STAT 3690 Lecture 17

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Multivariate linear regression

- Interested in the relationship between random q-vector $[Y_1, \dots, Y_p]^{\top}$ and random q-vector $[X_1, \dots, X_q]^{\top}$
- Model
 - Population version: $[Y_1,\ldots,Y_p]^{\top} \mid X_1,\ldots,X_q \sim (\mathbf{B}^{\top}[1,X_1,\ldots,X_q]^{\top},\sigma^2)$, where $\mathbf{B} =$ $[\beta_{kj}]_{(q+1)\times p}$, i.e.,

 - * $\mathrm{E}([Y_1,\ldots,Y_p]^{\top}\mid X_1,\ldots,X_q)=\mathbf{B}^{\top}[1,X_1,\ldots,X_q]^{\top}$ * $\mathrm{cov}([Y_1,\ldots,Y_p]^{\top}\mid X_1,\ldots,X_q)=\mathbf{\Sigma}>0,$ i.e., the conditional covariance of $[Y_1,\ldots,Y_p]^{\top}$ does not depend on X_1, \ldots, X_q
 - Sample version

$$\frac{\mathbf{Y}}{n \times p} = \frac{\mathbf{X}}{n \times (q+1)} \frac{\mathbf{B}}{(q+1) \times p} + \frac{\mathbf{E}}{n \times p}$$

- * $\mathbf{Y} = [Y_{ij}]_{n \times p}$
- * Design matrix

$$\mathbf{X} = \begin{bmatrix} 1 & X_{11} & \cdots & X_{q1} \\ \vdots & \vdots & \ddots & \vdots \\ 1 & X_{n1} & \cdots & X_{nq} \end{bmatrix}_{n \times (q+1)}$$

- $rk(\mathbf{X}) = q + 1$
- * $\mathbf{E} = [\mathbf{E}_1, \dots, \mathbf{E}_n]^{\top}$, where \mathbf{E}_i^{\top} is the *i*th row of \mathbf{E}
- $\ast\,$ Assume the independence across i, i.e.,
 - $\cdot [Y_{i1}, \dots, Y_{ip}, X_{i1}, \dots, X_{iq}]^{\top} \stackrel{\text{iid}}{\sim} [Y_1, \dots, Y_p, X_1, \dots, X_q]^{\top}$
 - $\mathbf{E}_{i\cdot} \overset{ ext{iid}}{\sim} (\mathbf{0}_p, \mathbf{\Sigma})$
- Relationship with univariate linear regression
 - If Σ is diagonal, the multivariate model reduces to $\mathbf{Y}_{\cdot j} = \mathbf{X}\mathbf{B}_{\cdot j} + \mathbf{E}_{\cdot j}, \ j = 1, \dots, p$
 - * $\mathbf{Y}_{.j}$: the jth column of \mathbf{Y}
 - * \mathbf{B}_{i} : the *j*th column of \mathbf{B}

 - $\begin{array}{ll} * \ \mathbf{E}_{.j} \sim (\mathbf{0}_n^{\cdot}, \sigma_{jj}^2 \mathbf{I}_n) \\ & \cdot \ \sigma_{jj}^2 \colon (j,j) \text{-entry of } \mathbf{\Sigma} \end{array}$
- Relationship with MANOVA
 - MANOVA models can be expressed as multivariate linear regression with carefully selected dummy (explanatory) variables.

Exercise: translate the following 1-way MANOVA model

$$\mathbf{Y}_{ij} = \boldsymbol{\mu} + \boldsymbol{\tau}_i + \mathbf{E}_{ij}, \quad i = 1, \dots, m, \quad j = 1, \dots, n_i$$

into a multivariate linear regression model, where $\mathbf{E}_{ij} \stackrel{\text{iid}}{\sim} MVN_p(\mathbf{0}, \boldsymbol{\Sigma})$ and $\sum_i \boldsymbol{\tau}_i = 0$.

- Least squares (LS) estimation (no need of normality)
 - $-\hat{\mathbf{B}}_{\mathrm{LS}} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{Y}$
 - $* E(\hat{\mathbf{B}}_{LS}) = \mathbf{B}$

*
$$\mathrm{E}(\mathbf{B}_{\mathrm{LS}}) = \mathbf{B}$$

- $\hat{\mathbf{\Sigma}}_{\mathrm{LS}} = (n-q-1)^{-1}(\mathbf{Y} - \mathbf{X}\hat{\boldsymbol{\beta}})^{\top}(\mathbf{Y} - \mathbf{X}\hat{\boldsymbol{\beta}}) = (n-q-1)^{-1}\mathbf{Y}^{\top}(\mathbf{I} - \mathbf{H})\mathbf{Y}$
* $\mathrm{E}(\hat{\mathbf{\Sigma}}_{\mathrm{LS}}) = \mathbf{\Sigma}$

- Maximum likelihood (ML) estimation (in need of (conditional) normality)

 - $-\hat{\mathbf{B}}_{\mathrm{ML}} = (\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{Y} = \hat{\mathbf{B}}_{\mathrm{LS}}$ $-\hat{\mathbf{\Sigma}}_{\mathrm{ML}} = n^{-1}\mathbf{Y}(\mathbf{I} \mathbf{H})\mathbf{Y} = n^{-1}(n q 1)\hat{\mathbf{\Sigma}}_{\mathrm{LS}}$
 - * Given \mathbf{X} , $n\hat{\boldsymbol{\Sigma}}_{\mathrm{ML}} \sim W_n(\boldsymbol{\Sigma}, n-q-1)$
- Inference (in need of (conditional) normality)
 - Inference on $\mathbf{B}^{\top} \boldsymbol{a}$, given $\boldsymbol{a} \in \mathbb{R}^{q+1}$

 - * Estimator $\hat{\mathbf{B}}_{\mathrm{ML}}^{\top} \boldsymbol{a}$ * 100(1 α)% confidence interval for $\mathbf{B}^{\top} \boldsymbol{a}$

$$\left\{ \boldsymbol{u} \in \mathbb{R}^p : (\boldsymbol{u} - \widehat{\mathbf{B}}_{\mathrm{ML}}^{\top} \boldsymbol{a})^{\top} \widehat{\boldsymbol{\Sigma}}_{\mathrm{LS}}^{-1} (\boldsymbol{u} - \widehat{\mathbf{B}}_{\mathrm{ML}}^{\top} \boldsymbol{a}) \leq \boldsymbol{a}^{\top} (\mathbf{X}^{\top} \mathbf{X})^{-1} \boldsymbol{a} \left\{ \frac{p(n-q-1)}{n-q-p} F_{1-\alpha,p,n-p-q} \right\} \right\}$$

- Inference on $Y_0 = \mathbf{X}_0^{\top} \boldsymbol{\beta} + \varepsilon_0$ with a new observation vector given $\mathbf{X}_0 = [1, X_{01}, \dots, X_{0q}]^{\top} \in \mathbb{R}^{q+1}$

 - * Prediction $\hat{Y}_0 = \mathbf{X}_0^{\top} \hat{\boldsymbol{\beta}}_{\mathrm{ML}}$ * $100(1-\alpha)\%$ prediction interval for Y_0

$$\left\{ \boldsymbol{u} \in \mathbb{R}^p : (\boldsymbol{u} - \widehat{\mathbf{B}}_{\mathrm{ML}}^{\top} \boldsymbol{a})^{\top} \widehat{\boldsymbol{\Sigma}}_{\mathrm{LS}}^{-1} (\boldsymbol{u} - \widehat{\mathbf{B}}_{\mathrm{ML}}^{\top} \boldsymbol{a}) \leq \{1 + \boldsymbol{a}^{\top} (\mathbf{X}^{\top} \mathbf{X})^{-1} \boldsymbol{a}\} \left\{ \frac{p(n-q-1)}{n-q-p} F_{1-\alpha,p,n-p-q} \right\} \right\}$$