HAMdetector: Combining information to detect HLA-associated mutations with a Bayesian regression model

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Abstract

Motivation
The human leucocyte antigen system (HLA) is of paramount importance to combat viral infections by presenting peptides on the cell surface via MHC I. Thus, CD8+ cytotoxic T-Lymphocytes exert a strong selection pressure towards virus variants that escape that immune recognition pathway, e.g. through point mutations that decreases binding of the respective peptide to MHC I.

Reliably identifying HLA-associated mutations is important for understanding viral evolution, but experimental methods like binding assays are prohibitively expensive for large-scale use and fail to recognize other mechanisms of immune escape like proteasomal processing.
One step in finding these mutations is through the statistical analysis of sequence data. However, existing methods are based on nullhypothesis significance testing and do not make use of all the available information and therefore have unsatisfactory real-world performance.

Results
Here, we present a Bayesian regression model that is easily extensible to include information from different sources (e.g. epitope prediction software) and makes use of recent advances in Bayesian inference, e.g. by using a sparsifying prior. We show that including this kind of information improves predictive performance considerably over state-of-the-art methods.

Availability and Implementation
The source code of this software is available at `http://github.com` under a permissive MIT license.

Supplementary information
Supplementary data are available at *Bioinformatics* online.

Keywords
human leucocyte antigen system, HLA, multiple sequence alignment, escape mutations, viral escape, Bayesian inference, sparsity, horseshoe, epitope prediction

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1. Introduction

1.1 The HLA system

One way how the human immune system is able to recognize intracellular viral infections is through the human leucocyte antigen system[1]: In cells with active protein biosynthesis, proteins are continuously synthesized and also degraded by a process called proteasomal degradation, which cleaves proteins into linear peptides of varying length[2]. A small subset of these peptides is presented on the cell surface via receptors called MHC class I (HLA-A, HLA-B and HLA-C in humans). The genomic region encoding for MHC I is known to be highly polymorphic, with more than 20000 different HLA alleles described today[3]. The resulting gene products differ in their binding properties, which means that cells from different individuals present a highly diverse set of peptides on their surface. Cytotoxic T cells are selected during maturation to only weakly bind to peptide/MHC I complexes when the peptide originated from proteins of the usual proteome, but might be able to strongly bind to complexes of MHC I with peptides which are generated from of a viral protein[4]. Upon activation, T cells induce cytolytic activity and recruit other immune cells [5].

1.2 HLA escape

In this way, the HLA system exerts strong selection pressure towards virus variants that escape T cell recognition[6], for example through a point mutation that results in reduced binding of an immunogenic peptide to MHC I or through a set of mutations that alters the viral protein in such a way that it is cleaved into different peptides that are not recognized by the host's T cell repertoire [7].

The evolutionary events are complex and occur not only on the level of individuals, where a virus adapts to specific features of the host, but also on the population level, because HLA alleles differ in their frequency across geographic regions[8].

Upon transmission to a new host, HLA escape mutations can revert to their wild type, as HLA escape mutations are associated with a reduction in viral replicative capacity [9]. Kawashima et al.[8] describe an escape mutation that is selected by HLA allele HLA-B*51, does not strongly affect viral replicative capacity, and therefore slowly enriches over time in Japan, where HLA-B*51 commonly occurs. How quickly a given escape mutation is selected upon transmission in a host depends on the magnitude of the reduction in viral replicative capacity, on the strength of selection pressure, and also on the genetic background, e.g. some escape mutations require compensatory mutations which partly attenuate the negative impact on viral replicative capacity.

Studying HLA escape provides an unique opportunity to gain insight into viral evolution, on the host level, but also on the population level. Unfortunately, identifying HLA escape mutations is difficult in practice.

1.3 Identifying HLA-escape mutations

There are several experimental methods available to study HLA escape: Recombinant MHC-I molecules can be used in binding assays. Upon complex formation with a peptide, a change in conformation can be detected with conformation-specific antibodies. This method is relatively fast, but only measures binding affinity of a peptide to MHC-I and does not account for antigen processing or immunodominance, which describes the observation that a peptide may be presented via MHC-I on the cell surface, but does not induce an immune response. An experimental setup that resembles the conditions in-vitro more closely but is also more time-consuming is to measure CD8+ T cell responses instead. This is usually done by stimulating peripheral blood mononuclear cells with prototype and variant peptides and measuring the secretion of IFN-γ by intracellular cytokine staining and fluorescence-assisted cell sorting. To analyze CD8+ T cell responses against endogenously processed antigens, it is necessary to generate cell-lines stably expressing the antigen in question and adding antigen-specific CD8+ T cells. This method scales poorly as it requires transfection of cell lines and antigen-specific expansion of CD8+ T cells.

1.4 Computational methods

Because the currently available experimental methods do not scale well enough to analyze whole viral genomes, a useful strategy might be to use annotated sequence data to identfiy candidate HLA escape mutations that can then be verified experimentally.

As the selection pressure exerted by cytotoxic T cells depends on successful recognition of viral peptides on the cell surface, escape mutations are often HLA-allele specific and can therefore be detected as HLA-allele dependent footprints in sequence alignments of viral proteins[10]: At certain alignment positions, a replacement

might be more frequently observed in sequences from hosts with a specific HLA allele than in sequences from hosts without that HLA allele. By quantifying this difference for all replacement and HLA allele pairs, it is possible to idenfify replacements that are enriched in sequences coming from hosts with certain HLA alleles, and thus are likely to be HLA escape mutations.

One way of quantifying this enrichment is Fisher's exact test[11]. For a given replacement R_i at alignment position i and HLA allele H, a 2-by-2 contingency table is constructed containing the absolute counts of the number of sequences in the four possible categories: (R_i, H), $(R_i, !H)$, $(!R_i, H)$ and $(!R_i, H)$, where $!R_i$ denotes any replacement except R_i, and $!H$ denotes any HLA allele except H. Fisher's exact test is a conventional nullhypothesis significance test (NHST) that generates p-values. In this case, the nullhypothesis is that HLA allele H and replacement R_i are independent, and the p-value is the probability of observing a deviation from independence that is at least as extreme as in the data at hand under the assumption that the nullhypothesis is true.

Fisher's exact test has the advantage of being easy to apply[12], but also has several disadvantages that are outlined in Carlson et al. [13]. The most striking one is that viral sequences share a common phylogenetic history, and, therefore, treating sequences as independent and identically distributed samples may under- or overestimate effect sizes. In the context of hypothesis testing, this leads to increased false positive and false negative rates. Another issue of applying Fisher's exact test is that HLA class I loci are located in close proximity on chromosome 6 and are therefore in linkage disquilibrium, which means that HLA alleles are not inherited completely independent of each other, i.e. inheritance of one HLA allele correlates with inheritance of another HLA allele. When using a statistical method that tests each HLA allele individually without considering the whole set of alleles present, spurious associations might occur: If HLA allele H_1 is associated with an amino acid replacement R, but H_1 is in linkage disequilibrium with another allele H_2, this also means that we observe an association between replacement R and H_2, even in absence of any underlying escape mechanism.

Correlations can not only occur between HLA alleles, but also between replacements. This kind of codon covariation occurs for example in compensatory mutations that attenuate the negative impact of immune escape mutations. For instance, a compensatory mutation might lead to a conformational change in such a way that the mutated protein resembles the original wild type more closely.

Carlson et al. [13] developed a method called Phylogenetic Dependency Network that accounts for phylogenetic bias, HLA linkage disequilibrium, and codon covariation and is based on nullhypothesis significance testing.

1.5 Issues with using p-values as a screening tool

In addition to the aforementioned biological reasons why Fisher's exact test is not suited for the analysis of sequence data, there are also more fundamental statistical issues that universally occur when using p-values as a screening tool[14].

In the presence of small effect sizes and high variance between measurements, as it is typically the case when working with biological data, statistically significant results can often be misleading and are likely to be in the wrong direction (a so-called type S error) or greatly overestimate an effect (a so-called type M error) [15]. This problem has recently been widely appreciated in the literature in the context of the current „replication crisis", which describes the circumstance that scientific claims with seemingly strong statistical evidence fail to replicate[16].

Another issue that occurs when using p-values as a screening tool is the problem of multiple comparisons. When applying a statistical test, the probability of obtaining a statistically significant result increases with each additional test, even in absence of any real effect. When using p-values as a filter, it is therefore likely to obtain significant effects that are in fact not real. To circumvent this problem, a common strategy is to control the false discovery rate, which is the expected proportion of false positives[17]. These adjustment procedures have the problem that, when performing many of such comparisons, none but the very largest effects remain. Instead of performing many hypothesis tests and trying to adjust for them, we instead prefer to fit a single, multilevel model that contains all comparisons of interest. When using multilevel models, the problem of multiple comparisons can disappear entirely and also yield more valid estimates[18].

2. Material and methods

When possible, we choose to fit Bayesian models. By using prior information, adding problem-specific structure and partial pooling, the accuracy of estimates can often be noticably improved[19]. Prior information does not necessarily mean to use external data, even a rough idea about the expected magnitude of estimates is often surprisingly effective. Additionally, Bayesian statistics provides an accessible way to test models: By comparing data generated under the model's assumptions to the actually observed data, it is possible to identify important aspects of the dataset that the models fails to capture and subsequently improve the model until it is consistent with the observed data [20].

In the context of identifying HLA-associated mutations we propose to improve existing methods by the following additions, which can be broadly divided into additional information and model structure:

Additional information

Binding affinity prediction

HLA-associated mutations are expected to lie more frequently in regions of known epitopes. There are vast experimental binding affinity data available of different peptides and MHC I molecule pairs, and there are well-established computational methods that use these data to extrapolate HLA binding for untested peptides. We show that, by including the outcome of these computational tools as input for a probabilistic model, the prediction of HLA-associated mutations can be improved.

Antigen processing prediction

Similarly, there are also HLA-allele independent effects like antigen processing that influence presentation on MHC I. Second generational tools like MHCFlurry 2.0 use both binding affinity and antigen processing data to improve epitope prediction. For our tool HAMdetector, we use the output of MHCFlurry to benefit from this binding affinity and antigen processing data in order to predict HLA-associated mutations.

Model structure

Sparsity-inducing priors

Recent advantages in Bayesian inference include so-called sparsity-promoting priors. Sparsity-promoting priors convey the apriori expectation that most coefficients in a regression model are close to 0. This assumption also applies to HLA-associated mutations, because the number of epitopes that are restricted by a given HLA allele is typically small compared to the number of all possible epitopes. If no epitope spanning a given alignment position is presented on the MHC I molecule, any association between a replacement and the respective HLA allele is likely to be due to random variation alone (or HLA linkage disequilibrium).

It has been shown that using a sparsity-promoting prior when non-zero coefficients are sparse can drastically improve predictive performance, because the model is better able to differentiate between signal and noise.

We show that a simple logistic regression model with a sparsity-promoting prior on the regression coefficients for the HLA alleles alone already performs roughly on-par with the much more elaborate Phylogenetic Dependency Network approach from Carlson et al. when using the fraction of identified escapes that lie in known HLA epitopes as a benchmark.

Partial pooling of 4-digit HLA alleles

It has recently been appreciated that binding specifities can vary drastically across HLA alleles from the same allele group, e.g. between HLA-B*51:01 and HLA-B*51:03. For predicting HLA associated mutations, there has consequently been a shift to use 4-digit resolution data whenever possible. This is not without downsides however, because overall, binding specifities are more similar for alleles in the same group, and, therefore, treating them as completely separate might unnecessarily fragment the available data.

In Bayesian statistics, it is not required to chose one extreme or the other: Instead of either treating all HLA alleles from the same allele group as identical or as completely separate, partial pooling of the estimates allows something inbetween: By partial pooling, estimates do share some information across each other, but they are also allowed to vary if necessary.

In the context of our model this means that observing a relevant association between a replacement and an HLA allele also makes the model more inclined to estimate relevant associations between that replacement and all other alleles from that allele group. The degree of this partial pooling can be estimated from the data, so if we do observe strong similarity of HLA alleles in a given

allele group the estimates are more influenced by each other than if they do not.

3. Implementation

Model backbone

- logistic regression model, because easily extensible
- coefficients are interpretable as expected increase in log-odds of observing a given replacement (also: average predictive comparisons to transform estimates to probability scale)
- probabilistic model allows us to perform posterior predictive checks
- disadvantages: none? maybe log-odds scale unintuitive. mention divide-by-4 rule
- establish notation, start simple and expand on that? Maybe confusing because some things change as models get more complex

Including phylogeny

- sequence data shares common phylogenetic history and are therefore not i.i.d, cite Brumme paper
- standard approach: multivariate normal intercept where covariance matrix is based on shared branch lengths
- this approach is computationally too expensive because of current limitations in Stan. Would be possible by fitting the model in Julia? Probably omit that.
- Our approach is close to the one in Carlson et al.
- slight tweak: use tree estimation tools for tree likelihoods p(tree | sequences), then bayes rule to invert that conditional probability. Optimize branch lengths for each replacement, but keep tree topology fixed
- advantages: faster (raxml well optimized)
- disadvantages: ad-hoc, close to conventional multivariate normal approach with off-diagonal elements set to 0, therefore not as flexible. But: should be good enough, phylogeny seems to be not that important (probably omit that?)
- how to include p(replacement | tree) in the model: just add additional intercept
- add coefficient for that information, so model can decide how to make use of that information.
- logit(p(replacement|tree)) used because it cancels out with the logistic link function. therefore: phylogeny information as baseline in absence of any HLA effects
- calibration plots to show that estimated phylogenetic probabilities are reasonably well calibrated.
- meaning: for sequences with an expected probability of observing the replacement of x%, we really do see that x% of the sequences have that replacement.

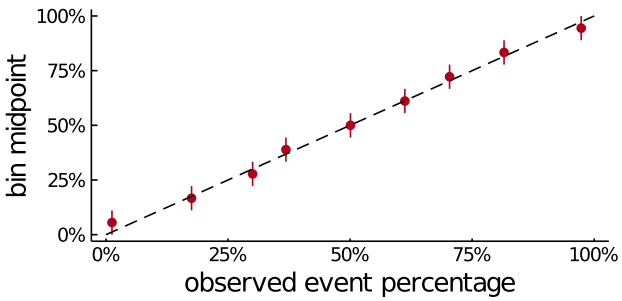

Figure 1. Calibration plot and how it works.

Including HLA linkage disequilibrium and codon covariation

- HLA linkage disequilibrium is accounted for by considering all HLA alleles of a sample at once.
- If strong correlations between HLA alleles: marginal posterior for both of them overlap with 0
- Codon covariation is not accounted for because we are mainly interested in HAMs.
- Model could be extended to include codon covariation by including them as additional predictors. This is possible because we use the horseshoe prior.

Including sparsity assumptions

- one of the recent advances in Bayesian inference: development of spacial priors to give model useful properties

- we expect most HLA coefficients to be 0, because no epitope spanning that region is presented on the cell surface, therefore no selection pressure and no association with possible replacements
- this is prior knowledge that should be incorporated in the model
- leads to better inferences because uncertainty of the regression coefficients does not propagate into observations.
- works by placing most probability mass very close to 0, but with large tails for non-zero coefficients.
- show formula
- global shrinkage parameter is very small and shrinks most coefficients to 0, local shrinkage parameter is very large and allows some coefficients to escape that shrinkage
- prior can be specified in terms of expected number of non-zero coefficients
- in our case, degree of sparsity can be well estimated from the data because all replacements share the same global shrinkage parameter.
- useful addition for logistic regression models: regularized horseshoe to have some regularization for non-zero coefficients. this helps to deal with issues of separability and collinearity.
- effect of horseshoe prior compared to standard logistic regression model: refer to model testing section. In our benchmark, the horseshoe prior alone is more effective than the model with phylogeny.
- maybe include comparison of marginal posteriors with and without horseshoe prior

Including epitope prediction software

- Vast experimental data available.
- HAMs are much more likely to lie in known epitopes
- epitope prediction software works reasonably well and can be included into as input to a probabilitic model
- MHCFlurry is used for epitope prediction and Antigen processing prediction: Input matrix (dimensions: #replacements x #HLA alleles) contains a 1 if that position is either expected to be inside a predicted epitope or related to antigen processing, 0 otherwise
- uses some thresholds, but this is not so bad because we just use it as input for a probabilistic model. Maybe play around with different options? The Brumme paper uses some offsets for defining the region of known epitopes, maybe this does make a difference.
- including this kind of information can drastically improve predictive performance, as it is relevant external data.
- Bayesian inference allows us to include informationen from different sources in a consistent manner.
- this information may be imprecise or wrong, but the model is able to learn how much to trust this data if parameterized correctly.
- epitope prediction / antigen processing escape is information about the expected degree of sparsity, i.e. some coefficients are more likely to be non-zero than others.
- can be included into the model by varying the local shrinkage parameter term.
- higher standard deviation of the local shrinkage parameters mean that it is more likely that this coefficient is non-zero.
- show formula

Partial pooling of 4-digit HLA alleles

- (this still has to be implemented, but should be a nice feature. If this is going to be too time consuming I can just leave it out)
- realized through allele group specific intercepts, e.g. HLA-B51:01 and HLA-B51:03, ..., share a common intercept.
- show formula

Full model specification

- show all formulas
- all models were implementen in Stan. Code is available online in two versions: one optimized

for readability, one optimized for speed with multithreading and GPU support

- data can be analyzed using a custom Julia package (tested on linux).

Prior justification

- justify priors, mostly based on the expected scales.

4. Results and Discussion

References

[1] Ronald N. Germain. MHC-dependent antigen processing and peptide presentation: Providing ligands for t lymphocyte activation. *Cell*, 76(2):287–299, jan 1994.

[2] Alfred L Goldberg, Paolo Cascio, Tomo Saric, and Kenneth L Rock. The importance of the proteasome and subsequent proteolytic steps in the generation of antigenic peptides. *Molecular Immunology*, 39(3-4):147–164, oct 2002.

[3] James Robinson, Jason A. Halliwell, James D. Hayhurst, Paul Flicek, Peter Parham, and Steven G. E. Marsh. The IPD and IMGT/HLA database: allele variant databases. *Nucleic Acids Research*, 43(D1):D423–D431, nov 2014.

[4] S. Murata, K. Sasaki, T. Kishimoto, S. i. Niwa, H. Hayashi, Y. Takahama, and K. Tanaka. Regulation of c8+ t cell development by thymus-specific proteasomes. *Science*, 316(5829):1349–1353, jun 2007.

[5] John T. Harty, Amy R. Tvinnereim, and Douglas W. White. CD8+ t cell effector mechanisms in resistance to infection. *Annual Review of Immunology*, 18(1):275–308, apr 2000.

[6] Persephone Borrow, Hanna Lewicki, Xiping Wei, Marc S. Horwitz, Nancy Peffer, Heather Meyers, Jay A. Nelson, Jean Edouard Gairin, Beatrice H. Hahn, Michael B.A. Oldstone, and George M. Shaw. Antiviral pressure exerted by HIV-l-specific cytotoxic t lymphocytes (CTLs) during primary infection demonstrated by rapid selection of CTL escape virus. *Nature Medicine*, 3(2):205–211, feb 1997.

[7] Jonathan W. Yewdell and Ann B. Hill. Viral interference with antigen presentation. *Nature Immunology*, 3(11):1019–1025, nov 2002.

[8] Yuka Kawashima, Katja Pfafferott, John Frater, Philippa Matthews, Rebecca Payne, Marylyn Addo, Hiroyuki Gatanaga, Mamoru Fujiwara, Atsuko Hachiya, Hirokazu Koizumi, Nozomi Kuse, Shinichi Oka, Anna Duda, Andrew Prendergast, Hayley Crawford, Alasdair Leslie, Zabrina Brumme, Chanson Brumme, Todd Allen, Christian Brander, Richard Kaslow, James Tang, Eric Hunter, Susan Allen, Joseph Mulenga, Songee Branch, Tim Roach, Mina John, Simon Mallal, Anthony Ogwu, Roger Shapiro, Julia G. Prado, Sarah Fidler, Jonathan Weber, Oliver G. Pybus, Paul Klenerman, Thumbi Ndung'u, Rodney Phillips, David Heckerman, P. Richard Harrigan, Bruce D. Walker, Masafumi Takiguchi, and Philip Goulder. Adaptation of HIV-1 to human leukocyte antigen class i. *Nature*, 458(7238):641–645, feb 2009.

[9] Philippa C. Matthews, Andrew Prendergast, Alasdair Leslie, Hayley Crawford, Rebecca Payne, Christine Rousseau, Morgane Rolland, Isobella Honeyborne, Jonathan Carlson, Carl Kadie, Christian Brander, Karen Bishop, Nonkululeko Mlotshwa, James I. Mullins, Hoosen Coovadia, Thumbi Ndung'u, Bruce D. Walker, David Heckerman, and Philip J. R. Goulder. Central role of reverting mutations in HLA associations with human immunodeficiency virus set point. *Journal of Virology*, 82(17):8548–8559, jul 2008.

[10] C. B. Moore. Evidence of HIV-1 adaptation to HLA-restricted immune responses at a population level. *Science*, 296(5572):1439–1443, may 2002.

[11] R. A. Fisher. On the interpretation of χ^2 from contingency tables, and the calculation of p. *Journal of the Royal Statistical Society*, 85(1):87, jan 1922.

[12] Bettina Budeus, Jörg Timm, and Daniel Hoffmann. SeqFeatR for the discovery of feature-sequence associations. *PLOS ONE*, 11(1):e0146409, jan 2016.

[13] Jonathan M. Carlson, Zabrina L. Brumme, Christine M. Rousseau, Chanson J. Brumme, Philippa Matthews, Carl Kadie, James I. Mullins, Bruce D. Walker, P. Richard Harrigan, Philip J. R. Goulder, and David Heckerman. Phylogenetic dependency networks: Inferring patterns of CTL escape and

codon covariation in HIV-1 gag. *PLoS Computational Biology*, 4(11):e1000225, nov 2008.

[14] Valentin Amrhein and Sander Greenland. Remove, rather than redefine, statistical significance. *Nature Human Behaviour*, 2(1):4–4, sep 2017.

[15] Andrew Gelman and John Carlin. Beyond power calculations. *Perspectives on Psychological Science*, 9(6):641–651, nov 2014.

[16] John P. A. Ioannidis. Why most published research findings are false. *PLoS Medicine*, 2(8):e124, aug 2005.

[17] Yoav Benjamini and Yosef Hochberg. Controlling the false discovery rate: A practical and powerful approach to multiple testing. *Journal of the Royal Statistical Society: Series B (Methodological)*, 57(1):289–300, jan 1995.

[18] Andrew Gelman, Jennifer Hill, and Masanao Yajima. Why we (usually) don't have to worry about multiple comparisons. *Journal of Research on Educational Effectiveness*, 5(2):189–211, 2012.

[19] Andrew Gelman. Bayesian statistics then and now. *Statistical Science*, 25(2):162–165, may 2010.

[20] Jonah Gabry, Daniel Simpson, Aki Vehtari, Michael Betancourt, and Andrew Gelman. Visualization in bayesian workflow. *Journal of the Royal Statistical Society: Series A (Statistics in Society)*, 182(2):389–402, jan 2019.