# PH 712 Probability and Statistical Inference

#### Part I: Random Variable

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## Probability (HMC Sec. 1.1–1.3)

- Sample space (denoted by  $\Omega$ ): the set of all the possible outcomes, e.g.,
  - $-\Omega = \mathbb{R}^+$  if investigating survival times of cancer patients
  - $-\Omega = \{yes, no\}$  if investigating whether a treatment is effective
- Event (denoted by capital Roman letters, e.g., A): a subset of the sample space, e.g., corresponding to the previous sample spaces,
  - (0, 10]: the survival time  $\leq 10$
  - {yes}: the treatment is effective
- Occurrence of event: the outcome is part of the event
- Probability (denoted by Pr): a function quantifying the occurrence likelihood of an event
  - E.g.,
    - \* Pr(A): the occurrence probability of event A
    - \*  $Pr(A^c)$ : the probability that event A does NOT occur  $(A^c = \Omega \setminus A \text{ denoting the complement set of } A)$
    - \*  $Pr(A \cup B)$ : the occurrence probability of either A or B
    - \*  $Pr(A \cap B)$ : the occurrence probability of both A and B
  - Input: an event
  - Output: a real number (the occurrence probability of the input event)
  - Requirements:
    - \*  $Pr(A) \ge 0$  for any event A
    - \*  $Pr(\Omega) = 1$  (i.e., the sample space as a special event always occurs)
    - \* (The probability of the union of mutually exclusive countably events is the sum of the probability of each event) If  $\{A_n\}_{n=1}^{\infty}$  is a sequence of events with  $A_{n_1} \cap A_{n_2} = \emptyset$  for all  $n_1 \neq n_2$ , then  $\Pr(\bigcup_{n=1}^{\infty} A_n) = \sum_{n=1}^{\infty} \Pr(A_n)$
  - More properties (deduced from the above requirements):
    - $* \operatorname{Pr}(A) = 1 \operatorname{Pr}(A^c)$
    - $* \Pr(\emptyset) = 0$
    - \* Pr(A) < Pr(B) if  $A \subset B$
    - \*  $0 \le \Pr(A) \le 1$  for each A
    - \*  $\lim_{n\to\infty} \Pr(A_n) = \Pr(\lim_{n\to\infty} A_n) = \Pr(\bigcup_{n=1}^{\infty} A_n)$  if  $\{A_n\}_{n=1}^{\infty}$  is nondecreasing (i.e.,  $A_1 \subset A_2 \subset \cdots$ )
    - \*  $\lim_{n\to\infty} \Pr(A_n) = \Pr(\lim_{n\to\infty} A_n) = \Pr(\bigcap_{n=1}^{\infty} A_n)$  if  $\{A_n\}_{n=1}^{\infty}$  is nonincreasing (i.e.,  $A_1 \supset A_2 \supset \cdots$ )
    - \*  $\Pr(A \cup B) = \Pr(A) + \Pr(B) \Pr(A \cap B)$  for any events A and B regardless if they are disjoint or not

\*  $\Pr(\bigcup_{n=1}^{\infty} A_n) \leq \sum_{n=1}^{\infty} \Pr(A_n)$  for arbitrary sequence  $\{A_n\}_{n=1}^{\infty}$ 

## Conditional probability and independence (HMC Sec. 1.4)

- Conditional probability of B given A (with Pr(A) > 0):  $Pr(B \mid A) = Pr(A \cap B) / Pr(A)$ 
  - Interpretation: the occurrence probability of B, given that A has already occurred.
  - Properties:
    - \*  $\Pr(B \mid A) \geq 0$
    - \*  $\Pr(A \mid A) = 1$

    - \*  $\Pr(\bigcup_{n=1}^{\infty} B_n \mid A) = \sum_{n=1}^{\infty} \Pr(B_n \mid A)$  if  $\{B_n\}_{n=1}^{\infty}$  are mutually exclusive \* (Law of total probability)  $\Pr(B) = \sum_{n=1}^{N} \Pr(A_n) \Pr(B \mid A_n)$  if  $\{A_n\}_{n=1}^{N}$  form a partition of  $\Omega$ (i.e.,  $\{A_n\}_{n=1}^{N}$  are mutually exclusive and  $\Omega = \bigcup_{n=1}^{N} A_n$ )
    - \* (Bayes' theorem)  $\Pr(A_i \mid B) = \Pr(A_i) \Pr(B \mid A_i) / \sum_{n=1}^{N} \Pr(A_n) \Pr(B \mid A_n)$  if  $\{A_n\}_{n=1}^{N}$  form a decomposition/partition of  $\Omega$
- Independence between two events B and A (i.e.,  $B \perp A$ ):  $\Pr(B \cap A) = \Pr(A) \Pr(B)$ 
  - $\Leftrightarrow B \perp A^c$
  - $\Leftrightarrow \Pr(B \mid A) = \Pr(B) \text{ (if } \Pr(A) \neq 0)$
- Mutual independence among N events  $A_1, \ldots, A_N$ : for arbitrary subset of  $\{A_1, \ldots, A_N\}$ , say  $\{A_{n_1}, \ldots, A_{n_K}\}$  with  $2 \le K \le N$ ,  $\Pr(\bigcap_{k=1}^K A_{n_k}) = \prod_{k=1}^K \Pr(A_{n_k})$

#### HMC Ex. 1.4.31

- A French writer, Chevalier de Méré, had asked a famous mathematician, Pascal, to explain why the following two probabilities were different (the difference had been noted from playing the game many times): (1) at least one six in four independent casts of a six-sided die; (2) at least a pair of sixes in 24 independent casts of a pair of dice. From proportions it seemed to Mr. de Méré that the two probabilities should be the same. Compute the probabilities of (1) and (2).
  - Hint: Pr(no six in one cast of a die) = 5/6, Pr(no six in one cast of a pair of dice) =  $(5/6)^2$ , and Pr(only one six in one cast of a pair of dice) =  $2 \times (1/6) \times (5/6)$ .

# Distribution of an RV (HMC Chp. 1.5–1.7)

- RV: a function encoding the entries of  $\Omega$ 
  - Input: arbitrary entry of  $\Omega$ , say  $\omega$
  - Output:  $X(\omega) \in \mathbb{R}$
- The cumulative distribution function (cdf) of RV X, say  $F_X$ , is defined as

$$F_X(t) = \Pr(X \le t), \quad t \in \mathbb{R}.$$

- $-\{X \leq t\}$ : short for the event  $\{\omega \in \Omega : X(\omega) \leq t\}$
- $F_X$  satisfies following three properties:
  - \* (Right continuous)  $\lim_{x \to t^+} F_X(x) = F_X(t)$  (p.s.,  $\lim_{x \to t^-} F_X(x) = \Pr(X < t)$ );
  - \* (Non-decreasing)  $F_X(t_1) \leq F_X(t_2)$  for  $t_1 \leq t_2$ ;
  - \* (Ranging from 0 to 1)  $F_X(-\infty) = 0$  and  $F_X(\infty) = 1$ .
- Reversely, a function satisfying the three above properties must be a cdf for certain RV.
  - \* Indicating an one-to-one correspondence between the set of all the RVs and the set of all the
- Knowing the distribution of an RV  $\Leftrightarrow$  knowing the cdf

# Example Lec1.1

• Given  $p \in (0,1)$ , suppose

$$F_X(x) = \begin{cases} 1 - (1-p)^{\lfloor x \rfloor}, & x \ge 1, \\ 0, & \text{otherwise,} \end{cases}$$

where |x| represents the integer part of real x.

- Show that  $F_X$  is a cdf. (Hint: Check all the three properties of cdf, especially the right-continuity of F at positive integers.)

# Distribution of an RV (con'd)

- Discrete RV
  - RV X merely takes countably different values
  - Probability mass function (pmf):  $p_X(t) = \Pr(X = t)$ 

    - \*  $F_X(t) = \sum_{x \le t} p_X(x)$ \*  $p_X(t) = F_X(t) \Pr(X < t) = F_X(t) \lim_{x \to t^-} F_X(x)$
  - Knowing the distribution of a discrete RV ⇔ knowing the pmf
  - Examples:
    - \* Bernoulli: a discrete RV with two possible outcomes, typically coded as 0 (failure) and 1 (success).
      - · https://en.wikipedia.org/wiki/Bernoulli distribution
    - \* Binomial: the number of successes in a fixed number of independent Bernoulli trials.
      - · https://en.wikipedia.org/wiki/Binomial\_distribution
      - · E.g., flipping a coin 10 times and counting the number of heads.
    - \* Negative binomial: the number of trials until a specified number of successes is achieved.
      - · https://en.wikipedia.org/wiki/Negative binomial distribution
      - · E.g., the number of coin flips until you get 3 heads.
    - \* Geometric: the number of trials until the first success in a series of independent Bernoulli trials.
      - · https://en.wikipedia.org/wiki/Geometric\_distribution
      - E.g., the number of coin flips needed until the first head appears.
    - \* Hypergeometric: the number of successes in a sample drawn without replacement from a finite population.
      - · https://en.wikipedia.org/wiki/Hypergeometric distribution
      - E.g., drawing a certain number of red balls from a bag containing both red and blue balls without replacement.
    - \* Poisson: the number of events that occur in a fixed interval of time or space, where events happen independently.
      - · https://en.wikipedia.org/wiki/Poisson\_distribution
      - · E.g., the number of emails you receive in an hour.
    - \* Uniform (the discrete version): each outcome in a finite set has an equal probability.
      - · https://en.wikipedia.org/wiki/Discrete\_uniform\_distribution
      - E.g., rolling a fair dice, where each of the six faces has an equal chance of landing.
- Continuous RV
  - RV X is continuous  $\Leftrightarrow$  its cdf  $F_X$  is absolutely continuous, i.e., there exists  $f_X$  such that

$$F_X(t) = \int_{-\infty}^t f_X(x) dx, \quad \forall t \in \mathbb{R}.$$

- \* Probability density function (pdf):  $f_X(t) = dF_X(t)/dt = \lim_{\delta \to 0^+} \Pr(t < X \le t + \delta)/\delta (\ge 0)$ . \*  $\int_{-\infty}^{\infty} f_X(x) dx = 1$  Knowing the distribution of a continuous RV  $\Leftrightarrow$  knowing the pdf
- Examples:
  - \* Uniform (the continuous version): all outcomes in a continuous range are equally likely.
    - · https://en.wikipedia.org/wiki/Uniform distribution (continuous)
  - \* Normal/Gaussian (denoted by  $\mathcal{N}(\mu, \sigma^2)$ ): the most important and widely used distributions, where data is symmetrically distributed around the mean.
    - · https://en.wikipedia.org/wiki/Normal\_distribution
  - \* Exponential: the time between events in a Poisson process, often used to describe waiting
    - https://en.wikipedia.org/wiki/Exponential distribution

- \* Chi-squared: sum of squared standard normal RVs; arising in hypothesis testing, particularly in tests of independence and goodness of fit.
  - · https://en.wikipedia.org/wiki/Chi-squared distribution
- \* Cauchy: known for its heavy tails and undefined mean and variance; used in robust statistics.
  - · https://en.wikipedia.org/wiki/Cauchy distribution
- \* Weibull: a generalization of the exponential distribution, used in reliability engineering and failure time analysis.
  - · https://en.wikipedia.org/wiki/Weibull distribution
- \* Log-normal:  $\exp(\mathcal{N}(0,1))$ ; commonly used to model stock prices and other financial data.
  - · https://en.wikipedia.org/wiki/Log-normal\_distribution
- \* (Student's) t: used in hypothesis testing, particularly for small sample sizes.
  - · https://en.wikipedia.org/wiki/Student%27s t-distribution

## Example Lec1.2

• Given  $p \in (0,1)$ , suppose

$$F_X(x) = \begin{cases} 1 - (1-p)^{\lfloor x \rfloor}, & x \ge 1, \\ 0, & \text{otherwise,} \end{cases}$$

where |x| represents the integer part of x.

- What is the type of X, discrete or continuous?

## Support of RV (CB pp. 50 & HMC pp. 46)

- For discrete RV X with pmf  $p_X$ 
  - $\text{ supp}(X) = \{ x \in \mathbb{R} : p_X(x) > 0 \}$
  - E.g., support of Binom(n, p) is  $\{0, \ldots, n\}$
  - $-\int_{\operatorname{supp}(X)} f_X(x) \mathrm{d}x = 1$
- For continuous RV X with pdf  $f_X$ 
  - $\text{ supp}(X) = \{ x \in \mathbb{R} : f_X(x) > 0 \}$
  - E.g., support of  $\mathcal{N}(0,1)$  is  $\mathbb{R}$
  - $-\sum_{x \in \text{supp}(X)} p_X(x) = 1$

#### Example Lec1.3

• Revisit  $F_X$  defined in Example Lec1.1, i.e.,

$$F_X(x) = \begin{cases} 1 - (1-p)^{\lfloor x \rfloor}, & x \ge 1, \\ 0, & \text{otherwise,} \end{cases}$$

where |x| represents the integer part of real x.

- What is the support of X?

## Expectations (HMC Sec. 1.8–1.9)

- Given RV X and function g, the expectation of g(X) is  $\mathrm{E}\{g(X)\}$ 
  - $\begin{array}{l} = \int_{-\infty}^{\infty} g(x) f_X(x) \mathrm{d}x \text{ for continuous } X \\ = \sum_{x \in \mathrm{supp}(X)} g(x) p_X(x) \text{ for discrete } X \\ \text{ Weighted average of values of } g(X) \end{array}$

  - $E\{a_1g_1(X) + a_2g_2(X)\} = a_1E\{g_1(X)\} + a_2E\{g_2(X)\}\$
- Mean of X (a.k.a. the 1st raw moment/moment about 0 of X): E(X)
- Variance of X (a.k.a. the 2nd central moment of X):  $Var(X) = E\{X E(X)\}^2$

$$- Var(X) = E(X^2) - {E(X)}^2$$

$$- \operatorname{Var}(aX + b) = a^2 \operatorname{Var}(X)$$

• Standard deviation of X: square root of the variance of X

## Example Lec1.4

- Find the mean and variance of  $X \sim \mathcal{N}(0,1)$ , i.e.,  $f_X(x) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{x^2}{2}\right)$
- Find the mean and variance of  $X \sim \mathcal{N}(\mu, \sigma^2)$  with  $\mu \in \mathbb{R}$  and  $\sigma \in \mathbb{R}^+$ , i.e.,  $f_X(x) =$  $\frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right) \text{ (p.s. } X \sim \mathcal{N}(\mu, \sigma^2) \Leftrightarrow Z = (X-\mu)/\sigma \sim \mathcal{N}(0, 1))$
- Find the mean and variance of Cauchy distribution, i.e.,  $f_X(x) = {\pi(1+x^2)}^{-1}, x \in \mathbb{R}$

# Distribution of an RV (con'd)

- Moment generating function (MGF, HMC Sec. 1.9/CB Sec. 2.3)
  - $M_X(t) = \mathbb{E}\{\exp(tX)\}\$ 
    - \* Continuous X:  $M_X(t) = \int_{-\infty}^{\infty} \exp(tx) f_X(x) dx$ \* Discrete X:  $M_X(t) = \sum_{u \in \text{supp}(X)} \exp(tx) p_X(x)$
  - The MGF of X is  $M_X(t)$ ,  $t \in A$ ,  $\Leftrightarrow M_X(t)$  is finite for t in a neighborhood of 0, say A; otherwise the MGF does NOT exist or is NOT well defined.
  - $-M_{aX+b}(t) = \exp(bt)M_X(at)$
  - Knowing the distribution of an RV ⇔ knowing the MGF (if any)
  - If MGF M(t) is well-defined, then the kth raw moment is the kth-order derivative of M(t) evaluated at 0, i.e.,  $E(X^k) = M^{(k)}(0)$
- Characteristic function (CF, optional)
  - $-\varphi_X(t) = \mathbb{E}\{\exp(itX)\}\$ 

    - \* Continuous X:  $\varphi_X(t) = \int_{-\infty}^{\infty} \exp(itu) f_X(u) du$ \* Discrete X:  $\varphi_X(t) = \sum_{u \in \text{supp}(X)}^{\infty} \exp(itu) p_X(u)$
  - Always well-defined
  - $-\varphi_{aX+b}(t) = \exp(bt)\varphi_X(at)$
  - Knowing the distribution of an RV  $\Leftrightarrow$  knowing the CF

#### Example Lec1.5

- Find the MGF of  $X \sim \mathcal{N}(\mu, \sigma^2)$  with  $\mu \in \mathbb{R}$  and  $\sigma \in \mathbb{R}^+$ , i.e.,  $f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$
- Find the MGF of Cauchy distribution, i.e.,  $f_X(x) = {\pi(1+x^2)}^{-1}, x \in \mathbb{R}$

#### Indicator function

Given a set A, the indicator function of A is

$$\mathbf{1}_{A}(x) = \begin{cases} 1, & x \in A, \\ 0, & \text{otherwise.} \end{cases}$$

#### Example Lec1.6

• Revisit  $F_X$  defined in Example Lec1.1, i.e.,

$$F_X(x) = \begin{cases} 1 - (1-p)^{\lfloor x \rfloor}, & x \ge 1, \\ 0, & \text{otherwise,} \end{cases}$$

where |x| represents the integer part of x.

- Please reformulate  $F_X$  with the indicator function of  $A = \{x : x \ge 1\}$ .