

# STAT 3690 Lecture 11

zhiyanggeezhou.github.io

Zhiyang Zhou (zhiyang.zhou@umanitoba.ca)

Feb 16, 2022

## Testing on $\mathbf{A}\boldsymbol{\mu}$ (J&W pp. 279)

- $\mathbf{A}$  is of  $q \times p$  and  $\text{rk}(\mathbf{A}) = q$ , i.e.,  $\mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^\top > 0$
- Model: iid  $\mathbf{A}\mathbf{X}_i \sim MVN_q(\mathbf{A}\boldsymbol{\mu}, \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^\top)$ .
- LRT for  $H_0 : \mathbf{A}\boldsymbol{\mu} = \boldsymbol{\nu}_0$  v.s.  $H_1 : \mathbf{A}\boldsymbol{\mu} \neq \boldsymbol{\nu}_0$ 
  - Test statistic:  $n(\mathbf{A}\bar{\mathbf{X}} - \boldsymbol{\nu}_0)^\top (\mathbf{A}\mathbf{S}\mathbf{A}^\top)^{-1}(\mathbf{A}\bar{\mathbf{X}} - \boldsymbol{\nu}_0) \sim T^2(q, n-1) = \frac{(n-1)q}{n-q} F(q, n-q)$  under  $H_0$
  - Rejection region at level  $\alpha$ :  $R = \{\mathbf{x}_1, \dots, \mathbf{x}_n : \frac{n(n-q)}{q(n-1)} (\mathbf{A}\bar{\mathbf{x}} - \boldsymbol{\nu}_0)^\top (\mathbf{A}\mathbf{S}\mathbf{A}^\top)^{-1}(\mathbf{A}\bar{\mathbf{x}} - \boldsymbol{\nu}_0) \geq F_{1-\alpha, q, n-q}\}$
  - $p$ -value:  $p(\mathbf{x}_1, \dots, \mathbf{x}_n) = 1 - F_{F(q, n-q)}\{\frac{n(n-q)}{q(n-1)} (\mathbf{A}\bar{\mathbf{x}} - \boldsymbol{\nu}_0)^\top (\mathbf{A}\mathbf{S}\mathbf{A}^\top)^{-1}(\mathbf{A}\bar{\mathbf{x}} - \boldsymbol{\nu}_0)\}$
- Multiple comparison
  - Interested in  $H_0 : \mu_1 = \dots = \mu_p$  v.s.  $H_1$  : Not all entries of  $\boldsymbol{\mu}$  are equal.
    - \*  $\mu_k$ : the  $k$ th entry of  $\boldsymbol{\mu}$
  - Take
$$\boldsymbol{\nu}_0 = \mathbf{0}_{(p-1) \times 1}, \quad \mathbf{A} = \begin{bmatrix} 1 & -1 & 0 & \cdots & 0 \\ 1 & 0 & -1 & \cdots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 1 & 0 & 0 & \cdots & -1 \end{bmatrix}_{(p-1) \times p}.$$
  - $p = 2$  (i.e.,  $\mathbf{A} = [1, -1]$ ): A/B testing

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## $(1 - \alpha) \times 100\%$ confidence region (CR) for $\boldsymbol{\mu}$ (J&W Sec. 5.4)

- $\Pr((1 - \alpha) \times 100\% \text{CR covers } \boldsymbol{\mu}) = 1 - \alpha$ 
  - CR is a set made of observations and is hence random
  - $\boldsymbol{\mu}$  is fixed
  - $(1 - \alpha) \times 100\%$  CR covers  $\boldsymbol{\mu}$  with probability  $(1 - \alpha) \times 100\%$
- Dual problem of testing  $H_0 : \boldsymbol{\mu} = \boldsymbol{\mu}_0$  v.s.  $H_1 : \boldsymbol{\mu} \neq \boldsymbol{\mu}_0$  at the  $\alpha$  level
  - Translated from rejection region. Steps:
    1. Take  $R$  as a function of  $\boldsymbol{\mu}_0$ ;
    2. Replace  $\boldsymbol{\mu}_0$  with  $\boldsymbol{\mu}$ ;
    3. Take the complement.
  - $(1 - \alpha) \times 100\%$  CR =  $\{\boldsymbol{\mu} : n(\bar{\mathbf{x}} - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1}(\bar{\mathbf{x}} - \boldsymbol{\mu}) < \chi_{1-\alpha, p}^2\}$  if  $\boldsymbol{\Sigma}$  is known
  - $(1 - \alpha) \times 100\%$  CR =  $\{\boldsymbol{\mu} : \frac{n(n-p)}{p(n-1)} (\bar{\mathbf{x}} - \boldsymbol{\mu})^\top \mathbf{S}^{-1}(\bar{\mathbf{x}} - \boldsymbol{\mu}) < F_{1-\alpha, p, n-p}\}$  if  $\boldsymbol{\Sigma}$  is not known

$(1 - \alpha) \times 100\%$  **CR for  $\nu = \mathbf{A}\mu$**

- $\mathbf{X}_1, \dots, \mathbf{X}_n \stackrel{\text{iid}}{\sim} MVN_p(\mu, \Sigma)$ 
  - Unknown  $\Sigma$
  - $n > p$
- $\mathbf{A}$  is of  $q \times p$  and  $\text{rk}(\mathbf{A}) = q$ , i.e.,  $\mathbf{A}\Sigma\mathbf{A}^\top > 0$
- Then iid  $\mathbf{A}\mathbf{X}_i \sim MVN_q(\nu, \mathbf{A}\Sigma\mathbf{A}^\top)$
- $(1 - \alpha) \times 100\%$  CR for  $\nu$  is  $\{\nu : \frac{n(n-q)}{q(n-1)}(\mathbf{A}\bar{\mathbf{x}} - \nu)^\top (\mathbf{A}\mathbf{S}\mathbf{A}^\top)^{-1}(\mathbf{A}\bar{\mathbf{x}} - \nu) < F_{1-\alpha, q, n-q}\}$
- Special case:  $\mathbf{A} = \mathbf{a} \in \mathbb{R}^p$ 
  - $(1 - \alpha) \times 100\%$  confidence interval (CI) for scalar  $\nu = \mathbf{a}^\top \mu$  is

$$\{\nu : n(\mathbf{a}^\top \bar{\mathbf{x}} - \nu)^2 (\mathbf{a}^\top \mathbf{S} \mathbf{a})^{-1} < F_{1-\alpha, 1, n-1}\} = \left( \mathbf{a}^\top \bar{\mathbf{x}} - t_{1-\alpha/2, n-1} \sqrt{\mathbf{a}^\top \mathbf{S} \mathbf{a} / n}, \mathbf{a}^\top \bar{\mathbf{x}} + t_{1-\alpha/2, n-1} \sqrt{\mathbf{a}^\top \mathbf{S} \mathbf{a} / n} \right)$$

## Simultaneous confidence intervals

- Construct  $(1 - \alpha_k)$  CI for scalars  $\mathbf{a}_k^\top \mu$ , say  $\text{CI}_k$ ,  $k = 1, \dots, m$ , simultaneously
- Make sure  $\Pr(\bigcap_k \{\mathbf{a}_k^\top \mu \in \text{CI}_k\}) \geq 1 - \alpha$
- Bonferroni correction
  - Bonferroni inequality:

$$\Pr\left(\bigcap_{k=1}^m \{\mathbf{a}_k^\top \mu \in \text{CI}_k\}\right) = 1 - \Pr\left(\bigcup_{k=1}^m \{\mathbf{a}_k^\top \mu \notin \text{CI}_k\}\right) \geq 1 - \sum_{k=1}^m \Pr(\mathbf{a}_k^\top \mu \notin \text{CI}_k) = 1 - \sum_{k=1}^m \alpha_k$$

- Taking  $\alpha_k$  such that  $\alpha = \sum_{k=1}^m \alpha_k$ , e.g.,  $\alpha_k = \alpha/m$ , i.e.,

$$(\mathbf{a}_k^\top \bar{\mathbf{x}} - t_{1-\alpha/(2m), n-1} \sqrt{\mathbf{a}_k^\top \mathbf{S} \mathbf{a}_k / n}, \mathbf{a}_k^\top \bar{\mathbf{x}} + t_{1-\alpha/(2m), n-1} \sqrt{\mathbf{a}_k^\top \mathbf{S} \mathbf{a}_k / n})$$

- Working for small  $m$
- Scheffé's method
  - Let  $\text{CI}_{\mathbf{w}} = (\mathbf{w}^\top \bar{\mathbf{x}} - c\sqrt{\mathbf{w}^\top \mathbf{S} \mathbf{w} / n}, \mathbf{w}^\top \bar{\mathbf{x}} + c\sqrt{\mathbf{w}^\top \mathbf{S} \mathbf{w} / n})$  for all  $\mathbf{w} \in \mathbb{R}^p$
  - Derive that  $c = \sqrt{p(n-1)(n-p)^{-1} F_{1-\alpha, p, n-p}}$ 
    - \* By Cauchy-Schwarz:  $\{\mathbf{w}^\top (\bar{\mathbf{x}} - \mu)\}^2 = [(\mathbf{S}^{1/2} \mathbf{w})^\top \{\mathbf{S}^{-1/2} (\bar{\mathbf{x}} - \mu)\}]^2 \leq \{(\mathbf{w}^\top \mathbf{S} \mathbf{w})^\top / n\} \{n(\bar{\mathbf{x}} - \mu)^\top \mathbf{S}^{-1} (\bar{\mathbf{x}} - \mu)\} \Rightarrow$

$$\begin{aligned} \Pr\left(\bigcap_{k=1}^m \{\mathbf{a}_k^\top \mu \in \text{CI}_k\}\right) &\geq \Pr\left(\bigcap_{\mathbf{w} \in \mathbb{R}^p} \{\mathbf{w}^\top \mu \in \text{CI}_{\mathbf{w}}\}\right) = 1 - \Pr\left(\bigcup_{\mathbf{w} \in \mathbb{R}^p} \{\mathbf{w}^\top \mu \notin \text{CI}_{\mathbf{w}}\}\right) \\ &= 1 - \Pr\left(\bigcup_{\mathbf{w} \in \mathbb{R}^p} [\{\mathbf{w}^\top (\bar{\mathbf{X}} - \mu)\}^2 / \{(\mathbf{w}^\top \mathbf{S} \mathbf{w})^\top / n\} > c^2]\right) \\ &\geq 1 - \Pr(\{n(\bar{\mathbf{X}} - \mu)^\top \mathbf{S}^{-1} (\bar{\mathbf{X}} - \mu) > c^2\}) \end{aligned}$$

- \*  $\Pr(\{n(\bar{\mathbf{X}} - \mu)^\top \mathbf{S}^{-1} (\bar{\mathbf{X}} - \mu) > c^2\}) = \alpha \Rightarrow c = \sqrt{p(n-1)(n-p)^{-1} F_{1-\alpha, p, n-p}}$
- Working for large even infinite  $m$

### Comparing two multivariate means (J&W Sec. 6.3)

- Two independent samples of (potentially) different sizes from two distributions with equal covariance

$$- \mathbf{X}_{11}, \dots, \mathbf{X}_{1n_1} \stackrel{\text{iid}}{\sim} MVN_p(\boldsymbol{\mu}_1, \boldsymbol{\Sigma})$$

$$- \mathbf{X}_{21}, \dots, \mathbf{X}_{2n_2} \stackrel{\text{iid}}{\sim} MVN_p(\boldsymbol{\mu}_2, \boldsymbol{\Sigma})$$

- Let  $\bar{\mathbf{X}}_i$  and  $\mathbf{S}_i$  be the sample mean and sample covariance for the  $i$ th sample
- Hypotheses  $H_0 : \boldsymbol{\mu}_1 = \boldsymbol{\mu}_2$  v.s.  $H_1 : \boldsymbol{\mu}_1 \neq \boldsymbol{\mu}_2$
- Test statistic following LRT

$$(\bar{\mathbf{X}}_1 - \bar{\mathbf{X}}_2)^\top \{(n_1^{-1} + n_2^{-1})\mathbf{S}_{\text{pool}}\}^{-1}(\bar{\mathbf{X}}_1 - \bar{\mathbf{X}}_2) \sim \frac{(n_1 + n_2 - 2)p}{n_1 + n_2 - p - 1} F(p, n_1 + n_2 - p - 1)$$

- Rejection region at level  $\alpha$

$$\left\{ x_{11}, \dots, x_{1n_1}, x_{21}, \dots, x_{2n_2} : (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)^\top \{(n_1^{-1} + n_2^{-1})\mathbf{S}_{\text{pool}}\}^{-1}(\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2) \geq \frac{p(n_1 + n_2 - 2)}{n_1 + n_2 - p - 1} F_{1-\alpha, p, n_1 + n_2 - p - 1} \right\}$$

$$- \mathbf{S}_{\text{pool}} = \frac{(n_1 - 1)\mathbf{S}_1 + (n_2 - 1)\mathbf{S}_2}{n_1 + n_2 - 2}$$

- $p$ -value

$$1 - F_{F_{1-\alpha, p, n_1 + n_2 - p - 1}} \left[ \frac{n_1 + n_2 - p - 1}{p(n_1 + n_2 - 2)} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)^\top \{(n_1^{-1} + n_2^{-1})\mathbf{S}_{\text{pool}}\}^{-1}(\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2) \right]$$