UVA CS 6316: Machine Learning

Lecture 14 Extra: More about Logistic Regression

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Today: Extra



- Bayes Classifier
- Expected Prediction Error
- 0-1 Loss function for Bayes Classifier
- ✓ Logistic regression

Parameter Estimation for LR

e BX 1+eBX B

- Treat each feature attribute and the class label as random variables.
- Given a sample **x** with attributes $(x_1, x_2, ..., x_p)$:
 - Goal is to predict its class C.
 - Specifically, we want to find the value of C_i that maximizes $p(C_i \mid x_1, x_2, ..., x_p)$.

• Can we estimate $p(C_i | \mathbf{x}) = p(C_i | x_1, x_2, ..., x_p)$ directly from data?

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- → MAP classification rule
 - Establishing a probabilistic model for classification
 - → MAP classification rule
 - MAP: Maximum A Posterior
 - Assign x to c^* if

$$P(C = c^* | \mathbf{X} = \mathbf{x}) > P(C = c | \mathbf{X} = \mathbf{x})$$

for $c \neq c^*$, $c = c_1, \dots, c_L$

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Bayes Classifiers – MAP Rule

Task: Classify a new instance X based on a tuple of attribute values $X = \langle X_1, X_2, ..., X_p \rangle$ into one of the classes

$$c_{MAP} = \underset{c_j \in C}{\operatorname{argmax}} P(c_j \mid x_1, x_2, \dots, x_p)$$

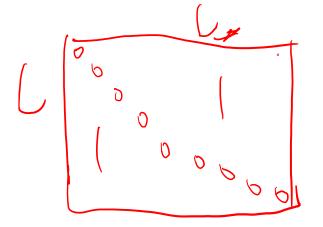
WHY?

MAP = Maximum Aposteriori Probability

0-1 LOSS for Classification

• Procedure for categorical output variable C

- Frequently, 0-1 loss function used: $L(k, \ell)$
- $L(k, \ell)$ is the price paid for misclassifying an element from class C_k as belonging to class C_ℓ
 - $\rightarrow L*L \ matrix$



C1, C2, ..., CL.

Expected prediction error (EPE)

• Expected prediction error (EPE), with expectation taken w.r.t. the joint distribution Pr(C,X)

•
$$Pr(C,X) = Pr(C \mid X) Pr(X)$$

$$\nearrow \emptyset.\emptyset. 0 - (\emptyset)$$

$$EPE(f) = E_{X,C}(L(C,f(X)))$$

$$E_{X,C}(X) = \sum_{X} (X) Pr(X)$$

$$E_{X,C}(X) = \sum_{X} (X) Pr(X)$$

$$= E_X \sum_{k=1}^{L} L[C_k, f(X)] Pr(C_k | X)$$

Consider sample population distribution

$$EPE(f) = E_{X,C} \left(L\left(C,f(X)\right) \right)$$

$$= E_{X} E_{C|X} \left[L\left(C,f(X)\right) \mid X \right]$$

$$\Rightarrow \int (X = X) = L\left[C_{X},f(X) \right] P_{Y}(C_{X} \mid X = X)$$

$$\Rightarrow \int (X = X) = a_{X} m_{X} P_{Y}\left(C_{X} \mid X = X\right)$$

$$\Rightarrow \int (X = X) = a_{X} m_{X} P_{Y}\left(C_{X} \mid X = X\right)$$

$$= P(C_{X} \mid X)$$

$$= P(C_{X} \mid X = X)$$

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Expected prediction error (EPE)

$$EPE(f) = E_{X,C}(L(C, f(X))) = E_X \sum_{k=1}^{K} L[C_k, f(X)]Pr(C_k | X)$$

Consider sample population distribution

• Pointwise minimization suffices

•
$$\Rightarrow$$
 simply $\hat{f}(X) = \operatorname{argmin}_{g \in C} \sum_{k=1}^{K} L(C_k, g) \Pr(C_k | X = X)$



$$\hat{f}(X) = C_k$$
 if

$$Pr(C_k | X = x) = \max_{g \in C} Pr(g | X = x)$$

SUMMARY: WHEN Expected prediction error (EPE) USES DIFFERENT LOSS

Loss Function	Estimator $\hat{f}(x)$
L_2 \leftarrow E_2	$\widehat{f}(x) = E[Y X = x]$
L_1 $\overset{L(\epsilon)}{\longleftarrow} \epsilon$	$\widehat{f}(x) = \text{median}(Y X=x)$
$\begin{array}{c c} \bullet L(\epsilon) \\ \hline 0-1 & \hline \\ \hline -\delta & \delta \end{array}$	$\widehat{f}(x) = \arg\max_{Y} P(Y X=x)$ (Bayes classifier / MAP)

Today: Extra

- ✓ Why Bayes Classification MAP Rule?
 - Expected Prediction Error
 - 0-1 Loss function for Bayes Classifier
- ✓ Logistic regression

Parameter Estimation for LR

 $\frac{e^{\beta x}}{1 + e^{\beta x}}$ of for LR

Newton's method for optimization

- The most basic second-order optimization algorithm
- Updating parameter with

Mewton:
$$oldsymbol{ heta}_{k+1} = oldsymbol{ heta}_k - \mathbf{H}_K^{-1} \mathbf{g}_k$$

Review: Hessian Matrix / n==2 case

Singlevariate

→ multivariate

f(x,y)

• 1st derivative to gradient,

$$g = \nabla f = \begin{pmatrix} \frac{\partial f}{\partial x} \\ \frac{\partial f}{\partial y} \end{pmatrix}$$

• 2nd derivative to Hessian

$$H = \begin{pmatrix} \frac{\partial^2 f}{\partial x^2} & \frac{\partial^2 f}{\partial x \partial y} \\ \frac{\partial^2 f}{\partial x \partial y} & \frac{\partial^2 f}{\partial y^2} \end{pmatrix}$$

Review: Hessian Matrix

Suppose that $f: \mathbb{R}^n \to \mathbb{R}$ is a function that takes a vector in \mathbb{R}^n and returns a real number. Then the **Hessian** matrix with respect to x, written $\nabla_x^2 f(x)$ or simply as H is the $n \times n$ matrix of partial derivatives,

$$\nabla_x^2 f(x) \in \mathbb{R}^{n \times n} = \begin{bmatrix} \frac{\partial^2 f(x)}{\partial x_1^2} & \frac{\partial^2 f(x)}{\partial x_1 \partial x_2} & \cdots & \frac{\partial^2 f(x)}{\partial x_1 \partial x_n} \\ \frac{\partial^2 f(x)}{\partial x_2 \partial x_1} & \frac{\partial^2 f(x)}{\partial x_2^2} & \cdots & \frac{\partial^2 f(x)}{\partial x_2 \partial x_n} \\ \vdots & \vdots & \ddots & \vdots \\ \frac{\partial^2 f(x)}{\partial x_n \partial x_1} & \frac{\partial^2 f(x)}{\partial x_n \partial x_2} & \cdots & \frac{\partial^2 f(x)}{\partial x_n^2} \end{bmatrix}.$$

Newton's method for optimization

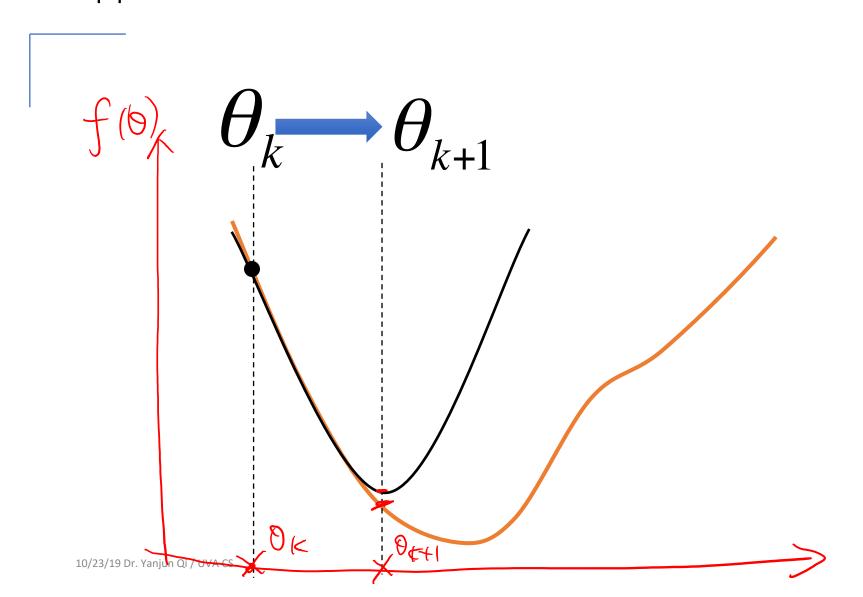
• Making a quadratic/second-order Taylor series approximation

$$\mathbf{f}_{quad}(\boldsymbol{\theta}) = f(\boldsymbol{\theta}_k) + \mathbf{g}_k^T(\boldsymbol{\theta} - \boldsymbol{\theta}_k) + \frac{1}{2}(\boldsymbol{\theta} - \boldsymbol{\theta}_k)^T\mathbf{H}_k(\boldsymbol{\theta} - \boldsymbol{\theta}_k)$$

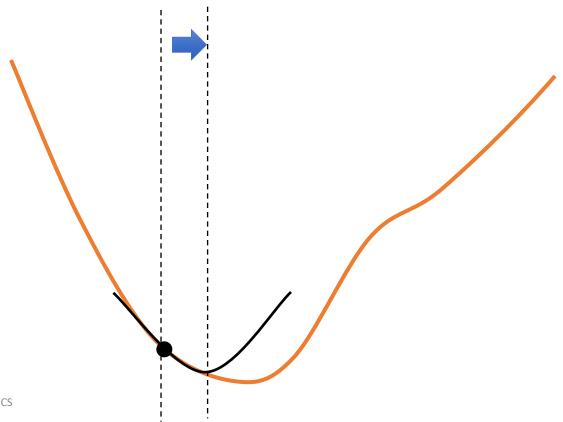
Finding the minimum solution of the above right quadratic approximation (quadratic function minimization is easy!)

$$\widehat{J}(0) = \widehat{J}(0) + \widehat{J}_{K}^{T}(0 - 0_{K}) + \frac{1}{2}(0 - 0_{K}) + \frac{1}{2}(0 - 0_{K})^{T} + \frac{$$

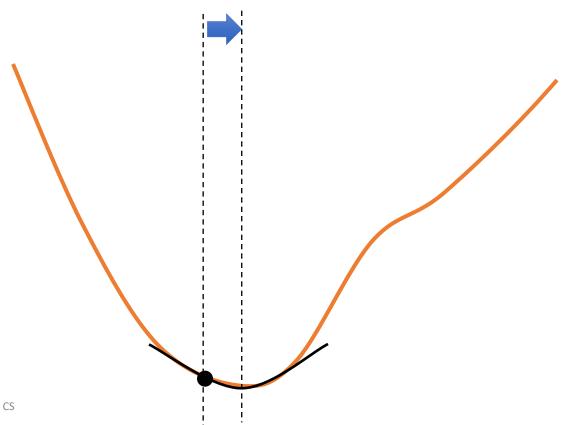
Newton's Method / second-order Taylor series approximation



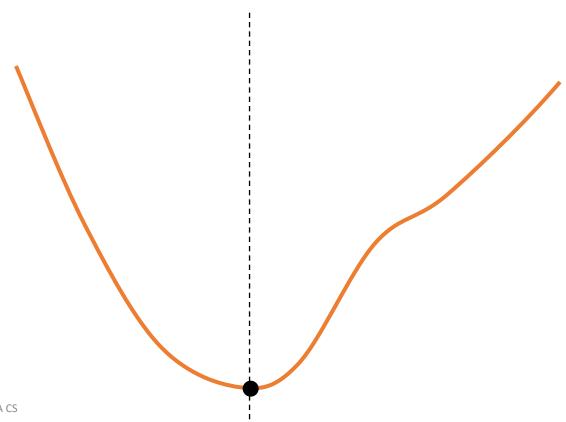
Newton's Method / second-order Taylor series approximation



Newton's Method / second-order Taylor series approximation



Newton's Method / second-order Taylor series approximation



Newton's Method

• At each step:

$$\theta_{k+1} = \theta_k - \frac{f'(\theta_k)}{f''(\theta_k)}$$

$$\theta_{k+1} = \theta_k - H^{-1}(\theta_k) \nabla f(\theta_k)$$

- Requires 1st and 2nd derivatives
- Quadratic convergence
- However, finding the inverse of the Hessian matrix is often expensive

Newton vs. GD for optimization

Newton: a quadratic/second-order Taylor series approximation

$$\mathbf{f}_{quad}(oldsymbol{ heta}) = f(oldsymbol{ heta}_k) + \mathbf{g}_k^T(oldsymbol{ heta} - oldsymbol{ heta}_k) + rac{1}{2}(oldsymbol{ heta} - oldsymbol{ heta}_k)^T \mathbf{H}_k(oldsymbol{ heta} - oldsymbol{ heta}_k)$$

Finding the minimum solution of the above right quadratic approximation (quadratic function minimization is easy!)

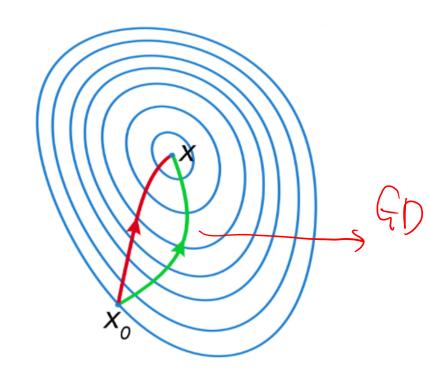
$$\begin{aligned} \mathbf{f}_{quad}(\boldsymbol{\theta}) &= f(\boldsymbol{\theta}_k) + \mathbf{g}_k^T(\boldsymbol{\theta} - \boldsymbol{\theta}_k) + \frac{1}{2}(\boldsymbol{\theta} - \boldsymbol{\theta}_k)^T \frac{1}{\alpha}(\boldsymbol{\theta} - \boldsymbol{\theta}_k) \\ & & & & & \\ \mathbf{g}(\boldsymbol{\theta}) &= \mathbf{g}(\boldsymbol{\theta}_k) \end{aligned}$$

Comparison

Newton's method vs. Gradient descent

A comparison of gradient descent (green) and Newton's method (red) for minimizing a function (with small step sizes).

Newton's method uses curvature information to get a more direct route ...



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1.0/23/19 . Tanjun Qi / UVA CS

MLE for Logistic Regression Training

Let's fit the logistic regression model for K=2, i.e., number of classes is 2

For Bernoulli distribution

Training set: (x_i, y_i) , i=1,...,N

 $p(y \mid x)^{y} (1-p)^{1-y}$

(conditional)
Log-likelihood:
$$l(\beta) = \sum_{i=1}^{N} \{ \log \Pr(Y = y_i | X = x_i) \}$$

$$= \sum_{i=1}^{N} y_i \log (\Pr(Y = 1 | X = x_i)) + (1 - y_i) \log (\Pr(Y = 0 | X = x_i))$$

$$= \sum_{i=1}^{N} (y_i \log \frac{\exp(\beta^T x_i)}{y_i}) + (1 - y_i) \log \frac{1}{y_i}$$

$$= \sum_{i=1}^{N} (y_i \log \frac{\exp(\beta^T x_i)}{1 + \exp(\beta^T x_i)}) + (1 - y_i) \log \frac{1}{1 + \exp(\beta^T x_i)})$$

$$= \sum_{i=1}^{N} (y_i \beta^T x_i - \log(1 + \exp(\beta^T x_i)))$$

 x_i are (p+1)-dimensional input vector with leading entry 1 \beta is a (p+1)-dimensional vector

$$I(\beta) = \sum_{i=1}^{N} \{ \log \Pr(Y = y_i | X = x_i) \}$$

$$\int_{0}^{N} \{ |Y = y_i | X = x_i \} \}$$

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Newton-Raphson for LR (optional)

$$\frac{\partial l(\beta)}{\partial \beta} = \sum_{i=1}^{N} (y_i - \frac{\exp(\beta^T x)}{1 + \exp(\beta^T x)}) x_i = 0$$

(p+1) Non-linear equations to solve for (p+1) unknowns

Solve by Newton-Raphson method:

$$\beta^{new} \leftarrow \beta^{old} - \left[\left(\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^T} \right) \right]^{-1} \frac{\partial l(\beta)}{\partial \beta},$$

where,
$$\left(\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^T}\right) = -\sum_{i=1}^N x_i x_i^T \left(\frac{\exp(\beta^T x_i)}{1 + \exp(\beta^T x_i)}\right) \left(\frac{1}{1 + \exp(\beta^T x_i)}\right)$$

minimizes a quadratic approximation to the function we are really interested in.

$$oldsymbol{ heta}_{k+1} = oldsymbol{ heta}_k - \mathbf{H}_K^{-1} \mathbf{g}_k$$

 $p(x_i; \beta)$ 1 - $p(x_i; \beta)$

Newton-Raphson for LR...

$$\frac{\partial l(\beta)}{\partial \beta} = \sum_{i=1}^{N} (y_i - \frac{\exp(\beta^T x)}{1 + \exp(\beta^T x)}) x_i = X^T (y - p)$$

$$(\frac{\partial^2 l(\beta)}{\partial \beta \partial \beta^T}) = -X^T W X$$

So, NR rule becomes:

$$\beta^{new} \leftarrow \beta^{old} + (X^T W X)^{-1} X^T (y - p),$$

$$X = \begin{bmatrix} x_1^T \\ x_2^T \\ \vdots \\ x_N^T \end{bmatrix}_{N-by-(p+1)}, y = \begin{bmatrix} y_1 \\ y_2 \\ \vdots \\ y_N \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \exp(\beta^T x_2)/(1 + \exp(\beta^T x_2)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \exp(\beta^T x_2)/(1 + \exp(\beta^T x_2)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_1)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp(\beta^T x_N)) \end{bmatrix}_{N-by-1}, p = \begin{bmatrix} \exp(\beta^T x_1)/(1 + \exp(\beta^T x_N)) \\ \vdots \\ \exp(\beta^T x_N)/(1 + \exp($$

 $X: N \times (p+1)$ matrix of x_i

 $y: N \times 1$ matrix of y_i

 $p: N \times 1$ matrix of $p(x_i; \beta^{old})$

$$W: N \times N \text{ diagonal matrix of } p(x_i; \beta^{old})(1 - p(x_i; \beta^{old}))$$

$$(\frac{\exp(\beta^T x_i)}{(1 + \exp(\beta^T x_i))})(1 - \frac{1}{(1 + \exp(\beta^T x_i))})$$

Newton-Raphson for LR...

Newton-Raphson

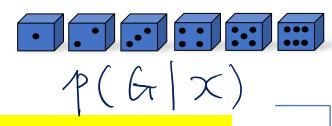
Iteratively reweighted least squares (IRLS)

 $z = X\beta^{old} + W^{-1}(v - p)$

$$\beta^{new} \leftarrow \arg\min_{\beta} (z - X\beta^{T})^{T} W (z - X\beta^{T})$$

$$\leftarrow \arg\min_{\beta} (y - p)^{T} W^{-1} (y - p)$$

Binary - Multinoulli Logistic Regression Model



Directly models the posterior probabilities as the output of regression

$$\Pr(G = k \mid X = x) = \frac{\exp(\beta_{k0} + \beta_k^T x)}{1 + \sum_{l=1}^{K-1} \exp(\beta_{l0} + \beta_l^T x)}, \ k = 1, ..., K-1$$

$$\Pr(G = K \mid X = x) = \frac{1}{1 + \sum_{l=1}^{K-1} \exp(\beta_{l0} + \beta_l^T x)}$$

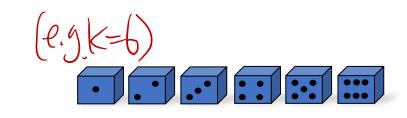
x is p-dimensional input vector

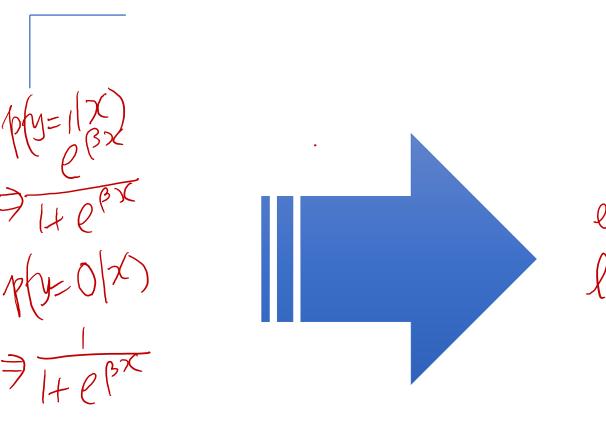
 β_{ν}^{T} is a p-dimensional vector for each class k

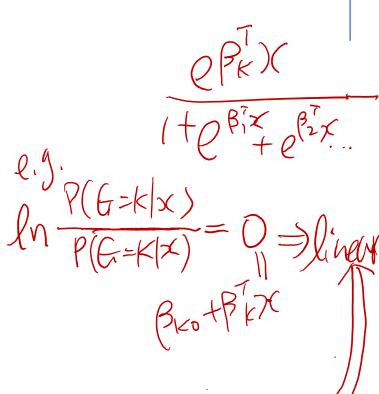
Total number of parameters is (K-1)(p+1) (K-1)(p+1) (K-1)(p+1)

Note that the class boundaries are linear

Binary Multinoulli Logistic Regression Model







Note that the class boundaries are linear

References

- Prof. Tan, Steinbach, Kumar's "Introduction to Data Mining" slide
- ☐ Prof. Andrew Moore's slides
- ☐ Prof. Eric Xing's slides
- ☐ Prof. Ke Chen NB slides
- ☐ Hastie, Trevor, et al. *The elements of statistical learning*. Vol. 2. No.
 - 1. New York: Springer, 2009.