



Rafael S. de Souza,

in collaboration with E. E. O. Ishida

Researcher at the University of North Carolina at Chapel Hill

Chapel Hill, USA

Astrophysics + Statistics + Machine Learning

Interdisciplinary science development

Chair of the Cosmostatistics Initiative (COIN)

Day 1: Overview

1. Interpretation

Bayesian x Frequentists

2. Parameter inference

Markov Chain Monte Carlo

3. Language/tools

JAGS, ...

https://github.com/RafaelSdeSouza/Bayes ESTEC

or

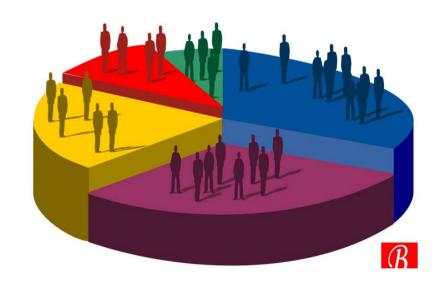
https://rafaelsdesouza.github.io/Bayes ESTEC/

"Begin at the beginning,"
the King said, very gravely,
"and go on till you come to the end:
then stop".

New Latin: Statisticum collegium

Italian: statista ("statesman")

German: Statistik (science of state)



Taxes

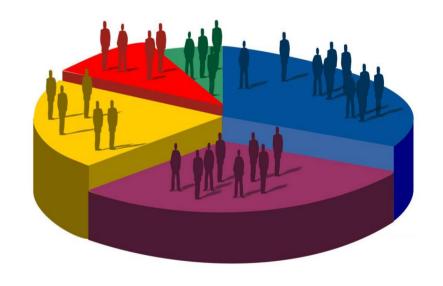
Medicine

Statistics ← State

Mean

Variance

Frequency



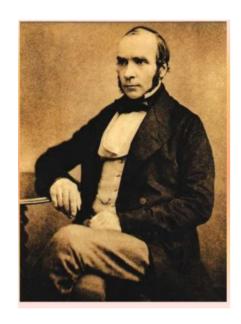
Taxes

Medicine

Statistics State **Taxes** Medicine

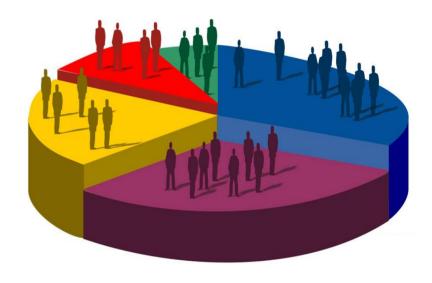
Statistics





John Snow (1813 - 1858) The father of epidemiology

Used data to prove that cholera was transmitted via contaminated water

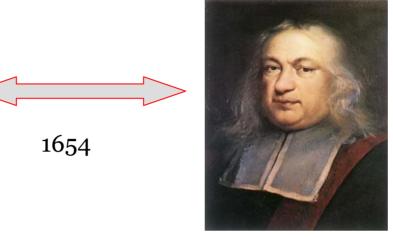


Taxes

Medicine







Pierre de Fermat



APS NEWS

July 2009 (Volume 18, Number 7)

This Month in Physics History

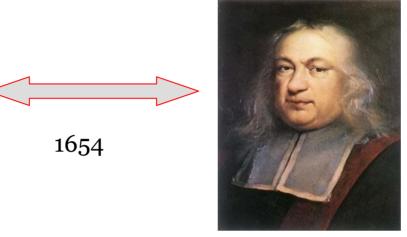
July 1654: Pascal's Letters to Fermat on the "Problem of Points"

Atoine Gambaud asked for help in games of chance for Pascal.

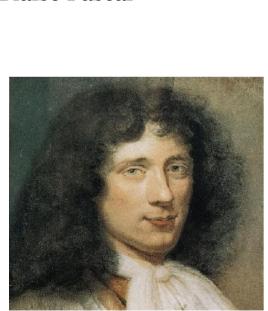
Gambling led, indirectly, to the birth of probability theory. In the mid-17th century, an exchange of letters between Blaise Pascal and Pierre de Fermat–laid the foundation for probability.







Pierre de Fermat

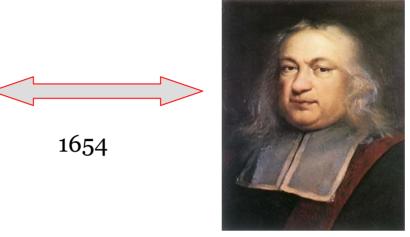


Christian Huygens

On Reasoning in Games of Chance, 1657



Blaise Pascal



Pierre de Fermat





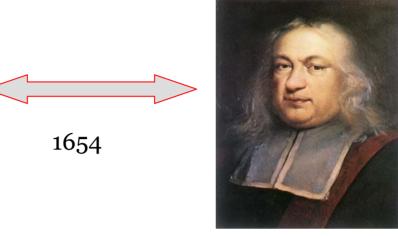
Christian Huygens



On Reasoning in Games of Chance, 1657







Pierre de Fermat





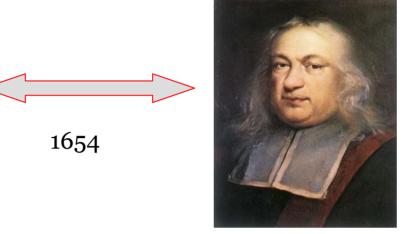
Christian Huygens

On Reasoning in Games of Chance, 1657

Probability





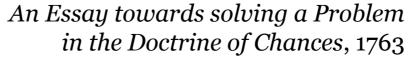


Pierre de Fermat



Christian Huygens

On Reasoning in Games of Chance, 1657





Thomas Bayes

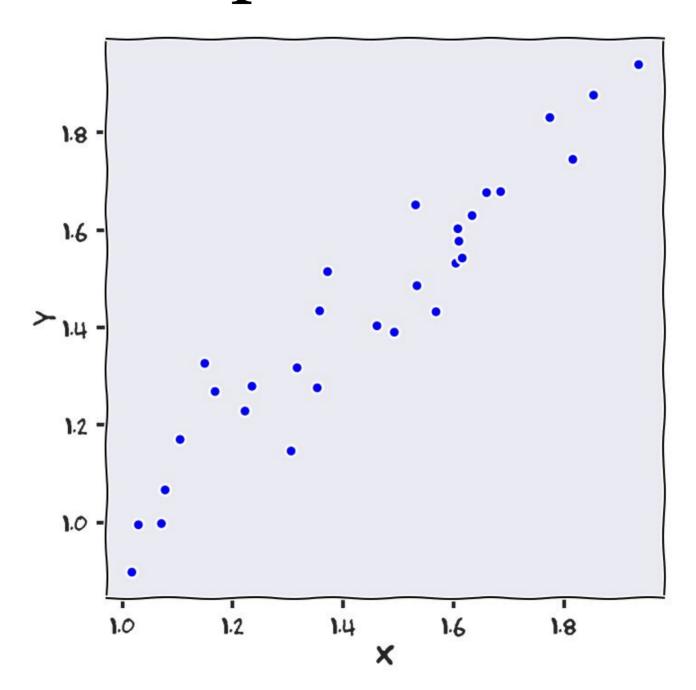
What is a probability?

Frequentist	Bayesian
Probability is a long-run average	Probability is a degree of belief
There is a true Model, the Data is a random realization	The Data is true/fixed, Models have probabilities
Probability of the data given a hypothesis (Likelihood)	Probability of a hypothesis given the data
Each repeated experiment/observation starts from ignorance	Can incorporate prior knowledge: probabilities can be updated

Jerzy Neyman

Harold Jeffereys

A first example:



$$y = a x + b$$

Ordinary Least Squares:

$$\{a,b\} \leftarrow \min \left[\sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2} \right]$$

$$y = a x + b$$

Ordinary Least Squares:

$$\{a,b\} \leftarrow \min \left[\sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2} \right]$$

Main assumptions:

- independent observations (iid) no correlations
- all the points are equally valid no outliers
- errors are Gaussian distributed homoscedasticity
- no exterior information no priors

$$y = a x + b$$
 "Physical" model

Ordinary Least Squares:

$$\{a,b\} \leftarrow \min \left[\sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2} \right]$$

Main assumptions:

- independent observations (iid) no correlations
- all the points are equally valid no outliers
- errors are Gaussian distributed homoscedasticity
- no exterior information no priors

Statistical model

y = ax + b

Model of the mean

Ordinary Least Squares:

$$\{a,b\} \leftarrow \min \left[\sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2} \right]$$

Main assumptions:

- independent observations (iid) no correlations
- all the points are equally valid no outliers
- errors are Gaussian distributed homoscedasticity
- no exterior information no priors

Model for the uncertainties

If $\{X, Y\}$ are random variables:

$$X \sim Uniform(1.0, 2.0)$$

 $Y = aX + b + \varepsilon$
 $\varepsilon \sim Normal(0, \sigma^2)$

or

 $Y \sim Normal(aX + b, \sigma^2)$

If
$$\{X, Y\}$$
 are random variables:

$$X \sim Uniform(1.0, 2.0)$$

$$Y = aX + b + \varepsilon$$

$$\varepsilon \sim Normal(0, \sigma^2)$$

$$Y \sim Normal(aX + b, \sigma^2)$$

$$\mathcal{L}(x,y|a,b,\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(y-(ax+b))^2}{2\sigma^2}\right]$$

If $\{X, Y\}$ are random variables:

$$X \sim Uniform(1.0, 2.0)$$

$$Y = aX + b + \varepsilon$$

$$\varepsilon \sim Normal(0, \sigma^2)$$

or

$$Y \sim Normal(aX + b, \sigma^2)$$

Likelihood for 1 point:

$$\mathcal{L}(x,y|a,b,\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(y-(ax+b))^2}{2\sigma^2}\right]$$

Likelihood for the complete data set:

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(y_i - (ax_i + b))^2}{2\sigma^2}\right]$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$

If $\{X, Y\}$ are random variables:

$$X \sim Uniform(1.0, 2.0)$$

$$Y = aX + b + \varepsilon$$

$$\varepsilon \sim Normal(0, \sigma)$$

or

$$Y \sim Normal(aX + b, \sigma)$$

Likelihood for 1 point:

$$\mathcal{L}(x,y|a,b,\sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(y-(ax+b))^2}{2\sigma^2}\right]$$

Likelihood for the complete data set:

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(y_i - (ax_i + b))^2}{2\sigma^2}\right]$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$

Maximum Likelihood

$$\theta_{\mathrm{ML}} \equiv \max_{\theta} \mathscr{L}(\theta)$$

If Gaussian

$$\mathscr{L}(\mu, \sigma) = p(\hat{x}|\mu, \sigma) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi}\sigma} \exp\left(-\frac{1}{2} \frac{(\hat{x}_i - \mu)^2}{\sigma^2}\right)$$

$$\mathscr{L} = L_0 \exp\left(-\chi^2/2\right)$$

$$\chi^2 = \sum_{i=1}^N \frac{(\hat{x}_i - \mu)^2}{\sigma^2}$$

Confidence intervals

Taylor series expansion around maximum of a general likelihood function:

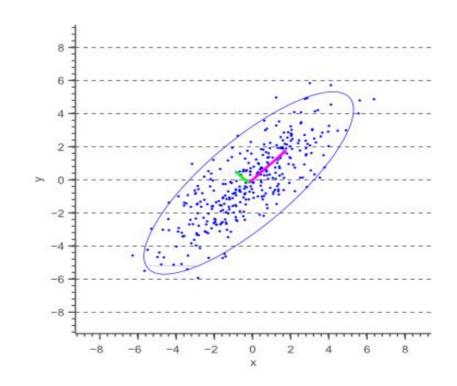
$$\ln \mathcal{L}(\theta) = \ln \mathcal{L}|_{\hat{\theta}} + (\theta - b) \frac{\partial \ln \mathcal{L}}{\partial \theta} \Big|_{\hat{\theta}} + \frac{1}{2} (\theta - \hat{\theta})^2 \frac{\partial^2 \ln \mathcal{L}}{\partial \theta^2} \Big|_{\hat{\theta}} + \dots \qquad \frac{\partial \ln \mathcal{L}}{\partial a} \Big|_{\hat{\theta}} = \frac{\partial \ln \mathcal{L}}{\partial b} \Big|_{\hat{\theta}} = 0$$

A general likelihood function can be approximated to second order as a Gaussian around the ML value

$$\mathcal{L}(\theta) \approx \mathcal{L}(\hat{\theta}) \exp\left(-\frac{1}{2} \frac{(\theta - \hat{\theta})^2}{C_{\hat{\theta}}}\right)$$

$$C_{\hat{\theta}} = F^{-1} = \left(-\frac{\partial^2 \ln \mathcal{L}}{\partial \theta^2} \Big|_{\hat{\theta}} \right)^{-1}$$

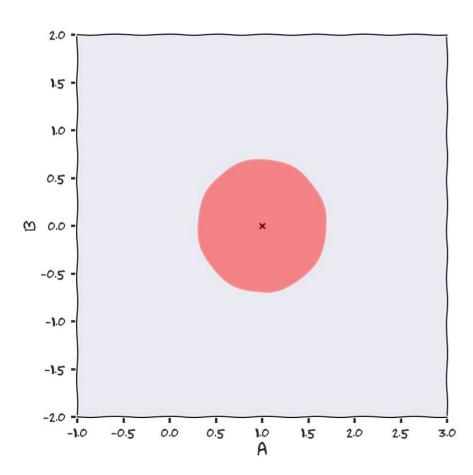
Fisher Matrix



$$y = a x + b$$

Frequentist:

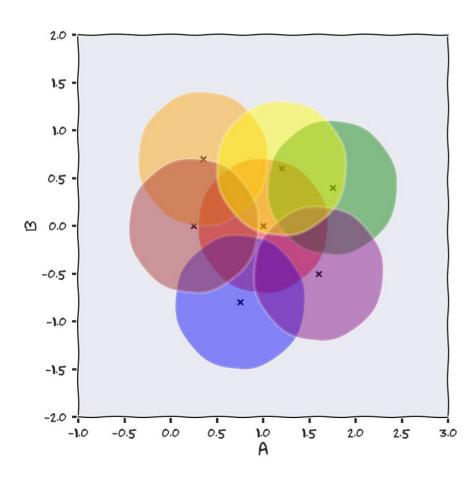
$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$



y = a x + b

Frequentist:

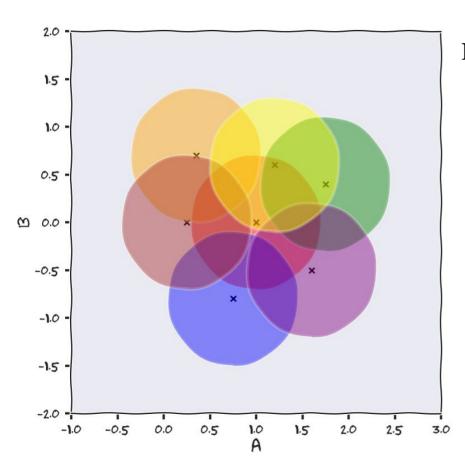
$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$



y = a x + b

Frequentist:

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$



Frequentist:

95% confidence interval → 95% of the intervals derived from possible data will contain the true value

If $\{X, Y\}$ are random variables:

$$X \sim Uniform(1.0, 2.0)$$

$$Y = aX + b + \varepsilon$$

$$\varepsilon \sim Normal(0, \sigma^2)$$

Y ~ Normal(aX + b,
$$\sigma^2$$
)
+
a ~ Normal(0, 5.0)
b ~ Normal(0, 5.0)

If {X, Y} are random variables:

Bayes theorem

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$

$$X \sim Uniform(1.0, 2.0)$$

$$Y = aX + b + \varepsilon$$

$$\varepsilon \sim Normal(0, \sigma^2)$$

Y ~ Normal(aX + b,
$$\sigma^2$$
)
+
a ~ Normal(0, 5.0)
b ~ Normal(0, 5.0)

If {X, Y} are random variables:

Bayes theorem

$$P(\vec{\theta}|\mathcal{D}) = \frac{P(\mathcal{D}|\vec{\theta})P(\vec{\theta})}{P(\mathcal{D})}$$

$$X \sim Uniform(1.0, 2.0)$$

 $Y = aX + b + \varepsilon$
 $\varepsilon \sim Normal(0, \sigma^2)$

Y ~ Normal(aX + b,
$$\sigma^2$$
)
+
a ~ Normal(0, 5.0)
b ~ Normal(0, 5.0)

If {X, Y} are random variables:

Bayes theorem

$$P(a, b | \vec{x}, \vec{y}, \sigma) = \frac{\mathcal{L}(\vec{x}, \vec{y} | a, b, \sigma) p(a) p(b)}{E(\vec{x}, \vec{y})}$$

$$X \sim Uniform(1.0, 2.0)$$

 $Y = aX + b + \varepsilon$
 $\varepsilon \sim Normal(0, \sigma^2)$

Y ~ Normal(aX + b,
$$\sigma^2$$
)
+

If $\{X, Y\}$ are random variables:

Bayes theorem

$$P(a, b|\vec{x}, \vec{y}, \sigma) = \frac{\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma)p(a)p(b)}{\int_{a} \int_{b} \mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma)p(a)p(b)dadb}$$

$$X \sim Uniform(1.0, 2.0)$$

$$Y = aX + b + \varepsilon$$

$$\varepsilon \sim Normal(0, \sigma^2)$$

or

Y ~ Normal(aX + b,
$$\sigma^2$$
)
+
a ~ Normal(0, 5.0)

 $b \sim Normal(0, 5.0)$

If {X, Y} are random variables:

Bayes theorem

$$P(a,b|\vec{x},\vec{y},\sigma) = \frac{\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)}{\int_a \int_b \mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)dadb}$$

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$

$$X \sim Uniform(1.0, 2.0)$$

 $Y = aX + b + \varepsilon$
 $\varepsilon \sim Normal(0, \sigma^2)$

Y ~ Normal(aX + b,
$$\sigma^2$$
)
+
a ~ Normal(0, 5.0)
b ~ Normal(0, 5.0)

If {X, Y} are random variables:

Bayes theorem

$$P(a,b|\vec{x},\vec{y},\sigma) = \frac{\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)}{\int_a \int_b \mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)dadb}$$

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$

$$p(a) = \frac{1}{\sqrt{2\pi(5)^2}} \exp\left[-\frac{(a-0)^2}{2\times(5^2)}\right]$$

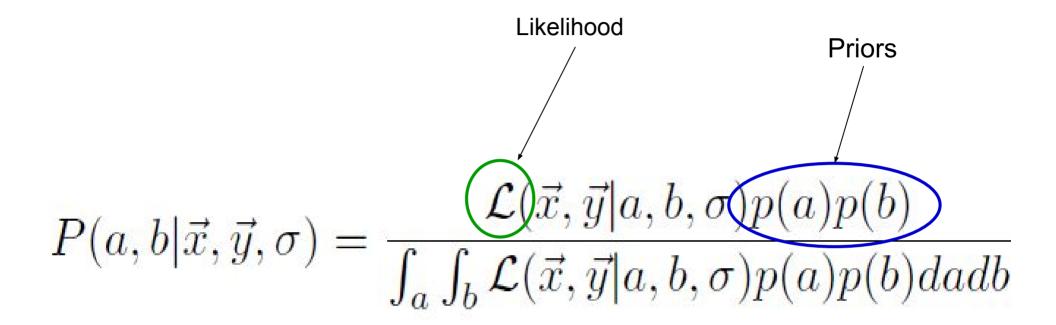
$$p(b) = \frac{1}{\sqrt{2\pi(5)^2}} \exp\left[-\frac{(b-0)^2}{2\times(5^2)}\right]$$

 $X \sim Uniform(1.0, 2.0)$ $Y = aX + b + \varepsilon$ $\varepsilon \sim Normal(0, \sigma^2)$

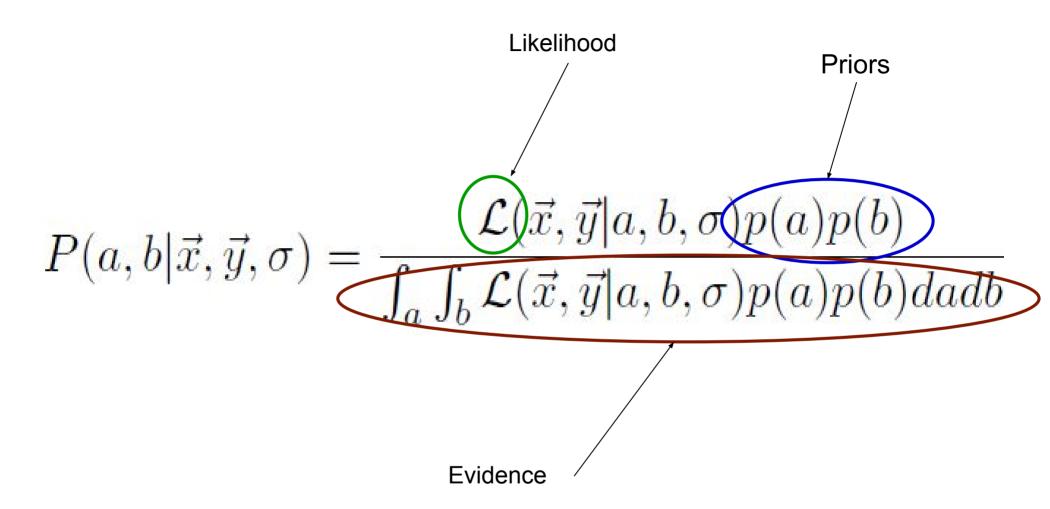
or

Y ~ Normal(aX + b, σ^2) + a ~ Normal(0, 5.0) b ~ Normal(0, 5.0)

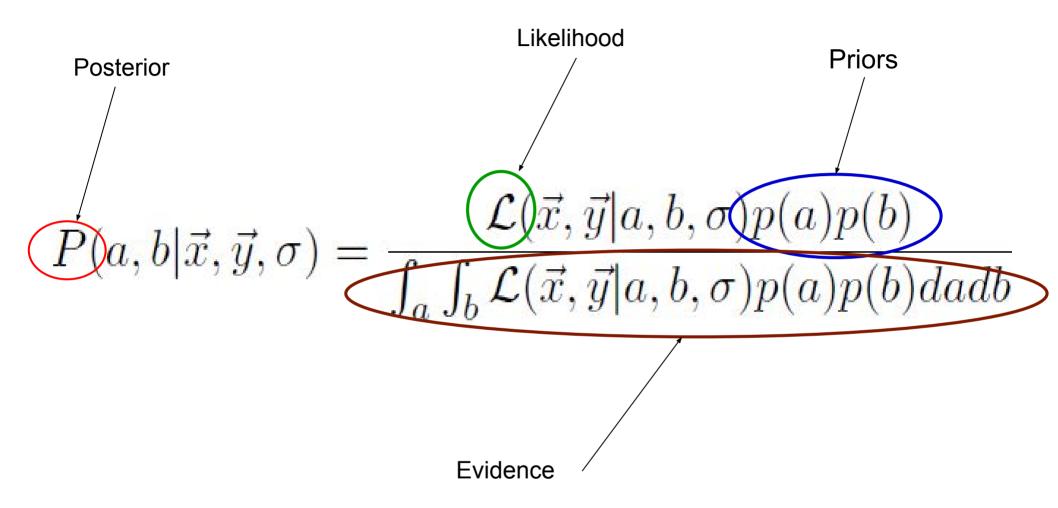
$$P(a,b|\vec{x},\vec{y},\sigma) = \frac{\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)}{\int_a \int_b \mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)dadb}$$



Bayesian approach



Bayesian approach



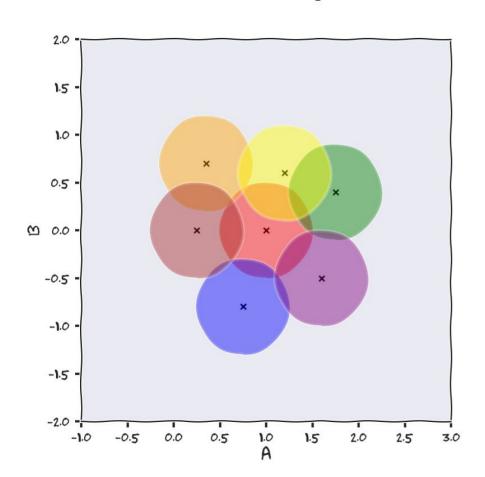
Bayesian approach

$$P(\boxed{a,b|\vec{x},\vec{y},\sigma}) = \frac{\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)}{\int_a \int_b \mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)dadb}$$

y = a x + b

Frequentist:

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$

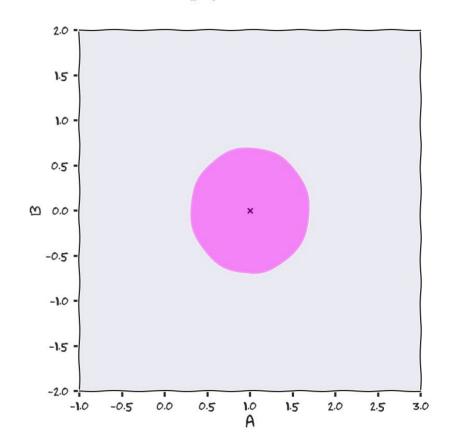


Frequentist:

95% confidence interval →
95% of the intervals derived from possible data
will contain the true value

Bayesian:

$$P(a,b|\vec{x},\vec{y},\sigma) = \frac{\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)}{\int_a \int_b \mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)dadb}$$



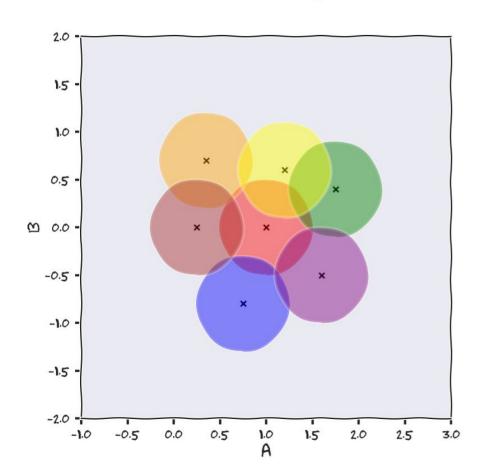
Bayesian:

95% credible interval →
There is a 95% chance that the true
value is within this interval

y = a x + b

Frequentist:

$$\mathcal{L}(\vec{x}, \vec{y}|a, b, \sigma) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{1}{2} \sum_{i=1}^{N} \frac{(y_i - (ax_i + b))^2}{\sigma^2}\right]$$

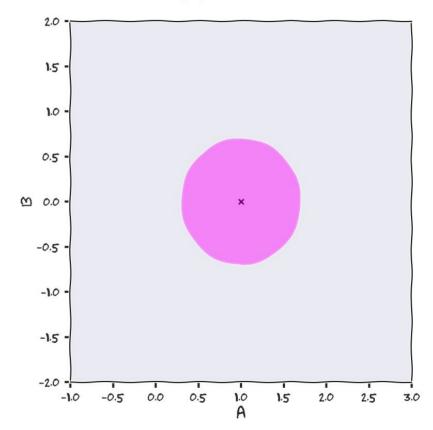


Frequentist:

95% confidence interval →
95% of the intervals derived from possible data
will contain the true value

Bayesian:

$$P(a,b|\vec{x},\vec{y},\sigma) = \frac{\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)}{\int_{a}\int_{b}\mathcal{L}(\vec{x},\vec{y}|a,b,\sigma)p(a)p(b)dadb}$$

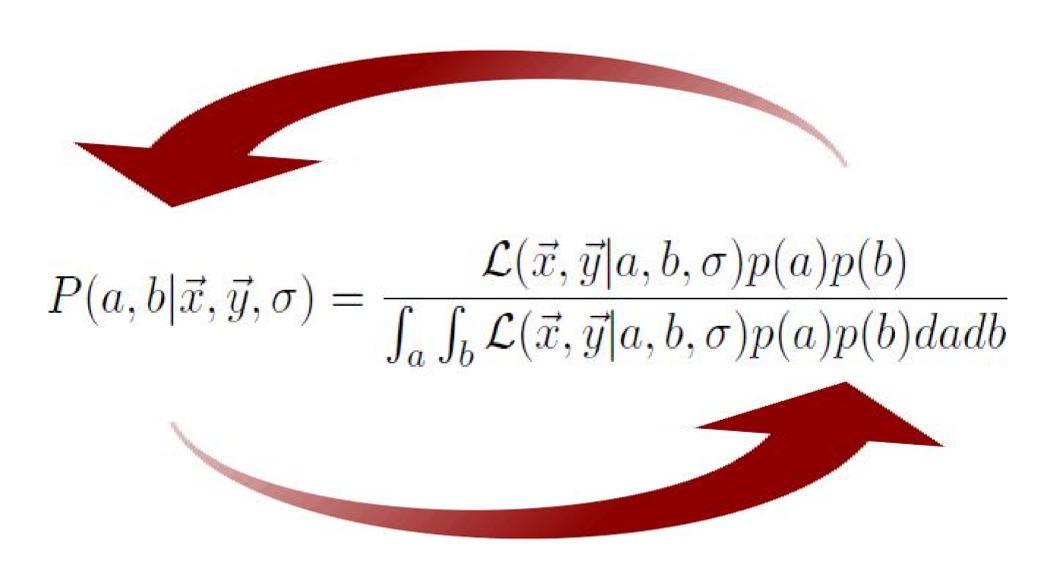


Bayesian:

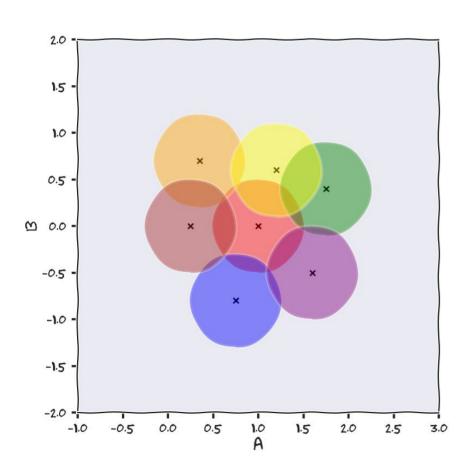
95% credible interval →
There is a 95% chance that the true
value is within this interval

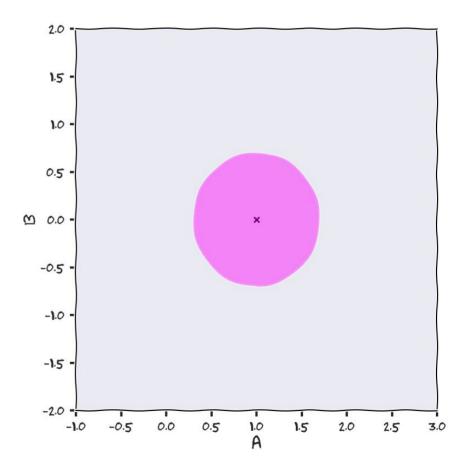
belief

The Bayesian approach is a process...



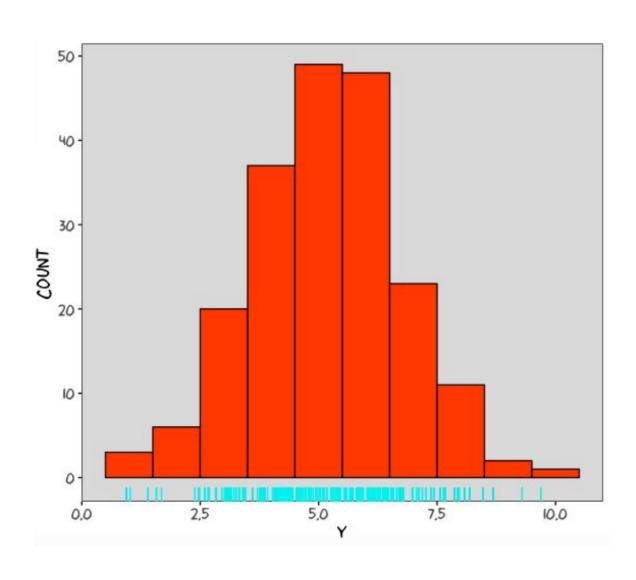
Confidence vs bayesian Intervals, practical examples





The mean of a Gaussian

 $Y \sim Normal(mu = 5, sd = 1.5)$



Frequentist Approach

Unbiased estimator of the mean:

$$\bar{x} = \frac{1}{N} \sum_{i=1}^{N} x_i$$

The **sampling distribution of** observed frequency of the x

$$f(\bar{x} \mid\mid \mu) \propto \exp\left[\frac{-(\bar{x} - \mu)^2}{2\sigma_{\mu}^2}\right]$$

Standard error of the mean,

$$\sigma_{\mu} = \sigma_{x}/\sqrt{N}$$

The 95% confidence interval (two standard deviations), which roughly covers 95% of the area $CI_{\mu}=\left(\bar{x}-2\sigma_{\mu},\ \bar{x}+2\sigma_{\mu}\right)$ under the Gaussian curve.

Bayesian Approach

Bayes' theorem:

$$P(\mu \mid D) = \frac{P(D \mid \mu)P(\mu)}{P(D)}$$

Prior:

$$P(\mu) \propto 1$$

Likelihood:

$$P(D \mid \mu) = \prod_{i=1}^{N} \frac{1}{\sqrt{2\pi\sigma_x^2}} \exp\left[\frac{(\mu - x_i)^2}{2\sigma_x^2}\right]$$

Posterior:

$$P(\mu \mid D) \propto \exp\left[\frac{-(\mu - \bar{x})^2}{2\sigma_u^2}\right]$$

The 95% credible interval, the shortest interval that contains 95% of the probability.

$$CR_{\mu} = \left(\bar{x} - 2\sigma_{\mu}, \ \bar{x} + 2\sigma_{\mu}\right)$$

Example II: Truncated exponential

The Enterprise force field operates without failure until depletion of dilithium.

If under attack by Klingons, full depletion may occur after a time θ , after which failures may happen following the exponential failure law.

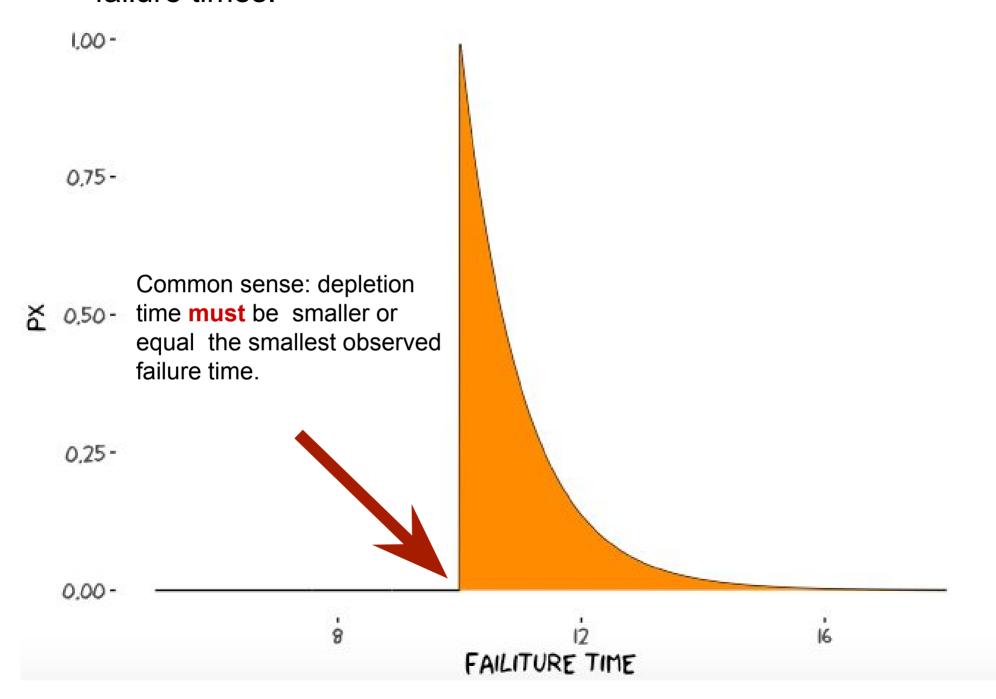


It is not feasible to observe the depletion rate directly, but one can observe the resulting failures after θ .

From previous data, one want to estimate the time θ of guaranteed safe operation... *

$$p(x \mid \theta) = \left\{ \begin{array}{ll} \exp(\theta - x) & , & x > \theta \\ 0 & , & x < \theta \end{array} \right\}$$

Goal: estimate the depletion time from a series of observed failure times.



Frequentist Approach

Unbiased estimator of the mean: $E(x) = \int_0^\infty x p(x) dx$ = $\theta + 1$

$$\hat{\theta} = \frac{1}{N} \sum_{i=1}^{N} x_i - 1$$

The 95% confidence intervals $CI_{\text{large N}} = (\hat{\theta} - 2N^{-1/2}, \hat{\theta} + 2N^{-1/2})$

For a sample of e.g. {14,18, 21} it will lead to CI = {15.53,17.8}

Bayesian approach: analytical

$$p(x \mid \theta) = \left\{ \begin{array}{ll} \exp(\theta - x) & , & x > \theta \\ 0 & , & x < \theta \end{array} \right\}$$

Likelihood:

$$p(D \mid \theta) = \prod_{i=1}^{N} p(x \mid \theta)$$

Posterior:

$$p(\theta \mid D) \propto \left\{ \begin{aligned} N \exp[N(\theta - \min(D))] &, & \theta < \min(D) \\ 0 &, & \theta > \min(D) \end{aligned} \right\}$$

 $\theta_2 = \min(D)$

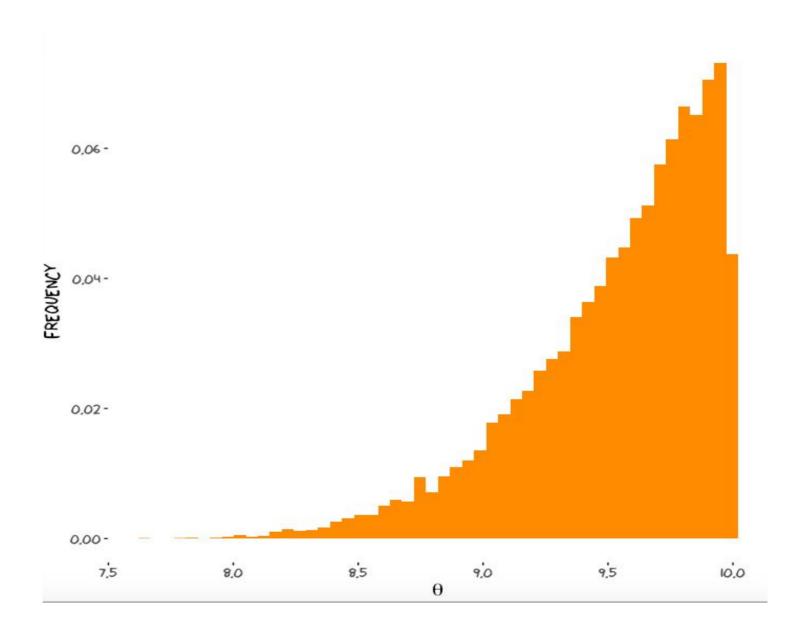
Leads for the same sample BCI = {13,14}

Interval:

$$\theta_1 = \theta_2 + \frac{\log(1-f)}{N}$$

$$\int_{\theta_1}^{\theta_2} N \exp[N(\theta - \theta_2)] d\theta = f$$

Bayesian approach: MCMC



Next: Markov Chain Monte Carlo

MCMC is a numerical technique data allow us to sample from a target distribution