# GIS-FA: an approach to integrating thematic maps, factor-analytic, and envirotyping for cultivar targeting

Theor Appl Genet (2024). https://doi.org/10.1007/s00122-017-3033-y

Saulo F. S. Chaves, Mauricio S. Araujo, Matheus D. Krause, Kaio O. G. Dias

## Contents

Introduction	2
Section 1: Phenotypic data	3
Data description and visualization	3
Factor analytic (FA) mixed models	4
Selection of the best-fit FA model	4
$\mathrm{FA}(4)$	6
Factor Analytic Selection Tools	10
Section 2: Environmental data	12
Download and compilation	12
Obtaining environmental covariates for the tested environments	13
Obtaining environmental covariates for the untested environments	15
Data management	19
Environmental similarity	20
Section 3: Prediction via PLS	<b>2</b> 4
GIS-FA	24
GIS-GGE	27
GIS-FA vs GIS-GGE	29
Prediction in untested environments	29

Section 4: Visualization using maps	<b>30</b>
Interpolation	30
Which-won-where	31
Pairwise comparisons	32
Genotypes' performance map	34
Closing remarks	36
References	37

## Introduction

This tutorial describes each step for performing the GIS-FA analysis. The pipeline was divided into four stages: 1- Phenotypic data; 2- Environmental data; 3- Prediction via PLS; and 4- Visualization using maps. The packages we will need are listed below.

```
# For data management, and building and customizing plots:
library(tidyverse)
library(gghighlight)
library(ggrepel)
library(ggpubr)
library(ggspatial)
# For building Heatmaps:
## this is the only package that is not from CRAN,
## so a previous step may be necessary for installing it:
if(!require("BiocManager", quietly = TRUE)){
  install.packages("BiocManager")}
library(ComplexHeatmap)
library(circlize)
# For fitting the factor analytic mixed model and extracting its outputs straightforwadly
library(asreml)
source("https://raw.githubusercontent.com/saulo-chaves/May_b_useful/main/fa.outs.R")
# For using the partial least square method
require(pls)
# For managing rasters, shapefiles and other spatial data
library(raster)
library(sf)
library(spatstat)
library(geodata)
```



```
# For parallelization
library(parallel)
library(foreach)
```

We will also use the EnvRtype package. However, there seems to be a conflict between EnvRtype and asreml, so that asreml::asreml() does not work when EnvRtype is loaded.

We express our gratitude to the developers of all packages used in this pipeline. The credits are properly given at the References section.

We recommend the creation of folders in the working directory to store the necessary data to perform the analyses. In this pipeline, there are four subfolders: data stores the phenotypic data, rasters stores some rasters that will be used to extract environmental data, shapefiles stores the shapefiles from where we will obtain coordinates and delimit the prediction area, and saves stores some .RDA and .RDS files that will be built throughout the document.

## Section 1: Phenotypic data

## Data description and visualization

It is important to know your data. The dataset used as example refers to soybean VCU (value of cultivation and use) trials. It has 182 genotypes and 49 environments laid out in randomized complete blocks with three blocks. Let's load the dataset into R memory and take a look in its structure:

```
data = read.csv("data/data_soy.csv")
str(data)
```

```
## 'data.frame':
                   6456 obs. of 10 variables:
                 "E01" "E01" "E01" "E01" ...
   $ env
          : chr
                 "2019-11-21" "2019-11-21" "2019-11-21" "2019-11-21" ...
   $ plant: chr
                 "2020-03-20" "2020-03-20" "2020-03-20" "2020-03-20" ...
##
   $ harv : chr
   $ lat
          : num
                 -22.2 -22.2 -22.2 -22.2 ...
                 -52.7 -52.7 -52.7 -52.7 ...
   $ lon
          : num
##
   $ block: int
                 1 2 3 1 2 3 1 2 3 1 ...
##
   $ row
          : int
                 32 13 13 6 18 18 27 8 8 36 ...
                 1 2 3 1 2 3 1 2 3 1 ...
##
   $ col
          : int
   $ gen
          : chr
                 "G001" "G001" "G001" "G002" ...
   $ yield: num 5525 4380 4795 5860 5185 ...
```

The environments are distinguished in the column env (E01, E02, ..., E49). plant, harv, lat and lon contain the planting date, harvesting date, latitude and longitude of each environment, respectively. block, row and col provide the information of the position (block, row and column, respectively) of each genotype (gen, G001, G002, ..., G182). Finally, yield is the grain yield, the trait we are assessing. Note that env and gen are string vectors, and block, row and col are



numeric vectors. To fit the model, we have to transform these columns into factors, so the model can treat them as such:

After defining the factors, we store the number and name of genotypes and environments into objects. This will be useful latter:

```
num.env = nlevels(data$env)
num.gen = nlevels(data$gen)
name.env = levels(data$env)
name.gen = levels(data$gen)
```

## Factor analytic (FA) mixed models

This is the FA (Piepho, 1997; Smith et al., 2001) part of the GIS-FA method.

#### Selection of the best-fit FA model

The best-fit FA models are selected considering the proportion of genetic covariance explained by k factors. An ad-hoc threshold of 75% for the genetic covariance was adopted. We assessed the explicative power of the models by using the Average Semivariances Ratio (ASR) (Chaves, Alves, et al., 2023; Piepho, 2019):

$$ASR = \frac{\frac{2}{J(J-1)} \sum_{j=1}^{J-1} \sum_{j'=j+1}^{J} \frac{1}{2} (\sum_{k=1}^{K} \hat{\lambda}_{k_{j}}^{\star^{2}} d_{k} + \sum_{k=1}^{K} \hat{\lambda}_{k_{j'}}^{\star^{2}} d_{k}) - \sum_{k=1}^{K} \hat{\lambda}_{k_{j}}^{\star} \hat{\lambda}_{k_{j'}}^{\star} d_{k}}{\frac{2}{J(J-1)} \sum_{j=1}^{J-1} \sum_{j'=j+1}^{J} \frac{1}{2} [(\sum_{k=1}^{K} \lambda_{k_{j}}^{\star^{2}} d_{k} + \hat{\psi}_{j}) + (\sum_{k=1}^{K} \hat{\lambda}_{k_{j'}}^{\star^{2}} d_{k} + \hat{\psi}_{j'})] - \sum_{k=1}^{K} \hat{\lambda}_{k_{j}}^{\star} \hat{\lambda}_{k_{j'}}^{\star} d_{k}} \times 100$$

where  $d_k$  is the  $k^{th}$  element of the diagonal of  $\mathbf{D}$ , a  $K \times K$  symmetric positive (semi)-definite factor score variance matrix;  $\hat{\lambda}_{k_j}$  is the rotated loading of the  $\mathbf{j}^{th}$  ( $\mathbf{j} = 1, 2, ..., \mathbf{J}$ ) environment in the  $\mathbf{k}^{th}$  factor ( $\mathbf{k} = 1, 2, ..., \mathbf{K}$ ), and  $\hat{\psi}_j$  is the specific variance of the  $\mathbf{j}^{th}$  environment.

First, let's fit the models. You will notice that in the paper, we fitted a model adjusting spatial trends in multi-environment in a single step, and considered heterogeneous variances. Nevertheless, for the sake of simplicity, we opted to show a regular FA model, for the consideration of spatial trends in linear mixed models is not the aim of this tutorial. You can find the fitted model in this github page.



```
asreml.options(maxit = 100, workspace = '1gb', pworkspace = '2gb')
mod1 = asreml(fixed = yield ~ block:env + env,
             random = ~ gen:fa(env, 1),
              data = data,
              na.action = na.method(x="include", y = "include"), maxit = 50)
# FA2
mod2 = asreml(fixed = yield ~ block:env + env,
              random = ~ gen:fa(env, 2),
              data = data,
              na.action = na.method(x="include", y = "include"), maxit = 50)
# FA3
mod3 = asreml(fixed = yield ~ block:env + env,
             random = ~ gen:fa(env, 3),
              data = data,
              na.action = na.method(x="include", y = "include"), maxit = 100)
# FA4
mod4 = asreml(fixed = yield ~ block:env + env,
              random = ~ gen:fa(env, 4),
              data = data,
              na.action = na.method(x="include", y = "include"), maxit = 100)
```

Note that we changed some options before running the models. Usually, FA models need more iterations to converge. The memory usage will depend on the dataset.

The results showed below are from the model we fitted in the paper.

```
load(file = 'saves/FAMM_soy.RData')
```

Now, we are going to compute the objective criteria to select the best-fit FA model:

```
diagnos = do.call(rbind,lapply(fa.res, function(x) x$diagnostics))
rownames(diagnos) = c("mod1", "mod2", "mod3", "mod4"); diagnos[,'ASVR']
```

```
## mod1 mod2 mod3 mod4
## 22.578 54.083 69.084 76.627
```

mod4 with K=4 reach our criterion regarding explicative power, and therefore it is the best-fit model for further analyses. The remaining models can be removed from the workspace, so we can save some memory.

```
fa4 = fa.res[[4]]
rm(mod1, mod2, mod3, fa.models, fa.res); gc()
```

## **FA(4)**

The best-fit model mod4 was previously saved in an object called fa4. Let's take a look at the needed parameters for downstream analyses. All parameters presented in this section were computed using the fa.outs() function, available here.

• Rotated factor loadings:

asreml does not rotate factors loading by default unless you set rotate.fa = TRUE. We personally like to rotate them manually using the singular value decomposition (SVD) method, as suggested by Smith et al. (2021). See the main manuscript for more details about the rotation process.

#### head(fa4\$rot.loads)

```
##
              fa1
                           fa2
                                       fa3
                                                    fa4
## E01 0.19332304
                   0.080907355 -0.20263846
                                            0.071132347
## E02 0.06860044 -0.199689358 -0.20180917 -0.001911237
## E03 0.13790138 0.131566541 -0.09134606
                                            0.023812729
## E04 0.02115538
                   0.041711945 -0.12819687
                                            0.217929631
## E05 0.16426246
                   0.104474372 -0.02068189 -0.153304493
## E06 0.15198068 -0.003096557 -0.02703350 -0.160782802
```

• Matrix of genetic covariances between environments:

Given by:

$$\mathbf{\Sigma}_q = (\hat{\mathbf{\Lambda}}^{\star} \mathbf{D} \hat{\mathbf{\Lambda}}^{\star'} + \hat{\mathbf{\Psi}}) \otimes \mathbf{I}_V$$

where  $\Sigma_g$  is the matrix of genetic covariances (49 × 49 in our example),  $\hat{\Lambda}^*$  is the matrix of rotated loadings (49 × 4, in our example), **D** was previously defined, and  $\hat{\Psi}$  is a diagonal matrix (49 × 49) of specific variances.

#### fa4\$Gvcov[1:5, 1:5]

```
##
             E01
                       E02
                                  E03
                                            E04
                                                       E05
## E01 198687.34
                  59471.13 123241.82 43676.588 123971.962
        59471.13 151421.29
                            18238.38 14260.969
                                                 20868.986
## E03 123241.82
                  18238.38 114447.89 27316.533
                                                 98004.270
        43676.59
                  14260.97
                            27316.53 57039.095
## E04
                                                  5674.497
## E05 123971.96
                  20868.99
                            98004.27
                                      5674.497 128023.205
```



• Matrix of genetic correlations between environments:

Given by Cullis et al. (2010):

$$\mathbf{\Upsilon} = \mathbf{\Delta} \mathbf{\Sigma}_q \mathbf{\Delta}$$

where  $\Upsilon$  is the matrix of genetic correlations (49 × 49 in our example), and  $\Delta$  is a diagonal matrix whose elements are the inverse of the square roots of the diagonal elements of  $\Sigma_q$ .

## fa4\$Gcor[1:5, 1:5]

```
## E01 E02 E03 E04 E05
## E01 1.0000000 0.3428685 0.8172767 0.41027711 0.77730959
## E02 0.3428685 1.0000000 0.1385444 0.15345084 0.14988685
## E03 0.8172767 0.1385444 1.0000000 0.33809237 0.80964895
## E04 0.4102771 0.1534508 0.3380924 1.00000000 0.06640438
## E05 0.7773096 0.1498868 0.8096489 0.06640438 1.00000000
```

• Proportion of variance explained by each factor in each environment:

Given by Smith et al. (2015):

$$v_{k_j} = \frac{\hat{\lambda}_{k_j}^{\star^2} d_k}{\sum_{k=1}^{K} \hat{\lambda}_{k_i}^{\star^2} d_k + \hat{\psi}_j} \times 100$$

where  $v_{k_j}$  is the proportion of variance of the j<sup>th</sup> environment explained by the  $\mathbf{k}^{th}$  factor.

#### fa4\$expvar\_j[,1:5]

```
## E01 E02 E03 E04 E05

## fa1 68.714769 18.893226505 73.0689444 4.679451 81.5810877

## fa2 3.694773 49.146333016 20.4181042 5.584742 10.1312094

## fa3 14.756477 31.958782266 6.2665899 33.586458 0.2527839

## fa4 1.051904 0.001658213 0.2463615 56.149350 8.0349190
```

• Rotated scores:

In the function, we opted to show the scores in a matrix form, rather than a vector form, for better visualization.

#### head(fa4\$rot.scores)

```
##
              fa1
                          fa2
                                     fa3
                                               fa4
        -962.4416 1111.87158 -1038.4198
                                         287.9607
## C006
## C049 1018.8383 390.53328
                              -884.6681 -212.1031
## C054
        1894.1170 -444.45912
                              -867.8695
                                         477.0996
## C084 2024.6305 -25.40594
                               631.5421
                                         169.3687
## G001 -1239.7424 -823.04425 -226.0714
                                         234.6343
## G002
         741.7032 -395.93031 -372.9204
                                        117.0934
```



• Conditional (accounting for the specific environmental effects, and summed to the intercept) and marginal empirical BLUPs:

The conditional EBLUPs estimated by the function is given by the following equation:

$$\mathbf{g} = (\mathbf{\Lambda}^{\star} \otimes \mathbf{I}_{V}) \mathbf{\tilde{f}}^{\star} + \mathbf{\tilde{\delta}} + \mu$$

where  $\tilde{\mathbf{f}}^{\star}$  is the vector of rotated scores,  $\mathbf{I}_g$  is an identity matrix whose dimension is the number of genotypes,  $\boldsymbol{\delta}$  is the vector of specific effect (lack of fit of the multiple regression), and  $\mu$  is the intercept of the model. The marginal EBLUPs are estimated using only the common effects, i.e.,  $\mathbf{g} = (\boldsymbol{\Lambda}^{\star} \otimes \mathbf{I}_V) \tilde{\mathbf{f}}^{\star}$ .

## head(fa4\$blups)

```
gen conditional std.error
                                       marginal
##
                  5347.383
                            409.6906
                                       134.8035
  1 E01 C006
  2 E01 C049
                  5566.562
                            110.7297
                                       392.7423
## 3 E01 C054
                  5696.263
                            108.2500
                                       540.0174
## 4 E01 C084
                  5494.399
                            174.7449
                                       273.4251
## 5 E01 G001
                  4946.074
                            119.5839
                                     -243.7602
## 6 E01 G002
                  5449.179
                            108.2943
                                       195.2518
```

Generalized heritabilities

Given by Cullis et al. (2006):

$$H_j^2 = \frac{\overline{V}(\Delta)}{2 \times \sigma_{g_j}^2}$$

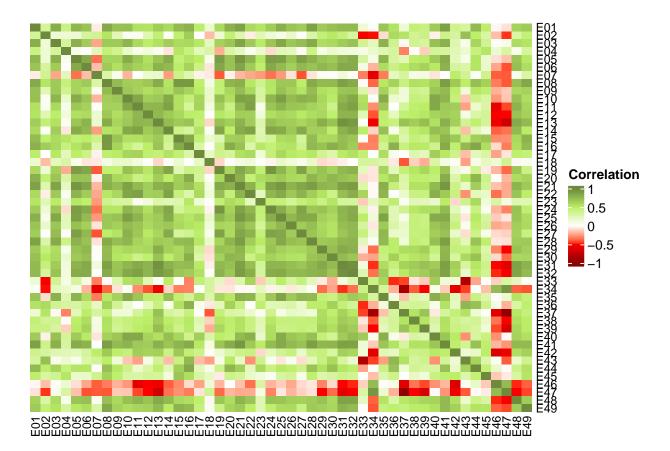
where  $H_j^2$  is the generalized heritability of the  $\mathbf{j}^{th}$  environment,  $\overline{V}(\Delta)$  is the mean pairwise prediction error variance, and  $\sigma_{g_j}^2$  is the genetic variance of the  $\mathbf{j}^{th}$  environment.

#### fa4\$H2

```
##
         E01
                    E02
                              E03
                                         E04
                                                    E05
                                                              E06
                                                                         E07
                                                                                    E08
## 0.6584196 0.4414755 0.6183688 0.3895006 0.6818437 0.6410480 0.4178423 0.7144013
         E09
                    E10
                                                    E13
                                                              E14
                              E11
                                         E12
                                                                         E15
## 0.5887136 0.7072477 0.7497581 0.5907105 0.7423759 0.6994666 0.6241437 0.6780266
##
         E17
                    E18
                                         E20
                                                    E21
                                                              E22
                                                                         E23
                                                                                    E24
                              E19
## 0.4124776 0.3154184 0.6548304 0.5778095 0.7506231 0.6829289 0.4064819 0.6510548
                                         E28
                                                    E29
##
         E25
                    E26
                              E27
                                                              E30
                                                                         E31
                                                                                    E32
## 0.7164063 0.5470520 0.6212530 0.6672332 0.6212460 0.5927928 0.7752457 0.7648092
         E33
                              E35
##
                    E34
                                         E36
                                                    E37
                                                              E38
                                                                         E39
                                                                                    E40
  0.6239515 0.6480647 0.6474860 0.5036621 0.6730913 0.4259277 0.5470619 0.6701653
         E41
                    E42
                              E43
                                         E44
                                                    E45
                                                              E46
                                                                         E47
                                                                                    E48
## 0.7218086 0.5268644 0.6278025 0.5878723 0.4188775 0.5678912 0.6744804 0.6809696
##
         E49
## 0.6000386
```



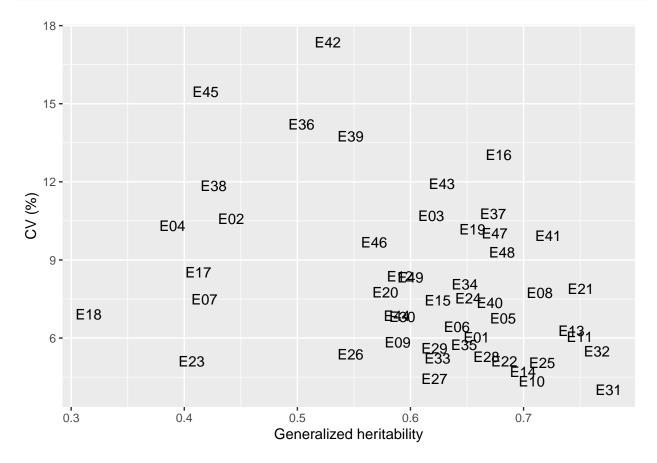
Plots can provide additional illustration of the results. For example, a heatmap of the genetic correlation between environment can be computed as follows:



Other useful information can be extracted from the model, such as the generalized heritabilities and the experimental coefficients of variation:



```
data.frame(
   H2 = fa4$H2,
   CV = (sqrt(summary(mod4)$varcomp[grep('!R', rownames(summary(mod4)$varcomp)),1])/
   tapply(data$yield, data$env, mean, na.rm = T)) * 100
) |>
   tibble::rownames_to_column("env") |>
   ggplot(aes(x = H2, y = CV)) +
   geom_text(aes(label = env)) +
   labs(x = "Generalized heritability", y = "CV (%)")
```



## Factor Analytic Selection Tools

We can now employ the Factor Analytic Selection Tools (FAST) to estimate overall performance (OP) (Smith & Cullis, 2018; Stefanova & Buirchell, 2010) and root mean square deviation (RMSD) (Smith & Cullis, 2018), which are a measure of performance and stability, respectively. We can associate these metrics with the reliability (Mrode, 2014) of each selection candidate and build a selection index to facilitate the selection process.

The overall performance is given by:

$$OP_i = \frac{1}{J} \sum_{j=1}^J \lambda_{1_j}^{\star} f_{1_j}^{\star}$$



where J is the number of environment.

The root mean square deviation is given by:

$$RMSD_{i} = \sqrt{\frac{1}{J} \sum_{j=1}^{J} (\lambda_{2_{j}}^{\star} f_{2_{j}}^{\star} + \lambda_{3_{j}}^{\star} f_{3_{j}}^{\star} + \lambda_{4_{j}}^{\star} f_{4_{j}}^{\star})^{2}}$$

The reliability is given by:

$$r = 1 - \frac{PEV_i}{\overline{\sigma_q^2}}$$

where  $PEV_i$  is the prediction error variance of the  $i^{th}$  genotype, and  $\overline{\sigma_g^2}$  is the genetic variance across environments.

When building the index, you can prioritize what you think is more important for reaching your goals (performance over stability, or the other way around). For example, if we define performance is two times more important than stability, the index will have a weight of 2 for OP, and a weight of 1 for RMSD:

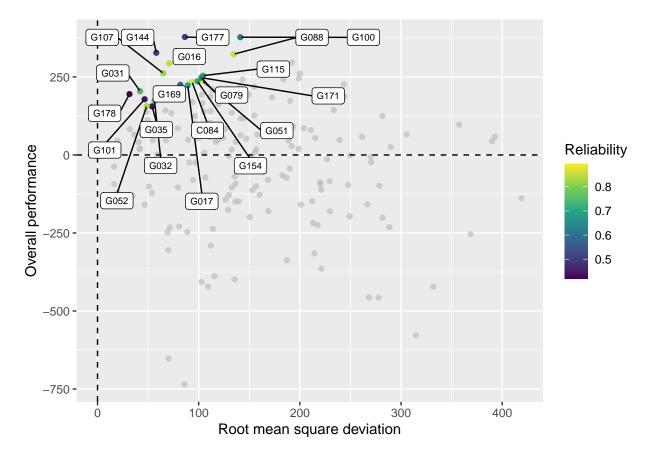
$$\mathtt{Index}_i = \left\lceil 2 \times \frac{OP_i - \overline{OP}}{\sqrt{V(OP)}} \right\rceil - \frac{RMSD_i - \overline{RMSD}}{\sqrt{V(RMSD)}}$$

The R code is presented below:

Now that we built our data frame, we can make the FAST plot. Here, we will highlight only the top 20 (11%) selection candidates:

```
FAST |> ggplot(aes(x = RMSD, y = OP)) +
  geom_point(aes(color = rel)) +
  geom_vline(xintercept = 0, linetype = 'dashed') +
  geom_hline(yintercept = 0, linetype = "dashed") +
```





Do not forget to check if most of the loadings in the first factor are positive. If this is not the case, FAST is not applicable. Smith et al. (2021) and Chaves, Evangelista, et al. (2023) provided some options for when this happens.

We will save the fa4 list for later. The other objects will be cleared from the memory:

```
save(fa4, file = "saves/FA4.RDA")
rm(list=ls())
```

## Section 2: Environmental data

## Download and compilation

There are several on-line platforms that have freely available environmental data such as soilgrids, nasapower, and GAEZ. We will only show two ways of downloading environmental data using R.

Also, keep in mind there are two separate procedures from now on: obtaining environmental data for the tested environments, and obtaining the environmental data for the geographical coordinates (i.e., points) at which you want to do the prediction (untested environments).

## Obtaining environmental covariates for the tested environments

First, let's load again the phenotypic data:

```
data = read.csv("data/data_soy.csv")
```

The function get\_weather() from the package EnvRtype downloads daily environmental data from the nasapower database according to the dates of planting and harvesting. After obtaining the weather data, we can use other functions such as param\_atmospheric(), param\_temperature(), and param\_radiation() to process it into derived environmental covariates. The arguments Tbase1 and Tbase2 in param\_temperature() below refers to the minimum and maximum cardinal temperature, respectively. The parameters Topt1 and Topt2 are the lower and upper bound temperature for phenological development, respectively. These values are species-dependent. See Costa-Neto et al. (2021) for more details.

Note we are overwriting the object covamb when using the processing functions. This will do no harm since these functions merge (merge = T) the columns with the new, derived environmental covariates into the original data.

We can now obtain soil-related environmental covariates from the soilgrids platform using the function soil\_word() from geodata. The data will be downloaded as rasters (.tif) in a path that must be specified (in our case, 'Rasters').



```
geodata::soil_world(var = 'bdod', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'clay', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'nitrogen', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'phh2o', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'sand', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'silt', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'soc', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'cfvo', depth = 15, stat = 'mean', path = 'Rasters')
geodata::soil_world(var = 'cfvo', depth = 15, stat = 'mean', path = 'Rasters')
```

Data from elevation (altitude) can be obtained from geodata:

```
geodata::elevation_30s(country = 'BRA', path = 'Rasters')
```

Note that you just have to download the rasters once. We also manually downloaded rasters from the platform of Lembrechts et al. (2022) regarding the soil's mean temperature, isothermality, temperature seasonality and mean diurnal range. In the platform, there is a raster of soil's mean temperature for each month of the year. Nevertheless, since the trials were conducted between October and March (Brazil's soybean first cropping season), we only downloaded raster corresponding to months between this range.

We can now use the extract\_GIS() from EnvRtype to extract the data for the corresponding coordinates:

A last step is to take the mean of the soil temperature across the months in which the trial was conducted:



Then, we just have to merge the data frame that contain climatic the environmental covariates with the one that has the soil environmental covariates:

```
covamb = left_join(covamb, soil[,-(2:5)], by = c('env'))
```

For saving memory, we save the covamb object as a .RDS data in our "saves" subdirectory and clear the internal memory:

```
saveRDS(covamb, file = "saves/covamb.RDS")
rm(list=ls())
```

## Obtaining environmental covariates for the untested environments

Up to this point, we only obtained environmental data for the observed target population of environments (TPE). This section will obtain the environmental data needed to predict the performance of the selection candidates in untested environments from several geographic coordinates.

The first step is to obtain the coordinates, which can be done from shapefiles freely available from several platforms. For example, in Brazil, the Institute of Geography and Statistics (IBGE in the Portuguese acronym) provides shapefiles with several details regarding administrative boundaries (from municipalities to the whole country). These are the shapefiles we used for obtaining the coordinates. Another useful way of obtaining shapefiles is using the function gadm(), from the geodata package.

It is important to consider how you want to sample the breeding region. Here, we opted to sample coordinates within each municipality to represent the whole area well. Thus, we have to extract the coordinates of each municipality. So, first, the shapefile with the administrative boundaries of each municipality must be loaded as follows:

```
br_mun = raster::shapefile('shapefiles/municipality/BR_Municipios_2021.shp')
br_mun.sf = sf::st_as_sf(br_mun)
```

The object br\_mun.sf will be used soon. Let's give an example first. Suppose you are interested in predicting the phenotypic performance in the city of Belém, Pará State. The steps would be the following:

1- Extract the coordinates of the whole country using coordinates() from the raster package. You can do this using the elevation raster, for example.

```
br_coord = as.data.frame(
  raster::coordinates(raster::raster('rasters/BRA_elv_msk.tif'))
)
```

2- Filter the shapefiles to the appropriate State (since we obtained the shapefiles from a Brazilian source, the words are in Portuguese).



```
mun = br_mun[br_mun$SIGLA == 'PA', ]
mun.sf = br_mun.sf[br_mun.sf$SIGLA == "PA", ]
```

3- Filter to the desired municipality and determine a boundary box based on the municipality's boundaries. This boundary box is filled with coordinates, but not all of them correspond to the desired municipality.

4- Extract the geometry of the desired municipality.

```
pts = unique(do.call(rbind, mun.sf[mun.sf$NM_MUN == "Belém",]$geometry[[1]][[1]]))
pts = as.data.frame(pts)
colnames(pts) = c('x','y')
```

5- Using the point.in.polygon() of the sp package, determine which points in the boundary box are within the municipality's geometry.

```
coord$filt = sp::point.in.polygon(coord$x, coord$y, pts$x, pts$y)
```

6- Exclude the points that are not within the municipality

```
coord = coord[which(coord$filt == 1),]
coord = subset(coord, select = -filt)
```

And voilà:

#### head(coord)

```
## x y
## 19 -48.47083 -1.020833
## 57 -48.47917 -1.029167
## 58 -48.47083 -1.029167
## 59 -48.46250 -1.029167
## 60 -48.45417 -1.029167
## 61 -48.44583 -1.029167
```

We have done the same procedure for all municipalities of Mato Grosso do Sul ('MS') State - which is the State where the soybean trials were conducted - using a for loop. The resulting data frame (mun.xy) has all coordinates of these municipalities.



```
mun = br_mun[br_mun$SIGLA == 'MS',]
mun.sf = br_mun.sf[br_mun.sf$SIGLA == "MS", ]
mun coords = list()
for (i in mun$NM_MUN) {
    coord = dplyr::filter(br coord, x > mun[mun$NM MUN == i,]@bbox['x','min'] &
                            x < mun[mun$NM_MUN == i,]@bbox['x','max'] &</pre>
                            y > mun[mun$NM_MUN == i,]@bbox['y','min'] &
                            y < mun[mun$NM_MUN == i,]@bbox['y','max'])</pre>
    pts = unique(do.call(rbind, mun.sf[mun.sf$NM_MUN == i,]$geometry[[1]][[1]]))
    pts = as.data.frame(pts)
    colnames(pts) = c('x', 'y')
    coord$filt = sp::point.in.polygon(coord$x, coord$y, pts$x, pts$y)
    if(all(coord$filt == 0)){
      cat('I was at', i, '-', 'MS', ', but found no points \n')
     next
    coord = coord[which(coord$filt == 1),]
    coord = subset(coord, select = -filt)
    coord$mun = i
    coord$est = 'MS'
    mun_coords[[i]] = coord
   rm(coord, pts)
    cat('I am at', i, '-', 'MS', '\n')
mun.xy = do.call(rbind, mun_coords)
colnames(mun.xy) = c('lon', 'lat', 'mun', 'sta')
rm(mun, mun_coords)
```

The object mun.xy contains the geographical coordinate of 79 cities in the Mato Grosso do Sul State.

After obtaining the geographical coordinates, we will randomly sample 50 points from each municipality to perform the predictions.

```
set.seed(100) #This seed guarantees the reproducibility of the results
samp = mun.xy %>% dplyr::group_by(mun) %>% dplyr::slice_sample(n = 50)
samp$env = paste0("P", 1:nrow(samp))
```

The next step is to download the environmental covariates. We had modified the original functions from the EnvRtype package to download monthly instead of daily averages. Thus, before extracting the climatic covariates, let's load the modified functions into R's memory:

```
source("https://raw.githubusercontent.com/saulo-chaves/May_b_useful/main/get_weather2")
```

Now, we can extract and process the covariates:



```
covamb_pred = get_weather2(env.id = samp$env[1:50], lat = samp$lat[1:50],
                           lon = samp$lon[1:50], start.year = '2000',
                           end.year = '2021', parallel = T)
covamb_pred = EnvRtype::extract_GIS(covraster = raster('rasters/BRA_elv_msk.tif'),
                                    Latitude = 'LAT', name.out = 'ALT',
                                    Longitude = 'LON', env.data = covamb_pred,
                                    env.id = 'env')
Tfun = c(8, 45, 30, 35)
names(Tfun) = c('Tbase1', 'Tbase2', 'Topt1', 'Topt2')
covamb_pred = covamb_pred |> dplyr::select(env, ALT, LON, LAT, PARAMETER, YEAR,
                                           OCT, NOV, DEC, JAN, FEB, MAR) |>
 tidyr::pivot_longer(cols = OCT:MAR, names_to = 'MONTH') |>
 tidyr::pivot_wider(names_from = PARAMETER, values_from = value) |>
 dplyr::mutate(
   T2RANGE = T2M MAX - T2M MIN
  ) |>
 dplyr::mutate(
   FRUE = F.RUE.Temperature(Tbase1 = Tfun['Tbase1'], Tmed = T2M,
                             Tbase2 = Tfun['Tbase2'], Topt1 = Tfun['Topt1'],
                             Topt2 = Tfun['Topt2']),
   GDD = ((adjust_for_Tbase2(adjust_for_Tbase(T2M_MAX, Tfun['Tbase1']),
                              Tfun['Tbase2']) +
              adjust_for_Tbase2(adjust_for_Tbase(T2M_MIN, Tfun['Tbase1']),
                                Tfun['Tbase2']))/2 -
             Tfun['Tbase1']),
   VPD = ((teten(T2M_MIN) + teten(T2M_MAX)/2)) - teten(T2MDEW),
   SPV = 4098 * (0.6108 * exp((17.27 * T2M)/(T2M + 237.3)))/(T2M + 237.2)^2
 ) |>
 dplyr::mutate(
   ETP = EToPT(alfa = 1.26, Srad = ALLSKY_SFC_SW_DWN, slope = SPV,
               psyc = psyco(AtmP(elevation = ALT)))
 ) |>
 dplyr::mutate(PETP = PRECTOTCORR - ETP)
```

To extract the soil covariates, we will use the same procedure previously described, as follows:



And finally, we merge the data frames with weather and soil environmental covariates:

```
covamb_pred = merge(covamb_pred, soil)

covamb_pred = covamb_pred |> dplyr::select(-lon, -lat) |>
  dplyr::relocate(sta, mun, .before = env) |>
  dplyr::relocate(YEAR, MONTH, ALT, LON, LAT, .before = T2M)
```

The previous step were computationally demanding and required several downloads. From time to time, we recommend saving the R environment [e.g., using save.image("GIS-FA.RDA)] so the pipeline can be continued at any point without the need to compute/download all the information again.

We will save the covamb\_pred object as a .RDS file and clean the memory for the next step:

```
saveRDS(covamb_pred, file = "saves/covamb_pred.RDS")
rm(list=ls())
```

### Data management

Now that we have all environmental covariates downloaded for the observed and unobserved environments, let's compute some important results. First, let's load the .RDS files we saved from the previous step:

```
covamb = readRDS("saves/covamb.RDS")
covamb_pred = readRDS("saves/covamb_pred.RDS")
```

Now, we will build matrices that contain the mean of each environmental covariate (columns) in each environment (rows):



```
matcovamb = tibble::column_to_rownames(matcovamb, var = 'env')

# Unobserved environments

covamb_pred = covamb_pred |>
    dplyr::filter(YEAR != 2000 | !MONTH %in% c('JAN','FEB','MAR')) |>
    dplyr::filter(YEAR != 2021 | !MONTH %in% c('OCT','NOV','DEC')) |>
    dplyr::select(-YEAR, -MONTH) |>
    dplyr::reframe(across(ALT:soiltemp, mean), .by = c(mun, env)) |>
    dplyr::rename(T2M_RANGE = T2RANGE, PRECTOT = PRECTOTCORR)
```

## Environmental similarity

The prediction will surely be more accurate in the environments where trials were conducted since we are able to observe the relationship between selection candidates and environmental features. Nevertheless, we cannot be sure about how reliable the prediction will be for the untested environments. A metric that can give us a hint about this reliability is the environmental similarity. If a given unobserved environment has similar environmental conditions to any observed environment, it is likely that this environment's prediction will be more accurate.

To calculate the environmental similarity, we must first sort the columns of the matrices that contain the environmental information for the observed and unobserved environments, in a way that they follow the same order:

```
# Observed environments
matcovamb = matcovamb[, sort(colnames(matcovamb))]

# Yet-to-be-seen environments (sampled points)
matcovamb_pred = covamb_pred |>
    tibble::column_to_rownames('env') |>
    dplyr::select(-mun)
matcovamb_pred = matcovamb_pred[, sort(colnames(matcovamb_pred))]
matcovamb_pred = as.matrix(matcovamb_pred)

# Let's confirm both matrices have the same order of columns:
all(tolower(colnames(matcovamb_pred)) == tolower(colnames(matcovamb)))
```

#### ## [1] TRUE

Then, the next step is calculating the Euclidean distances between the observed and unobserved environments to quantify the environmental similarity. Let  $\mathbf{W}$  be a  $J \times P$  matrix that contains the scaled values of P environmental covariates with the J observed environments, and let  $\Omega$  be the matrix with the same information but for the U unobserved environments. The Euclidean distance between an observed environment j and an unobserved environment u is given by the distances between the rows of  $\mathbf{W}$  and  $\Omega$  that correspond to j and u, respectively:

$$D_{ju} = \sqrt{\sum_{p=1}^{P} (w_{jp} - \omega_{up})^2}$$



where  $w_{jp}$  is an entry of **W** that contains the value of the  $p^{th}$  environmental covariate for the  $j^{th}$  environment, and  $\omega_{up}$  is the same for the unobserved environments.

```
eucl = as.matrix(pdist::pdist(scale(matcovamb), scale(matcovamb_pred)))
rownames(eucl) = rownames(matcovamb)
colnames(eucl) = rownames(matcovamb_pred)
```

In the next step, we transform the distance matrix to a long format and divide it into data frames stored in a list. In total, there will be 49 data frames, each one representing an observed environment from the target population of environments (TPE).

Each data frame has 3,891 sampled prediction points. But recall that we want to predict the genotypes' performance to the whole state. In other words, we have to expand the results of each sampled point by doing an interpolation process. Here, we used the inverse distance weighted (IDW) interpolation method. IDW interpolation operates on the principle that objects nearby exhibit greater similarity than those farther apart. When estimating the value for an unobserved environment, IDW considers the measured values surrounding that environment (i.e., observation window). The measured values closest to the prediction environment carry more weight in determining the predicted value than those farther away. The weights are defined using the Power function (see below). For example, the Euclidean distance between the trial network and a given point (say,  $u^*$ ) is given by:

$$D_{u^*j} = \frac{\sum_{u=1}^{U} \frac{1}{\|u^* - x_u\|^{\tau}} D_{uj}}{\sum_{u=1}^{U} \frac{1}{\|u^* - x_u\|^{\tau}}}$$

where  $x_u$  refers to a given sampled point u in the observation window, and  $\tau$  is a power defined by cross-validation (see below).

For interpolating, we must first define an observation window. This window is defined by a bounding box, that corresponds to the administrative limits of the state. We will extract these limits from a shapefile downloaded at IBGE that has the administrative limits for all Brazilian states, as follows:

```
br_est = raster::shapefile("shapefiles/states/BR_UF_2021.shp")
est_sf = sf::st_as_sf(br_est)

# Selecting only the state of Mato Grosso do Sul
MS = est_sf[est_sf$SIGLA == "MS",]
MS_pol = br_est[br_est$SIGLA == 'MS',]
```



```
# Building the observation window
obs_window = spatstat.geom::owin(xrange = MS_pol@bbox[1,], yrange = MS_pol@bbox[2,])
```

After defining the observation window, we have to perform an iterative process to define an optimum power to define the weights for the IDW interpolation. We define the optimum power as the one that provides the lower mean square prediction error (MSPE). After defining the optimum power, we perform the interpolation using this value and create a mask to overlap the area of the state. This process is shown below using parallel processing to speed up computational time.

```
cl = parallel::makeCluster(parallel::detectCores(logical = F)-1)
parallel::clusterEvalQ(cl, {library(spatstat)
                             library(raster)
                             library(sf)})
parallel::clusterExport(cl, varlist = c('obs_window', 'MS'))
dist.list = parallel::parLapply(cl = cl, dist.env, function(x){
 ppp_test = spatstat.geom::ppp(x = x$LON, y = x$LAT,
                 marks = x$dist, window = obs_window)
 powers \leftarrow seq(0.1, 5, 0.1)
 mse_result <- NULL</pre>
 for(power in powers){
    CV_idw <- spatstat.explore::idw(ppp_test, power=power, at="points")</pre>
    mse_result <- c(mse_result, mean((ppp_test$marks-CV_idw)^2, na.rm=T))</pre>
  }
  optimal_power <- powers[which.min(mse_result)]</pre>
 tr_idw = spatstat.explore::idw(ppp_test, power = optimal_power, at = 'pixels')
  idw_raster = as(raster::mask(raster::raster(tr_idw, crs = crs(br_est)), MS),
                   "SpatialPixelsDataFrame")
  idw_raster.df = as.data.frame(idw_raster)
})
parallel::stopCluster(cl)
rm(cl)
```

This is a computationally intensive process, so we will save the outcome for avoiding doing it again:

```
save(dist.list, file = 'saves/dist_list.RData')
```

By doing that, we can always reach dist.list using the following command:

```
load(file = 'saves/dist_list.RData')
```



Once we performed the interpolation for all list elements (i.e., for the observed 49 environments), the next step is binding the lists by columns and selecting the minimum value of each row. In other words, the environmental similarity between the observed and unobserved environments ( $S_u$  and  $S_u^*$ ) is given by the minimum distance between the points to be predicted and the TPE:

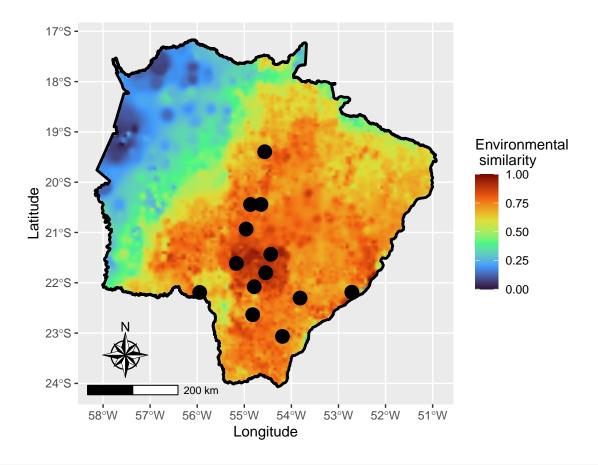
```
S_u = min(D_{uj}) S_{u^*} = min(D_{u^*j}) mindist = apply(do.call(cbind, lapply(dist.list, function(x) x[,-c(2,3)])),
```

Finally, we can build a map for illustrating how reliable the prediction will be across the whole state:

1, function(x) x[which.min(x)])

```
cbind(dist.list$E01[,-1],
      layer = abs((mindist - min(mindist))/(max(mindist) - min(mindist))-1)) |>
  ggplot() +
  ggspatial::annotation_scale(location = "bl", width_hint = 0.3) +
  ggspatial::annotation_north_arrow(location = "bl", which_north = "true",
                         pad_x = unit(0.5, "cm"), pad_y = unit(0.7, "cm"),
                         style = ggspatial::north_arrow_nautical())+
  geom_raster(aes(x = x, y = y , fill = layer), interpolate = T) +
  geom_sf(data = MS, fill = NA, linewidth = 1, color = "black") +
  labs(x = 'Longitude', y = 'Latitude', fill = 'Environmental \n similarity') +
  scale fill viridis c(option = 'turbo', direction = 1,
                       limits = c(0, 1) +
  theme(legend.position = 'right') +
  geom_point(
    data = data.frame("lat" = c(-22.187, -22.191, -21.434, -22.634,
                                 -20.442, -22.078, -22.304, -21.614,
                                -23.065, -21.801, -19.395, -20.931, -20.443),
                      "lon" = c(-52.717, -55.947, -54.438, -54.822, -54.646,
                                 -54.789, -53.815, -55.168, -54.190, -54.546,
                                 -54.566, -54.961, -54.868)),
    aes(x = lon, y = lat, size = 1.2), show.legend = F, color = 'black',
  )
```





rm(dist.list, dist.env)

## Section 3: Prediction via PLS

After obtaining the environmental data and the factor loadings, it is time to predict the expected performance of selection candidates in untested environments using partial least squares (PLS). Initially, we will train a model with environmental features to predict the loadings for all environments (i.e., observed and unobserved). With the predicted loadings and the previously estimated scores for each genotype from the FA model, we can predict the empirical BLUPs of the genotypes in untested environments.

## **GIS-FA**

The first step is to fit the PLS using the function plsr(), of the pls package. Initially, we must define the optimum number of components for prediction using leave-one-out cross-validation. In this scheme, we delete the data of one environment and predict it using the PLS model. Then, we compute the correlation between the predicted and observed values, which indicates the model's accuracy. Naturally, we will choose the model with a number of components that provides the highest prediction accuracy.

Let's load into R memory the object we used to save the outputs of the best-fit FA model we previously selected:

24

```
load('saves/FA4.RDA')
```

The computations in R work as follows. The object data.pls.lamb below is a data frame with the left-hand side composed of dependent variables, i.e., the matrix of rotated loadings, and the right-hand side composed of predictors, i.e. ,the matrix of environmental covariates in the tested environments of the PLS model. Note that we are scaling the matrix of predictors.

Then, an initial PLS model is fitted to find out the maximum number of components that can be used for prediction:

```
pls1.lamb = pls::plsr(lambda ~ CovAmb, validation = 'CV', data = data.pls.lamb)
```

This information will be used in the cross-validation process, where we will defined the best PLS model based on the prediction accuracy:

```
ncomp = list()
for (j in 1:pls1.lamb$ncomp) {
  acc_eblup = vector()
  acc_lamb = vector()
  for (i in rownames(matcovamb)) {
    train = data.pls.lamb[rownames(data.pls.lamb) != i,]
    test = data.pls.lamb[rownames(data.pls.lamb) == i,]
    cvloo = pls::plsr(lambda ~ CovAmb, validation = 'CV', data = train,
                 ncomp = j)
    acc_lamb = c(
      acc_lamb,
      cor(
        t(test$CovAmb %*% coef(cvloo, intercept = T)[-1,,1] +
            coef(cvloo, intercept = T)[1,,1]),
        t(test$lambda)
      )
    )
    acc_eblup = c(
      acc_eblup,
      cor(
        t(tcrossprod(
          test$CovAmb %*% coef(cvloo, intercept = T)[-1,,1] +
            coef(cvloo, intercept = T)[1,,1],
          fa4$rot.scores
        )),
```



```
fa4$blups[fa4$blups$env == i, 'marginal'], method = 'spearman'
      )
    )
    rm(cvloo, train, test)
 ncomp[[j]] = data.frame(ncomp = j, lambda = mean(acc_lamb),
                          eblup = mean(acc_eblup))
}
cv.results = do.call(rbind, ncomp)
cv.results[which.max(cv.results$lambda),] # Best results for the prediction of loadings
              lambda
##
     ncomp
                         eblup
## 8
         8 0.5722463 0.7399924
cv.results[which.max(cv.results$eblup),] # Best results for the prediction of eBLUPs
##
    ncomp
              lambda
                         eblup
## 3
         3 0.5603842 0.7436089
```

Thus, the PLS model with 3 components had the best outcomes. We will now fit the proper prediction model, as follows:

This model explains 61.57% of the total variance.

Using the same model, we can investigate its prediction ability within environments using eBLUEs. Let's first fit a model per environment considering fixed genetic and block main effects:



```
data = x)
predict(mod, classify = 'gen')$pvals[,1:2]
})
```

Now, we can perform the cross-validation:

```
acc_eblue = vector()
for (i in rownames(matcovamb)) {
 train = data.pls.lamb[rownames(data.pls.lamb) != i,]
  test = data.pls.lamb[rownames(data.pls.lamb) == i,]
 cvloo = pls::plsr(lambda ~ CovAmb, validation = 'CV', data = train,
               ncomp = cv.results[which.max(cv.results$eblup), 'ncomp'])
 pred_blup = tcrossprod(
    test$CovAmb %*% coef(cvloo, intercept = T)[-1,,1] +
      coef(cvloo, intercept = T)[1,,1],
    fa4$rot.scores
  )
 blue = eblues[[i]] |>
    dplyr::left_join(data.frame(t(pred_blup)) |>
                       tibble::rownames_to_column('gen'), by = 'gen')
  acc_eblue = c(acc_eblue, cor(blue[,2], blue[,3], use = "na.or.complete",
                method = "spearman"))
 rm(cvloo, train, test, blue, pred_blup)
}
mean(acc_eblue)
```

## [1] 0.5490957

## **GIS-GGE**

The GIS-GGE is a PLS model whose training dataset uses the estimated eBLUP values of genotypic means within observed environments (Costa-Neto et al., 2022). We hypothesize the GIS-FA model has a higher prediction accuracy than the GIS-GGE model.

In the dataset used for this tutorial, there are 182 genotypes and 49 observed environments. Let's now apply the GIS-GGE model and compare its results with the GIS-model, as follows:

```
# Building the data frame
GGE = fa4$blups
GGE = reshape(GGE[,1:3], idvar = 'gen', timevar = 'env', direction = 'wide')
colnames(GGE)[-1] = sub("conditional.","",colnames(GGE)[-1])
```



```
rownames(GGE) = levels(fa4$blups$gen)
GGE = t(GGE[,-1])
data.pls.gge = data.frame(GGE = I(GGE), CovAmb = I(scale(matcovamb)))
# Fitting the first PLS model
pls1.gge = pls::plsr(GGE ~ CovAmb, validation = 'CV', data = data.pls.gge)
# Performing the cross-validation
ncomp = list()
for (j in 1:pls1.gge$ncomp) {
  acc_eblue.gge = vector()
 acc_eblup.gge = vector()
 for (i in rownames(matcovamb)) {
   train = data.pls.gge[rownames(data.pls.gge) != i,]
   test = data.pls.gge[rownames(data.pls.gge) == i,]
   cvloo = pls::plsr(GGE ~ CovAmb, validation = 'CV', data = train,
                 ncomp = j
   pred_blup = test$CovAmb %*% coef(cvloo, intercept = T)[-1,,1] +
      coef(cvloo, intercept = T)[1,,1]
   blue = eblues[[i]] |>
      dplyr::left_join(data.frame(t(pred_blup)) |>
                         tibble::rownames_to_column('gen'), by = 'gen')
   blup = fa4$blups[fa4$blups$env == i, c('gen', 'conditional')] |>
       dplyr::left_join(data.frame(t(pred_blup)) |>
                          tibble::rownames_to_column('gen'), by = 'gen')
   acc_eblue.gge = c(acc_eblue.gge, cor(blue[,2], blue[,3],
                                         use = "na.or.complete", method = 'spearman'))
   acc_eblup.gge = c(acc_eblup.gge, cor(blup[,2], blup[,3],
                                         use = "na.or.complete", method = 'spearman'))
   rm(cvloo, train, test, blue, pred_blup, blup)
 ncomp[[j]] = data.frame(ncomp = j, eblue = mean(acc_eblue.gge),
                          eblup = mean(acc_eblup.gge))
cv.results.gge = do.call(rbind, ncomp)
# Choosing the number of components
cv.results.gge[which.max(cv.results.gge$eblup),] #Highest accuracy for predicting eBLUPs
##
               eblue
                         eblup
    ncomp
## 8
        8 0.5137178 0.7068775
```



```
cv.results.gge[which.max(cv.results.gge$eblue),] #Highest accuracy for predicting eBLUEs
## ncomp eblue eblup
## 1 1 0.5305078 0.6896101
```

#### GIS-FA vs GIS-GGE

The results are presented in the table below. The predictive ability of the GIS-FA model was 5.2% and 3.5% superior than the GIS-GGE model when predicting eBLUP and eBLUE values, respectively.

Models	m eBLUPs	m eBLUEs
GIS-FA	0.7384869	0.5481807
GIS-GGE	0.7071189	0.5306063

Let's clean R memory before jumping into the next section:

```
rm(cv.results.gge, cv.results, data.pls.gge, data.pls.lamb, eblues, GGE, ncomp,
    pls1.gge, pls1.lamb, acc_eblue, acc_eblue.gge, acc_eblup, acc_eblup.gge,
    acc_lamb, i, j)
```

#### Prediction in untested environments

Given the GIS-FA model was superior to the GIS-GGE model, let's predict the phenotypic performance of the 182 genotypes in untested environments using the GIS-FA model. The first step is extracting the regression coefficients from the fitted model:

```
coef_pls = coef(pls2.lamb, intercept = T)[,,1]
```

Then, we will use the environmental data from the unobserved environments to estimate their factor loadings:

```
pred_lambda = scale(matcovamb_pred) %*% coef_pls[-1, ]

pred_lambda = cbind(
   fa1 = pred_lambda[,'fa1'] + coef_pls["(Intercept)", "fa1"],
   fa2 = pred_lambda[,'fa2'] + coef_pls["(Intercept)", "fa2"],
   fa3 = pred_lambda[,'fa3'] + coef_pls["(Intercept)", "fa3"],
   fa4 = pred_lambda[,'fa4'] + coef_pls["(Intercept)", "fa4"]
)
```

And finally, let's combine multiple regression with the FA model (i.e., the GIS-FA) to predict the eBLUP values in unobserved environments:



## Section 4: Visualization using maps

After performing the predictions, we can build thematic maps to aid the visualization. In addition to the shapefile of the Brazilian state's administrative boundaries already loaded into R, we will also load a shapefile with the boundaries of Brazilian biomes. We intend to highlight in the map the portion corresponding to Pantanal, a protected area. Pantanal is represented by a contour with the same colour as the background in maps.

```
br_bio = raster::shapefile("shapefiles/biomes/lm_bioma_250.shp")
bio_sf = sf::st_as_sf(br_bio)
Pant = bio_sf[bio_sf$Bioma == 'Pantanal',]
Pant_crop = sf::st_intersection(Pant, MS)
```

## Interpolation

Recall that we made the prediction only for sampled points. To expand the prediction to the whole state, we have to perform an interpolation process. The procedure is similar to what was previously described.

```
# Building a list with data frames containing the eBLUPs of each sampled point
pred.list = lapply(apply(EBLUPs_pred, 2, function(x) reshape2::melt(x)),
                          function(x) x |>
                     dplyr::rename(blup = value) |>
                     tibble::rownames to column('point') |>
                     dplyr::left_join(
                       unique(covamb_pred[,c('env', 'LON', 'LAT')]),
                                      by = c('point' = 'env'))
                   )
# Initializing the parallelization
cl = parallel::makeCluster(parallel::detectCores(logical = F)-1)
parallel::clusterEvalQ(cl, {library(spatstat)
                            library(raster)
                            library(sf)})
parallel::clusterExport(cl, varlist = c('obs_window', 'MS'))
pred.list = parallel::parLapply(cl, pred.list, function(x){
```



As before, this was computationally intensive process, so we will save the outcome as follows:

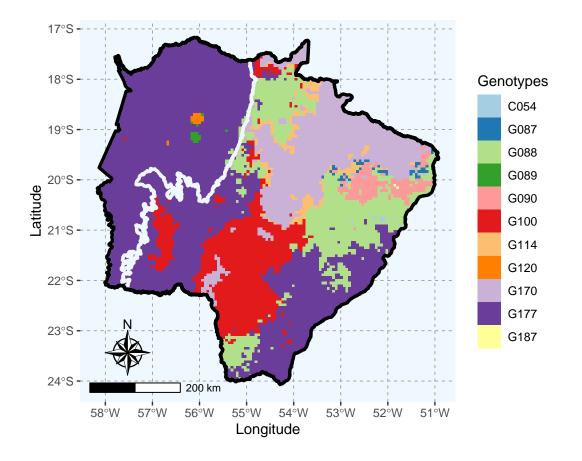
```
save(pred.list, file = "saves/pred_blups.RData")
```

Now everything is set for building the maps.

#### Which-won-where

The "which-won-where" map shows which genotype performed best on each unobserved environment:



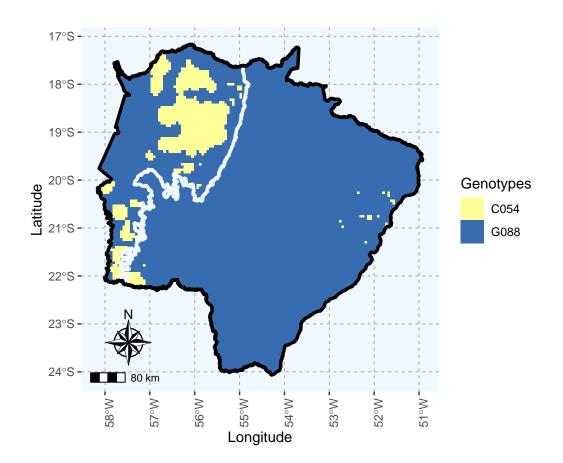


### Pairwise comparisons

The following map presents the pairwise comparison between an experimental genotype (G088) and a check cultivar (C054). The winner is represented by its colour, as in the "which-won-where" plot.



```
)) |>
ggplot() +
ggspatial::annotation_scale(location = "bl", width_hint = .1) +
ggspatial::annotation_north_arrow(location = "bl", which_north = "true",
                       pad_x = unit(0.5, "cm"), pad_y = unit(0.7, "cm"),
                       style = ggspatial::north_arrow_nautical())+
geom_raster(aes(x = x, y = y , fill = winner)) +
geom_sf(data = Pant_crop, fill = NA, lwd=1.2, linetype = 'solid',
        color = 'aliceblue', show.legend = T, na.rm = T)+
geom_sf(data= MS, fill = NA, linewidth = 1.2, color = 'black')+
labs(x = 'Longitude', y = 'Latitude',
     fill = "Genotypes", color = "")+
theme(panel.grid.major = element_line(color = gray(.6), linetype = "dashed",
                                      linewidth = 0.3),
      panel.background = element_rect(fill = "aliceblue"),
      legend.position = 'right',
      \#legend.box.margin = margin(.05, .05, .05, .05, 'cm'),
      legend.spacing = unit(.02,'cm'),
      axis.text.x = element_text(angle = 90)) +
scale_fill_manual(values = c("C054" = "#ffff99", "G088" = "#386cb0")) +
guides(fill = guide_legend(order = 1, override.aes = list(colour = 0)))
```





## Genotypes' performance map

The following map depict the expected performance of genotypes in untested environments, which might be useful for recommendation purposes.

Let's first compile the pred.list object into a single data frame:

```
EBLUPs_pred = do.call(rbind, pred.list) |>
  rownames_to_column('gen') |>
  separate(gen, into = c('gen', 'trash')) |>
  dplyr::select(-trash) |>
  rename(blup = layer)

str(EBLUPs_pred)
```

```
## 'data.frame': 1842750 obs. of 4 variables:
## $ gen : chr "C006" "C006" "C006" "C006" ...
## $ blup: num 3368 3386 3401 3342 3362 ...
## $ x : num -56.2 -56.1 -56 -56.2 -56.2 ...
## $ y : num -17.2 -17.2 -17.2 -17.2 ...
```

Then, we will divide the genotypes' predicted performance into groups to facilitate interpretation. For example, performances superior to  $4000 \text{ kg ha}^{-1}$  are classified as group 8, between 3750 and  $4000 \text{ kg ha}^{-1}$  as group 7, and so on. This classification helps to determine if a given genotype will have a minimum predicted performance in a specific environment or region.

```
EBLUPs_pred = EBLUPs_pred %>% mutate(
    class = case_when(
        blup > min(blup, na.rm = T) & blup <= 2500 ~ 1,
        blup > 2500 & blup <= 2750 ~ 2,
        blup > 2750 & blup <= 3000 ~ 3,
        blup > 3000 & blup <= 3250 ~ 4,
        blup > 3250 & blup <= 3500 ~ 5,
        blup > 3500 & blup <= 3750 ~ 6,
        blup > 3750 & blup <= 4000 ~ 7,
        blup > 4000 & blup <= max(blup, na.rm = T) ~ 8
    )
)</pre>
```

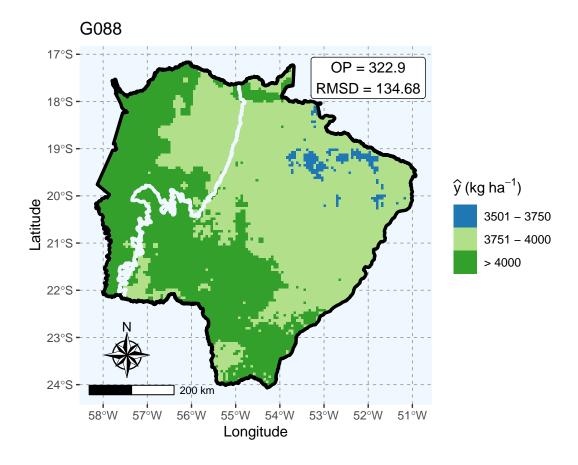
Lastly, we will calculate the Overall Performance (OP) and the root mean square deviation (RMSD) using the factor analytic selection tools. These metrics can be employed as a filter for choosing which genotypes the breeder should focus on.



Finally, we can build the adaptation maps. For example, for G088:

```
g = "G088"
EBLUPs_pred |> filter(gen == g) |>
  ggplot() +
  ggspatial::annotation_scale(location = "bl", width_hint = 0.3) +
  ggspatial::annotation_north_arrow(location = "bl", which north = "true",
                         pad_x = unit(0.5, "cm"), pad_y = unit(0.7, "cm"),
                         style = ggspatial::north_arrow_nautical())+
  geom_raster(aes(x = x, y = y , fill = factor(class))) +
  geom_sf(data = Pant_crop, fill = NA, lwd=1.2, linetype = 'solid',
          color = 'aliceblue', show.legend = T, na.rm = T)+
  geom_sf(data= MS, fill = NA, linewidth = 1.2, color = 'black')+
  labs(x = 'Longitude', y = 'Latitude',
       fill = expression(widehat(y)~"("*kg~ha^{-1}*")"), color = "",
       title = g)+
  theme(panel.grid.major = element_line(color = gray(.6), linetype = "dashed",
                                        linewidth = 0.3),
        panel.background = element_rect(fill = "aliceblue"),
        legend.position = 'right',
        legend.spacing = unit(.02,'cm')) +
  scale_fill_manual(values = c("1" = '#e31a1c',
                                 '2' = '#fb9a99',
                                 '3' = '#fdbf6f',
                                 '4' = '#ff7f00',
                                 '5' = '#a6cee3'.
                                 '6' = '#1f78b4'
                                 '7' = '#b2df8a',
                                 '8' = '#33a02c'),
                      labels = c("1" = 'Non-adapted',
                                 '2' = '2500 - 2750',
                                 '3' = '2751 - 3000',
                                 '4' = '3001 - 3250'
                                 '5' = '3251 - 3500',
                                 '6' = '3501 - 3750'
                                 '7' = '3751 - 4000',
                                 '8' = '> 4000')) +
```





# Closing remarks

In this tutorial, we presented the main ideas from our paper. We hope the GIS-FA model to be routinely incorporated in breeding programs as an important tool for cultivar recommendation in MET.



## References

- Akaike, H. (1974). A new look at the statistical model identification. *IEEE Transactions on Automatic Control*, 19(6), 716–723. https://doi.org/10.1109/TAC.1974.1100705
- Butler, D. (2022). Asreml: Fits the linear mixed model. www.vsni.co.uk
- Chaves, S. F. S., Alves, R. S., Dias, L. A. S., Alves, R. M., Dias, K. O. G., & Evangelista, J. S. P. C. (2023). Analysis of repeated measures data through mixed models: An application in *Theobroma grandiflorum* breeding. *Crop Science*, 63(4), 2131–2144. https://doi.org/10.1002/csc2.20995
- Chaves, S. F. S., Evangelista, J. S. P. C., Trindade, R. S., Dias, L. A. S., Guimarães, P. E., Guimarães, L. J. M., Alves, R. S., Bhering, L. L., & Dias, K. O. G. (2023). Employing factor analytic tools for selecting high-performance and stable tropical maize hybrids. *Crop Science*, 63(3), 1114–1125. https://doi.org/10.1002/csc2.20911
- Corporation, M., & Weston, S. (2022). doParallel: Foreach parallel adaptor for the 'parallel' package. https://CRAN.R-project.org/package=doParallel
- Costa-Neto, G., Crespo-Herrera, L., Fradgley, N., Gardner, K., Bentley, A. R., Dreisigacker, S., Fritsche-Neto, R., Montesinos-López, O. A., & Crossa, J. (2022). Envirome-wide associations enhance multi-year genome-based prediction of historical wheat breeding data. *G3: Genes/Genomes/Genetics*, 13(2), jkac313. https://doi.org/10.1093/g3journal/jkac313
- Costa-Neto, G., Gilli, G., Carvalho, H. F., Crossa, J., & Fritsche-Neto, R. (2021). EnvR-type: A software to interplay environics and quantitative genomics in agriculture. G3 Genes/Genomes/Genetics, 11, jkab040. https://doi.org/10.1093/g3journal/jkab040
- Cullis, B. R., Smith, A. B., Beeck, C. P., & Cowling, W. A. (2010). Analysis of yield and oil from a series of canola breeding trials. Part II. Exploring variety by environment interaction using factor analysis. *Genome*, 53(11), 1002–1016. https://doi.org/10.1139/G10-080
- Cullis, B. R., Smith, A. B., & Coombes, N. E. (2006). On the design of early generation variety trials with correlated data. *Journal of Agricultural, Biological, and Environmental Statistics*, 11(4), 381–393. https://doi.org/10.1198/108571106X154443
- Dunnington, D. (2023). Ggspatial: Spatial data framework for ggplot2. https://CRAN.R-project.org/package=ggspatial
- Gu, Z. (2022). Complex heatmap visualization. iMeta.
- Gu, Z., Gu, L., Eils, R., Schlesner, M., & Brors, B. (2014). Circlize implements and enhances circular visualization in r. *Bioinformatics*, 30, 2811–2812.
- Hijmans, R. J. (2023). Raster: Geographic data analysis and modeling. https://CRAN.R-project.org/package=raster
- Hijmans, R. J., Barbosa, M., Ghosh, A., & Mandel, A. (2023). Geodata: Download geographic data. https://CRAN.R-project.org/package=geodata
- Kassambara, A. (2023). Ggpubr: 'ggplot2' based publication ready plots. https://CRAN.R-project.org/package=ggpubr
- Lembrechts, J. J., Hoogen, J. van den, Aalto, J., Ashcroft, M. B., De Frenne, P., Kemppinen, J., Kopecký, M., Luoto, M., Maclean, I. M. D., Crowther, T. W., Bailey, J. J., Haesen, S., Klinges, D. H., Niittynen, P., Scheffers, B. R., Van Meerbeek, K., Aartsma, P., Abdalaze, O., Abedi, M., . . . Lenoir, J. (2022). Global maps of soil temperature. *Glob Chang. Biol.*, 28(9), 3110–3144. https://doi.org/10.1111/gcb.16060
- Liland, K. H., Mevik, B.-H., & Wehrens, R. (2022). Pls: Partial least squares and principal component regression. https://CRAN.R-project.org/package=pls
- Microsoft, & Weston, S. (2022). Foreach: Provides foreach looping construct. https://CRAN.R-project.org/package=foreach
- Mrode, R. A. (2014). Linear models for the prediction of animal breeding values (3rd ed). CABI.



- Pebesma, E., & Bivand, R. (2023). Spatial Data Science: With applications in R (p. 352). Chapman and Hall/CRC. https://r-spatial.org/book/
- Piepho, H.-P. (1997). Analyzing genotype-environment data by mixed models with multiplicative terms. *Biometrics*, 53(2), 761–766. https://doi.org/10.2307/2533976
- Piepho, H.-P. (2019). A coefficient of determination (r<sup>2</sup>) for generalized linear mixed models. Biometrical Journal, 61(4), 860–872. https://doi.org/10.1002/bimj.201800270
- R Core Team. (2023). R: A language and environment for statistical computing. R Foundation for Statistical Computing. https://www.R-project.org/
- Slowikowski, K. (2023). Ggrepel: Automatically position non-overlapping text labels with 'ggplot2'. https://CRAN.R-project.org/package=ggrepel
- Smith, A. B., & Cullis, B. R. (2018). Plant breeding selection tools built on factor analytic mixed models for multi-environment trial data. *Euphytica*, 214(8), 143. https://doi.org/10.1007/s10681-018-2220-5
- Smith, A. B., Cullis, B. R., & Thompson, R. (2001). Analyzing variety by environment data using multiplicative mixed models and adjustments for spatial field trend. *Biometrics*, 57(4), 1138–1147. https://doi.org/10.1111/j.0006-341X.2001.01138.x
- Smith, A. B., Ganesalingam, A., Kuchel, H., & Cullis, B. R. (2015). Factor analytic mixed models for the provision of grower information from national crop variety testing programs. *Theoretical and Applied Genetics*, 128(1), 55–72. https://doi.org/10.1007/s00122-014-2412-x
- Smith, A. B., Norman, A., Kuchel, H., & Cullis, B. R. (2021). Plant variety selection using interaction classes derived from factor analytic linear mixed models: Models with independent variety effects. Frontiers in Plant Science, 12, 1857. https://doi.org/10.3389/fpls.2021.737462
- Stefanova, K. T., & Buirchell, B. (2010). Multiplicative mixed models for genetic gain assessment in lupin breeding. *Crop Science*, 50(3), 880–891. https://doi.org/10.2135/cropsci2009.07.0402
- Wickham, H., Averick, M., Bryan, J., Chang, W., McGowan, L. D., François, R., Grolemund, G., Hayes, A., Henry, L., Hester, J., Kuhn, M., Pedersen, T. L., Miller, E., Bache, S. M., Müller, K., Ooms, J., Robinson, D., Seidel, D. P., Spinu, V., ... Yutani, H. (2019). Welcome to the tidyverse. *Journal of Open Source Software*, 4(43), 1686. https://doi.org/10.21105/joss.01686
- Wong, J. (2022). Pdist: Partitioned distance function. https://CRAN.R-project.org/package=pdist
- Yutani, H. (2022). Gghighlight: Highlight lines and points in 'ggplot2'. https://CRAN.R-project.org/package=gghighlight

