

# STAT3602 Statistical Inference

## Example Class 5

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- 1 Midterm Exam Review: Frequentist and Bayesian Approach
- 2 Exercise 1
- 3 Exercise 2
- 4 Exercise 3

# Chapter 1: Frequentist Approach I

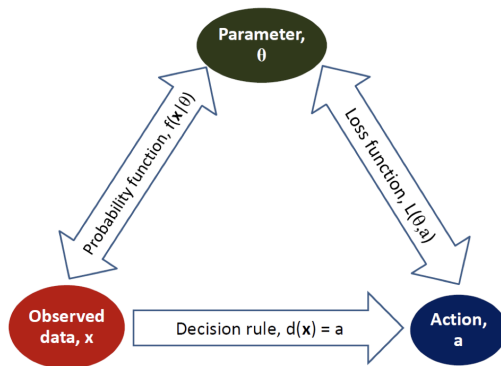


Figure: Frequentist Approach: Formulation

## Frequentist Approach: Formulation

- Formulation and Risk Function
- Criteria for a good decision rule
  - Admissibility
  - Minimavity
  - Unbiased rule
- Randomised decision rules

# Chapter 2: Bayesian Approach I

## Bayesian Decision

- Prior and Posterior Probability Function.
- Expected Posterior Loss
- Bayesian Decision

## Bayesian Statistical Inference

- Point Estimation
- Hypothesis Testing about  $\theta$
- Interval Estimation of  $\theta$ :
  - Given fixed interval length, maximise posterior coverage probability.
  - Given fixed posterior coverage probability, minimise interval length.
  - Given fixed posterior coverage probability, require “equal-tailed” interval.
- Predictive Distribution

# Outline

- 1 Midterm Exam Review: Frequentist and Bayesian Approach
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## Exercise 1 I

Papet, an old greedy man living in village A, wants to purchase a plot of land in his village. The land is currently owned by Jean, a hunchback recently arriving from Village B, who inherits the land from his deceased mother Florette, Papet's ex-lover.

The price of the land depends on the availability of water. Papet knows that there is a water source 1 mile from the land, and is considering to block it.

If Papet does not block the source (action  $a_0$ ), then Jean will have access to water from this source and will sell his land at a price of 1 unit. If Papet blocks the source (action  $a_1$ ), then Jean will need to search for an alternative water source, at an unknown distance of  $\theta$  miles away, and will sell his land at a price of  $e^{1-\theta}$  units. The following questions concern

Papet's decision problem of whether to block the water source.

## Exercise 1 II

- 1 Papet's loss function  $L(\theta, a)$  can be defined as the price of the land. Write down the expressions for  $L(\theta, a_0)$  and  $L(\theta, a_1)$ , for any  $\theta > 0$
- 2 Papet tries to make his decision based on an observation  $X$ , which is the growth rate of a plant found near the land plot. It is known that  $X$  is distributed under the density function

$$f(x | \theta) = \theta e^{-\theta x}, \quad x > 0.$$

Show that the risk function of Papet's decision rule  $d$  is given by

$$R(\theta, d) = \mathbb{P}_\theta(d(X) = a_1) \left( e^{1-\theta} - 1 \right) + 1, \quad \theta > 0$$

- 3 Find a minimax decision rule for Papet.



## Exercise 1 III

- ④ Suppose that Papet has a prior belief about  $\theta$ , represented by the prior density function

$$\pi(\theta) = ce^{-c\theta}, \quad \theta > 0,$$

for some fixed constant  $c > 0$ . Show that Papet's Bayes rule is to block the water source if and only if

$$c + X < \frac{1}{\sqrt{e} - 1}$$

**[ You may find the following identity useful:  $\int_0^\infty \theta e^{-k\theta} d\theta = k^{-2}$  for any  $k > 0$ . ]**

# Exercise 1: Solution I

- ① Write down the expressions for  $L(\theta, a_0)$  and  $L(\theta, a_1)$ , for any  $\theta > 0$ .

$$L(\theta, a_0) = 1, \quad L(\theta, a_1) = e^{1-\theta}$$

- ② The risk function of Papet's decision rule  $d$  is

$$\begin{aligned} R(\theta, d) &= \mathbb{E}_\theta[L(\theta, d(X))] \\ &= L(\theta, a_0) \mathbb{P}_\theta(d(X) = a_0) + L(\theta, a_1) \mathbb{P}_\theta(d(X) = a_1) \\ &= 1 - \mathbb{P}_\theta(d(X) = a_1) + e^{1-\theta} \mathbb{P}_\theta(d(X) = a_1) \\ &= \mathbb{P}_\theta(d(X) = a_1) (e^{1-\theta} - 1) + 1 \end{aligned}$$

# Exercise 1: Solution II

## 3 Find a minimax decision rule for Papet.

Denote by  $d_0$  the rule of always choosing action  $a_0$ . For any rule  $d$ ,

$$\begin{aligned}\sup_{\theta > 0} R(\theta, d) &\geq R(\theta = 0, d) \\ &= \mathbb{P}_{\theta}(d(X) = a_1) (e^{1-0} - 1) + 1 \\ &\geq 1 \\ &= \sup_{\theta > 0} R(\theta, d_0)\end{aligned}$$

so that  $d_0$  is minimax.

## Exercise 1: Solution III

- ④ Suppose that Papet has a prior belief about  $\theta$ , represented by the prior density function  $\pi(\theta) = ce^{-c\theta}$ ,  $\theta > 0$  for some fixed constant  $c > 0$ . Show that Papet's Bayes rule is to block the water source if and only if  $c + X < \frac{1}{\sqrt{e}-1}$

Expected posterior loss:

$$\begin{aligned}\mathbb{E}[L(\theta, a) \mid X] &\propto \int_0^\infty L(\theta, a) (\theta e^{-\theta X}) (ce^{-c\theta}) d\theta \\ &= \begin{cases} c \int_0^\infty \theta e^{-\theta(c+X)} d\theta = c(c+X)^{-2}, & a = a_0 \\ ce \int_0^\infty \theta e^{-\theta(1+c+X)} d\theta = ce(1+c+X)^{-2}, & a = a_1 \end{cases}\end{aligned}$$

Bayes rule is to choose  $a_1$  iff

$$c(c+X)^{-2} > ce(1+c+X)^{-2} \Leftrightarrow c+X < \frac{1}{\sqrt{e}-1}$$

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## Exercise 2

Each customer buys one of the  $K$  products in a shop, buying product  $i$  with probability  $\theta_i$ . Suppose there are  $n$  customers.  $X_i$  of them buy product  $i$ .  $X = (X_1, \dots, X_K)$  follows multinomial distribution with probability mass function

$$f(x | \theta) = \frac{n!}{x_1! \dots x_K!} \theta_1^{x_1} \dots \theta_K^{x_K}.$$

(Mean for  $X_i$  is  $n\theta_i$ ).

Our prior distribution for  $\theta = (\theta_1, \dots, \theta_K)$  is  $\theta \sim \text{Dirichlet}(\alpha_1, \dots, \alpha_K)$

$$\pi(\theta) = \frac{\Gamma(\alpha_1 + \dots + \alpha_K)}{\Gamma(\alpha_1) \dots \Gamma(\alpha_K)} \theta_1^{\alpha_1-1} \dots \theta_K^{\alpha_K-1}.$$

- 1 Find the posterior distribution of  $\theta | x$
- 2 Suppose there are  $m$  customers independent of the first  $n$  customers.  $Y_i$  of the  $m$  customers buy product  $i$ . Find the mean of the posterior predictive distribution of  $Y_i | x, i = 1, \dots, K$

## Exercise 2: Solution I

- 1 Please refer to the page 28 – 29 in the note of Chapter 2

$$\begin{aligned}\pi(\theta \mid \mathbf{x}) &\propto \pi(\theta) f(\mathbf{x} \mid \theta) \\ &\propto \theta_1^{\alpha_1-1} \dots \theta_K^{\alpha_K-1} \theta_1^{x_1} \dots \theta_K^{x_K} \\ &= \theta_1^{x_1+\alpha_1-1} \dots \theta_K^{x_K+\alpha_K-1}\end{aligned}$$

which follows  $\text{Dir}(x_1 + \alpha_1, \dots, x_K + \alpha_K)$

2

$$\begin{aligned}E(Y_i \mid \mathbf{x}) &= E(E(Y_i \mid \theta) \mid \mathbf{x}) \\ &= E(m\theta_i \mid \mathbf{x}) = \frac{m(x_i + \alpha_i)}{\sum_{j=1}^K (x_j + \alpha_j)}\end{aligned}$$

[It can be derived that the mean of  $\theta_i$  is  $\frac{\alpha_i}{\sum_{j=1}^K \alpha_j}$  for  
 $\theta \sim \text{Dirichlet}(\alpha_1, \dots, \alpha_K)$  ]

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## Exercise 3 I

Let  $\theta$  be an unknown parameter in  $[0, 1]$ , representing the pollution index of a city. In a sample of air collected from the city, the content of a particular pollutant is observed to be  $100X\%$ , where  $X \in [0, 1]$  is a random variable with the density function

$$f(x | \theta) = 1 + 2\theta \left( x - \frac{1}{2} \right), \quad x \in [0, 1]$$

Before the measurement of  $X$ , non-informative prior distribution is given to  $\theta$ , so that  $\theta$  is uniformly distribution over  $[0, 1]$

- 1 Plot the functions  $f(x | 0)$  and  $f(x | 1)$  against  $x \in [0, 1]$  on the same graph. Comment briefly on the likely magnitude of the observation  $X$  if the pollution index  $\theta$  is high in the city.

## Exercise 3 II

- ② Show that the posterior distribution of  $\theta$  given  $X = x$  has the density function

$$\pi(\theta | x) = \left(x + \frac{1}{2}\right)^{-1} (1 + (2x - 1)\theta), \quad \theta \in [0, 1]$$

Describe briefly how the observation  $x$  affected our posterior belief about  $\theta$  in the two cases where  $x < 1/2$  and  $x > 1/2$

- ③ Based on the observation  $X = x$ , find the Bayes interval for  $\theta$  which has length  $1/3$  and the maximum posterior coverage probability if
- ①  $x < 1/2$
  - ②  $x > 1/2$

## Exercise 3 III

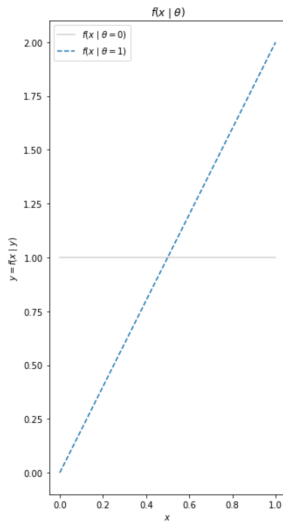
- ④ If a second sample of air is to be collected from the city independently of the first sample, and its pollutant content observed to be  $100Y\%$ , show that the posterior predictive distribution of  $Y$  given  $X = x$  has the density function

$$g^*(y | x) = \frac{1}{3} \left( x + \frac{1}{2} \right)^{-1} (x + 1 + (4x + 1)y), \quad y \in [0, 1].$$

- ⑤ If the first sample is not polluted, that is  $X = 0$ , predict the pollutant content  $Y$  of the second sample using the mean of the posterior predictive distribution  $g^*$  given in (d) above.

# Exercise 3: Solution I

- 1 The density functions are plotted as



## Exercise 3: Solution II

- 2 The posterior density is

$$\pi(\theta | x) \propto f(x | \theta)\pi(\theta) = 1 + 2\theta \left(x - \frac{1}{2}\right)$$

with normalizing constant

$$\int_0^1 f(x | \theta)\pi(\theta)d\theta = x + \frac{1}{2}$$

That is

$$\pi(\theta | x) = \left(x + \frac{1}{2}\right)^{-1} \left(1 + 2\theta \left(x - \frac{1}{2}\right)\right)$$

If  $x < 1/2$ ,  $\pi(\theta | x)$  is a decreasing function in  $\theta$ , so we have a stronger belief in small  $\theta$ . If  $x > 1/2$ ,  $\pi(\theta | x)$  is an increasing function in  $\theta$  and we have a stronger belief in large  $\theta$ .

## Exercise 3: Solution III

- 3 For  $x < 1/2$ ,  $\pi(\theta | x)$  is a downward sloping straight line from 0 to 1, so the Bayes interval with length  $1/3$  is  $[0, 1/3]$ ; For  $x > 1/2$ ,  $\pi(\theta | x)$  is an upward sloping straight line from 0 to 1, so the Bayes interval with length  $1/3$  is  $[2/3, 1]$ .
- 4 The predictive density is

$$\begin{aligned} g^*(y | x) &= \int_0^1 f(y | \theta) \pi(\theta | x) d\theta \\ &= \int_0^1 [1 + (2y - 1)\theta](x + 1/2)^{-1} [1 + (2x - 1)\theta] d\theta \\ &= (x + 1/2)^{-1} \left\{ 1 + 2(x + y - 1)\frac{1}{2} + \frac{1}{3}(2y - 1)(2x - 1) \right\} \\ &= \frac{4xy + x + y + 1}{3(x + 1/2)} \end{aligned}$$

## Exercise 3: Solution IV

- ⑤ Given the observed  $X = 0$ , the mean of  $Y$  is

$$\int_0^1 yg^*(y | x)dy = \frac{1}{3}(x + 1/2)^{-1} \left[ \frac{1}{2}(x + 1) + \frac{1}{3}(4x + 1) \right], \quad x = 0$$

Final answer is  $\frac{5}{9}$ .