

ANNOUNCEMENTS

- EVAL on PLEACH
- LATE HW & SOLUTIONS → DON'T WAIT TO VERY LAST MINUTE

THE CENTRAL LIMIT THEOREM

in the limit of a large # of MEASUREMENTS,
the distribution of a measured parameter
is GAUSSIAN

$$p(x) dx = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$

\uparrow \uparrow
 true val.

Variance = (standard dev.)²
 \uparrow
 measure of significance

SKETCH PROOF recall from LEC 17 (DISPERSION REL)
 the convolution theorem

$$\text{if } f(X) \equiv \int dx_1 f_1(x_1) f_2(X-x_1) = \int dx_1 dx_2 f_1(x_1) f_2(x_2) \delta(X-x_1-x_2)$$

\uparrow
 ARGUMENTS SUM TO X

- then the FOURIER TRANSFORMS ARE RELATED BY

$$\tilde{f}(k) = \tilde{f}_1(k) \tilde{f}_2(k)$$

$\uparrow \quad \uparrow \quad \uparrow$
 SAME MOMENTUM VAR

⚠ not for class: TRICK IS TO INSERT $1 = \int \delta(\dots) dy$ cleverly
 $f(X) = \int dx_1 \int dk e^{-ikx_1} \tilde{f}_1(k) \int dk' e^{-ik'(X-x_1)} \tilde{f}_2(k')$

\uparrow
 THIS INTEGRAL

$$= \int dk \int dk' e^{-ikX} \tilde{f}_1(k) \tilde{f}_2(k') \left[\int dx_1 e^{-i(k-k')x_1} \right]$$

$$= \int dk e^{-ikX} \tilde{f}_1(k) \tilde{f}_2(k)$$

= 4

this lec:
from
PESKIN
SLAC
STATS
WEEK
2011

THIS RESULT GENERALIZES ("it is obvious")

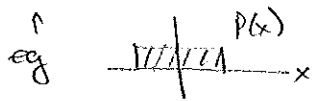
$$f(X) = \int dx_1 \dots dx_N f_1(x_1) f_2(x_2) \dots f_N(x_N) \delta(X - \sum_{i=1}^N x_i)$$

$$\Rightarrow \hat{f}(k) = \prod_{i=1}^N \hat{f}_i(k)$$

SOME k

DIFFERENT x 'S
SUBJECT TO: SUM
IS FIXED

SO LET'S PROVE IT. ASSUME SOME NICE pdf $p(x) dx$,
NOT NEC. GAUSSIAN.



finite $\langle x \rangle$

CHOOSE VARIABLES SUCH THAT $\langle x \rangle = 0$.

ie if $\langle x \rangle \neq 0$, then use $y = x - \langle x \rangle$; SHAPE OF pdf UNCHANGED.

NOW DRAW N numbers from this distribution.

$$x_1, \dots, x_N \quad \text{WITH SUM } X = \sum_{i=1}^N x_i$$

WANT: SHOW THAT $p(X)$ IS GAUSSIAN

from which you may deduce that
 $p(x_i)$ IS ALSO GAUSSIAN.

$$P(X) = \int dx_1 \dots dx_N \left[\prod_{i=1}^N p(x_i) \right] \delta(X - \sum_{i=1}^N x_i)$$

↑ to dist.
from p

↑ PRODUCT OF INDIVIDUAL
PROBABILITIES OF
EACH MEASUREMENT

ALLOWING ANY VALUES OF
THE $\{x_i\}$ SUBJECT TO
SUMMING TO X .

this is just an N -dim. convolution
USE CONVOLUTION THM

$$p(x_i) = \int dk e^{-ikx_i} \tilde{p}(k)$$

$$\tilde{p}(k) = \int dx e^{ikx} p(x)$$

REMEMBER MOMENTS OF A PDF? (EXPECTATION VALS)

$$\langle x^n \rangle = \int dx x^n p(x) \quad \text{of.} \quad \tilde{p}(k) = \int dx e^{ikx} p(x)$$

OBSERVE: $\tilde{p}(0) = \langle 1 \rangle = 1$

that was boring. but we can write any exp. val:

$$\langle x \rangle = \left(-i \frac{d}{dk} \right) \tilde{p}(k) \Big|_{k=0} \quad \leftarrow \langle x \rangle = 0 \text{ by assump.}$$

$$\langle x^n \rangle = \left(-i \frac{d}{dk} \right)^n \tilde{p}(k) \Big|_{k=0}$$

↑

but these are just n^{th} derivatives
of a function $\tilde{p}(k)$
(TAYLOR EXPANSION?!)

without loss of generality (given $\tilde{p}(0) = 1$)

instead of $\tilde{p}(k) = A + Bk + Ck^2 + \dots$

write this as

$$\tilde{p}(k) = e^{a + ib_1 k + \frac{1}{2} C_1 k^2 + \dots}$$

then:

$$\tilde{p}(0) = e^a \rightarrow \boxed{a = 1}$$

$$-i \frac{d}{dk} \tilde{p}(k) \Big|_0 = \boxed{b_1 = \langle x \rangle = 0}$$

$$-\frac{d^2}{dk^2} \tilde{p}(k) \Big|_0 = \boxed{C_1 = \langle x^2 \rangle} \quad \text{etc.}$$

then the Fourier transform of $p(x)$ is

$$\tilde{P}(k) = \prod_i P(k) = e^{-\frac{1}{2} G k^2 - i D k^3 + \dots}$$

↑
BY CONVOLUTION
THEOREM

$$a = b = 0$$

$$G = Nc$$

"LARGE N" UNIT : HOW DO THINGS SCALE W/ N?

C, D, E, \dots are all sum over N terms

\Rightarrow all $\mathcal{O}(N)$

$$e^{-\frac{1}{2}CK^2 - iDK^3 + \dots}$$

exp. suppress. $\mathcal{O}(N)$

so this gets small when k is at least $\mathcal{O}(1/\sqrt{N})$

\rightarrow at this point,

CK^2 is $\mathcal{O}(1)$

these terms get super tiny

$\left\{ \begin{array}{l} DK^3 \text{ is } \mathcal{O}(1/\sqrt{N}) \\ EK^4 \text{ is } \mathcal{O}(1/N) \text{ etc.} \end{array} \right.$

* this is where we use $\langle x^2 \rangle$ is finite
 ** also nb that every other term is oscillatory, not exponential

As $N \rightarrow \infty$,

$$\tilde{P}(k) = e^{-\frac{1}{2}CK^2}$$

← GAUSSIAN!

FACT: FOURIER TRANSFORM OF A GAUSSIAN IS ALSO GAUSSIAN

$$\tilde{P}(x) = \frac{1}{\sqrt{2\pi C}} e^{-x^2/2C}$$

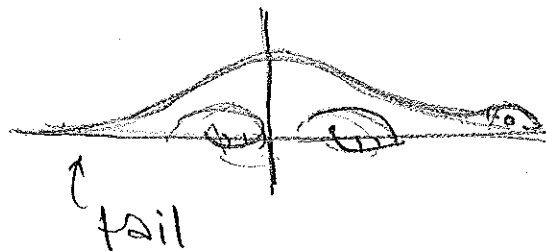
✓

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SO WHAT:

- CUTE PROOF, REMINDS US OF FOURIER-SLODY
- REMINDER OF WHERE CLT FAILS

Convergence to a Gaussian is slowest @ the tails of the distribution



STATISTICAL INFERENCE IS CRUCIALLY TIED TO BEHAVIOR @ TAILS.

THIS IS WHY WE SOMETIMES FALL INTO THE TRAP OF MISQUOTING "SIGMAS"

WE DO EXPERIMENTS, THEN DO A STATISTICAL TEST

↳ wave hands @ central limit theorem like it is some holy mantration, then quote some value for σ

BUT σ ONLY MAKES SENSE IF THE DISTRIBUTION IS ACTUALLY DESCRIBED BY A GAUSSIAN

↑
for that value of x
↑ the finite value of N
↑ the shape of $p(x)$

CHEAT: SIMULATE EXPERIMENTS ON A COMPUTER (pseudo experiments)

↳ do statistics i make sure it aligns w/ what you expect.

(computer time is cheap, data isn't)

HYPOTHESIS TEST

GIVEN DATA x \leftrightarrow eg some vector of measurements
 THEORY (hypothesis) H \leftrightarrow eg some set of params.

HOW COMPATIBLE IS A w/ x ? QUANTIFY

$$L(H) \equiv P(x|H)$$

\uparrow LIKELIHOOD

prob. of obtaining data
ASSUMING H is true.

this is easy to calculate.

small $L(H) \rightarrow H$
 is unlikely... but
how unlikely?

HARDER: how to use $L(H)$ to say something
 quantifiable about H .

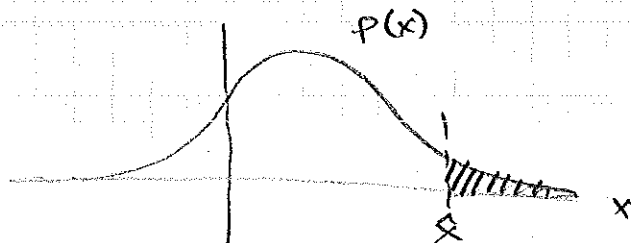
WRITE: \hat{x} AS EXPERIMENTAL VALUE (DATA)

in contrast to x , which is the
 quantity being measured

DEFINE A "TAIL MEASURE"

$$\alpha = \int_{\hat{x}}^{\infty} dx P(x|H)$$

UPPER CUMULATIVE DIST.

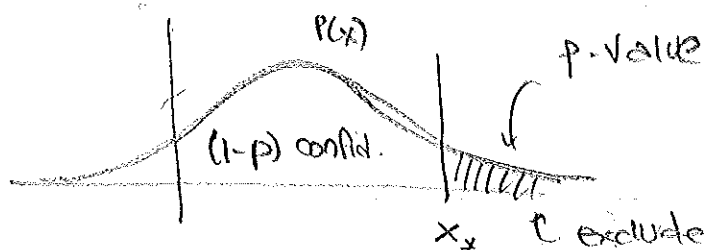


(assume \hat{x} is on upper tail)

if H IS TRUE: α MEASURES PROBABILITY THAT THE
 OUTCOME OF AN EXPERIMENT IS AT LEAST AS
 CRAZY AS \hat{x} .

if $\alpha = 1\%$, then: H excluded @ 1% confidence.

CAN DEFINE CRITERION FOR EXCLUSION:



NB "p" in "p-value"
 is not a prob.

x_c \uparrow exclude H if $\hat{x} > x_c$

WHEN $p(x)$ IS GAUSSIAN

↑ i.e. "sufficiently close to bump" (vs tail)
 † N is large enough that central limit thm. holds

then: can relate p-values to standard deviations

eg.

| | |
|-------|----|
| 16% | 1σ |
| 2.3% | 2σ |
| 0.14% | 3σ |

nb: if you underestimate $\sigma \rightarrow$ too small p-value!

philosophical question:

what is the probability of H given \hat{x} ?

FREQUENTIST: impossible to answer.
 can give p-value, but this is not a measure of truth of H .

BAYESIAN:

$$P(H|\hat{x}) = \frac{P(\hat{x}|H) \cdot P(H)}{P(\hat{x})} \quad \left\{ \begin{array}{l} \text{what is this?!} \\ \leftarrow = 1 \text{ since we definitely measured } \hat{x} \end{array} \right.$$

$P(H)$: PRIOR PROBABILITY

"probability prior to measurement that H is correct"

↑ SUBJECTIVE AS HELL. ← incomplete knowledge

$P(H|\hat{x})$: POSTERIOR PROBABILITY

↑ euphemism for butt.

SUPPOSE YOU HAVE A FAMILY OF HYPOTHESES

$H(a)$

↑ theory depends on a parameter
eg lifetime of unstable particle

prior: no idea what a is.

$P(H(a))$ or $p(a)$ is constant.

posterior: $p(a) \sim p(\hat{x} | a)$, makes sense.

NORMALIZING:

$$P(a | \hat{x}) = \frac{P(\hat{x} | a)}{\int da' P(\hat{x} | a')}$$

← flat dist. is
a const. that
cancels.

A PROBLEM: WE ASSUMED a HAS FLAT PRIOR.
) THAT THIS IS CONSERVATIVE.

BUT $b = a^2$ IS A PERFECTLY FINE
ALTERNATIVE PARAMETER TO a
THAT CONTAINS THE SAME INFO.

BUT IF a HAS A FLAT PRIOR,
 b DOES NOT!!

↑ even choice of param is
subjective.

↑
RELATED TO THE MEASURE PROBLEM
IN MULTIVERSE THEORIES

PARTICLE THEORIST: want to answer

"What is prob. that $M_H \in [124, 126] \text{ GeV}$?"

\uparrow 1 PARAM

\rightarrow WANT TO BE BAYESIAN

ASTROPHYSICS: so many uncertainties
... impossible to make
useful conclusion w/o
priors.

\rightarrow BAYESIAN (typically)

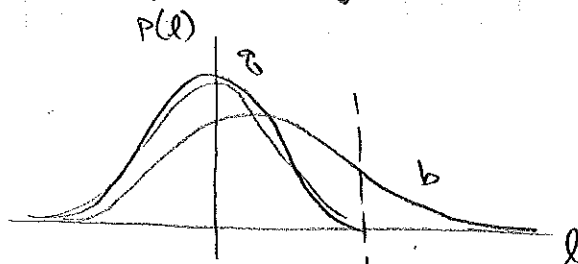
PARTICLE EXPERIMENT:

given ~ 1000 collaborators,
will never agree on a prior.

\rightarrow FREQUENTIST $\hat{\theta}$

eg. @ LHC, you produce jets from $\left\{ \begin{array}{l} \text{LIGHT QUARK, } q \\ \text{BOTTOM QUARK, } b \end{array} \right.$ or give

distinguish by: position of Z-vertex, l



$l_x \leftarrow$ if $l > l_x$, probably b-jet!

EFFICIENCY: given a b-jet, how likely are we to tag it?

$$\epsilon_b = \int_{l_x}^{\infty} dl \, p(l|b)$$

PURITY: given a tagged b-jet, how likely is it actually a q?

$$\epsilon_q = \int_{l_x}^{\infty} dl \, p(l|q)$$

EMERGENCY: choose l_v according to balance of high E_b , low E_e .

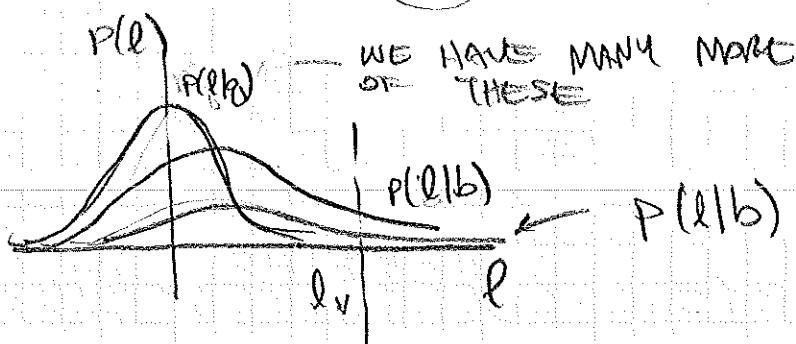
Bayesian: actually, it is much more likely that a given proton collision produces a q jet rather than b .

posterior: $p(q) \gg p(b)$

so RELATIVE PROBABILITIES ARE

$$\frac{p(l|b)}{p(l|q)} \frac{p(b)}{p(q)}$$

because $p(b)/p(q)$ SMALL,
WANT TO PICK l_v
LARGE



$$p(l|b) \frac{p(b)}{p(q)}$$

RATIO OF PRIORS

so THIS LINE SHOULD BE FURTHER TO RIGHT
→ tails !!

BAYESIAN INFERENCE EXAMPLE
from How Not to be Wrong, Ellenberg p.169

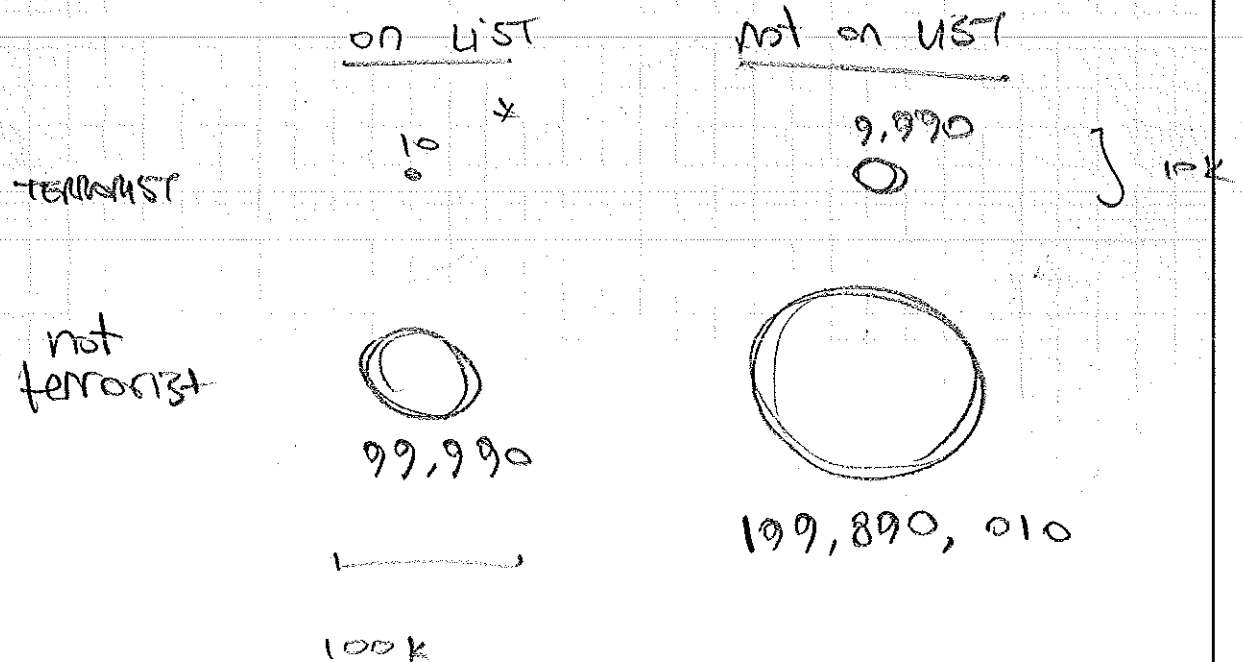
Minority Report + Facebook

IMAGINE FB. CAN DEV. AN ALGORITHM TO SAY:

here's a list of users (subset) $\leftarrow 100k$
(or 200M)
people on this list are 2x as likely
as avg user to be a terrorist

SUPPOSE one of your friends is on
this list. 2 how do you feel?

LET'S BE PARANOID } MAKE UP A #: 10K
"persons of interest"



↓ : on list: 2x AS LIKELY TO BE TERRORIST

AVG USER: $\frac{10k}{200M} = \frac{1}{20k}$ so on list: $\frac{1}{10k}$

even if on list, 99.99% chance innocent.

Better q: If NOT terrorist, what
is chance of being on list?

$$\frac{99,900}{199,899,010}$$

}

$$\boxed{0.05\%}$$

1

$$\boxed{1/2k}$$

chance of
wrong ID!

TWO DIFF Q

1. CHANCE PERSON PUT ON LIST (1/2000)
GIVEN NOT TERRORIST

2. CHANCE NOT TERRORIST (1/10k)
GIVEN ON LIST