



Clinical and Epidemiological Virology,
Rega Institute, Department of Microbiology
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Estimating evolutionary rates and divergence times....

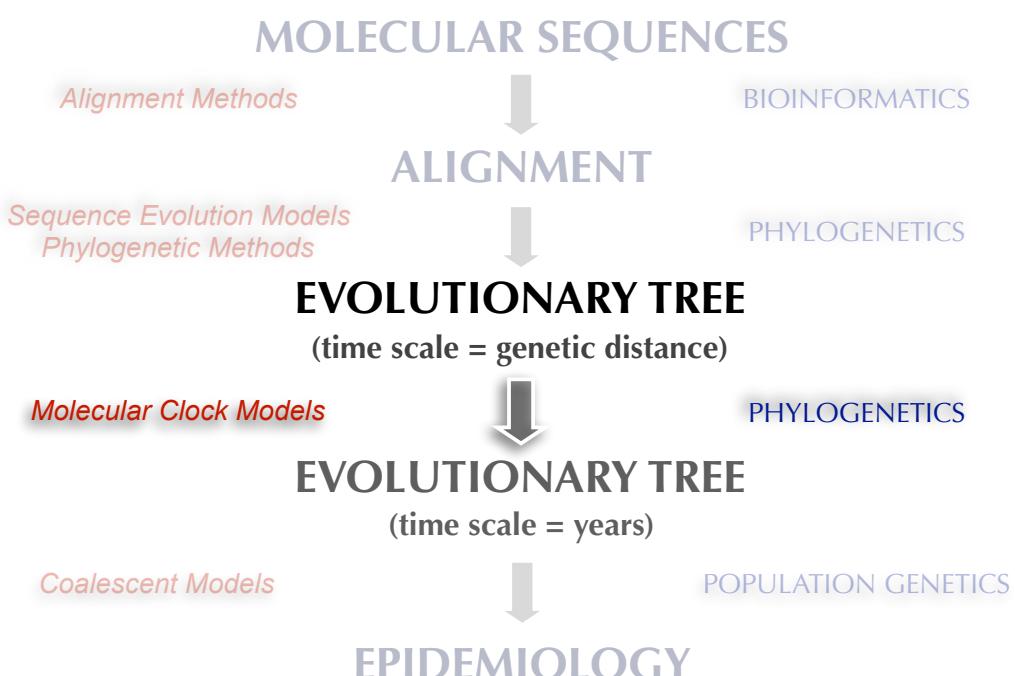
...and a bit of model testing

Philippe Lemey¹ and Marc Suchard²

1.Regia Institute, Department of Microbiology and Immunology, K.U.
Leuven, Belgium.

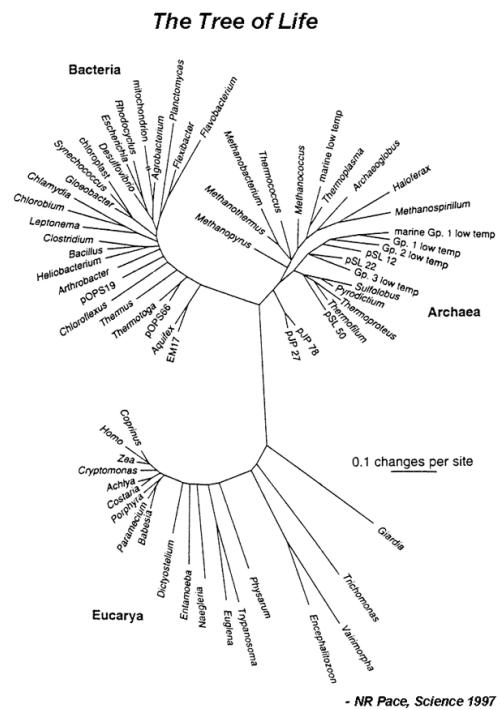
2.Departments of Biomathematics and Human Genetics, David
Geffen School of Medicine at UCLA. Department of Biostatistics,
UCLA School of Public Health

SISMID, July 10-12, 2019



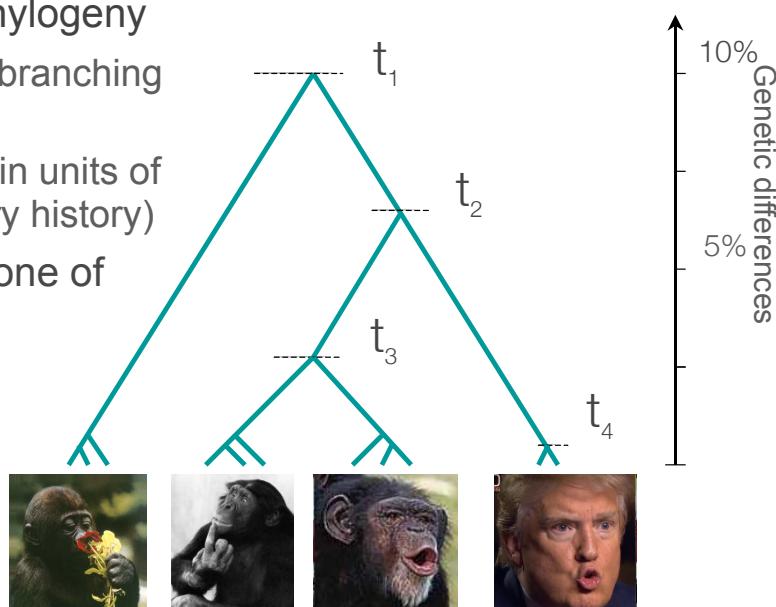
Molecular phylogenies

- most molecular phylogenies
 - ▶ are unrooted (or the rooting is due to prior information)
 - ▶ have branch lengths representing genetic change



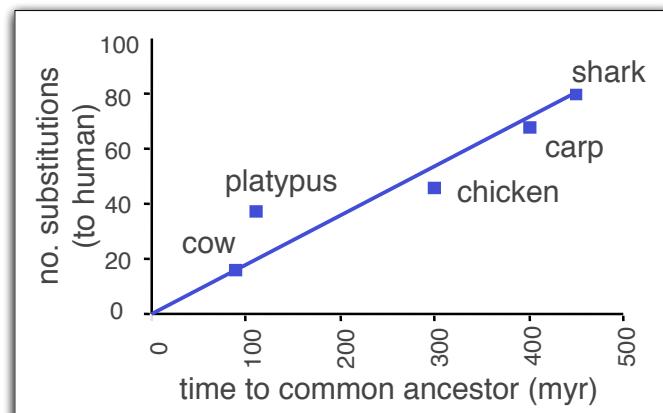
Molecular phylogenies

- the ideal molecular phylogeny
 - ▶ is rooted (implies a branching order)
 - ▶ has branch lengths in units of time (an evolutionary history)
- how do we construct one of these trees?



A constant evolutionary rate through time

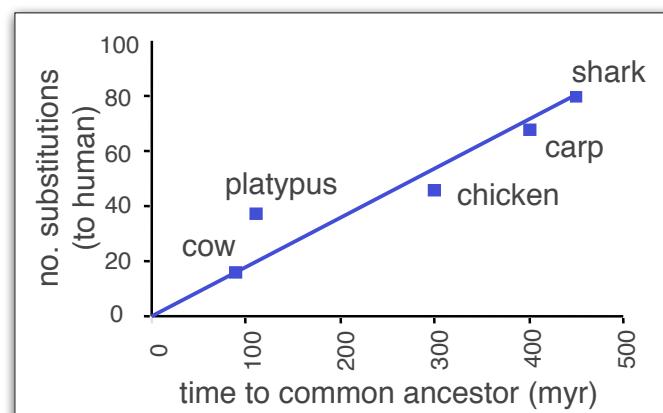
- to obtain a timed phylogeny, the evolutionary model must assume a relationship between the accumulation of genetic diversity and time



- Zuckerkandl and Pauling (1962): the rate of amino acid replacements in animal haemoglobins was roughly proportional to real time, as judged against the fossil record

A constant evolutionary rate through time

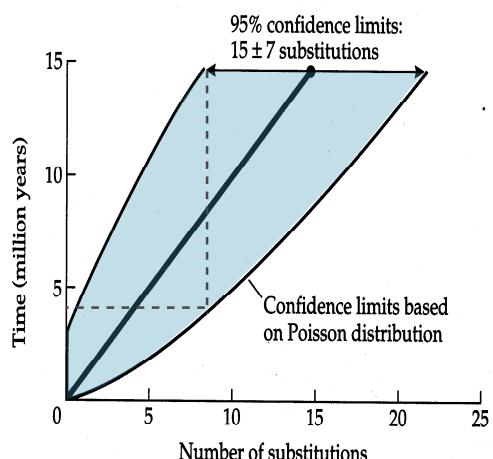
- the *molecular clock* is particularly striking when compared to the obvious differences in rates of morphological evolution...



The molecular clock is not a metronome

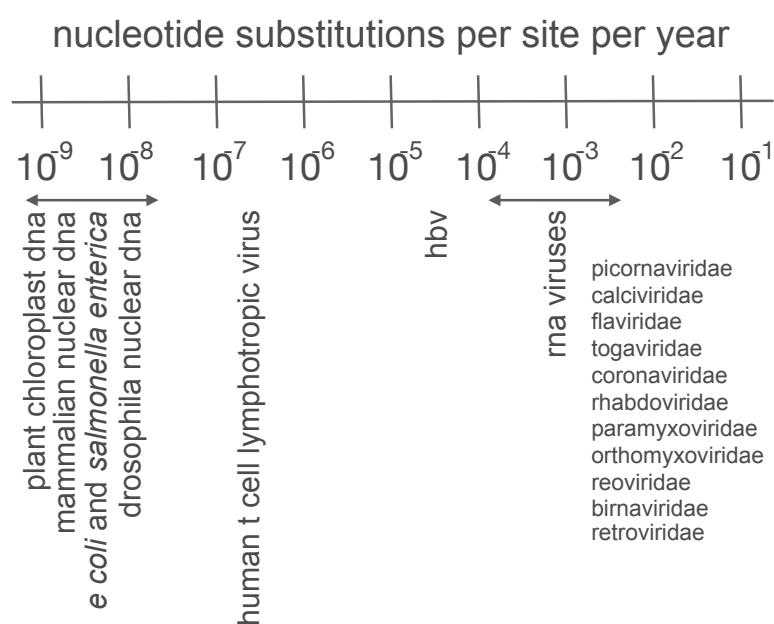
- if mutation every MY with Poisson variance

- 95% of the lineages 15MY old have 8-22 substitutions
- 8 substitutions also could be < 5 MY old



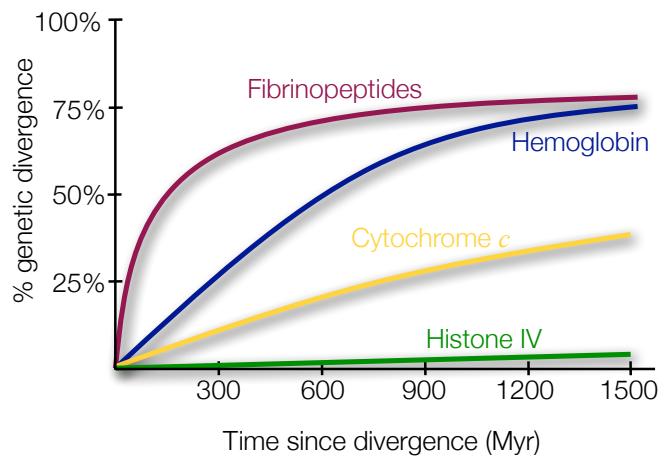
‣ Molecular Systematics, p532.

And there is no global molecular clock



And there is no global molecular clock

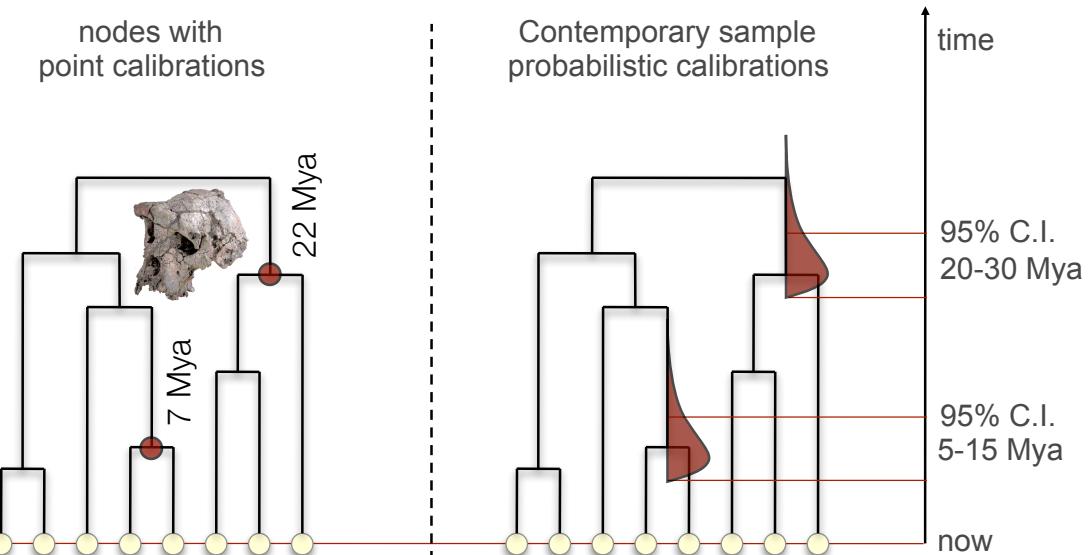
- different genes, different profiles
- variation in mutation rate?
- variation in selection genes coding for some molecules under very strong stabilizing selection



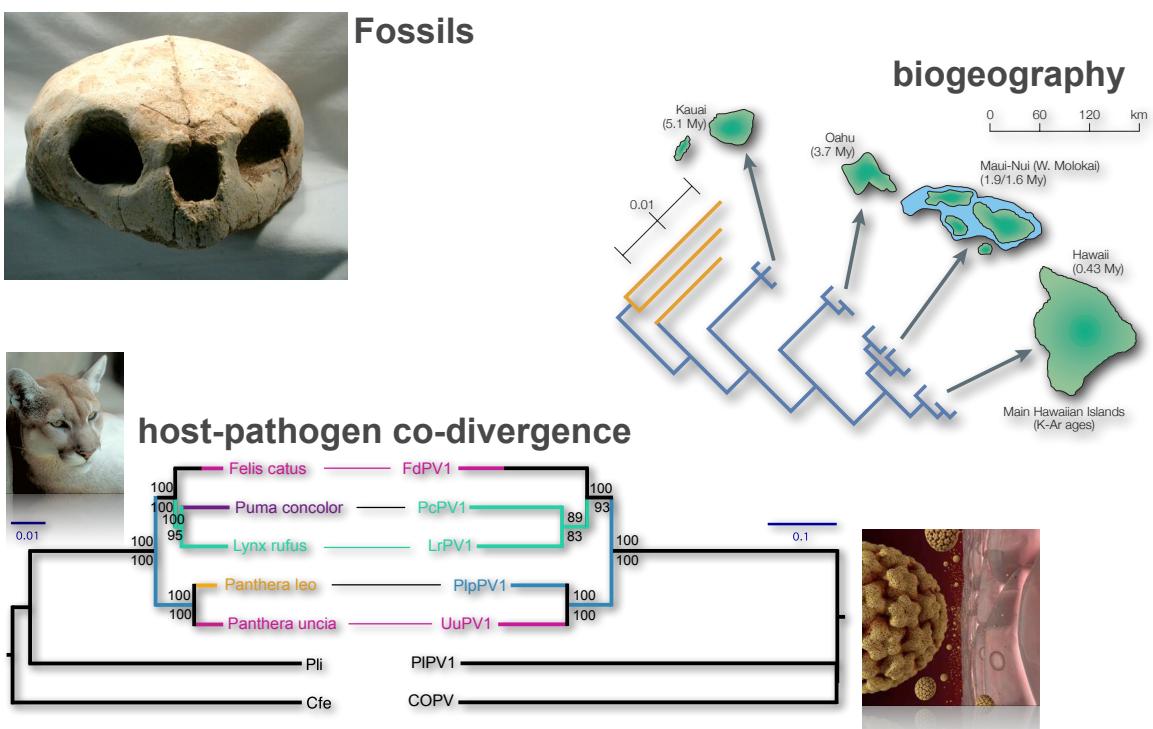
calibrating the molecular clock



From substitution units to time units

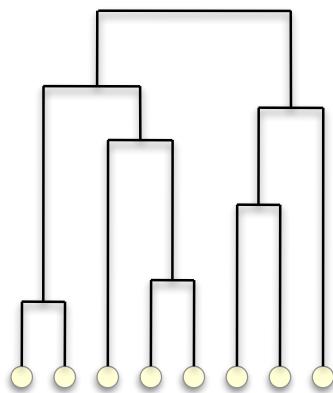


Node Calibrations

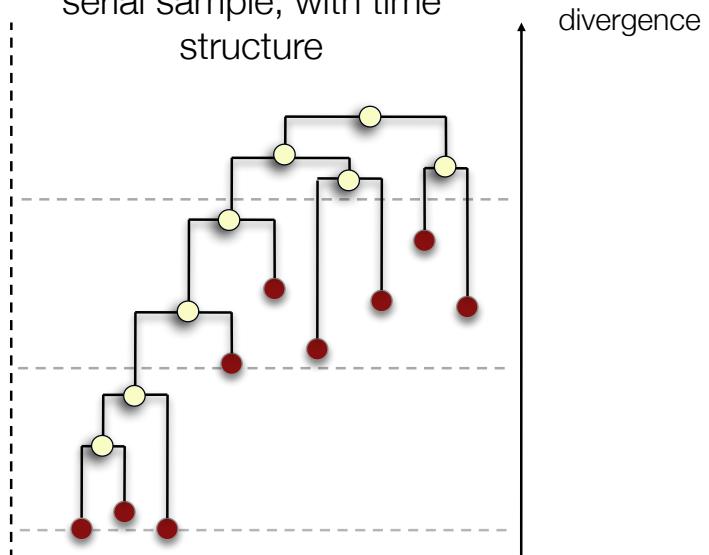


Calibration using sampling times

contemporary sample,
no time structure



serial sample, with time
structure



Tip calibration: two major applications



RNA viruses
evolve quickly:
 10^{-3} - 10^{-5}
substitutions per
site per year.

- Substitutions accumulate
between the times of sampling



ancient DNA
data sets of
radiocarbon-dated
specimens

- Serially sampled sequences or
heterochronous sequences

LETTER

<https://doi.org/10.1038/s41586-018-0097-z>

Ancient hepatitis B viruses from the Bronze Age to the Medieval period

Barbara Mühlemann¹, Terry C. Jones^{1,2}, Peter de Barros Damgaard³, Morten E. Allentoft¹, Irina Shevchenko⁴, Andrey Logvin⁴, Emma Usmanova⁵, Irina P. Panysheksina⁶, Bazartseren Boldgiv⁷, Tsevel Bazartseren⁸, Kadicha Tashbaeva⁹, Victor Merz¹⁰, Nina Lau¹¹, Václav Smrká¹², Dmitry Vovk¹³, Egor Kitov¹⁴, Andrej Epimakhov¹⁵, Dalia Pokutta¹⁶, Magdalena Vizec¹⁷, T. Douglas Price¹⁸, Vyacheslav Moiseyev¹⁹, Alexander J. Hansen²⁰, Ludovic Orlando^{1,20}, Simon Rasmussen²¹, Martin Sikora²², Lasse Viitner²³, Albert D. M. E. Osterhaus²⁴, Derek J. Smith²⁵, Dieter Glebe^{26,27}, Ron A. M. Fouchier²⁸, Christian Drosten^{29,30}, Karl-Göran Sjögren¹, Kristian Kristiansen¹⁸ & Eske Willerslev^{1,7,29,30}

incorporating sampling time: naive method

observed number of substitutions
or genetic divergence

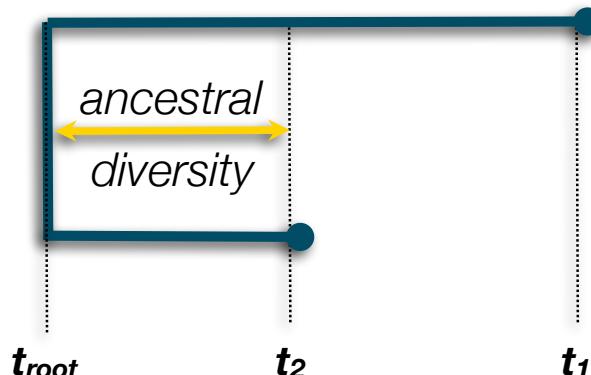


sampling time 1
 t_1

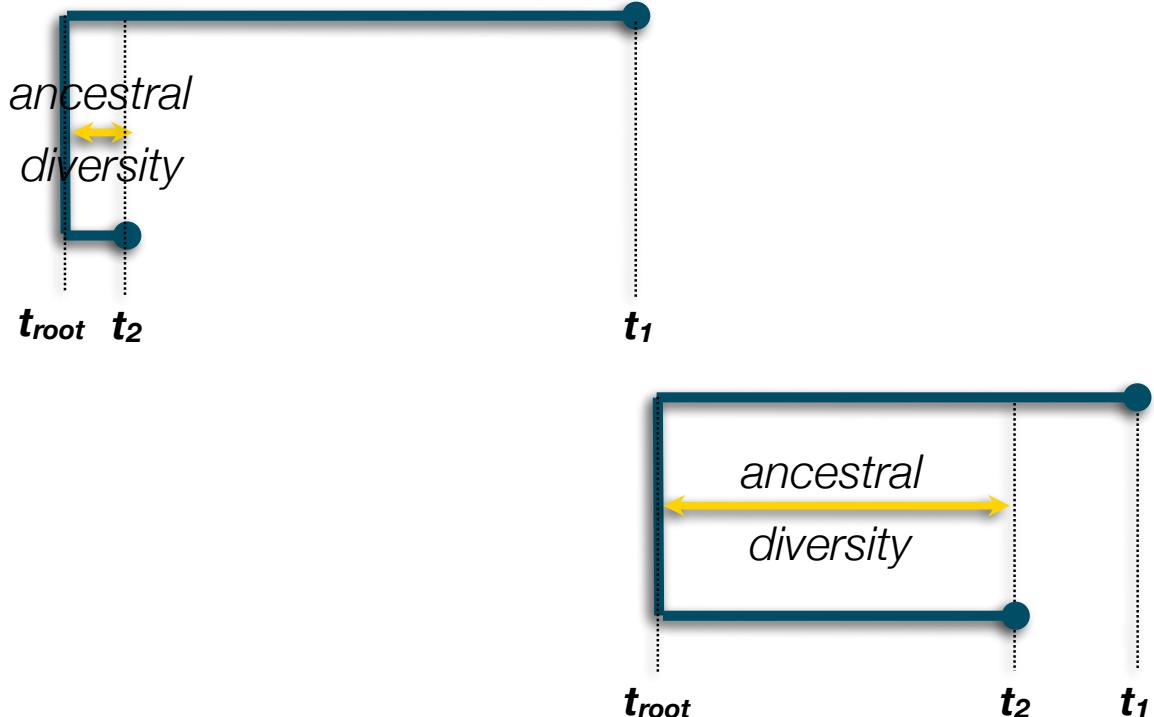
sampling time 2
 t_2

$$\text{substitution rate, } \mu = d / |t_1 - t_2|$$

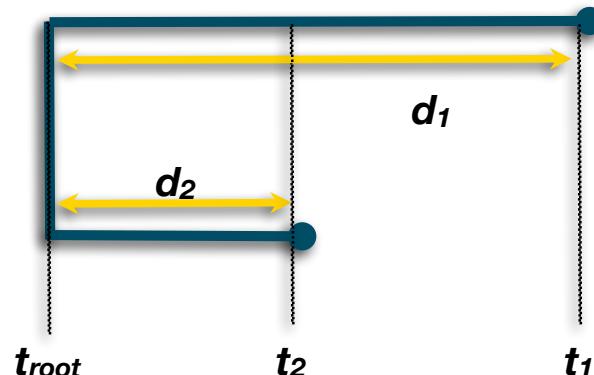
incorporating sampling time: naive method



incorporating sampling time: naive method



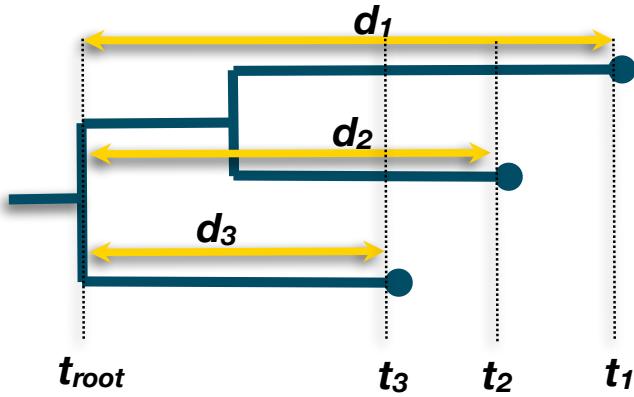
incorporating sampling time: naive method



$$\mu = (d_1 - d_2) / (t_1 - t_2)$$

linear regression

$$\mu = d_i / (t_i - t_{root})$$



can be rearranged:

$$d_i = \mu (t_i - t_{root})$$

$$E[d_i] = \mu \cdot t_i - \mu \cdot t_{root}$$

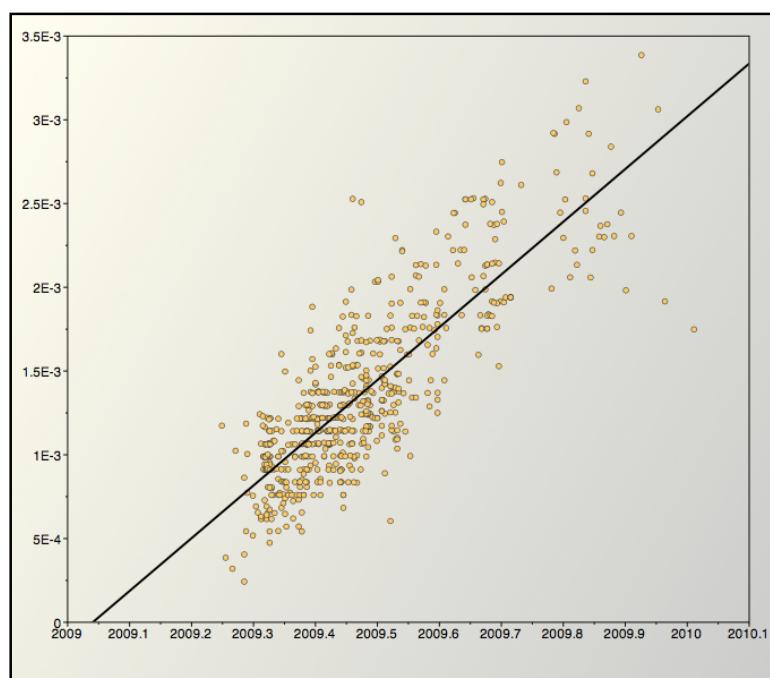
gradient is: μ

y-intercept is: $-\mu \cdot t_{root}$

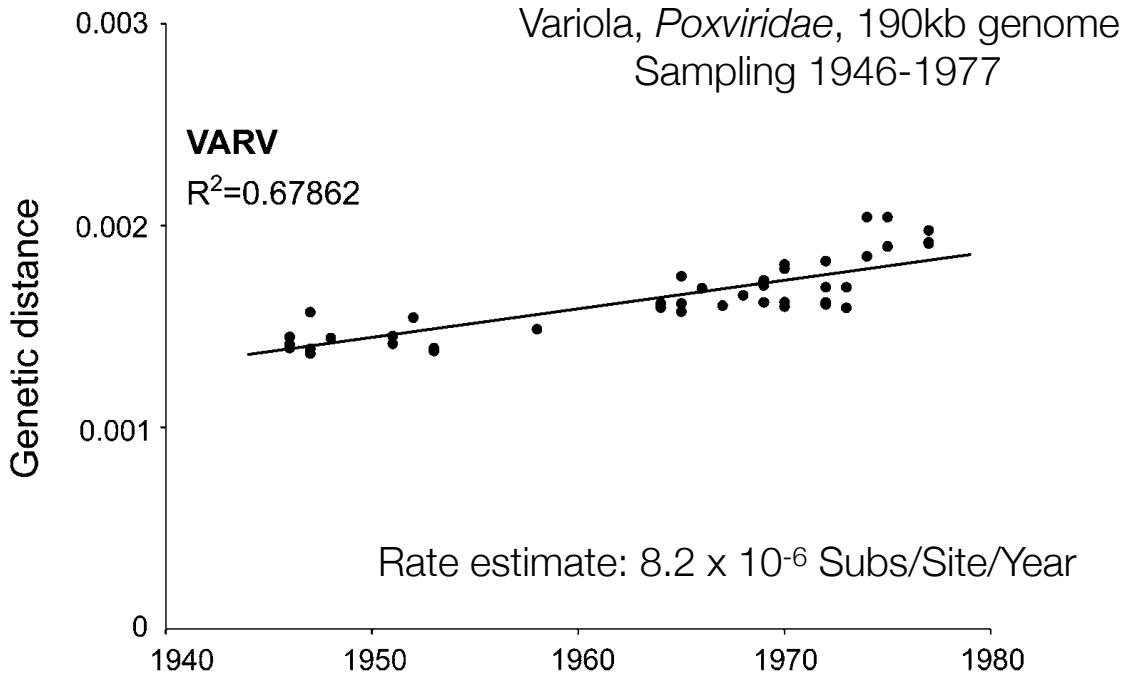
x-intercept is: t_{root}

Estimating the time-scale

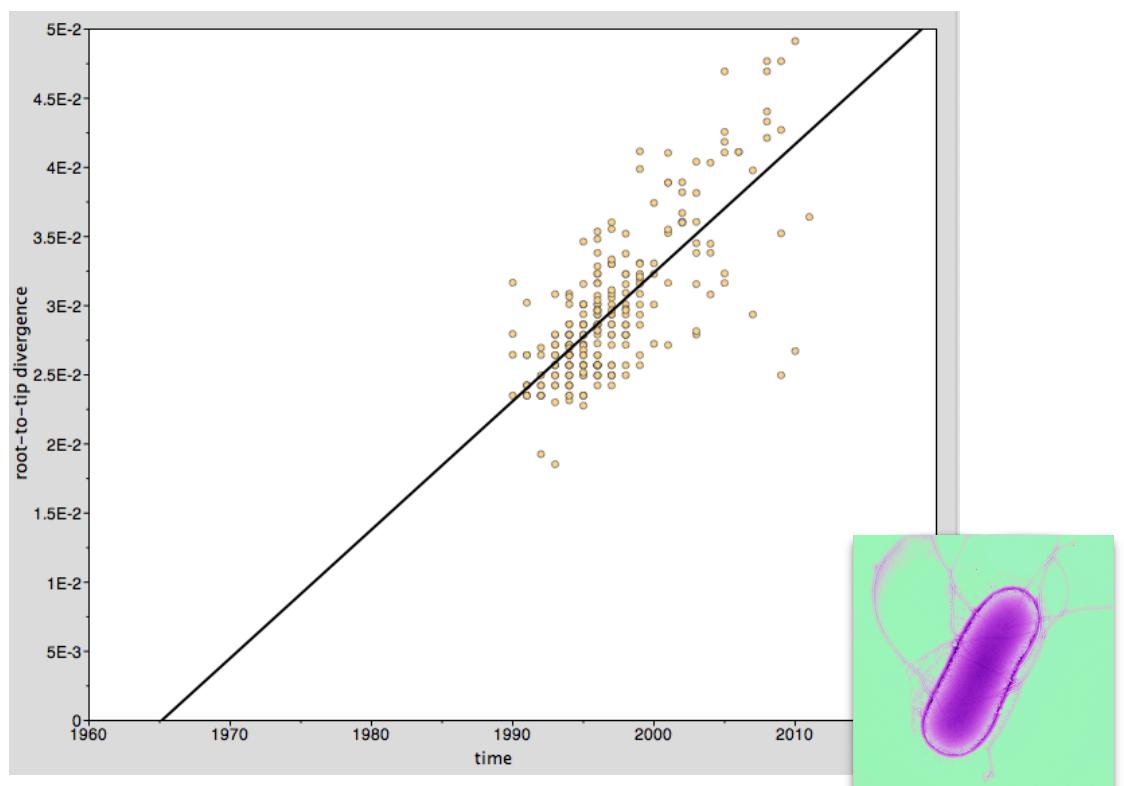
- H1N1/09 ‘Swine Flu’
- Rate: $3.14E^{-3}$ mutations/genomic site/year
- tMRCA: 2009.041 (15-Jan-2009)
- Correlation: 0.83
- R^2 : 0.69



A DNA virus (smallpox)



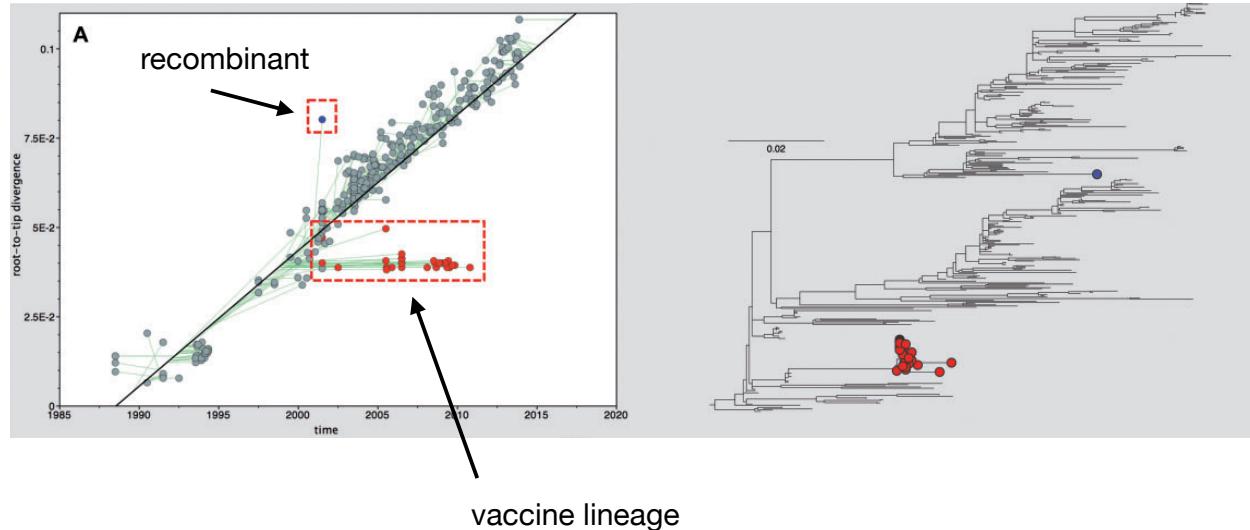
Salmonella Typhimurium



Diagnostic tool

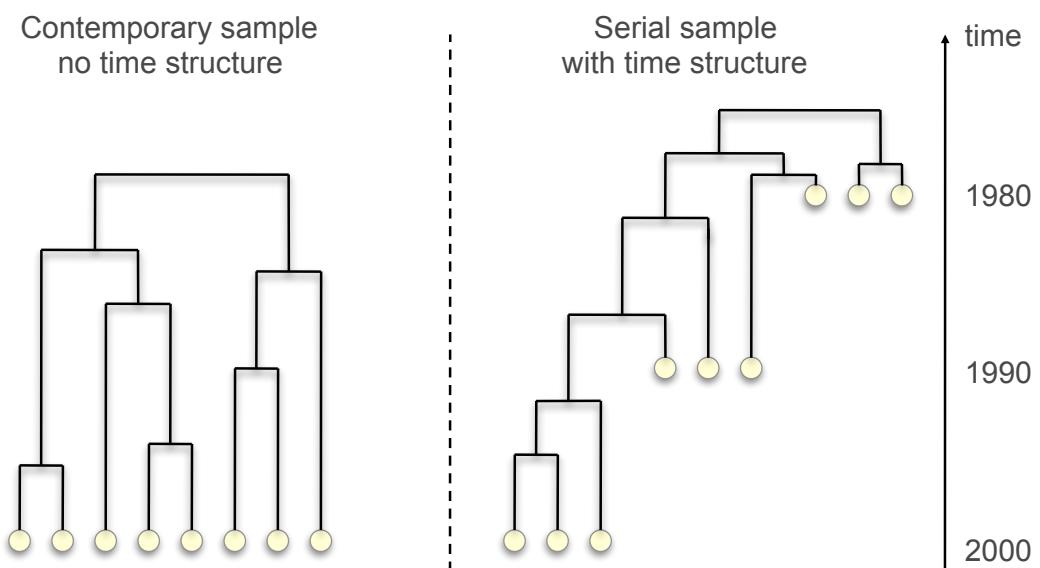


- divergence accumulation
- outliers



► Rambaut A. et al. (2016) *Virus Evolution*, **2(1)**, vew07.

Time structure via tip calibration



► Rambaut A. (2000) *Bioinformatics*, **16**, 395-399.

Relaxing the molecular clock

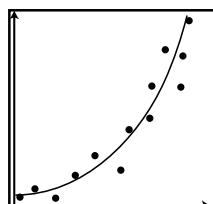
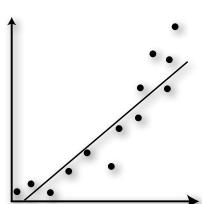


Clock versus non-clock

- unconstrained (unrooted) Felsenstein model:
Felsenstein (1981) *JME*, 17: 368 - 376
 - each branch has its own rate independent of all others
 - time and rate are confounded and can only be estimated as a compound parameter (branch lengths)
- strict molecular clock:
Zuckerkandl & Pauling (1962) in *Horizons in Biochemistry*, pp. 189–225
 - all lineages evolve at the same rate
 - allows the estimation of the root of the tree and dates of individual nodes

Need for a relaxed molecular clock

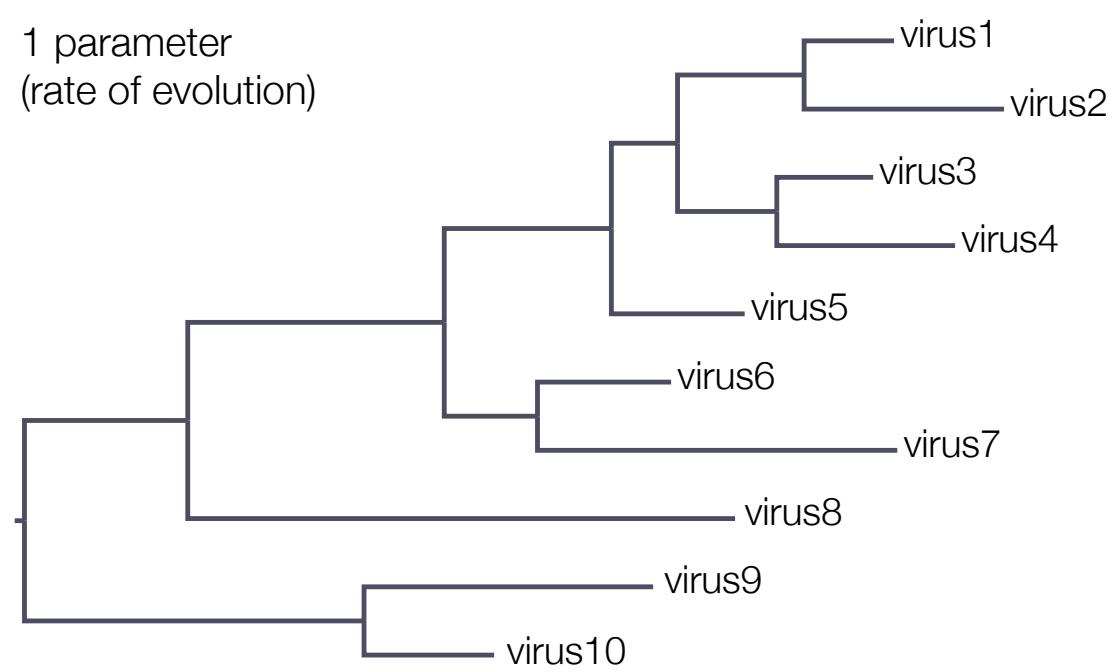
- the unrooted model of phylogeny and the strict molecular clock model are two extremes of a continuum.
- dominate phylogenetic inference
- but both are biologically unrealistic:
 - the real evolutionary process lies between these two extremes
 - model misspecification can produce positively misleading results



‣ Pybus (2006) *Genome Biol.* **4**, e151

'strict' molecular clock

1 parameter
(rate of evolution)

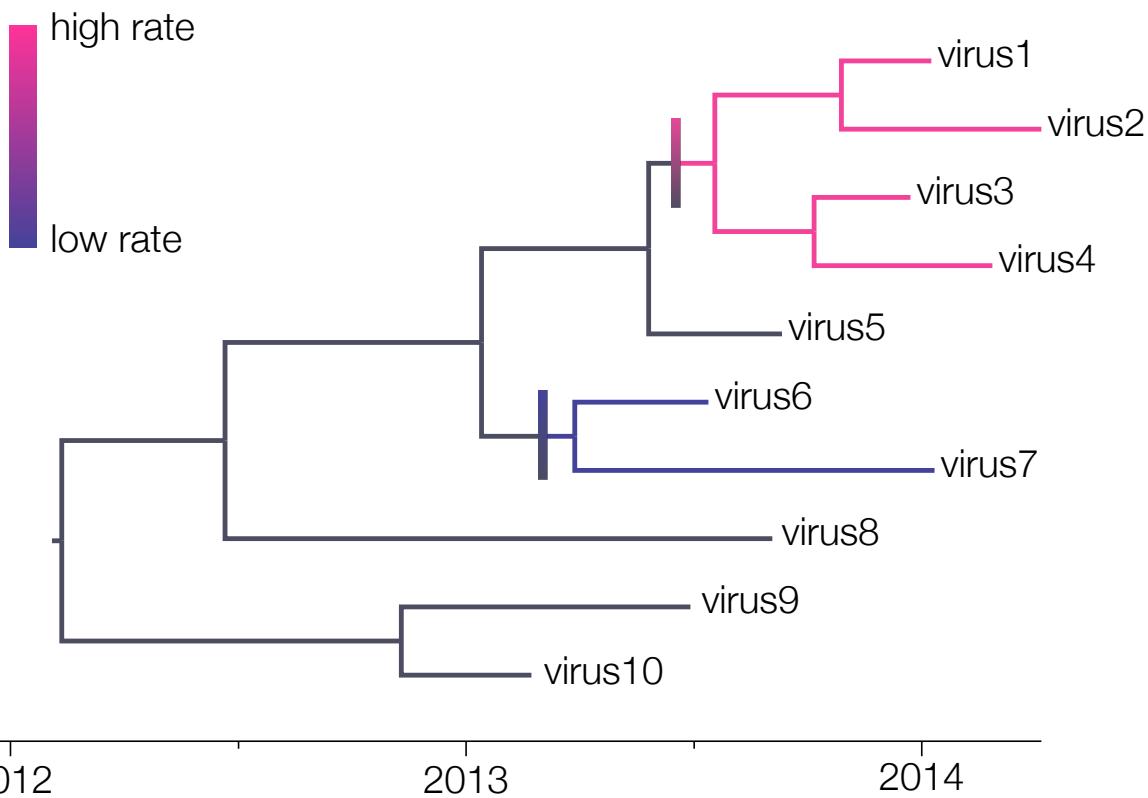


2012

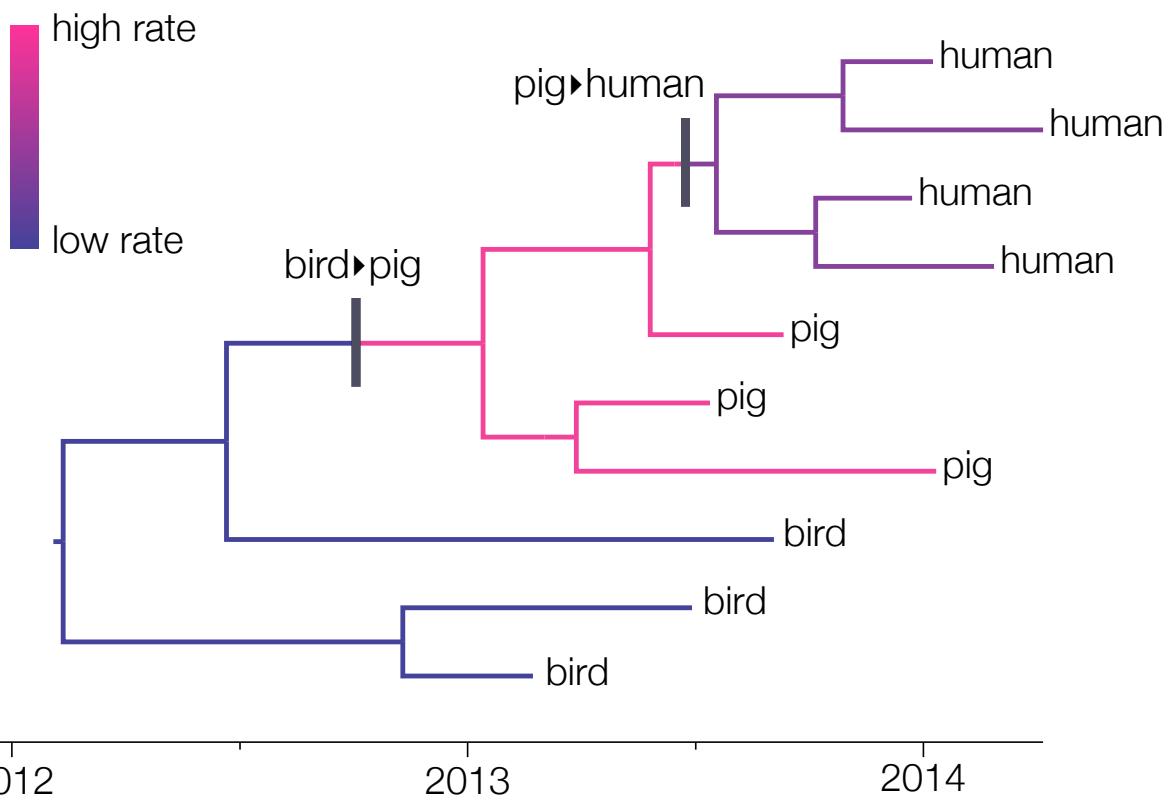
2013

2014

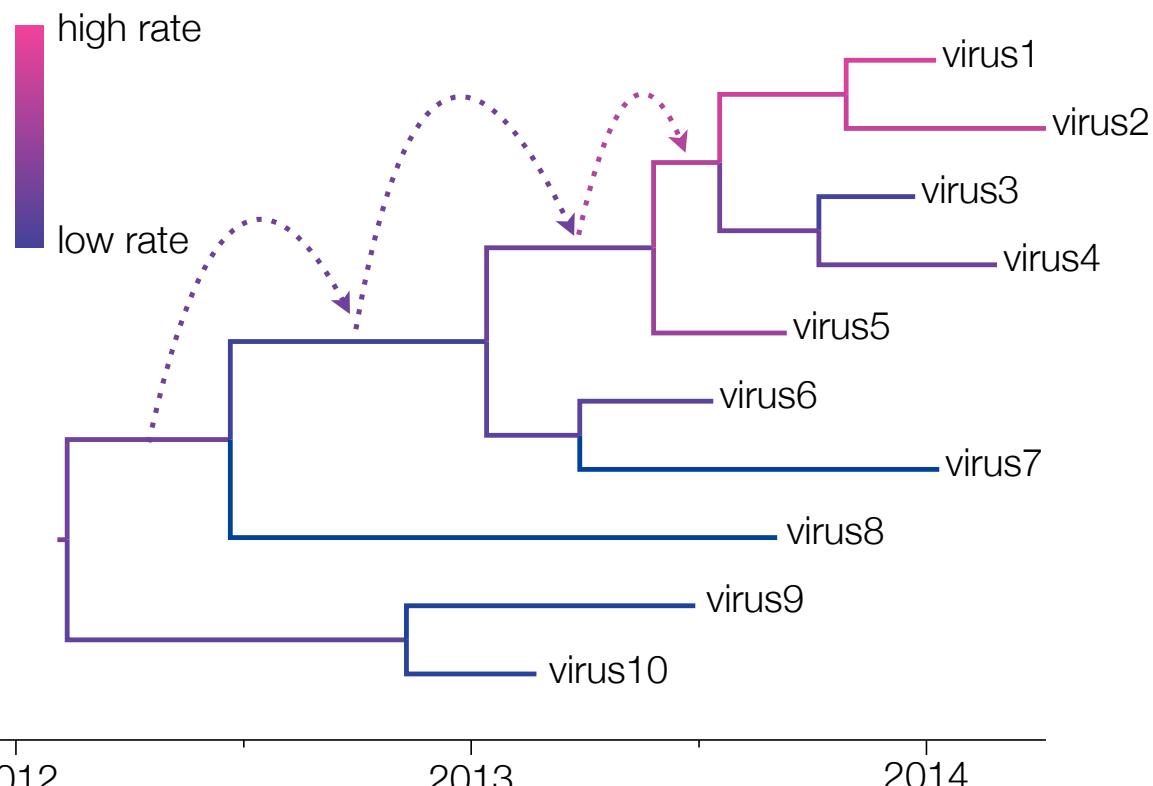
'local' molecular clock



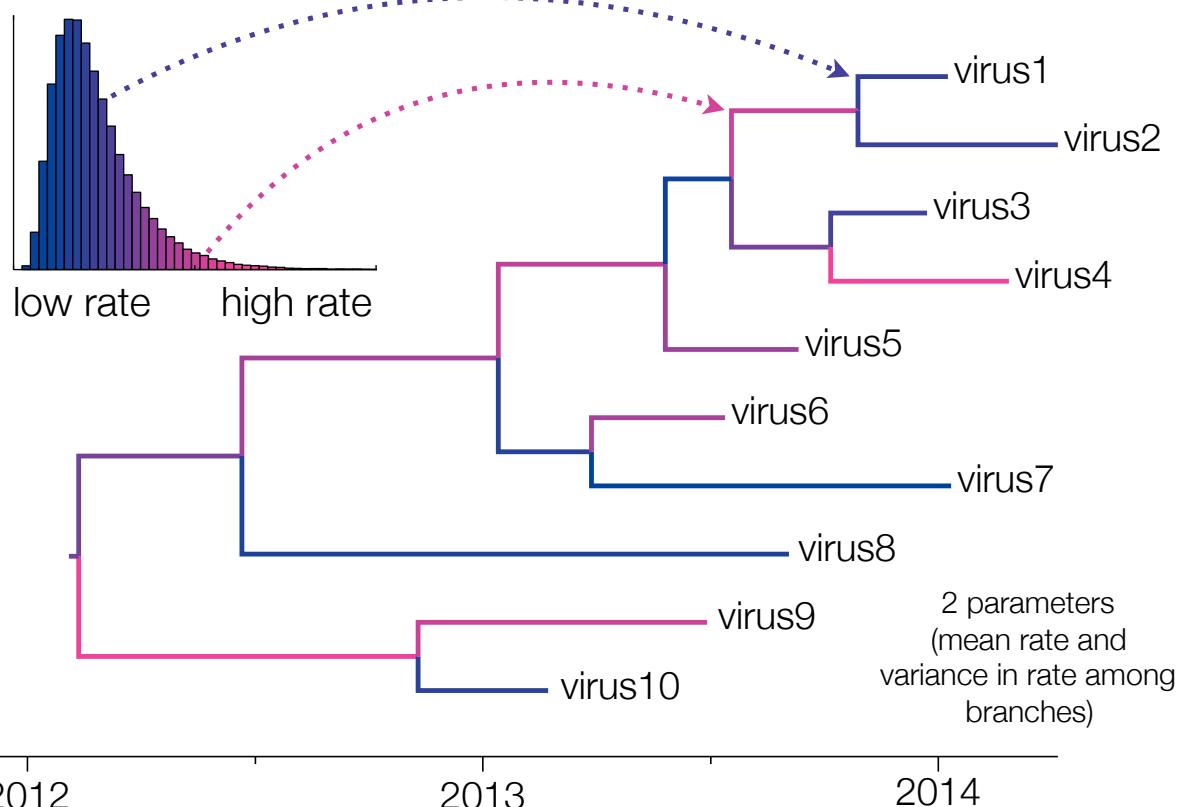
host-specific local clock



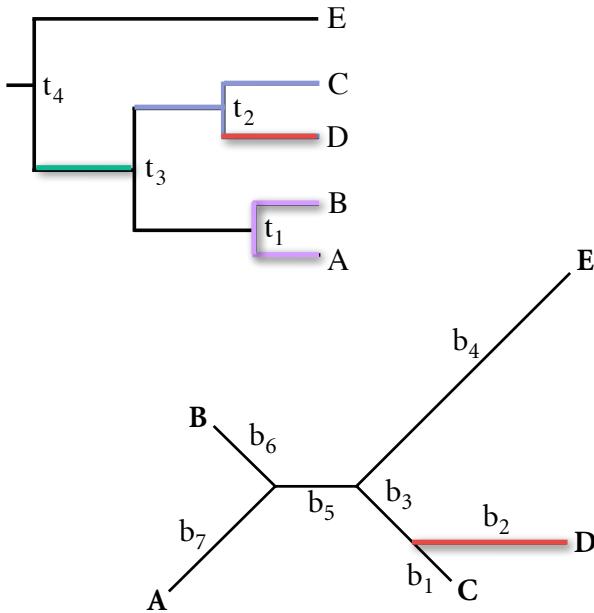
autocorrelated relaxed clock



lognormal uncorrelated relaxed clock



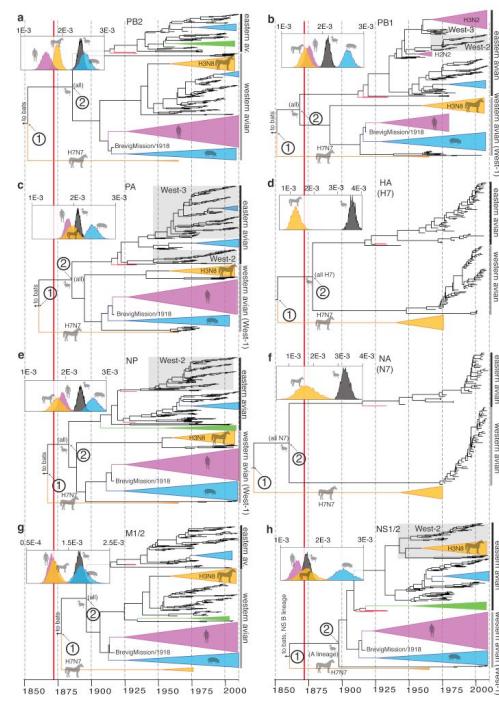
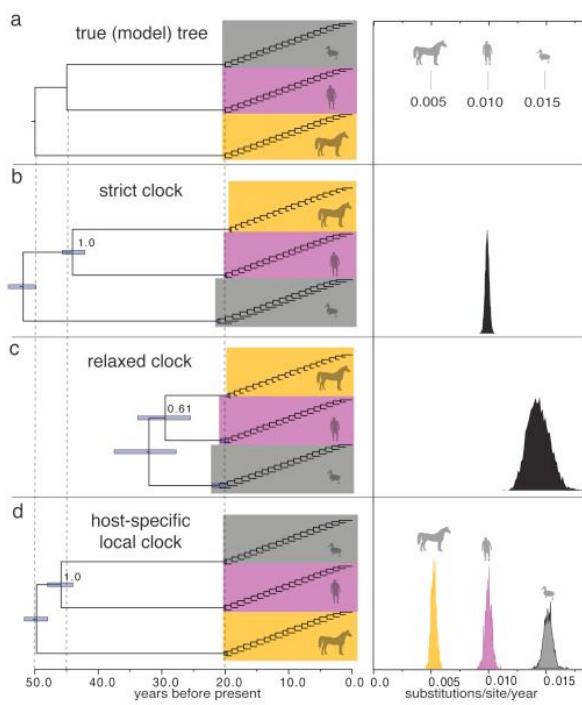
Relaxed clocks: (1) local molecular clocks



- specify H_0 beforehand
- problem of identifiability

‣ Yoder and Yang (2000) Mol Biol & Evol 17: 1081-1090.

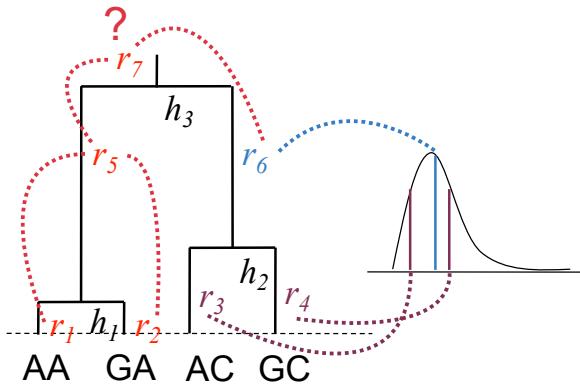
Bayesian local clocks



Worobey et al., Nature, 2014; 508(7495): 254–257

Autocorrelated relaxed clocks

- rates for each branch are drawn from a distribution centred on the rate of the ancestor
 - but what is the rate at the root?
 - A prior degree of autocorrelation?
 - (not currently possible to do phylogenetic inference)

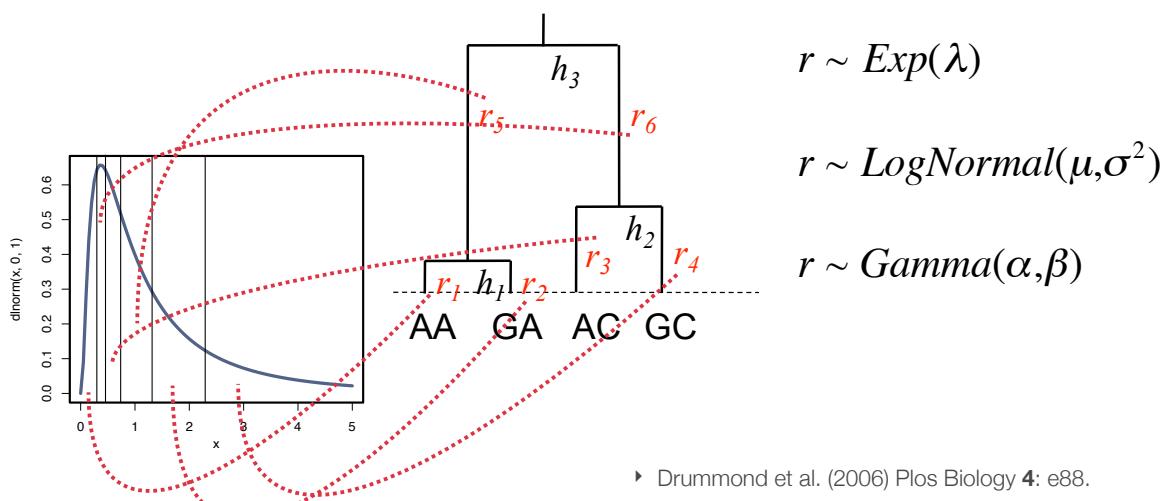


$$r_i \sim \text{LogNormal}(r_{A(i)}, \sigma^2 \Delta t_i)$$

‣ e.g., Thorne JL, Kishino H, Painter IS (1998) Mol Biol & Evol **15**: 1647-1657.

Uncorrelated relaxed clocks

- rates for each branch are drawn independently from an identical distribution:



Bayesian evolutionary analysis sampling trees

- Given sequence data that is temporally spaced estimate true values of:

- substitution parameters (μ and Q)
- ancestral genealogy ($g = E_g, t_y$)
- tree topology
- dates of divergence
- population history (θ)

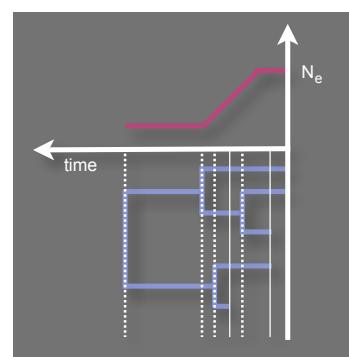
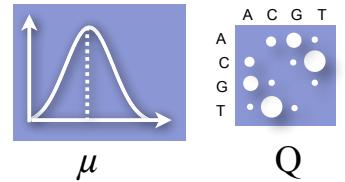
- Bayesian inference

$$P(g, \mu, \theta, Q | D) = \frac{1}{Z} \Pr\{D|g, \mu, Q\} f_g(g|\theta) f_\mu(\mu) f_\theta(\theta) f_Q(Q)$$

“relaxed phylogenetics and dating with confidence”

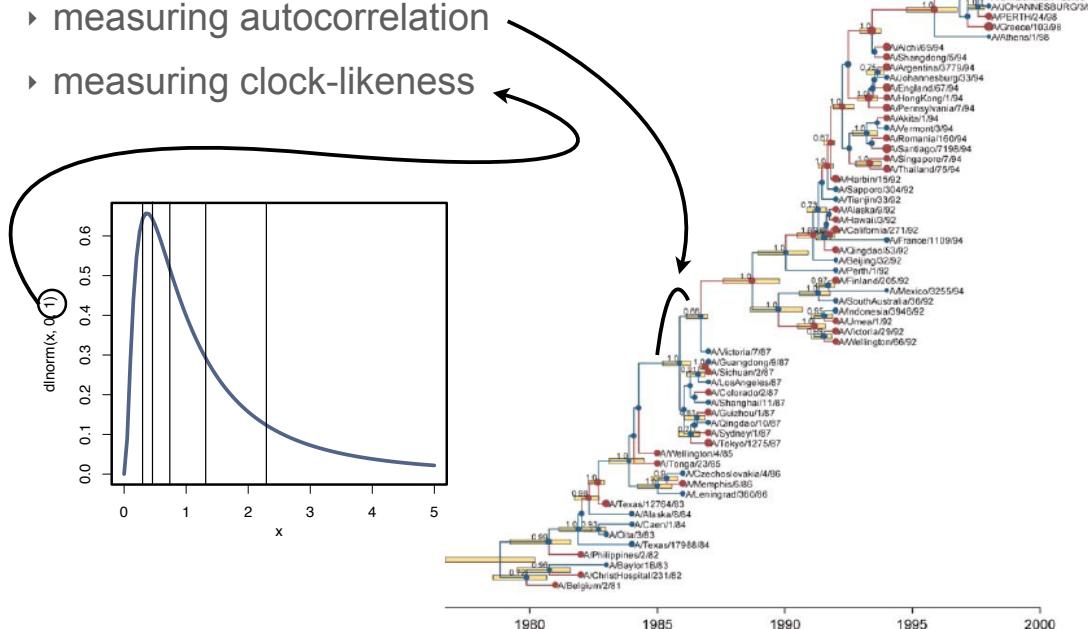
$$t = \{t_1, t_2, \dots, t_{2n-1}\}$$

$$R = \{r_1, r_2, \dots, r_{2n-1}\} \quad f(R|g) = f(R) = \prod_{i=1}^n \lambda e^{-\lambda r_i}$$

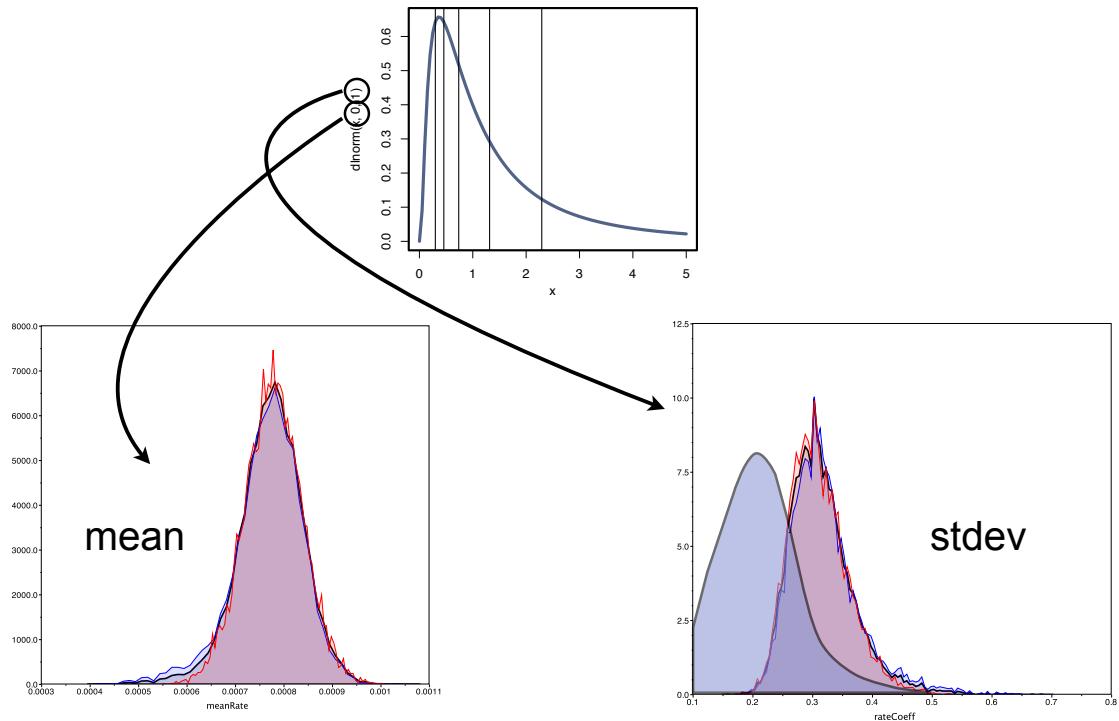


Uncorrelated relaxed clocks: example

- Phylogenetic inference
- measuring autocorrelation
- measuring clock-likeness



Evaluating clock-like behaviour?



Model testing using Bayes factors

- A Bayesian alternative to classical hypothesis testing: the Bayes factor (a summary of the evidence provided by the data in favor of one scientific theory, represented by a statistical model, as opposed to another; Kass & Raftery, 1995).

- Bayes factor
$$B_{01} = \frac{p(Y|M_1)}{p(Y|M_0)}$$

- When two models M_0 and M_1 are being compared, one defines the Bayes factor in favor of M_1 over M_0 as the **ratio of their respective marginal likelihoods**
- When there are unknown parameters, the Bayes Factor B_{01} has in a sense the form of a likelihood ratio

Model testing using Bayes factors

- However, the densities are obtained by integrating over parameter space:

$$p(Y|M) = \int_{\theta} p(Y|\theta, M) p(\theta|M) d\theta$$

- Posterior:

$$p(\theta|Y, M) =$$

$$\frac{p(Y|\theta, M) p(\theta|M)}{p(Y|M)}$$

← p(Y|M)

- So for model fit, the marginal likelihood $p(Y|M)$ or integrated likelihood, i.e. the normalizing constant (cancels out in the calculation of the MH acceptance ratio), is of primary importance, but awfully hard to calculate.

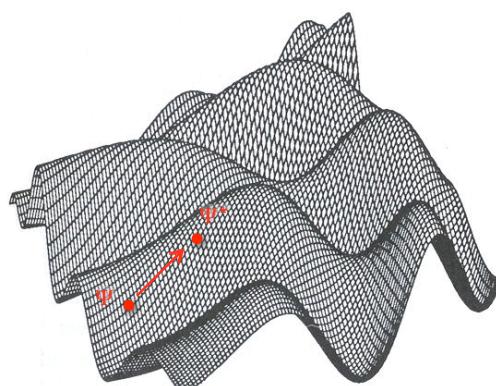
Reminder: MHG MCMC Sampling

The algorithm starts from a random state (θ) and 'proposes' a new state (θ^*)

The new state is accepted with probability:

$$\begin{aligned} R &= \min \left(1, \frac{p(\theta^*|D)}{p(\theta|D)} \times \frac{p(\theta|\theta^*)}{p(\theta^*|\theta)} \right) \\ &= \min \left(1, \frac{p(D|\theta^*) p(\theta^*)}{p(D|\theta) p(\theta)} \times \frac{f(\theta|\theta^*)}{f(\theta^*|\theta)} \right) \\ &\quad \xrightarrow{\text{the two marginal likelihoods cancel out and don't have to be computed !}} \\ &= \min \left(1, \frac{f(D|\theta^*)}{f(D|\theta)} \times \frac{f(\theta^*)}{f(\theta)} \times \frac{f(\theta|\theta^*)}{f(\theta^*|\theta)} \right) \end{aligned}$$

Likelihood ratio Prior ratio Proposal ratio



Calculating marginal likelihoods

Methods of general applicability:

- the posterior arithmetic mean estimator (pAME; Aitkin, 1991)
- the arithmetic mean estimator (AME/ILP; but a misnomer)
- the importance sampling estimators, and particularly the harmonic mean estimator (HME) (Newton and Raftery, 1994)
- the stabilized harmonic mean estimator (sHME) (Redelings and Suchard, 2005)

No additional analysis required

- path sampling (Gelman, 1998; Ogata, 1989), applied in phylogenetics (Lartillot and Philippe, 2006)
- stepping-stone sampling (Xie et al., 2011)
- generalised stepping-stone sampling (Fan et al., 2011; Baele et al., 2016)

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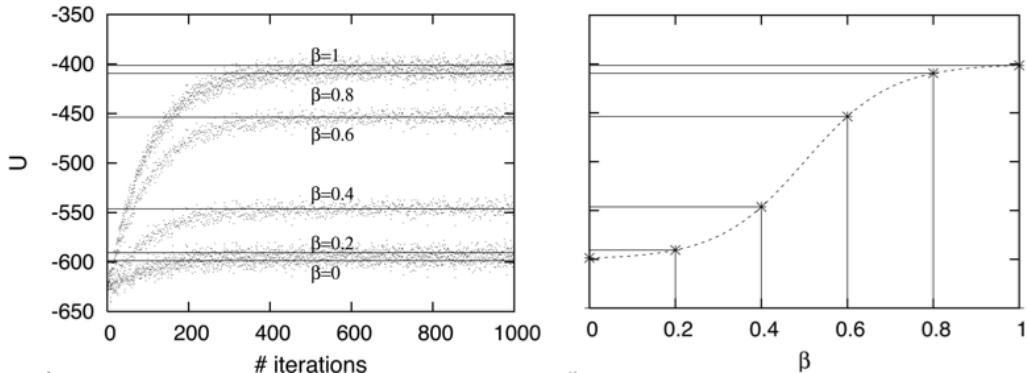
Additional analysis required

path sampling and stepping-stone sampling

- requires samples from a series of power posteriors, along a path between prior and posterior:

$$q_{\beta}(\theta) = p(Y | \theta, M)^{\beta} p(\theta | M)$$

reduces to the posterior when $\beta = 1$
reduces to the prior when $\beta = 0$



path sampling and stepping-stone sampling

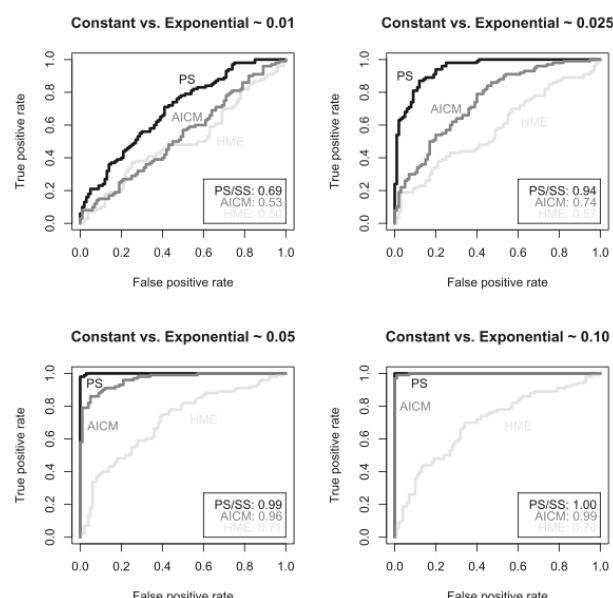


FIG. 2. Evaluation of log BF estimates using PS (SS yields an undistinguishable plot), AICM, and the HME to compare model fit, with four pairwise comparisons being shown: a constant population size versus an exponential population size with growth rates of 0.01, 0.025, 0.05, and 0.10. An increasingly strong discriminatory behavior (low false positive rates and high true positive rates) can be seen for PS (and SS) up to a growth rate of 0.10, whereas the HME retains questionable performance. AICM performance lies in between that of the HME and PS/SS. Color-coded area under the curve values are given at the bottom right of each plot.

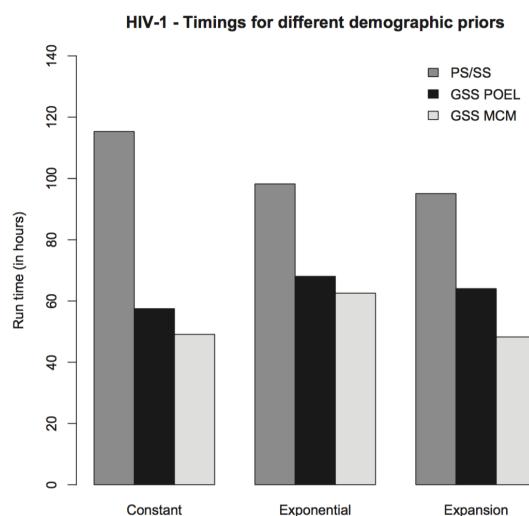
Generalised stepping-stone sampling

requires samples from a series of power posteriors, along a path between reference/working distribution and posterior:

$$q_\beta(\theta) = [p(Y | \theta, M)p(\theta | M)]^\beta p_0(\theta | M)^{1-\beta}$$

- reduces to the original SS method if the reference/working distribution is equal to the actual prior
- in practice, samples from the posterior distribution ($\beta = 1$) are used to parameterize the joint reference/working distribution $p_0(\theta|M)$
- we will use kernel density estimation (KDE) to construct reference/working priors for each of the parameters being estimated

GSS: decreased run time



- GSS does not need to explore the prior, which avoids computing the likelihood for highly unlikely parameter values, which may lead to numerical instabilities
- combined with a “shorter” path to be traversed, this leads to a considerable performance increase (dependent on the actual reference/working prior)

Bayesian model testing

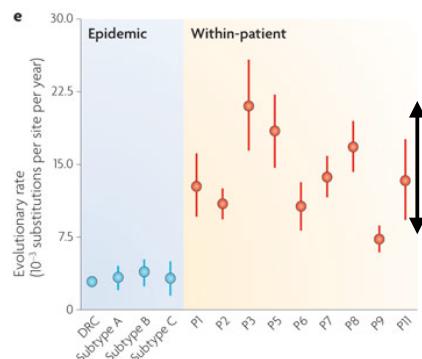
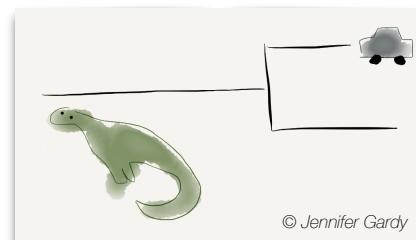
- Don't compare all possible model combinations (evolutionary model, clock models, coalescent tree prior, ...) to one another!
- Test/compare those models if
 - it is part of the hypothesis your testing,
 - or if your hypothesis test is sensitive to the model choice

Bayesian model selection vs model averaging

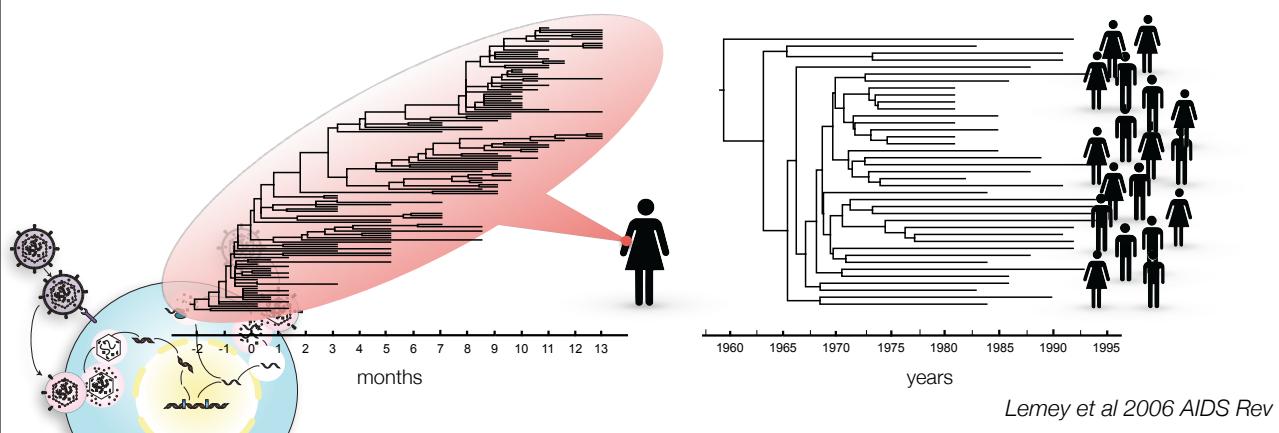
Model selection refers to the problem of using the data to select one model from the list of candidate models

Model averaging refers to the process of estimating some quantity under each model and then averaging the estimates according to how likely each model is.

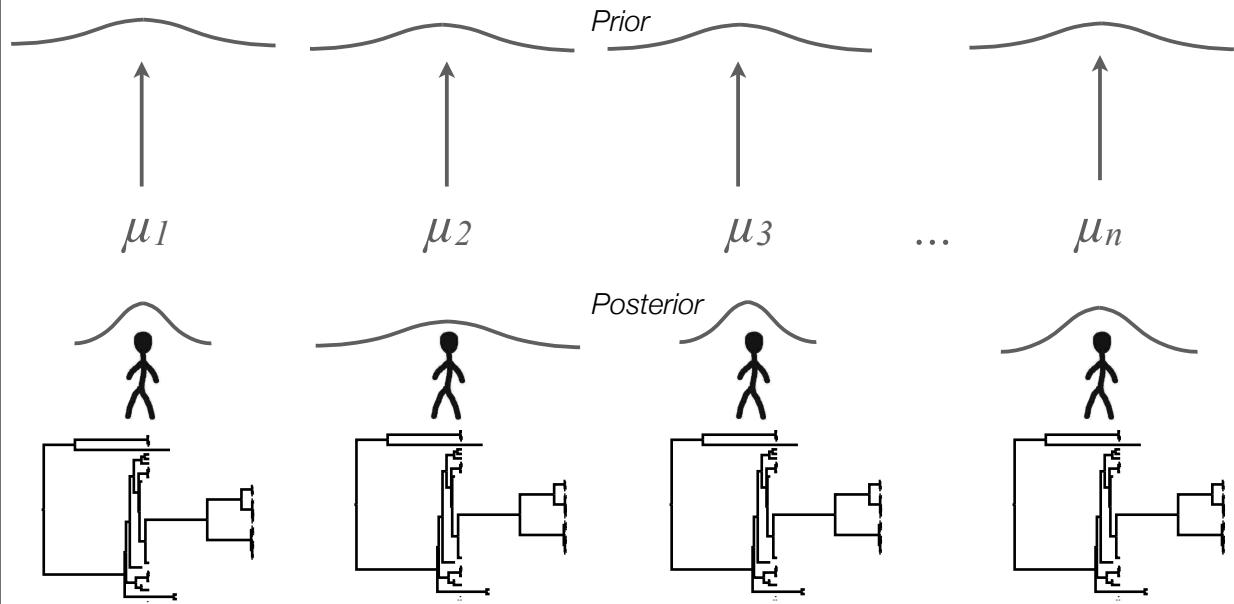
Extensions for testing evolutionary rate hypotheses



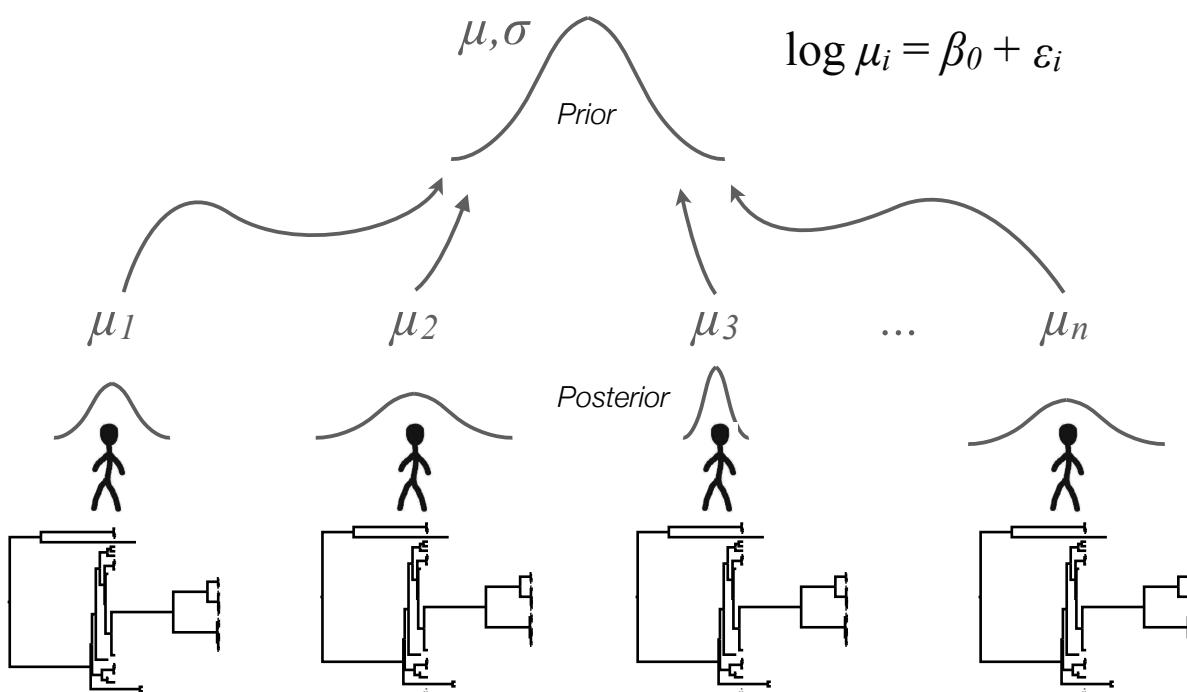
Pybus and Rambaut, NGR, 2009



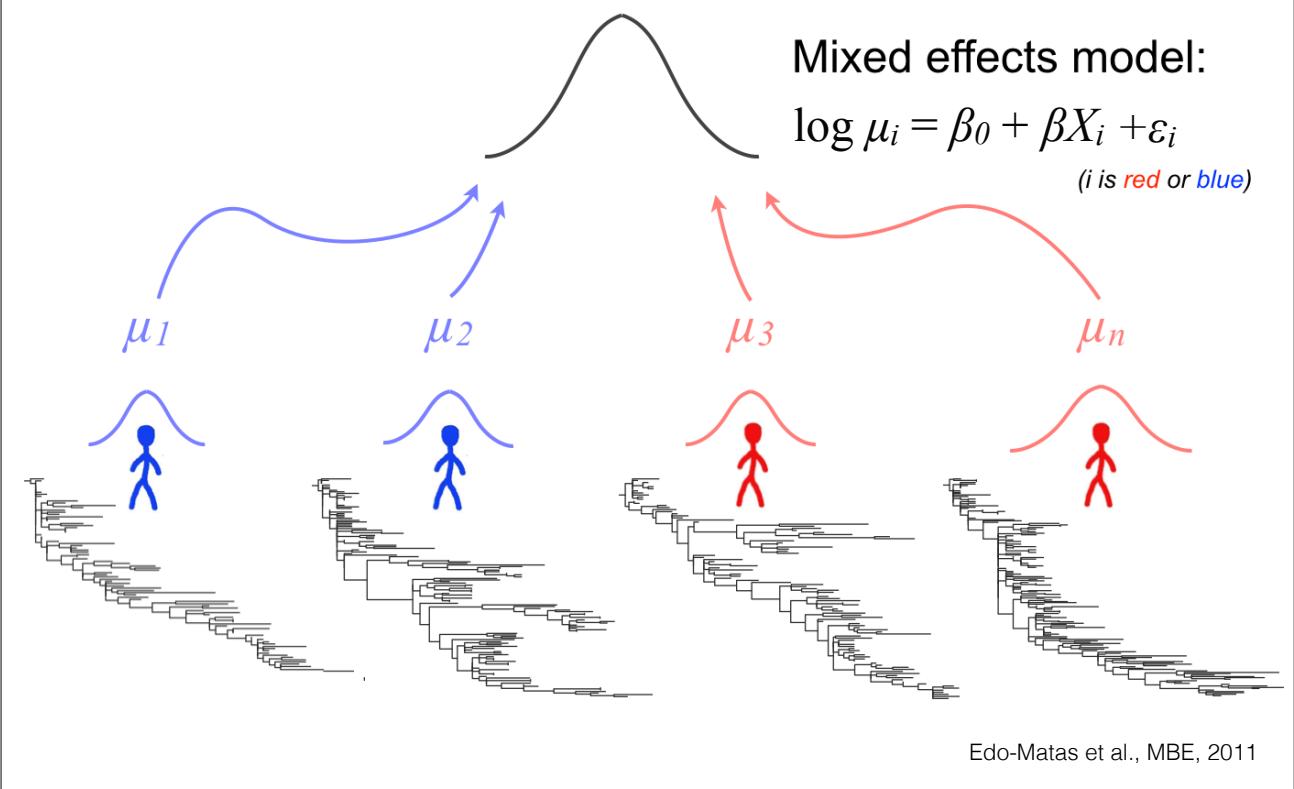
Independent parameter estimation



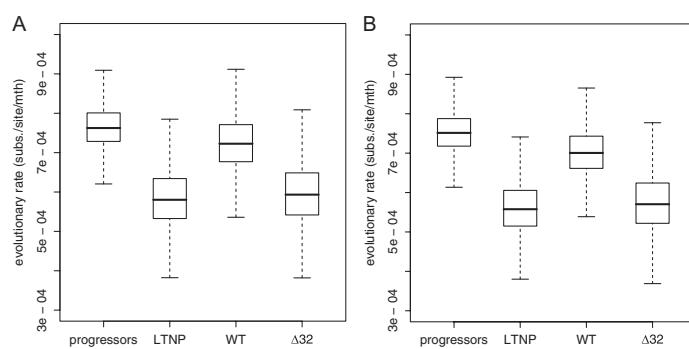
Hierarchical phylogenetic models



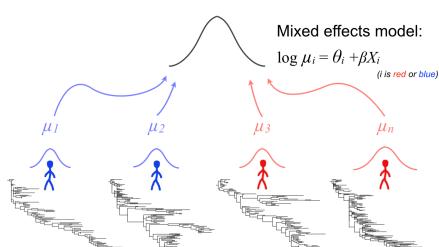
Hierarchical model with fixed effects



Hierarchical model with fixed effects



$$\log \theta_i = \beta_0 + \delta_{LTNP} \beta_{LTNP} LTNP_i + \delta_{Δ32} \beta_{Δ32} Δ32_i + \varepsilon_i$$



Evolutionary Parameter	Effect Support/Size	LTNP Effect
Nucleotide substitution rate	Posterior probability $\delta_{\text{effect}} = 1$	0.72
	BF _{effect}	2.6
Codon substitution rate	$\beta_{\text{effect}} \delta_{\text{effect}} = 1^a$	-0.275 (-0.524, -0.016)
	Posterior probability $\delta_{\text{effect}} = 1$	0.726
d_N/d_S	BF _{effect}	2.6
	$\beta_{\text{effect}} \delta_{\text{effect}} = 1^a$	-0.265 (-0.523, 0.019)
	Posterior probability $\delta_{\text{effect}} = 1$	0.502
	BF _{effect}	1.0
	$\beta_{\text{effect}} \delta_{\text{effect}} = 1^a$	0.083 (-0.101, 0.25)

Edo-Matas et al., MBE, 2011

beast-users ›
Comparing evolutionary rates using a t-test?

7 posts by 3 authors



Joseph Hughes

Jul 16

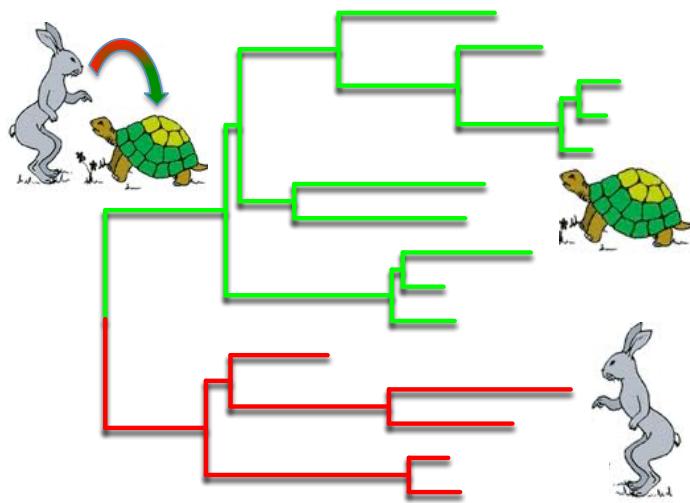


Hi all,

We have run beast analysis on a set of sequences of Feline Immunodeficiency Virus for a number of cats. Some cats are kept in nice conditions, others are in a cat home that could be considered "stressful". The sequences from each cat are monophyletic and I have estimated the evolutionary rate of FIV in each cat. BEAST was run estimating independent trees for each cat.
Can I use the estimated ucl.d.mean from each cat to compare the rates between the cats kept in good conditions versus those under stress?



What drives the tempo of pathogen evolution?



Pathogen factors

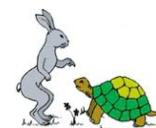


Mutation rate



Life cycle/replication dynamics

Host factors

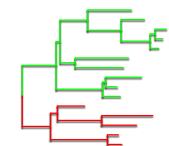


Life history

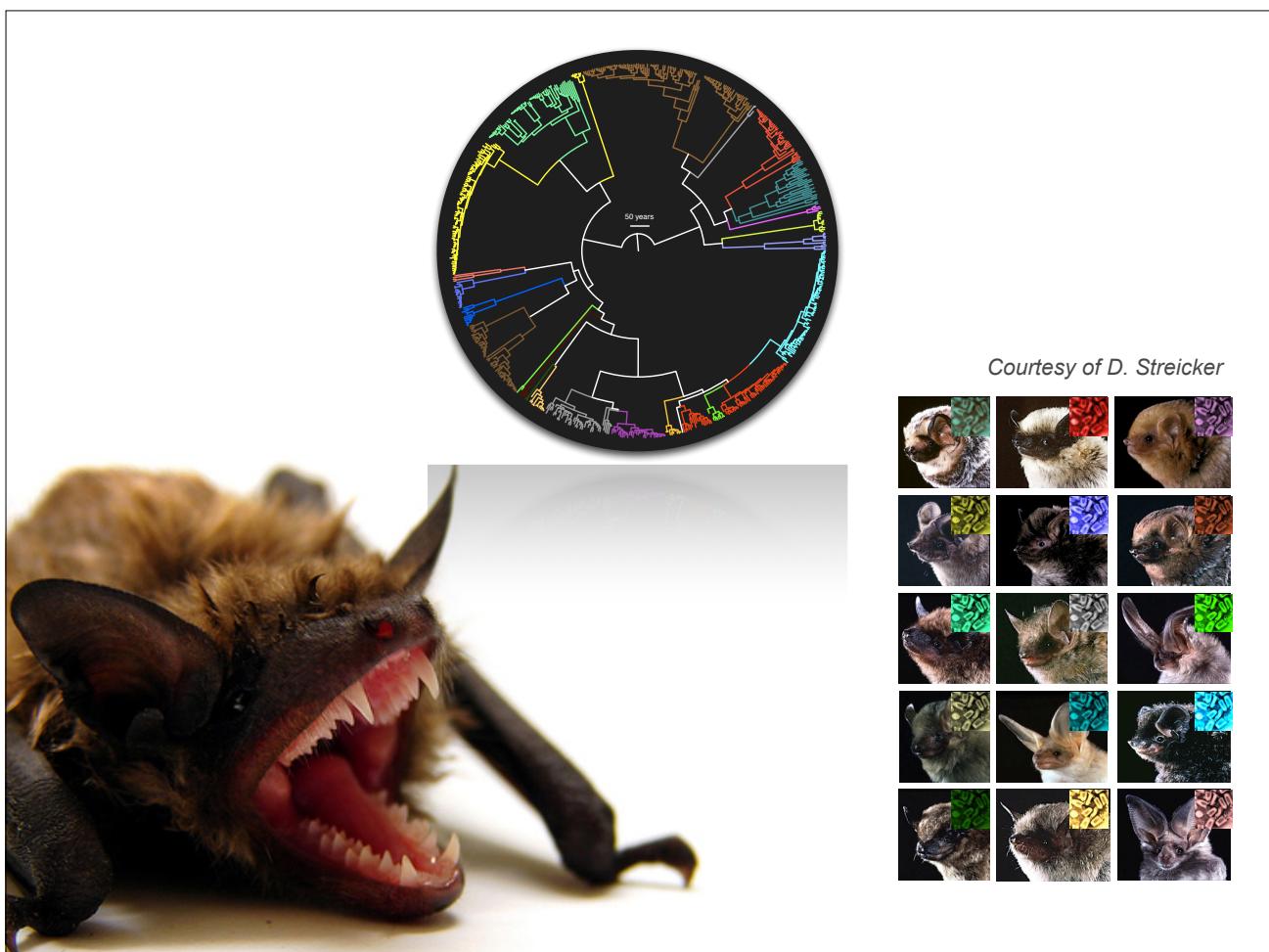
Seasonality

Metabolic rate etc.

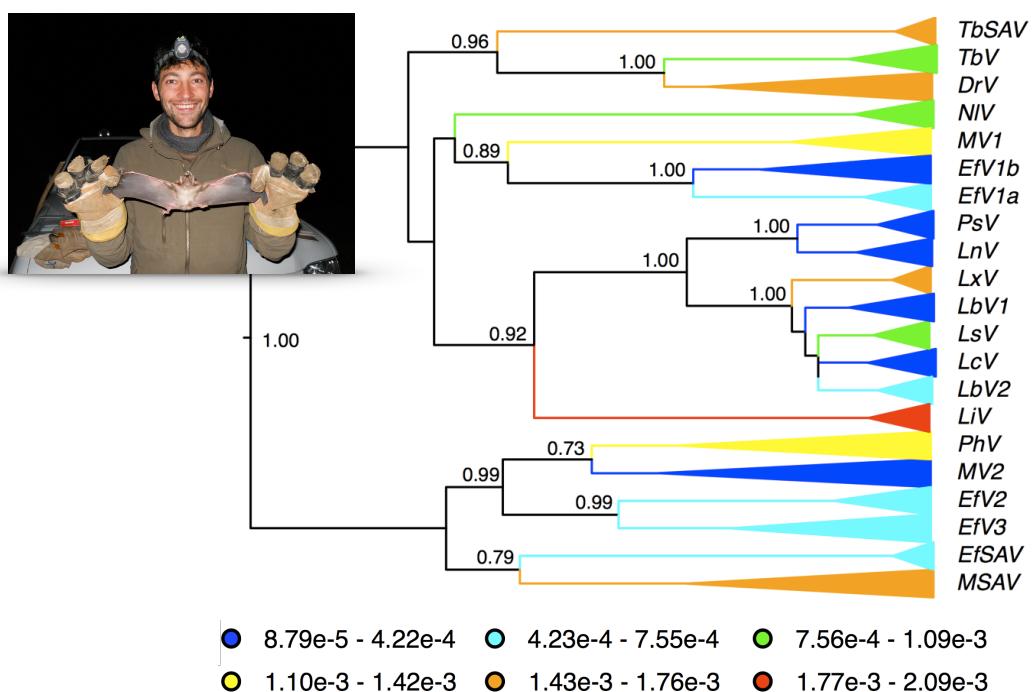
Historical factors



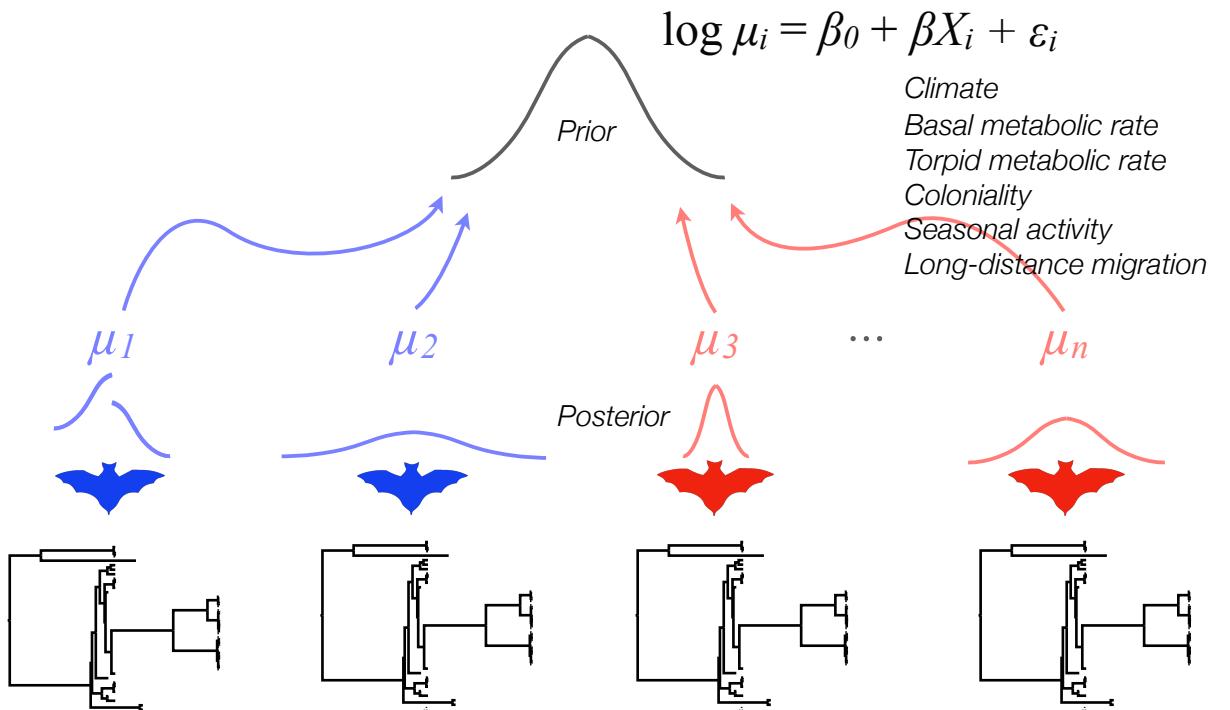
Pathogen phylogeny



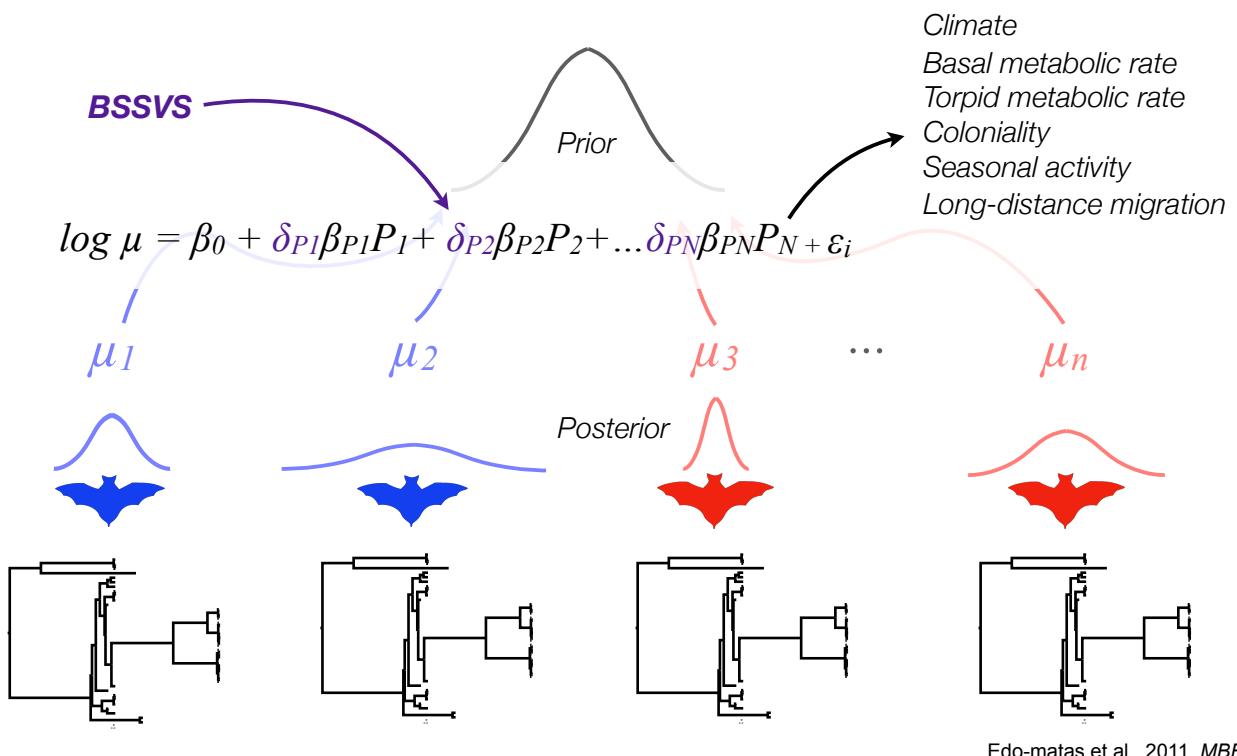
Bat rabies virus evolutionary rates



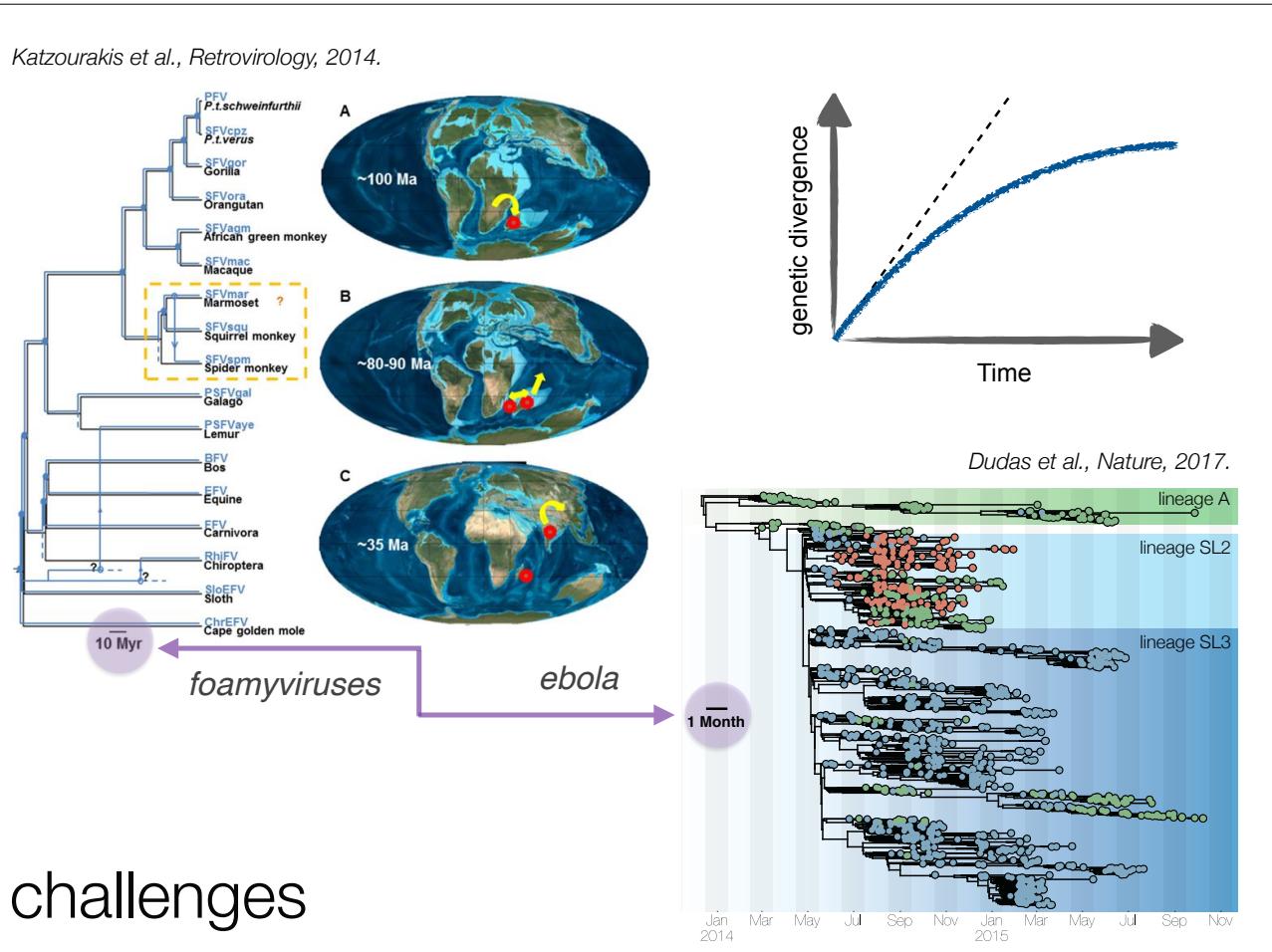
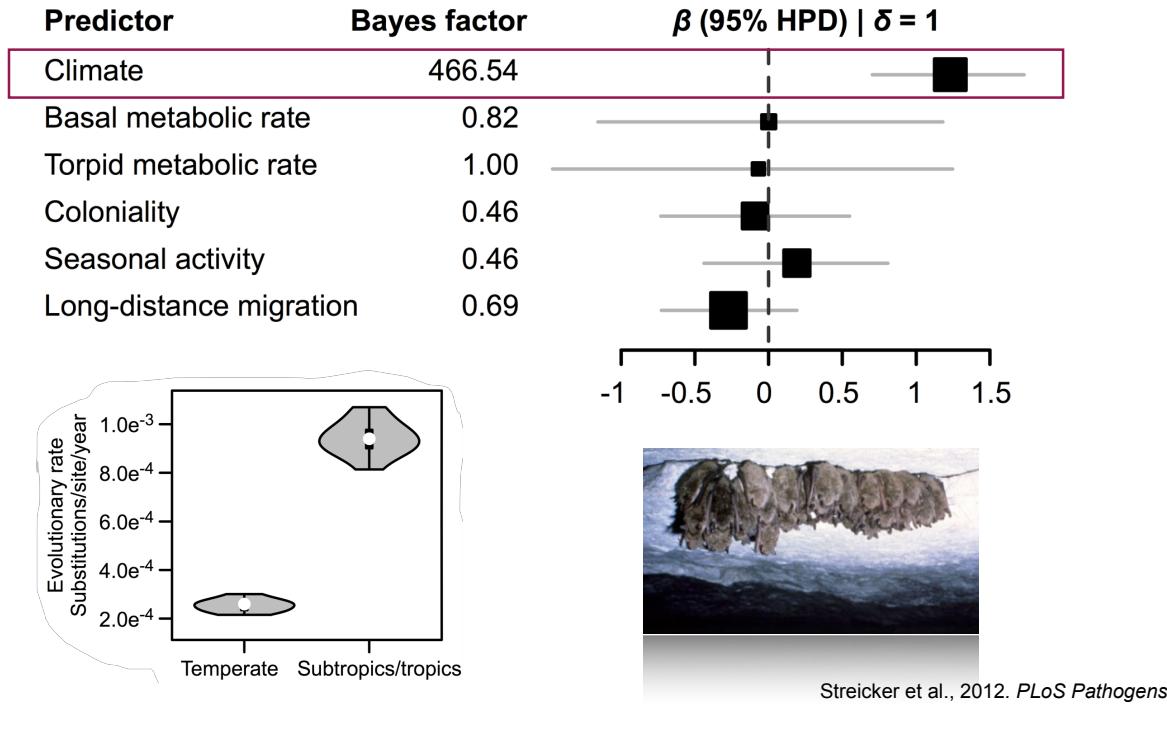
Fixed-effect hierarchical phylogenetic models



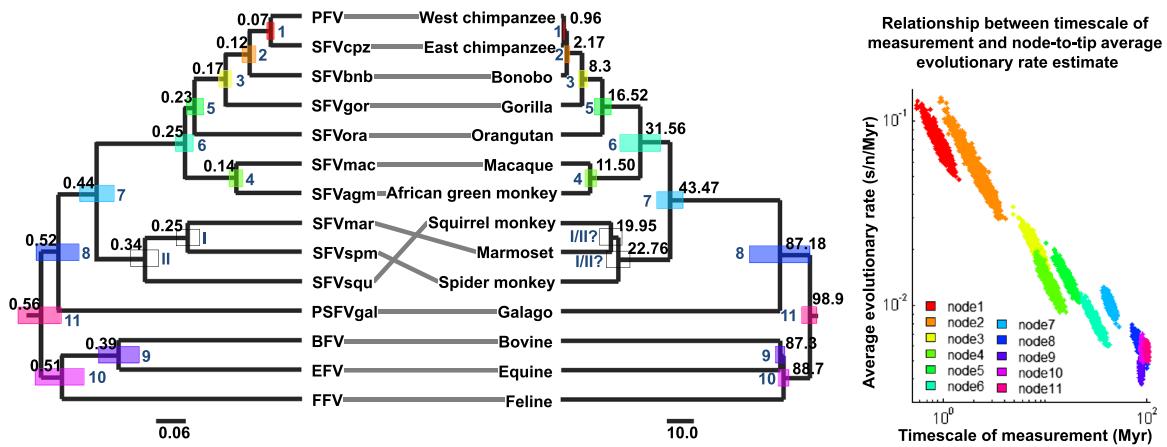
Fixed-effect hierarchical phylogenetic models



Bat rabies virus evolutionary rates

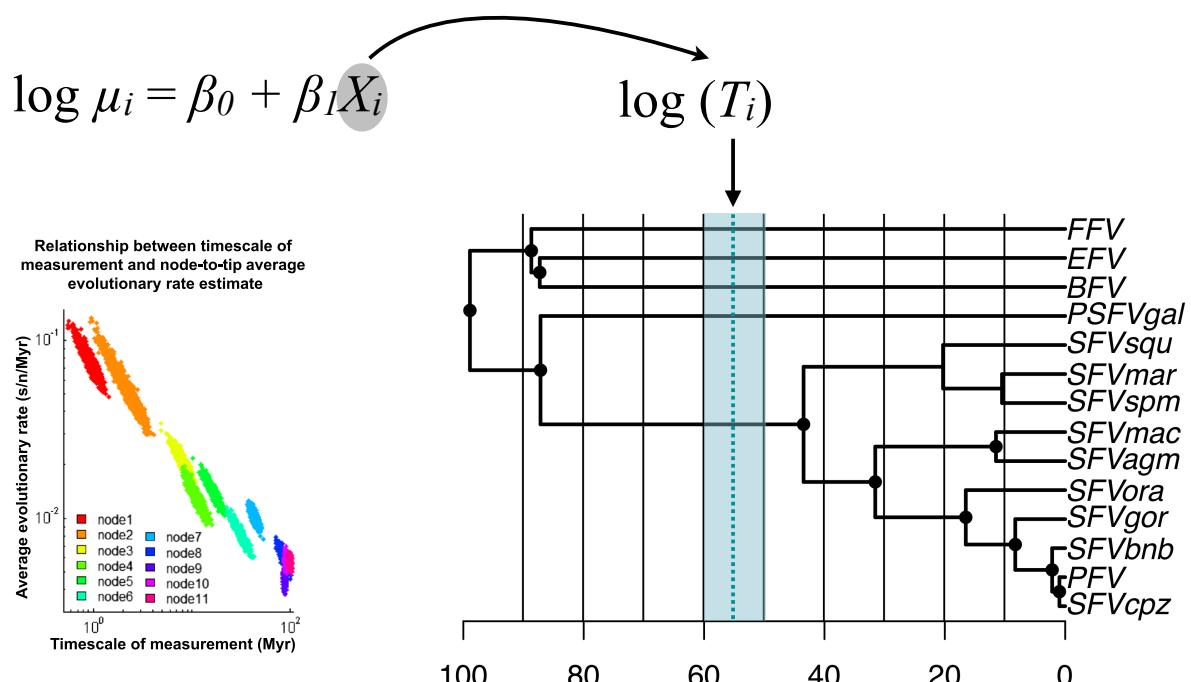


Time-dependent evolutionary rates

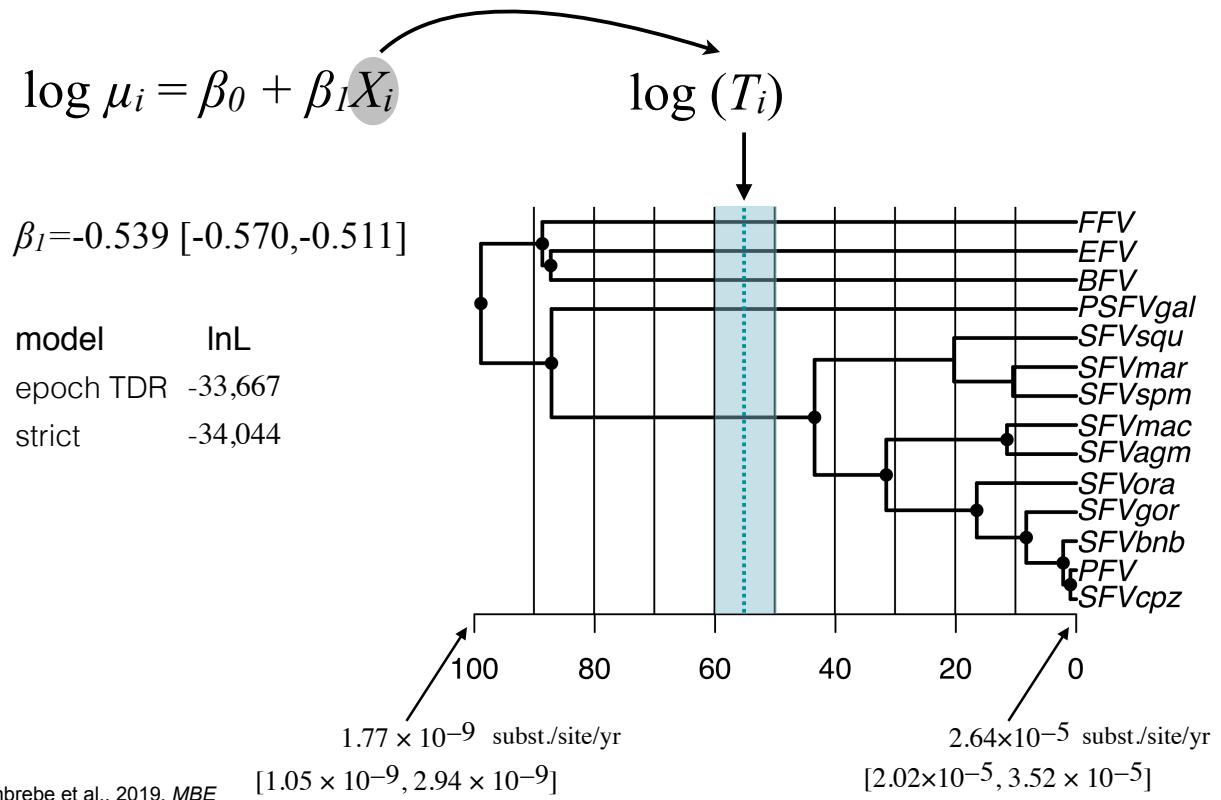


Aiewsakun et al., BMC Evol Biol, 2015.

epoch modelling with TDR



epoch modelling with TDR



conclusions

- molecular clocks: rate constancy assumption and tick rate calibration
- unconstrained \leftrightarrow strict molecular clock
- relaxed clocks
- model testing: use wisely
- hypotheses \rightarrow incorporate them into your model

