Week 1: Potential Outcomes

PLSC 30600 - Causal Inference

### Welcome!

#### Course Overview

- Instructor: Anton Strezhnev
- TA: Cindy Wang
- Logistics:
  - Lectures T/Th 2:00pm 3:20pm; Sections F 2:30pm 3:20pm
  - 3 Problem Sets (~ 2 weeks)
  - Take-home midterm (~ 1 week)
  - In-class final (exact day/time TBA)
  - My office hours: Tuesdays 4pm-6pm (Pick 328)
  - Cindy's office hours: TBA
- What is this course about?
  - Defining causal effects w/ a structured statistical framework (potential outcomes)
  - Outlining assumptions necessary to identify causal effects from data
  - Estimating and conducting inference on causal effects
- Goals for the course
  - Give you the tools you need to develop your own causal research designs and comment on others' designs.
  - Equip you with an understanding of the fundamentals of causal inference to enable you to learn new methods.
  - Professionalization -- how did causal inference emerge as a field and how do different disciplines approach the topic.

#### Course workflow

#### Lectures

- Topics organized by week
- Tuesday lecture: Introduction + higher-level overview
- Thursday lecture: More in-depth details + applications
- You should do the readings prior to each week, but definitely be sure to do them before the Thursday.

#### Readings

- Mix of textbooks and papers theory + application.
- Organized on the syllabus by priority. Try to do all of them, but if you need to prioritize, start with the first ones on the list for the week.
- All readings available digitally on Canvas

#### Course workflow

- Problem sets (25% of your grade)
  - Main day-to-day component of the class -- meant to get you working with your colleagues and thinking hard about the material.
  - Collaboration is strongly encouraged -- you should ask and answer questions on our class Ed
     Discussion board
  - Grading is on a plus/check/minus scale.
  - Conversion to grade point is somewhat hollistic.
  - o Majority plusses is an A, Majority checks is an A-, Majority minus is a B
  - Solutions will be posted after the due-date.

#### Course workflow

- Midterm Exam (30% of your grade)
  - The midterm will be structured like the problem sets with two main differences:
    - 1. You have about 1 week to complete them instead of 2
    - 2. You may **not** collaborate with one another on the exams.
- Final Exam (35% of your grade)
  - The final exam will be in-person and written.
  - Combination of theoretical questions and code/results analysis.
- Participation (10% of your grade)
  - It is important that you actively engage with lecture and section -- ask and answer questions.
  - Do the reading!
  - Participating on Ed counts towards this as well.

# Class Requirements

- Overall: An interest in learning and willingness to ask questions.
- Generally assume some background in probability theory and statistics (e.g. an intro course taught in most departments)
  - Regression a plus but not required -- if you know what  $(X'X)^{-1}X'Y$  is, then that is great, but not strictly necessary.
- Main concepts to be familiar with:
  - properties of random variables
  - estimands and estimators
  - bias, variance, consistency
  - o central limit theorem
  - o confidence intervals; hypothesis testing

#### A brief overview

- Week 1: Introduction to potential outcomes
  - Defining causal estimands
- Week 2-3: Experiments
  - Why they work, design, inference
- Week 4-5: Selection-on-observables
  - Key assumptions, graphical models, how to think about confounding
  - Estimation via regression, weighting, matching + combinations
- Week 6-8: Designs for dealing with unobserved confounding
  - Instrumental variables
  - Differences-in-differences
  - Regression discontinuity
- Week 9: Causal mediation; Sensitivity analyses

### Random variables and estimators

#### Random variables

- Understanding the behavior and properties of random variables is at the core of statistical theory.
- (Simply put) a random variable X is a mapping from a **sample space** to the real number line Random variables have a *distribution* (which we may or may not assume we know) defined by the cumulative distribution function (CDF)

$$F(x) = Pr(X \le x)$$

#### Random variables

• Discrete random variables take on a countable number of values (e.g. Bernoulli r.v. can take on 0 or 1) and have a probability mass function (PMF)

$$p(x) = Pr(X = x)$$

• Continuous random variables take on an uncountable number of values (e.g. the Normal distribution on  $(-\infty, \infty)$ ). No PMF, but have a *density* function (PDF) that integrates to a probability

$$Pr(X \in \mathcal{A}) = \int_{\mathcal{A}} f(x) dx$$

Remember: PMFs (and PDFs) sum (integrate) to 1 over the support of the random variable.

### Expectations

- One important property of a random variable is its expectation E[X]. We'll often make assumptions about the expectation of an R.V. while remaining agnostic about its true distribution.
- The expectation is a *weighted average*. For a discrete r.v. X, we sum over the support of the random variable  $\mathcal{X}$ .

$$E[X] = \sum_{x \in \mathcal{X}} x Pr(X = x)$$

For continuous r.v. we have an integral

$$E[X] = \int_{x\mathcal{X}} x f(x) dx$$

• Fun fact: we can get the expectation of any function of g(X) just by plugging it into the integral

$$E[g(X)] = \int_{x\mathcal{X}} g(x) f(x) dx$$

### Expectations

- You'll probably not have to do any integration in this class. Rather, we'll derive expectations of functions of random variables via known properties
- Most important. Linearity. For any two random variables X and Y and constants a and b

$$E[aX + bY] = aE[X] + bE[Y]$$

- Note that for any generic function g(),  $E[g(X)] \neq g(E[X])$ . If g() is convex, by Jensen's inequality  $E[g(X)] \geq g(E[X])$
- ullet For a binary r.v.  $X \in \{0,1\}$ , it's helpful to remember the "fundamental bridge" between expectations and probability

$$E[X] = Pr(X = 1)$$

#### Variance

• We also care about the *spread* of a random variable -- how far is the average draw of X from its mean E[X]. One measure of this is the variance.

$$Var(X) = E[(X - E[X])^2]$$

Also written as

$$Var(X) = E[X^2] - E[X]^2$$

- Note that the square is a convex function. Which means that by Jensen's inequality  $E[X^2] \ge E[X]^2$ . Variances *cannot* be negative!
- We also can define a *covariance* between two variables (does *X* take high values when *Y* takes high values?)

$$Cov(X,Y) = E[(X - E[X])(Y - E[Y])] = E[XY] - E[X]E[Y]$$

#### Variance

- Again, you won't have to do the integral. Variances have some useful properties.
- For a constant *a*

$$Var(aX) = a^2 Var(X)$$

ullet For any two random variables X and Y

$$Var(X + Y) = Var(X) + Var(Y) + 2Cov(X, Y)$$

$$Var(X-Y) = Var(X) + Var(Y) - 2Cov(X,Y)$$

ullet For independent random variables X and Y

$$Var(X + Y) = Var(X) + Var(Y)$$

$$Var(X - Y) = Var(X) + Var(Y)$$

### Conditional probabilities

- We will also spend *a lot of time* with conditional distributions and conditional expectations of random variables.
  - What's the probability that an individual enrolls in a job training program given their income?
  - We represent the conditioning set using a vertical bar with the right-hand side denoting what is being conditioned on.
    - For example:  $Pr(D_i = 1 | X_i = x)$

### Conditional probabilities

- **Key concept** Dependence and independence. If two variables are *independent*, the distribution of one does not change conditional on the other. We'll write this using the  $\bot$  notation.
- $Y_i \perp \!\!\! \perp D_i$  implies

$$Pr(Y_i = 1 | D_i = 1) = Pr(Y_i = 1 | D_i = 0) = Pr(Y_i = 1)$$

• Otherwise, the two variables are dependent:  $Y_i \perp \!\!\! \perp D_i$ .

$$Pr(Y_i=1|D_i=1)
eq Pr(Y_i=1|D_i=0)$$

• Two variables can be *conditionally independent* in that they are independent only when conditioning on a third variable. For example, we can have  $Y_i \perp \!\!\! \perp D_i$  but  $Y_i \perp \!\!\! \perp D_i | X_i$ . This implies

$$Pr(Y_i = 1 | D_i = 1, X_i = x) = Pr(Y_i = 1 | D_i = 0, X_i = x) = Pr(Y_i = 1 | X_i = x)$$

• Remember: Conditional independence does not imply independence or vice-versa!

## Conditional expectations

- A central object of interest in statistics is the **conditional expectation** function (CEF) E[Y|X].
  - $\circ$  Given a particular value of X, what is the expectation of Y?
  - $\circ$  The CEF is a function of X.
- All the usual properties of expectations apply to conditional expectations.
- We also will often make use of the law of total expectation

$$E[Y] = E[E[Y|X]]$$

Easiest to think about this in terms of discrete rvs

$$E[Y] = \sum_{x \in \mathcal{X}} E[Y|X=x] Pr(X=x)$$

- One critical use of statistical theory is understanding how to learn about things we don't observe using things that we do observe. We call this estimation.
  - e.g. What is the share of voters in Wisconsin who will turn out in the 2022 election?
  - What is the share of voters who turn out among those assigned to receive a GOTV phone call?
- Estimand: The unobserved quantity that we want to learn about. Often denoted via a greek letter (e.g.  $\mu$ ,  $\pi$ )
  - o Often a "population" characteristic that we want to learn about via a sample.
    - But you'll learn another reason why we sometimes can't observe a quantity of interest even in a sample!
  - Important to define your estimand well. (Lundberg, Johnson and Stewart, 2022)

- Estimator: The function of random variables that we will use to try to estimate the quantity of interest. Often denoted with a hat on the parameter of interest (e.g.  $\hat{\mu}$ ,  $\hat{\pi}$ )
  - Why are the variables random?
    - Classic inference: We have a random sample from the population -- if we took another sample, we would obtain a different realization of our estimator.
    - Randomization inference: We have a randomly assigned treatment -- if we were to re-run the experiment, we would observe a different treatment/control allocation.
- Estimate: A single realization of our estimator (e.g. 0.3, 9.535)
  - We often report both point estimates ("best guess") and interval estimates (e.g. confidence intervals).
  - Careful not to confuse properties of estimators with properties of the estimates themselves.



estimand

estimator



from the heat.

chocolate has all melted remove

200g light muscovado

2 large eggs

estimate

- The classic estimation problem in statistics is to estimate some unknown population mean  $\mu$  from an i.i.d. sample of n observations  $Y_1, Y_2, \ldots, Y_n$ .
  - $\circ$  We assume that each  $Y_i$  is a draw from the target population with mean  $\mu$ . (identically distributed) -- therefore  $E[Y_i] = \mu$
  - $\circ$  We'll also assume that conditioning on  $Y_i$  tells us nothing about any other  $Y_j Y_i \perp \!\!\! \perp Y_j$  (independently distributed) -- this implies  $Cov(Y_i,Y_j)=0$
- Our estimand:  $\mu$
- Our estimator: The sample mean  $\hat{\mu} = \bar{Y} = \frac{1}{n} \sum_{i=1}^{n} Y_i$
- ullet Our estimate: A particular realization of that estimator based on our observed sample (e.g. 0.4)
- Note that our estimator is a random variable -- it's a function of  $Y_i$ s which are random variables.
  - $\circ$  Therefore it has an expectation  $E[\hat{\mu}]$  (assuming  $Y_i$  has an expectation)
  - $\circ$  It has a variance  $Var(\hat{\mu})$  (again, under regularity conditions)
  - It has a distribution (which we may or may not know).

- How do we know if we've picked a good estimator? Will it be close to the truth? Will it be systematically higher or lower than the target?
- We want to derive some of its properties
  - $\circ$  Bias:  $E[\hat{\mu}] \mu$
  - $\circ$  Variance:  $Var(\hat{\mu})$
  - $\circ$  Consistency: Does  $\hat{\mu}$  converge in probability to  $\mu$  as n goes to infinity?
  - Asymptotic distribution: Is the sampling distribution of  $\hat{\mu}$  well approximated by a known distribution?

#### Unbiasedness

- Is the expectation of  $\hat{\mu}$  equal to  $\mu$ ?
- First we pull out the constant.

$$E[\hat{\mu}] = E\left[rac{1}{n}\sum_{i=1}^n Y_i
ight] = rac{1}{n}E\left[\sum_{i=1}^n Y_i
ight].$$

Next we use linearity of expectations

$$\left[rac{1}{n}E\left[\sum_{i=1}^{n}Y_{i}
ight]=rac{1}{n}\sum_{i=1}^{n}E\left[Y_{i}
ight]$$

• Finally, under our i.i.d. assumption

$$rac{1}{n}\sum_{i=1}^n E\left[Y_i
ight] = rac{1}{n}\sum_{i=1}^n \mu = rac{n\mu}{n} = \mu$$

• Therefore, the bias,  $\mathrm{Bias}(\hat{\mu}) = E[\hat{\mu}] - \mu = 0$ 

#### Variance

• What is the variance of  $\hat{\mu}$ ? Again, start by pulling out the constant.

$$Var(\hat{\mu}) = Var\left[rac{1}{n}\sum_{i=1}^{n}Y_i
ight] = rac{1}{n^2}Var\left[\sum_{i=1}^{n}Y_i
ight]$$

• We can further simplify by using our i.i.d. assumption. The variance of a sum of i.i.d. random variables is the sum of the variances.

$$\left[rac{1}{n^2}Var\left[\sum_{i=1}^nY_i
ight] = rac{1}{n^2}\sum_{i=1}^nVar\left[Y_i
ight]$$

• "identically distributed"

$$rac{1}{n^2}\sum_{i=1}^n Var\left[Y_i
ight] = rac{1}{n^2}\sum_{i=1}^n \sigma^2 = rac{n\sigma^2}{n^2} = rac{\sigma^2}{n}$$

• Therefore, the variance is  $\frac{\sigma^2}{n}$ 

## Asymptotic behavior

- As n gets large, what can we say about the estimator  $\hat{\mu}$ .
- First, we can show that it is **consistent** -- it converges in probability to the true parameter  $\mu$ 
  - $\circ$  Unbiasedness + Variance that goes to 0 as n gets large.
  - $\circ$  Some estimators may be biased but have bias terms that go to 0 -- if variance also goes to 0 these are still consistent.
- Second, we can say something about the distribution of  $\hat{\mu}$ .
  - $\circ$  Remember, we've only made assumptions about  $E[Y_i]$  and  $Var(Y_i)$  (that they exist). We have made no assumptions on the distribution of  $Y_i$ .  $Y_i$  can be normal, poisson, bernoulli, or whatever!
  - However, we know something about sums and means of random variables -- they are well-approximated by a normal distribution. The Central Limit Theorem!
  - $\circ$  So in large samples, the **sampling distribution** of  $\hat{\mu}$  is close to normal. This lets us construct confidence intervals and do inference with this approximation and be confident that we won't be far off!

The potential outcomes model

## Thinking about causal effects

- Two types of causal questions (Gelman and Rubin, 2013)
- Causes of effects
  - What are the factors that cause some outcome *Y*?
  - "Why?" questions: Why do states go to war? Why do politicians get re-elected?
- Effects of causes
  - If *X* were to change, what might happen to *Y*?
  - "What if?" questions: What if a pair of states were democratic, would that change their chances of going to war? What if a politician were an incumbent, would that affect their re-election probability?
- We'll spend more time on the effects of causes
  - Why? Because we can connect them to well-defined statistical quantities of interest (e.g. an "average treatment effect")
  - "Causes of effects" are still important questions, and theoretical inquiry can lead you towards possible causes to evaluate.

# Defining a causal effect

- Historically, causality was seen as a deterministic process.
  - Hume (1740): Causes are regularities in events of "constant conjunctions"
  - Mill (1843): Method of difference
- This became problematic -- empirical observation alone does not demonstrate causality.
  - Russell (1913): Scientists aren't interested in causality!
- How do we talk about causation that both incorporates *uncertainty* in measurement and clearly defines what we mean by a "causal effect"?
- Rubin (1974) formalizes a framework for understanding causation from a statistical perspective. Inspired by earlier Neyman (1923) and Fisher (1935) on randomized experiments.
- We'll spend most of our time with this approach, often called the Rubin Causal Model or potential outcomes framework.

### Causality and interventions

- Causal effects are effects of interventions
  - What happens to an outcome when a treatment is changed.
- It's very difficult to learn about vague causal statements:
  - What is the effect of growth on democracy or conflict?
  - What is the effect of exercise on health?
- The potential outcomes framework clarifies:
  - 1. What action is doing the causing?
  - 2. Compared to **what** alternative action?
  - 3. On what outcome metric?
  - 4. How would we learn about the effect from data?

### Statistical setup.

- Population of units
  - Finite population or infinite super-population
- Sample of N units from the population indexed by i Observed outcome  $Y_i$
- Binary treatment indicator  $D_i$ .
  - $\circ$  Units receiving "treatment":  $D_i = 1$
  - $\circ$  Units receiving "control":  $D_i=0$
- Covariates (observed prior to treatment)  $X_i$

### Potential outcomes

- Let  $D_i$  be the value of a treatment assigned to each individual.
- $Y_i(d)$  is the value that the outcome would take if  $D_i$  were set to d.
  - $\circ$  For binary  $D_i$ :  $Y_i(1)$  is the value we would observe if unit i were treated.
  - $\circ Y_i(0)$  is the value we would observe if unit i were under control
- We model potential outcomes as fixed attributes of the units.
- Notation alert! -- Sometimes you'll see potential outcomes written as:
  - $\circ \ Y_i^1$ ,  $Y_i^0$  or  $Y_i^{d=1}$ ,  $Y_i^{d=0}$
  - $\circ \ Y_{i0}$ ,  $Y_{i1}$
  - $\circ \ Y_1(i)$ ,  $Y_0(i)$
- Causal effects are contrasts in potential outcomes.
  - $\circ$  Individual treatment effect:  $au_i = Y_i(1) Y_i(0)$
  - $\circ$  Can consider ratios or other transformations (e.g.  $rac{Y_i(1)}{Y_i(0)}$ )

# Consistency/SUTVA

- How do we link the potential outcomes to observed ones?
- Consistency/Stable Unit Treatment Value (SUTVA) assumption

$$Y_i(d) = Y_i \text{ if } D_i = d$$

• Sometimes you'll see this w/ binary  $D_i$  (often in econometrics)

$$Y_i = Y_i(1)D_i + Y_i(0)(1 - D_i)$$

- Implications
  - 1. No interference -- other units' treatments don't affect i's potential outcomes.
  - 2. Single version of treatment
  - 3. D is in principle manipulable -- a "well-defined intervention"
  - 4. The means by which treatment is assigned is irrelevant (a version of 2)

# Positivity/Overlap

- We also need some assumptions on the treatment assignment mechanism  $D_i$ .
- In order to be able to observe *some units'* values of  $Y_i(1)$  or  $Y_i(0)$  treatment can't be deterministic. For all i:

$$0 < Pr(D_i = 1) < 1$$

- If no units could ever receive treatment or control it would be impossible to learn about  $E[Y_i|D_i=1]$  or  $E[Y_i|D_i=0]$
- This is sometimes called a **positivity** or overlap assumption.
  - $\circ$  Pretty trivial in a randomized experiment, but can be tricky in observational studies when  $D_i$  is perfectly determined by some covariates  $X_i$

# A missing data problem

• It's useful to think of the causal inference problem in terms of *missingness* in the complete table of potential outcomes.

| Unit i | Treatment $D_i$ | $Y_i(1)$ | $Y_i(0)$ | Observed $Y_i$ |
|--------|-----------------|----------|----------|----------------|
| 1      | 1               | 5        | ?        | 5              |
| 2      | 0               | ?        | -3       | -3             |
| 3      | 1               | 9        | ?        | 9              |
| •      | :               | •        | •        | :              |
| N      | 0               | ?        | 8        | 8              |

- If we could observe both  $Y_i(1)$  and  $Y_i(0)$  for each unit, then this would be easy!
- ullet But we can't we only observe what we're given by  $D_i$
- Holland (1986) calls this "The Fundamental Problem of Causal Inference"

### Causal Estimands

- All causal inference starts with a definition of the estimand.
- The individual causal effect:  $\tau_i$

$$au_i = Y_i(1) - Y_i(0)$$

- Problem: Can't identify this without extremely strong assumptions!
- "The Fundamental Problem of Causal Inference"
- The average treatment effect (ATE): au

$$au=E[Y_i(1)-Y_i(0)]$$

What is the expected effect of the treatment?

#### Causal Estimands

• The conditional average treatment effect (CATE):  $\tau(x)$ 

$$au(x)=E[Y_i(1)-Y_i(0)|X_i=x]$$

- $\circ$  The average treatment effect for the sub-population with a particular set of covariate values  $X_i = x$
- The average treatment effect on the treated (ATT):  $au_{
  m ATT}$

$$au_{ ext{ATT}} = E[Y_i(1) - Y_i(0)|D_i=1]$$

- The average treatment effect among those units that were assigned treatment.
- Why do we care about this?
  - Sometimes more policy-relevant.
  - Often easier to identify in selection-on-observables designs.

### Causal Estimands

• The **risk ratio** (RR) for binary outcomes:

$$RR=rac{Pr(Y_i(1)=1)}{Pr(Y_i(0)=1)}$$

- Political science and social sciences rarely use these.
- Medicine *loves* risk ratios (e.g. Vaccine effectiveness = 1 RR)

### Finite Population Estimands

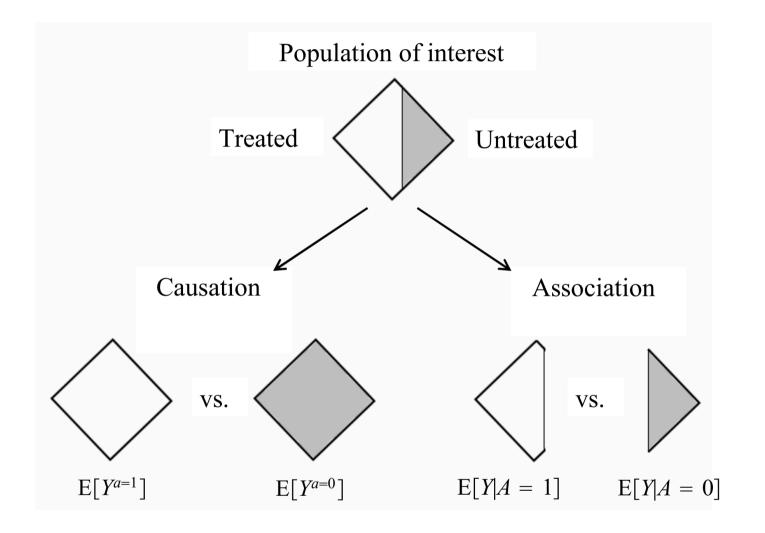
- For these estimands we have two sources of uncertainty
  - Sampling from a population.
  - Random assignment of treatment (unobserved P.O.s)
- What if we're just interested in doing inference on the treatment effect in the sample, or there is no population from which we sample from?
  - We can define analogues of the ATE for the sample -- the sample average treatment effect

$$au_{ ext{SATE}} = rac{1}{N} \sum_{i=1}^N Y_i(1) - Y_i(0)$$

### Finite Population Estimands

- There's *still* uncertainty -- we don't observe all of the potential outcomes. But now we don't have to justify generalizing to the population.
  - o Under random sampling, the SATE is unbiased for the (population) ATE or PATE.
  - Large literature on when and how we can generalize absent random sampling ("transportability")
- Inference justified on the basis of *randomization* of the treatment.
  - A lot of the same tools for inference in classical super-population sampling transfer over nicely.

### Causal vs. Associational Estimands



#### Identification

- In statistics, we often talk about whether a parameter is *identifiable*.
  - A parameter is "point identified" if, having access to infinite data, the parameter could only take on a single value.
  - In other words, no other value of the parameter could generate the same observable data.
- In statistical models, non-identifiability arises when different values of the parameter could give rise to the same observable data
  - In classical regression: More parameters than observations.
  - Or perfectly collinear regressors

#### Causal Identification

- Causal identification: Can we learn about the value of a causal effect from the observed data.
  - $\circ$  Can we express the causal estimand (e.g.  $au = E[Y_i(1) Y_i(0)]$ ) entirely in terms of *observable* quantities
- Causal identification comes prior to questions of estimation
  - It doesn't matter whether you're using regression, weighting, matching, doubly-robust estimation, double-LASSO, etc...
  - If you can't answer the question "What's your identification strategy?" then no amount of fancy stats will solve your problems.
- Identification requires assumptions about the connection between the observed data  $Y_i$ ,  $D_i$  and the unobserved counterfactuals  $Y_i(d)$ 
  - Under what assumptions will the observed difference-in-means identify the average treatment effect?

#### Conclusion

- Causal effects are contrasts in counterfactuals
- The potential outcomes framework gives us a tool for defining statistical *estimands* in terms of functions of individual effects of treatment.
  - $\circ$  The potential outcome  $Y_i(d)$  denotes the outcome  $Y_i$  we would observe for unit i if they were assigned treatment d
  - Implicitly making a consistency/SUTVA assumption to link observed to counterfactual
- Fundamental problem of causal inference: We're going to need more assumptions -- don't get to directly observe individual treatment effects.
- Consistency and positivity only get us part of the way there.
- Next week: What additional assumptions do we need to **identify** causal effects from the observed data. Why randomized experiments satisfy those assumptions.