STAT 4100 Lecture Note

Week Twelve (Nov 28, 30 & Dec 2, 2022)

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Approximation to the variance of $\hat{\theta}_n$

- Why?
  - Reflect the variation or dispersion of $\hat{\theta}_n$
  - Help approximate the distribution of $\hat{\theta}_n$ (and further construct the confidence region for θ) if assuming normality
- How?
  - Utilizing the asymptotic variance of $\hat{\theta}_n$
  - Resampling methods, e.g., bootstraping

CB Example 10.1.17 & Ex. 10.9 (con'd)

- iid $X_1, \ldots, X_n \sim p(x \mid \lambda) = \lambda^x \exp(-\lambda)/x!$, $x \in \mathbb{Z}^+$, $\lambda > 0$. Define $\theta = \Pr(X_i = 2 \mid \theta) = \lambda^2 \exp(-\lambda)/2$. Approximate the variance of $\hat{\theta}_{\mathrm{ML}} = \bar{X}_n^2 \exp(-\bar{X}_n)/2$ by delta methods.

CB Example 10.1.15

- Holding iid $X_i \sim \text{Bernoulli}(p)$, the variance of Bernoulli(p) is $\tau(p) = p(1-p)$ whose MLE is $\tau(\hat{p}_{\mathrm{mle}}) = \bar{X}_n(1-\bar{X}_n)$. Approximate $\operatorname{var}\{\tau(\hat{p}_{\mathrm{mle}})\}$ by delta methods.

Bootstraping the variance of $\hat{\theta}_n$ (CB Sec. 10.1.4)

- Nonparametric bootstrap:
  1. For j in $1:B$, do steps 2–3.
  2. Draw the jth resample $\boldsymbol{x}_j^*$ of size n from the original sample $\boldsymbol{x} = \{x_1, \ldots, x_n\}$, with replacement.
  3. Let $\hat{\theta}_j^* = \hat{\theta}(\boldsymbol{x}_j^*)$.
  4. $\operatorname{var}(\hat{\theta}) \approx$ the sample variance of $\{\hat{\theta}_1^*, \ldots, \hat{\theta}_B^*\}$.
- (Optional, see, e.g., www.stat.columbia.edu/~bodhi/Talks/Emp-Proc-Lecture-Notes.pdf) Empirical process: theoretical foundation for nonparametric bootstrap
  - (Glivenko-Cantelli) $\sup_{x\in\mathbb{R}} |F_n(x) - F(x)| \xrightarrow{\text{a.s.}} 0$
  - (Donsker) $\sqrt{n}(F_n - F) \xrightarrow{d} \mathrm{BB} \circ F$, i.e., $\mathrm{E}[g\{\sqrt{n}(F_n - F)\}] \to \mathrm{E}[g(\mathrm{BB} \circ F)]$ for all bounded, continuous and real-valued g
    - BB is a standard Brownian bridge process on $[0,1]$, i.e.,
      - $\mathrm{BB}(0) = \mathrm{BB}(1) = 0$ but $\mathrm{BB}(t) \sim \mathcal{N}(0, t(1-t))$ for $t \in (0,1)$;
      - fixing $t_1, \ldots, t_p \in (0,1)$, $[\mathrm{BB}(t_1), \ldots, \mathrm{BB}(t_p)]^\top$ is of multivariate normal with $\operatorname{cov}(\mathrm{BB}(s), \mathrm{BB}(t)) = \min(s,t) - st$;

    - $\cdot$ $\mathrm{BB}(t)$ is continuous in t.
- Parametric bootstrap:
  1. For j in $1:B$, do steps 2–3.
  2. Draw the jth resample $\boldsymbol{x}_j^*$ of size n from a fitted model $f(x \mid \hat{\theta})$.
  3. Let $\hat{\theta}_j^* = \hat{\theta}(\boldsymbol{x}_j^*)$.
  4. $\mathrm{var}(\hat{\theta}) \approx$ the sample variance of $\{\hat{\theta}_1^*, \ldots, \hat{\theta}_B^*\}$.

CB Example 10.1.15

- Holding iid $X_i \sim \mathrm{Bernoulli}(p)$, the variance of $\mathrm{Bernoulli}(p)$ is $\tau(p) = p(1-p)$ for which the MLE is $\tau(\hat{p}_{\mathrm{mle}}) = \bar{X}_n(1-\bar{X}_n)$. Approximate $\mathrm{var}\{\tau(\hat{p}_{\mathrm{mle}})\}$ by the bootstrap.

Large-sample hypothesis testing

Recall the LRT

- $H_0 : \boldsymbol{\theta} \in \boldsymbol{\Theta}_0$ v.s. $H_1 : \boldsymbol{\theta} \in \boldsymbol{\Theta}_1$, where $\boldsymbol{\Theta} = \boldsymbol{\Theta}_0 \cup \boldsymbol{\Theta}_1$
- LRT statistic
$$\lambda(\boldsymbol{x}) = \frac{\sup_{\boldsymbol{\theta} \in \boldsymbol{\Theta}_0} L(\boldsymbol{\theta}; \boldsymbol{x})}{\sup_{\boldsymbol{\theta} \in \boldsymbol{\Theta}} L(\boldsymbol{\theta}; \boldsymbol{x})} = \frac{L(\hat{\boldsymbol{\theta}}_{0,\mathrm{ML}}; \boldsymbol{x})}{L(\hat{\boldsymbol{\theta}}_{\mathrm{ML}}; \boldsymbol{x})}$$
  - $\hat{\boldsymbol{\theta}}_{0,\mathrm{ML}}$: constrained MLE for $\boldsymbol{\theta} \in \boldsymbol{\Theta}_0$
  - $\hat{\boldsymbol{\theta}}_{\mathrm{ML}}$: unconstrained MLE for $\boldsymbol{\theta} \in \boldsymbol{\Theta}$
- $\{\boldsymbol{x} : \lambda(\boldsymbol{x}) \le c_\alpha\}$: rejection region of level α LRT
  - c_α is such defined that $\sup_{\boldsymbol{\theta} \in \boldsymbol{\Theta}_0} \Pr(\lambda(\mathbf{X}) \le c_\alpha \mid \boldsymbol{\theta}) = \alpha$

Asymptotic distribution of LRT statistic (CB Thm 10.3.1 & 10.3.3)

- As $n \to \infty$,
$$-2 \ln \lambda(\mathbf{X}) \xrightarrow{d} \chi^2(\nu),$$
  where $\nu =$ difference of numbers of free parameters in $\boldsymbol{\Theta}_0$ and $\boldsymbol{\Theta}$.
- (CB Thm 10.3.3) $\{\boldsymbol{x} : -2 \ln \lambda(\boldsymbol{x}) \ge \chi^2_{\nu, 1-\alpha}\}$: asymptotic rejection region of level α LRT
  - $\chi^2_{\nu,1-\alpha}$ is the $1-\alpha$ quantile of $\chi^2(\nu)$.

CB Example 10.3.4

- iid $X_1, \ldots, X_n \sim \Pr(X_i = j) = p_j$, $j = 1, \ldots, 5$. Specify the $1-\alpha$ LRT rejection region for $H_0 : p_1 = p_2 = p_3$ and $p_4 = p_5$ versus H_1 : Otherwise.

Wald test (CB pp. 493)

- $H_0 : \theta = \theta_0$ versus $H_1 : \theta \ne \theta_0$
  - Wald statistic: $(\hat{\theta}_{\mathrm{ML}} - \theta_0)\sqrt{I_n(\theta_0)}$ ($\xrightarrow{d} \mathcal{N}(0,1)$ under H_0 as $n \to \infty$)
  - Level α rejection region: $\{\boldsymbol{x} : |\hat{\theta}_{\mathrm{ML}} - \theta_0|\sqrt{I_n(\theta_0)} \ge \Phi^{-1}_{1-\alpha/2}\}$
- Asymptotically equivalent to LRT Under H_0, but merely involve the (unconstrained) MLE.

Score test (CB pp. 494)

- Merely involve the constrained MLE
- $H_0 : \theta = \theta_0$ vs. $H_1 : \theta \neq \theta_0$
  - Score statistic: $S(\theta_0; \mathbf{X})/\sqrt{I_n(\theta_0)}$ ($\xrightarrow{d} \mathcal{N}(0,1)$ under H_0 as $n \to \infty$)
  - Level α rejection region: $\{\boldsymbol{x} : |S(\theta_0; \boldsymbol{x})|/\sqrt{I_n(\theta_0)} \geq \Phi^{-1}_{1-\alpha/2}\}$.
- If $\boldsymbol{\Theta}_0$ contains more than one points, then substitute $\hat{\theta}_{0,\mathrm{ML}}$ for θ_0.

CB Examples 10.3.5 & 10.3.6

- iid $X_1, \ldots, X_n \sim \text{Bernoulli}(p)$, $p \in (0,1)$. Derive LRT, Wald and score tests for $H_0 : p = p_0$ versus $H_1 : p \neq p_0$.

Asymptotic confidence regions

- Constructed by reverting rejection regions
- Examples
  - $1-\alpha$ LRT confidence region for θ: $\{\theta : -2\ln\{L(\theta; \boldsymbol{x})/L(\hat{\theta}_{\mathrm{ML}}; \boldsymbol{x})\} < \chi^2_{1,1-\alpha}\}$
  - $1-\alpha$ Wald confidence region for θ: $\{\theta : |\hat{\theta}_{\mathrm{ML}} - \theta|\sqrt{I_n(\theta)} < \Phi^{-1}_{1-\alpha/2}\}$
  - $1-\alpha$ score confidence region for θ: $\{\theta : |S(\theta; \boldsymbol{x})|/\sqrt{I_n(\theta)} < \Phi^{-1}_{1-\alpha/2}\}$

CB Examples 10.4.2, 10.4.3 & 10.4.5

- iid $X_1, \ldots, X_n \sim \text{Bernoulli}(p)$, construct $1-\alpha$ confidence intervals for p.

Take-home exercises (NOT to be submitted; to be potentially covered in labs)

- CB Ex. 10.17(a–c), 10.36, 10.38
- HMC Ex. 6.3.16–6.3.18