

# Updating a Movements Account

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June 8, 2017

## 1 Definitions

Assume that account contains age, and that eage age group or period has length 1.

- Let  $N$  be population.
- Let  $C_1, \dots, C_K$  be components of the demographic accounts such as migration, births, or deaths. All  $C_k$  are compatible with  $N$ .
- Let  $Q = \{N, C_1, \dots, C_K\}$  denote the demographic account.
- Let  $A$  be accession.  $A$  does not include age 0 (i.e. births).
- Let  $E$  be exposure, derived deterministically from  $N$ .
- Let  $\gamma^N$  be expected counts from the system model for  $N$ .
- Let  $\gamma_k^C$  be rates or expected counts from the system model for  $C_k$ .
- Let  $X_1, \dots, X_M$  be the datasets used to estimate the demographic account. If  $X_m$  is used to estimate  $N$ , then  $N$  can be collapsed to be exactly compatible with  $X_m$ ; let  $N_m = N^{[X_m]}$  denote the collapsed version of  $N$ . If  $X_m$  is used to estimate  $C_k$ , then  $C_k$  can be collapsed to be exactly compatible with  $X_m$ ; let  $C_{km} = C_k^{[X_m]}$  denote the collapsed version of  $C_k$ .
- Age group  $a$  is the age interval between exact ages  $a$  and  $a + 1$ .
- Period  $t$  is the interval between exact times  $t - 1$  and  $t$ .
- Let  $\hat{N} = \gamma^N$  be expected population.
- Let  $\hat{E}$  be expected exposure.  $\hat{E}$  is calculated from  $\hat{N}$  using the formulas for calculating  $E$  from  $N$ .

## 2 System model

$$p(\mathbf{Q}) = p(\mathbf{N}) \prod_{k=1}^K p(\mathbf{C}_k | \mathbf{N}) \quad (1)$$

Population: Poisson, no exposure

Net migration, internal or external: Normal integer-only

Everything else: Poisson, typically with exposure, but can be without exposure

## 3 Quantities Required

### 3.1 Updating starting values for $\mathbf{N}$

Let  $\mathbf{N} = (\mathbf{N}_0, \mathbf{N}_{1+})$ , where  $\mathbf{N}_0$  is  $\mathbf{N}$  at time 0. Proposals for  $\mathbf{N}_0$  do not depend on the current value of  $\mathbf{N}_0$ , so

$$\frac{J(\mathbf{Q}^{(h)} | \mathbf{Q}^*)}{J(\mathbf{Q}^* | \mathbf{Q}^{(h)})} = \frac{p(n^{(h)})}{p(n^*)}. \quad (2)$$

We then have

$$\frac{p(\mathbf{Q}^*)}{p(\mathbf{Q}^{(h)})} \frac{J(\mathbf{Q}^{(h)} | \mathbf{Q}^*)}{J(\mathbf{Q}^* | \mathbf{Q}^{(h)})} = \frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} \left( \prod_{k=1}^K \frac{p(\mathbf{C}_k^{(h)} | \mathbf{N}^*)}{p(\mathbf{C}_k^{(h)} | \mathbf{N}^{(h)})} \right) \frac{p(n^{(h)})}{p(n^*)} \quad (3)$$

$$= \frac{p(\mathbf{N}_{1+}^*)}{p(\mathbf{N}_{1+}^{(h)})} \prod_{k=1}^K \frac{p(\mathbf{C}_k^{(h)} | \mathbf{N}^*)}{p(\mathbf{C}_k^{(h)} | \mathbf{N}^{(h)})}. \quad (4)$$

### 3.2 Updating component $\mathbf{C}_u$

Suppress dependence on  $\gamma$ .

Use expected population  $\hat{\gamma}^N$ , rather than actual population  $\mathbf{N}$ , to generate proposals:

$$\frac{J(\mathbf{Q}^{(h)} | \mathbf{Q}^*)}{J(\mathbf{Q}^* | \mathbf{Q}^{(h)})} = \frac{p(c^{(h)} | \hat{\mathbf{N}})}{p(c^* | \hat{\mathbf{N}})}. \quad (5)$$

Then

$$\frac{p(\mathbf{Q}^*)}{p(\mathbf{Q}^{(h)})} \frac{J(\mathbf{Q}^{(h)}|\mathbf{Q}^*)}{J(\mathbf{Q}^*|\mathbf{Q}^{(h)})} = \frac{p(\mathbf{N}^*) \prod_{k=1}^K p(\mathbf{C}_k^*|\mathbf{N}^*)}{p(\mathbf{N}^{(h)}) \prod_{k=1}^K p(\mathbf{C}_k^{(h)}|\mathbf{N}^{(h)})} \quad (6)$$

$$= \frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} \frac{p(\mathbf{C}_u^*|\mathbf{N}^*)}{p(\mathbf{C}_u^{(h)}|\mathbf{N}^{(h)})} \left( \prod_{k \neq u} \frac{p(\mathbf{C}_k^{(h)}|\mathbf{N}^*)}{p(\mathbf{C}_k^{(h)}|\mathbf{N}^{(h)})} \right) \frac{p(c^{(h)}|\hat{\mathbf{N}})}{p(c^*|\hat{\mathbf{N}})} \quad (7)$$

$$= \frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} \frac{p(\mathbf{C}_u^*|\mathbf{N}^*)}{p(\mathbf{C}_u^{(h)}|\mathbf{N}^{(h)})} \frac{p(\mathbf{C}_u^{(h)}|\mathbf{N}^{(h)})}{p(\mathbf{C}_u^{(h)}|\mathbf{N}^*)} \frac{p(\mathbf{C}_u^{(h)}|\mathbf{N}^*)}{p(\mathbf{C}_u^{(h)}|\mathbf{N}^{(h)})} \\ \times \left( \prod_{k \neq u} \frac{p(\mathbf{C}_k^{(h)}|\mathbf{N}^*)}{p(\mathbf{C}_k^{(h)}|\mathbf{N}^{(h)})} \right) \frac{p(c^{(h)}|\hat{\mathbf{N}})}{p(c^*|\hat{\mathbf{N}})} \quad (8)$$

$$= \frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} \frac{p(\mathbf{C}_u^*|\mathbf{N}^*)}{p(\mathbf{C}_u^{(h)}|\mathbf{N}^*)} \frac{p(c^{(h)}|\hat{\mathbf{N}})}{p(c^*|\hat{\mathbf{N}})} \prod_{k=1}^K \frac{p(\mathbf{C}_k^{(h)}|\mathbf{N}^*)}{p(\mathbf{C}_k^{(h)}|\mathbf{N}^{(h)})} \quad (9)$$

## 4 Cells updated directly in proposals

### 4.1 Starting population

Cell updated is  $\mathbf{N}[i, a, 0]$ . Let  $n^{(h)}$  denote current value and  $n^*$  the proposed value, and define  $\Delta = n^* - n^{(h)}$ .

### 4.2 Generic $\mathbf{C}_k$

Cell updated is  $\mathbf{C}_k[i, a, l, t]$ . Let  $c^{(h)}$  denote the current value and  $c^*$  the proposed value. Let  $\Delta = c^* - c^{(h)}$ .

### 4.3 Births

Cell updated is  $\mathbf{B}[i, a, l, t]$ . Let  $c^{(h)}$  denote the current value and  $c^*$  the proposed value. Let  $\Delta = c^* - c^{(h)}$ .

### 4.4 Internal migration with origin-destination format

Cell updated is  $\mathbf{M}[i, r_{\text{orig}}, r_{\text{dest}}, a, l, t]$ . Let  $c^{(h)}$  denote the current value and  $c^*$  the proposed value, and let  $\Delta = c^* - c^{(h)}$ .

### 4.5 Internal migration with pool format

Cells updated are  $\mathbf{M}[i, r, \text{Out}, a, l, t]$  and  $\mathbf{M}[i, r', \text{In}, a, l, t]$ ,  $r' \neq r$ . Let  $c_{\text{out}}^{(h)}$  and  $c_{\text{in}}^{(h)}$  denote the current values and  $c_{\text{out}}^*$  and  $c_{\text{in}}^*$  the proposed values. Let  $\Delta = c_{\text{out}}^* - c_{\text{out}}^{(h)} = c_{\text{in}}^* - c_{\text{in}}^{(h)}$ .

#### 4.6 Internal migration with net format

Cells updated are  $\mathbf{M}[i, r, a, l, t]$  and  $\mathbf{M}[i, r', a, l, t]$ ,  $r' \neq r$ . Let  $c_1^{(h)}$  and  $c_2^{(h)}$  denote the current values and  $c_1^*$  and  $c_2^*$  the proposed values. Let  $\Delta = c_1^* - c_1^{(h)} = -(c_2^* - c_2^{(h)})$ .

### 5 Calculating $\frac{p(\mathbf{N}_{1+}^*)}{p(\mathbf{N}_{1+}^{(h)})}$

Define cohort quantities

$$n^{(h)}(s) = \mathbf{N}^{(h)}[i, \min(a + s, A), s] \quad (10)$$

$$n^*(s) = \mathbf{N}^*[i, \min(a + s, A), s] = n^{(h)}(s) + \Delta \quad (11)$$

$$\lambda(s) = \gamma_N[i, \min(a + s, A), s] \quad (12)$$

Then

$$\frac{p(\mathbf{N}_{1+}^*)}{p(\mathbf{N}_{1+}^{(h)})} = \prod_{s=1}^T \frac{\text{Poisson}(n^*(s)|\lambda(s))}{\text{Poisson}(n^{(h)}(s)|\lambda(s))} \quad (13)$$

### 6 Calculating $\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})}$

#### 6.1 Generic component $C_k$

Define cohort quantities

$$n^{(h)}(s) = \mathbf{N}^{(h)}[i, \min(a + 1(l = \text{U}) + s, A), t + s] \quad (14)$$

$$n^*(s) = n^{(h)}(s) + S_k \Delta \quad (15)$$

$$\lambda(s) = \gamma_N[i, \min(a + 1(l = \text{U}) + s, A), t + s] \quad (16)$$

Then

$$\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} = \prod_{s=0}^{T-t} \frac{\text{Poisson}(n^*(s)|\lambda(s))}{\text{Poisson}(n^{(h)}(s)|\lambda(s))} \quad (17)$$

#### 6.2 Births

Define cohort quantities starting at age 0

$$n^{(h)}(s) = \mathbf{N}^{(h)}[i, \min(s, A), t + s] \quad (18)$$

$$n^*(s) = n^{(h)}(s) + \Delta \quad (19)$$

$$\lambda(s) = \gamma_N[i, \min(s, A), t + s] \quad (20)$$

Then

$$\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} = \prod_{s=0}^{T-t} \frac{\text{Poisson}(n^*(s)|\lambda(s))}{\text{Poisson}(n^{(h)}(s)|\lambda(s))} \quad (21)$$

### 6.3 Internal migration with origin-destination format

$$n_{\text{orig}}^{(h)}(s) = \mathbf{N}^{(h)}[i, r_{\text{orig}}, \min(a + 1(l = \text{U}) + s, A), t + s] \quad (22)$$

$$n_{\text{dest}}^{(h)}(s) = \mathbf{N}^{(h)}[i, r_{\text{dest}}, \min(a + 1(l = \text{U}) + s, A), t + s] \quad (23)$$

$$n_{\text{orig}}^*(s) = n^{(h)}(s) - \Delta \quad (24)$$

$$n_{\text{dest}}^*(s) = n^{(h)}(s) + \Delta \quad (25)$$

$$\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} = \prod_{s=0}^{T-t} \frac{\text{Poisson}(n_{\text{orig}}^*(s)|\lambda(s))}{\text{Poisson}(n_{\text{orig}}^{(h)}(s)|\lambda(s))} \frac{\text{Poisson}(n_{\text{dest}}^*(s)|\lambda(s))}{\text{Poisson}(n_{\text{dest}}^{(h)}(s)|\lambda(s))} \quad (26)$$

### 6.4 Internal migration with pool format

$$n_{\text{out}}^{(h)}(s) = \mathbf{N}^{(h)}[i, r, \min(a + 1(l = \text{U}) + s, A), t + s] \quad (27)$$

$$n_{\text{in}}^{(h)}(s) = \mathbf{N}^{(h)}[i, r', \min(a + 1(l = \text{U}) + s, A), t + s] \quad (28)$$

$$n_{\text{out}}^*(s) = n^{(h)}(s) - \Delta \quad (29)$$

$$n_{\text{in}}^*(s) = n^{(h)}(s) + \Delta \quad (30)$$

$$\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} = \prod_{s=0}^{T-t} \frac{\text{Poisson}(n_{\text{out}}^*(s)|\lambda(s))}{\text{Poisson}(n_{\text{out}}^{(h)}(s)|\lambda(s))} \frac{\text{Poisson}(n_{\text{in}}^*(s)|\lambda(s))}{\text{Poisson}(n_{\text{in}}^{(h)}(s)|\lambda(s))} \quad (31)$$

### 6.5 Internal migration with net format

$$n_1^{(h)}(s) = \mathbf{N}^{(h)}[i, r, \min(a + 1(l = \text{U}) + s, A), t + s] \quad (32)$$

$$n_2^{(h)}(s) = \mathbf{N}^{(h)}[i, r', \min(a + 1(l = \text{U}) + s, A), t + s] \quad (33)$$

$$n_1^*(s) = n^{(h)}(s) + \Delta \quad (34)$$

$$n_2^*(s) = n^{(h)}(s) - \Delta \quad (35)$$

$$\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} = \prod_{s=0}^{T-t} \frac{\text{Poisson}(n_1^*(s)|\lambda(s))}{\text{Poisson}(n_1^{(h)}(s)|\lambda(s))} \frac{\text{Poisson}(n_2^*(s)|\lambda(s))}{\text{Poisson}(n_2^{(h)}(s)|\lambda(s))} \quad (36)$$

## 7 Calculating $\frac{p(\mathbf{C}_k^*|\mathbf{N}^*)p(c^{(h)}|\hat{\mathbf{N}})}{p(\mathbf{C}_k^{(h)}|\mathbf{N}^*)p(c^*|\hat{\mathbf{N}})}$

### 7.1 Generic $\mathbf{C}_k$

No exposure Cancels

**With exposure** Let

$$e^{(h)} = \mathbf{E}^{(h)}[i, a, l, t] \quad (37)$$

$$e^* = e^{(h)} + 1(l = L)S_k \frac{1}{2}\Delta \quad (38)$$

$$\gamma = \gamma_k[i, a, l, t] \quad (39)$$

Then

$$\frac{p(\mathbf{C}_k^* | \mathbf{N}^*)p(c^* | \hat{\mathbf{N}})}{p(\mathbf{C}_k^{(h)} | \mathbf{N}^*)p(c^{(h)} | \hat{\mathbf{N}})} = \frac{\text{Poisson}(c^* | \gamma e^*) \text{Poisson}(c^{(h)} | \gamma \hat{e})}{\text{Poisson}(c^{(h)} | \gamma e^*) \text{Poisson}(c^* | \gamma \hat{e})} \quad (40)$$

$$= (e^* / \hat{e})^{(c^* - c^{(h)})} \quad (41)$$

## 7.2 Internal migration with origin-destination format

**No exposure** Cancels

**With exposure**

$$e^{(h)} = \mathbf{E}^{(h)}[i, r_{\text{orig}}, a, l, t] \quad (42)$$

$$e^* = e^{(h)} - 1(l = L)\frac{1}{2}\Delta \quad (43)$$

$$\gamma = \gamma_k[i, r_{\text{orig}}, a, l, t] \quad (44)$$

Remainder as for generic  $\mathbf{C}_k$ .

### 7.2.1 Internal migration with pool format

Two cells change, so don't get complete cancellations.

**Without exposure**

$$\gamma_{\text{out}} = \gamma_k[i, r, \text{Out}, a, l, t] \quad (45)$$

$$\gamma_{\text{in}} = \gamma_k[i, r', \text{In}, a, l, t] \quad (46)$$

$$\frac{p(\mathbf{C}_k^* | \mathbf{N}^*)p(c^{(h)} | \hat{\mathbf{N}})}{p(\mathbf{C}_k^{(h)} | \mathbf{N}^*)p(c^* | \hat{\mathbf{N}})} = \frac{\text{Poisson}(c_{\text{out}}^* | \gamma_{\text{out}}) \text{Poisson}(c_{\text{in}}^* | \gamma_{\text{in}}) \text{Poisson}(c_{\text{out}}^{(h)} | \gamma_{\text{out}})}{\text{Poisson}(c_{\text{out}}^{(h)} | \gamma_{\text{out}}) \text{Poisson}(c_{\text{in}}^{(h)} | \gamma_{\text{in}}) \text{Poisson}(c_{\text{out}}^* | \gamma_{\text{out}})} \quad (47)$$

$$= \frac{\text{Poisson}(c_{\text{in}}^* | \gamma_{\text{in}})}{\text{Poisson}(c_{\text{in}}^{(h)} | \gamma_{\text{in}})} \quad (48)$$

**With exposure**

$$e_{\text{out}}^{(h)} = \mathbf{E}^{(h)}[i, r, a, l, t] \quad (49)$$

$$e_{\text{in}}^{(h)} = \mathbf{E}^{(h)}[i, r', a, l, t] \quad (50)$$

$$e_{\text{out}}^* = e_{\text{out}}^{(h)} - 1(l = L)\frac{1}{2}\Delta \quad (51)$$

$$e_{\text{in}}^* = e_{\text{in}}^{(h)} + 1(l = L)\frac{1}{2}\Delta \quad (52)$$

$$\hat{e}_{\text{out}} = \hat{\mathbf{E}}[i, r, a, l, t] \quad (53)$$

$$\hat{e}_{\text{in}} = \hat{\mathbf{E}}[i, r', a, l, t] \quad (54)$$

$$\gamma_{\text{out}} = \gamma_k[i, r, \text{Out}, a, l, t] \quad (55)$$

$$\gamma_{\text{in}} = \gamma_k[i, r', \text{In}, a, l, t] \quad (56)$$

$$\frac{p(\mathbf{C}_k^* | \mathbf{N}^*) p(c^{(h)} | \hat{\mathbf{N}})}{p(\mathbf{C}_k^{(h)} | \mathbf{N}^*) p(c^* | \hat{\mathbf{N}})} = \frac{\text{Poisson}(c_{\text{out}}^* | \gamma_{\text{out}} e_{\text{out}}^*) \text{Poisson}(c_{\text{in}}^* | \gamma_{\text{in}} e_{\text{in}}^*) \text{Poisson}(c_{\text{out}}^{(h)} | \gamma_{\text{out}} \hat{e}_{\text{out}})}{\text{Poisson}(c_{\text{out}}^{(h)} | \gamma_{\text{out}} e_{\text{out}}^*) \text{Poisson}(c_{\text{in}}^{(h)} | \gamma_{\text{in}} e_{\text{in}}^*) \text{Poisson}(c_{\text{out}}^* | \gamma_{\text{out}} \hat{e}_{\text{out}})} \quad (57)$$

$$= (e_{\text{out}}^* / \hat{e}_{\text{out}})^{(c_{\text{out}}^* - c_{\text{out}}^{(h)})} \frac{\text{Poisson}(c_{\text{in}}^* | \gamma_{\text{in}} e_{\text{in}}^*)}{\text{Poisson}(c_{\text{in}}^{(h)} | \gamma_{\text{in}} e_{\text{in}}^*)} \quad (58)$$

### 7.2.2 Internal migration with net format

Never uses exposure, but does sometimes use weights.

$$w_{\text{out}} = \mathbf{W}[i, r, a, l, t] \quad (59)$$

$$w_{\text{in}} = \mathbf{W}[i, r', a, l, t] \quad (60)$$

$$\gamma_{\text{out}} = \gamma_k[i, r, a, l, t] \quad (61)$$

$$\gamma_{\text{in}} = \gamma_k[i, r', a, l, t] \quad (62)$$

$$\frac{p(\mathbf{C}_k^* | \mathbf{N}^*) p(c^{(h)} | \hat{\mathbf{N}})}{p(\mathbf{C}_k^{(h)} | \mathbf{N}^*) p(c^* | \hat{\mathbf{N}})} = \frac{\text{N}(c_{\text{out}}^* | \gamma_{\text{out}}, \frac{\phi^2}{w_{\text{out}}}) \text{N}(c_{\text{in}}^* | \gamma_{\text{in}}, \frac{\phi^2}{w_{\text{in}}}) \text{N}(c_{\text{out}}^{(h)} | \gamma_{\text{out}}, \frac{\phi^2}{w_{\text{out}}})}{\text{N}(c_{\text{out}}^{(h)} | \gamma_{\text{out}}, \frac{\phi^2}{w_{\text{out}}}) \text{N}(c_{\text{in}}^{(h)} | \gamma_{\text{in}}, \frac{\phi^2}{w_{\text{in}}}) \text{N}(c_{\text{out}}^* | \gamma_{\text{out}}, \frac{\phi^2}{w_{\text{out}}})} \quad (63)$$

$$= \frac{\text{N}(c_{\text{in}}^* | \gamma_{\text{in}}, \frac{\phi^2}{w_{\text{in}}})}{\text{N}(c_{\text{in}}^{(h)} | \gamma_{\text{in}}, \frac{\phi^2}{w_{\text{in}}})} \quad (64)$$

## 8 Calculating $\prod_{k=1}^K \frac{p(\mathbf{C}_k^{(h)}|\mathbf{N}^*)}{p(\mathbf{C}_k^{(h)}|\mathbf{N}^{(h)})}$

Does not depend on which component is being updated. Instead look at form of individual  $\frac{p(\mathbf{C}_k^{(h)}|\mathbf{N}^*)}{p(\mathbf{C}_k^{(h)}|\mathbf{N}^{(h)})}$

### 8.1 Generic $\mathbf{C}_k$

**No exposure** Drops out

**With exposure** Define cohort quantities

$$c^{(h)}(s) = \mathbf{C}_k^{(h)}[i, \min(a + s, A), t + s] \quad (65)$$

$$e^*(0) = e^{(h)}(s) + 1(l = L)S_k \frac{1}{2}\Delta \quad (66)$$

$$e^*(s) = e^{(h)}(s) + S_k \frac{1}{2}\Delta, \quad s > 0 \quad (67)$$

$$\lambda(s) = \gamma_N[i, \min(a + s, A), t + s] \quad (68)$$

Then

$$\frac{p(\mathbf{N}^*)}{p(\mathbf{N}^{(h)})} = \prod_{s=0}^{T-t} \frac{\text{Poisson}(n^*(s)|\lambda(s))}{\text{Poisson}(n^{(h)}(s)|\lambda(s))} \quad (69)$$

## 8.2 Algorithm for updating starting population

### 8.2.1 Choose cell to update

Choose a cell  $u$  in  $\mathbf{N}$  from the first time point. All valid cells have equal selection probabilities.

### 8.2.2 Identify subsequent population cells changed by update

Let  $\mathcal{N}$  be a set of indices for  $\mathbf{N}$  consisting of  $u$  plus indices for all cells for later time points whose values would need to change to satisfy the accounting identities if the value of cell  $u$  changed. For all  $i \in \mathcal{N}$ , let  $n_i^{(h)} = \mathbf{N}[i]$  and  $\lambda_i = \gamma^N[i]$ . Also let  $\lambda = \lambda_u$ .

### 8.2.3 Identify accession cells changed by update

If the account uses age, then let  $\mathcal{A}$  be the indices for cells in  $\mathbf{A}$  whose values would need to change to satisfy the accounting identities if the value of cell  $u$  in  $\mathbf{N}$  changed. For all  $i \in \mathcal{A}$ , let  $a_i^{(h)} = \mathbf{A}[i]$ .

### 8.2.4 Identify exposure cells changed by update

Let  $\mathcal{E}$  be the indices for cells in  $\mathbf{E}$  whose values would need to change to satisfy the accounting identities if the values of cell  $u$  in  $\mathbf{N}$  changed. For all  $i \in \mathcal{E}$ , let  $e_i^{(h)} = \mathbf{E}[i]$ .



### 8.2.5 Identify cells of components where expected values changed by update

If the system model for  $\mathbf{C}_k$  uses exposure, and if at least one cell in  $\mathcal{E}$  maps on to  $\mathbf{C}_k$ , then let  $\mathcal{C}_k$  be the indices for cells in  $\mathbf{C}$  whose expected values were to change if the value of cell  $u$  in  $\mathbf{N}$  changed; otherwise let  $\mathcal{C}_k = \{\}$ . For all  $i \in \mathcal{C}_k$ , let  $c_{ik} = \mathbf{C}_k[i]$  and  $\lambda_{ik} = \gamma_k^C[i]$ . Let  $e_{ik}^{(h)}$  denote the element of  $\mathbf{E}$  associated with cell  $i$  of  $\mathbf{C}_k$ . One element of  $\mathbf{E}$  may be associated with more than one cell of  $\mathbf{C}_k$ , such as when  $\mathbf{C}_k$  measures migrations using an origin-destination format.

### 8.2.6 Identify corresponding cells in datasets measuring population

If  $\mathbf{X}_m$  is used to estimate  $\mathbf{N}$ , and at least one cell in  $\mathcal{N}$  maps on to a cell in  $\mathbf{N}_m$ , then let  $\mathcal{X}_m$  be the cells that  $\mathcal{N}$  maps on to; otherwise let  $\mathcal{X}_m = \{\}$ . For all  $i \in \mathcal{X}_m$ , let  $x_{im}^{(h)} = \mathbf{X}_m[i]$ .

### 8.2.7 Find minimum value for population and accession

Let

$$v = \begin{cases} \min \left( \min_i n_i^{(h)}, \min_i a_i^{(h)} \right) & \text{if } \mathbf{N} \text{ has age} \\ \min_i n_i^{(h)} & \text{otherwise.} \end{cases} \quad (70)$$

### 8.2.8 Generate proposed values for cell

Draw proposed value  $n^*$  from the left-truncated Poisson distribution

$$p(y) \propto \begin{cases} e^{-\lambda} \lambda^y & \text{if } y \geq n^{(h)} - v \\ 0 & \text{otherwise.} \end{cases} \quad (71)$$

Let

$$\Delta = n^* - n^{(h)}. \quad (72)$$

### 8.2.9 Calculate ratio of likelihoods

For all  $i \in \mathcal{N}_m$ , let  $n_{km}^* = n_{km}^{(h)} + \Delta$  and let  $x_{km} = \mathbf{X}_m[i]$ . Then

$$r_{X|N} = \prod_{m \in \mathcal{M}_X} \prod_{i \in \mathcal{N}_m} \frac{p(x_{km} | c_{km}^*)}{p(x_{km} | c_{km}^{(h)})}. \quad (73)$$

### 8.2.10 Calculate ratio of prior probabilities

For all  $i \in \mathcal{N}$ , let  $n_i^* = n_i^{(h)} + \Delta$ . Then

$$r_{N|\lambda} = \prod_{i \in \mathcal{N}} \frac{p(n_i^* | \lambda_i)}{p(n_i^{(h)} | \lambda_i)}. \quad (74)$$

Let  $e_{ik}^*$  denote an updated version of  $e_{ik}^{(h)}$ . For all age group other than the final age group,  $e_{ik}^* = e_{ik}^{(h)} + \Delta/2$  (REFERENCE TO DATA STRUCTURES CHAPTER). Then

$$r_{C_k|\lambda, E} = \prod_{i \in \mathcal{C}_k} \frac{p(c_{ik}|\lambda_{ik}e_{ik}^*)}{p(c_{ik}|\lambda_{ik}e_{ik}^{(h)})}. \quad (75)$$

The term inside the multiplication simplifies, on a log scale, to

$$-\lambda_{ik}(e_{ik}^* - e_{ik}^{(h)}) + c_{ik}(\log e_{ik}^* - \log e_{ik}^{(h)}). \quad (76)$$

.

### 8.2.11 Calculate ratio of proposal probabilities

Let

$$r_J = \frac{p(n^{(h)}|\lambda)}{p(n^*|\lambda)}. \quad (77)$$

### 8.2.12 Calculate combined Metropolis-Hastings ratio, determine acceptance, and update

Let

$$r = r_{X|N} \times r_{N|\lambda} \times r_J. \quad (78)$$

If  $r > 1$  or if  $U \sim \text{Unif}(0, 1) < r$ , then accept the proposal, and update  $\mathbf{N}$ ,  $\mathbf{E}$ , and, if  $\mathbf{N}$  has age,  $\mathbf{A}$ . Otherwise leave all arrays at their current values.

The calculation of the Metropolis-Hasting ratio can be simplified slightly by taking advantage of the fact that  $r_J$  and the first element of  $r_{N|\lambda}$  cancel.

## 8.3 Algorithm for updating component $\mathbf{C}$

### 8.3.1 Choose cells in component to update

Choose a cell  $u$  in  $\mathbf{C}$  to update. If  $\mathbf{C}$  has a pool format, then  $u$  must come from the ‘ins’ part of the array, with every ‘ins’ cell having an equal selection probability; otherwise it can come from any cell in  $\mathbf{C}$ , with every cell having an equal selection probability. If  $\mathbf{C}$  is internal migration has a pool format, then select a second cell  $u_{\text{oth}}$  that is identical to  $u$  except that it comes from the ‘outs’ part of  $\mathbf{C}$ , and belongs to different categories for the ‘between’ dimensions. If  $\mathbf{C}$  is internal migration and has a net format, then choose a second cell  $u_{\text{oth}}$  that is identical to  $u$  except that belongs to different categories for the ‘between’ dimensions. In either case, all valid cells have equal selection probabilities. Let  $\mathcal{J}$  denote the set of cells in  $\mathbf{C}$  that are being updated. Let  $c^{(h)} = \mathbf{C}[u]$  and  $c_{\text{oth}}^{(h)} = \mathbf{C}[u_{\text{oth}}]$ . It is also convenient, for all  $j \in \mathcal{J}$ , to define  $c_j^{(h)} = \mathbf{C}[j]$ .

Let  $\mathcal{Q}$  denote the set of  $m$  such that (i)  $\mathbf{X}_m$  is used to estimate  $\mathbf{C}$ , and (ii) at least one cell in  $\mathcal{J}$  maps on to a cell in  $\mathbf{C}_m$ . If  $m \in \mathcal{Q}$ , then let  $\mathcal{J}_m$  be the cells that  $\mathcal{J}$  maps on to. If  $\mathbf{C}$  is internal migration, or if  $\mathbf{C}$  has parent-child dimensions (implying that it is births), then  $\mathcal{J}_m$  contains one or two cells; otherwise it contains one cell. For all  $j \in \mathcal{J}_m$ , let  $c_{jm}^{(h)} = \mathbf{C}_m[j]$ .

Table 1: Cells involved in update

Component	Format	Update two cells in $\mathbf{C}$	Update two states in $\mathbf{N}$
Internal migration	Origin-destination	No	Yes
Internal migration	Pool	Yes	Yes
Internal migration	Net	Yes	Yes
All other	–	No	No

### 8.3.2 Identify population cells involved in update

Let  $\mathcal{K}$  denote the set of cells in  $\mathbf{N}$  whose values would need to change to satisfy the accounting identities if the values of the cells in  $\mathcal{J}$  changed. If  $\mathbf{C}$  is internal migration, then  $\mathcal{K} = \{\mathcal{K}^{\text{orig}}, \mathcal{K}^{\text{dest}}\}$ , where  $\mathcal{K}^{\text{orig}}$  contains cells in the origin and  $\mathcal{K}^{\text{dest}}$  contains cells in the destination. Similarly, if  $\mathbf{C}$  has parent-child dimensions, then  $\mathcal{K} = \{\mathcal{K}^{\text{parent}}, \mathcal{K}^{\text{child}}\}$ . For all  $i \in \mathcal{K}$ , let  $n_i^{(h)} = \mathbf{N}[i]$ . If  $\mathbf{C}$  is internal migration, then, for all  $i \in \mathcal{K}^{\text{orig}}$ , let  $n_i^{\text{orig}(h)} = \mathbf{N}[i]$ . Define  $n_i^{\text{dest}(h)}$ ,  $n_i^{\text{parent}(h)}$ , and  $n_i^{\text{child}(h)}$  similarly.

Let  $\mathcal{M}_X$  denote the set of  $m$  such that (i)  $\mathbf{X}_m$  is used to estimate  $\mathbf{N}$ , and (ii) at least one cell in  $\mathcal{K}$  maps on to a cell in  $\mathbf{N}_m$ . If  $m \in \mathcal{M}_X$ , then let  $\mathcal{K}_m$  be the cells that  $\mathcal{K}$  maps on to. If  $\mathbf{C}$  is internal migration, then  $\mathcal{K}_m = \{\mathcal{K}_m^{\text{orig}}, \mathcal{K}_m^{\text{dest}}\}$ . Similarly, if  $\mathbf{C}$  has parent-child dimensions, then  $\mathcal{K}_m = \{\mathcal{K}_m^{\text{parent}}, \mathcal{K}_m^{\text{child}}\}$ . For all  $i \in \mathcal{K}_m$ , let  $n_{km}^{(h)} = \mathbf{C}_m[i]$ . If  $\mathbf{C}$  is internal migration, then for all  $i \in \mathcal{K}_m^{\text{orig}}$ , let  $n_{km}^{\text{orig}(h)} = \mathbf{N}[i]$ . Define  $n_{km}^{\text{dest}(h)}$ ,  $n_{km}^{\text{parent}(h)}$ , and  $n_{km}^{\text{child}(h)}$  similarly.

The scope of the updates of  $\mathbf{C}$  and  $\mathbf{N}$  is summarised in Table 1.

### 8.3.3 Identify accession cells involved in update

If  $\mathbf{C}$  has an age dimension, then let  $\mathcal{K}_A$  be the set of cells in  $\mathbf{A}$  whose values would need to change to satisfy the accounting identities if the values of the cells in  $\mathcal{J}$  changed. Sets  $\mathcal{K}_A^{\text{orig}}$ ,  $\mathcal{K}_A^{\text{dest}}$ ,  $\mathcal{K}_A^{\text{parent}}$ , and  $\mathcal{K}_A^{\text{child}}$  are defined analogously to their equivalents for population, as are values  $a_i^{\text{orig}(h)}$ ,  $a_i^{\text{dest}(h)}$ ,  $a_i^{\text{parent}(h)}$ , and  $a_i^{\text{child}(h)}$ .

### 8.3.4 Find minimum values for population and accession

If  $\mathbf{C}$  is not internal migration and does not have parent-child dimensions, then let

$$v = \begin{cases} \min \left( \min_i n_i^{(h)}, \min_i a_i^{(h)} \right) & \text{if } \mathbf{C} \text{ has age} \\ \min_i n_i^{(h)} & \text{otherwise.} \end{cases} \quad (79)$$

If  $\mathbf{C}$  is internal migration, then let

$$v^{\text{orig}} = \begin{cases} \min \left( \min_i n_i^{\text{orig}(h)}, \min_i a_i^{\text{orig}(h)} \right) & \text{if } \mathbf{C} \text{ has age} \\ \min_i n_i^{\text{orig}(h)} & \text{otherwise.} \end{cases} \quad (80)$$

$$v^{\text{dest}} = \begin{cases} \min \left( \min_i n_i^{\text{dest}(h)}, \min_i a_i^{\text{dest}(h)} \right) & \text{if } \mathbf{C} \text{ has age} \\ \min_i n_i^{\text{dest}(h)} & \text{otherwise.} \end{cases} \quad (81)$$

If  $\mathbf{C}$  has parent-child dimensions, then define  $v^{\text{parent}}$  and  $v^{\text{child}}$  similarly.

### 8.3.5 Obtain current exposures

If the system model for  $\mathbf{C}$  uses exposure, then, for  $j \in \mathcal{J}$ , let  $e_j^{(h)} = \mathbf{E}[j]$ .

### 8.3.6 Obtain current expected values for component

If the demographic model for  $\mathbf{C}$  uses exposure, then, for  $j \in \mathcal{J}$ , let  $\lambda_j^{C(h)} = \gamma^C[j]e_j^{(h)}$ . Otherwise, for  $j \in \mathcal{J}$ , let  $\lambda_j^{C(h)} = \gamma^C[j]$ .

### 8.3.7 Obtain expected counts for population

For  $i \in \mathcal{K}$ , let  $\lambda_i^N = \gamma^N[i]$ .

### 8.3.8 Generate proposed values for cells in component

Draw proposed value  $c^*$ . If  $\mathbf{C}$  is a type of increment and does not have parent-child dimensions, then the value is drawn from the left-truncated Poisson distribution

$$p(y) \propto \begin{cases} e^{-\lambda} \lambda^{(h)y} & \text{if } y \geq c^{(h)} - v \\ 0 & \text{otherwise.} \end{cases} \quad (82)$$

If  $\mathbf{C}$  a type of decrement, then  $c^*$  is drawn from the right-truncated Poisson distribution

$$p(y) \propto \begin{cases} e^{-\lambda} \lambda^{(h)y} & \text{if } y \leq v - c^{(h)} \\ 0 & \text{otherwise.} \end{cases} \quad (83)$$

If  $\mathbf{C}$  is internal migration, then  $c^*$  is drawn from the truncated Poisson distribution

$$p(y) \propto \begin{cases} e^{-\lambda} \lambda^{(h)y} & \text{if } c^{(h)} - v^{\text{dest}} \leq y \leq v^{\text{orig}} - c^{(h)} \\ 0 & \text{otherwise.} \end{cases} \quad (84)$$

If  $\mathbf{C}$  has parent-child dimensions, then  $c^*$  is drawn from the truncated Poisson distribution

$$p(y) \propto \begin{cases} e^{-\lambda} \lambda^{(h)y} & \text{if } c^{(h)} - v^{\text{parent}} \leq y \leq v^{\text{child}} - c^{(h)} \\ 0 & \text{otherwise.} \end{cases} \quad (85)$$

Let

$$\Delta_c = c^* - c^{(h)}. \quad (86)$$

If  $\mathbf{C}$  is internal migration, or if  $\mathbf{C}$  has parent-child dimensions, then let  $c_{\text{oth}}^* = c_{\text{oth}}^{(h)} + \Delta_c$ .

### 8.3.9 Calculate ratio of likelihoods for component $\mathbf{C}$

For all  $j \in \mathcal{J}_m$ , let  $c_{jm}^* = c_{jm}^{(h)} + \Delta_c$ , and let  $x_{jm} = X_m[j]$ . The ratio of the likelihoods for component  $\mathbf{C}$  can then be calculated as

$$r_{X|\mathbf{C}} = \prod_{m \in \mathcal{Q}} \prod_{j \in \mathcal{J}_m} \frac{p(x_{jm}|c_{jm}^*)}{p(x_{jm}|c_{jm}^{(h)})}. \quad (87)$$

### 8.3.10 Calculate ratio of likelihoods for population $\mathbf{N}$

If  $\mathbf{C}$  is an increment and does not have parent-child dimensions, then, for all  $i \in \mathcal{K}_m$ , let

$$n_{km}^* = n_{km}^{(h)} + \Delta_c. \quad (88)$$

If  $\mathbf{C}$  is a decrement, then, for all  $i \in \mathcal{K}_m$ , let

$$n_{km}^* = n_{km}^{(h)} - \Delta_c. \quad (89)$$

If  $\mathbf{C}$  is internal migration, then for all  $i \in \mathcal{K}_m^{\text{orig}}$ , let

$$n_{km}^{\text{orig}*} = n_{km}^{\text{orig}(h)} - \Delta_c, \quad (90)$$

and for all  $i \in \mathcal{K}_m^{\text{dest}}$ , let

$$n_{km}^{\text{dest}*} = n_{km}^{\text{dest}(h)} + \Delta_c. \quad (91)$$

If  $\mathbf{C}$  has parent-child dimensions, then for all  $i \in \mathcal{K}_m^{\text{parent}}$ , let

$$n_{km}^{\text{parent}*} = n_{km}^{\text{parent}(h)} - \Delta_c, \quad (92)$$

and for all  $i \in \mathcal{K}_m^{\text{child}}$ , let

$$n_{km}^{\text{child}*} = n_{km}^{\text{child}(h)} + \Delta_c. \quad (93)$$

In addition, for all  $i \in \mathcal{K}_m$  let  $x_{km} = X_m[i]$ . The ratio of likelihoods for population  $\mathbf{N}$  can then be calculated as

$$r_{X|\mathbf{N}} = \prod_{m \in \mathcal{M}_X} \prod_{i \in \mathcal{K}_m} \frac{p(x_{km}|c_{km}^*)}{p(x_{km}|c_{km}^{(h)})}. \quad (94)$$

**8.3.11 Calculate ratio of prior probabilities for component  $C$**

Let

$$r_{C|\lambda} = \prod_{j \in \mathcal{J}} \frac{p(c_j^* | \lambda_j^{C*})}{p(c_j^{(h)} | \lambda_j^{C(h)})}. \quad (95)$$

**8.3.12 Calculate ratio of prior probabilities for population  $N$**

Let

$$r_{N|\lambda} = \prod_{i \in \mathcal{K}} \frac{p(n_i^* | \lambda_i^N)}{p(n_i^{(h)} | \lambda_i^N)}. \quad (96)$$

**8.3.13 Calculate ratio of proposal probabilities**

Let

$$r_J = \frac{p(c_u^{(h)} | \lambda_u^{C*})}{p(c_u^* | \lambda_u^{C(h)})}. \quad (97)$$

**8.3.14 Calculate combined ratio, determine acceptance, and update**

Let

$$r = r_{X|C} \times r_{X|N} \times r_{C|\lambda} \times r_{N|\lambda} \times r_J. \quad (98)$$

If  $r > 1$  or if  $U \sim \text{Unif}(0, 1) < r$ , then accept the proposal, and update  $C$  and  $N$ , and possibly  $A$  and  $E$ . Otherwise leave all elements at their current values.