

# Gaussian Mixture Models and the EM Algorithm

Machine Learning (CE 40717) — Spring 2025

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# 1 Overview

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# Mixture Models

$$\mathbb{P}(\mathbf{x} | \boldsymbol{\theta}) = \sum_{j=1}^K \pi_j \mathbb{P}(\mathbf{x} | z=j; \boldsymbol{\theta}_j), \quad 0 \leq \pi_j \leq 1, \sum_j \pi_j = 1$$

Gaussian Mixture Model (GMM):

$$\mathbb{P}(\mathbf{x}) = \sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j),$$

where

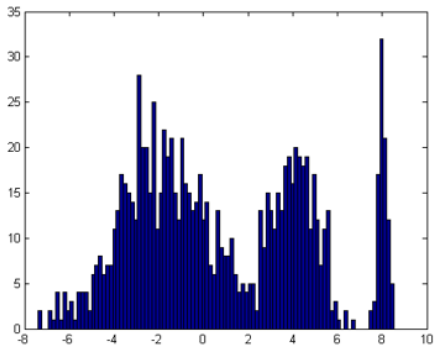
$$\mathcal{N}(\mathbf{x} | \boldsymbol{\mu}, \boldsymbol{\Sigma}) = \frac{1}{(2\pi)^{d/2} |\boldsymbol{\Sigma}|^{1/2}} \exp\left(-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\right).$$

# Why GMMs?

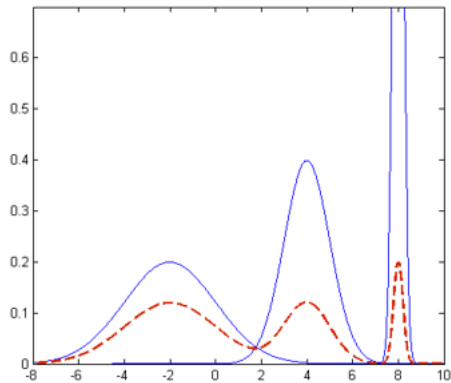
- Can model multi-modal densities beyond k-means.
- Provide a probabilistic generative framework.
- Parameters:  $\boldsymbol{\theta} = \{\pi_j, \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j\}_{j=1}^K$ .
- Maximum likelihood has no closed-form  $\rightarrow$  use EM.

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# 1-D GMM Examples

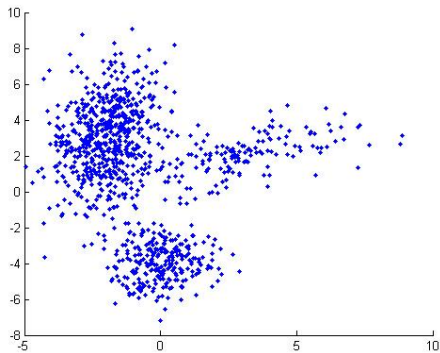


(1)

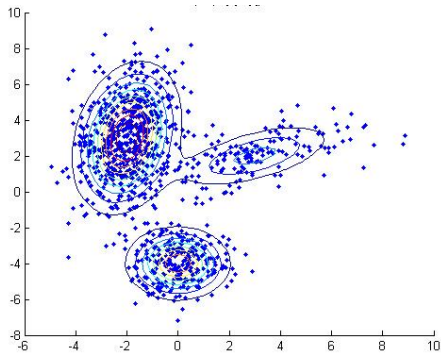


(2)

## 2-D GMM Examples



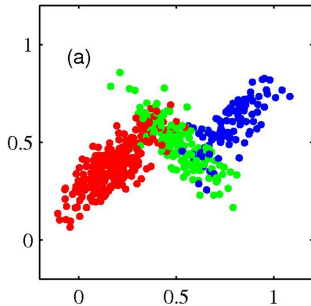
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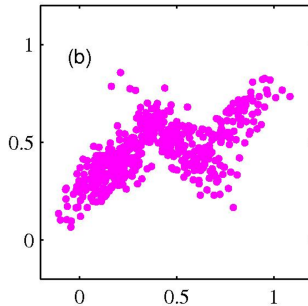
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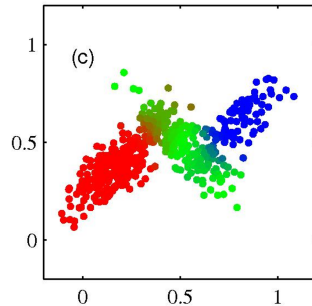
# EM & GMM Examples (Set 1)



(a)

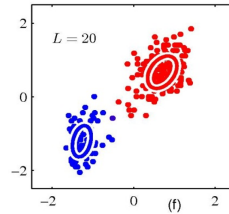
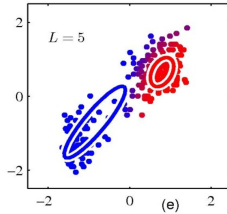
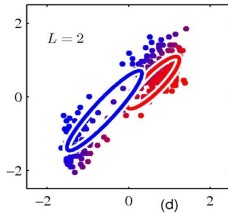
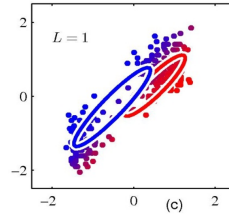
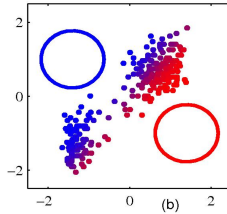
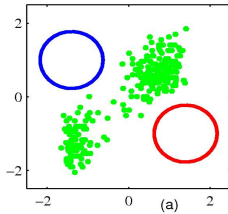


(b)

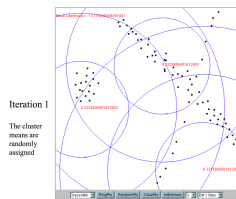


(c)

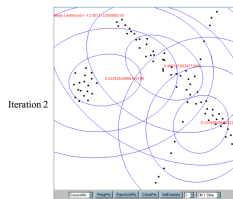
# EM & GMM Examples (Set 2)



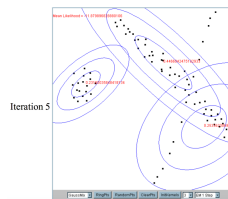
# EM & GMM Iterations



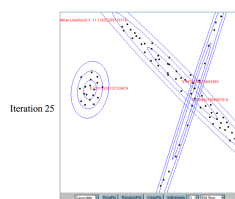
Iter 1



Iter 2



Iter 5



Iter 25

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# Maximum Likelihood for GMM

Given data  $\mathcal{X} = \{\mathbf{x}^{(i)}\}_{i=1}^N$ :

$$\ell(\boldsymbol{\theta}) = \ln \mathbb{P}(\mathcal{X} \mid \boldsymbol{\theta}) = \sum_{i=1}^N \ln \left( \sum_{j=1}^K \pi_j \mathcal{N}(\mathbf{x}^{(i)} \mid \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j) \right).$$

No closed form due to the inner summation inside the log.

# Complete-data Likelihood

Introduce latent variables  $Z = \{z^{(i)}\}$ :

$$\mathbb{P}(\mathcal{X}, Z | \boldsymbol{\theta}) = \prod_{i=1}^N \prod_{j=1}^K [\pi_j \mathcal{N}(\mathbf{x}^{(i)} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j)]^{z_j^{(i)}}.$$

Thus:

$$\ln \mathbb{P}(\mathcal{X}, Z | \boldsymbol{\theta}) = \sum_{i=1}^N \sum_{j=1}^K z_j^{(i)} \left( \ln \pi_j + \ln \mathcal{N}(\mathbf{x}^{(i)} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j) \right).$$

# E-step: Responsibilities

$$\gamma_j^{(i)} = \mathbb{P}(z^{(i)} = j | \mathbf{x}^{(i)}, \boldsymbol{\theta}^{(t)}) = \frac{\pi_j^{(t)} \mathcal{N}(\mathbf{x}^{(i)} | \boldsymbol{\mu}_j^{(t)}, \boldsymbol{\Sigma}_j^{(t)})}{\sum_{k=1}^K \pi_k^{(t)} \mathcal{N}(\mathbf{x}^{(i)} | \boldsymbol{\mu}_k^{(t)}, \boldsymbol{\Sigma}_k^{(t)})}.$$

Interpretation: soft assignment of each  $\mathbf{x}^{(i)}$  to component  $j$ .

# M-step: Maximization of Expected Log-likelihood

Define expected complete log-likelihood:

$$Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)}) = \mathbb{E}_{Z|\mathcal{X}, \boldsymbol{\theta}^{(t)}} [\ln \mathbb{P}(\mathcal{X}, Z | \boldsymbol{\theta})].$$

Substitute expectations:

$$Q(\boldsymbol{\theta}; \boldsymbol{\theta}^{(t)}) = \sum_{i=1}^N \sum_{j=1}^K \gamma_j^{(i)} \left( \ln \pi_j + \ln \mathcal{N}(\mathbf{x}^{(i)} | \boldsymbol{\mu}_j, \boldsymbol{\Sigma}_j) \right).$$



# Update for $\mu_j$

Maximize  $Q$  w.r.t.  $\mu_j$ :

$$\frac{\partial Q}{\partial \mu_j} = \sum_{i=1}^N \gamma_j^{(i)} \Sigma_j^{-1} (\mathbf{x}^{(i)} - \mu_j) = 0$$

giving:

$$\mu_j^{\text{new}} = \frac{1}{N_j} \sum_{i=1}^N \gamma_j^{(i)} \mathbf{x}^{(i)}, \quad N_j = \sum_{i=1}^N \gamma_j^{(i)}.$$

# Update for $\Sigma_j$

$$\Sigma_j^{\text{new}} = \frac{1}{N_j} \sum_{i=1}^N \gamma_j^{(i)} (\mathbf{x}^{(i)} - \boldsymbol{\mu}_j^{\text{new}})(\mathbf{x}^{(i)} - \boldsymbol{\mu}_j^{\text{new}})^\top.$$

(In practice add small  $\epsilon I$  for numerical stability.)

# Update for $\pi_j$

Maximize under  $\sum_j \pi_j = 1$ :

$$\pi_j^{\text{new}} = \frac{N_j}{N}.$$

Using Lagrange multiplier ensures normalization.

# Algorithm Summary

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## Algorithm 1 EM for Gaussian Mixture Models

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- 1: Initialize  $\{\pi_j^{(0)}, \boldsymbol{\mu}_j^{(0)}, \boldsymbol{\Sigma}_j^{(0)}\}$
  - 2: **repeat**
  - 3:     **E-step:** Compute  $\gamma_j^{(i)}$
  - 4:     **M-step:** Update parameters as above
  - 5: **until** log-likelihood convergence
-

# Variational View of EM

$$\ln \mathbb{P}(\mathcal{X} \mid \boldsymbol{\theta}) = F(\boldsymbol{\theta}, Q) + \text{KL}(Q(Z) \parallel \mathbb{P}(Z \mid \mathcal{X}, \boldsymbol{\theta})),$$

where

$$F(\boldsymbol{\theta}, Q) = \sum_Z Q(Z) \ln \frac{\mathbb{P}(\mathcal{X}, Z \mid \boldsymbol{\theta})}{Q(Z)}.$$

EM maximizes  $F$  alternately over  $Q$  and  $\boldsymbol{\theta}$ .

# EM Monotonicity

$$\ln \mathbb{P}(\mathcal{X} \mid \boldsymbol{\theta}^{(t+1)}) \geq F(\boldsymbol{\theta}^{(t+1)}, Q) \geq F(\boldsymbol{\theta}^{(t)}, Q) = \ln \mathbb{P}(\mathcal{X} \mid \boldsymbol{\theta}^{(t)}).$$

Thus the likelihood never decreases.

# Practical Notes

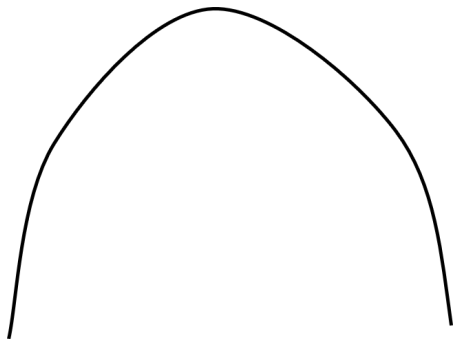
- Initialization via k-means.
- Regularize  $\Sigma_j \leftarrow \Sigma_j + \epsilon I$ .
- Convergence: monitor  $\Delta\ell < 10^{-6}$ .

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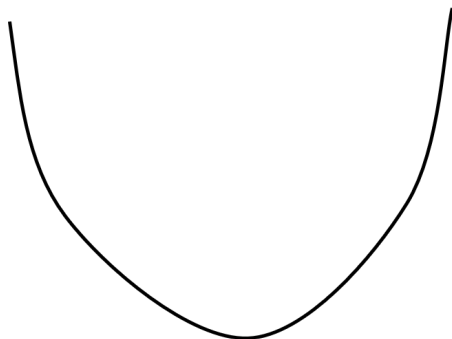


# Convergence and Local Optima

- EM converges to a stationary (local) optimum.
- Run multiple initializations.



(Left) Concave function



(Right) Convex function

# k-means vs EM+GMM

- k-means: hard assignments, equal variance clusters.
- EM: soft probabilistic assignments.
- EM generalizes k-means when covariances are isotropic.

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# Summary

- GMM = flexible probabilistic mixture model.
- EM = iterative ML method with guaranteed non-decrease.
- Use careful initialization and covariance regularization.

# References

- C. M. Bishop, *Pattern Recognition and Machine Learning*, Ch. 9.
- Lecture slides: Hamid R. Rabiee & Zahra Dehghanian, Spring 2025.

# Contributions

- **This slide deck was prepared thanks to:**
  - Soheil Sayah Varg