

1 Within group variable selection

Let's focus on only one group for now. Our model is as follows:

$$P(Y_i = k | X_i, \beta, \gamma) = \frac{\exp \left(X_i \cdot \beta_k \cdot 1_{k \neq K} \right)}{1 + \sum_{\ell=1}^{K-1} \exp \left(X_i \cdot \beta_\ell \right)}, \quad k = 1, \dots, K. \quad (1)$$

We also have for $j = 1, \dots, p$ and $k = 1, \dots, K$,

$$\begin{aligned} \rho_j &\sim \text{Beta}(\alpha_1, \alpha_2) \\ \sigma_{0j}^2 &\sim \text{IG}(\lambda_0, \nu_0) \\ \sigma_{1j}^2 &\sim \text{IG}(\lambda_1, \nu_1) \\ \gamma_j | \rho_j &\sim \text{Bernoulli}(1 - \rho_j) \\ \beta_{jk} | \gamma_j, \sigma_{0j}, \sigma_{1j} &\sim \text{Normal}(0, \tau_j^2), \end{aligned}$$

where $\tau_j^2 = \sigma_{0j}^{2(1-\gamma_j)} \sigma_{1j}^{2\gamma_j}$. We can sample from the posterior $P(\gamma, \beta, \sigma_0^2, \sigma_1^2, \rho | Y_i = k, X)$:

Step 1: Sample from

$$\begin{aligned} P(\gamma | \beta, \rho, \sigma_0^2, \sigma_1^2, Y, X) &\propto P(Y | \beta, \sigma_0^2, \sigma_1^2, X) \cdot P(\beta | \gamma, \sigma_0^2, \sigma_1^2) \cdot P(\gamma | \rho) \\ &\propto P(\beta | \gamma, \sigma_0^2, \sigma_1^2) \cdot P(\gamma | \rho). \end{aligned}$$

Since the γ_j are binary, we can directly sample from

$$P(\gamma | \beta, \rho, \sigma_0^2, \sigma_1^2, Y, X) \propto \exp \left(\sum_{\ell=1}^{K-1} \sum_{j=1}^p -\frac{\beta_{j\ell}^2}{2\tau_j^2} \right) \cdot \prod_{j=1}^p (1 - \rho_j)^{\gamma_j} \rho_j^{(1-\gamma_j)} \prod_{\ell=1}^{K-1} \tau_j^{-1}$$

to produce B samples $\hat{\gamma}^{(1)}, \dots, \hat{\gamma}^{(B)}$. This expression can be written as

$$\begin{aligned} \log P(\gamma | \beta, \rho, \sigma_0^2, \sigma_1^2, Y, X) &= (1 - K) \sum_{j=1}^p \log \tau_j - \frac{1}{2} \sum_{\ell=1}^{K-1} \sum_{j=1}^p \frac{\beta_{j\ell}^2}{\tau_j^2} + \\ &\quad \sum_{j=1}^p \left(\gamma_j \log(1 - \rho_j) + (1 - \gamma_j) \log \rho_j \right). \end{aligned}$$

To sample γ , we can sample one γ_j at a time while holding the remaining γ_{-j} fixed; the log likelihood needed here is

$$\log P(\gamma_j | \beta_j, \rho_j, \sigma_{0j}^2, \sigma_{1j}^2, Y, X) = (1 - K) \log \tau_j - \frac{1}{2\tau_j^2} \sum_{\ell=1}^{K-1} \beta_{j\ell}^2 + \gamma_j \log(1 - \rho_j) + (1 - \gamma_j) \log \rho_j.$$

Step 2: Sample from

$$P(\beta|\gamma, \sigma_0^2, \sigma_1^2, \rho, Y_1 = k_1, \dots, Y_n = k_n, X) \propto P(Y_1 = k_1, \dots, Y_n = k_n|\beta, X) \cdot P(\beta|\gamma, \sigma_0^2, \sigma_1^2).$$

We can use MCMC to sample from

$$P(\beta|\gamma, \sigma_0^2, \sigma_1^2, \rho, Y_1 = k_1, \dots, Y_n = k_n, X) \propto \exp\left(\sum_{\ell=1}^{K-1} \sum_{j=1}^p -\frac{\beta_{j\ell}^2}{2\tau_j^2}\right) \prod_{i=1}^n \frac{\exp\left(X_i \cdot \beta_{k_i} \cdot 1_{k_i \neq K}\right)}{1 + \sum_{\ell=1}^{K-1} \exp\left(X_i \cdot \beta_{\ell}\right)}$$

to produce B samples $\hat{\beta}^{(1)}, \dots, \hat{\beta}^{(B)}$. This expression can be written as

$$\begin{aligned} \log P(\beta|\gamma, \sigma_0^2, \sigma_1^2, \rho, Y_1 = k_1, \dots, Y_n = k_n, X) = & -\frac{1}{2} \sum_{\ell=1}^{K-1} \sum_{j=1}^p \frac{\beta_{j\ell}^2}{\tau_j^2} + \sum_{i=1}^n X_i \cdot \beta_{k_i} \cdot 1_{k_i \neq K} \\ & - \sum_{i=1}^n \log\left(1 + \sum_{\ell=1}^{K-1} \exp(X_i \cdot \beta_{\ell})\right). \end{aligned}$$

To sample β , we can sample one β_j vector at a time while holding the remaining β_{-j} fixed; define $U(\beta_j)$ as

$$\begin{aligned} U(\beta_j) = & -\log P(\beta_j|\gamma_j, \sigma_{0j}^2, \sigma_{1j}^2, \rho_j, Y_1 = k_1, \dots, Y_n = k_n, X) \\ = & \frac{1}{2\tau_j^2} \sum_{\ell=1}^{K-1} \beta_{j\ell}^2 - \sum_{i=1}^n X_i \cdot \beta_{k_i} \cdot 1_{k_i \neq K} + \sum_{i=1}^n \log\left(1 + \sum_{\ell=1}^{K-1} \exp(X_i \cdot \beta_{\ell})\right). \end{aligned}$$

Then

$$\frac{\partial}{\partial \beta_j} U(\beta_j) = \frac{1}{\tau_j^2} \beta_j + \sum_{i=1}^n \left(\frac{X_{ij} \exp(X_i \cdot \beta)}{1 + \sum_{\ell=1}^{K-1} \exp(X_i \cdot \beta_{\ell})} \right) - X_{Y,j},$$

where

$$\exp(X_i \cdot \beta) = [\exp(X_i \cdot \beta_1), \dots, \exp(X_i \cdot \beta_{K-1})]$$

$$X_{Y,j} = \left[\sum_{i \in A_1} X_{ij}, \dots, \sum_{i \in A_{K-1}} X_{ij} \right]$$

for

$$A_{\ell} = \{i; Y_i = \ell, i = 1, \dots, n\}, \quad \ell = 1, \dots, K-1.$$

We can use Langevin dynamics to sample β_j :

$$\beta_j^{(t+1)} = \beta_j^{(t)} - \frac{\epsilon^2}{2} \frac{\partial}{\partial \beta_j} U(\beta_j^{(t)}) + \epsilon \cdot \mathbf{Z}^{(t)},$$

where $\epsilon > 0$ is the step size and $\mathbf{Z}^{(t)} \sim N(0, I_{K-1})$.

Step 3: For $j = 1, \dots, p$, sample

$$\sigma_{0j}^2 \sim IG\left(\lambda_0 + (1 - \gamma_j)\frac{K}{2}, \nu_0 + \frac{(1 - \gamma_j)}{2} \sum_{k=1}^K \beta_{jk}^2\right)$$

$$\sigma_{1j}^2 \sim IG\left(\lambda_1 + \gamma_j\frac{K}{2}, \nu_1 + \frac{\gamma_j}{2} \sum_{k=1}^K \beta_{jk}^2\right).$$

Step 4: For $j = 1, \dots, p$, sample

$$\rho_j \sim Beta\left(\alpha_1 + 1 - \gamma_j, \alpha_2 + \gamma_j\right).$$