

Interference Modeling in Small Cell Networks

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1 Mathematical Background

The purpose of this document is to provide the mathematical foundation to understand analytical network capacity analysis under spatial structures using point processes (PP). The network we are modeling contains an infinite number of transmitters denoted by the homogeneous Poisson PP (PPP) Φ , with intensity measure λ . A point process is simply a collection of random variables spatial distributed on a plain according to specified distribution. A PPP can be informally understood as a PP with uniform spaces points in an area B with the number of points in areas of size B being Poisson distributed with mean λ . For simplicity we will only be working with two dimensional plains, points $x \in \Phi$ and $x \in \mathbb{R}^2$, avoiding much of the confusing introduced by heavy measure theory based work. λ is a first order statistical property of a PP, defined as the mean number of points/events N per unit area/volume v in any set B [1]. Equation (1) show this relation, where \mathbf{E} represents the expectation operation. This is analogous to density of points in region or area.

$$\lambda = \frac{\mathbf{E}[N(B)]}{v(B)} \quad (1)$$

1.1 Palm Probability

Palm characteristics are probabilities or means that refer to individual points in a PP. Meaning that we want statistics from individual points perspectives. The usual approach considers a point at the origin $\mathbf{0}$. However, the probability that a stationary point process has a point exactly at $\mathbf{0}$ is zero. Therefore, the probability that a PP Φ has some property provided that it has a point at $\mathbf{0}$ is a difficult quantity.

We determine the mean and probability related to an event of having a point at $\mathbf{0}$ as follows. Consider an observation window W in which Φ has $\Phi(W) = n$ points. These points x_1, \dots, x_n are taken in turn and Φ is shifted such that the relevant point lies at the origin $\mathbf{0}$. This process is repeated for each x_i . For each shift, the number of points is counted within r i.e. $b(\mathbf{0}, r)$. Their average yields an estimate of the mean number of points in a sphere of radius r centered at a PP point, where in all cases the point x_i itself is never counted. Second, each point is then marked. If within the radius r , it will receive a mark 1 and 0 otherwise. Then all the marks are then added and divided by n , the number of points in Φ . This is an estimate of the probability that a point in the PP has its nearest neighbor at a distance less than r [1].

Mathematically in the Palm sense, the mean and probabilities are show in equations (2) and (3) respectively.

$$\mathbf{E}_{\mathbf{0}}(\Phi(b(\mathbf{0}, r) \setminus \{\mathbf{0}\})) \quad (2)$$

$$\mathbf{P}_{\mathbf{0}}(\Phi(b(\mathbf{0}, r) \setminus \{\mathbf{0}\}) > 0) \quad (3)$$

2 Summary of Existing Work

[2]

- For CSMA: provides probability of transmission (media access), coverage probabilities, and rough estimates for successful transmission density
- For Aloha: provides

3 CSMA

Representation of a CSMA network requires use of the Matérn hard-core PP. This PP requires a minimum distance r between points, or in our case transmitting nodes. A Matérn PP is generated from a marked PPP, where each point x uniformly selects a mark $m(x) \in [0, 1]$. Points with marks larger than their neighboring point's marks are deleted if their positions are within range r . Mathematically if $m(x) < m(y)$ and $\|x - y\| \leq r$, point y is deleted. This is process called conservative, since points that are already deleted can delete other points with larger marks. This is inaccurate in CSMA networks, but it is a common technique in the literature literature [3].

To analyze the intensity of such PP, requires conditioning on existence of points at specific locations. Palm theory formalizes the notion of the conditional distribution of a PP given it has a point at some location. Note that for a PP without a fixed *atom* at this particular location, the probability of the condition is equal to zero and the basic discrete definition of the conditional probability does not apply [4].

3.1 Slotted Aloha

Before we considered the complicated case of CSMA, let us first look at a simpler case of slotted aloha with PPP distributed transmitters. In this case we will define the coverage, or non-outage, probability for a *typical* node. A transmit receive pair is able to communicate if the SINR at the receiver is above some threshold T . To say it in another way, transmitter X_i *covers* receiver y_i in the reference time slot if equation (4) is satisfied.

$$\text{SINR}_i = \frac{F_i^i / l(\|X_i - y_i\|)}{W + I_i^1} \geq T \quad (4)$$

Where $I_i^1 = \sum_{X_j \in \hat{\Phi}^1, j \neq i} F_j^i / l(\|X_j - y_i\|)$ is the shot noise associated with the marked PP $\hat{\Phi}^1$. Now under this condition let us apply the mark δ_i to each transmitter x_i meeting this condition. The newly marked points form the PP $\hat{\Phi}$.

More introduction needed on next step to explain more on ideas be-

We denote \mathbf{P}^0 as the Palm probability associated to the δ marked stationary PP $\hat{\Phi}$. By probability of coverage of a typical node given it is a transmitter, we show is equation (5)

$$\begin{aligned}\mathbf{P}^0\{\delta_0 = 1|e_0 = 1\} &= \mathbf{E}^0[\delta_0|e_0 = 1] \\ &= \frac{1}{\lambda_1\|B\|}\mathbf{E}[\sum_i \delta_i \mathbb{1}(x_i \in B)]\end{aligned}\quad (5)$$

Equation (5) line one is understood as given the central point $\mathbf{0}$ is selected/able to transmit ($e_1 = 1$), the Palm Probability that the receiver y_0 has significant SINR (equation (5) is met) from transmitter x_0 located at $\mathbf{0}$. This probability is equal to the average δ_0 mark of the central located points of transmitters which are selected to transmit, seen by the r.h.s. of (5).

Unlike in Aloha type MAC schemes, in this Matérn CSMA model, the probability of medium access of a typical node $p = \mathbf{E}^0[e_i]$ is not given *a priori* and it has to be determined. e_i are simply the medium access indicators, marking the points that survived the Matérn process creation from the original PPP Φ . When calculating the probability of access to the channel, we must first calculate the expected number of neighbors of a typical node, which we denote as $\bar{N}(\lambda)$. From [2] we have Equation (6).

$$\begin{aligned}\bar{N} &= \mathbf{E}^0\left[\sum_{(x_j, m_j, \mathbf{F}_j^0) \in \hat{\Phi}} \mathbb{1}\left(\frac{F_j^0}{l(\|x_i - x_j\|)} \geq P_0\right)\right] \\ &= \lambda \int_{\mathbb{R}^2} \mathbf{P}\{F \geq P_0 l(\|x\|)\} dx \\ &= 2\pi\lambda \int_0^\infty (1 - G(P_0 l(r))) r dr\end{aligned}\quad (6)$$

In equation (6) F_j^0 denotes the power emitted by node j , given that it is selected to transmit, towards the origin $\mathbf{0}$. $l(\cdot)$ is the pathloss model used in the scenario (Ex: Power Law $l(r) = r^{-\alpha}$), P_0 is the detection threshold for the receiver at the $\mathbf{0}$. G is simply the CDF of the fading of the transmitters. If F is Rayleigh fading (exponential F with mean $\frac{1}{\mu}$), \bar{N} becomes:

$$\bar{N} = 2\pi\lambda \int_0^\infty e^{-P_0\mu l(r)} r dr \quad (7)$$

Taking equation (6) and rewriting the probability of medium access, we get equation (8) based on void probability.

$$p = \mathbf{E}[e_0] = \int e^{-\bar{N}t} dt = \frac{1 - e^{-\bar{N}}}{\bar{N}} \quad (8)$$

Note by the ergodicity of the model that p represents some spatial average, namely an average over all nodes of the PPP realizations. More precisely λp is the density of the Matérn PP of nodes authorized to transmit at a given time slot. However, it does not mean that any given node is selected by the MAC with the time frequency p , even if in different time slots the marks of each point are re-sampled independently. This p is better understood as a network wide proportion, rather than a local average.

4 System Model

x

5 Notation

e_i	the medium access indicator of node i ; ($e_i = 1$ if node i is allowed to transmit in the considered time slot and 0 otherwise). The random variables e_i are hence <i>i.i.d.</i> and independent of everything else, with $\mathbf{P}(e_i = 1) = p$ (p is the MAP).
F_i^j	denotes the virtual power emitted by node i (provided $e_i = 1$) towards receiver y_j

6 LTE Interference

In this section we will discuss the application of LTE femtocells (FC) in the stochastic geometry framework. We will only be considering co-channel interference among FC, assuming they are operating in disjoint channels from other LTE base stations such as Macrocells (MC). FC's are assumed to be uncoordinated nodes that are uniformly spread across the environment. Such nodes are uncoordinated because they lack the low-latency dedicated links that exists at the higher tiered nodes, known as X2 links. MC

FC's utilize an orthogonal frequency division multiple access (OFDMA) medium access (MAC) scheme to transmit data to mobile users called UE's in the downlink.

We extend the notions from the slotted Aloha based analysis into this framework. First we will provide a simplified model assuming a fixed resource usage or fractional usage of resources. Then we will add additional complexity by providing a stochastic traffic model to the resource allocation of each FC. This traffic model will be discussed first, and derivations provided into the geometry framework.

At a high level if we ignore UE feedback, we can model resource selection as an independent thinning of the network intensity λ_f by the fraction of channels used p , creating a new PPP with intensity λ_{tf} .

Add more background about LTE and femto-cells

Add info about resource block terminology

References

- [1] J. Illian, *Statistical analysis and modelling of spatial point patterns*, 2008. [Online]. Available: <http://books.google.com/books?id=i9LAm7QCtu8C>
- [2] F. Baccelli and B. laszczyszyn, "Stochastic geometry and wireless networks, volume," in *II-Applications. Foundations and Trends in Networking. NoW Publishers*. Citeseer, 2009.
- [3] H. Q. Nguyen, F. Baccelli, and D. Kofman, "A Stochastic Geometry Analysis of Dense IEEE 802.11 Networks," *IEEE INFOCOM 2007 - 26th IEEE International Conference on Computer Communications*, pp. 1199–1207, 2007. [Online]. Available: <http://ieeexplore.ieee.org/lpdocs/epic03/wrapper.htm?arnumber=4215725>

- [4] F. Baccelli and B. Błaszczyszyn, “Stochastic geometry and wireless networks, volume,” in *I-Theory. Foundations and Trends in Networking. Now Publishers*. Citeseer, 2009.