

# Formularius

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## 1 Binary SBM

### 1.1 Undirected/**Directed**

#### 1.1.1 MAR

In the directed case, only  $\mathcal{D}^\circ$  is changing in the following (i.e don't forget that the matrix is not symmetric:  $Y_{ij}$  and  $Y_{ji}$  are two different random variables).

**Proposition 1.** *The complete log-likelihood restricted to the observed variables is*

$$\log p_\theta(Y^\circ, Z) = \sum_{(i,j) \in \mathcal{D}^\circ} \sum_{(q,\ell) \in \mathcal{Q}^2} Z_{iq} Z_{j\ell} \log b(Y_{ij}, \pi_{q\ell}) + \sum_{i \in \mathcal{N}^\circ} \sum_{q \in \mathcal{Q}} Z_{iq} \log(\alpha_q),$$

with  $b(x, \pi) = \pi^x (1 - \pi)^{1-x}$  the Bernoulli probability density function.

**Proposition 2.** *The variationnal approximation is*

$$J_{\tau, \theta}(Y^\circ) = \sum_{(i,j) \in \mathcal{D}^\circ} \sum_{(q,\ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log b(Y_{ij}, \pi_{q\ell}) + \sum_{i \in \mathcal{N}^\circ} \sum_{q \in \mathcal{Q}} \tau_{iq} \log(\alpha_q / \tau_{iq}).$$

**Proposition 3.** *Consider the lower bound  $J_{\tau, \theta}(Y^\circ)$ .*

1. *The parameters  $\theta = (\alpha, \pi)$  maximizing  $J_\theta(Y^\circ)$  when  $\tau$  is held fixed are*

$$\hat{\alpha}_q = \frac{\sum_{i \in \mathcal{N}^\circ} \hat{\tau}_{iq}}{\text{card}(\mathcal{N}^\circ)}, \quad \hat{\pi}_{q\ell} = \frac{\sum_{(i,j) \in \mathcal{D}^\circ} \hat{\tau}_{iq} \hat{\tau}_{j\ell} Y_{ij}}{\sum_{(i,j) \in \mathcal{D}^\circ} \hat{\tau}_{iq} \hat{\tau}_{j\ell}}.$$

2. *The variational parameters  $\tau$  maximizing  $J_\tau(Y^\circ)$  when  $\theta$  is held fixed are obtained thanks to the following fixed point relation:*

$$\hat{\tau}_{iq} \propto \alpha_q \left( \prod_{(i,j) \in \mathcal{D}^\circ} \prod_{\ell \in \mathcal{Q}} b(Y_{ij}; \pi_{q\ell})^{\hat{\tau}_{j\ell}} \right).$$

$$\hat{\tau}_{iq} \propto \alpha_q \left( \prod_{(i,j) \in \mathcal{D}^\circ} \prod_{\ell \in \mathcal{Q}} (b(Y_{ij}; \pi_{q\ell}) b(Y_{ji}; \pi_{\ell q}))^{\hat{\tau}_{j\ell}} \right).$$

**Proposition 4.** For an SBM with  $Q$  blocks and for  $\hat{\theta} = \arg \max \log p_\theta(Y^\circ, Z)$ , the ICL criterion is given by

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_\tau} [\log p_{\hat{\theta}}(Y^\circ, Z; Q)] + \frac{Q(Q+1)}{2} \log \text{card}(\mathcal{D}^\circ) + (Q-1) \log \text{card}(\mathcal{N}^\circ).$$

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_\tau} [\log p_{\hat{\theta}}(Y^\circ, Z; Q)] + Q^2 \log \text{card}(\mathcal{D}^\circ) + (Q-1) \log \text{card}(\mathcal{N}^\circ).$$

#### Likelihood of Random-Pair sampling

$$p_\rho(R) = \prod_{1 \leq i < j \leq n} p_\rho(R_{ij}) = \left( \prod_{Y^\circ} \rho \right) \left( \prod_{Y^\mathfrak{m}} (1 - \rho) \right) = \rho^{\text{card}(Y^\circ)} (1 - \rho)^{\text{card}(Y^\mathfrak{m})}$$

$$p_\rho(R) = \prod_{1 \leq i \neq j \leq n} p_\rho(R_{ij}) = \left( \prod_{Y^\circ} \rho \right) \left( \prod_{Y^\mathfrak{m}} (1 - \rho) \right) = \rho^{\text{card}(Y^\circ)} (1 - \rho)^{\text{card}(Y^\mathfrak{m})}$$

**Likelihood of star sampling** The likelihood is the same in symmetric and non-symmetric context.

$$p_\rho(S) = \prod_{1 \leq i \leq n} p_\rho(S_i) = \left( \prod_{i, S_i=1} \rho \right) \left( \prod_{i, S_i=0} (1 - \rho) \right) = \rho^{\sum_i S_i} (1 - \rho)^{\sum_i (1-S_i)}$$

#### ICL in random-pair sampling case

**Proposition 5.** For an SBM with  $Q$  blocks and for  $\hat{\theta} = \arg \max \log p_\theta(Y^\circ, Z, R)$ , the ICL criterion is given by

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_\tau} [\log p_{\hat{\theta}}(Y^\circ, Z, R; Q)] + (1 + \frac{Q(Q+1)}{2}) \log \text{card}(\mathcal{D}^\circ) + (Q-1) \log \text{card}(\mathcal{N}^\circ).$$

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_\tau} [\log p_{\hat{\theta}}(Y^\circ, Z, R; Q)] + (1 + Q^2) \log \text{card}(\mathcal{D}^\circ) + (Q-1) \log \text{card}(\mathcal{N}^\circ).$$

#### ICL in star sampling case

**Proposition 6.** For an SBM with  $Q$  blocks and for  $\hat{\theta} = \arg \max \log p_\theta(Y^\circ, Z, R)$ , the ICL criterion is given by

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_\tau} [\log p_{\hat{\theta}}(Y^\circ, Z, R; Q)] + \frac{Q(Q+1)}{2} \log \text{card}(\mathcal{D}^\circ) + Q \log \text{card}(\mathcal{N}^\circ).$$

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_\tau} [\log p_{\hat{\theta}}(Y^\circ, Z, R; Q)] + Q^2 \log \text{card}(\mathcal{D}^\circ) + Q \log \text{card}(\mathcal{N}^\circ).$$

### 1.1.2 NMAR

Notice that  $\mathcal{D}^\circ$  and  $\mathcal{D}^m$  are changing in the following (i.e don't forget that the matrix is not symmetric:  $Y_{ij}$  and  $Y_{ji}$  are two different random variables).

**Proposition 7.** *The complete log-likelihood is*

$$\log p_{\theta, \psi}(Y^\circ, R, Y^m, Z) = \log p_\psi(R|Y^\circ, Y^m, Z) + \log p_\theta(Y^\circ, Y^m, Z),$$

where

$$\log p_\theta(Y, Z) = \sum_{1 \leq i < j \leq n} \sum_{(q, \ell) \in \mathcal{Q}^2} Z_{iq} Z_{j\ell} \log b(Y_{ij}, \pi_{q\ell}) + \sum_{i \in \{1, \dots, n\}} \sum_{q \in \mathcal{Q}} Z_{iq} \log(\alpha_q),$$

$$\log p_\theta(Y, Z) = \sum_{1 \leq i \neq j \leq n} \sum_{(q, \ell) \in \mathcal{Q}^2} Z_{iq} Z_{j\ell} \log b(Y_{ij}, \pi_{q\ell}) + \sum_{i \in \{1, \dots, n\}} \sum_{q \in \mathcal{Q}} Z_{iq} \log(\alpha_q),$$

with  $b(x, \pi) = \pi^x (1 - \pi)^{1-x}$  the Bernoulli probability density function.

**Proposition 8.** *The variationnal approximation is*

$$\begin{aligned} J_{\tau, \nu, \theta, \psi}(Y^\circ, R) &= \mathbb{E}_{\tilde{p}_{\tau, \nu}} [\log p_{\theta, \psi}(Y^\circ, R, Y^m, Z)] - \mathbb{E}_{\tilde{p}_{\tau, \nu}} [\log \tilde{p}_{\tau, \nu}(Z, Y^m)] \\ &= \mathbb{E}_{\tilde{p}_{\tau, \nu}} [\log p_\psi(R|Y^\circ, Y^m, Z)] \\ &\quad + \sum_{(i, j) \in \mathcal{D}^\circ} \sum_{(q, \ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log b(Y_{ij}, \pi_{q\ell}) + \sum_{(i, j) \in \mathcal{D}^m} \sum_{(q, \ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log b(\nu_{ij}, \pi_{q\ell}) \\ &\quad + \sum_{i \in \mathcal{N}} \sum_{q \in \mathcal{Q}} \tau_{iq} \log(\alpha_q / \tau_{iq}) - \sum_{(i, j) \in \mathcal{D}^m} \nu_{ij} \log(\nu_{ij}) + (1 - \nu_{ij}) \log(1 - \nu_{ij}). \end{aligned} \tag{1}$$

**Proposition 9.** *Consider the lower bound  $J_{\tau, \nu, \theta, \psi}(Y^\circ, R)$  given by (1).*

1. *The parameters  $\theta = (\alpha, \pi)$  maximizing (1) when all other parameters are held fixed are*

$$\hat{\alpha}_q = \frac{1}{n} \sum_{i \in \mathcal{N}} \hat{\tau}_{iq}, \quad \hat{\pi}_{q\ell} = \frac{\sum_{(i, j) \in \mathcal{D}^\circ} \hat{\tau}_{iq} \hat{\tau}_{j\ell} Y_{ij} + \sum_{(i, j) \in \mathcal{D}^m} \hat{\tau}_{iq} \hat{\tau}_{j\ell} \hat{\nu}_{ij}}{\sum_{(i, j) \in \mathcal{D}} \hat{\tau}_{iq} \hat{\tau}_{j\ell}}.$$

2. *The optimal  $\tau$  in (1) when all other parameters are held fixed verifies*

$$\begin{aligned} \hat{\tau}_{iq} &\propto \lambda_{iq} \alpha_q \left( \prod_{(i, j) \in \mathcal{D}^\circ} \prod_{\ell \in \mathcal{Q}} b(Y_{ij}; \pi_{q\ell})^{\hat{\tau}_{j\ell}} \right) \left( \prod_{(i, j) \in \mathcal{D}^m} \prod_{\ell \in \mathcal{Q}} b(\nu_{ij}; \pi_{q\ell})^{\hat{\tau}_{j\ell}} \right). \\ \hat{\tau}_{iq} &\propto \lambda_{iq} \alpha_q \left( \prod_{(i, j) \in \mathcal{D}^\circ} \prod_{\ell \in \mathcal{Q}} (b(Y_{ij}; \pi_{q\ell}) b(Y_{ji}; \pi_{\ell q}))^{\hat{\tau}_{j\ell}} \right) \left( \prod_{(i, j) \in \mathcal{D}^m} \prod_{\ell \in \mathcal{Q}} (b(\nu_{ij}; \pi_{q\ell}) b(\nu_{ji}; \pi_{\ell q}))^{\hat{\tau}_{j\ell}} \right). \end{aligned}$$

with  $\lambda_{iq}$  a simple constant depending on the sampling design.

**Double-standard sampling.** Recall that  $S^\circ = \sum_{(i, j) \in \mathcal{D}^\circ} Y_{ij}$ ,  $\bar{S}^\circ = \sum_{(i, j) \in \mathcal{D}^\circ} (1 - Y_{ij})$ , and denote  $S^m = \sum_{(i, j) \in \mathcal{D}^m} \nu_{ij}$ ,  $\bar{S}^m = \sum_{(i, j) \in \mathcal{D}^m} (1 - \nu_{ij})$ . Regarding the likelihood of the double standard sampling, we have

$$\mathbb{E}_{\tilde{p}} \log p_\psi(R|Y) = S^\circ \log \rho_1 + \bar{S}^\circ \log \rho_0 + S^m \log(1 - \rho_1) + \bar{S}^m \log(1 - \rho_0).$$

Based on this expression, we easily derive the following proposition:

**Proposition 10.** Consider the maximization of the lower bound (1) in the double standard sampling.

1. The parameters  $\psi = (\rho_0, \rho_1)$  maximizing (1) when all other parameters are held fixed are

$$\hat{\rho}_0 = \frac{\bar{S}^\circ}{\bar{S}^\circ + \bar{S}^\mathfrak{m}}, \quad \hat{\rho}_1 = \frac{S^\circ}{S^\circ + S^\mathfrak{m}}. \quad (1)$$

2. The optimal  $\nu$  in (1) when all other parameters are held fixed are

$$\hat{\nu}_{ij} = \text{logistic} \left( \log \left( \frac{1 - \rho_1}{1 - \rho_0} \right) + \sum_{(q, \ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log \left( \frac{\pi_{q\ell}}{1 - \pi_{q\ell}} \right) \right).$$

Moreover,  $\lambda_{iq} = 1 \ \forall (i, q) \in \mathcal{N} \times \mathcal{Q}$  for optimization of  $\tau$  in Proposition 9.2).

**Class sampling.** According to the likelihood of the class sampling, we derive the following expression of the conditional expectation under the variational approximation:

$$\mathbb{E}_{\tilde{p}} \log p_\psi(R|Y) = \sum_{i \in \mathcal{N}^\circ} \sum_{q \in \mathcal{Q}} \tau_{iq} \log(\rho_q) + \sum_{i \in \mathcal{N}^\mathfrak{m}} \sum_{q \in \mathcal{Q}} \tau_{iq} \log(1 - \rho_q),$$

from which we derive the maximization of the remaining parameters for class sampling.

**Proposition 11.** Consider the maximization of the lower bound (1) in the class sampling.

1. The parameters  $\psi = (\rho_1, \dots, \rho_Q)$  maximizing (1) when all other parameters are held fixed are

$$\hat{\rho}_q = \frac{\sum_{i \in \mathcal{N}^\circ} \tau_{iq}}{\sum_{i \in \mathcal{N}} \tau_{iq}}. \quad (2)$$

2. The optimal  $\nu$  in (1) when all other parameters are held fixed verify

$$\hat{\nu}_{ij} = \text{logistic} \left( \sum_{(q, \ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log \left( \frac{\pi_{q\ell}}{1 - \pi_{q\ell}} \right) \right).$$

Moreover  $\lambda_{iq} = \rho_q^{\mathbb{1}_{\{i \in \mathcal{N}^\circ\}}} (1 - \rho_q)^{\mathbb{1}_{\{i \in \mathcal{N}^\mathfrak{m}\}}}$  for optimization of  $\tau$  in Proposition 9.2).

**Star degree sampling.** From Expression of the likelihood for star degree sampling, one has

$$\mathbb{E}_{\tilde{p}} \log p_\psi(R|Y) = - \sum_{i \in \mathcal{N}^\mathfrak{m}} \left( a + b \tilde{D}_i \right) + \sum_{i \in \mathcal{N}} \mathbb{E}_{\tilde{p}} \left[ -\log(1 + e^{-(a+bD_i)}) \right],$$

where  $\tilde{D}_i = \mathbb{E}_{\tilde{p}} [D_i] = \sum_{j \in \mathcal{N}^\mathfrak{m}} \nu_{ij} + \sum_{j \in \mathcal{N}^\circ} Y_{ij}$  is the approximation of the degrees.

Because  $\mathbb{E}_{\tilde{p}} [-\log(1 + e^{-(a+bD_i)})]$  has no explicit form, we rely on an additional variational approximation. The principle is as follows: since  $g(x) = -\log(1 + e^{-x})$  is a convex function, we have from Taylor expansion

$$g(x) \geq g(\zeta) + \frac{x - \zeta}{2} + h(\zeta)(x^2 - \zeta^2), \ \forall (x, \zeta) \in \mathbb{R} \times \mathbb{R}^+,$$

where  $h(x) = \frac{-1}{2\zeta} [\text{logistic}(\zeta) - \frac{1}{2}]$ . This leads to a lower bound of the initial lower bound:

$$\log p_{\theta, \psi}(Y^\circ, R) \geq J_{\tau, \nu, \theta, \psi}(Y^\circ, R) \geq J_{\tau, \nu, \zeta, \theta, \psi}(Y^\circ, R), \quad (3)$$

with  $\zeta = (\zeta_i, i \in \mathcal{N})$  such that  $\zeta_i > 0$  is an additional set of variational parameters used to approximate  $-\log(1 + e^{-x})$ .

By gathering all this, we have

$$\begin{aligned} J_\zeta &= C_{miss} + \sum_{i=1}^n \mathbb{E}_{\tilde{p}} \left[ \text{logistic}(\zeta_i) + \frac{(a + bD_i) - \zeta_i}{2} + h(\zeta_i)((a + bD_i)^2 - \zeta_i^2) \right], \\ &= C_{miss} + \sum_{i=1}^n \left[ \text{logistic}(\zeta_i) + \frac{(a + b\tilde{D}_i) - \zeta_i}{2} + h(\zeta_i)(a^2 + 2ab\tilde{D}_i + b^2\hat{D}_i - \zeta_i^2) \right], \end{aligned}$$

where  $C_{miss} = \sum_{i \in \mathcal{N}^m} (a + b\tilde{D}_i)$  and  $\hat{D}_i = \mathbb{E}_{\tilde{p}}[D_i^2] = \mathbb{V}_{\tilde{p}}(D_i) + \mathbb{E}_{\tilde{p}}[D_i]^2 = \sum_{j \in \mathcal{N}^m} \nu_{ij}(1 - \nu_{ij}) + (\sum_{i \in \mathcal{N}^m} \nu_{ij} + \sum_{i \in \mathcal{N}^\circ} Y_{ij})^2$ .

**Proposition 12.** Consider the maximization of the lower bound (3) in the star degree sampling. Let us denote  $\hat{D}_i = \mathbb{E}_{\tilde{p}}[D_i^2]$  and  $\tilde{D}_k^{-\ell} = \tilde{D}_k - \nu_{k\ell}$ .

1. The parameters  $\psi = (a, b)$  maximizing  $J_{\tau, \nu, \zeta, \theta, \psi}(Y^\circ, R)$  when all other parameters are held fixed are

$$\begin{aligned} \hat{b} &= \frac{2(\frac{n}{2} - \text{card}(\mathcal{N}^m)) \sum_{i=1}^n (h(\zeta_i)\tilde{D}_i) - \left(\frac{1}{2} \sum_{i=1}^n \tilde{D}_i - \sum_{i \in \mathcal{N}^m} \tilde{D}_i\right) \times \sum_{i=1}^n h(\zeta_i)}{2 \sum_{i=1}^n (h(\zeta_i)\hat{D}_i) \times \sum_{i=1}^n h(\zeta_i) - \left(2 \sum_{i=1}^n h(\zeta_i)\tilde{D}_i\right)^2}, \\ \hat{a} &= -\frac{\hat{b} \sum_{i=1}^n (h(\zeta_i)\tilde{D}_i) + \frac{n}{2} - \text{card}(\mathcal{N}^m)}{\sum_{i=1}^n h(\zeta_i)}. \end{aligned} \quad (4)$$

2. The parameters  $\zeta$  maximizing  $J_{\tau, \nu, \zeta, \theta, \psi}(Y^\circ, R)$  when all other parameters are held fixed are

$$\hat{\zeta}_i = \sqrt{a^2 + b^2\hat{D}_i + 2ab\tilde{D}_i}, \quad \forall i \in \mathcal{N}.$$

3. The optimal  $\nu$  in  $J_{\tau, \nu, \zeta, \theta, \psi}(Y^\circ, R)$  when all other parameters are held fixed verify

$$\begin{aligned} \hat{\nu}_{ij} &= \text{logistic} \left( \sum_{(q, \ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log \left( \frac{\pi_{q\ell}}{1 - \pi_{q\ell}} \right) - b \right. \\ &\quad \left. + 2h(\zeta_i) \left( ab + b^2(1 + \tilde{D}_i^{-j}) \right) + 2h(\zeta_j) \left( ab + b^2(1 + \tilde{D}_j^{-i}) \right) \right). \end{aligned} \quad (5)$$

Moreover,  $\lambda_{iq} = 1 \quad \forall (i, q) \in \mathcal{N} \times \mathcal{Q}$  for optimization of  $\tau$  in Proposition 9.2).

**Proposition 13.** For a model with  $Q$  blocks, a sampling design with a vector of parameters  $\psi$  with dimension  $K$  and  $(\hat{\theta}, \hat{\psi}) = \arg \max_{(\theta, \psi)} \log p_{\theta, \psi}(Y^\circ, Y^m, R, Z)$ , the ICL criterion is

$$\text{ICL} = -2\mathbb{E}_{\hat{p}_{\tau, \nu; \hat{\theta}, \hat{\psi}}} \left[ \log p_{\hat{\theta}, \hat{\psi}}(Y^\circ, Y^m, R, Z | Q, K) \right] + \text{pen}_{\text{ICL}},$$

where

$$\text{pen}_{\text{ICL}} = \begin{cases} \left( K + \frac{Q(Q+1)}{2} \right) \log \left( \frac{n(n-1)}{2} \right) + (Q-1) \log(n) & \text{if the sampling design} \\ & \text{is dyad-centered,} \\ \frac{Q(Q+1)}{2} \log \left( \frac{n(n-1)}{2} \right) + (K+Q-1) \log(n) & \text{otherwise.} \end{cases}$$

$$\text{pen}_{\text{ICL}} = \begin{cases} (K+Q^2) \log(n(n-1)) + (Q-1) \log(n) & \text{if the sampling design} \\ & \text{is dyad-centered,} \\ Q^2 \log(n(n-1)) + (K+Q-1) \log(n) & \text{otherwise.} \end{cases}$$

## 2 Poisson SBM

### 2.1 Undirected/**Directed**

#### 2.1.1 MAR

In the directed case, only  $\mathcal{D}^\circ$  is changing in the following (i.e don't forget that the matrix is not symmetric:  $Y_{ij}$  and  $Y_{ji}$  are two different random variables).

**Proposition 14.** The complete log-likelihood restricted to the observed variables is

$$\log p_\theta(Y^\circ, Z) = \sum_{(i,j) \in \mathcal{D}^\circ} \sum_{(q,\ell) \in \mathcal{Q}^2} Z_{iq} Z_{j\ell} \log p(Y_{ij}, \lambda_{q\ell}) + \sum_{i \in \mathcal{N}^\circ} \sum_{q \in \mathcal{Q}} Z_{iq} \log(\alpha_q),$$

with  $p(x, \lambda) = \frac{\lambda^x}{x!} e^{-\lambda}$  the Poisson probability density function.

**Proposition 15.** The variationnnal approximation is

$$J_{\tau, \theta}(Y^\circ) = \sum_{(i,j) \in \mathcal{D}^\circ} \sum_{(q,\ell) \in \mathcal{Q}^2} \tau_{iq} \tau_{j\ell} \log p(Y_{ij}, \lambda_{q\ell}) + \sum_{i \in \mathcal{N}^\circ} \sum_{q \in \mathcal{Q}} \tau_{iq} \log(\alpha_q / \tau_{iq}).$$

**Proposition 16.** Consider the lower bound  $J_{\tau, \theta}(Y^\circ)$ .

1. The parameters  $\theta = (\alpha, \lambda)$  maximizing  $J_\theta(Y^\circ)$  when  $\tau$  is held fixed are

$$\hat{\alpha}_q = \frac{\sum_{i \in \mathcal{N}^\circ} \hat{\tau}_{iq}}{\text{card}(\mathcal{N}^\circ)}, \quad \hat{\lambda}_{q\ell} = \frac{\sum_{(i,j) \in \mathcal{D}^\circ} \hat{\tau}_{iq} \hat{\tau}_{j\ell} Y_{ij}}{\sum_{(i,j) \in \mathcal{D}^\circ} \hat{\tau}_{iq} \hat{\tau}_{j\ell}}.$$

2. The variational parameters  $\tau$  maximizing  $J_\tau(Y^\circ)$  when  $\theta$  is held fixed are obtained thanks to the following fixed point relation:

$$\hat{\tau}_{iq} \propto \alpha_q \left( \prod_{(i,j) \in \mathcal{D}^\circ} \prod_{\ell \in \mathcal{Q}} p(Y_{ij}; \lambda_{q\ell})^{\hat{\tau}_{j\ell}} \right).$$

$$\hat{\tau}_{iq} \propto \alpha_q \left( \prod_{(i,j) \in \mathcal{D}^\circ} \prod_{\ell \in \mathcal{Q}} (p(Y_{ij}; \lambda_{q\ell}) p(Y_{ji}; \lambda_{\ell q}))^{\hat{\tau}_{j\ell}} \right).$$

**Proposition 17.** *For an SBM with  $Q$  blocks and for  $\hat{\theta} = \arg \max \log p_{\theta}(Y^{\circ}, Z)$ , the ICL criterion is given by*

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_{\tau}} [\log p_{\hat{\theta}}(Y^{\circ}, Z; Q)] + \frac{Q(Q+1)}{2} \log \text{card}(\mathcal{D}^{\circ}) + (Q-1) \log \text{card}(\mathcal{N}^{\circ}).$$

$$\text{ICL}(Q) = -2\mathbb{E}_{\tilde{p}_{\tau}} [\log p_{\hat{\theta}}(Y^{\circ}, Z; Q)] + Q^2 \log \text{card}(\mathcal{D}^{\circ}) + (Q-1) \log \text{card}(\mathcal{N}^{\circ}).$$