

STAT 3690 Lecture 08

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Assumptions

- Model: $\mathbf{X}_1, \dots, \mathbf{X}_n \stackrel{\text{iid}}{\sim} MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, $n > p$
- Parameter space: $\Theta = \{(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \mid \boldsymbol{\mu} \in \mathbb{R}^p, \boldsymbol{\Sigma} \in \mathbb{R}^{p \times p}, \boldsymbol{\Sigma} > 0\}$

Method of moments (MM) estimators for $(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

- No requirement on normality
- Steps
 1. Equate raw moments to their sample counterparts:

$$\begin{cases} E(\mathbf{X}) = \bar{\mathbf{X}} \\ E(\mathbf{X}\mathbf{X}^\top) = n^{-1} \sum_i \mathbf{X}_i \mathbf{X}_i^\top \end{cases} \Leftrightarrow \begin{cases} \boldsymbol{\mu} = \bar{\mathbf{X}} \\ \boldsymbol{\Sigma} + \boldsymbol{\mu}\boldsymbol{\mu}^\top = n^{-1} \sum_i \mathbf{X}_i \mathbf{X}_i^\top \end{cases}$$

2. Solve the above equations w.r.t. $\boldsymbol{\mu}$ and $\boldsymbol{\Sigma}$ and obtain estimators

$$\begin{cases} \hat{\boldsymbol{\mu}}_{\text{MM}} = \bar{\mathbf{X}} \\ \hat{\boldsymbol{\Sigma}}_{\text{MM}} = n^{-1} \sum_i \mathbf{X}_i \mathbf{X}_i^\top - \bar{\mathbf{X}} \bar{\mathbf{X}}^\top = n^{-1}(n-1)\mathbf{S}, \end{cases}$$

where $\mathbf{S} = (n-1)^{-1} \sum_{i=1}^n (\mathbf{X}_i - \bar{\mathbf{X}})(\mathbf{X}_i - \bar{\mathbf{X}})^\top$

Maximum likelihood (ML) estimation for parameters of MVN (J&W Sec 4.3)

- Likelihood function

$$\begin{aligned} L(\boldsymbol{\mu}, \boldsymbol{\Sigma}) &= \prod_{i=1}^n \left[\frac{1}{\sqrt{(2\pi)^p \det(\boldsymbol{\Sigma})}} \exp \left\{ -\frac{1}{2} (\mathbf{X}_i - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{X}_i - \boldsymbol{\mu}) \right\} \right] \\ &= \frac{1}{\sqrt{(2\pi)^{np} \{\det(\boldsymbol{\Sigma})\}^n}} \exp \left\{ -\frac{1}{2} \sum_{i=1}^n (\mathbf{X}_i - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{X}_i - \boldsymbol{\mu}) \right\} \end{aligned}$$

- Log likelihood

$$\ell(\boldsymbol{\mu}, \boldsymbol{\Sigma}) = \ln\{L(\boldsymbol{\mu}, \boldsymbol{\Sigma})\} = -\frac{np}{2} \ln(2\pi) - \frac{n}{2} \ln\{\det(\boldsymbol{\Sigma})\} - \frac{1}{2} \sum_{i=1}^n (\mathbf{X}_i - \boldsymbol{\mu})^\top \boldsymbol{\Sigma}^{-1} (\mathbf{X}_i - \boldsymbol{\mu})$$

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options(digits = 4)
set.seed(1)
n = 1e3L
Mu = matrix(c(3, 6), ncol = 1, nrow = 2)
Sigma = matrix(c(10, 2, 2, 5), ncol = 2, nrow = 2)
X = MASS::mvrnorm(n, Mu, Sigma)

loglik <- function(Mu, Sigma, data = X) {
  n = nrow(data)
  p = length(Mu)
  X_center = sweep(X, 2, Mu)
  return(
    (-n*p/2)*log(2*pi)+
    (-n/2)*log(det(Sigma))+
    (-1/2)*sum(diag(X_center %*% solve(Sigma) %*% t(X_center)))
  )
}

grid_xy <- expand.grid(
  seq(Mu[1]-2*Sigma[1,1]^0.5, Mu[1]+2*Sigma[1,1]^0.5, length.out = 32),
  seq(Mu[2]-2*Sigma[2,2]^0.5, Mu[2]+2*Sigma[2,2]^0.5, length.out = 32))

contours <- purrr::map_df(
  seq_len(nrow(grid_xy)),
  function(i) {
    # Where we will evaluate loglik
    mu_obs <- as.numeric(grid_xy[i,])
    # Evaluate at the pop covariance
    z <- loglik(mu_obs, Sigma, X)
    # Output data.frame
    data.frame(x = mu_obs[1],
              y = mu_obs[2],
              z = z)
  })

# Contour plot
library(tidyverse)
library(ggrepel)
data_means <- data.frame(
  x = c(Mu[1], mean(X[,1])),
  y = c(Mu[2], mean(X[,2])),
  label = c("Mu", "Sample Mean"))

contours %>%
  ggplot(aes(x, y)) +
  geom_contour(aes(z = z)) +
  geom_point(data = data_means, aes(color = label)) +
  geom_label_repel(data = data_means, aes(label = label))

# 3d scatter plot
library(scatterplot3d)
with(contours, scatterplot3d(x, y, z))

```

- ML estimator

$$(\hat{\mu}_{ML}, \hat{\Sigma}_{ML}) = \arg \max_{(\mu, \Sigma) \in \Theta} \ell(\mu, \Sigma) = (\bar{X}, n^{-1}(n-1)S)$$

Derive $(\hat{\mu}_{ML}, \hat{\Sigma}_{ML})$

$$\ell(\mu, \Sigma) = \text{const} - \frac{n}{2} \ln \{\det(\Sigma)\} - \frac{1}{2} \underbrace{\sum_{i=1}^n (X_i - \mu)^T \Sigma^{-1} (X_i - \mu)}_{\textcircled{1}}$$

$$\begin{aligned} \textcircled{1} &= \sum_{i=1}^n \text{tr} \{ (X_i - \mu)^T \Sigma^{-1} (X_i - \mu) \} \\ &= \sum_{i=1}^n \text{tr} \left\{ \Sigma^{-1} (X_i - \mu) (X_i - \mu)^T \right\} \quad (\because \text{tr}(ABC) = \text{tr}(BCA)) \\ &= \text{tr} \left\{ \sum_{i=1}^n \Sigma^{-1} (X_i - \bar{X} + \bar{X} - \mu) (X_i - \bar{X} + \bar{X} - \mu)^T \right\} \\ &= \text{tr} \left[\Sigma^{-1} \left\{ \sum_{i=1}^n (X_i - \bar{X}) (X_i - \bar{X})^T + (\bar{X} - \mu) \sum_{i=1}^n \underbrace{(X_i - \bar{X})^T}_{\textcircled{2}} + \underbrace{\sum_{i=1}^n (X_i - \bar{X}) (\bar{X} - \mu)^T}_{\textcircled{2}} + n(\bar{X} - \mu) (\bar{X} - \mu)^T \right\} \right] \\ &= \text{tr} \left[\Sigma^{-1} \left\{ (n-1)S + n(\bar{X} - \mu) (\bar{X} - \mu)^T \right\} \right] \\ &= (n-1) \text{tr}(\Sigma^{-1}S) + n \underbrace{(\bar{X} - \mu)^T \Sigma^{-1} (\bar{X} - \mu)}_{\textcircled{2}} \quad (\because \text{tr}(ABC) = \text{tr}(BCA)) \end{aligned}$$

So, no matter what $\Sigma > 0$ is, $\ell(\mu, \Sigma)$ is maximised at $\hat{\mu}_{ML} = \bar{X}$. We then have

$$\begin{aligned} \ell(\hat{\mu}_{ML}, \Sigma) &= \text{const} - \frac{n}{2} \ln \{\det(\Sigma)\} - \frac{n-1}{2} \text{tr}(\Sigma^{-1}S) \\ \Rightarrow \frac{\partial \ell(\hat{\mu}_{ML}, \Sigma)}{\partial \Sigma} &= -\frac{n}{2} \{\det(\Sigma)\}^{-1} \det(\Sigma) (\Sigma^{-1})^T \quad (\because \partial \det(A) / \partial A = \det(A) (A^{-1})^T) \\ &\quad - \frac{n-1}{2} (-S^T \Sigma^{-2}) \quad (\because \partial \text{tr}(AB) / \partial A = B^T, \frac{\partial A^{-1}}{\partial A} = -A^{-2}) \\ &= -\frac{n}{2} \Sigma^{-1} + \frac{n-1}{2} S \Sigma^{-2} \end{aligned}$$

Let $\partial \ell(\hat{\mu}_{ML}, \Sigma) / \partial \Sigma = 0$. Then $\hat{\Sigma}_{ML} = \frac{n-1}{n} S$

NOT trivial to verify that $(\hat{\mu}_{ML}, \hat{\Sigma}_{ML})$ is the only maximizer.

- Properties of $(\hat{\mu}_{ML}, \hat{\Sigma}_{ML})$

- Consistency: $(\hat{\mu}_{ML}, \hat{\Sigma}_{ML}) \xrightarrow{P} (\mu, \Sigma)$.
- Efficiency: As $n \rightarrow \infty$, the covariance of $(\hat{\mu}_{ML}, \hat{\Sigma}_{ML})$ achieves the Cramer-Rao lower bound.
- Invariance: For any function g , the ML estimator of $g(\mu, \Sigma)$ is $g(\hat{\mu}_{ML}, \hat{\Sigma}_{ML})$.