STAT 4100 Lecture Note

Week Seven (Oct 17, 19, & 21, 2022)

Zhiyang Zhou (zhiyang.zhou@umanitoba.ca, zhiyanggeezhou.github.io)

2022/Oct/22 17:31:17

Review for midterm (con'd)

To cover CB Ex. 7.66, 7.58, & 7.57

Hypothesis Testing

Binary classification

- Assume $\mathbf{X} = [X_1, \dots, X_n]^\top \sim f(\boldsymbol{x} \mid \boldsymbol{\theta}^*) \in \{f(\boldsymbol{x} \mid \boldsymbol{\theta}) : \boldsymbol{\theta} \in \boldsymbol{\Theta}\}$
 - Fixed unknown $\boldsymbol{\theta}^*$ to be inferred
- Make a decision on θ^* between two hypotheses $H_0: \theta^* \in \Theta_0$ and $H_1: \theta^* \in \Theta_1$
 - $\mathbf{\Theta}_0 \cup \mathbf{\Theta}_1 = \mathbf{\Theta}$
 - $\mathbf{\Theta}_0 \cap \mathbf{\Theta}_1 = \emptyset$
- Decision and correctness
 - True positive (TP) = H_0 correctly rejected
 - False positive (FP, type I error) = H_0 incorrectly rejected
 - True negative (TN) = H_0 is correctly accepted
 - False negative (FN, type II error) = H_0 incorrectly accepted
- E.g., H_0 : healthy vs H_1 : sick
 - TP: sick people identified as sick
 - FP: healthy people identified as sick
 - TN: healthy people identified as healthy
 - FN: sick people identified as healthy

	Accept H_0	Reject H_0
H_0 is true	True negative (TN)	False positive (FP, type I error)
H_0 is false	False negative (FN, type II error)	True positive (TP)

- Misclassification rate = Pr(FP) + Pr(FN)
- False discovery rate $(FDR) = Pr(FP)/\{Pr(FP) + Pr(TP)\}$

- FDR controlling for sequential/simultaneous testing
- Receiver operating characteristic curve (ROC curve): plot of TPR vs FPR
 - True positive rate (TPR, sensitivity) = $Pr(TP)/\{Pr(TP) + Pr(FN)\}$
 - False positive rate $(FPR) = Pr(FP)/\{Pr(FP) + Pr(FN)\}$
 - Area under the ROC curve (AUC)
- True negative rate (TNR, specificity) = $Pr(TN)/\{Pr(TN) + Pr(FP)\}$
- The optimal hypothesis testing seeking to minimize Pr(FN) subject to capped Pr(FP), i.e.,

 $\min \Pr(\text{type II error}) \text{ subject to } \Pr(\text{type I error}) \leq \alpha$

Power function

- Rejection/critical region: $R = \{x : \text{data } x \text{ corresponding to the rejection of } H_0\}$
 - Typically specified in terms of a function of the sample (called the *test statistic*); e.g., if $R = \{x : \bar{x} \geq 3\}$, then \bar{X} is the test statistic.
- Test function $\phi : \text{supp}(\mathbf{X}) \to \{0,1\}$ defined as $\phi(\mathbf{x}) = \mathbf{1}_R(\mathbf{x})$
 - $-\phi(\mathbf{x})=1$ implying the rejection of H_0
- Each test function ϕ corresponds to a unique rejection region $R_{\phi} = \{x : \phi(x) = 1\}$
 - Two tests considered to be equivalent if they correpond to the same rejection region/test function
- Power function (for ϕ): $\beta_{\phi}(\boldsymbol{\theta}) = \Pr(\mathbf{X} \in R_{\phi} \mid \boldsymbol{\theta}) = \mathrm{E}\{\phi(\mathbf{X}) \mid \boldsymbol{\theta}\}$
 - Pr(type I error) = $\beta_{\phi}(\boldsymbol{\theta}^*)$ if H_0 is correct $(\boldsymbol{\theta}^* \in \boldsymbol{\Theta}_0)$
 - Pr(type II error) = $1 \beta_{\phi}(\boldsymbol{\theta}^*)$ if H_1 is correct $(\boldsymbol{\theta}^* \in \boldsymbol{\Theta}_1)$
- Prefer larger $\beta_{\phi}(\boldsymbol{\theta})$ for all $\boldsymbol{\theta} \in \boldsymbol{\Theta}_1$ and smaller $\beta_{\phi}(\boldsymbol{\theta})$ for all $\boldsymbol{\theta} \in \boldsymbol{\Theta}_0$ (because $\boldsymbol{\theta}^*$ is unknown)

Example Lec14.2

- iid $X_1, \ldots, X_n \sim N(\theta, \sigma_0^2)$ with known σ_0 . Consider a test for $H_0: \theta = \theta_0$ vs $H_1: \theta \neq \theta_0$ with rejection region $\{x: \sqrt{n}|\bar{x} \theta_0|/\sigma_0 > c\}$.
 - a. Elaborate the power function.
 - b. Find sample size n and threshold c if one desires that the type I error rate is 5% and the type II error rate at $\theta_0 + \sigma_0$ is 25%.

Uniformly most powerful (UMP) level α test (CB Sec 8.3.2)

- ϕ is of level α iff $\sup_{\theta \in \Theta_0} \beta_{\phi}(\theta) \leq \alpha$
 - $-\phi$ is of size α iff $\sup_{\boldsymbol{\theta}\in\boldsymbol{\Theta}_0}\beta_{\phi}(\boldsymbol{\theta})=\alpha$
- Let ϕ is a level α test for $H_0: \boldsymbol{\theta}^* \in \boldsymbol{\Theta}_0$ vs $H_1: \boldsymbol{\theta}^* \in \boldsymbol{\Theta}_1$. If $\beta_{\phi}(\boldsymbol{\theta}) \geq \beta_{\phi'}(\boldsymbol{\theta})$ for all $\boldsymbol{\theta} \in \boldsymbol{\Theta}_1$ and all ϕ' of level α , then ϕ is a UMP level α test.
- If ϕ is a UMP level α test, then $\beta_{\phi}(\boldsymbol{\theta}) \geq \alpha \geq \sup_{\boldsymbol{\theta} \in \boldsymbol{\Theta}_0} \beta_{\phi}(\boldsymbol{\theta})$ for all $\boldsymbol{\theta} \in \boldsymbol{\Theta}_1$ (unbiasedness for testing, CB Def 8.3.9)

UMP level α test for simple hypotheses $(H_0: \theta^* = \theta_0 \text{ vs } H_1: \theta^* = \theta_1)$

- To maximize $\beta_{\phi}(\boldsymbol{\theta}_1)$ with respect to ϕ subject to $\beta_{\phi}(\boldsymbol{\theta}_0) \leq \alpha$
- Neymann-Pearson (NP) Lemma (CB Thm 8.3.12): ϕ is the UMP test of level α for simple hypotheses $\iff \exists c > 0$ such that $\beta_{\phi}(\boldsymbol{\theta}_0) = \mathbb{E}\{\phi_c(\mathbf{X}) \mid \boldsymbol{\theta}^* = \boldsymbol{\theta}_0\} = \alpha$, where

$$\phi_c(\boldsymbol{x}) = \begin{cases} 1, & f(\boldsymbol{x} \mid \boldsymbol{\theta}_1) > cf(\boldsymbol{x} \mid \boldsymbol{\theta}_0), \\ 0, & f(\boldsymbol{x} \mid \boldsymbol{\theta}_1) < cf(\boldsymbol{x} \mid \boldsymbol{\theta}_0). \end{cases}$$

In practice (especially for discrete distributions), c is taken as the largest real number such that

$$\Pr\{f(\mathbf{X} \mid \boldsymbol{\theta}_1) / f(\mathbf{X} \mid \boldsymbol{\theta}_0) \ge c \mid \boldsymbol{\theta}^* = \boldsymbol{\theta}_0\} \ge \alpha$$

and

$$\Pr\{f(\mathbf{X} \mid \boldsymbol{\theta}_1) / f(\mathbf{X} \mid \boldsymbol{\theta}_0) \le c \mid \boldsymbol{\theta}^* = \boldsymbol{\theta}_0\} \ge 1 - \alpha.$$

- According to NP Lemma, for simple hypotheses, UMP test at level $\alpha \iff$ UMP test at size α .
- UMP test and sufficiency (CB Coro 8.3.13): sufficient statistics can be taken as test statistics for UMP test.