STAT 4100 Lecture Note

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Univariate transformation (con'd)

Find pdf of Y = g(X) given the distribution of X

- 1. Figure out supp $(Y) = \{y : y = g(x), x \in \text{supp}(X)\}$
- 2. (Generically) If the cdf F_Y is known OR pdf f_X is easy to be integrated, then

$$f_Y(y) = \frac{\mathrm{d}}{\mathrm{d}y} F_Y(y) = \frac{\mathrm{d}}{\mathrm{d}y} \int_{\{x:g(x) \le y\}} f_X(x) \mathrm{d}x$$

• The integration of f_X is often avoidable by employing the Leibniz Rule (CB Thm. 2.4.1):

$$\frac{\mathrm{d}}{\mathrm{d}y} \int_{a(y)}^{b(y)} f(x) \mathrm{d}x = f\{b(y)\} \frac{\mathrm{d}}{\mathrm{d}y} b(y) - f\{a(y)\} \frac{\mathrm{d}}{\mathrm{d}y} a(y)$$

with a(y) and b(y) both differentiable with respect to y.

2. (Alternatively) According to CB Ex. 2.7(b), i.e., an extension of CB Thm. 2.1.5 & 2.1.8 and HMC Thm 1.7.1.

$$f_Y(y) = \sum_{k=1}^K f_X\{g_k^{-1}(y)\} \left| J_{g_k^{-1}} \right| \mathbf{1}_{B_k}(y)$$

- Partition supp(X) into K intervals A_1, \ldots, A_K such that $\bigcup_{k=1}^K A_k = \text{supp}(X)$ and $A_k \cap A_{k'} = \emptyset$
- g_k is strictly monotonic on A_k and $g(x)=g_k(x)$ for all $x\in A_k$ g_k^{-1} is continuously differentiable on $B_k=\{g_k(x):x\in A_k\}$ Jacobian of transformation g_k^{-1}

$$J_{g_k^{-1}} = \frac{\mathrm{d}}{\mathrm{d}y} g_k^{-1}(y)$$

Example Lec2.2

Let X have the uniform pdf $f_X(x) = \pi^{-1} \mathbf{1}_{(-\pi/2,\pi/2)}(x)$. Find the pdf of $Y = \tan X$.

Example Lec2.3

 $X \sim \text{Weibull}(\text{shape} = \alpha, \text{scale} = \beta), \text{ viz. } f_X(x) = (\alpha/\beta)(x/\beta)^{\alpha-1} \exp\{-(x/\beta)^{\alpha}\}\mathbf{1}_{(0,\infty)}(x). \text{ Find the pdf of } f_X(x) = (\alpha/\beta)(x/\beta)^{\alpha-1} \exp\{-(x/\beta)^{\alpha}\}\mathbf{1}_{(0,\infty)}(x).$ $Y = \ln(X)$.

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Example Lec2.4

Let X have the pdf $f_X(x) = 2^{-1}\mathbf{1}_{(0,2)}(x)$. Find the pdf of $Y = X^2$.

Example Lec2.5

Let $f_X(x) = 3^{-1} \mathbf{1}_{(-1,2)}(x)$. Find the pdf of $Y = X^2$.

Multivariate Transformation

Multivariate distribution

- Random vector $\mathbf{X} = (X_1, \dots, X_n)$ with realization $\mathbf{x} = (x_1, \dots, x_n)$ $-\operatorname{cdf} F_{\mathbf{X}}(\boldsymbol{x}) = \Pr(X_1 \le x_1, \dots, X_n \le x_n)$
- Discrete
 - Joint pmf

$$p_{\mathbf{X}}(\mathbf{x}) = \Pr(X_1 = x_1, \dots, X_n = x_n)$$

- $-\operatorname{supp}(\mathbf{X}) = \operatorname{supp}(p_{\mathbf{X}}) = \{ \boldsymbol{x} \in \mathbb{R}^n : p_{\mathbf{X}}(\boldsymbol{x}) > 0 \}$
- Marginal pmf of (X_1, \ldots, X_k)

$$p_{X_1,...,X_k}(x_1,...,x_k) = \sum_{(x_{k+1},...,x_n) \in \mathbb{R}^{n-k}} p_{\mathbf{X}}(\mathbf{x})$$

- Continuous
 - Joint pdf

$$f_{\mathbf{X}}(\mathbf{x}) = (\partial^n/\partial x_1 \cdots \partial x_n) F_{\mathbf{X}}(\mathbf{x})$$

- * $\Pr(\mathbf{X} \in B) = \int_B f_{\mathbf{X}}(\boldsymbol{x}) d\boldsymbol{x}$ for each Borel set $B \subset \mathbb{R}^n$ $\operatorname{supp}(\mathbf{X}) = \operatorname{supp}(f_{\mathbf{X}}) = \{\boldsymbol{x} \in \mathbb{R}^n : f_{\mathbf{X}}(\boldsymbol{x}) > 0\}$
- Marginal pdf of (X_1, \ldots, X_k)
 - $* f_{X_1,\ldots,X_k}(x_1,\ldots,x_k) = \int_{\mathbb{R}^{n-k}} f_{\mathbf{X}}(\mathbf{x}) \mathrm{d}x_{k+1} \cdots \mathrm{d}x_n$

Find the joint pdf of random vector $\mathbf{Y} = g(\mathbf{X})$ by multivariate transformation (CB Sec. 4.3 & 4.6)

- Conditions
 - **X** and **Y** both of *n* dimensions
 - $-g(\cdot)=(g_1(\cdot),\ldots,g_n(\cdot)):\operatorname{supp}(\mathbf{X})\to\operatorname{supp}(\mathbf{Y})$ is one-to-one, i.e.,

*
$$\mathbf{y} = (y_1, \dots, y_n) = (q_1(\mathbf{x}), \dots, q_n(\mathbf{x})) = \mathbf{q}(\mathbf{x})$$

*
$$\mathbf{y} = (y_1, \dots, y_n) = (g_1(\mathbf{x}), \dots, g_n(\mathbf{x})) = \mathbf{g}(\mathbf{x})$$

* $\mathbf{x} = (x_1, \dots, x_n) = \mathbf{g}^{-1}(\mathbf{y}) = (g_1^{-1}(\mathbf{y}), \dots, g_n^{-1}(\mathbf{y}))$

- Jacobian matrices
 - Jacobian matrix of transformation g^{-1}

$$\mathbf{J}_{\boldsymbol{g}^{-1}} = \mathbf{J}_{\boldsymbol{g}^{-1}}(\boldsymbol{y}) = \begin{bmatrix} \frac{\partial g_i^{-1}(\boldsymbol{y})}{\partial y_j} \end{bmatrix}_{n \times n} = \begin{bmatrix} \frac{\partial g_1^{-1}(\boldsymbol{y})}{\partial y_1} & \cdots & \frac{\partial g_1^{-1}(\boldsymbol{y})}{\partial y_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial g_n^{-1}(\boldsymbol{y})}{\partial y_1} & \cdots & \frac{\partial g_n^{-1}(\boldsymbol{y})}{\partial y_n} \end{bmatrix}$$

- Jacobian matrix of transformation g

$$\mathbf{J}_{m{g}} = \mathbf{J}_{m{g}}(m{x}) = \begin{bmatrix} rac{\partial g_i(m{x})}{\partial x_j} \end{bmatrix}_{n imes n} = \begin{bmatrix} rac{\partial g_1(m{x})}{\partial x_1} & \dots & rac{\partial g_1(m{x})}{\partial x_n} \\ \vdots & \ddots & \vdots \\ rac{\partial g_n(m{x})}{\partial x_1} & \dots & rac{\partial g_n(m{x})}{\partial x_n} \end{bmatrix}$$

– Alternative way to reach $\mathbf{J}_{g^{-1}}(y)$: $\mathbf{J}_{g^{-1}}(y) = {\{\mathbf{J}_g(g^{-1}(y))\}}^{-1}$

• Then

$$f_{\mathbf{Y}}(\boldsymbol{y}) = f_{\mathbf{X}}\{g^{-1}(\boldsymbol{y})\}|\det\{\mathbf{J}_{\boldsymbol{g}^{-1}}(\boldsymbol{y})\}|\mathbf{1}_{\text{supp}(\mathbf{Y})}(\boldsymbol{y}).$$

- Never miss $\mathbf{1}_{\text{supp}(\mathbf{Y})}(\boldsymbol{y})$
- If g is NOT one-to-one, one may figure out the cdf of Y and then differentiate it.

Example Lec3.1

 X_1 and X_2 are iid from $\mathcal{N}(0,1)$. Find the joint pdf of $Y_1=(X_1+X_2)/\sqrt{2}$ and $Y_2=(X_1-X_2)/\sqrt{2}$ and show their independence.

Note: the sample mean and standard deviation are respectively $\bar{X}=(X_1+X_2)/2=Y_1/\sqrt{2}$ and S= $\sqrt{(X_1 - \bar{X})^2 + (X_2 - \bar{X})^2} = |Y_2|.$

Find the marginal pdf

- 1. Figure out the joint pdf first
- 2. Taking the Integral

Example Lec3.2

 X_1 and X_2 are iid from $\mathcal{N}(0,1)$. Find the pdf of $U=\sqrt{X_1^2+X_2^2}$.

Basics on square matrices

Eigen-decomposition

- **A** is a real $n \times n$ matrix
- Eigenvalues of **A**, say $\lambda_1 \geq \cdots \geq \lambda_n$: n roots of characteristic equation $\det(\lambda \mathbf{I}_n \mathbf{A}) = 0$
- The *i*th (Right) eigenvector v_i : $\mathbf{A}v_i = \lambda_i v_i$
- Eigen-decomposition: $\mathbf{A} = \mathbf{V}\Lambda\mathbf{V}^{-1}$
 - $-\mathbf{V} = [\mathbf{v}_1, \dots, \mathbf{v}_n]$ and $\Lambda = \operatorname{diag}(\lambda_1, \dots, \lambda_n)$ both $n \times n$ matrices - Specifically $\mathbf{V}^{-1} = \mathbf{V}^{\top}$ for symmetric \mathbf{A}
- Numerical implementation in R: eigen()
- Connection to determinant and trace
 - Determinant
 - * $\det \mathbf{A} = \prod_{i=1}^{n} \lambda_i$ * $\det(\mathbf{A}^{\top}) = \det \mathbf{A}$

 - $* \det(\mathbf{A}^{-1}) = (\det \mathbf{A})^{-1}$
 - * $\det(c\mathbf{A}) = c^n \det \mathbf{A}$ for $n \times n$ matrix \mathbf{A} and scalar c
 - * $\det(\mathbf{AB}) = (\det \mathbf{A})(\det \mathbf{B})$ for squared \mathbf{A} and \mathbf{B}
 - Trace
 - * $\operatorname{tr}(\mathbf{A}) = \sum_{i=1}^{n} \lambda_i$
 - * $tr(c\mathbf{A}) = ctr(\mathbf{A})$ for scalar c
 - * $tr(\mathbf{A} + \mathbf{B}) = tr(\mathbf{A}) + tr(\mathbf{B})$ for squared **A** and **B**
 - $* \operatorname{tr}(\mathbf{AB}) = \operatorname{tr}(\mathbf{BA})$

Square root of matrices

- $\mathbf{A}^{1/2} = \mathbf{V} \Lambda^{1/2} \mathbf{V}^{\top}$ if for semi-positive definite \mathbf{A} - Semi-positive/non-negative definite: symmetric **A** with eigenvalues all non-negative $-\Lambda^{1/2} = diag(\lambda_1^{1/2}, \dots, \lambda_n^{1/2})$ $-\mathbf{A}^{1/2}\mathbf{A}^{1/2} = \mathbf{A}$ • $\mathbf{A}^{-1/2} = \mathbf{V} \Lambda^{-1/2} \mathbf{V}^{\top}$ for positive definite \mathbf{A}
 - Positive definite: symmetric **A** with eigenvalues all positive
 - $-\Lambda^{-1/2} = \operatorname{diag}(\lambda_1^{-1/2}, \dots, \lambda_n^{-1/2})$ $-\mathbf{A}^{-1/2}\mathbf{A}^{-1/2} = \mathbf{A}^{-1} \text{ and } \mathbf{A}^{1/2}\mathbf{A}^{-1/2} = \mathbf{I}$

Multivariate normal (MVN) distribution

 $MVN(\mathbf{0}, \mathbf{I}_n)$

- Random p-vector $\mathbf{Z} = (Z_1, \dots, Z_p)^{\top} \sim \text{MVN}(\mathbf{0}, \mathbf{I}_p) \Leftrightarrow \text{iid } Z_1 \dots, Z_p \sim \mathcal{N}(0, 1).$
- pdf of $MVN(0, \mathbf{I}_p)$:

$$f_{\mathbf{Z}}(\boldsymbol{z}) = \prod_{i=1}^{p} (2\pi)^{-1/2} \exp(-z_i^2/2)$$
$$= (2\pi)^{-p/2} \exp(-\boldsymbol{z}^{\top} \boldsymbol{z}/2), \quad \boldsymbol{z} \in \mathbb{R}^p$$

 $MVN(\boldsymbol{\mu}, \boldsymbol{\Sigma})$

• pdf of MVN(μ, Σ), $\Sigma > 0$:

$$f_{\mathbf{X}}(\boldsymbol{x}) = (2\pi)^{-p/2} (\det \boldsymbol{\Sigma})^{-1/2} \exp\{-(\boldsymbol{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu})/2\}, \quad \boldsymbol{x} \in \mathbb{R}^p$$

- $\mathbf{X} \sim \text{MVN}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \Rightarrow \mathbf{A}\mathbf{X} + \boldsymbol{a} \sim \text{MVN}(\mathbf{A}\boldsymbol{\mu} + \boldsymbol{a}, \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^{\top})$ for arbitrary $\mathbf{A} \in \mathbb{R}^{q \times p}$ and $\boldsymbol{a} \in \mathbb{R}^q$
- $\mathbf{X} \sim \text{MVN}(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \Leftrightarrow \mathbf{X} = \boldsymbol{\Sigma}^{1/2} \mathbf{Z} + \boldsymbol{\mu} \text{ with } \mathbf{Z} = \boldsymbol{\Sigma}^{-1/2} (\mathbf{X} \boldsymbol{\mu}) \sim \text{MVN}(0, \mathbf{I}_n)$

Marginals of MVN

- Suppose p-vector $\mathbf{X} = (X_1, \dots, X_p)^{\top}$ and q-vector $\mathbf{Y} = (Y_1, \dots, Y_q)^{\top}$ are jointly normally distributed. Then, **X** and **Y** are independent $\Leftrightarrow \text{cov}(\mathbf{X}, \mathbf{Y}) = \mathbf{0}_{p \times q}$.
- If X is of MVN, then its all margins are still of MVN. The inverse proposition does NOT hold.
 - Cautionary example: Let Y = XZ, where $X \sim \mathcal{N}(0,1)$; Z is independent of X with $\Pr(Z=1) =$ Pr(Z=-1)=.5. X and Y both turn out to be of standard normal, but they are not jointly normal.

```
if (!("plot3D" %in% rownames(installed.packages())))
  install.packages("plot3D")
set.seed(1)
xsize = 1e4L
X = rnorm(xsize)
Z = rbinom(n = xsize, 1, .5)
Y = (2 * Z - 1) * X
# 3d histogram of (X, Y)
plot3D::hist3D(z=table(cut(X, 100), cut(Y, 100)), border = "black")
# plot the support of joint pdf of (X, Y)
plot3D::image2D(z=table(cut(X, 100), cut(Y, 100)), border = "black")
```