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9. Multivariate Gaussian Distribution

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Exercises due Jun 29, 2021 19:59 EDT

Note: Now is a good time to review Gaussian random variables from [Lecture 2](#).

Video Note: In the slide of the video below, there is a typo in the formula of the pdf of the multivariate Gaussian distribution: the exponent d in overall scaling factor should apply only to 2π , rather than $2\pi\det\Sigma$. The correct version is in the note below the video. (The unannotated slides in the resource section have also been corrected).

Multivariate Gaussian Distribution: Definition



multivariate PDF by heart.

What I'm really trying to say is if you give me

a covariance matrix and a vector of expectations--

the expectation vector, I guess--

then I actually can fully characterize a Gaussian PDF.

I have just an entire distribution.

Just like if you give me a expectation and a variance,

I have a Gaussian PDF in one dimension.

Video

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Multivariate Gaussian Random Variable

A random vector $\mathbf{X} = (X^{(1)}, \dots, X^{(d)})^T$ is a **Gaussian vector** , or **multivariate Gaussian or normal variable** , if any linear combination of its components is a (univariate) Gaussian variable or a constant (a "Gaussian" variable with zero variance), i.e., if $\alpha^T \mathbf{X}$ is (univariate) Gaussian or constant for any constant non-zero vector $\alpha \in \mathbb{R}^d$.

The distribution of \mathbf{X} , the **d -dimensional Gaussian or normal distribution** , is completely specified by the vector mean $\mu = \mathbb{E}[\mathbf{X}] = (\mathbb{E}[X^{(1)}], \dots, \mathbb{E}[X^{(d)}])^T$ and the $d \times d$ covariance matrix Σ . If Σ is invertible, then the pdf of \mathbf{X} is

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{\sqrt{(2\pi)^d \det(\Sigma)}} e^{-\frac{1}{2}(\mathbf{x}-\mu)^T \Sigma^{-1}(\mathbf{x}-\mu)}, \quad \mathbf{x} \in \mathbb{R}^d$$

where $\det(\Sigma)$ is the determinant of the Σ , which is positive when Σ is invertible.

If $\mu = \mathbf{0}$ and Σ is the identity matrix, then \mathbf{X} is called a **standard normal random vector** .

Note that when the covariant matrix Σ is diagonal, the pdf factors into pdfs of univariate Gaussians, and hence the components are independent.

Linear Transformation of a Multivariate Gaussian Random Vector

1 point possible (graded)

($X^{(1)}$)

(1 2)

(0)

Consider the **2**-dimensional Gaussian $\mathbf{X} = \begin{pmatrix} X^{(1)} \\ X^{(2)} \end{pmatrix}$ with covariance matrix $\Sigma_{\mathbf{X}} = \begin{pmatrix} 1 & 2 \\ 2 & 5 \end{pmatrix}$ and mean $\mu_{\mathbf{X}} = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$.

Consider the vector $\alpha = \begin{pmatrix} 1 \\ -1 \end{pmatrix}$, so that $Y = \alpha^T \mathbf{X}$ is a **1**-dimensional Gaussian.

What is the variance **Var**(Y) of Y ?

Var(Y) =

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You have used 0 of 3 attempts

Singular Covariance Matrices

1 point possible (graded)

Consider again a **2**-dimensional Gaussian $\mathbf{X} = \begin{pmatrix} X^{(1)} \\ X^{(2)} \end{pmatrix}$. But instead, $\Sigma_{\mathbf{X}}$ is $\begin{pmatrix} 1 & 2 \\ 2 & 4 \end{pmatrix}$ and $\alpha = \begin{pmatrix} 2 \\ -1 \end{pmatrix}$, what is the variance **Var**(Y) of $Y = \alpha^T \mathbf{X}$?

Var(Y) =

This result tells us that the Gaussian $(X^{(1)}, X^{(2)})^T$ is actually a one-dimensional Gaussian, orthogonal to the direction of $\begin{pmatrix} 2 \\ -1 \end{pmatrix}$.

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(Optional) Diagonalization of the Covariance Matrix

Let Σ be a covariance matrix of size $d \times d$. Note that its entries are all real numbers with diagonal elements being non-negative. Σ has the following properties:

- Σ is symmetric. That is, $\Sigma = \Sigma^T$.
- Σ is diagonalizable to a diagonal matrix D via a transformation $D = U \Sigma U^T$, where U is an orthogonal matrix (recall that a square matrix A is orthogonal if $AA^T = A^T A = I$, where I is the identity matrix). This implies that $\Sigma = U^T D U$.
- Moreover, Σ is positive semidefinite. That is, the diagonal matrix D has diagonal entries that are all non-negative.
- Σ has a unique positive semidefinite square root matrix. That is, there exists a positive semi-definite matrix $\Sigma^{\frac{1}{2}}$ that is unique such that $\Sigma^{\frac{1}{2}} \cdot \Sigma^{\frac{1}{2}} = \Sigma$.
- If Σ is of size $d \times d$, then it has d orthonormal eigenvectors (even if there are repeated eigenvalues). Furthermore, if U is a matrix with rows corresponding to the orthonormal eigenvectors, then the diagonal matrix $D = U \Sigma U^T$ contains the eigenvalues of Σ along its diagonal. Therefore, diagonalization of a symmetric matrix involves finding its eigenvalues and the orthonormal eigenvectors.
- If Σ is positive definite, i.e. the diagonal matrix $D = U \Sigma U^T$ has diagonal entries that are all strictly positive, then it is invertible and the inverse Σ^{-1} satisfies the following: $\Sigma^{-\frac{1}{2}} \cdot \Sigma^{-\frac{1}{2}} = \Sigma^{-1}$, where $\Sigma^{-\frac{1}{2}}$ is the inverse of the square root of Σ .

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(Optional) Gaussian Random Vectors I

0 points possible (ungraded)
Recall from an earlier part of this lecture that the covariance between two random variables being 0 does not necessarily imply that the random variables are independent. However, this is true if the random variables are multivariate Gaussian.

Let \mathbf{X} be a Gaussian random vector with mean μ and covariance Σ . Assume that Σ is positive definite. Determine if the following statement is true or false.

“There exists a vector B and a matrix A such that $A(\mathbf{X} + B)$ is a Gaussian random vector whose components are independent and each of mean 0 ”.

☐ True

☐ False

Hint: Refer to the note above on diagonalization of the covariance matrix.

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