

Homework 3Exercise 8.1

let P the ideal patch

\tilde{P} the observed patch

N the noise patch

$C_{\tilde{P}}$ the covariance matrix of the patches similar to \tilde{P}

C_P the covariance matrix of the patches similar to P .

We assume that the patches and noise are independent.

We assume that if \tilde{P}' is a similar patch of \tilde{P} then, P' is a similar patch to P .

For any $\tilde{P}_i, \tilde{P}_j \in \tilde{Q}$, the set of similar patch to \tilde{P} ,

$$\begin{aligned} \text{Cov}(\tilde{P}_i, \tilde{P}_j) &= \text{Cov}_{\tilde{P}}(i, j) \\ &= \text{Cov}(P_i + N_i, P_j + N_j) \end{aligned}$$

with $\tilde{P} = P + N$

$$= E[(P_i + N_i - E[P_i])(P_j + N_j - E[P_j])]$$

bc $E[N] = 0$

$$= E[(P_i - E[P_i])(P_j - E[P_j])]$$

$$+ E[(P_i - E[P_i])N_j] = 0$$

$$+ E[(P_j - E[P_j])N_i] = 0$$

$$+ E[N_i N_j]$$

bc noise and patch are indep.

$$= \text{Cov}(P_i, P_j) + \sigma^2 \delta_{ij}$$

In the end,

$$\boxed{C_{\tilde{P}} = C_P + \sigma^2 I}$$

We define \bar{P} as the center of similar patches to P .

$$E[\tilde{Q}] = \sum_{P \in \tilde{Q}} \tilde{P} P(\tilde{P}) = \sum_{P \in \tilde{Q}} (P + N) P(P + N), \quad N \text{ noise}$$

$$= \sum_{P \in \tilde{Q}} P P(P) \quad \text{bc } P, N \text{ indep.}$$

$$= E[\tilde{Q}] = \bar{P}$$

ie

$$\boxed{E[\tilde{Q}] = \bar{P}}$$

Exercice 8.3

We want to compare:

$$\begin{aligned}\hat{P}_1 &= \bar{P} + [C_p - \sigma^2 I] C_p^{-1} (\hat{P} - \bar{P}) \\ \hat{P}_2 &= \bar{P}' + C_p' [C_p' + \sigma^2 I]^{-1} (\hat{P} - \bar{P}')\end{aligned}$$

Let G_i the eigen vector of C_p , with eigen value $\mu_i > 0$
 or G_i the eigen vector of C_p^{-1} , with eigen value $\frac{1}{\mu_i}$ (psd)
 So

$$\begin{aligned}\hat{P}_1 &= \bar{P} + [C_p - \sigma^2 I] C_p^{-1} (\hat{P} - \bar{P}) \\ &= \sum_{i=1}^n \langle \bar{P} + [C_p - \sigma^2 I] C_p^{-1} (\hat{P} - \bar{P}), G_i \rangle G_i \\ &= \sum_{i=1}^n \langle \bar{P}, G_i \rangle G_i + \underbrace{\langle \hat{P} - \bar{P}, ([C_p - \sigma^2 I] C_p^{-1})^T G_i \rangle}_{\text{bc } \{G_i\} \text{ orthon basis.}} G_i \\ &= \sum_{i=1}^n \langle \bar{P}, G_i \rangle G_i + \langle \hat{P} - \bar{P}, C_p^{-1} [C_p - \sigma^2 I] G_i \rangle G_i\end{aligned}$$

bc C_p and $[C_p - \sigma^2 I]$ symmetric.

$$= \sum_{i=1}^n \langle \bar{P}, G_i \rangle G_i + \frac{\mu_i}{\mu_i} \langle \hat{P} - \bar{P}, G_i \rangle G_i$$

$$\text{we } \hat{P}_1 = \sum_{i=1}^n \frac{\sigma^2}{\mu_i} \langle \bar{P}, G_i \rangle G_i + \sum_{i=1}^n \underbrace{\frac{\mu_i - \sigma^2}{\mu_i}}_{= a(\mu_i) = \frac{\mu_i - \sigma^2}{\mu_i}} \langle \hat{P}, G_i \rangle G_i$$

The second part of the formula is the Wiener filter!

Let H_i the eigen vector of C_p' , with eigen value μ_i'

We have

$$\begin{aligned}\hat{P}_2 &= \bar{P}' + C_p' [C_p' + \sigma^2 I]^{-1} (\hat{P} - \bar{P}') \\ &= \sum_{i=1}^n \langle \bar{P}' + C_p' [C_p' + \sigma^2 I]^{-1} (\hat{P} - \bar{P}'), H_i \rangle H_i \\ &= \sum_{i=1}^n \langle \bar{P}', H_i \rangle H_i + \underbrace{\frac{\mu_i'}{\mu_i' + \sigma^2}}_{\text{bc } \{H_i\} \text{ orthon basis.}} \langle \hat{P} - \bar{P}', H_i \rangle H_i\end{aligned}$$

$$u \quad \hat{P}_2 = \sum_{i=1}^n \frac{\sigma^2}{p_i^1 + \sigma^2} \langle \bar{P}^1 | H_i \rangle H_i + \sum_{i=1}^n \underbrace{\frac{p_i^1}{p_i^1 + \sigma^2}}_{b(i)} \langle \hat{P}^1 | H_i \rangle H_i$$

The second part is the Wiener filtering!

$$b(i) = \frac{p_i^1}{p_i^1 + \sigma^2}$$

Thus, the two steps Bayesian method corresponds to a Wiener and a scalar method. However, there is a small difference.

In this case, we denote the patch \tilde{P} , but add a weighted average patch $\frac{\sigma^2}{p_i^1} \bar{P} \propto \frac{\sigma^2}{p_i^1 + \sigma^2} \bar{P}^1$.

We keep patch \bar{P} coordinates with the smallest p_i elongation. This allows not to erase the high frequency pixels.

Exercice 8.4.

Recall Fubini-Tonelli's theorem:

$$\bullet P \mapsto \int |P(\tilde{P}) P(P|\tilde{P})| \|P - \hat{P}\|^2 dP < +\infty$$

(because finite number of patch)

$$\bullet \tilde{P} \mapsto \int |P(\tilde{P}) P(P|\tilde{P})| \|P - \hat{P}\|^2 d\tilde{P} < +\infty$$

(because finite number of patch).

$$\text{Then } \iint P(\tilde{P}) P(P|\tilde{P}) \|P - \hat{P}\|^2 dP d\tilde{P}$$

$$= \iint P(\tilde{P}) P(P|\tilde{P}) \|P - \hat{P}\|^2 d\tilde{P} dP$$

Formula 8.16 gives:

$$\text{MSE} = \int P(P) \int P(\tilde{P}|P) \|P - \hat{P}\|^2 d\tilde{P} dP$$

$$= \iint P(\tilde{P}, P) \|P - \hat{P}\|^2 dP d\tilde{P} \quad \text{Fubini Thm}$$

$$= \iint P(P|\tilde{P}) \times P(\tilde{P}) \|P - \hat{P}\|^2 dP d\tilde{P}$$

$$u \quad \boxed{\text{MSE} = \int P(\tilde{P}) \int P(P|\tilde{P}) \|P - \hat{P}\|^2 dP d\tilde{P}} \quad \text{Bayes formula}$$

Fubini Thm.

Exercice 8.5:

let

$$\hat{p} = \underset{\tilde{p}}{\text{argmin}} \text{MMSE}(\tilde{p}) \underset{\tilde{p}}{\text{argmin}} \int \underbrace{P(p|\tilde{p})}_{\text{convex function}} (p - \hat{p})^2 dp.$$

$$\frac{\partial \text{MMSE}(\tilde{p})}{\partial \tilde{p}} = -2 \int P(p|\tilde{p}) (p - \hat{p}) dp = 0$$

$$\text{ie } \int P(p|\tilde{p}) p dp = \int P(p|\tilde{p}) \hat{p} dp$$

$$= \hat{p} \int P(p|\tilde{p}) dp$$

$$= \hat{p} \frac{\int P(p, \tilde{p}) dp}{P(\tilde{p})}$$

$$= \hat{p} \frac{P(\hat{p})}{P(\tilde{p})}$$

$$\text{so } \hat{p} = \int P(p|\tilde{p}) p dp$$

$$\text{ie } \boxed{\hat{p} = E[p|\tilde{p}]}$$

This is the Bayesian estimator on MSE!