + 2 - N = (N + 1) + 1$$

which is indeed compatible with (\*). So, substituting back in (15) we get:

$$I_1^{(N)} = \frac{1}{\sqrt{4\pi D\epsilon}} \frac{1}{\sqrt{N+1}} \stackrel{(a)}{=} \frac{1}{\sqrt{4\pi Dt}}$$

where in (a) we used  $\epsilon = t/(N + 1) \Rightarrow N + 1 = t/\epsilon$  from the discretization. Note that this result is *constant* with respect to  $\epsilon$  or  $N$  (recall that  $t$  is fixed beforehand) and so taking the *continuum* limit leads immediately to  $I_1$  (13):

$$I_1 \equiv \lim_{\substack{\epsilon \rightarrow 0 \\ N \rightarrow \infty}} \frac{1}{\sqrt{4\pi Dt}} = \frac{1}{\sqrt{4\pi Dt}}$$

which is coherent with the result we previously computed (11).