

# STAT 3100 Lecture Note

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## Asymptotic properties of MLE (con'd)

### Consistency (or consistence, CB Sec 10.1.1)

- $T_n = T_n(X_1, \dots, X_n)$  is consistent for  $\theta$  iff  $T_n \xrightarrow{p} \theta$  as  $n \rightarrow \infty$ 
  - A sufficient condition for consistency:  $E(T_n | \theta) \rightarrow \theta$  and  $\text{var}(T_n | \theta) \rightarrow 0$  as  $n \rightarrow \infty$

### CB Example 5.5.3

- Suppose that iid  $X_1, \dots, X_n \sim \mathcal{N}(\mu, \sigma^2)$ . Prove that
  - $S_n^2 = (n-1)^{-1} \sum_i (X_i - \bar{X}_n)^2$  is consistent for  $\sigma^2$ ;
  - $\widehat{\sigma}_{\text{ML}}^2 = n^{-1} \sum_i (X_i - \bar{X}_n)^2$  is consistent for  $\sigma^2$  too.

to  $\theta$ ; - Violated by, e.g.,  $\text{Unif}(0, \theta)$ ; -  $\theta_0$  is an interior point of parameter space  $\Theta$ .

### Example of inconsistent MLE

There are independent  $X_{i1}, X_{i2} \sim \mathcal{N}(\mu_i, \sigma^2)$ ,  $i = 1, \dots, n$ . Then  $\widehat{\sigma}_{\text{ML}}^2$  is NOT consistent for  $\sigma^2$ .

### Examples of consistent MLE with the regularity conditions violated

- iid  $X_1, \dots, X_n \sim \text{Ber}(1)$
- iid  $X_1, \dots, X_n \sim \text{Unif}(0, \theta)$

### Efficiency

- (HMC Def 6.2.2) For an estimator, say  $T_n$ , unbiased for  $\tau(\theta)$ , the (finite-sample) *efficiency* of  $T_n$  is the ratio of the CRLB to  $\text{var}(T_n)$ , i.e.,  $[\{\tau'(\theta)\}^2 / I_n(\theta)] / \text{var}(T_n | \theta)$ .
  - The higher efficiency the better;
  - the efficiency = 1  $\iff$  an efficient estimator.
- (CB Def 10.1.9) If  $k_n\{T_n - \tau(\theta)\} \xrightarrow{d} \mathcal{N}(0, \sigma^2)$ , then  $\sigma^2$  is the *asymptotic variance* of  $T_n$ .
- (CB Def 10.1.11)  $T_n$  is *asymptotically efficient* for  $\tau(\theta)$   $\iff \sqrt{n}\{T_n - \tau(\theta)\} \xrightarrow{d} \mathcal{N}(0, \{\tau'(\theta)\}^2 / I_1(\theta))$ , where

$I_1(\theta) = -E \left\{ \frac{\partial^2}{\partial \theta^2} \ln f(X_i | \theta) \mid \theta \right\} = -E\{H(\theta; X_i) \mid \theta\}$  is the Fisher information of one single observation.

- i.e., the asymptotic variance of  $T_n$  is  $\{\tau'(\theta)\}^2 / I_1(\theta)$ , attaining the CRLB

- (CB Def 10.1.16 & HMC Def 6.2.3(c)) Denote by  $T_n$  and  $W_n$  two estimators for  $\tau(\theta)$ . Suppose that  $\sqrt{n}\{T_n - \tau(\theta)\} \xrightarrow{d} \mathcal{N}(0, \sigma_T^2)$  and  $\sqrt{n}\{W_n - \tau(\theta)\} \xrightarrow{d} \mathcal{N}(0, \sigma_W^2)$ . The *asymptotic relative efficiency* (ARE) of  $T_n$  with respect to  $W_n$  is defined as

$$\text{ARE}(T_n, W_n) = \sigma_W^2 / \sigma_T^2.$$

- $T_n$  is asymptotically more efficient than  $W_n \iff \text{ARE}(T_n, W_n) > 1$
- $T_n$  is asymptotically efficient  $\iff \{\tau'(\theta)\}^2 / \{I_1(\theta)\sigma_T^2\} = 1$

### CB Example 10.1.17 & Ex. 10.9

- iid  $X_1, \dots, X_n \sim p(x | \lambda) = \lambda^x \exp(-\lambda) / x!$ ,  $x \in \mathbb{Z}^+$ ,  $\lambda > 0$ . To estimate  $\Pr(X_i = 0) = \exp(-\lambda)$ .
  - a. Consider  $T_n = n^{-1} \sum_i \mathbf{1}_{\{0\}}(X_i)$  and MLE  $W_n = \exp(-\bar{X}_n)$ . Compute  $\text{ARE}(T_n, W_n)$ , the ARE of  $T_n$  with respect to  $W_n$ .
  - b. Find the UMVUE for  $\Pr(X_i = 0)$ , say  $U_n$ , and then calculate  $\text{ARE}(U_n, W_n)$ .
    - Hint:  $\sqrt{n}(U_n - W_n) \xrightarrow{P} 0$  (derived from S. Portnoy, *The Annals of Statistics*, 1977, Vol. 5, pp. 522–529, Theorem 1) and  $\sum_{i=1}^n X_i \sim \text{Poisson}(n\lambda)$

### Asymptotic efficiency of MLE (CB Thm 10.1.12 & Ex. 10.7)

- $\sqrt{n}\{\tau(\hat{\theta}_{\text{ML}}) - \tau(\theta_0)\} \xrightarrow{d} \mathcal{N}(0, \{\tau'(\theta_0)\}^2 / I_1(\theta_0))$ , provided that  $\hat{\theta}_{\text{ML}}$  is the MLE for  $\theta_0$ ,  $\tau$  is differentiable and we have the previous four regularity conditions (for the consistency of MLE) plus the following two more (CB Sec 10.6.2):
  - For each  $x \in \text{supp}(X)$ ,  $f(x | \theta)$  is three times continuously differentiable with respect to  $\theta$ ; and  $\int f(x | \theta) dx$  can be differentiated three times under the integral sign;
  - for each  $\theta \in \Theta$ , there exists  $c(\theta) > 0$  and  $M(x, \theta)$  such that  $|\frac{\partial^3}{\partial \theta^3} \ln f_X(x | \theta)| \leq M(x, \theta)$  for all  $x \in \text{supp}(X)$  and  $\theta \in (\theta - c(\theta), \theta + c(\theta))$ .
- In practice,
  - $nI_1(\theta_0) = I_n(\theta_0) \approx I_n(\hat{\theta}_{\text{ML}}) \approx \hat{I}_n(\hat{\theta}_{\text{ML}})$ 
    - \* (Expected) Fisher information (number)  $I_n(\theta_0) = -\mathbb{E}\{H(\theta_0; \mathbf{X})\}$
    - \* Observed Fisher information (number)  $\hat{I}_n(\hat{\theta}_{\text{ML}}) = -\frac{\partial^2}{\partial \theta^2} \ln L(\theta; \mathbf{x})|_{\theta=\hat{\theta}_{\text{ML}}} = -H(\hat{\theta}_{\text{ML}}; \mathbf{x})$
  - Hence  $\text{var}\{\tau(\hat{\theta}_{\text{ML}})\} \approx \{\tau'(\theta_0)\}^2 / I_n(\theta_0) \approx \{\tau'(\hat{\theta}_{\text{ML}})\}^2 / I_n(\hat{\theta}_{\text{ML}}) \approx \{\tau'(\hat{\theta}_{\text{ML}})\}^2 / \hat{I}_n(\hat{\theta}_{\text{ML}})$

### Take-home exercises (NOT to be submitted; to be potentially covered in labs)

- CB Ex. 10.3, 10.17(a-c)