STAT 3690 Lecture 04

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Block/partitioned matrix

• A partition of matrix: Suppose A_{11} is of $p \times r$, A_{12} is of $p \times s$, A_{21} is of $q \times r$ and A_{22} is of $q \times s$. Make a new $(p+q) \times (r+s)$ -matrix by organizing A_{ij} 's in a 2 by 2 way:

$$\mathbf{A} = \begin{bmatrix} \mathbf{A}_{11} & \mathbf{A}_{12} \\ \mathbf{A}_{21} & \mathbf{A}_{22} \end{bmatrix}$$

e.g.,

$$\mathbf{A} = \begin{bmatrix} 1 & 0 & 2 \\ 0 & 1 & 3 \\ \hline 4 & 5 & 6 \end{bmatrix}$$

if

$$\mathbf{A}_{11} = \left[\begin{array}{cc} 1 & 0 \\ 0 & 1 \end{array} \right], \quad \mathbf{A}_{12} = \left[\begin{array}{c} 2 \\ 3 \end{array} \right], \quad \mathbf{A}_{21} = \left[\begin{array}{cc} 4 & 5 \end{array} \right], \quad \text{and} \quad \mathbf{A}_{22} = \left[\begin{array}{cc} 6 \end{array} \right].$$

- Operations with block matrices
 - Working with partitioned matrices just like ordinary matrices
 - Matrix addition: if dimensions of A_{ij} and B_{ij} are quite the same, then

$$\mathbf{A} + \mathbf{B} = \left[egin{array}{ccc} \mathbf{A}_{11} & \mathbf{A}_{12} \ \mathbf{A}_{21} & \mathbf{A}_{22} \end{array}
ight] + \left[egin{array}{ccc} \mathbf{B}_{11} & \mathbf{B}_{12} \ \mathbf{B}_{21} & \mathbf{B}_{22} \end{array}
ight] = \left[egin{array}{ccc} \mathbf{A}_{11} + \mathbf{B}_{11} & \mathbf{A}_{12} + \mathbf{B}_{12} \ \mathbf{A}_{21} + \mathbf{B}_{21} & \mathbf{A}_{22} + \mathbf{B}_{22} \end{array}
ight]$$

- Matrix multiplication: if $A_{ij}B_{jk}$ makes sense for each i, j, k, then

$$\mathbf{AB} = \left[\begin{array}{ccc} \mathbf{A}_{11} & \mathbf{A}_{12} \\ \mathbf{A}_{21} & \mathbf{A}_{22} \end{array} \right] \left[\begin{array}{ccc} \mathbf{B}_{11} & \mathbf{B}_{12} \\ \mathbf{B}_{21} & \mathbf{B}_{22} \end{array} \right] = \left[\begin{array}{ccc} \mathbf{A}_{11} \mathbf{B}_{11} + \mathbf{A}_{12} \mathbf{B}_{21} & \mathbf{A}_{11} \mathbf{B}_{12} + \mathbf{A}_{12} \mathbf{B}_{22} \\ \mathbf{A}_{21} \mathbf{B}_{11} + \mathbf{A}_{22} \mathbf{B}_{21} & \mathbf{A}_{21} \mathbf{B}_{12} + \mathbf{A}_{22} \mathbf{B}_{22} \end{array} \right]$$

- Inverse: if \mathbf{A} , \mathbf{A}_{11} and \mathbf{A}_{22} are all invertible, then

$$\mathbf{A}^{-1} = \left[\begin{array}{cc} \mathbf{A}_{11.2}^{-1} & -\mathbf{A}_{11.2}^{-1} \mathbf{A}_{12} \mathbf{A}_{22}^{-1} \\ -\mathbf{A}_{22}^{-1} \mathbf{A}_{21} \mathbf{A}_{11.2}^{-1} & \mathbf{A}_{22.1}^{-1} \end{array} \right]$$

- $* \mathbf{A}_{11.2} = \mathbf{A}_{11} \mathbf{A}_{12} \mathbf{A}_{22}^{-1} \mathbf{A}_{21}$ $* \mathbf{A}_{22.1} = \mathbf{A}_{22} \mathbf{A}_{21} \mathbf{A}_{11}^{-1} \mathbf{A}_{12}$
- Conditional mean vectors and covariance matrices: If $\mathbf{X} \sim (\boldsymbol{\mu}, \boldsymbol{\Sigma})$ and

$$\mathbf{X} = \left[\begin{array}{c} \mathbf{X}_1 \\ \mathbf{X}_2 \end{array} \right], \quad \boldsymbol{\mu} = \left[\begin{array}{c} \boldsymbol{\mu}_1 \\ \boldsymbol{\mu}_2 \end{array} \right] \quad \text{and} \quad \boldsymbol{\Sigma} = \left[\begin{array}{cc} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{array} \right] > 0,$$

where $E(\mathbf{X}_i) = \boldsymbol{\mu}_i$ and $cov(\mathbf{X}_i, \mathbf{X}_j) = \boldsymbol{\Sigma}_{ij}$, then

$$-\operatorname{E}(\mathbf{X}_{i} \mid \mathbf{X}_{j} = \boldsymbol{x}_{j}) = \boldsymbol{\mu}_{i} + \boldsymbol{\Sigma}_{ij} \boldsymbol{\Sigma}_{jj}^{-1} (\boldsymbol{x}_{j} - \boldsymbol{\mu}_{j}) \text{ for } i \neq j \text{ and } \boldsymbol{\Sigma}_{jj} > 0$$

$$-\operatorname{cov}(\mathbf{X}_{i} \mid \mathbf{X}_{j} = \boldsymbol{x}_{j}) = \boldsymbol{\Sigma}_{ii} - \boldsymbol{\Sigma}_{ij} \boldsymbol{\Sigma}_{jj}^{-1} \boldsymbol{\Sigma}_{ji} \text{ for } i \neq j \text{ and } \boldsymbol{\Sigma}_{jj} > 0$$

$$-\operatorname{cov}(\mathbf{X}_i \mid \mathbf{X}_j = \mathbf{x}_j) = \mathbf{\Sigma}_{ii} - \mathbf{\Sigma}_{ij} \mathbf{\Sigma}_{jj}^{-1} \mathbf{\Sigma}_{ji} \text{ for } i \neq j \text{ and } \mathbf{\Sigma}_{jj} > 0$$

Multivariate normal (MVN) distribution

• Standard normal random vector

$$-\mathbf{Z} = [Z_1, \dots, Z_p]^{\top} \sim MVN_p(\mathbf{0}, \mathbf{I}) \Leftrightarrow Z_1, \dots, Z_p \overset{\text{iid}}{\sim} N(0, 1) \Leftrightarrow$$

$$\phi_{\mathbf{Z}}(\mathbf{z}) = (2\pi)^{-p/2} \exp(-\mathbf{z}^{\mathsf{T}}\mathbf{z}/2), \quad \mathbf{z} = [z_1, \dots, z_p]^{\mathsf{T}} \in \mathbb{R}^p$$

- (General) normal random vector
 - Def: The distribution of **X** is MVN iff there exists $q \in \mathbb{Z}^+$, $\boldsymbol{\mu} \in \mathbb{R}^q$, $\mathbf{A} \in \mathbb{R}^{q \times p}$ and $\mathbf{Z} \sim MVN_p(\mathbf{0}, \mathbf{I})$ such that $\mathbf{X} = \mathbf{AZ} + \boldsymbol{\mu}$
 - * Limit the discussion to non-degenerate cases, i.e., $rk(\mathbf{A}) = q$
 - * $\mathbf{X} \sim MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, i.e.,

$$f_{\mathbf{X}}(\boldsymbol{x}) = \frac{1}{\sqrt{(2\pi)^p \mathrm{det}(\boldsymbol{\Sigma})}} \exp\{-(\boldsymbol{x} - \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\boldsymbol{x} - \boldsymbol{\mu})/2\}, \quad \boldsymbol{x} \in \mathbb{R}^p$$

$$\Sigma = \operatorname{var}(\mathbf{X}) = \mathbf{A}\mathbf{A}^{\top} > 0$$

- Exercise:
 - 1. $\Sigma = \mathbf{A}\mathbf{A}^{\top} > 0 \Leftrightarrow \operatorname{rk}(\mathbf{A}) = q \text{ (Hint: SVD of } \mathbf{A});$
 - 2. there exists a $p \times p$ positive definite matrix, say $\Sigma^{1/2}$, such that $\Sigma = \Sigma^{1/2}\Sigma^{1/2}$ and $\Sigma^{-1} = \Sigma^{-1/2}\Sigma^{-1/2}$ (Hint: spectral decomposition of Σ).
- Useful properties of MVN
 - $-\mathbf{X} \sim MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma}) \Leftrightarrow \mathbf{Z} = \boldsymbol{\Sigma}^{-1/2}(\mathbf{X} \boldsymbol{\mu}) \sim MVN_p(\mathbf{0}, \mathbf{I})$. So, we have a stochastic representation of arbitrary $\mathbf{X} \sim MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$: $\mathbf{X} = \boldsymbol{\Sigma}^{1/2}\mathbf{Z} + \boldsymbol{\mu}$, where $\mathbf{Z} \sim MVN_p(\mathbf{0}, \mathbf{I})$.
 - $-\mathbf{X} \sim MVN$ iff, for all $a \in \mathbb{R}^p$, $a^{\top}\mathbf{X}$ has a (univariate) normal distribution.
 - If $\mathbf{X} \sim MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$, then $\mathbf{A}\mathbf{X} + \boldsymbol{b} \sim MVN_q(\mathbf{A}\boldsymbol{\mu} + \boldsymbol{b}, \mathbf{A}\boldsymbol{\Sigma}\mathbf{A}^{\top})$ for $\mathbf{A} \in \mathbb{R}^{q \times p}$ and $\mathrm{rk}(\mathbf{A}) = q$.
- Exercise: Generate six iid samples following bivariate normal $MVN_2(\mu, \Sigma)$ with

$$\boldsymbol{\mu} = [3, 6]^{\top}, \quad \boldsymbol{\Sigma} = \left[\begin{array}{cc} 10 & 2 \\ 2 & 5 \end{array} \right].$$

- Exercise:
 - 1. Prove that $(\mathbf{X} \boldsymbol{\mu})^{\top} \boldsymbol{\Sigma}^{-1} (\mathbf{X} \boldsymbol{\mu}) \sim \chi^2(p)$ if $\mathbf{X} \sim MVN_p(\boldsymbol{\mu}, \boldsymbol{\Sigma})$.
 - 2. Suppose $X_1 \sim N(0,1)$ and $\mathbf{X} = [X_1, X_2]^{\top}$. Does \mathbf{X} follow an MVN in the following two cases? a. $X_2 = -X_1$;
 - b. $X_2 = (2Y 1)X_1$, where $Y \sim Ber(p)$ is independent of **X**.

Joint, marginal and conditional MVN

• If $\mathbf{X} \sim MVN_n(\boldsymbol{\mu}, \boldsymbol{\Sigma})$ and

$$\mathbf{X} = \left[\begin{array}{c|c} \mathbf{X}_1 \\ \hline \mathbf{X}_2 \end{array} \right], \quad \boldsymbol{\mu} = \left[\begin{array}{c|c} \boldsymbol{\mu}_1 \\ \hline \boldsymbol{\mu}_2 \end{array} \right] \quad \text{and} \quad \boldsymbol{\Sigma} = \left[\begin{array}{c|c} \boldsymbol{\Sigma}_{11} & \boldsymbol{\Sigma}_{12} \\ \hline \boldsymbol{\Sigma}_{21} & \boldsymbol{\Sigma}_{22} \end{array} \right]$$

with $\Sigma_{11} > 0$ and $\Sigma_{22} > 0$, then

 $-\mathbf{X}_i \sim MVN_{p_i}(\boldsymbol{\mu}_i, \boldsymbol{\Sigma}_{ii})$, i.e., marginals of MVN are MVN.

$$\begin{split} & - \mathbf{X}_i \mid \mathbf{X}_j = \boldsymbol{x}_j \sim MVN_{p_i}(\boldsymbol{\mu}_{i|j}, \boldsymbol{\Sigma}_{i|j}), \text{ i.e., conditionals of MVN ar MVN.} \\ & * \boldsymbol{\mu}_{i|j} = \boldsymbol{\mu}_i + \boldsymbol{\Sigma}_{ij}\boldsymbol{\Sigma}_{jj}^{-1}(\boldsymbol{x}_j - \boldsymbol{\mu}_j) \\ & * \boldsymbol{\Sigma}_{i|j} = \boldsymbol{\Sigma}_{ii} - \boldsymbol{\Sigma}_{ij}\boldsymbol{\Sigma}_{jj}^{-1}\boldsymbol{\Sigma}_{ji} \\ & - \mathbf{X}_i \perp \mathbf{X}_j \Leftrightarrow \boldsymbol{\Sigma}_{ij} = 0 \end{split}$$