

An empirical investigation of peremptory challenge

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1 Introduction

The Gerald Stanley murder trial was noteworthy for all of the wrong reasons. The first reason was the crime itself. The rural region around Biggar, Saskatchewan[8] is not known for crime, indeed, the crime statistics collected by Statistics Canada suggest it is one of the safest in the province[11]. Any murder at all would be worthy of attention and subject to plenty of drama. But beyond the damage this trial has done to the community, this trial is noteworthy because it led to a significant re-examination of the legal jurisprudence surrounding the jury selection process culminating in the proposition of Bill C-75 by the Canadian government in March of 2018[1], less than two months after the trial's verdict[9].

Bill C-75, in part, aims to ameliorate one of the critical points of contention about the Gerald Stanley case: the use of peremptory challenges in jury selection. The outsized impact of the case was due, in large part, to its racial aspect. Gerald Stanley, a white man, was accused of second degree murder in the killing of Colten Boushie, a First Nations man. Given Canada's troubled history with First Nations groups, this alone would have been enough to make the trial a flash point for race issues, but that was not the worst aspect of the trial. Rather, it was the alleged use of peremptory challenges to strike five potential jurors who "appeared" to be First Nations, resulting in an all-white jury, that proved to be the most controversial and influential facet of the entire affair[5] [7].

With Bill C-75 currently moving through the Canadian parliamentary system, having completed its second reading in June 2018[2], a close re-examination of the practice of peremptory challenge is warranted. A great deal of ink has already been spilled on both sides of the debate [6] [13] [10], but startlingly little of this discussion has been based on any hard evidence on the impact of peremptory challenge in jury selection. This paper aims to provide analysis and evidence to illuminate the topic further by analyzing three separate peremptory challenge data sets collected in the United States [12] [4] [3]. While this data cannot tell us if challenges were racially motivated in the Stanley trial, stepping back from this fraught legal episode to take a wider view of the practice of peremptory challenge provides a more sober place to start the discussion of its place in modern jury trials.

This paper will proceed in five parts. Section 2 provides a brief history of the practice of peremptory challenges in jury trials, in particular explaining their original motivation and past implementations in 2.1 and how they have developed in the United States, the United Kingdom, and Canada in 2.2. Section 3 proceeds to discuss the three data sets obtained, with 3.1 – 3.3 discussing the sources and collection methods before the cleaning and preprocessing are explained in 3.4. Section 4 then provides the details and results of the different analyses performed on the different data sets, before these results are compared to previous works in Section 5. Finally, the results and findings are summarized in 6, and recommendations based on the observations obtained here are provided.

2 Background

2.1 History of Peremptory Challenge

2.2 Modern Practice

3 Data

3.1 Jury Sunshine Project

3.2 North Carolina Data

3.3 Philadelphia Data

3.4 Data Cleaning

Jury Sunshine Data

The data collected in North Carolina proved invaluable to this project [12].

Problem: some columns of the data contained only NA values Solution: `lapply` to remove these uninformative columns

Problem: relational database provided did not have all data in one joined table Solution: creation of `CleaningMerge` function: a wrapper for `merge` which provides information about the mismatches which may be present in the two merged tables

Problem: inconsistently coded levels, e.g. inconsistent case or “?” instead of “U” for unknowns Solution: forcing levels to be uppercase and the replacement of obvious mis-specified levels

Problem: some columns seem to have swapped values, e.g. the gender column should be one of “M”, “F”, or “U” and the political affiliation column should be one of “D”, “R”, “I”, or “U”, but some individuals have the gender recorded as “R” and political affiliation as “M” Solution: the creation of the `IdentifySwap` function, which has two arguments: a data set and the acceptable or correct levels for the variables in the data set. It then identifies rows which have candidate swaps and presents them for review

4 Analysis

4.1 Modelling

In order to create a single model to test the statistical significance of the differences observed for strike rates by race, defendant race, and party doing the striking, a saturated poisson regression model was fit to the data. Letting i denote the level of the venire member race, j the defendant race, and k the disposition, the numbers of observed venire members in each ijk combination, y_{ijk} were modelled as Poisson-distributed random variables with expectation λ_{ijk} . A saturated model was then fit to the data, that is a model described by the equation:

$$\log E[y_{ijk}] = \mathbf{x}_{ijk}\beta = \beta_o + \beta_R x_{i..} + \beta_D x_{.j.} + \beta_S x_{..k} + \beta_{R:D} x_{i..} x_{.j.} + \beta_{R:S} x_{i..} x_{..k} + \beta_{D:S} x_{.j.} x_{..k} + \beta_{R:D:S} x_{i..} x_{.j.} x_{..k}$$

Where $x_{i..}$ indicates the race level of the ijk cell, and $x_{.j.}, x_{..k}$ are defined analogously for the defendant race and disposition. The interaction terms then serve to answer questions about the racial pattern of strikes which is utilized by each party given the defendant race. Most interesting to this investigation is the third order interaction term. This term indicates a significant difference in racial strike patterns given the party striking and the defendant race. In other words, this term accounts for different patterns for the different parties which are not independent of the defendant race.

When this term is tested using a nested model without the third order interaction, the third order interaction is found to be significant. This suggests that not only do the patterns present in the different parties vary, but they vary differently for different defendant races. This dependence can be viewed using a novel graphic presented in Figure 1.

The conditional probability of a particular disposition given the racial combination of venire person and defendant is displayed on the y-axis, that is the count of individuals for a particular race, defendant race, and disposition

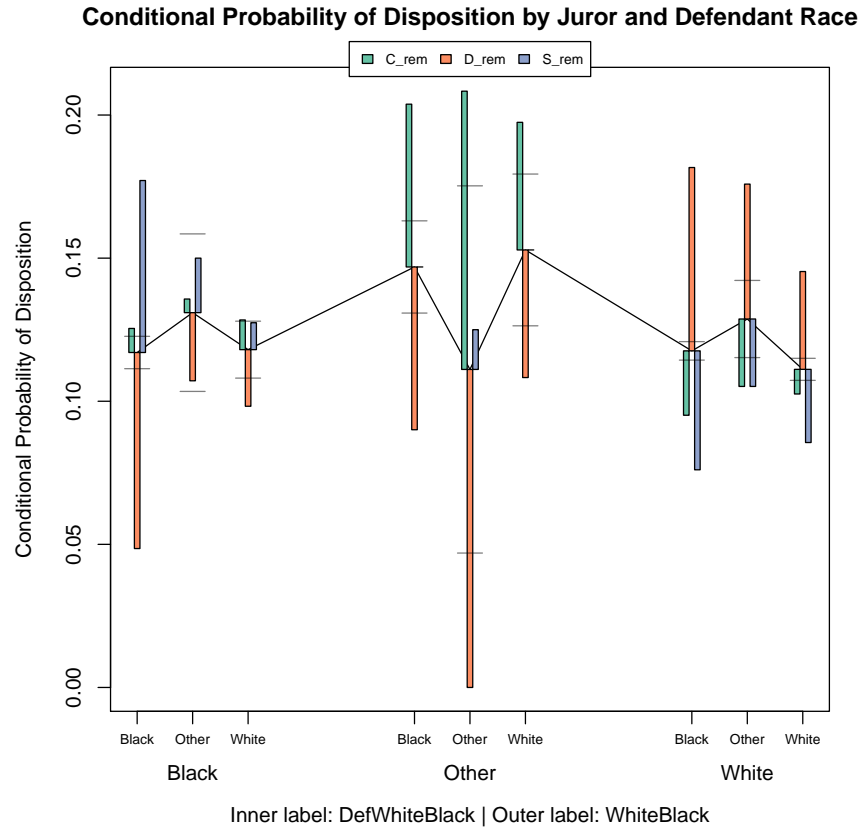


Figure 1: Parallel coordinate plot of racial strike tendencies

combination divided by the number of individuals with the racial combination across all dispositions. The x-axis then displays the combinations, grouped by the venire member race to show the dominant pattern in the data.

The black line running across the plot is the mean, or expected, rejection probability that all parties would have if they acted identically. That is, the relative level of this line provides the relative strike rate on aggregate for a particular racial combination. The bars extending from this line at each point go from this line to the corresponding value of the party represented by the bar. Finally, the horizontal lines provide approximate confidence intervals for each combination¹.

The dominant pattern to these strikes is a tendency of the defense to preferentially reject white venire members and keep black venire members, and of the prosecution to do the opposite. It was already noted in the literature[12], but the addition of defendant race allows us to make a stronger statement, as this pattern remains across defendant races. It also adds nuance, however, as the race of the defendant has a clear impact on the lengths of the bars for both the defense and prosecution. The prosecution seems to favour a jury which does not match the race of the defendant, while the defense seems to favour a jury which does.

While this second tendency seems to have no justification beyond race, the dominant tendency may have other justification than simply skin colour. As was noted by “Ideological Imbalance and Peremptory Challenge”, black individuals are more consistently aligned with the democratic party, and as a consequence a lawyer which suspects this political bias will impact the trial outcome would preferentially strike or keep black jurors in order to keep as many left wing individuals as possible. In this data, this political imbalance is incredibly prevalent, as can be seen in Figure 2 **Add the plot of this effect here, elaborate on this pattern more based on the plot.**

Perhaps more interestingly, the prosecution and judge seem to match in their tendency from the mean at every combination. This suggests that both challenges with cause and the prosecution tend to have the same effect on the

¹Generated assuming a binomial distribution of struck (by any party) against kept, as when this data is modelled with a poisson distribution, the distribution of sub-processes given the overall count will be binomially distributed

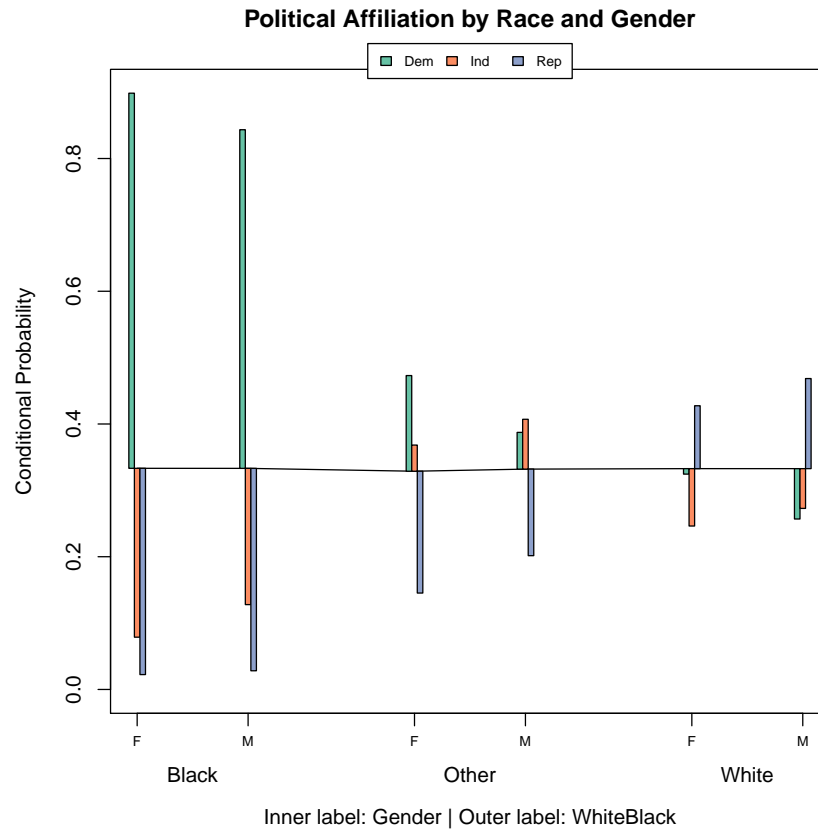


Figure 2: Conditional probabilities of political affiliation by race and gender

jury composition, though the magnitudes can differ greatly for these two strikes.

5 Comparison to Previous Work

6 Conclusions and Recommendations

7 Ideas

- look at the CSI from StatsCan, or an analogous US value, to assess the severity of a crime
- Kullback-Leibler divergence of accepted jury distribution to the venire distribution
- Look at guilty verdict tendencies based on jury race vs. defendant race

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