

# PH 712 Probability and Statistical Inference

## Part IX: Confidence Set/Interval

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### Confidence set

- Called a confidence interval (CI) If the set is indeed an interval
- $(1 - \alpha) \times 100\%$  confidence set for  $\theta$ , say  $C(X_1, \dots, X_n)$ :  $C(X_1, \dots, X_n)$  covers  $\theta$  with probability  $\geq (1 - \alpha) \times 100\%$ , i.e.,  $\Pr\{\theta \in C(X_1, \dots, X_n)\} \geq (1 - \alpha) \times 100\%$ 
  - $C(X_1, \dots, X_n)$  is a set defined on sample  $X_1, \dots, X_n$  and hence is randomized, while  $\theta$  is fixed

### Construction of a confidence set by inverting a level $\alpha$ two-sided test

- Implementation
  1. Construct the rejection region  $R(\theta_0)$  for a level  $\alpha$  test for hypotheses  $H_0 : \theta = \theta_0$  vs.  $H_1 : \theta \neq \theta_0$ .
  2.  $C(x_1, \dots, x_n) = \{\theta_0 : (x_1, \dots, x_n) \in \text{supp}(X_1, \dots, X_n)/R(\theta_0)\}$ ,
    - $\text{supp}(X_1, \dots, X_n)/R(\theta_0)$ : the complementary set of  $R(\theta_0)$ .
- If  $C(X_1, \dots, X_n)$  is constructed by inverting a level  $\alpha$  a level  $\alpha$  test, then
  - $C(X_1, \dots, X_n)$  does NOT cover  $\theta_0 \Leftrightarrow$  reject  $H_0 : \theta = \theta_0$  (vs.  $H_1 : \theta \neq \theta_0$ ) at level  $\alpha$
- Special cases:
  - $(1 - \alpha) \times 100\%$  asymptotic LRT confidence set for  $\theta$ :  $\{\theta_0 : -2(\ell(\theta_0) - \ell(\hat{\theta}_{\text{ML}})) < \chi^2_{1,1-\alpha}\}$
  - $(1 - \alpha) \times 100\%$  Wald confidence set for  $\theta$ :  $\{\theta_0 : |\hat{\theta}_{\text{ML}} - \theta_0| / \sqrt{\widehat{\text{var}}(\hat{\theta}_{\text{ML}})} < \Phi^{-1}_{1-\alpha/2}\}$ , i.e.,
$$\left( \hat{\theta}_{\text{ML}} - \Phi^{-1}_{1-\alpha/2} \sqrt{\widehat{\text{var}}(\hat{\theta}_{\text{ML}})}, \quad \hat{\theta}_{\text{ML}} + \Phi^{-1}_{1-\alpha/2} \sqrt{\widehat{\text{var}}(\hat{\theta}_{\text{ML}})} \right)$$

### Construction of a confidence set via (nonparametric) bootstrap

- Implementation
  1. Suppose you observe  $x_1, \dots, x_n$  for an iid sample of size  $n$ .
  2. Set a seed to make your result reproducible.
  3. For  $b$  in  $1 : B$ , do steps a–b.
    - a. Generate a bootstrap sample  $x_1^{(b)}, \dots, x_n^{(b)}$  by drawing a sample of size  $n$  with replacement from  $\{x_1, \dots, x_n\}$ .
    - b. Generate a estimate  $\hat{\theta}^{(b)}$  from  $x_1^{(b)}, \dots, x_n^{(b)}$ .
  4.  $(q_{\alpha/2}, q_{1-\alpha/2})$  is the  $(1 - \alpha)$  bootstrap confidence interval for  $\theta$ , where  $q_{\alpha/2}$  and  $q_{1-\alpha/2}$  are  $\alpha/2$  and  $(1 - \alpha/2)$  sample quantiles of  $\{\hat{\theta}^{(1)}, \dots, \hat{\theta}^{(B)}\}$ , respectively.

### Example Lec9.1

- Assume  $X_1, \dots, X_{10} \stackrel{\text{iid}}{\sim} f(x | p) = p^x(1-p)^{1-x}$ ,  $x = 0, 1$ ,  $0 < p < 1$ , and observe  $0, 0, 0, 1, 1, 1, 1, 1, 1, 1$ . Construct a 95% confidence set for  $p$ .

```

options(digits = 4)
alpha = .05
xs = c(rep(0,3), rep(1,7))
n = length(xs)
ell = function(p, xs){
  sum(xs*log(p) + (1-xs)*log(1-p))
}
# Asymptotic LRT CI
ml = optim(
  par = 0, lower = 0.000001, upper = .999999,
  fn = ell, xs = xs,
  method="L-BFGS-B",
  control=list(fnscale=-1), hessian=TRUE
)
ml$value # log-likelihood at MLE
qchisq(1-alpha, df=1)
# Wald CI
fish_info = -ml$hessian
var_p_hat = 1/fish_info
ci_wald = c(
  p_hat - qnorm(1-alpha/2)*sqrt(var_p_hat),
  p_hat + qnorm(1-alpha/2)*sqrt(var_p_hat)
); ci_wald
# Bootstrap CI
set.seed(712)
B = 1e4L
p_hat_bs = numeric(B)
for (b in 1:B) {
  xbs = sample(xs, size=n, replace=TRUE)
  p_hat_bs[b] = optim(
    par = 0, lower = 0.000001, upper = .999999,
    fn = ell, xs = xbs,
    method="L-BFGS-B",
    control=list(fnscale=-1))$par
}
ci_boot = quantile(p_hat_bs, probs = c(alpha/2, 1-alpha/2)); ci_boot

```