Statistics of Solutions to Test Models for SPEKF

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Outline

Introduction

Signals from nature can be modeled by Langevin equation:

Langevin Equation

$$\frac{\mathrm{d}u(t)}{\mathrm{d}t} = -\gamma(t)u(t) + \mathrm{i}\omega u(t) + \sigma \dot{W}(t) + f(t),$$

where $\dot{W}(t)$ is a Brownian motion, and $f(t)$ is the external force.

B. Gershgorin, *et, al* proposed the Stochastic Parameterization Extended Kalman Filter (SPEKF) to cope with model errors.

Introduction

Test Model

$$\begin{cases} \dfrac{du(t)}{dt} = (-\gamma(t) + \mathrm{i}\omega)u(t) + b(t) + f(t) + \sigma W(t), \\[2mm] \dfrac{db(t)}{dt} = (-\gamma_b + \mathrm{i}\omega_b)(b(t) - \hat{b}) + \sigma_b W_b(t), \\[2mm] \dfrac{d\gamma(t)}{dt} = -d_\gamma(\gamma(t) - \hat{\gamma}) + \sigma_\gamma W_\gamma(t) \end{cases}$$

The initial values are complex random variables, with their first-order and second-order statistics known.

Introduction

> **Solution**
>
> With knowledge of ODEs, solution of the SDE set is
>
> $$\begin{cases} b(t) = \hat{b} + (b_0 - \hat{b})e^{\lambda_b(t-t_0)} + \sigma_b \int_{t_0}^t e^{\lambda_b(t-s)}dW_b(s) \\[2mm] \gamma(t) = \hat{\gamma} + (\gamma_0 - \hat{\gamma})e^{-d_\gamma(t-t_0)} + \sigma_\gamma \int_{t_0}^t e^{-d_\gamma(t-s)}dW_\gamma(s) \\[2mm] u(t) = e^{-J(t_0,t)+\hat{\lambda}(t-t_0)}u_0 + \int_{t_0}^t (b(s)+f(s))e^{-J(s,t)+\hat{\lambda}(s-t_0)}ds \\[2mm] \qquad + \sigma \int_{t_0}^t e^{-J(s,t)+\hat{\lambda}(s-t_0)}dW(s) \end{cases}$$

with $\lambda_b = -\gamma_b + i\omega_b$, $\hat{\lambda} = -\hat{\gamma} + i\omega$, $J(s,t) = \int_s^t (\gamma(s') - \hat{\gamma})ds'$.

Introduction

Itô Isometry

$\forall f \in \mathscr{V}(S,T)$, B_t is a standard Brownian motion,

$$\mathrm{E}\left[\left(\int_S^T f(t,\omega)dB_t\right)^2\right] = \mathrm{E}\left[\int_S^T f^2(t,\omega)dt\right].$$

Linear property of Itô integration

(1) $\quad \int_S^T f dB_t = \int_S^U f dB_t + \int_U^T f dB_t$, a.e.

(2) $\quad \int_S^T (cf+g)dB_t = c\int_S^T f dB_t + \int_S^T g dB_t$, a.e.

(3) $\quad \mathrm{E}\left[\int_S^T f dB_t\right] = 0.$

Outline

Mean of $b(t)$, $\gamma(t)$

Solution

$$
\begin{cases}
b(t) = \hat{b} + (b_0 - \hat{b})e^{\lambda_b(t-t_0)} + \sigma_b \int_{t_0}^t e^{\lambda_b(t-s)} dW_b(s) \\[4mm]
\gamma(t) = \hat{\gamma} + (\gamma_0 - \hat{\gamma})e^{-d_\gamma(t-t_0)} + \sigma_\gamma \int_{t_0}^t e^{-d_\gamma(t-s)} dW_\gamma(s)
\end{cases}
$$

With property of Itô integration (3), the last term in $E[b(t)]$ is 0.

Thus we find

$$
E[b(t)] = \hat{b} + (E[b_0] - \hat{b})e^{\lambda_b(t-t_0)}
$$

$$
E[\gamma(t)] = \hat{\gamma} + (E[\gamma_0] - \hat{\gamma})e^{-d_\gamma(t-t_0)}
$$

Variance of $b(t)$, $\gamma(t)$

According to definition,

$$\text{Var}(b(t)) = \text{E}\left[(b(t) - \text{E}[b(t)])(b(t) - \text{E}[b(t)])^*\right]$$

$$= e^{-2\gamma_b(t-t_0)}\text{Var}(b_0) + \text{E}\left[\sigma_b^2 \int_{t_0}^t e^{\lambda_b(t-s)}dW_b(s)\left(\int_{t_0}^t e^{\lambda_b(t-s)}dW_b(s)\right)^*\right]$$

$$= e^{-2\gamma_b(t-t_0)}\text{Var}(b_0) + \sigma_b^2\text{E}\left[\int_{t_0}^t e^{-2\gamma_b(t-s)}ds\right]$$

The last step takes use of Itô isometry. Compute the integration in last term, we have

$$\text{Var}(b(t)) = e^{-2\gamma_b(t-t_0)}\text{Var}(b_0) + \frac{\sigma_b^2}{2\gamma_b}(1 - e^{-2\gamma_b(t-t_0)})$$

$$\text{Var}(\gamma(t)) = e^{-2d_\gamma(t-t_0)}\text{Var}(\gamma_0) + \frac{\sigma_\gamma^2}{2d_\gamma}(1 - e^{-2d_\gamma(t-t_0)})$$

Covariance of $b(t)$, $\gamma(t)$

According to definition,

$$\mathrm{Cov}(b(t), b(t)^*) = \mathrm{E}\left[(b(t) - \mathrm{E}[b(t)])(b(t)^* - \mathrm{E}[b(t)^*])\right]$$

$$= \mathrm{E}\left[(b_0 - \mathrm{E}[b_0])(b_0^* - \mathrm{E}[b_0^*])e^{2\lambda_b(t-t_0)}\right] + \sigma_b \mathrm{E}\left[(b_0 - \mathrm{E}[b_0])\int_{t_0}^t e^{\lambda_b(t-s)}dW_b(s)\right]$$

$$+ \sigma_b \mathrm{E}\left[(b_0^* - \mathrm{E}[b_0^*])e^{\lambda_b(t-t_0)}\int_{t_0}^t e^{\lambda_b(t-s)}dW_b(s)\right] + \sigma_b^2 \mathrm{E}\left[(\int_{t_0}^t e^{\lambda_b(t-s)}dW_b(s))^2\right]$$

With property of Itô integration (3), the second and third term are both 0; with Itô isometry we know the last term is also 0. Accordingly,

$$\mathrm{Cov}(b(t), b(t)^*) = \mathrm{E}[(b_0 - \mathrm{E}[b_0])(b_0^* - \mathrm{E}[b_0^*])]e^{2\lambda_b(t-t_0)} = \mathrm{Cov}(b_0, b_0^*)e^{2\lambda_b(t-t_0)}$$

$$\mathrm{Cov}(b(t), \gamma(t)) = \mathrm{E}[(b(t) - \mathrm{E}[b(t)])(\gamma(t) - \mathrm{E}[\gamma(t)])] = \mathrm{Cov}(b_0, \gamma_0)e^{(\lambda_b - d_\gamma)(t-t_0)}$$

Outline

Mean of $u(t)$

Using the same properties, it is easy to find

$$\mathrm{E}[u(t)] = e^{\hat{\lambda}(t-t_0)} \mathrm{E}\left[e^{-J_0(t_0,t)} u_0\right] + \int_{t_0}^{t} e^{\hat{\lambda}(t-s)} \mathrm{E}\left[b(s) e^{-J(s,t)}\right] ds$$

$$+ \sigma \int_{t_0}^{t} e^{\hat{\lambda}(t-s)} f(s) \mathrm{E}\left[e^{-J(s,t)}\right] ds.$$

We find it necessary to compute expectations of terms like

$$\mathrm{E}[z e^{bx}],$$

where z is a complex-valued Gaussian random variable and x is a real-valued Gaussian variable. We propose two lemmas here.

Mean of $u(t)$

Lemma (1)

$$E\left[ze^{ibx}\right] = (E[z] + ibCov(z,x))e^{ibE[x]-\frac{1}{2}b^2Var(x)},$$

with z being a complex-valued Gaussian, and x a real-valued Gaussian.

Corollary (1)

Under the condition of Lemma 1,

$$E\left[ze^{bx}\right] = (E[z] + bCov(z,x))e^{bE[x]+\frac{1}{2}b^2Var(x)}.$$

Proof of lemma 1 takes advantage of characteristic function of multivariable Gaussian distribution.

Proof

Let $z = y + \mathrm{i}w$, $y, w \in \mathbb{R}$. Denote $\mathbf{v} = (x, y, w)^\top$, then $\mathbf{v}$ satisfies the multivariable Gaussian distribution, with its characteristic function

$$\phi_{\mathbf{v}}(\mathbf{s}) = \exp(\mathrm{i}\mathbf{s}^\top \mathrm{E}[\mathbf{v}] - \frac{1}{2}\mathbf{s}^\top \Sigma \mathbf{s}).$$

Let $g(\mathbf{v})$ being the PDF of $\mathbf{v}$, then one knows from that char. func. being Fourier transform of PDF,

$$\phi_{\mathbf{v}}(\mathbf{s}) = \frac{1}{(2\pi)^3} \int e^{\mathrm{i}\mathbf{s}^\top \mathbf{v}} g(\mathbf{v}) d\mathbf{v}. \text{`}'$$

According to the differential property of Fourier transform (Proposition 2.10),

$$\frac{\partial \phi_{\mathbf{v}}(\mathbf{s})}{\partial s_2} = \frac{1}{(2\pi)^3} \int \mathrm{i}y_0 e^{\mathrm{i}\mathbf{s}^\top \mathbf{v}} g(\mathbf{v}) d\mathbf{v} = \mathrm{i}\mathrm{E}\left[y_0 e^{\mathrm{i}\mathbf{s}^\top \mathbf{v}} \right].$$

Let $\mathbf{v} = (b, 0, 0)^\top$,

$$\mathrm{E}[y_0 e^{\mathrm{i}bx_0}] = -\mathrm{i}\frac{\partial \phi_{\mathbf{v}}(s)}{\partial s_2}\bigg|_{\mathbf{s}=(b,0,0)^\top}, \quad \mathrm{E}[w_0 e^{\mathrm{i}bx_0}] = -\mathrm{i}\frac{\partial \phi_{\mathbf{v}}(s)}{\partial s_3}\bigg|_{\mathbf{s}=(b,0,0)^\top}$$

Proof

From PDF of multivariable Gaussian distribution, one knows

$$\frac{\partial \phi_{\mathbf{v}}(\mathbf{s})}{\partial s_2} = (\mathrm{i}\mathrm{E}[y_0] - \mathrm{Var}(y_0)s_2 - \mathrm{Cov}(x_0, y_0)s_1 - \mathrm{Cov}(y_0, w_0)s_3)\phi_{\mathbf{v}}(s)$$

$$\frac{\partial \phi_{\mathbf{v}}(\mathbf{s})}{\partial s_3} = (\mathrm{i}\mathrm{E}[w_0] - \mathrm{Var}(w_0)s_3 - \mathrm{Cov}(x_0, w_0)s_1 - \mathrm{Cov}(y_0, w_0)s_2)\phi_{\mathbf{v}}(s)$$

Compute the partial derivatives at $\mathbf{s} = (b, 0, 0)^{\top}$,

$$\mathrm{E}\left[y_0 e^{\mathrm{i}b x_0}\right] = (\mathrm{E}[y_0] + \mathrm{i}\mathrm{Cov}(x_0, y_0)b)\exp(\mathrm{i}b\mathrm{E}[x_0] - \frac{1}{2}\mathrm{Var}(x_0)b^2)$$

$$\mathrm{E}\left[w_0 e^{\mathrm{i}b x_0}\right] = (\mathrm{E}[w_0] + \mathrm{i}\mathrm{Cov}(x_0, w_0)b)\exp(\mathrm{i}b\mathrm{E}[x_0] - \frac{1}{2}\mathrm{Var}(x_0)b^2)$$

Then

$$\mathrm{E}\left[z e^{\mathrm{i}b x}\right] = (\mathrm{E}[z] + \mathrm{i}b\mathrm{Cov}(z, x))e^{\mathrm{i}b\mathrm{E}[x] - \frac{1}{2}b^2\mathrm{Var}(x)}.$$

$\square$

Mean of $u(t)$

Lemma (2)

$$E\left[zwe^{bx}\right] = [E[z]E[w] + Cov(z, w^*) + b(E[z]Cov(w, x)) + E[w]Cov(z, x) +$$

$$b^2 Cov(z, x) Cov(w, x)] e^{bE[x] + \frac{b^2}{2} Var(x)}.$$

with z, w being complex-valued Gaussian, and x real-valued Gaussian.

The proof of this lemma is the same as Lemma 1.

Mean of $u(t)$

We now make use of Lemma 1 to obtain mean of $u(t)$.

$$\mathrm{E}[u(t)] = e^{\hat{\lambda}(t-t_0)}(\mathrm{E}[u_0] - \mathrm{Cov}(u_0, J(t_0, t)))e^{-\mathrm{E}[J(t_0,t)] + \frac{1}{2}\mathrm{Var}(J(t_0,t))}$$

$$+ \int_{t_0}^t e^{\hat{\lambda}(t-s)}(\hat{b} + e^{\lambda_b(s-t_0)}(\mathrm{E}[b_0] - \hat{b} - \mathrm{Cov}(b_0, J(s,t))))e^{-\mathrm{E}[J(s,t)] + \frac{1}{2}\mathrm{Var}(J(s,t))}ds$$

$$+ \int_{t_0}^t e^{\hat{\lambda}(t-s)}f(s)e^{-\mathrm{E}[J(s,t)] + \frac{1}{2}\mathrm{Var}(J(s,t))}ds$$

The terms $\mathrm{Cov}(u_0, J(s,t))$, $\mathrm{Cov}(b_0, J(s,t))$, $\mathrm{E}[J(s,t)]$ and $\mathrm{Var}(J(s,t))$ can be found using Itô isometry.

Variance of $u(t)$

Denote $u(t) = A + B + C$,

$$\begin{cases} A = e^{-J(t_0,t)+\hat{\lambda}(t-t_0)}u_0, \\[2mm] B = \displaystyle\int_{t_0}^{t}(b(s)+f(s))e^{-J(s,t)+\hat{\lambda}(t-s)}ds, \\[2mm] C = \sigma\displaystyle\int_{t_0}^{t}e^{-J(s,t)+\hat{\lambda}(t-s)}dW(s). \end{cases}$$

By definition we find $\mathrm{Var}(u(t)) = \mathrm{E}\left[|u(t)|^2\right] - |\mathrm{E}[u(t)]|^2$, with

$$\mathrm{E}\left[|u(t)|^2\right] = \mathrm{E}\left[|A|^2\right] + \mathrm{E}\left[|B|^2\right] + \mathrm{E}\left[|C|^2\right] + 2\mathrm{Re}\{\mathrm{E}[A^*B]\}.$$

We can obtain $\mathrm{E}\left[|A|^2\right]$ by Lemma 2, and $\mathrm{E}\left[|B|^2\right]$ by Itô isometry. Noticing that

$$\mathrm{Cov}(J(s,t),J(r,t)) = \mathrm{Var}(J(s,t)) + \mathrm{Cov}(J(s,t),J(r,s)).$$

$\mathrm{E}\left[|C|^2\right]$ and $\mathrm{Re}\{\mathrm{E}[A^*B]\}$ can also be computed by Itô isometry and property of Itô integration.

Covariances

By definition,

$$\text{Cov}(u(t), u^*(t)) = \text{E}\left[u(t)^2\right] - \text{E}[u(t)]^2$$

$$\text{Cov}(u(t), \gamma(t)) = \text{E}[u(t)(\gamma(t) - \hat{\gamma})] + \text{E}[u(t)](\hat{\gamma} - \text{E}[\gamma(t)])$$

$$\text{Cov}(u(t), b(t)) = \text{E}[u(t)b^*(t)] - \text{E}[u(t)]\text{E}[b(t)]^*$$

$$\text{Cov}(u(t), b^*(t)) = \text{E}[u(t)b(t)] - \text{E}[u(t)]\text{E}[b(t)].$$

Each term can be obtained by Lemma 3 and Itô isometry.

Outline

Parameters

As a real-world example, we choose a periodic external force

$$f(t) = \frac{3}{2} e^{\frac{1}{10}it},$$

and parameters of the equation set are given

$$\begin{cases} d = 1.5, & d_\gamma = 0.01d \\ \sigma = 0.1549, & \omega = 1.78 \\ \sigma_\gamma = 5\sigma, & \gamma_b = 0.1d \\ \sigma_b = 5\sigma, & \omega_b = \omega \\ \hat{b} = 0, & \hat{\gamma} = 0 \end{cases}$$

We assume that initial values satisfy

$$\begin{cases} \mathrm{Re}(u_0), \mathrm{Im}(u_0) \sim \mathscr{N}(0, \frac{1}{\sqrt{2}}), \text{ i.i.d.} \\ \mathrm{Re}(b_0), \mathrm{Im}(b_0) \sim \mathscr{N}(0, \frac{1}{\sqrt{2}}), \text{ i.i.d.} \\ \gamma_0 \sim \mathscr{N}(0, 1) \end{cases}$$

Euler-Maruyama Scheme

Itô integration can be simulated by E-M scheme:

$$X_j = X_{j-1} + f(X_{j-1})\Delta t + g(X_{t-1})(W(\tau_j) - W(\tau_{j-1})),$$

with

$$W(\tau_j) - W(\tau_{j-1}) = \sum_{k=jR-R+1}^{jR} \mathrm{d}W_k,$$

and R being the step length of E-M scheme,

$$\mathrm{d}W = \sqrt{\Delta t} \times \texttt{randn}().$$

Revised Euler Scheme

Since we need to simulate a SDE set with three relevant equations, we choose the Revised Euler Scheme from numerical solution to ODE for simulating $u(t)$:

$$x_n = x_{n-1} + \Delta t \, \phi\left(t_{n-1} + \frac{\Delta t}{2}, \, x_{n-1} + \frac{\Delta t}{2} \phi(t_{n-1}, \, x_{n-1})\right),$$

it is a second-order Runge-Kutta scheme with a locally quadratic convergence.

Result

Simulate 10^6 times with $R = 1$, the results are as follows:

(a) E(Re(u))

(b) E(Im(u))

(c) E(Re(b))

(d) E(Im(b))

(e) E(γ)

(f) Var(u)

(g) Var(b)

(h) Var(γ)
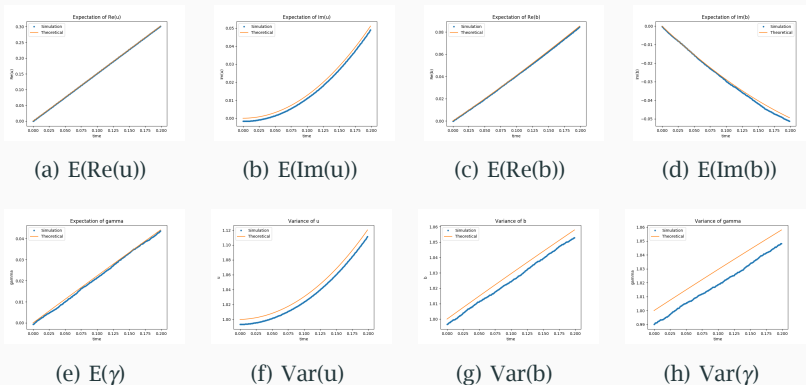
Figure 1: Simulation of Expectations and Variances, $n = 10^6$

Result
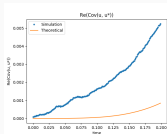
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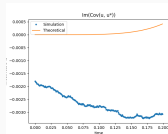
(b) Im(Cov(u, u^*))
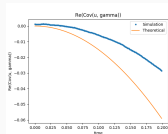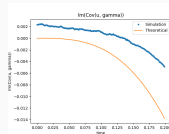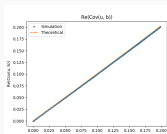
(c) Re(Cov(u, γ))
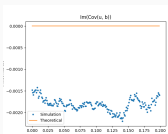
(d) Im(Cov(u, γ))
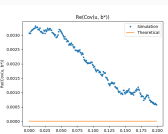
(e) Re(Cov(u, b))
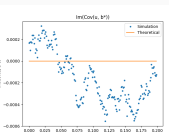
(f) Im(Cov(u, b))

(g) Re(Cov(u, b^*))

(h) Im(Cov(u, b^*))

Figure 2: Simulation of Covariances, $n = 10^6$

Result

Simulate 10^7 times with $R = 1$, the results are as follows:

(a) E(Re(u))

(b) E(Im(u))

(c) E(Re(b))

(d) E(Im(b))

(e) E(γ)

(f) Var(u)

(g) Var(b)

(h) Var(γ)
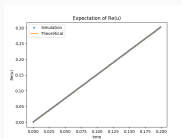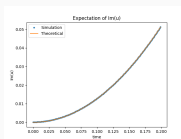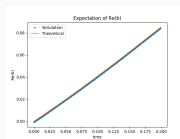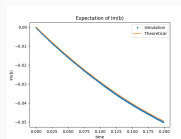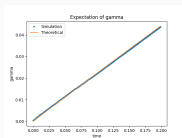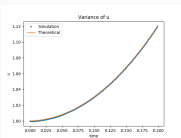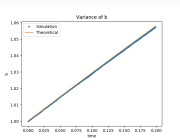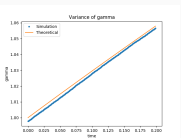
Figure 3: Simulation of Expectations and Variances, $n = 10^7$

Result
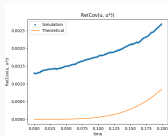
(a) Re(Cov(u, u^*))   (b) Im(Cov(u, u^*))   (c) Re(Cov(u, γ))   (d) Im(Cov(u, γ))
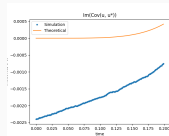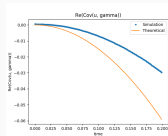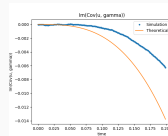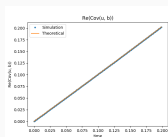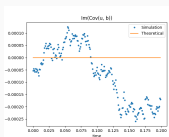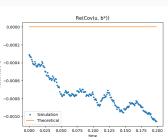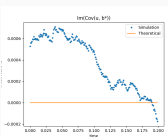
(e) Re(Cov(u, b))   (f) Im(Cov(u, b))   (g) Re(Cov(u, b^*))   (h) Im(Cov(u, b^*))

Figure 4: Simulation of Covariances, $n = 10^7$

Discussion

From the simulation results we find that the simulation of expectations fit the theoretical results satisfyingly.

For the results of Variances and Covariances,

- When $n = 10^6$, the error of variances is about $O(10^{-2}) \sim O(10^{-3})$, error of covariances is about $O(10^{-2}) \sim O(10^{-3})$.

- When $n = 10^7$, the error of $\text{Var}(u)$, $\text{Var}(b)$ is about $O(10^{-5})$, and error of $\text{Var}(\gamma)$ is about $O(10^{-3}) \sim O(10^{-4})$, error of $\text{Cov}(u, u^*)$, $\text{Cov}(u, \gamma)$ is about $O(10^{-3})$, error of $\text{Cov}(u, b)$, $\text{Cov}(u, b^*)$ is $O(10^{-4})$.

Discussion

One can find the simulation errors decrease with the growth of simulations. Thus one could believe that when $n \to \infty$, the simulation would converge to the theoretical result. In fact, error of simulation comes from that initial values are all random variables; according to Law of large numbers, the errors would converge to 0 when $n \to \infty$.

We can get third-order and fourth-order statistics of u, b, γ with the same methods, which could be written in a form of multiple integrals of statistics of initial values.

Thank you!