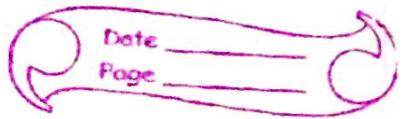


Tute - 3 Soln



Sol 1 There are n types of coupons.

P_i = Probability that randomly chosen coupon is of type i .

→ Here $\forall i, P_i$ is constant. Means you have to assume that there are infinitely many coupons from which probability for coupon with type i to be chosen is P_i and that is constant, even after many trials. ⇒ Or "Assume the case of "With Replacement"."

→ Given that
$$\sum_{i=1}^n P_i = 1$$

"Independent Events".
Or

k coupons are tube collected.

A_i : At least one coupon of type " i " is collected among those k coupons.

& $i \neq j$.

(A) $P(A_i) = ?$

A_i^c = None of the coupons are of type i .

$$\therefore P(A_i^c) = (1 - P_i)^k$$

$$\therefore P(A_i) = 1 - P(A_i^c)$$

$$= 1 - (1 - P_i)^k$$

* Long method to understand.

A_i can be seen as there are 1 or 2 or 3 ... k types of coupons.

$$\therefore P(A_i) = \binom{k}{1} P_i \cdot (1 - P_i)^{k-1} + \binom{k}{2} P_i^2 (1 - P_i)^{k-2} + \dots + \binom{k}{k} P_i^k$$

↓ ↓ ↓ ↓
1 'i' type 2 'i' type ... k 'i' type

$$\begin{aligned} P(A_i) &= \binom{k}{1} p_i (1-p_i)^{k-1} + \binom{k}{2} p_i^2 (1-p_i)^{k-2} + \dots + \binom{k}{k} p_i^k (1-p_i)^0 \\ &\quad \text{for 1 only} \qquad \qquad \text{for 2 only} \quad \dots \quad \text{for all} \\ &= \binom{k}{0} (1-p_i)^k + \binom{k}{1} p_i (1-p_i)^{k-1} + \dots + \binom{k}{k} p_i^k (1-p_i)^0 \\ &= (p_i + 1 - p_i)^k - 1 \cdot (1-p_i)^k \quad \boxed{\frac{1 - (1-p_i)^k}{1}} \quad \rightarrow \binom{k}{0} (1-p_i)^k \\ &\quad \text{Using binomial theorem} \end{aligned}$$

(B) $P(A_i \cup A_j) = ?$

here., $(A_i \cup A_j)^c = A_i^c \cap A_j^c$

\downarrow

No coupon is of either type 'i' or type 'j'!

$\therefore P(A_i \cup A_j)^c = (1 - p_i - p_j)^k$

$\boxed{P(A_i \cup A_j) = 1 - (1 - p_i - p_j)^k}$

(C) $P(A_i | A_j) = ?$

Using formula.

$$\begin{aligned}
 P(A_i | A_j) &= \frac{P(A_i \cap A_j)}{P(A_j)} \\
 &= \frac{-P(A_i \cup A_j) + P(A_i) + P(A_j)}{P(A_j)} \\
 &= \frac{-(1 - (1 - P_i - P_j)^k) + 1 - (1 - P_i)^k + 1 - (1 - P_j)^k}{1 - (1 - P_j)^k} \\
 &= \frac{1 - (1 - P_i)^k - (1 - P_j)^k + (1 - P_i - P_j)^k}{1 - (1 - P_j)^k}
 \end{aligned}$$

SIM 2]

Independent trials.

$$P(\text{Success}) = p$$

$$P(\text{Failure}) = 1-p.$$

What is the probability that n successes occur before m failures?

→ Here it is not specified that how many trials will be taken.

⇒ the question can be seen as "What is the probability that at least n successes occur before m failures."

⇒ There must be at least $n+m-1$ successes in first $n+m-1$ trials. Here game can be over before $n+m-1$ trials! And in that case we will assume that additional trials are taken to complete $n+m-1$ trials.

Now the probability for k successes in l trials

$$P_{S,l} = \binom{l}{k} p^k (1-p)^{l-k}$$

p : Probability of success

$$\text{here } k = n, n+1, \dots, n+m-1$$

$$l = n+m-1$$

∴ Probability of n successes before m failures

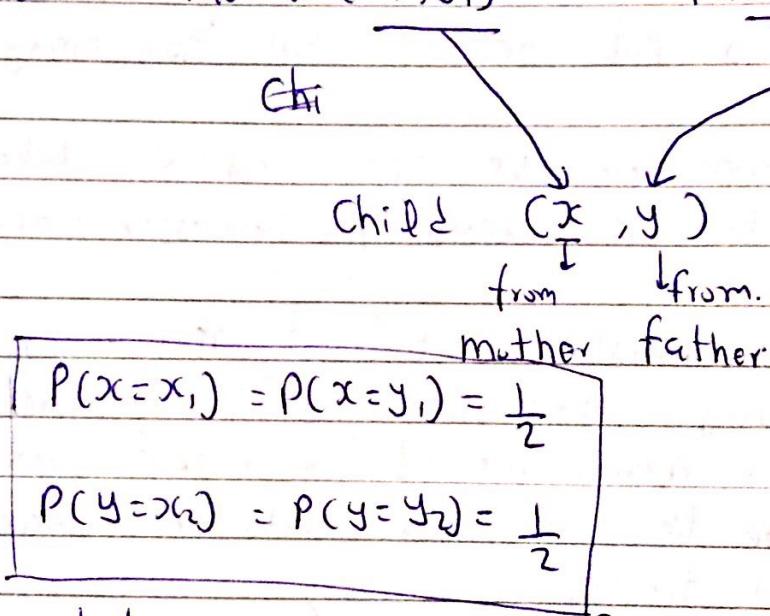
$$\text{is } P_{n,m} = \sum_{i=n}^{n+m-1} (n+m-1) p^i (1-p)^{n+m-1-i}$$

→ The game discussed in the question is just to give you a brief idea about the question.

Solution 3]

Let's learn some biology !!!

Every living things have a gene pair in the DNA and that can be seen as (x, y) . Now x & y are determined by its mother and father. \rightarrow Mother (x_1, y_1) Father (x_2, y_2)



- Now Let's move on to the question.
A female chimp has given birth to a child.
 M_1 : Male no. 1 is the father.

$$\boxed{P(M_1) = P}$$

M_2 : Male no. 2 is the father.

$$\boxed{P(M_2) = (1-P)}$$

pair of gene.

Female (A, A)

Child/Baby. (A, a)

Male 1 (a, a)

Male 2 (A, a)

$BGA \in B(a, a) \rightarrow$ Baby has gene pair (A, a)

→ What is the probability that Male 1 is a father?

$$\therefore P(M_1 | B_{(A,4)}) = ?$$

$$\begin{aligned} \therefore P(M_1 | B_{(A,4)}) &= \frac{P(B_{(A,4)} | M_1) P(M_1)}{P(B_{(A,4)} | M_1) P(M_1) + P(B_{(A,4)} | M_2) P(M_2)} \\ &= \frac{1 \cdot P}{1 \cdot P + (1-P) \frac{1}{2}} \quad (\because M_1 \text{ has } (A,4)) \end{aligned}$$

$$\boxed{= \frac{2P}{P+1}}$$

Sol'n 4] n men in a party. Their hats are removed and mixed up. Now a match occurs if a man selects his own hat.

(a) What is the probability that no matches occur.

Method 1 $\rightarrow P(C) = \frac{\text{Nu. of ways in which no matches occur.}}{\text{total ways of arrangement.}}$

here total ways of arrangement = $n!$

Now to find total ways in which no matches occur, we will use concept of Derangement and Inclusion - Exclusion.

~~S₀: #ways in which 0 derangements occurs.
S₁: #ways in which 1 derangement occurs.
S₂: #ways in which 2 derangements occurs.
S₃: #ways in which 3 derangements occurs.
S_n: #ways in which n derangements occurs.~~

$d_n = \# \text{ways in which no one gets their hats.}$

Now using inclusion exclusion argument

$$d_n = S_0 - S_1 + S_2 - S_3 + \dots + (-1)^n S_n \quad (1)$$

$S_i = \binom{n}{i} (n-i)!$ → i things go in correct place and rest of the things go "anywhere"!

to explain in better way

M_i : Men i gets his correct hat.

$$\therefore d_n = N(M_0^c \cap M_1^c \cap M_2^c \cap \dots \cap M_n^c)$$

- Now using eqn 2 in eqn 1.

$$\therefore d_n = \binom{n}{0} n! - \binom{n}{1} (n-1)! + \binom{n}{2} (n-2)! + \dots + (-1)^n \binom{n}{n} 0!$$

$$= n! - \frac{n! (n-1)!}{(n-1)! 1!} + \frac{n! (n-2)!}{(n-2)! 2!} + \dots + \frac{(-1)^n n! 0!}{0! n!}$$

$$\boxed{d_n = n! \left(1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!} + \dots + \frac{(-1)^n}{n!} \right)}$$

now expected probability

$$p(E) = \frac{d_n}{n!}$$

$$\boxed{= \left(1 - \frac{1}{1!} + \frac{1}{2!} - \frac{1}{3!} + \dots + \frac{(-1)^n}{n!} \right)}$$

→ If you don't get this, plz try to understand the inclusion & exclusion argument.

* Method - 2 / Approach 2.

E: Event that no matches occur.

M: First person got his correct hat.

$P(E) = P_n \rightarrow$ to understand or to make it explicit the dependence on n .

$$\text{Now, } P_n = P(E) = P(E|M) P(M) + P(E|M^c) P(M^c)$$

$$\text{here } P(E|M) = 0$$

$$\therefore P_n = P(E|M^c) P(M^c)$$

$$\therefore P_n = P(E|M^c) \frac{(n-1)}{n}$$

→ ①

→ Now, $P(E|M^c) \rightarrow$ The probability of no matches when $(n-1)$ men select from a set of $(n-1)$ hats, that does not contain the hat of one of these men. This can happen in two ways

1: There are no matches and the extra man does not select that extra hat.

2: There are no matches and the extra man selects the extra hat.

→ Here case 1 can be seen as $n-1$ men and none chooses their hats. (for our convenience that extra hat is now belonging to that extra man)
so the probability of first case is P_{n-1}

→ Now the probability of second case is

$$\frac{1}{(n-1)} \quad P_{n-2}$$

→ no matches for rest of the $n-2$ men.
extra man chooses extra hat.

$$\therefore P_n = \frac{(n-1)}{n} \left(P_{n-1} + \frac{1}{n-1} P_{n-2} \right)$$

$$\therefore P_n = \left(\frac{n-1}{n} \right) P_{n-1} + \frac{1}{n} P_{n-2}$$

$$\therefore P_n - P_{n-1} = \frac{1}{n} (P_{n-1} - P_{n-2})$$

→ ②

now $P_1 = 0$ (1 man 1 hat)

$$P_2 = \frac{1}{2}$$

Putting this in eqn ②

$$P_3 - P_2 = -\frac{(P_2 - P_1)}{3} = -\frac{1}{3!} \Rightarrow P_3 = \frac{1}{2!} - \frac{1}{3!}$$

$$P_4 - P_3 = -\frac{(P_3 - P_2)}{4} = \frac{1}{4!} \Rightarrow P_4 = \frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!}$$

$$\therefore P_n = \frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} + \dots + \frac{(-1)^n}{n!}$$

Which is same as answer we got earlier using method 1.

- (B) What is the probability that ~~no matches~~ exactly k matches occur?
 → there are $\binom{n}{k}$ ways to choose k people

who will get their correct hat. And for them, fixed group of k men will get correct hat.

$$\therefore P(B) = \binom{n}{k} \times \frac{1}{n} \times \frac{1}{(n-1)} \times \dots \times \frac{1}{(n-(k-1))} \times P_{n-k}$$

$\binom{n}{k}$ ways probability that k rest of them to choose. men will get correct hat. will get wrong hat.

$$= \frac{n!}{(n-k)! k!} \times \frac{1}{n \times (n-1) \times (n-2) \times \dots \times (n-k+1)} \times P_{n-k}$$

$$= \frac{n!}{n! k!} \times P_{n-k}$$

$$\therefore P(B) = \frac{P_{n-k}}{k!}$$

$$= \frac{\frac{1}{2!} - \frac{1}{3!} + \frac{1}{4!} + \dots + \frac{(-1)^{(n-k)}}{(n-k)!}}{k!}$$