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# HH+met reinterpretation exercise

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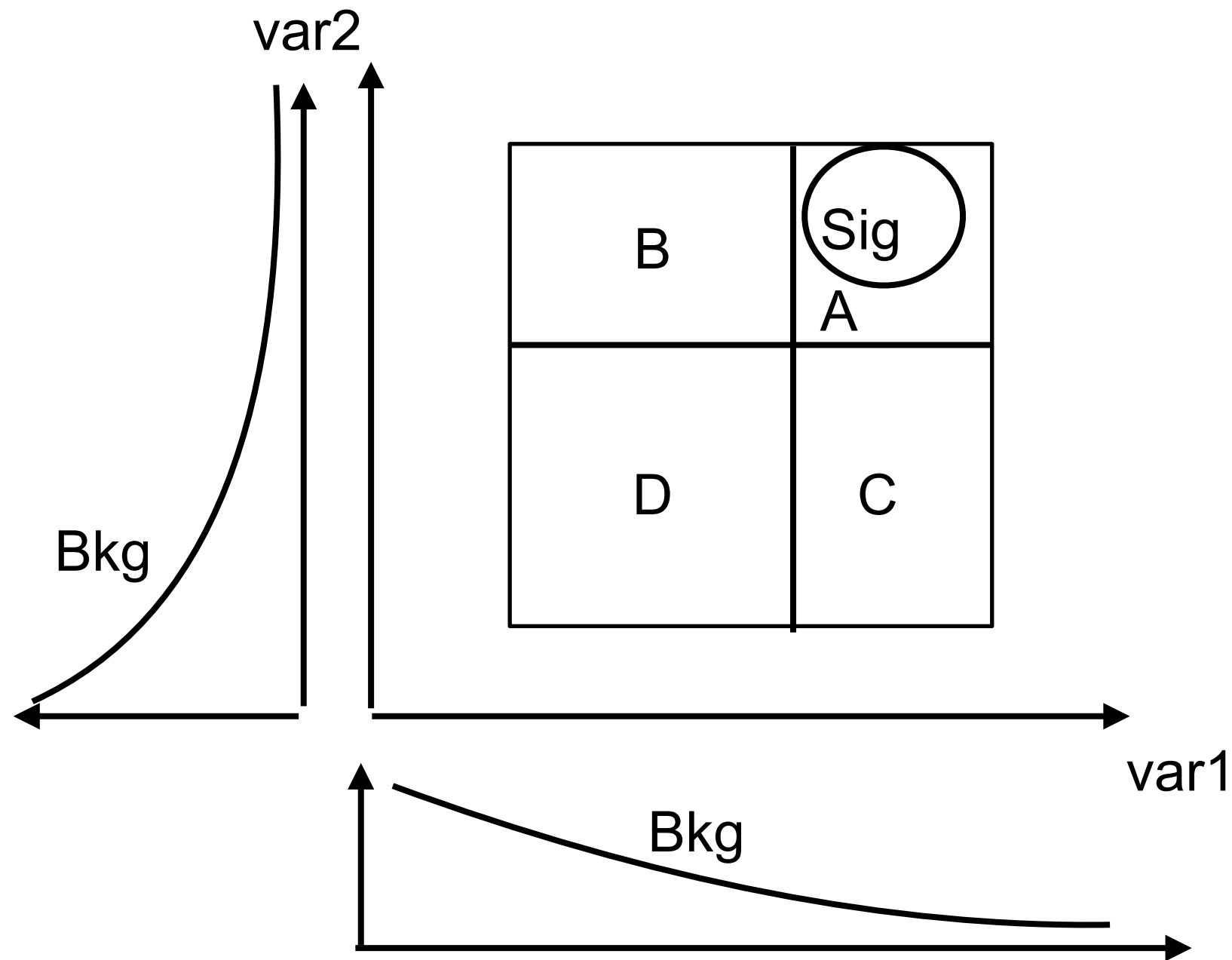


# Methods for calculating limits from the data in search experiments

- Higgs combine tool (used in Higgsino search [1], RA2/b [2], ...)
- Question: how well can one reproduce our results from HEPData?
- Chisquare fit
- Improved version: simplified likelihood
  - Described by Waltenberger at the June SUSY workshop:
- Results, comparisons
- Observations

- [1] CMS-SUS-20-004
- [2] CMS-SUS-19-006

# Data-driven background prediction, “ABCD”



- Predicted  $N_{\text{bkg}}$  in A =  $N_B (N_C / N_D)$
- All N's are event counts, so Poisson distributed

# Higgs combine tool method

## ■ Inputs are

- $N^{\text{obs}}_i$  in each SR (signal region, A)
- $N^{\text{obs}}_i$  in each SB, CSR, CSB (B, C, D) region for the ABCD estimate of  $N^{\text{bkg}}_i$  in A
- $N^{\text{sig}}_i$  (MC) in each SR, SB, CSR, CSB (for each  $(m_{\text{NLSP}}, m_{\text{LSP}})$  scan point)
- $\kappa_i$  for each SR, with uncertainty nuisance parameter
- Uncertainty nuisances for other systematics (scale factors, JEC, ...)

## ■ Likelihood is built from

- Poisson pdfs for  $N^{\text{obs}}_i$  in all SR, SB, CSR, CSB regions
- Constraints  $N^{\text{bkg}} = A = \kappa B C / D$  (rateParams)
- Gaussian pdfs for  $\kappa$  uncertainties
- Log-normal pdfs for other nuisances

## ■ The expected yields $N^{\text{exp}}_i$ are given by

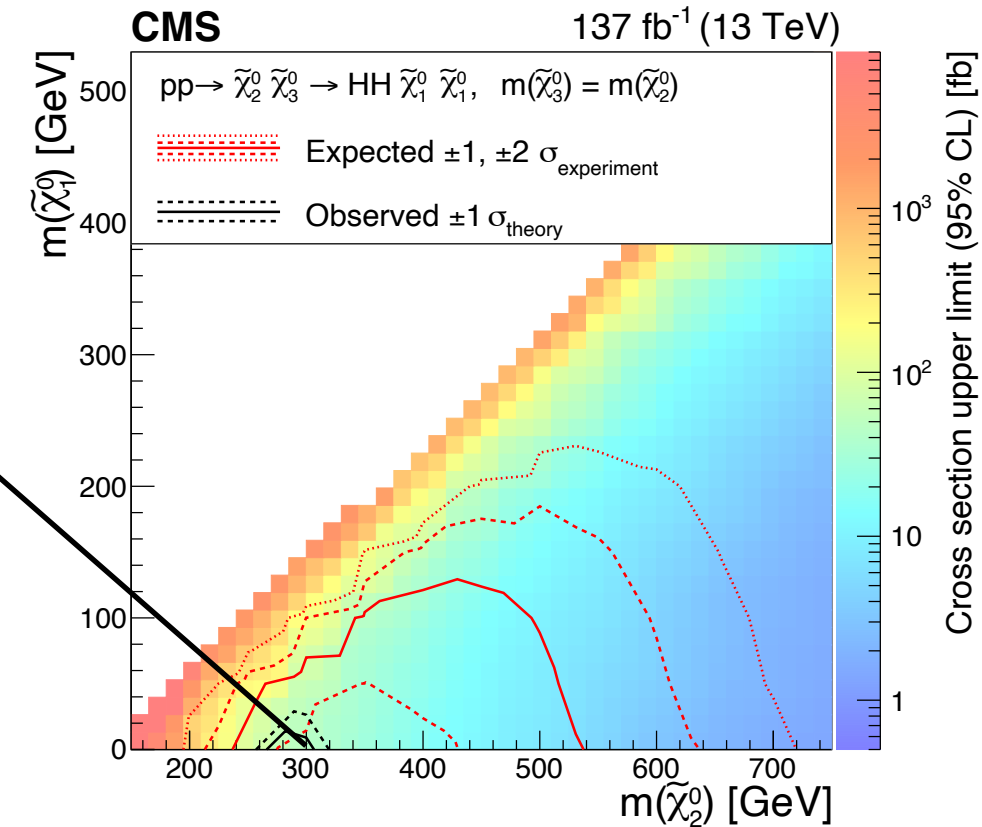
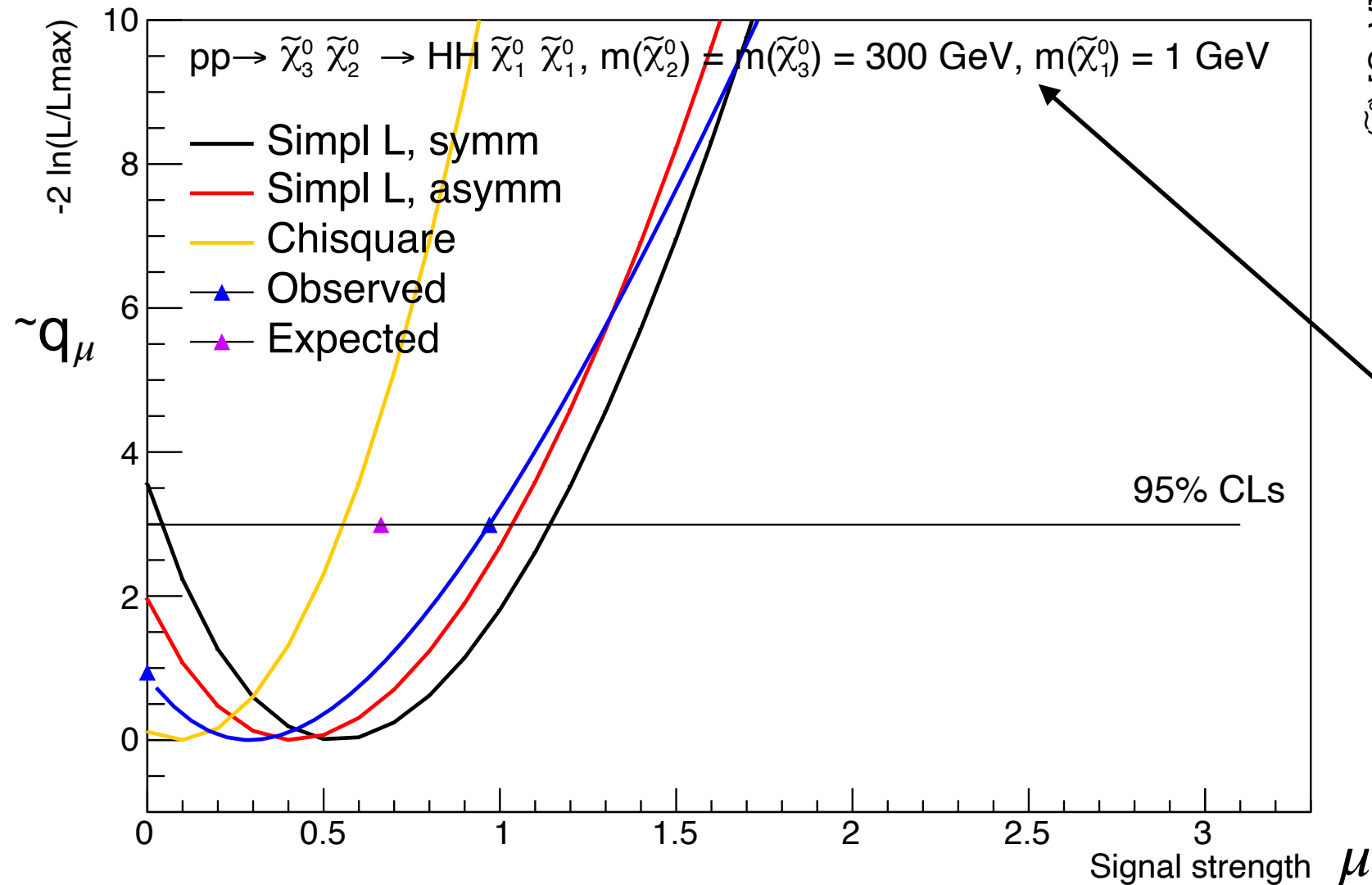
- $N^{\text{exp}}_i = N^{\text{bkg}}_i + \mu N^{\text{sig}}_i$ , where  $\mu$  (r in combine) is the signal strength

# Extraction of limit & profile likelihood

- Running “combine -M AsymptoticLimits” on the datacard for a model scan point gives:
  - 95% CL upper limit from the data
  - 95% CL upper limit expected, with +/- 1 and 2-sigma deviations
- The criterion for 95% CL is that  $CL_s = 0.05$ 
  - $CL_s = CL_{s+b} / CL_b$
  - $CL_{s+b} = 1 - \Phi(\sqrt{\tilde{q}_\mu})$ , where  $\tilde{q}_\mu$  is the profile likelihood test statistic:
$$\tilde{q}_\mu = -2 \ln \frac{\mathcal{L}(\text{data}|\mu, \hat{\theta}_\mu)}{\mathcal{L}(\text{data}|\hat{\mu}, \hat{\theta})}, \quad 0 \leq \hat{\mu} \leq \mu$$
, and  $\Phi$  is the normal cumulative distribution function (cdf)
  - $CL_b$  measured with the Asimov data set ( $N^{\text{obs}}$  set to  $N^{\text{expected}}$ )
  - Details in [CMS-NOTE-2011/005](#) (joint w ATLAS)
  - Adding `—verbose 1` to `AsymptoticLimits` gives  $CL_b$  at limit
- Running “combine -M MultiDimFit --algo grid” gives  $\tilde{q}_\mu$  vs  $\mu$

# Profile likelihood vs $\mu$

Profile likelihood vs signal strength



- Blue curve from combine tool method
  - blue triangles: significance, 95% CLs limit
    - $\mu < 1$  @ 95% CLs  $\Rightarrow$  this (300, 1) point is (barely) excluded
  - purple triangle: expected limit

# Reproducing the result from HEPData

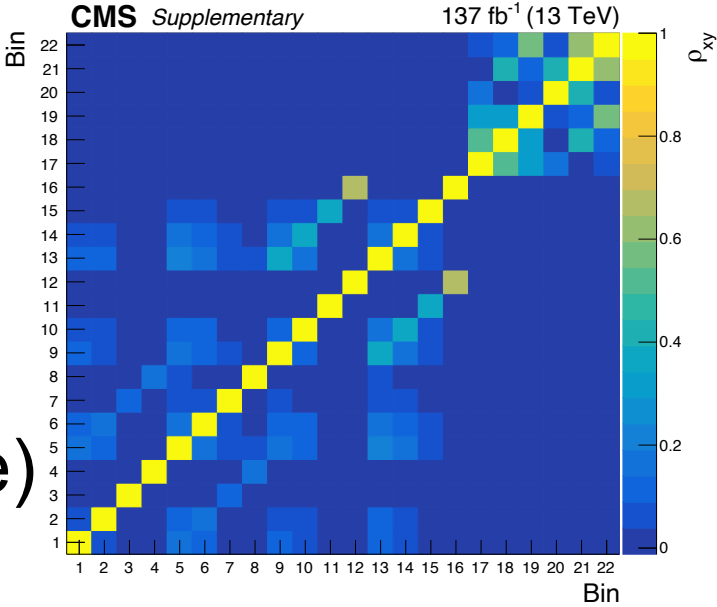
- The (relevant) available information:
- From Tables 4 & 5
  - Observed yields
  - Estimated pre-fit background yields
    - upper and lower uncertainties
- From figs. 11 (TChiHH-G), 13 (TChiHH), & 14 (T5HH)
  - Cross sections vs scan point ( $m_{\text{NLSP}}$ ,  $m_{\text{LSP}}$ )
    - Observed 95% CL upper limit, +/- 1 sigma
    - Expected 95% CL upper limit, +/- 1, 2 sigma
    - Theory (vs  $m_{\text{NLSP}}$ )
- Covariance matrix
- Efficiency vs scan point and SR bin
- Significance vs scan point

# From HEPData

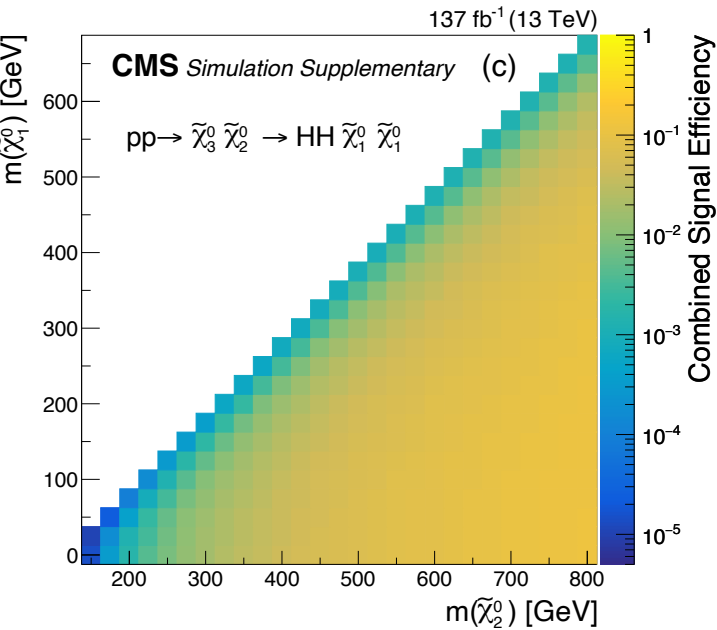
Bin	$\Delta R_{\max}$	$N_b$	$p_T^{\text{miss}}$ [GeV]	$\kappa$	$N_{\text{SR}}^{\text{pred}}$	$N_{\text{SR}}^{\text{fit}}$	$N_{\text{SR}}^{\text{obs}}$
1	1.1–2.2	3	150–200	$1.09 \pm 0.04 \pm 0.02$	$161^{+14}_{-13}$	$149.7^{+8.9}_{-8.5}$	138
2			200–300	$0.92 \pm 0.04 \pm 0.02$	$90.4^{+9.7}_{-9.0}$	$91.5^{+6.9}_{-6.5}$	91
3			300–400	$0.94 \pm 0.09 \pm 0.01$	$11.5^{+3.4}_{-2.7}$	$12.8^{+2.6}_{-2.2}$	14
4			>400	$0.98^{+0.19}_{-0.16} \pm 0.02$	$2.8^{+2.3}_{-1.4}$	$2.8^{+1.4}_{-1.0}$	3
5		4	150–200	$1.13 \pm 0.09 \pm 0.08$	$53.5^{+8.8}_{-7.8}$	$54.1^{+5.6}_{-5.2}$	54
6			200–300	$0.96 \pm 0.07 \pm 0.07$	$28.3^{+5.6}_{-4.8}$	$33.2^{+4.2}_{-3.9}$	38
7			300–400	$0.89^{+0.16}_{-0.15} \pm 0.05$	$2.6^{+1.5}_{-1.1}$	$3.2^{+1.3}_{-1.0}$	4
8			>400	$0.92^{+0.27}_{-0.22} \pm 0.07$	$2.6^{+2.4}_{-1.4}$	$1.27^{+0.98}_{-0.63}$	0
9	<1.1	3	150–200	$1.05^{+0.18}_{-0.15} \pm 0.12$	$5.1^{+1.6}_{-1.3}$	$5.9^{+1.4}_{-1.2}$	8
10			200–300	$1.04^{+0.14}_{-0.13} \pm 0.11$	$2.17^{+0.79}_{-0.60}$	$2.31^{+0.73}_{-0.57}$	2
11			300–400	$0.72^{+0.33}_{-0.22} \pm 0.08$	$0.06^{+0.11}_{-0.04}$	$0.72^{+0.53}_{-0.33}$	4
12			>400	$1.24^{+0.67}_{-0.45} \pm 0.10$	$0.89^{+1.42}_{-0.60}$	$0.52^{+0.65}_{-0.35}$	0
13		4	150–200	$1.26^{+0.21}_{-0.20} \pm 0.23$	$2.68^{+1.06}_{-0.79}$	$2.58^{+0.85}_{-0.67}$	1
14			200–300	$1.21^{+0.22}_{-0.21} \pm 0.22$	$1.26^{+0.62}_{-0.44}$	$1.62^{+0.65}_{-0.48}$	3
15			300–400	$2.35^{+0.88}_{-0.72} \pm 0.34$	$0.42^{+0.61}_{-0.27}$	$1.16^{+0.87}_{-0.55}$	1
16			>400	$0.94^{+0.53}_{-0.36} \pm 0.13$	$0.67^{+1.10}_{-0.46}$	$0.78^{+0.76}_{-0.43}$	1

Bin	$N_H$	$p_T^{\text{miss}}$ [GeV]	$N_{\text{SR, tot}}^{\text{pred}}$	$f_{\text{pTmiss}}$	$N_{\text{SR}}^{\text{pred}}$	$N_{\text{SR}}^{\text{fit}}$	$N_{\text{SR}}^{\text{obs}}$
17	1	300–500		$0.789 \pm 0.030$	$33.6^{+6.1}_{-5.2}$	$37.0^{+4.2}_{-4.0}$	42
18		500–700	$42.6 \pm 4.2$	$0.172 \pm 0.028$	$7.3^{+2.0}_{-1.6}$	$7.2^{+1.5}_{-1.3}$	6
19		>700		$0.039 \pm 0.014$	$1.65^{+1.04}_{-0.66}$	$1.50^{+0.75}_{-0.53}$	1
20	2	300–500		$0.789 \pm 0.030$	$4.0^{+1.5}_{-1.1}$	$4.0^{+1.2}_{-1.0}$	4
21		500–700	$5.1 \pm 1.0$	$0.172 \pm 0.028$	$0.88^{+0.40}_{-0.28}$	$0.74^{+0.29}_{-0.21}$	0
22		>700		$0.039 \pm 0.014$	$0.20^{+0.21}_{-0.10}$	$0.14^{+0.13}_{-0.07}$	0

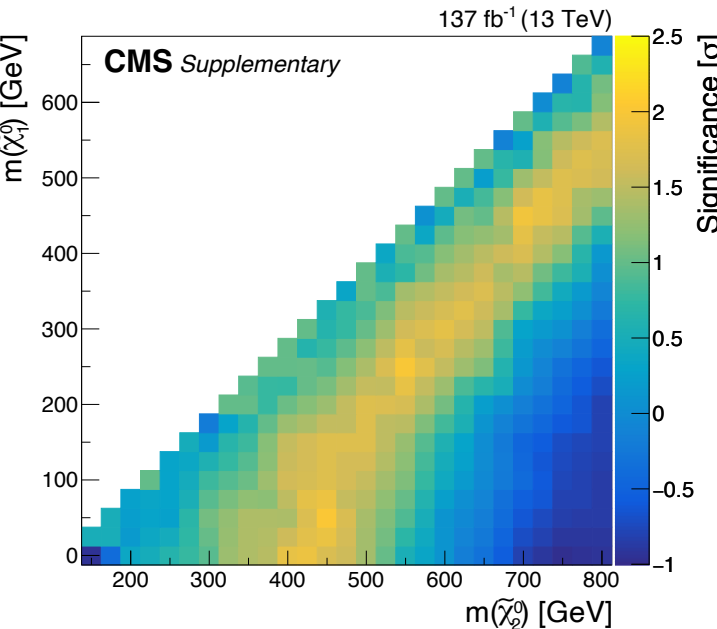
covariance  
(correlation  
shown here)



efficiency  
(by bin  
available)



significance



- HEPData tables downloadable as csv, yaml, ...



# Chisquare method

- Simplest approximation: build the likelihood as a product of Gaussians.
- Then minimize  $\chi^2 = -2 \ln(L)$ :

- Define the residual  $\Delta_i$  for each bin  $i$ :

$$\Delta_i \equiv N_i^{\text{obs}} - N_i^{\text{pred}},$$

$$N_i^{\text{pred}} \equiv N_i^{\text{bkg}} + \mu N_i^{\text{sig}},$$

$$N_i^{\text{sig}} = \epsilon_i \mathcal{B}^2(H \rightarrow b\bar{b}) \sigma \mathcal{L}$$

- $\mu$  is the signal strength
  - $\epsilon$ ,  $\sigma$ , and  $\mathcal{L}$  are the efficiency, cross section, and luminosity
- Given the covariance matrix  $V$ , we find

$$\chi^2 = \Delta_i V_{ij}^{-1} \Delta_j,$$

$$\rightarrow \sum \Delta_i^2 / \sigma_i^2$$

- in the limit of uncorrelated bins.

# Accounting for signal uncertainty in chisq

- The covariance matrix is given for the background.
- In the chisq formulation we need to add the contribution of the signal uncertainty:

$$V = V^{\text{bkg}} + \text{diag}(N^{\text{obs}})$$

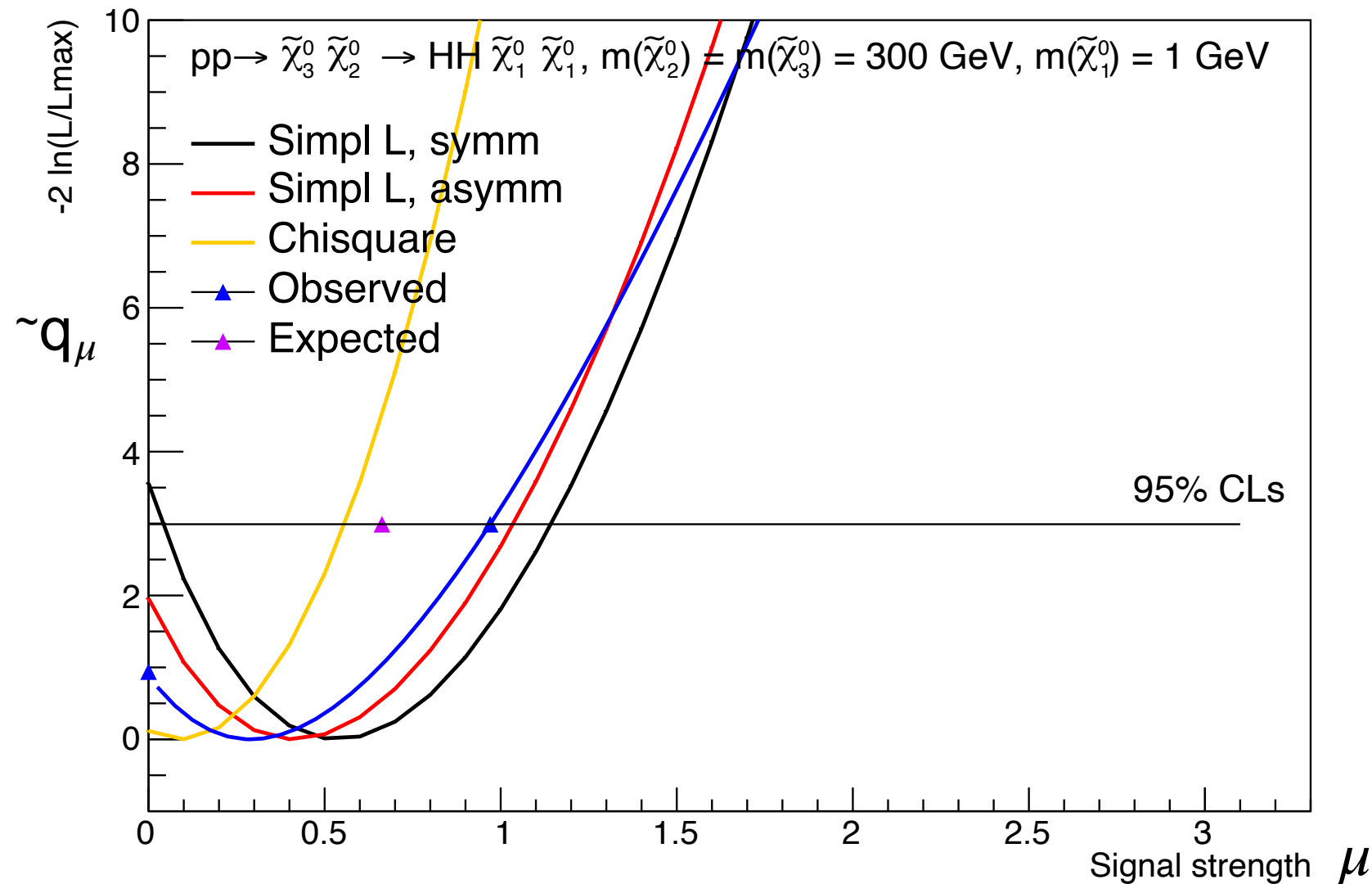
- $N^{\text{obs}}$  being the square of the Poisson uncertainty  $\sqrt{N^{\text{obs}}}$
- The resulting function  $\chi^2(\mu)$  is a parabola, with minimum at the best-fit value of  $\mu$ ,  $\mu_0$ .
- The square root of  $\chi^2(\mu) - \chi^2(\mu_0)$  at any other value of  $\mu$  gives the Z-value (number of standard deviations of a Gaussian pdf).

# Limitations of the chisq method

- All uncertainties assumed Gaussian
  - and symmetric
  - not so good for small yields
  - $V^{\text{bkg}}$  reflects the combined effects of all sources (CR yields, scale factors, JECs, etc., whereas some contributions are statistical (sideband yields).
  - $V^{\text{obs}}$  represents Poisson stat uncertainties with a Gaussian approximation
- Signal uncertainties are not accounted for.
- The contribution of each bin to chisq represents the consistency of background and observation in that bin.
  - But the real question is, given  $N^{\text{bkg}}$ , what's the probability of observing  $N^{\text{obs}}$ ?
  - E.g., the bin 11 contribution before squaring is (very nearly)  $(4 - 0)/\sqrt{4}$ , which is 2 sigma, vs the detailed study giving 3.3 sigma local significance.

# Chisquare method for HH+met

Profile likelihood vs signal strength



- Orange curve from the  $\chi^2$  method
  - underestimates  $\mu_0$
  - underestimates the high-side uncertainty



# Simplified likelihood method

- Feedback from phenomenology team (Wolfgang Waltenberger): need efficiencies by SR bin
  - We now provide these in HEPData (and public twiki)
- Wolfgang also spoke at the June SUSY workshop
- Introduced the simplified likelihood method, SM:
  - <https://arxiv.org/abs/1809.05548>, [https://link.springer.com/article/10.1007/JHEP04\(2019\)064](https://link.springer.com/article/10.1007/JHEP04(2019)064)
- Reproduces one of RA2/b Run 2 limits quite well
- Improvements over chisq method:
  - Poisson pdfs for the observed yields
  - Option to include asymmetric background uncertainties:

# Building the simplified likelihood

- The predicted yield in bin  $i$  is

$$N_i^{\text{pred}} \equiv N_i^{\text{bkg}} + \mu N_i^{\text{sig}},$$

$$N_i^{\text{bkg}} = a_i + b_i \theta_i + c_i \theta_i^2$$

- $a_i$  is the central value of the bkg prediction
- $\theta_i$  is a nuisance parameter drawn from a unit Gaussian
- $b_i$  is the effective sigma of the bkg uncertainty,  $\sqrt{V_{ii}}$  in the limit of symmetric uncertainties
- $c_i$  gives the asymmetry of the bkg uncertainty

- The simplified likelihood is

$$L_S(\mu, \theta) \propto \prod_i \text{Pois}(N_i^{\text{obs}} | N_i^{\text{pred}}(\mu)) \exp\left(-\frac{1}{2} \theta_i \rho_{ij}^{-1} \theta_j\right)$$

- where  $\rho \rightarrow$  correlation matrix in the symmetric limit

- The code, and example input, are in gitHub:

- <https://gitlab.cern.ch/SimplifiedLikelihood/SLtools>

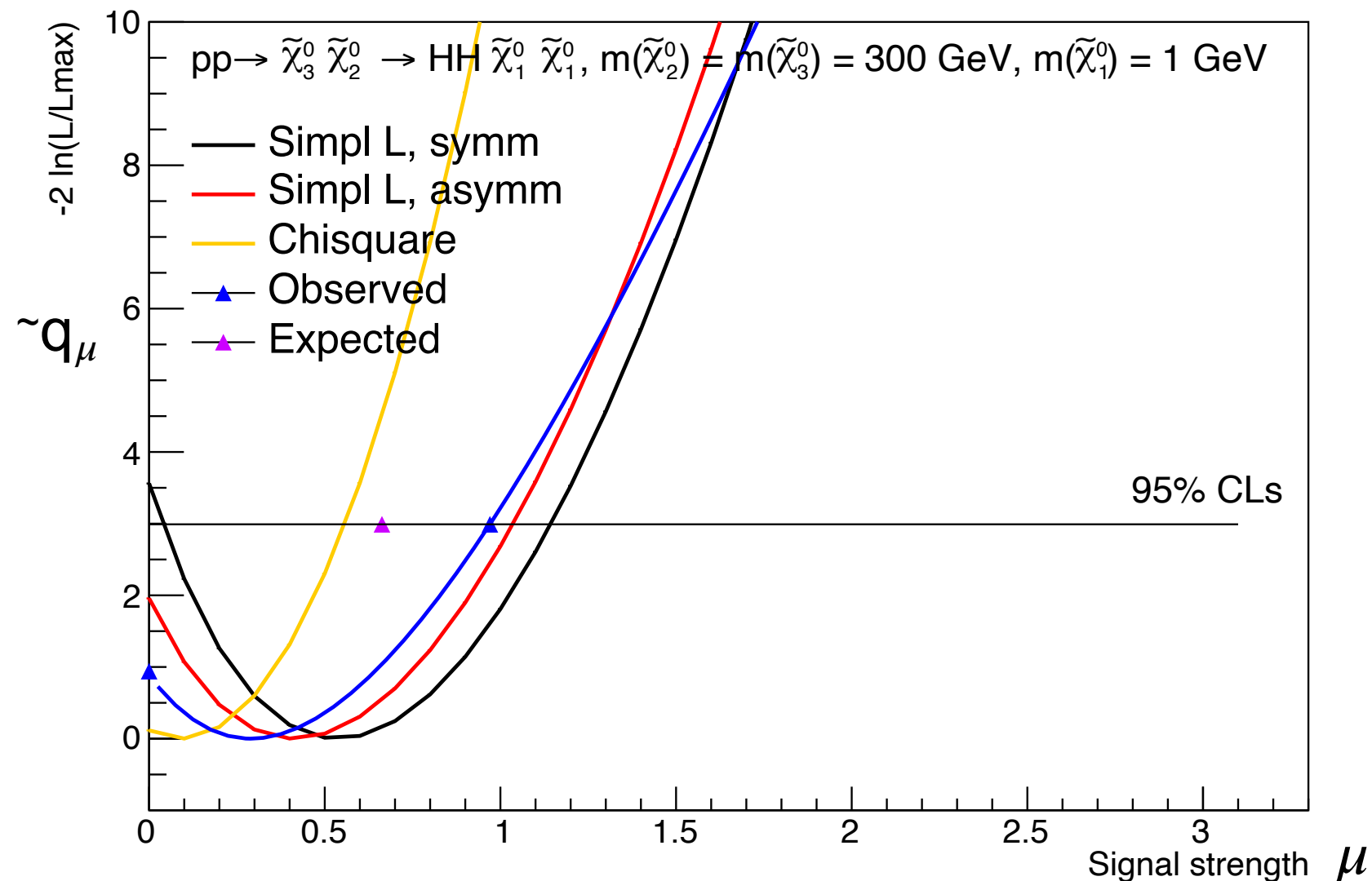
# SL: asymmetric bkg uncertainties

- The covariance matrix gives second moments, i.e.,  $\sigma^2$ , on the diagonal, and correlations, on off-diagonal elements
- To incorporate asymmetric uncertainties, SL uses the diagonal elements of the 3<sup>rd</sup> moment  $m_3$  of the background nuisances.
  - Maybe these could be extracted from combine tool?
- Compute  $m_3$  from a bifurcated Gaussian using the asymmetric uncertainties  $\sigma_{1,2}$ :

$$m_3 = \frac{2}{\sigma_1 + \sigma_2} \left[ \sigma_1 \int_{-\infty}^0 x^3 G(x; 0, \sigma_1) dx + \sigma_2 \int_0^{+\infty} x^3 G(x; 0, \sigma_2) dx \right]$$

# Compare SL with full L

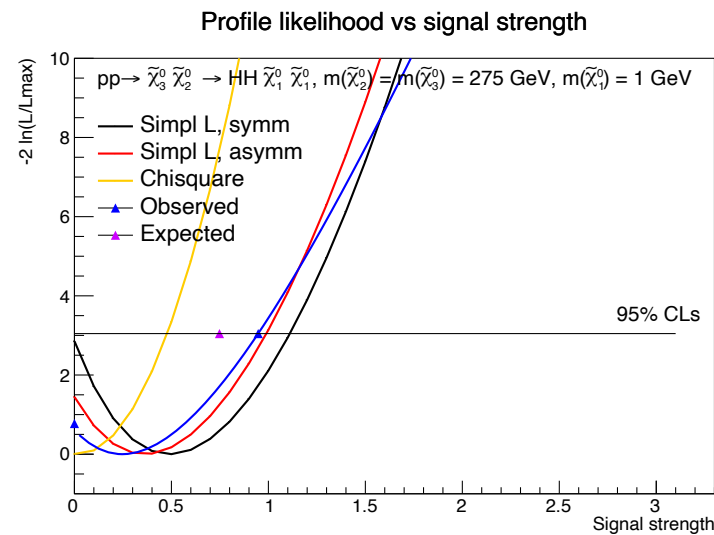
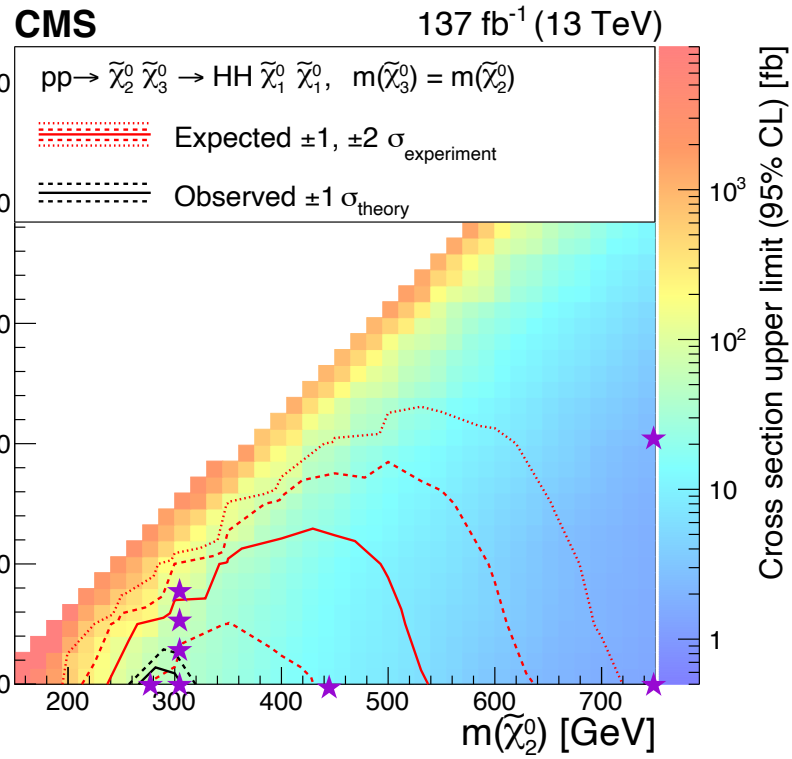
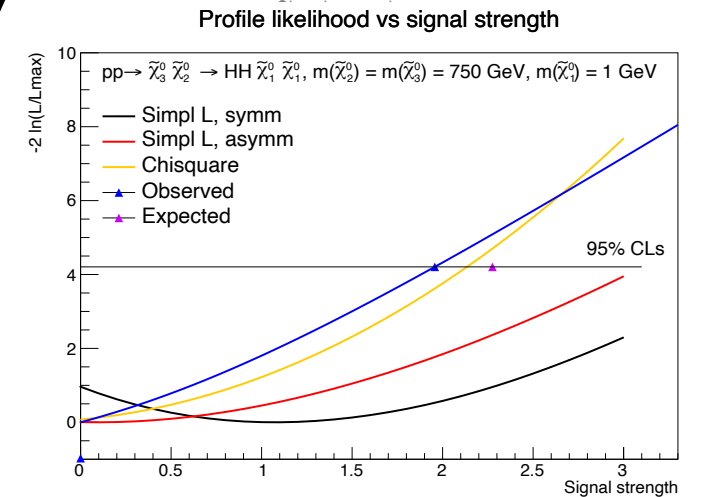
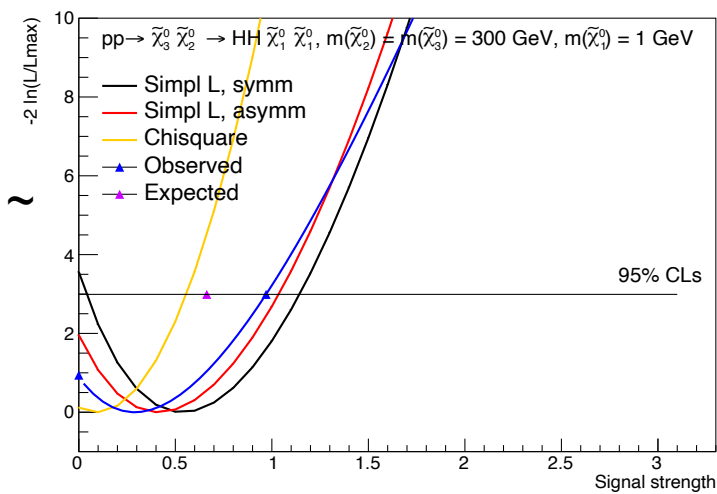
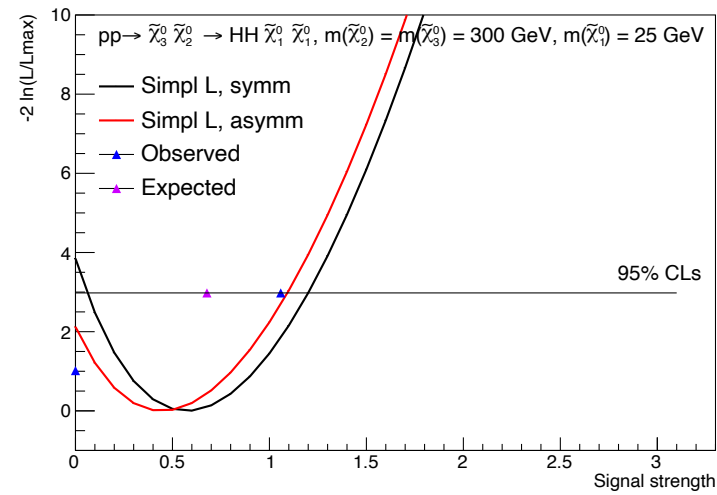
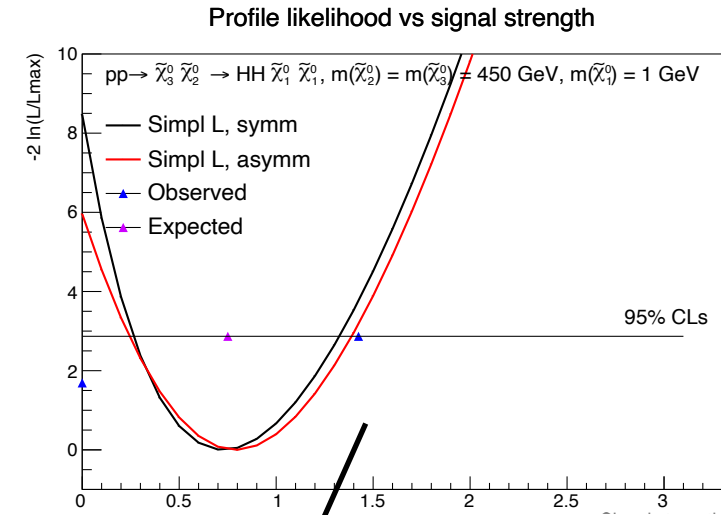
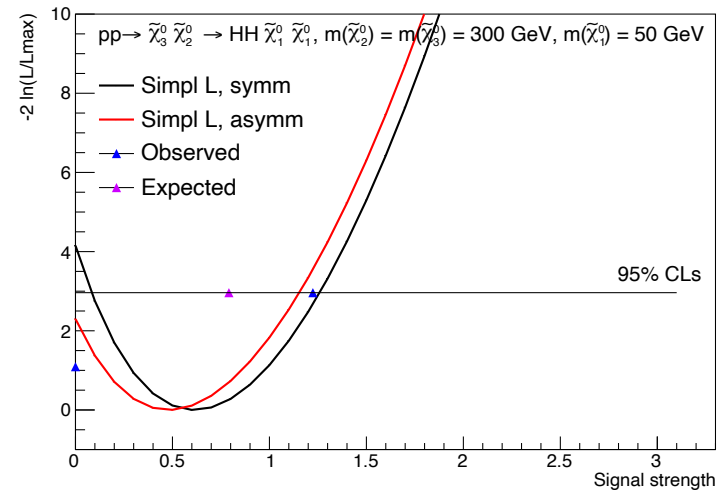
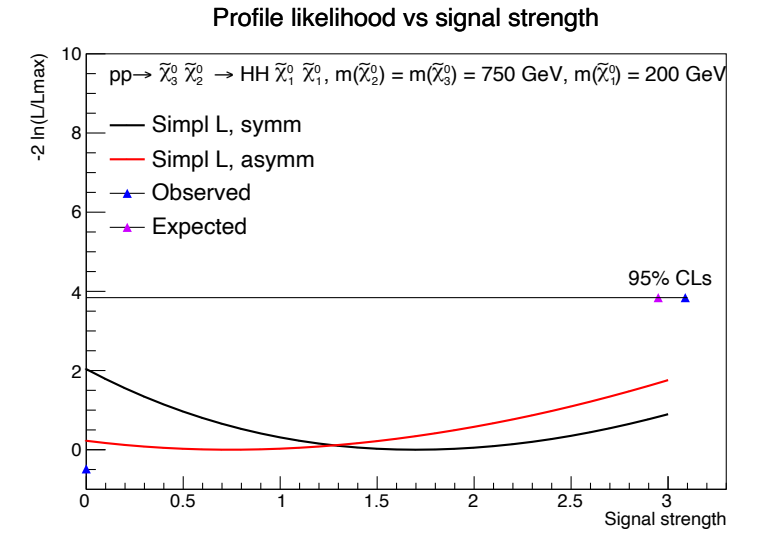
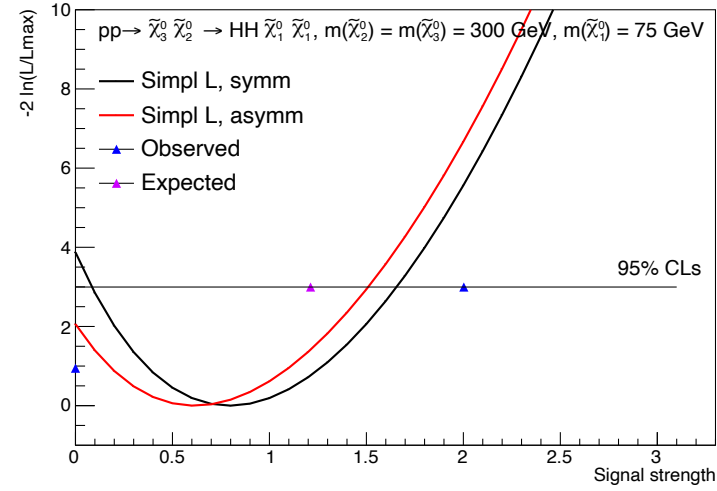
Profile likelihood vs signal strength



- The red (black) curve is SL with(out) asymmetric uncertainties
- SL treats  $N^{\text{obs}}$  as Poisson,
- $N^{\text{bkg}}$  as (asymmetric) Gaussian
  - doesn't fully account for Poisson fluctuations of low-stats CR yields



# L & SL vs scan pt



# Some observations

- Phenomenologists have rather sophisticated tools for incorporating published data from experiments
- For the HH+met and RA2/b papers, the provided information is adequate for reasonably closely reproducing the results
- How can one test other models with our HH+met data?
  - Reweight bin efficiencies?
    - The b quark content would have to be the same, else sorting of the model into 3b, 4b, or 1H, 2H bins would be impossible
    - Sorting the model into resolved/boosted, maybe with a cut on  $\Delta R$  between the H daughter quarks?
    - MET OK
    - $\Delta R_{\max}$  OK?
- For some ATLAS papers, “likelihood” is provided.
  - What would this entail?

# Update: combine with InN nuisances

- Jaebak presented a fit made with combine, but with
  - backgrounds entered directly from the paper
  - bkg uncertainties from the paper, entered into the datacard as (asymmetric) InN nuisances
  - no correlations
- I've modified this approach by adding the bkg correlations
  - Same InN nuisances, but multiplied by the correlation matrix elements to extend to the off-diagonal terms
  - Results shown as green plots in the profile plots (following page)
- Compared with Waltenberger's simplified likelihood
  - these do much better estimating the significance
  - not necessarily better in predicting the limits

# L & SLs vs scan point

