

Computational statistics, lecture math 2

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Topics today

- Probability distributions (reading: GH or Gentle, chapter 1.3) and integration
- Multivariate normal distribution
- Markov chains (reading: GH, chapter 1.7)

Exponential distribution

- The density of the exponential distribution $Exp(\lambda)$ is

$$f(x) = \lambda \exp(-\lambda x) \mathbf{1}\{x \geq 0\} = \begin{cases} \lambda \exp(-\lambda x), & \text{if } x \geq 0, \\ 0, & \text{if } x < 0 \end{cases}$$

dexp(x, rate=0.8)

- Compute the cumulative distribution function (CDF):

$$F(x) = P(X \leq x) = \int_{-\infty}^x f(t) dt = \begin{cases} 1 - \exp(-\lambda x), & \text{if } x \geq 0, \\ 0, & \text{if } x < 0 \end{cases}$$

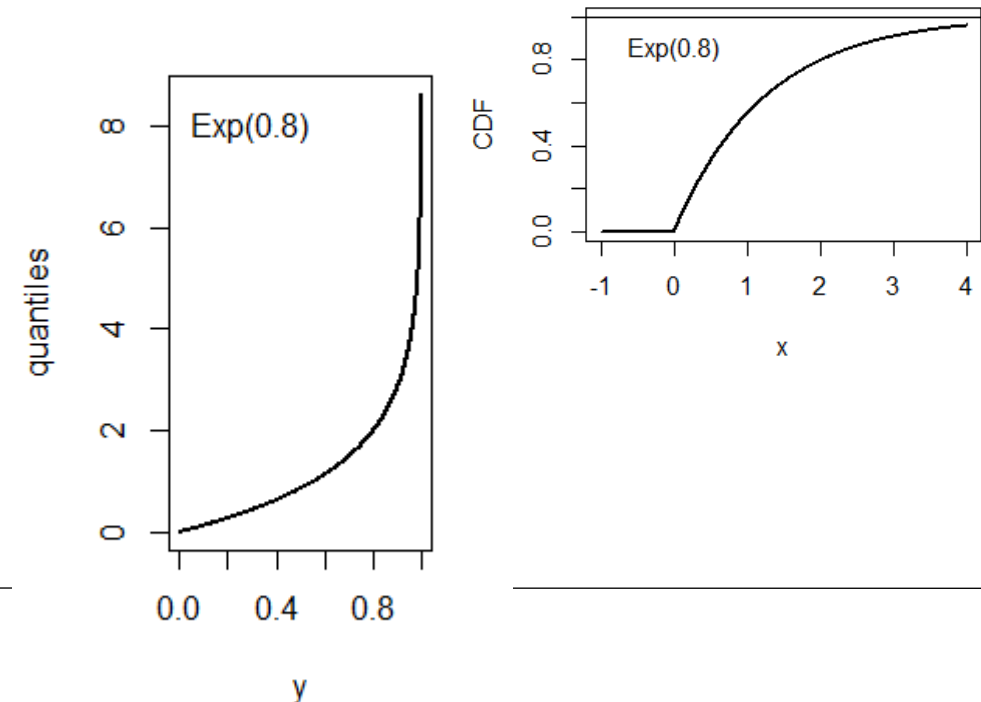
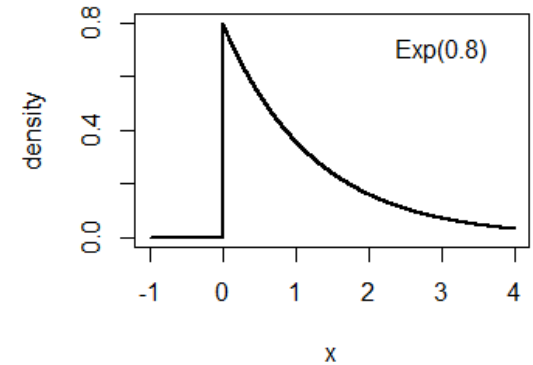
pexp(x, rate=0.8)

- The inverse cumulative distribution function is

$$F^{-1}(y) = -\log(1 - y)/\lambda$$

$$\begin{aligned} \text{since } y &= 1 - \exp(-\lambda x) \Leftrightarrow \exp(-\lambda x) = 1 - y \Leftrightarrow \\ &-\lambda x = \log(1 - y) \Rightarrow x = -\log(1 - y)/\lambda \end{aligned}$$

qexp(x, rate=0.8)



Expected value and variance of a distribution

- The probability for $X \in [a, b]$ can be written as integral according to

$$P(X \in [a, b]) = \int_a^b f(x) dx = F(b) - F(a)$$

- The expected value of a distribution with density $f(x)$ is:

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx$$

- The variance is

$$\text{Var}(X) = \int_{-\infty}^{\infty} (x - EX)^2 f(x) dx$$

- What is $E(X)$ and $\text{Var}(X)$ for the exponential distribution?

Integration

- The expected value of the exponential distribution is:

$$E(X) = \int_{-\infty}^{\infty} x f(x) dx = \int_{-\infty}^{\infty} x \lambda \exp(-\lambda x) \mathbf{1}\{x \geq 0\} dx =$$

Poisson distribution

- The density of the Poisson distribution $Po(\lambda)$ is

$$f(x) = \frac{\lambda^x}{x!} \exp(-\lambda) \text{ for } x = 0, 1, 2, \dots$$

dpois(x, lambda=2)

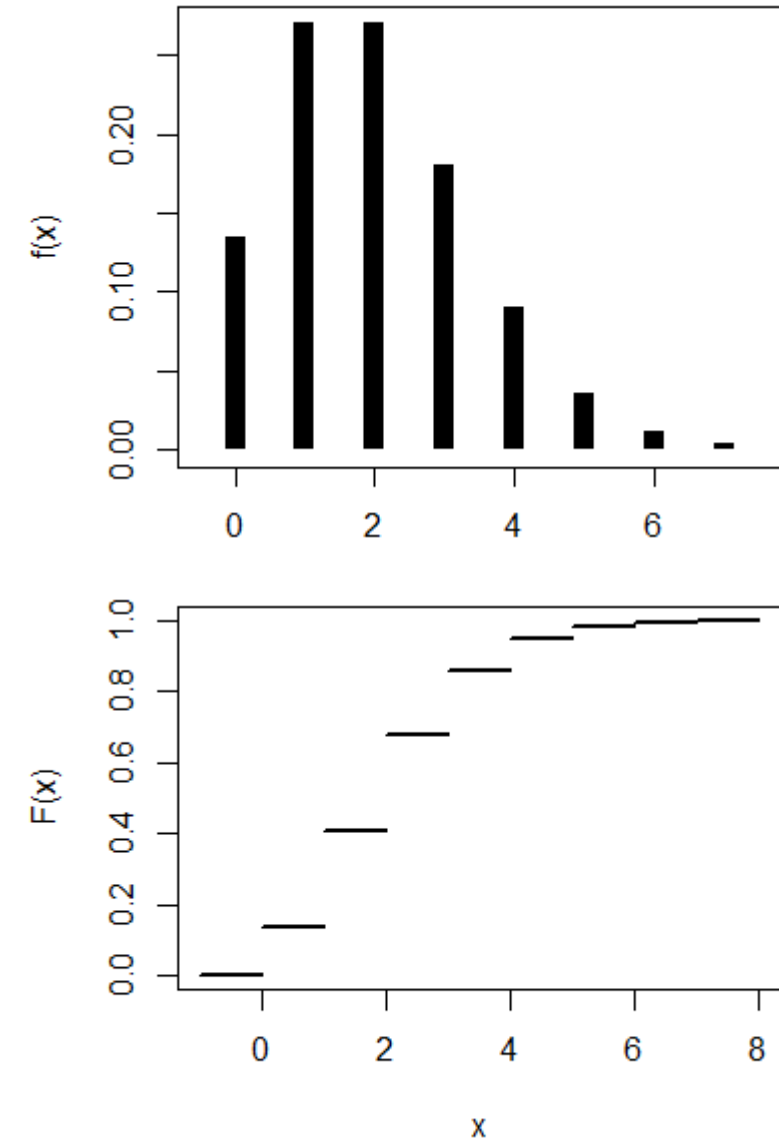
- The cumulative distribution function is:

$$F(x) = P(X \leq x) = \sum_{j=0}^x f(j) = \exp(-\lambda) \sum_{j=0}^x \frac{\lambda^j}{j!}$$

ppois(x, lambda=2)

- The inverse CDF (quantile function) can be defined as
 $F^{-1}(y) = \min\{x: F(x) \geq y\}$

qpois(x, lambda=2)



Exponential family (GH or Gentle, chapter 1.3)

- Exponential distribution: $f(x) = \lambda \exp(-\lambda x) \mathbf{1}\{x \geq 0\}$
- Poisson distribution $f(x) = \frac{\lambda^x}{x!} \exp(-\lambda)$
- Exponential, Poisson, normal, ... distributions belong to the exponential family of distributions
- Density of exponential family distribution:

$$f(x) = c_1(x)c_2(\gamma)\exp\left(\sum_{i=1}^k y_i(x)\theta_i(\gamma)\right)$$

Multivariate normal distribution

- A multivariate normal distribution $N(\mu, \Sigma)$ with mean vector μ and covariance matrix Σ has density

$$f(x) = \frac{1}{((2\pi)^p |\Sigma|)^{1/2}} e^{-(x-\mu)^T \Sigma^{-1} (x-\mu)/2}$$

- If X is a p -dimensional vector of independent standard normal variables (i.e., X is multivariate normal $N(0, I)$, $X \sim N(0, I)$), then $Y = A^T X + \mu$ has a $N(\mu, \Sigma)$ -distribution with $\Sigma = A^T A$
- If Σ is given, A with $\Sigma = A^T A$ can be computed, e.g., with the Cholesky decomposition)

Generate a bivariate normal distribution

- Example: Let Σ be the covariance matrix of a bivariate normal distribution with each component having variance 1 and correlation being ρ , $-1 \leq \rho \leq 1$, i.e., $\Sigma = \begin{pmatrix} 1 & \rho \\ \rho & 1 \end{pmatrix}$
- One can show that $\Sigma = A^T A$ with $A = \begin{pmatrix} 1 & \rho \\ 0 & \sqrt{1 - \rho^2} \end{pmatrix}$
- Therefore:
 - Let $X_1 \sim N(0,1)$; $X_2 \sim N(0,1)$ be independent, $X = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix}$
 - Let $Y = A^T X$, which means:
 $Y_1 = X_1$,
 $Y_2 = \rho X_1 + \sqrt{1 - \rho^2} X_2$.
 - Then, $\mu + Y \sim N(\mu, \Sigma)$

Generate multivariate normal distribution in R

- Package **mvtnorm** supplies (besides densities, probabilities, and quantile functions) random generator **rmvnorm** for multivariate normal:

$$\Sigma = \begin{pmatrix} 1 & 0.5 & 0.25 \\ 0.5 & 1 & 0.5 \\ 0.25 & 0.5 & 1 \end{pmatrix}$$

```
library(mvtnorm)
```

```
Sig <- matrix(c(1,0.5,0.25, 0.5,1,0.5, 0.25,0.5,1), nrow = 3)
```

```
rmvnorm(5, mean = c(2, -1, 0), sigma = Sig)
```

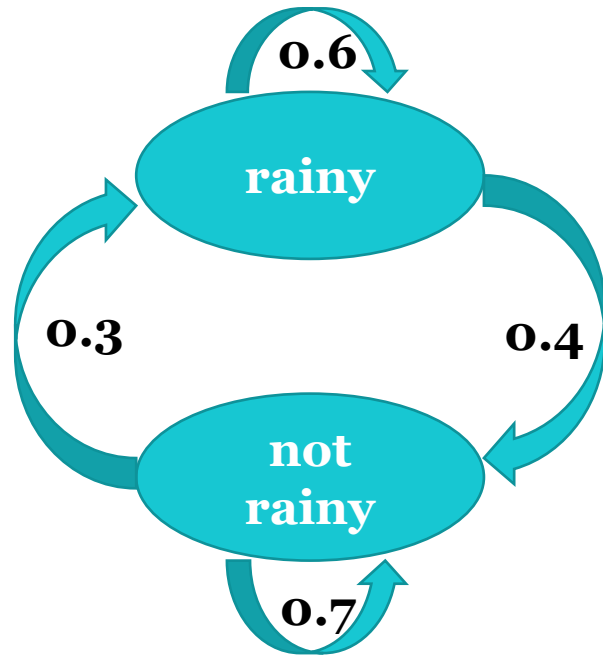
```
      [,1]      [,2]      [,3]  
[1,] 2.447613 -1.7390355 -0.43272668  
[2,] 1.933296 -1.9893701  0.47161126  
[3,] 2.872344 -0.6396215  0.61383902  
[4,] 2.505797 -0.8080744 -0.51696135  
[5,] 2.614221 -1.9280517  0.04471815
```

Each row is one
generated random vector

- Optionally, method of decomposition can be chosen
- Multivariate t distribution (same package): **rmvt**

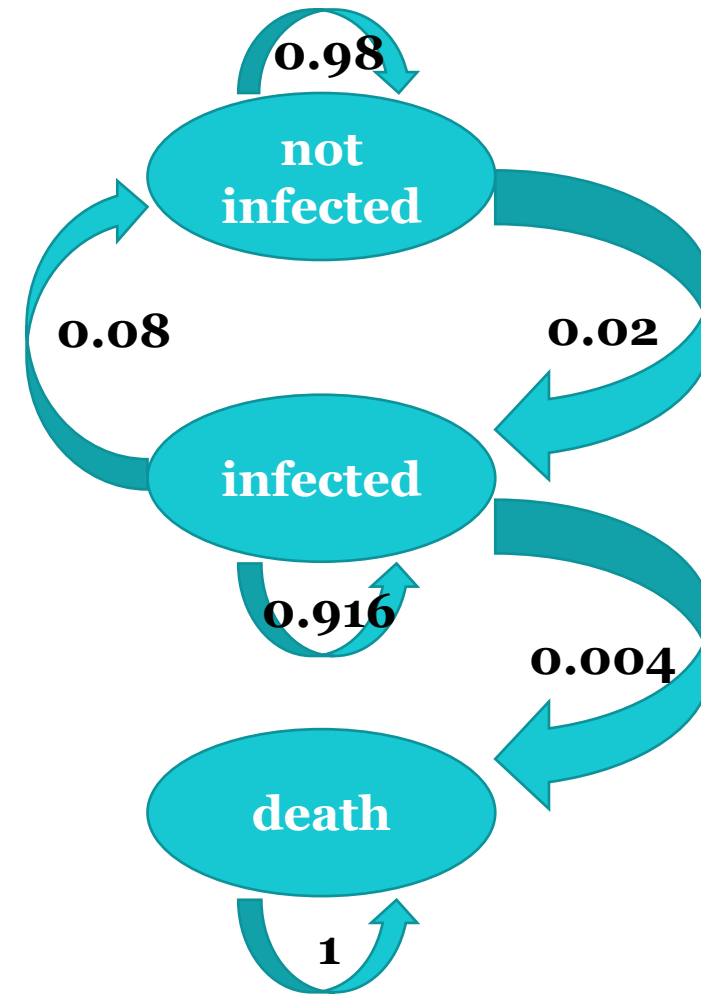
Markov chains (reading: GH chapter 1.7)

- Consider a sequence of random variables $(X^{(t)}), t = 0, 1, \dots$
- Each $X^{(t)}$ can have values in so-called *state space* S (can be discrete; we will later use a continuous one- or multidimensional S)
- A general random sequence (process) is described by specification of $P(X^{(t)} | X^{(t-1)}, \dots, X^{(0)})$ for all t
- A **Markov chain** is a specific process with *Markov property*
$$P(X^{(t)} | X^{(t-1)}, \dots, X^{(0)}) = P(X^{(t)} | X^{(t-1)})$$
- The state of $X^{(t)}$ depends only on the state before and not earlier history ($X^{(t)}$ is memoryless)
- $P(X^{(t)} | X^{(t-1)})$ is called *transition probability*



state space
 $S = \{\text{rainy}, \text{not rainy}\}$

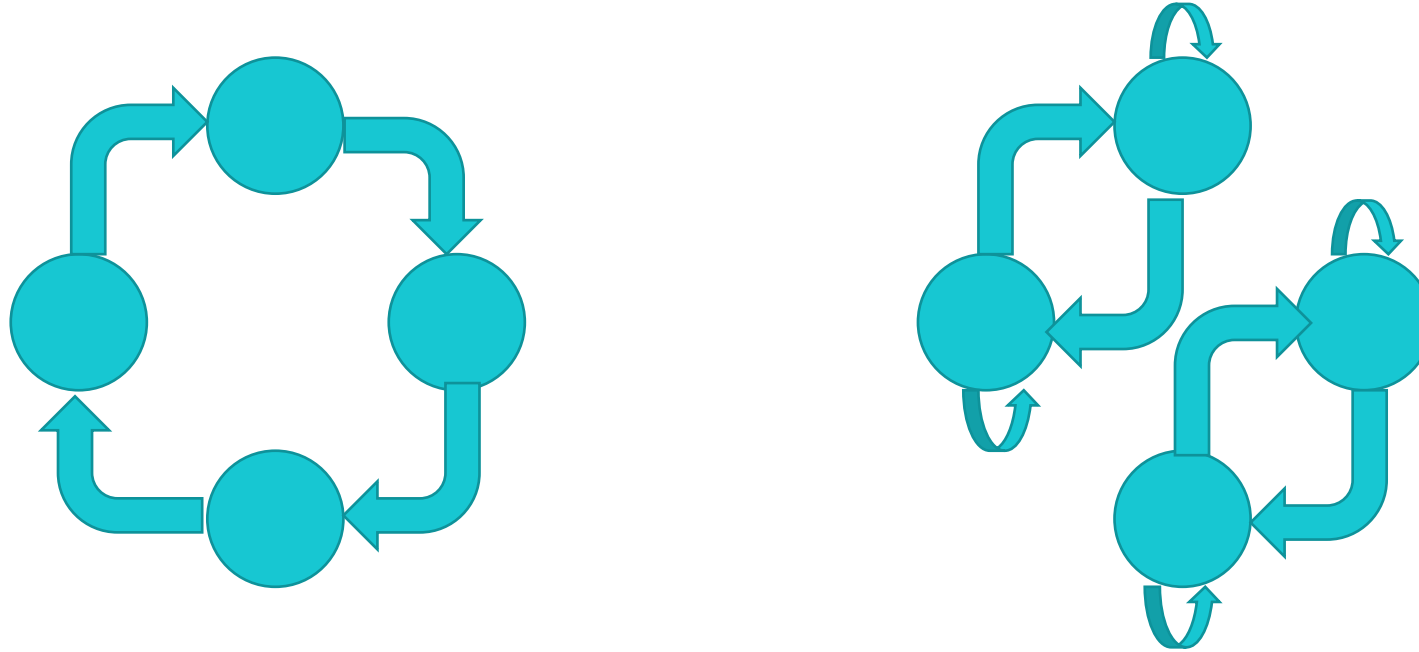
numbers are
transition probabilities



$S = \{\text{not infected}, \text{infected}, \text{death}\}$

Markov chains

- If S is discrete, we can write $p_{ij}^{(t)} = P(X^{(t)} = j | X^{(t-1)} = i)$
- If S is continuous, $P(X^{(t)} | X^{(t-1)})$ is represented by a cumulative distribution function or density
- A Markov chain is *(time-)homogenous* if distribution of $(X^{(t)} | X^{(t-1)})$ is equal for all t
- A Markov chain is *irreducible* if every state in S can be reached from any state
- A state has *period* k if multiples of k steps are necessary to return to it
- A Markov chain is *aperiodic* if each state has period 1



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