

Shift-Share IV

MIXTAPE TRACK



Roadmap

Motivation

Intuition

- Market Access Effects

- Medicaid Eligibility Effects

Formal Framework

Applications

- Market Access Effects

- Medicaid Eligibility Effects

Concluding Thoughts

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How can we just leverage the exogenous shocks to such z_i ?

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3. Recentering solution also can have attractive efficiency properties
 - Leverages non-random exposure to best predict shock effects

(Some) Other Settings where these Points are Relevant

- Linear shift-share IV (Autor et al. 2013, Borusyak et al. 2022)
- Nonlinear shift-share IV (Boustan et al. 2013, Berman et al. 2015, Chodorow-Reich and Wieland 2020, Derenoncourt 2021)
- IV based on centralized school assignment mechanisms (Abdulkadiroğlu et al. 2017, 2019, Angrist et al. 2020)
- Model-implied optimal IV (Adão-Arkolakis-Esposito 2021)
- Weather instruments (Gomez et al. 2007, Madestam et al. 2013)
- “Free space” instruments for mass media access (Olken 2009, Yanagizawa-Drott 2014)

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Example 1: Market Access Effects via RCT

Theory suggests transportation upgrades affect local outcomes (e.g. land value) of regions i by increasing their market access (MA):

$$\Delta \log V_i = \beta \Delta \log MA_i + \varepsilon_i,$$

$$\text{where } MA_{it} = \sum_j \tau(g_t, loc_i, loc_j)^{-1} pop_j,$$

for road network g_t in periods $t = 1, 2$, region locations loc_j (co-determining travel cost τ), and regional population pop_j

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No. Randomizing roads \nrightarrow randomizing MA due to them!

Illustration: Market Access on a Square Island

Start from no roads, assume equal population everywhere

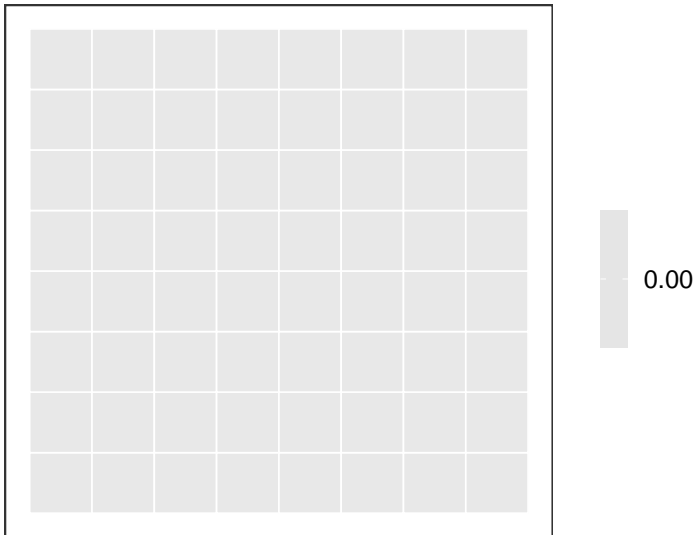


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Randomly connect adjacent regions by road

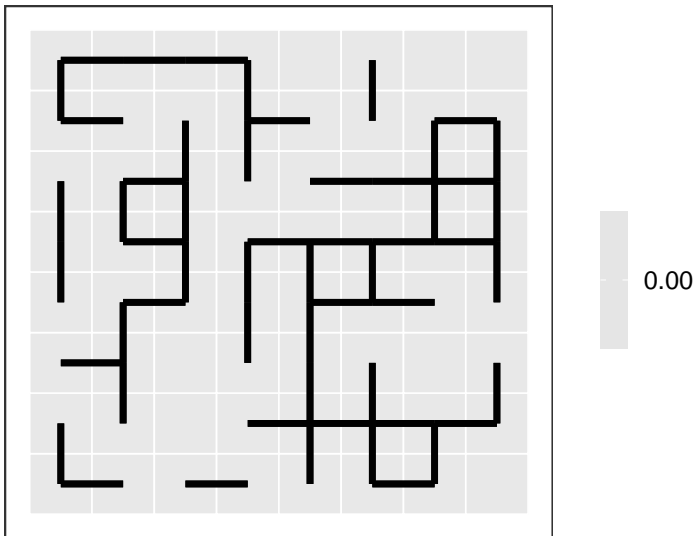


Illustration: Market Access on a Square Island

Randomly connect adjacent regions by road and compute MA growth

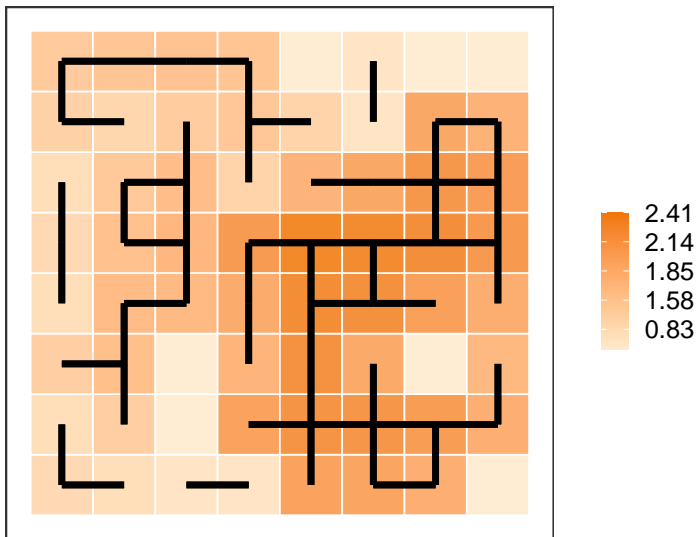


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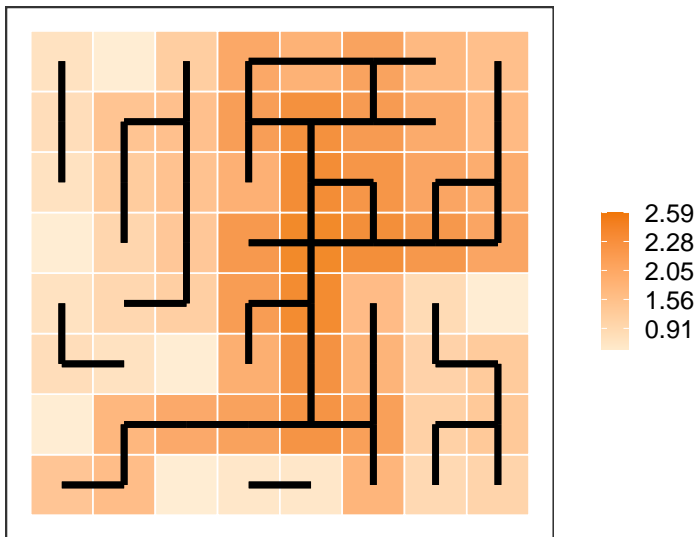
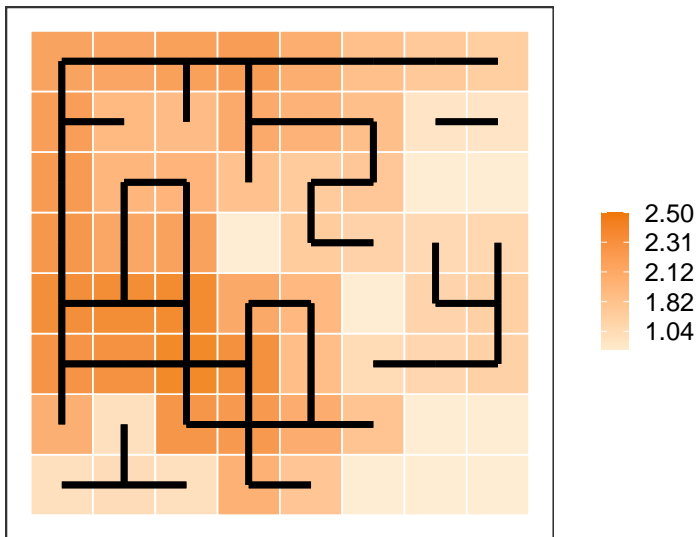


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Expected Market Access Growth μ_i

Some regions get systematically more MA

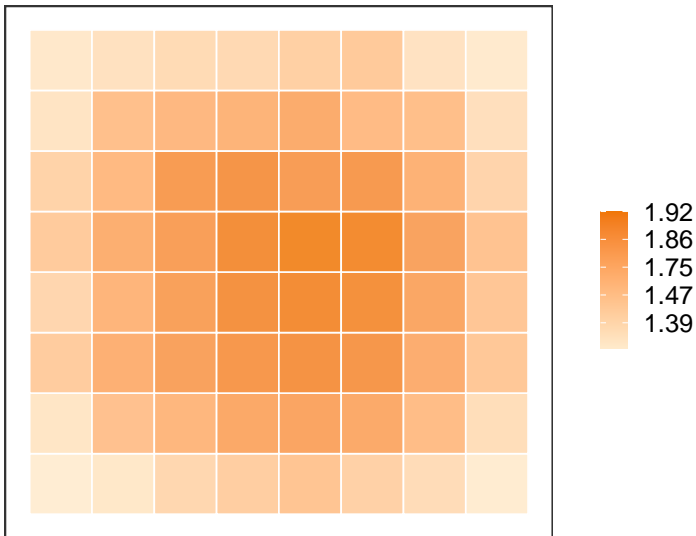


Illustration: High-Speed Rail in China

149 lines were built or planned (as of April 2019)

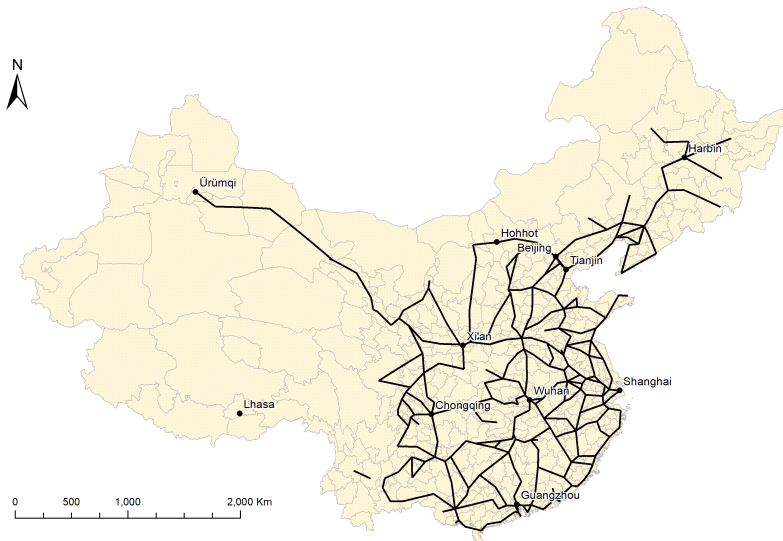


Illustration: High-Speed Rail in China

The 83 lines actually built by 2016. Suppose timing is random

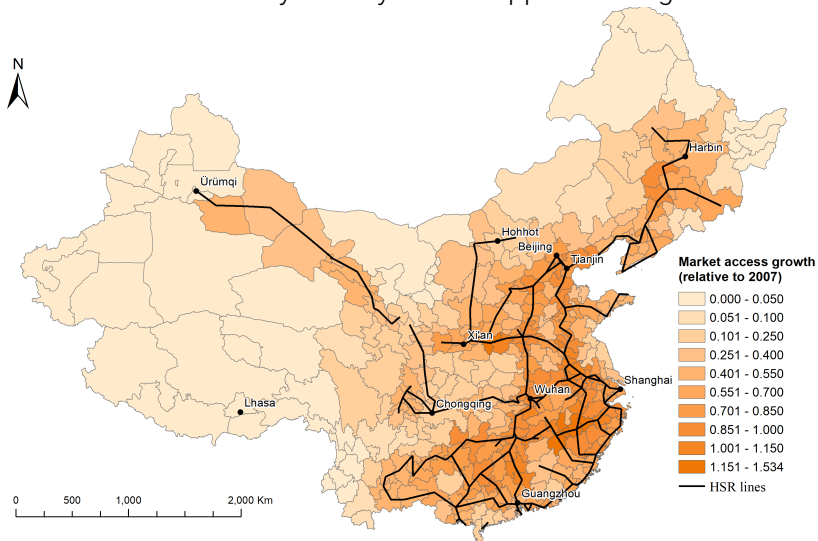


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A counterfactual draw of 83 lines by 2016

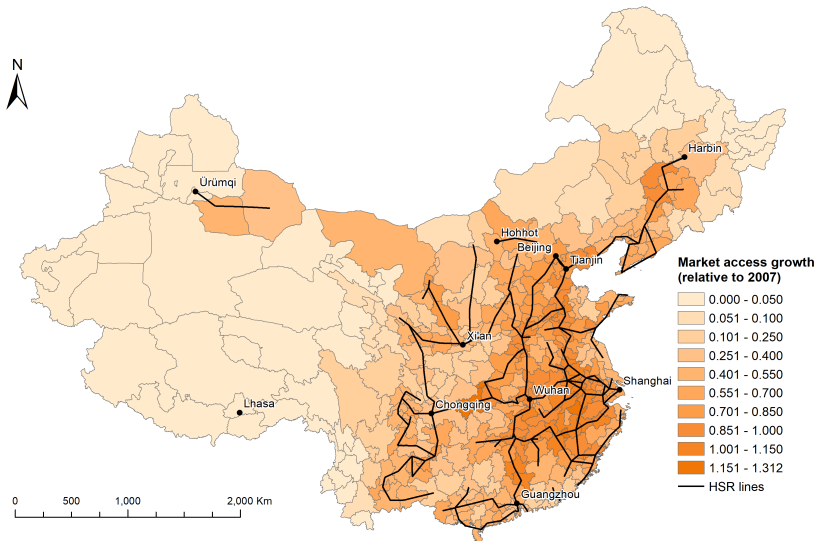
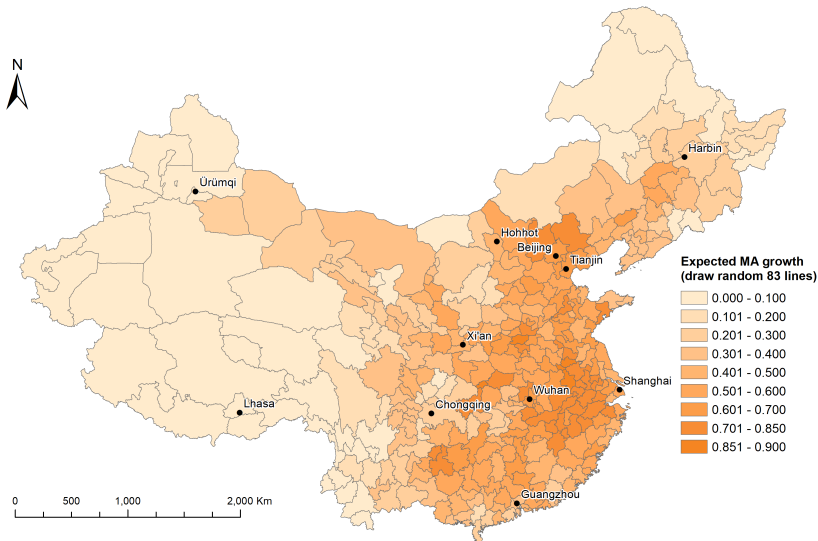


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Expected MA growth, μ_i



OVB and Recentering Solution

Systematic variation in MA growth can generate OVB

- E.g. land values fall in the periphery because of rising sea levels
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Systematic variation can be removed via “recentering”:

$$\begin{array}{ccccc} \text{Recentered} & & \text{Realized} & & \text{Expected} \\ \text{MA growth} & = & \text{MA growth} & - & \text{MA growth} \end{array}$$

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Systematic variation can be removed via “recentering”:

$$\text{Recentered MA growth} = \text{Realized MA growth} - \text{Expected MA growth}$$

Recentered MA is a valid instrument for realized MA growth

- Compares MA from actual and counterfactual shocks

Example 2: Effects of Program Eligibility

Consider the effects of individual eligibility x_i for Medicaid:

$$y_i = \beta x_i + \varepsilon_i$$

where x_i is determined by i 's state policy g_{state_i} and demographics

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Standard “simulated instruments” solution (Currie and Gruber (1996)):
use state-level variation (a measure of policy generosity) as IV for x_i

- This works, but is likely inefficient: policy shocks have heterogeneous effects across individuals

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- Yields efficiency gain by better first-stage prediction, e.g. by removing i who are always or never eligible

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General Setup

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- Applies to any z_i which can be constructed from observed data
- Nests reduced-form regressions: $x_i = z_i$
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Assumptions:

1. Shocks are exogenous: $g \perp \varepsilon \mid w$
2. Conditional distribution $G(g \mid w)$ is known (e.g. via randomization protocol or uniform across permutations of g)

Main Results

The expected instrument, $\mu_i = E[f_i(g, w) \mid w] \equiv \int f_i(g, w) dG(g \mid w)$, is the sole confounder generating OVB:

$$E \left[\frac{1}{N} \sum_i z_i \varepsilon_i \right] = E \left[\frac{1}{N} \sum_i \mu_i \varepsilon_i \right] \neq 0, \text{ in general}$$

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Regressions which control for μ_i also identify β (implicitly recenter, by the FWL theorem)

Extensions

Consistency: follows when \tilde{z}_i is weakly mutually dependent across i

Robustness to heterogeneous treatment effects: \tilde{z}_i identifies a convex avg. of β_i under appropriate first-stage monotonicity

Randomization inference provides exact confidence intervals for β (under constant effects) and falsification tests

BH also characterize the **asy. efficient** recentered IV among all $f_i(\cdot)$

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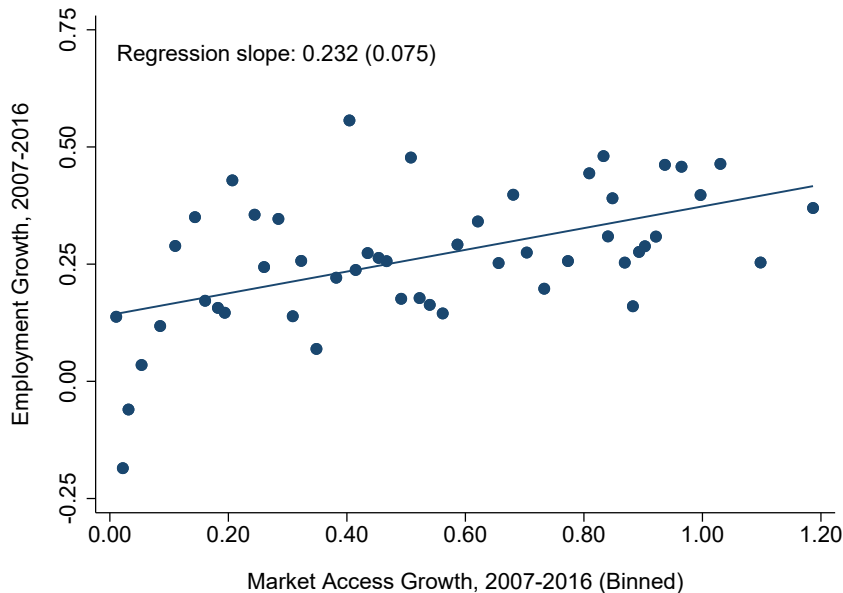
App. 1: Market Access from Chinese High-Speed Rail

We first show how instrument recentering can address OVB when estimating the effects of market access growth

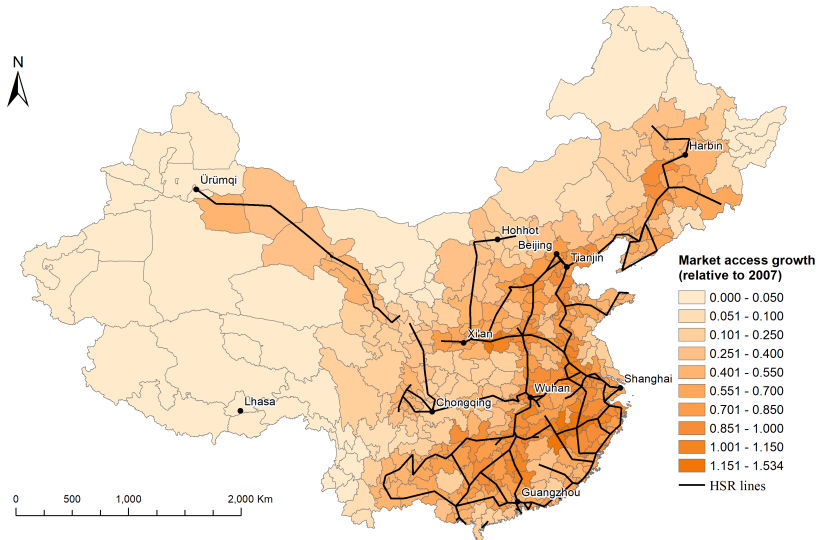
Setting: Chinese HSR; 83 lines built 2008–2016, 66 yet unbuilt

- Market access: $MA_{it} = \sum_k \exp(-0.02\tau_{ikt}) p_{k,2000}$, where τ_{ikt} is HSR-affected travel time between prefecture capitals (Zheng and Kahn, 2013) and $p_{i,2000}$ is prefecture i 's population in 2000
- Relate to employment growth in 274 prefectures, 2007-2016

Simple OLS Regressions Suggest a Large MA Effect



High vs. Low MA Growth is Not a Convincing Contrast!



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Add controls (province FE, longitude, etc...)

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Find valid contrasts for *one* source of variation—a natural experiment

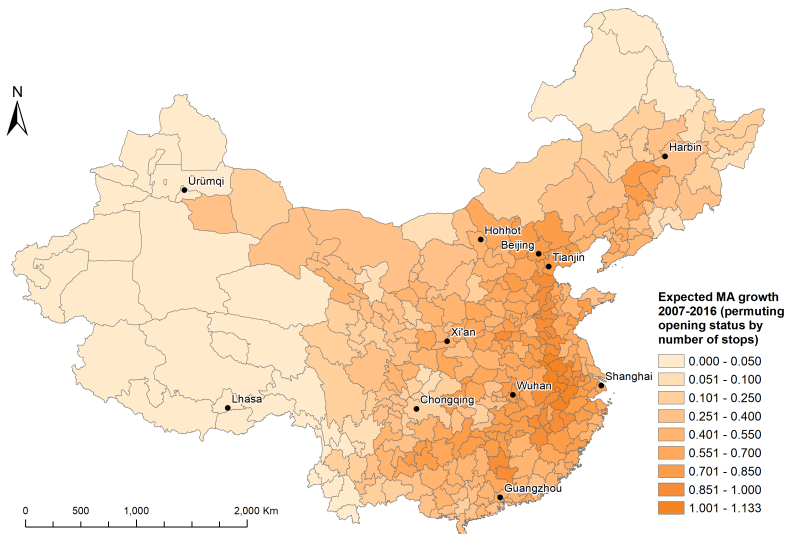
- Bartelme (2018): shocks affecting market size
- Donaldson (2018): built vs unbuilt lines
- BH application: assume random timing of observably similar lines

Built and Planned HSR Lines

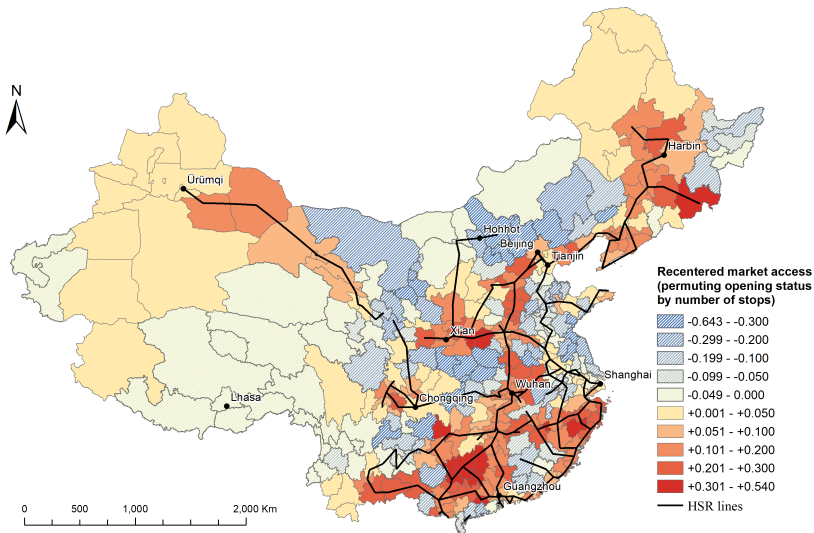
BH reshuffle built & planned lines connecting the same # of regions



Expected Market Access Growth (2007–2016), μ_i



Recentered Market Access Growth (2007–2016), \tilde{z}_i

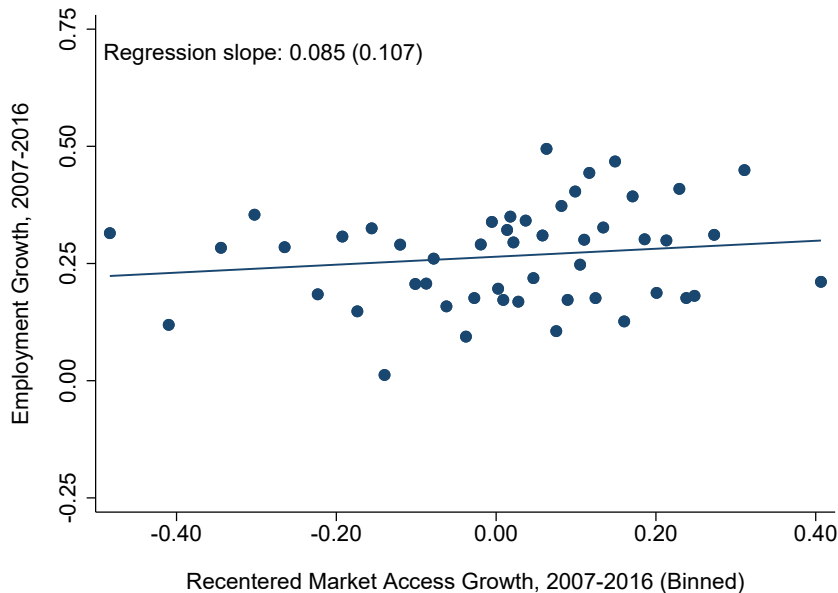


Market Access Balance Regressions

	Unadjusted	Recentered		
	(1)	(2)	(3)	(4)
Distance to Beijing	-0.292 (0.063)	0.069 (0.040)		0.089 (0.045)
Latitude/100	-3.323 (0.648)	-0.325 (0.277)		-0.156 (0.320)
Longitude/100	1.329 (0.460)	0.473 (0.239)		0.425 (0.242)
Expected Market Access Growth			0.027 (0.056)	0.056 (0.066)
Constant	0.536 (0.030)	0.014 (0.018)	0.014 (0.020)	0.014 (0.018)
Joint RI p-value		0.489	0.807	0.536
R^2	0.823	0.079	0.007	0.082
Prefectures	274	274	274	274

Regressions of unadjusted and recentered market access growth on geographic features. Spatial-clustered standard errors in parentheses.

Recentered MA Doesn't Predict Employment Growth!



Adjusted Estimates of Market Access Effects

	Unadjusted OLS (1)	Recentered IV (2)	Controlled OLS (3)
<i>Panel A. No Controls</i>			
Market Access Growth	0.232 (0.075)	0.081 (0.098) [-0.315, 0.328]	0.069 (0.094) [-0.209, 0.331]
Expected Market Access Growth			0.318 (0.095)
<i>Panel B. With Geography Controls</i>			
Market Access Growth	0.132 (0.064)	0.055 (0.089) [-0.144, 0.278]	0.045 (0.092) [-0.154, 0.281]
Expected Market Access Growth			0.213 (0.073)
Recentered	No	Yes	Yes
Prefectures	274	274	274

Regressions of log employment growth on log market access growth in 2007–2016.

Spatial-clustered standard errors in parentheses; permutation-based 95% CI in brackets

App. 2: Efficient Estimation of Medicaid Effects

Setting: U.S. Medicaid, partially expanded in 2014 under the ACA

- 19 of 43 states with low Medicaid coverage expanded to 138% FPL
- View expansion decisions as random across states with same-party governors, but not household demographics or pre-2014 policy
- Outcomes: Medicaid takeup and private insurance crowdout

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- Simulated IV: use state-level variation only (i.e. expansion dummy)
- Recentered IV: predict eligibility from expansion decisions & non-random demographics, and recenter

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We compare two estimators, both valid under the same assumptions:

- Simulated IV: use state-level variation only (i.e. expansion dummy)
- Recentered IV: predict eligibility from expansion decisions & non-random demographics, and recenter

Via non-random variation, recentered IV has ≈ 3 times smaller SEs

Estimates with Simulated vs. Recentered IV

	Has Medicaid		Has Private Insurance		Has Employer-Sponsored Insurance	
	Simulated IV (1)	Recentered IV (2)	Simulated IV (3)	Recentered IV (4)	Simulated IV (5)	Recentered IV (6)
<i>Panel A. Eligibility Effects</i>						
Eligibility	0.132 (0.028) [0.080,0.216]	0.072 (0.010) [0.051,0.093]	-0.048 (0.023) [-0.110,0.009]	-0.023 (0.007) [-0.040,-0.007]	0.009 (0.014) [-0.034,0.052]	-0.009 (0.005) [-0.021,0.004]
<i>Panel B. Enrollment Effects</i>						
Has Medicaid			-0.361 (0.165) [-0.813,0.082]	-0.321 (0.092) [-0.566,-0.108]	0.068 (0.111) [-0.232,0.421]	-0.125 (0.061) [-0.263,0.070]
P-value: SIV=RIV			0.719		0.104	
Exposed Sample	N	Y	N	Y	N	Y
States	43	43	43	43	43	43
Individuals	2,397,313	421,042	2,397,313	421,042	2,397,313	421,042

1% ACS sample of non-disabled adults in 2013–14, diff-in-diff IV regressions using one of the two instruments. Controls include state and year fixed effects and an indicator for Republican governor interacted with year. State-clustered standard errors in parentheses; wild score bootstrap 95% CI in brackets

Roadmap

Motivation

Intuition

- Market Access Effects

- Medicaid Eligibility Effects

Formal Framework

Applications

- Market Access Effects

- Medicaid Eligibility Effects

Concluding Thoughts

Conclusions

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Much more work to be done on the various econometrics here!

Keep Calm and SSIV On!

Good luck on your future adventures with SSIV!

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