

BRIGHAM YOUNG UNIVERSITY

Applied Bayesian Statistics

STAT 451

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January 8, 2013

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1 Class Introduction

Office Hours Monday

Keep a copy of the code

Preliminary assignment email now.

2 Bayesian Methodology

Parameterize state of knowledge

Estimate parameters

$$f(\underline{\theta}|x) = \frac{L(x|\underline{\theta}) \prod(\underline{\theta})}{\int \int_{\underline{\theta}} L(x|\underline{\theta}) \prod(\underline{\theta}) d\underline{\theta}} \propto L(x|\underline{\theta}) \prod(\underline{\theta})$$

2.1 Metropolis Sampler

$$\begin{aligned} & \pi \left[\frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} \exp\left(-\frac{x}{\beta}\right) \right] + (\pi - 1) \left[\frac{1}{\Gamma(\alpha)\beta^\alpha} x^{\alpha-1} \exp\left(-\frac{x}{\beta}\right) \right] \\ & \quad \frac{1}{4} \left[\frac{1}{\Gamma(2)3^2} x^{2-1} \exp\left(-\frac{x}{3}\right) \right] + \frac{3}{4} \left[\frac{1}{\Gamma(4)3^4} x^{4-1} \exp\left(-\frac{x}{3}\right) \right] \\ & \quad k = 4\Gamma(4)3^3 \end{aligned}$$

$$g(x) = k \cdot f(x) = 18x \exp\left(-\frac{x}{3}\right) + x^3 \exp\left(-\frac{x}{3}\right)$$

2.1.1 Metropolis Sampler Steps

- Initial $x = x_0$
- for $i = 1$ to n
 - generate candidate c
 - $\frac{g(c)}{g(x_{i-1})} = r$
 - let $x_i = c$ with $P = \min(1, r)$

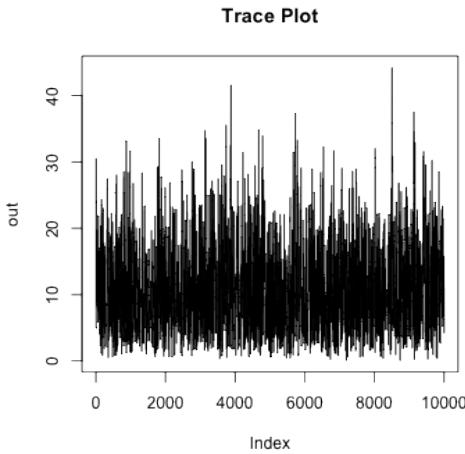
2.1.2 Metropolis Code

```
# Define g function
g <- function(x) {
  18 * x * exp(-x/3) + x^3 * exp(-x/3)
}

# Write Sampler
out <- NULL
out[1] <- 5
candidate.sigma <- 20
counter <- 0
samples <- 10^4
for (i in 2:samples) {
  out[i] <- out[i - 1]
  candidate <- rnorm(1, out[i - 1], candidate.sigma)
  if (candidate > 0) {
    r <- g(candidate)/g(out[i - 1])
    if (r > runif(1, 0, 1)) {
      out[i] <- candidate
      counter <- counter + 1
    }
  }
}
message(counter/samples) # should be about 24-40

## 0.3285

plot(out, type = "l", main = "Trace Plot")
```



Trace Plot looks good

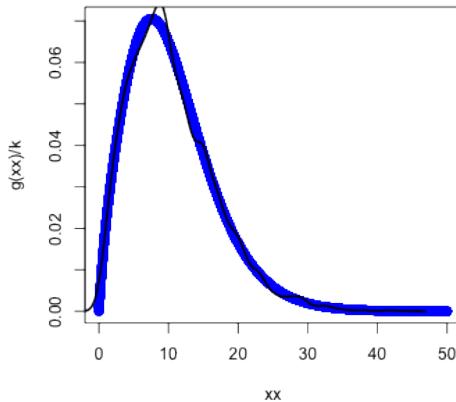
```
xx <- seq(0, 50, length = 1e+05)
k <- 4 * gamma(4) * 3^3
# Plot the exact distribution
plot(xx, g(xx)/k, lwd = 2, col = "blue")
# curve(.25*dgamma(x, shape=2, scale=3)+.75*dgamma(x, shape=4, scale=3),
# add=TRUE, col='blue') Both of these should produce the same line

lines(density(out), lwd = 2)

# Exact Mean
0.25 * (2 * 3) + 0.75 * (4 * 3)
## [1] 10.5

# Approximate Mean
```

```
mean(out)  
## [1] 10.35
```



* * January 10, 2013 *

Paul Sabin — Stat 451 TA
Tuesday 2:30-4:00
Friday: 9:30-11:00
Office: 233
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```
# For Winbugs  
install.packages("ARM", "R2WinBUGS")  
  
# For jags  
install.packages("R2jags")
```

3 Two Sample t test

$$y \sim \mathcal{N}(\mu, \sigma^2) \quad i = 1, 2$$

$$f(\mu_1) = \mathcal{N}(\mu = 0, \sigma^2 = 1000) f(\sigma^2) = \text{Unif}(0, 5000)$$

PROC MCMC two group t test

```
1 data examp2;
  infile dataset;
  input tmt response;
  run;
5
/*
nmc: Number of Marcov Chains
nbi: Number Burn in
```

```

10 dic: information criteria
*/
proc mcmc data=examp2 outpost=post2 seed=1234 nmc=100000 nbi=5000
            statistics (alpha=.05)=(summary interval) thin=10
            moniter=( parms mudif varratio ) propcov=quannew
            diagnostics=(all) dic ;
15 array mu[2] mu1-mu2;
array sig2 [2] sig21-sig22;
parms: mu: sig2 :;
prior mu: ~ normal(0, sd=10000);
20 prior sig21 ~ unif(9, 5000); /* gamma(shape=30, scale=50) */
prior sig22 ~ unif(9, 5000);
mudif = mu1 - mu2;
varratio = sig21/sig22;
mm = mu[tmt];
25 vv = sig2[tmt];
model response ~ normal(mm, var=vv);
run;

proc export data=post2 outfile='twogroups.csv' dbms=csv replace;
30 run;

```

HPD — Highest Posterior Density the smallest interval

Fischer variance problem — a two sample t test with unequal variances

Check the posterior autocorrelations

* * * January 15, 2013 *

```

library(rjags)
library(R2jags)

```

JAGS uses precisions, not variances. Precisions are $\frac{1}{\sigma^2}$

```

1 library(R2jags)
mdl <- "
  model {
    for (i in 1:33) {
      y[i] ~ dnorm(mu[tmt[i]], prec[tmt[i]])
    }

    for (i in 1:2) {
      mu[i] ~ dnorm(0, 0.000001)
      vr[i] ~ dunif(0, 5000)
      prec[i] <- 1/vr[i]
    }
  }
",
15 mdl <- "
  model {
    for (i in 1:33) {
      y[i] ~ dnorm(mu[tmt[i]], prec[tmt[i]])
    }
"
20

```

```

    for (i in 1:2) {
      mu[i] ~ dnorm(100,0.01)
      vr[i] ~ dgamma(20,.05)
      prec[i] <- 1/vr[i]
5   }
}
,
mdl <- "
model {
  for (i in 1:33) {
    y[i] ~ dnorm(mu[tmt[i]], prec[tmt[i]])
  }

  for (i in 1:2) {
    mu[i] ~ dnorm(100,0.01)
    vr[i] ~ dgamma(10,200)
    prec[i] <- 1/vr[i]
  }
}
"
40 groups <- read.table("../data/01twogroups.dat", col.names=c("tmt", "y"))
writeLines(mdl, "twogroups.jags")
tmt <- groups$tmt
y <- groups$y
45 # Data to go into jags
data.jags <- c('tmt','y')
# what parameters to keep track of
parms <- c('mu','vr')
# Initial Values
50 innts <- function() {list('mu' = rnorm(2,125,5), 'vr' = runif(2,0,5000))}
twogroups.sim <- jags(data=data.jags, parameters.to.save=parms, inits=innts,
  model.file="twogroups.jags",
  n.iter=11000, n.burnin=1000, n.chains=1, n.thin=1)

twogroups.sim
55 sims <- as.mcmc(twogroups.sim)
plot(sims)

plot(sims[,2], type='l')

```

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January 17, 2013

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You need to think about priors

4 Diagnostics

```

library(R2jags)
mdl <- "
model {
  for (i in 1:33) {
    y[i] ~ dnorm(mu[tmt[i]], prec[tmt[i]])
  }
}

```

```

}

for (i in 1:2) {
  mu[i] ~ dnorm(0,0.000001)
  vr[i] ~ dunif(0,5000)
  prec[i] <- 1/vr[i]
}
}

groups <- read.table("data/01twogroups.dat", col.names=c("tmt", "y"))
writeLines(mdl, "code/twogroups.jags")
tmt <- groups$tmt
y <- groups$y
# Data to go into jags
data.jags <- c('tmt', 'y')
# what parameters to keep track of
parms <- c('mu', 'vr')
# Initial Values
innts <- function() {list('mu' = rnorm(2,125,5), 'vr' = runif(2,0,5000))}
twogroups.sim <- jags(data=data.jags, parameters.to.save=parms, inits=innts, model.file="code/twogroups.jags",
n.iter=6000, n.burnin=1000, n.chains=1, n.thin=1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 76
##
## Initializing model

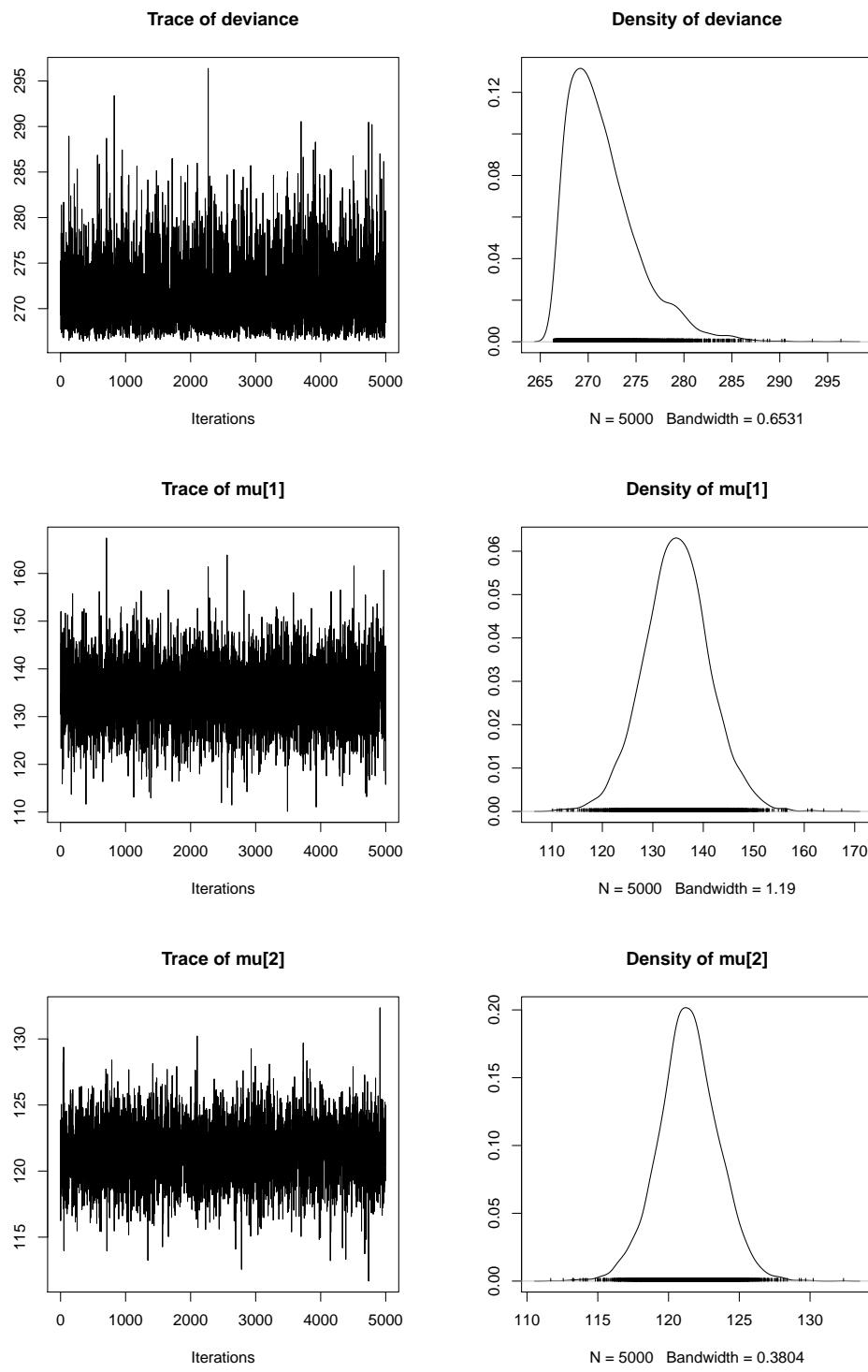
twogroups.sim

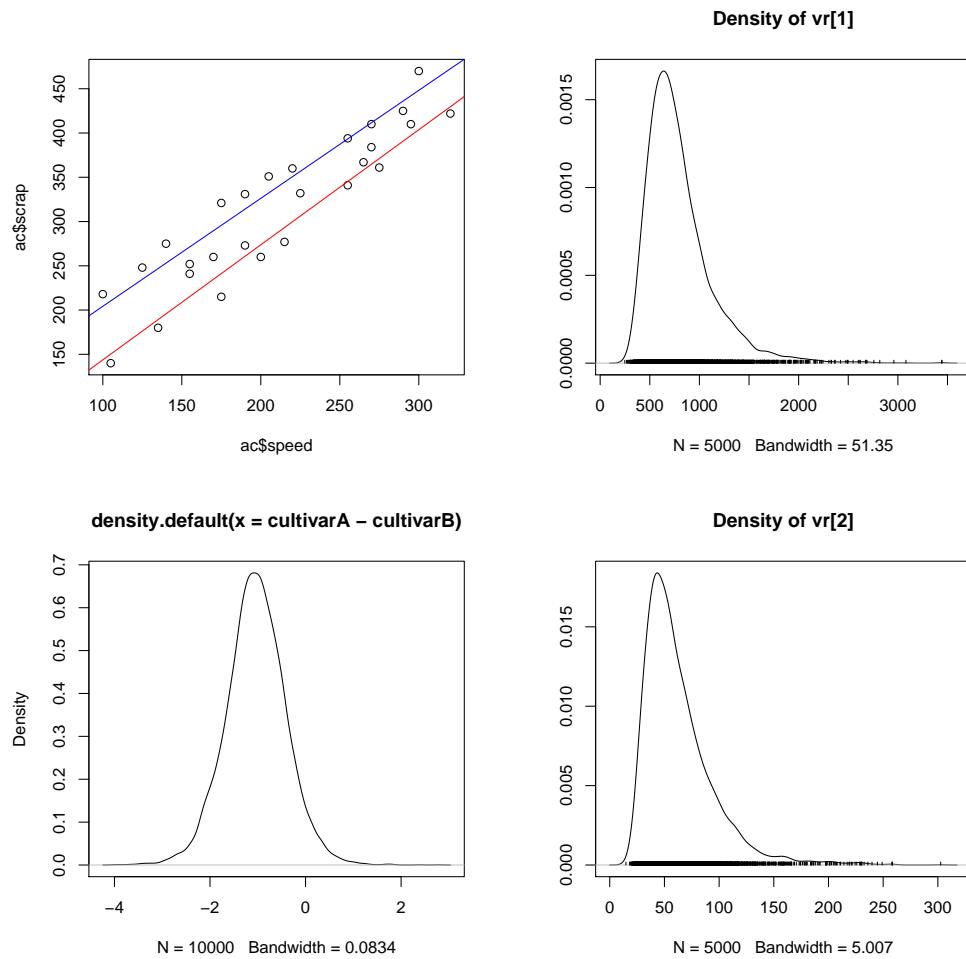
## Inference for Bugs model at "code/twogroups.jags", fit using jags,
## 1 chains, each with 6000 iterations (first 1000 discarded)
## n.sims = 5000 iterations saved
##      mu.vect sd.vect 2.5%   25%   50%   75% 97.5%
## mu[1]    134.75  6.499 122.11 130.61 134.70 138.87 147.8
## mu[2]    121.40  2.110 117.16 120.09 121.38 122.73 125.6
## vr[1]    797.45 333.228 381.36 573.92 723.20 930.48 1659.5
## vr[2]     63.12 31.741 26.22 41.34 55.22 76.11 146.2
## deviance 271.55  3.652 266.96 268.84 270.74 273.37 280.5
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 6.7 and DIC = 278.2
## DIC is an estimate of expected predictive error (lower deviance is better).

sims <- as.mcmc(twogroups.sim)
plot(sims, auto.layout=FALSE, ask=FALSE)

# plot(sims[,2], type='l')

```





First diagnostic: Look at your trace plot

Second diagnostic: look at auto corilation

```
autocorr(sims)
```

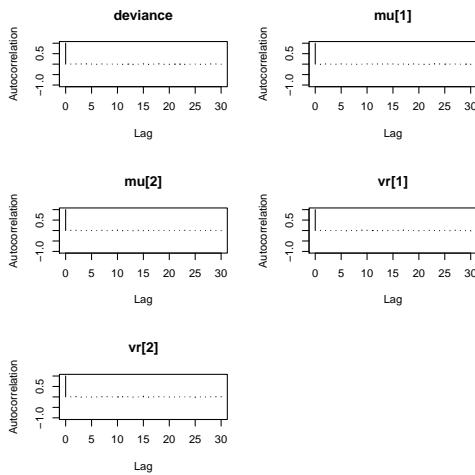
```
## , , deviance
##
##      deviance     mu[1]     mu[2]      vr[1]      vr[2]
## Lag 0  1.000000  0.0397216 -0.025591  0.5294713  0.5567647
## Lag 1  0.009287  0.0178228 -0.001470 -0.0079258  0.0045190
## Lag 5  0.008346  0.0222125 -0.004006  0.0258781  0.0005049
## Lag 10 -0.003025  0.0082836 -0.023008  0.0244412 -0.0188622
## Lag 50 -0.004687 -0.0007478 -0.009660 -0.0006706  0.0165176
##
## , , mu[1]
##
##      deviance     mu[1]     mu[2]      vr[1]      vr[2]
## Lag 0  0.039722  1.0000000 -0.019048  0.0242092  0.013122
## Lag 1  0.010297 -0.0176032 -0.018843 -0.0115282 -0.001713
## Lag 5  0.009260  0.0203083 -0.010197  0.0290402 -0.001748
## Lag 10 0.002206  0.0183513  0.048567 -0.0005999  0.015041
## Lag 50 0.001726 -0.0001966  0.002256  0.0190355 -0.014390
##
## , , mu[2]
##
##      deviance     mu[1]     mu[2]      vr[1]      vr[2]
## Lag 0 -0.025591 -0.0190482  1.000000 -0.002025 -0.017025
```

```

## Lag 1 -0.007725 -0.0083044 0.016805 -0.021768 0.005684
## Lag 5  0.030497 -0.0280916 -0.008886 0.010412 0.008702
## Lag 10 0.015048 -0.0139730 0.026550 0.019316 -0.003001
## Lag 50 0.010831 0.0004535 -0.001931 -0.006375 0.012813
##
## , , vr[1]
##
##      deviance    mu[1]    mu[2]    vr[1]    vr[2]
## Lag 0  0.529471 0.024209 -0.002025 1.000000 0.006924
## Lag 1  0.012432 0.008644 0.003187 0.013754 0.002647
## Lag 5  -0.006325 -0.007802 -0.013119 -0.001014 0.001037
## Lag 10 -0.001276 -0.022398 -0.016429 0.018451 -0.021251
## Lag 50  0.020901 -0.010891 -0.008754 0.011124 0.031264
##
## , , vr[2]
##
##      deviance    mu[1]    mu[2]    vr[1]    vr[2]
## Lag 0  0.556765 0.013122 -0.01702 0.006924 1.0000000
## Lag 1  -0.002083 0.002571 -0.00108 0.004606 0.0085805
## Lag 5  0.007769 0.018526 -0.01204 0.033278 -0.0131159
## Lag 10 0.020122 0.035722 -0.01813 0.019459 -0.0003303
## Lag 50 -0.019064 0.023849 -0.01956 0.001463 -0.0069104

autocorr.plot(sims)

```



A good autocorrelation plot shows the first bar being tall, but the rest being relatively small.

Fourth diagnostic: Effective Sample size

```
effectiveSize(sims)
```

```

## deviance    mu[1]    mu[2]    vr[1]    vr[2]
##      5000     5000     5000     5000     4840

```

these should be relatively close to the actual sample size

fifth diagnostic: Raftery Lewis Diagnostic

```
raftery.diag(sims)
```

```

##
## Quantile (q) = 0.025
## Accuracy (r) = +/- 0.005
## Probability (s) = 0.95
##
##      Burn-in Total Lower bound Dependence
##      (M)      (N)  (Nmin)   factor (I)

```

```
## deviance 2      3741  3746      0.999
## mu[1]    2      3741  3746      0.999
## mu[2]    2      3930  3746      1.050
## vr[1]    2      3803  3746      1.020
## vr[2]    2      3561  3746      0.951
```

Check if you are within .005 (in the tails) with high probability

Dependence Factor should be less than 5

These five are the most commonly used, but there are others

```
geweke.diag(sims)

##
## Fraction in 1st window = 0.1
## Fraction in 2nd window = 0.5
##
## deviance   mu[1]   mu[2]   vr[1]   vr[2]
## -0.6146 -1.7786 -0.8627  0.7567  0.1023
```

Tests for drift in the samples. does something like a t test on the first 10% and the last 50% of the chain.

Heidel

```
heidel.diag(sims)

##
##           Stationarity start      p-value
##           test          iteration
## deviance passed      1      0.822
## mu[1]    passed      1      0.140
## mu[2]    passed      1      0.398
## vr[1]    passed      1      0.684
## vr[2]    passed      1      0.636
##
##           Halfwidth Mean  Halfwidth
##           test
## deviance passed     271.5 0.1012
## mu[1]    passed     134.7 0.1802
## mu[2]    passed     121.4 0.0585
## vr[1]    passed     797.5 9.2366
## vr[2]    passed      63.1 0.8942
```

test of stationarity, or convergence

both the stationarity and half width mean test should pass.

Gelman needs chains started at different points

* * January 22, 2013 *

5 Posterior Predictives

Likelihood $g_{i1} \sim \mathcal{N}(\mu_1, \sigma_1^2)$, $g_{i2} \sim \mathcal{N}(\mu_2, \sigma_2^2)$

Get a new y by plugging the sampled values of the parameters into a random sampler for the likelihood

```
1 ynew_1 <- rnorm(1 ,mu1[1] ,sigma1[1])
```

```

 $\mu_1 \quad \sigma_1^2 \quad \mu_2 \quad \sigma_2^2 \quad y_{\text{new1}}$ 
 $\vdots \quad \vdots \quad \vdots \quad \vdots \quad \vdots$ 

ratios <- read.table("../Homework/01/monkeyratio.dat")[,1]
library(R2jags)
N <- length(ratios)

mdl <- "
model {
  for (i in 1:N) {
    ratios[i] ~ dbeta(alpha, beta)
  }

  alpha ~ dgamma(5, .025)
  beta ~ dgamma(5, .025)
}
"
writeLines(mdl, "code/hw1.jags")

jags.data <- c("ratios", "N")
jags.params <- c("alpha", "beta")
jags.inits <- function() {
  list('alpha' <- rgamma(5,.025), 'beta' <- rgamma(5,.025))
}
set.seed(5487)

ratios.sim <- jags(data=jags.data,
                     parameters.to.save=jags.params,
                     # inits=jags.inits,
                     model.file="code/hw1.jags",
                     n.iter=6000, n.burnin=1000, n.chains=1, n.thin=1)

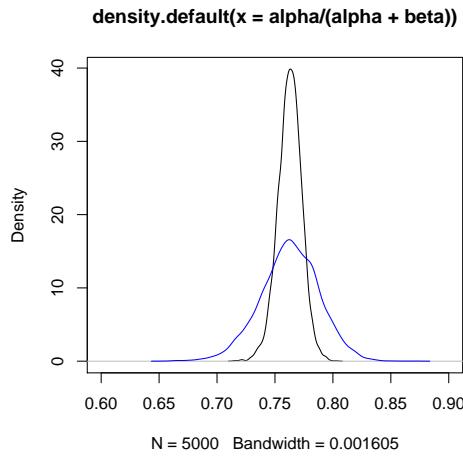
## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 10
##
## Initializing model

sims <- as.mcmc(ratios.sim)

# autocorr.plot(sims)
alpha <- sims[,1]
beta <- sims[,2]
ynew <- rbeta(10000,alpha,beta)

# Plot of the density of the mean and the posterior predictive
plot(density(alpha/(alpha+beta)), xlim=c(.6,.9))
lines(density(ynew), col='blue') # Plot of the posterior predictive
```



We don't make up data we get it from "Pedagogically flexible data acquisition."—Dr Tolly.
You can use other priors, they just may not be as easy for example:

6 Linear Regression

Likelihood

$$y_i \sim \mathcal{N}(\mu_i, \sigma^2)$$

The mean is different for every data point, but the variance is constant.

$$\begin{aligned}\mu_1 &\leftarrow \beta_0 + \beta_1 x \\ \beta_0 &\sim \mathcal{N}(, \tau = .001) \\ \beta_1 &\sim \mathcal{N}(\mu = 0, \tau = .001) \\ \sigma^2 &\sim\end{aligned}$$

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January 24, 2013

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```
jags.modelfile <- "code/linreg.jags"
mdl <- "
model {for (i in 1:N) {y[i] ~ dnorm(mu[i], prec)
mu[i] <- beta_0 + beta_1 * x[i]}
prec ~ invgamma(0.001)
beta_0 ~ normal(0, 0.001)
beta_1 ~ normal(0, 0.001)
sigma ~ uniform(0, 1000)
}
writeLines(mdl, "code/linreg.jags")

linregdata <- read.table("data/02linreg.dat", col.names = c("lotsize", "manhours"))
N <- nrow(linregdata)
y <- linregdata[, 2]
x <- linregdata[, 1]
jags.data <- c("y", "x", "N")
jags.params <- c("beta_0", "beta_1", "sigma")

set.seed(5487)
linreg.sim <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/linreg.jags",
n.iter = 10000, n.burnin = 5000, n.chains = 1, n.thin = 1)
```

```
## module glm loaded
## Error:
## Error parsing model file:
## syntax error on line 7 near """"

linreg.sim

## Error: object 'linreg.sim' not found
```

Let's check for convergence

```
sims <- as.mcmc(linreg.sim)

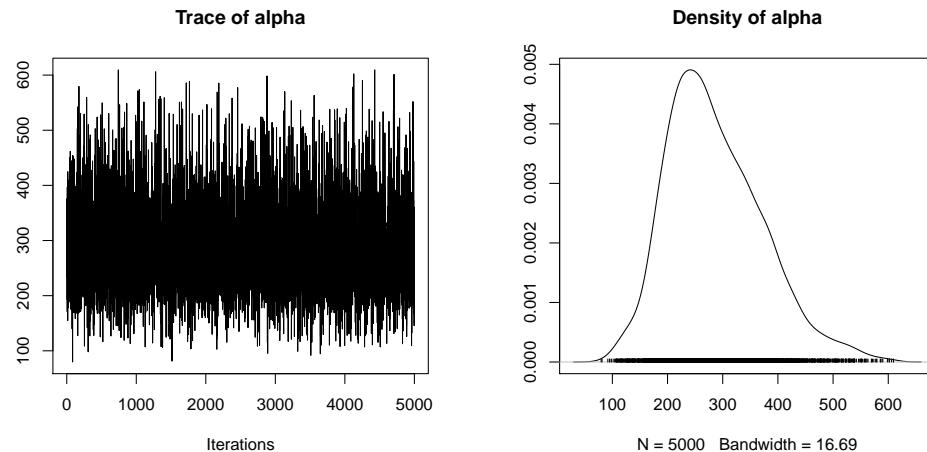
## Error: object 'linreg.sim' not found

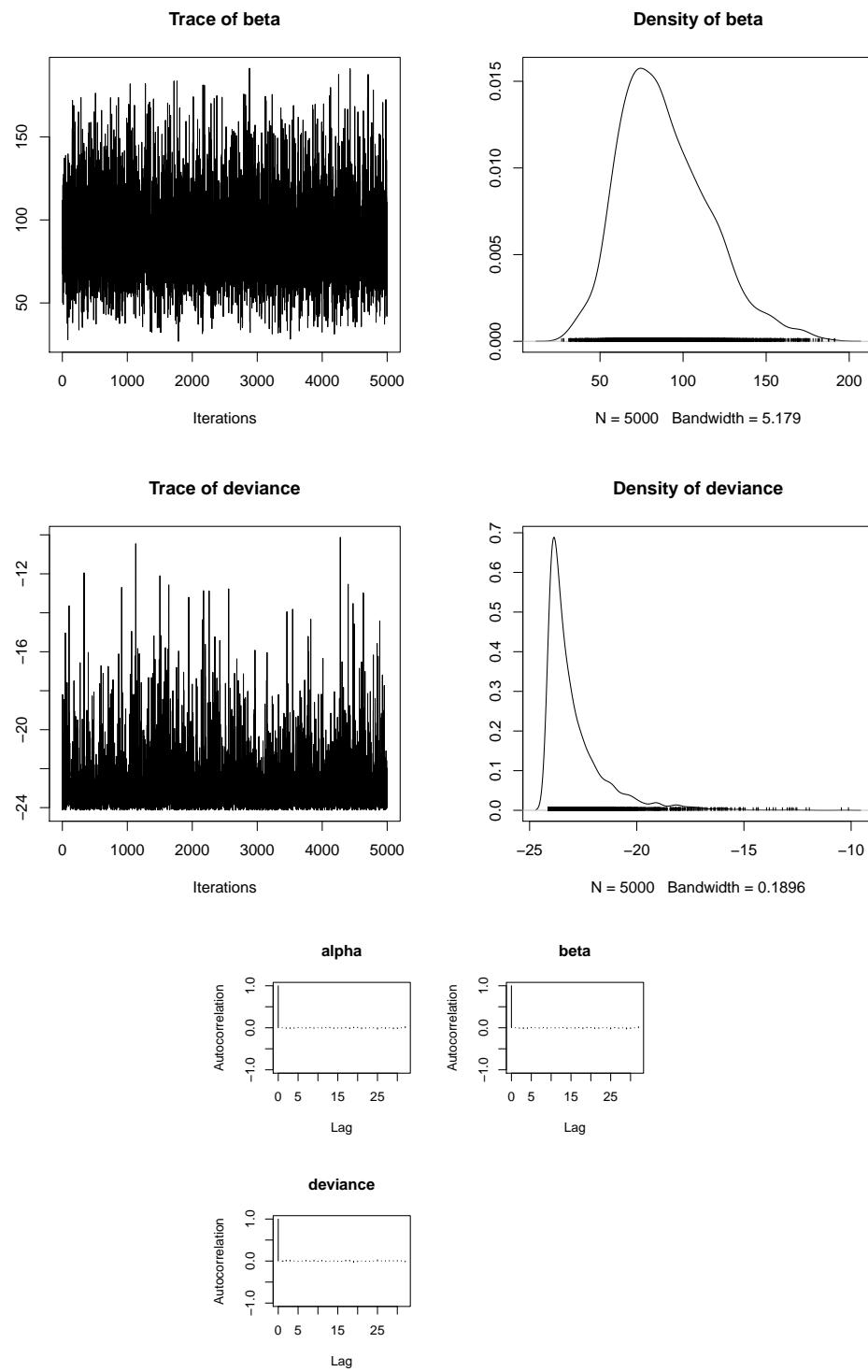
# Trace Plots
plot(sims, auto.layout = FALSE, ask = FALSE)
# Autocorrelation plots
autocorr.plot(sims, ask = FALSE)
effectiveSize(sims)

##      alpha      beta deviance
##      5000      5000      5000

raftery.diag(sims)

##
## Quantile (q) = 0.025
## Accuracy (r) = +/- 0.005
## Probability (s) = 0.95
##
##      Burn-in Total Lower bound Dependence
##            (M)     (N)   (Nmin)       factor (I)
##  alpha    2      3741  3746      0.999
##  beta     2      3680  3746      0.982
##  deviance 2      3680  3746      0.982
```





Here's the frequentest method for an ANOVA

```
fit.1 <- lm(y ~ x)
anova(fit.1)

## Analysis of Variance Table
##
## Response: y
```

```
##           Df Sum Sq Mean Sq F value Pr(>F)
## x          1 13600   13600    1813 1e-10 ***
## Residuals  8      60       7
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

to check the value of the slope

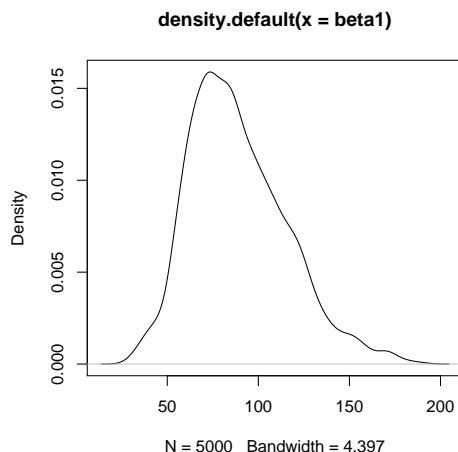
```
beta0 <- sims[, 1]
beta1 <- sims[, 2]
sigma2 <- sims[, 4]

## Error: subscript out of bounds

# Probability slope is greater than 0
mean(beta1 > 0)

## [1] 1

# Plot the density of the slope of the line
plot(density(beta1))
```



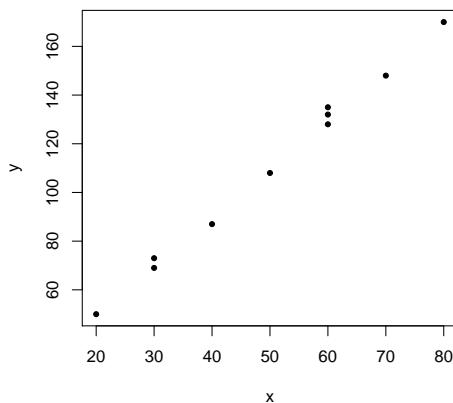
Plot of regression

```
plot(x, y, pch = 20)
abline(mean(beta0), mean(beta1))
xx <- seq(min(x), max(x), by = 0.01)

ans <- beta0 + outer(beta1, xx, "*")
lines(xx, apply(ans, 2, quantile, 0.025), col = "blue")
lines(xx, apply(ans, 2, quantile, 0.975), col = "blue")

# generate a new data set with the first set of
points(xx, beta0[1] + beta1[1] * xx + rnorm(length(xx), 0, sqrt(sigma2[1])), pch = 20)

## Error: object 'sigma2' not found
```



Homework Plot the prediction interval from the problem in class

generate a normal using ans each ans and sigma

* * January 29, 2013 *

7 Multiple Linear Regression

```

1 %let datadir = Z:\Dropbox\Active\STAT451\Notes\data;
  data vo2;
    infile "&datadir.\03vo2.dat" firstobs=2;
    input id gender age bmi mph hr rpe maxvo2;
5 run;

ods _all_ close;
ods html;

10 proc mcmc data=vo2 outpost=vo2out seed=1234
    nmc=300000
    nbi=60000
    thin=30
    statistics=(summary interval)
15   diagnostics=(rl ess)
    dic
    propcov=quane
    monitor=(-parms_);
20 parms b1 50 b2 10 b3 0 b4 0 b5 0 b6 0 s2f 20 s2m 30;
25   prior b1 ~ normal(50, var=1000);
    prior b2 ~ normal(10, var=1000);
    prior b3 ~ normal(0, var=1000);
    prior b4 ~ normal(0, var=1000);
    prior b5 ~ normal(0, var=1000);
    prior b6 ~ normal(0, var=1000);

```

```

prior s2f ~ gamma(2, scale=20);
prior s2m ~ gamma(2, scale=30);

30
mu= b1 + b2*gender + b3*bmi + b4*mph + b5*hr + b6*rpe;
if gender=0 then vv=s2f;
else vv=s2m;

35
model maxvo2 ~ normal(mu, var=vv);
run;

```

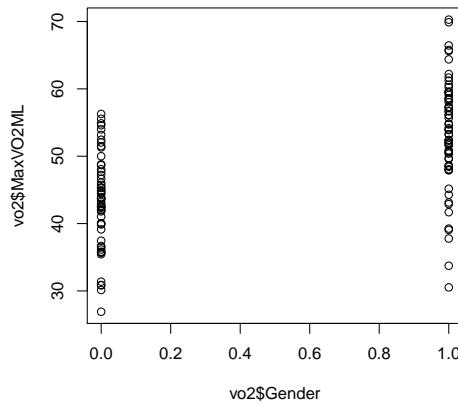
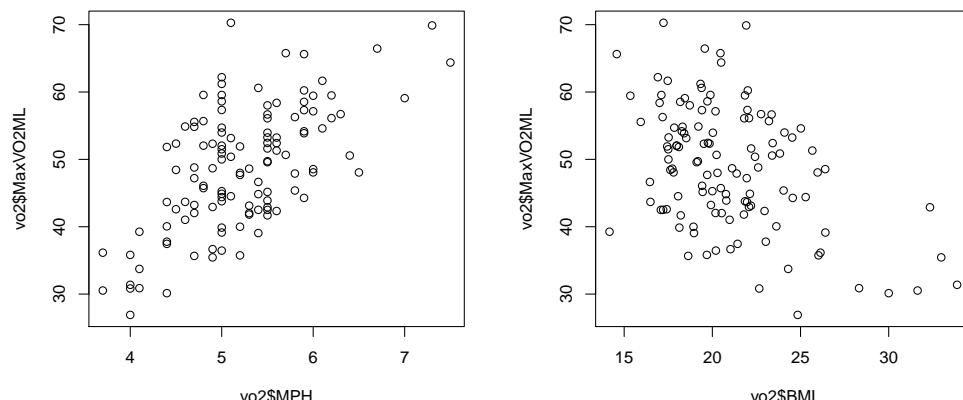
Raftery-lewis diagnostics should be less than 5.

Let the variance increase proportionaly with the *xs*

```

vo2 <- read.table("data/03vo2.dat", skip = 1, col.names = c("ID", "Gender", "Age1",
  "BMI", "MPH", "HR", "RPE", "MaxVO2ML"))
plot(vo2$MPH, vo2$MaxVO2ML)
plot(vo2$BMI, vo2$MaxVO2ML)
plot(vo2$Gender, vo2$MaxVO2ML)

```



DIC Deviance information criteria (calculated in formula 1)

$$\text{deviance} = -2 \log(\mathcal{L}) \quad (1)$$

* * January 31, 2013 *

```
# Using the VO2 Data
N <- nrow(vo2)
mdl <- '
model {
  for( i in 1:N){
    y[i] ~ dnorm(mu[i], prec)
    mu[i] <- b_0 +
      b_gen*gender[i] +
      b_bmi*bmi[i] +
      b_mph*mph[i] +
      b_hr*hr[i] +
      b_rpe*rpe[i] +
      # b_gen_bmi*gen_bmi[i] +
      # b_gen_mph*gen_mph[i] +
      # b_gen_hr*gen_hr[i] +
      # b_gen_rpe*gen_rpe[i] +
      # b_bmi_mph*bmi_mph[i] +
      # b_bmi_hr*bmi_hr[i] +
      # b_bmi_rpe*bmi_rpe[i] +
      b_mph_hr*mph_hr[i] # +
      # b_mph_rpe*mph_rpe[i] +
      # b_hr_rpe*hr_rpe[i]
  }
  b_0 ~ dnorm(0, .00001);
  b_gen ~ dnorm(0, .001);
  b_bmi ~ dnorm(0, .001);
  b_mph ~ dnorm(0, .001);
  b_hr ~ dnorm(0, .001);
  b_rpe ~ dnorm(0, .001);

  # b_gen_bmi ~ dnorm(0, .001);
  # b_gen_mph ~ dnorm(0, .001);
  # b_gen_hr ~ dnorm(0, .001);
  # b_gen_rpe ~ dnorm(0, .001);
  # b_bmi_mph ~ dnorm(0, .001);
  # b_bmi_hr ~ dnorm(0, .001);
  # b_bmi_rpe ~ dnorm(0, .001);
  b_mph_hr ~ dnorm(0, .001);
  # b_mph_rpe ~ dnorm(0, .001);
  # b_hr_rpe ~ dnorm(0, .001);

  vr ~ dgamma(2, .05)
  prec <- 1/vr
}
'
writeLines(mdl, "code/multipleRegression.jags")

y <- vo2[,8]
gender <- vo2[,2]
bmi <- vo2[,4]
hr <- vo2[,6]
mph <- vo2[,5]
rpe <- vo2[,7]

# gen_bmi <- gender*bmi
# gen_hr <- gender*hr
# gen_mph <- gender*mph
# gen_rpe <- gender*rpe
# bmi_mph <- bmi*mph
# bmi_hr <- bmi*hr
# bmi_rpe <- bmi*rpe
mph_hr <- mph*hr
# mph_rpe <- mph*rpe
# hr_rpe <- hr*rpe
```

```

jags.data <- c(
  "N",
  "y",
  "gender",
  "bmi",
  "hr",
  "mph",
  "rpe",
  # "gen_bmi",
  # "gen_hr",
  # "gen_mph",
  # "gen_rpe",
  # "bmi_mph",
  # "bmi_hr" #,
  # "bmi_rpe",
  # "mph_hr" #,
  # "mph_rpe",
  # "hr_rpe",
)
jags.params <- c(
  "b_0",
  "b_gen",
  "b_bmi",
  "b_hr",
  "b_mph",
  "b_rpe",
  # "b_gen_bmi",
  # "b_gen_hr",
  # "b_gen_mph",
  # "b_gen_rpe",
  # "b_bmi_mph",
  # "b_bmi_hr",
  # "b_bmi_rpe",
  "b_mph_hr",
  # "b_mph_rpe",
  "vr"
)

set.seed(1234)
multr.sim <- jags(data=jags.data,
                     parameters.to.save=jags.params,
                     model.file="code/multipleRegression.jags",
                     n.iter=10000, n.burnin=5000, n.chains=1, n.thin=1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 1294
##
## Initializing model

```

Full	756.4
-hr	764.4
-rpe	759.2
+ all two way interactions	766.7

$$\begin{aligned}
 IC &= -2 \log(\mathcal{L}) + \text{penalty(parameters)} \\
 \text{Deviance} &= -2 \log(\mathcal{L}) \\
 DIC &= -2 \log(\mathcal{L}) + pD \\
 pD &= \frac{\text{VAR(deviance)}}{2}
 \end{aligned}$$

8 ANOVA

Cell means model of ANOVA

Treatment 1	Treatment 2	Treatment 3	Treatment 4
y_{11}	y_{21}	y_{31}	y_{41}
y_{12}	y_{22}	y_{32}	y_{42}
y_{13}	y_{23}	y_{33}	y_{43}
y_{14}	y_{24}	y_{34}	y_{44}

so given this data we say that $y_i \sim N(\mu, \sigma^2)$

```

anova <- read.table("data/04anova.dat", col.names=c("tmt", "response", "tmt1"))
N <- nrow(anova)
mdl <- 'model {
  # Likelihood
  for (i in 1:N){
    response[i] ~ dnorm(mu[tmt[i]], prec)
  }

  # Create priors for each treatment
  for(i in 1:4){
    mu[i] ~ dnorm(15,.0001)
  }
  prec <- 1/vv
  vv ~ dgamma(1.1,.1)
}

writeLines(mdl, "code/ANOVAmodel.jags")
response <- anova$response
tmt <- anova$tmt

jags.data <- c("N", "response", "tmt")
jags.params <- c("mu", "vv")
anova.sim <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/ANOVAmodel.jags",
n.iter = 12000, n.chains = 1, n.thin = 1)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 68
##
## Initializing model

sims <- as.mcmc(anova.sim)
plot(sims, auto.layout = FALSE, ask = FALSE)
autocorr.plot(sims)
raftery.diag(sims)

##
## Quantile (q) = 0.025

```

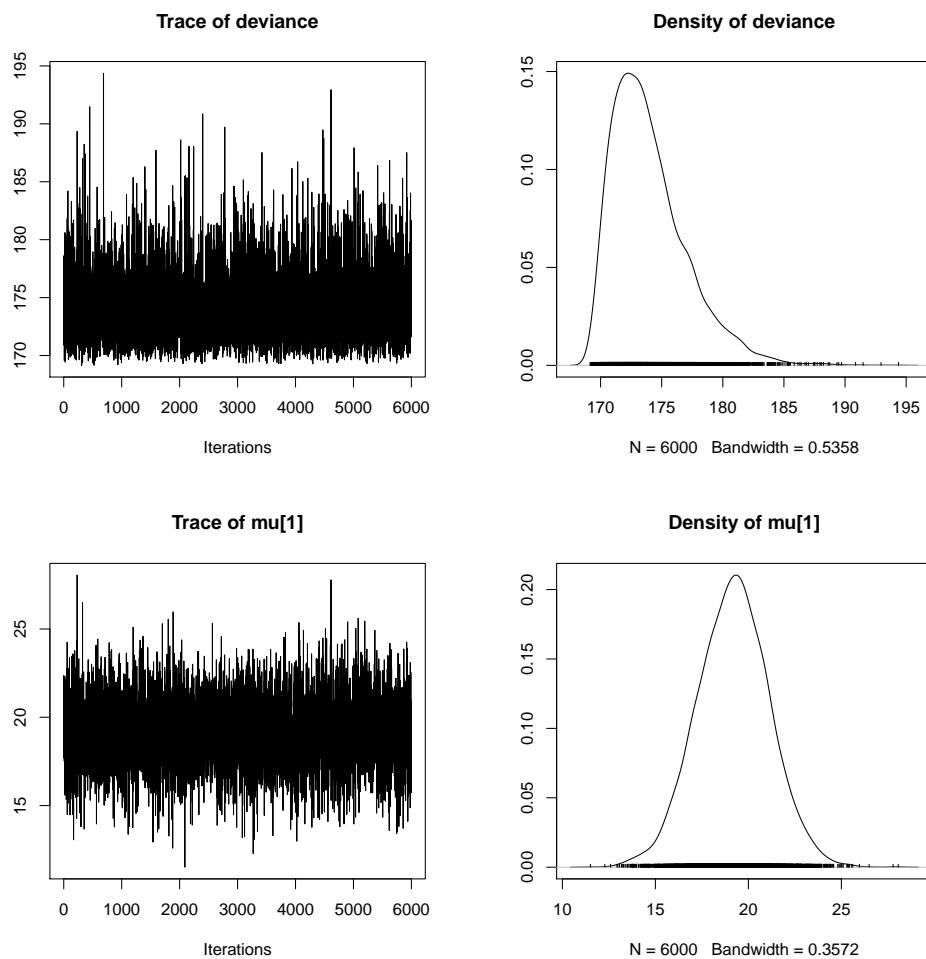
```

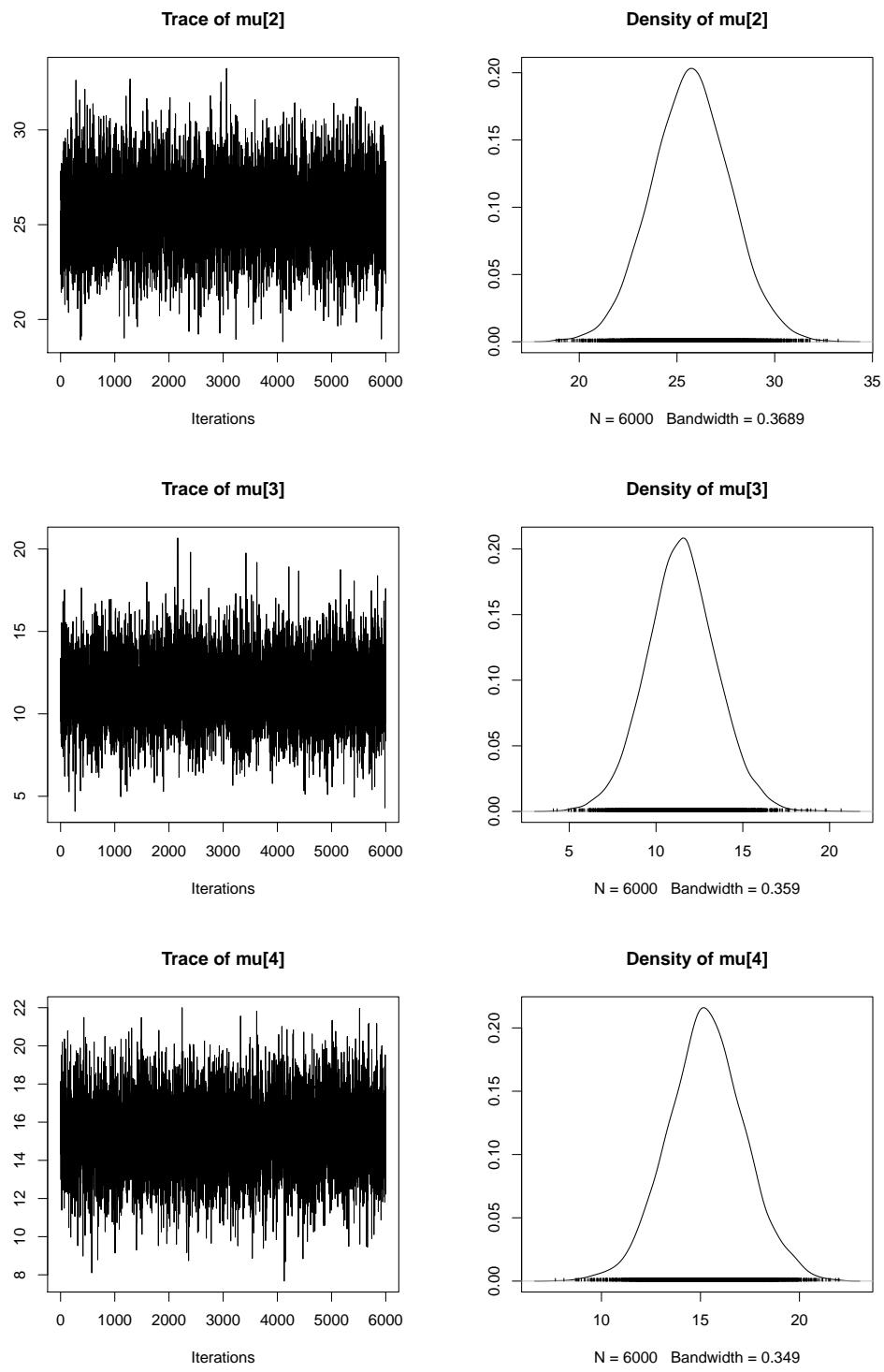
## Accuracy (r) = +/- 0.005
## Probability (s) = 0.95
##
##          Burn-in   Total Lower bound Dependence
##          (M)      (N)    (Nmin)   factor (I)
## deviance 2       3761   3746      1.00
## mu[1]    2       3710   3746      0.99
## mu[2]    2       3761   3746      1.00
## mu[3]    2       3761   3746      1.00
## mu[4]    2       3813   3746      1.02
## vv       3       4028   3746      1.08

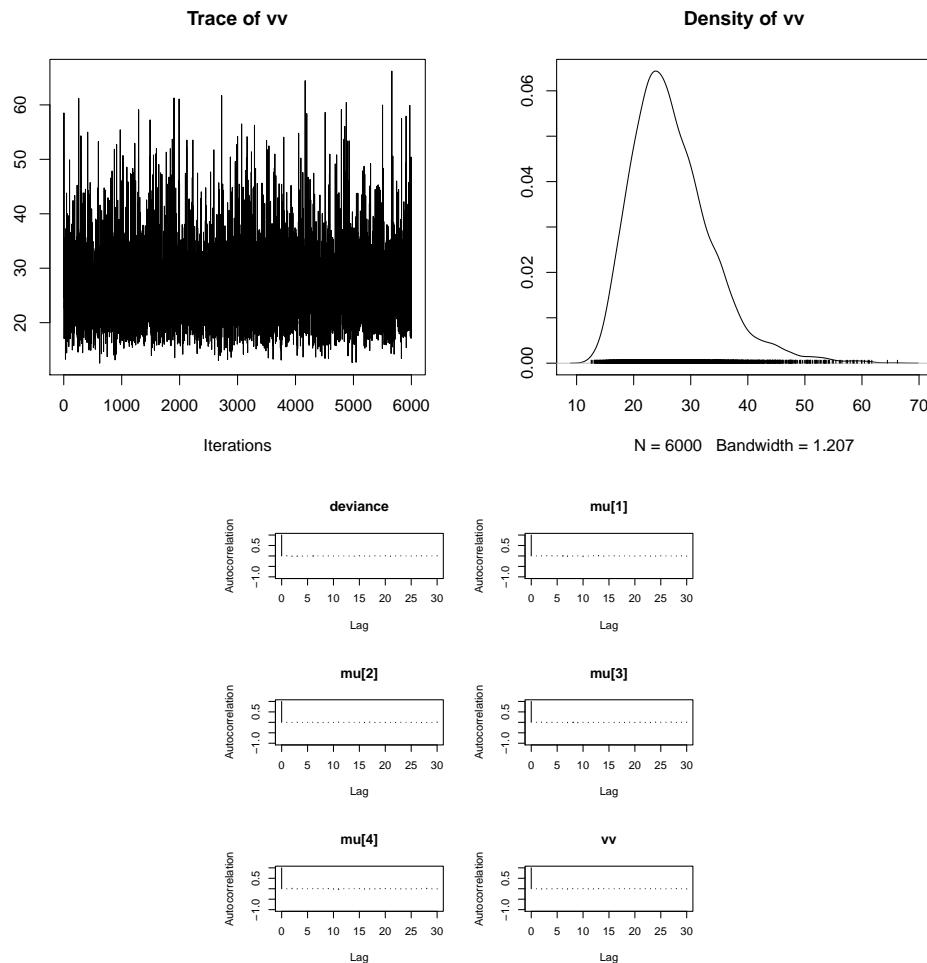
effectiveSize(sims)

## deviance   mu[1]     mu[2]     mu[3]     mu[4]       vv
## 6448       6000     6000     6000     5800     6000

```







```

mdl <- 'model {
  # Likelihood
  for (i in 1:N){
    response[i] ~ dnorm(mu[tmt[i]], prec[tmt[i]])
  }

  # Create priors for each treatment
  for(i in 1:4){
    mu[i] ~ dnorm(15,.0001)
    prec[i] <- 1/vv[i]
    vv[i] ~ dgamma(1.1,.1)
  }
}

writeLines(mdl, "code/ANOVAmodelNCvariance.jags")
si

## Error: object 'si' not found'

```

Devience is the $-2\log$ likelihood

8.1 Compare Means

We don't care about multiple test because we are not under the constraint of a null hypothesis

```
diff12 <- sims[, 2] - sims[, 3]
mean(dif12 > 0)

## Error: object 'dif12' not found

dif13 <- sims[, 2] - sims[, 4]
dif14 <- sims[, 2] - sims[, 5]
dif23 <- sims[, 3] - sims[, 4]
dif24 <- sims[, 3] - sims[, 5]
dif34 <- sims[, 4] - sims[, 5]
mean(dif13 > 0)

## [1] 0.9958

mean(dif14 > 0)

## [1] 0.9277

mean(dif23 > 0)

## [1] 1

mean(dif24 > 0)

## [1] 0.9995

mean(dif34 > 0)

## [1] 0.08267
```

Two by three factorial design.

```
twoway <- read.table("data/05twoway.dat", col.names = c("block", "seedType", "inoculate",
"yield"))
```

Using cell means method

		a	b	
Table should be rotated	dea	μ_{11}	μ_{12}	$\mu_{1\cdot}$
	con	μ_{21}	μ_{22}	$\mu_{2\cdot}$
	liv	μ_{31}	μ_{22}	$\mu_{3\cdot}$
		$\mu_{\cdot 1}$	$\mu_{\cdot 2}$	

Degrees of freedom is 5, one for type, two for inoculate, and two for interaction

```
y <- as.numeric(twoway$seedType)

mdl <- '
model {
  for (i in 1:24) {
    yield[i] ~ dnorm(mu[ncult[i],ninoc[i]], 1/vv)
  }

  for (i in 1:2) {
    for (j in 1:3) {
      mu[i,j] ~ dnorm(0, .000001)
    }
  }

  vv ~ dgamma(1.5, .1)
}'
```

```

writeLines(mdl, 'code/CultModel.jags')

yield <- twoway$yield
ncult <- as.numeric(twoway$seedType)
ninoc <- as.numeric(twoway$inoculate)

jags.data <- c('yield', 'ncult', 'ninoc')
jags.params <- c('mu', 'vv')

innits <- function(){list('mu'= matrix(0,2,3), 'vv', 15)}

cult1.jags <- jags( data=jags.data,
                     # inits=innits,
                     parameters.to.save=jags.params,
                     model.file='code/CultModel.jags',
                     n.ITER=12000, n.burnin=2000,
                     n.chains=1, n.thin=1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 85
##
## Initializing model

cult1.sim <- as.mcmc(cult1.jags)

```

Effects model and cell means model are two different ways to do analysis of variance. BYU was a pioneer of the cell means model.

8.2 ANOVA Marginal Means

marginals are calculated as $\mu_{1\cdot} = \frac{\mu_{11} + \mu_{12} + \mu_{13}}{3}$ its just a mean of the means

```

mu1dot <- (cult1.sim[, 2] + cult1.sim[, 4] + cult1.sim[, 6])/3
mu2dot <- (cult1.sim[, 3] + cult1.sim[, 5] + cult1.sim[, 7])/3

plot(mu1dot - mu2dot)
mean(mu2dot - mu1dot > 0)

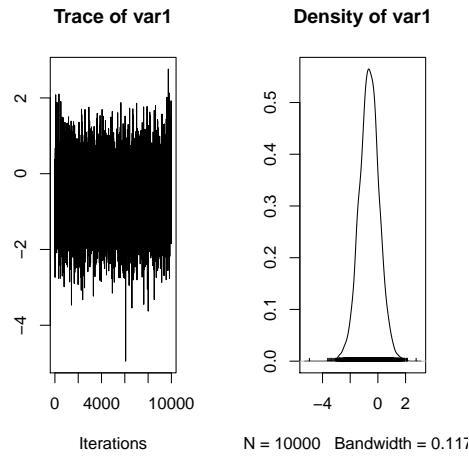
## [1] 0.8171

mudot1 <- (cult1.sim[, 2] + cult1.sim[, 3])/2
mudot2 <- (cult1.sim[, 4] + cult1.sim[, 5])/2
mudot3 <- (cult1.sim[, 6] + cult1.sim[, 7])/2

mean(mudot1 - mudot2 < 0) # Probability mu_dot2 is bigger than mu_dot1
## [1] 0.9858

mean(mudot3 - mudot2 > 0) # Probability mu_dot3 is bigger than mu_dot2
## [1] 0.9991

```



8.3 ANOVA Interactions

interactions are tested in four cell groups the formula is

$$\begin{aligned}\mu_{11} - \mu_{21} &= \mu_{12} - \mu_{22} \\ \mu_{11} - \mu_{21} - \mu_{12} + \mu_{22} &= 0\end{aligned}$$

```
int1 <- cult1.sim[, 2] - cult1.sim[, 4] - cult1.sim[, 3] - cult1.sim[, 4]
int2 <- cult1.sim[, 4] - cult1.sim[, 6] - cult1.sim[, 5] + cult1.sim[, 7]
```

Finally, compare to seed types that with the best inoculate, does it make sense to pay for the better seed or does the $\mu_{11} - \mu_{21} = \mu_{12} - \mu_{22}$ come from the inoculate

```
se1 <- cult1.sim[, 6] - cult1.sim[, 7]
mean(se1 < 0)
## [1] 0.6194
```

8.4 Linear Algebra Solution

$$\hat{\mu} = (W'W)^{-1}W'y$$

Least Squares Solution

```
w <- rbind(diag(1, 6), diag(1, 6), diag(1, 6), diag(1, 6))
muhat <- solve(t(w) %*% w) %*% t(w) %*% twoway$yield
```

$$\begin{aligned}y &= W\mu \\ y &= WI\mu \\ y &= \underbrace{WA^{-1}}_X \underbrace{A\mu}_\beta\end{aligned}$$

$$\begin{bmatrix} \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} \\ \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} & \frac{1}{6} \\ \frac{1}{2} & -\frac{1}{2} & 0 & \frac{1}{2} & -\frac{1}{2} & 0 \\ 0 & \frac{1}{2} & -\frac{1}{2} & 0 & \frac{1}{2} & -\frac{1}{2} \\ 1 & -1 & 0 & -1 & 1 & 0 \\ 0 & 1 & -1 & 0 & -1 & 1 \end{bmatrix} \begin{bmatrix} \mu_{11} \\ \mu_{12} \\ \mu_{13} \\ \mu_{21} \\ \mu_{22} \\ \mu_{23} \end{bmatrix} = \begin{bmatrix} \mu_{..} \\ \mu_{1.} - \mu_{2.} \\ \mu_{.1} - \mu_{.2} \\ \mu_{.2} - \mu_{.3} \\ \mu_{11} - \mu_{12} - \mu_{21} + \mu_{22} \\ \mu_{12} - \mu_{13} - \mu_{22} + \mu_{23} \end{bmatrix}$$

* * February 12, 2013 *

```
dev <- function(x) {
  fp <- length(y) * log(2*pi)
  sp <-
}
```

* * February 14, 2013 *

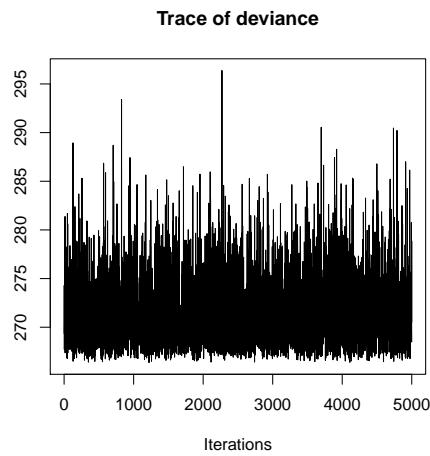
9 Analysis of Covariance ANCOVA

Join Me 777-129-200

```
ac <- read.table("data/06ancova.dat", header = TRUE)
```

First thing you do for analysis is plot the data

```
plot(ac$speed, ac$scrap, pch = ac$lines)
```



Let's try and come up with a model

$$y_{ij} = \beta_0 + \beta_1 \cdot \text{Line} + \beta_2 \cdot \text{Speed} + \beta_3 \cdot \text{Line} \cdot \text{Speed} \quad (2)$$

Or we could do this

$$y_{ij} = \beta_{0i} + \beta_{1i} \cdot \text{Speed}$$

```

mdl <- 'model {
  for(i in 1:27) {
    scrap[i] ~ dnorm(mu[i], prec);
    mu[i] <- b_0[line[i]] + b_1[line[i]]*speed[i];
  }

  for (i in 1:2) {
    b_0[i] ~ dnorm(30, .001);
    b_1[i] ~ dnorm(0, .01);
  }

  vr ~ dgamma(1.5, .0125);
  prec <- 1/vr;
}

writeLines(mdl, "code/ANCOVAModel.jags")

```

prior for beta not covers 30×100

```

speed <- ac$speed
line <- ac$line
scrap <- ac$scrap
jags.data <- c('speed', 'line', 'scrap')
jags.params <- c('b_0', 'b_1', 'vr')
ancova.jags <- jags(
  data=jags.data,
  # inits=innits,
  parameters.to.save=jags.params,
  model.file='code/ANCOVAModel.jags',
  n.iter=12000, n.burnin=2000,
  n.chains=1, n.thin=1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 148
##
## Initializing model

ancova.jags

## Inference for Bugs model at "code/ANCOVAModel.jags", fit using jags,
## 1 chains, each with 12000 iterations (first 2000 discarded)
## n.sims = 10000 iterations saved
##      mu.vect sd.vect   2.5%    25%    50%    75%   97.5%
## b_0[1]    81.738 15.570 50.441 71.518 81.938 92.185 111.585
## b_0[2]    13.553 16.417 -18.292  2.490 13.616 24.485 45.839
## b_1[1]     1.218  0.074  1.079  1.168  1.217  1.266  1.365
## b_1[2]     1.297  0.074  1.149  1.248  1.297  1.347  1.440
## vr       363.557 85.792 229.893 302.708 351.820 411.422 569.708
## deviance 241.252  3.179 236.996 238.870 240.626 242.949 249.002
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 5.1 and DIC = 246.3
## DIC is an estimate of expected predictive error (lower deviance is better).

```

Diagnostics

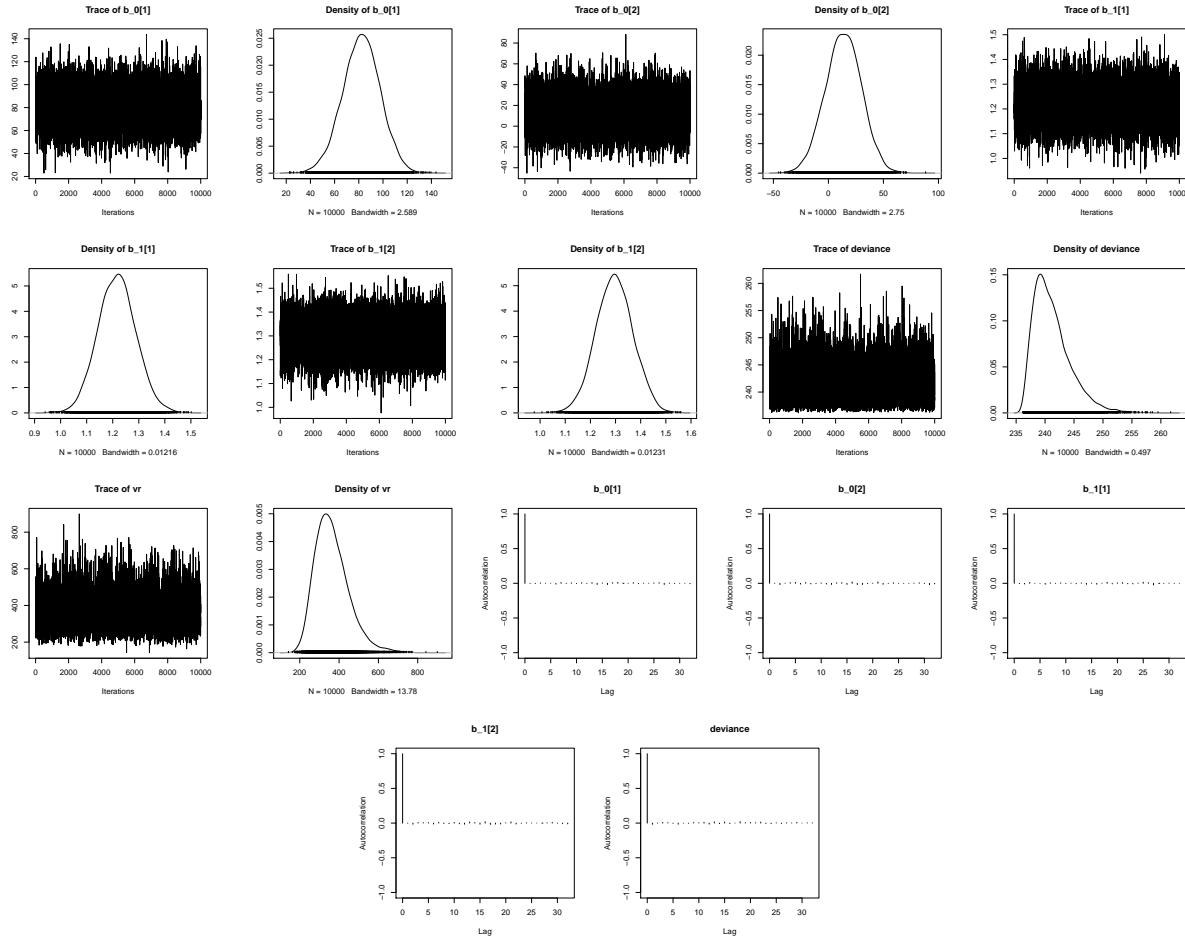
```

ancova.sim <- as.mcmc(ancova.jags)
plot(ancova.sim, auto.layout = FALSE, ask = FALSE)
autocorr.plot(ancova.sim, auto.layout = FALSE, ask = FALSE)
rafthery.diag(ancova.sim)

##
## Quantile (q) = 0.025
## Accuracy (r) = +/- 0.005
## Probability (s) = 0.95
##

```

```
##          Burn-in Total Lower bound Dependence
##          (M)      (N)  (Nmin)   factor (I)
## b_0[1]    2     3650  3746     0.974
## b_0[2]    2     3680  3746     0.982
## b_1[1]    2     3680  3746     0.982
## b_1[2]    2     3834  3746     1.020
## deviance  2     3680  3746     0.982
## vr        2     3620  3746     0.966
```



Test the probability that the β_1 are different

```
b11 <- ancova.sim[, 3]
b12 <- ancova.sim[, 4]
slopdif <- b11 - b12
mean(slopdif < 0)

## [1] 0.7814

mean(slopdif > 0)

## [1] 0.2186
```

Try a new model without the different β_1 then compare the DIC

```
mdl <- 'model {
  for(i in 1:27) {
    scrap[i] ~ dnorm(mu[i], prec);
    mu[i] <- b_0[line[i]] + b_1*speed[i];
  }
}'
```

```

b_1 ~ dnorm(0, .01);
for (i in 1:2) {
  b_0[i] ~ dnorm(30, .001);
}

vr ~ dgamma(1.5, .0125);
prec <- 1/vr;
}
'
writeLines(mdl, 'code/ANCOVAModel_1.jags')

```

prior for beta not covers

```

speed <- ac$speed
line <- ac$line
scrap <- ac$scrap
jags.data <- c('speed', 'line', 'scrap')
jags.params <- c('b_0', 'b_1', 'vr')
ancova1.jags <- jags( data=jags.data,
  # inits=innits,
  parameters.to.save=jags.params,
  model.file='code/ANCOVAModel_1.jags',
  n.iter=12000, n.burnin=2000,
  n.chains=1, n.thin=1)

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 142
##
## Initializing model

ancova1.jags

## Inference for Bugs model at "code/ANCOVAModel_1.jags", fit using jags,
## 1 chains, each with 12000 iterations (first 2000 discarded)
## n.sims = 10000 iterations saved
##      mu.vect sd.vect   2.5%    25%    50%    75% 97.5%
## b_0[1]    74.154 11.580 50.987 66.643 74.236 81.949 96.90
## b_0[2]    21.933 12.213 -2.044 13.621 22.005 30.055 45.95
## b_1       1.257  0.052  1.156  1.222  1.256  1.291  1.36
## vr       369.910 85.910 233.121 308.243 359.554 419.347 564.88
## deviance 242.053  2.761 238.636 240.016 241.464 243.406 248.97
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 3.8 and DIC = 245.9
## DIC is an estimate of expected predictive error (lower deviance is better).

```

Diagnostics

```

ancova1.sim <- as.mcmc(ancova1.jags)
# plot(ancova.sim, auto.layout=FALSE, ask=FALSE) autocorr.plot(ancova.sim,
# auto.layout=FALSE, ask=FALSE) raftery.diag(ancova.sim)

```

Lets test differences in line by testing difference in intercept

```

b_0_1 <- ancova1.sim[, 1]
b_0_2 <- ancova1.sim[, 2]
line_diff <- b_0_1 - b_0_2
mean(line_diff < 0)

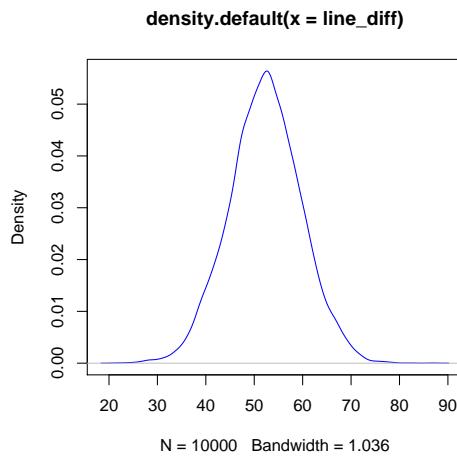
## [1] 0

mean(line_diff > 0)

## [1] 1

plot(density(line_diff), col = "blue")

```



```

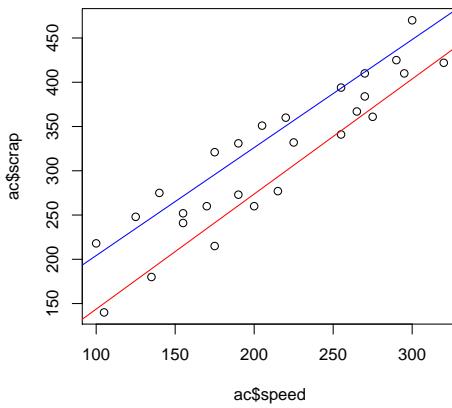
plot(ac$speed, ac$scrap, pch = ac$lines)
abline(82.2, 1.22, col = "blue")
abline(13.7, 1.3, col = "red")

b_0_1 <- ancova.sim[, 1]
b_0_2 <- ancova.sim[, 2]
b_1_1 <- ancova.sim[, 3]
b_1_2 <- ancova.sim[, 4]

y1_300 <- 300 * b_1_1 + b_0_1
y2_300 <- 300 * b_1_2 + b_0_2
diff_300 <- y1_300 - y2_300
mean(diff_300 > 0)

## [1] 1

```



* * February 21, 2013 *

		B				
		1	2	3	4	
A	1	μ_{11}	μ_{12}	μ_{13}	μ_{14}	$\mu_{1\bullet}$
	2	μ_{21}	μ_{22}	μ_{23}	μ_{24}	$\mu_{2\bullet}$
	3	μ_{31}	μ_{32}	μ_{33}	μ_{34}	$\mu_{3\bullet}$
	4	μ_{41}	μ_{42}	μ_{43}	μ_{44}	$\mu_{4\bullet}$

10 Multiple Sources of Variation

$$\begin{aligned}\underline{y} &= X\underline{\beta} + \varepsilon \\ \varepsilon &\sim \mathcal{N}(\underline{0}, \sigma^{2T}) \\ \underline{y} &= X\underline{\beta} + Z\underline{u} + e \\ e &\sim \mathcal{N}(\underline{0}, R) \\ \underline{u} &\sim \mathcal{N}(\underline{0}, G)\end{aligned}$$

$$\begin{aligned}\text{E}(Y) &= \text{E}(X\beta + Zu + e) = \text{E}(X\beta) + \text{E}(Zu) + \text{E}(e) = X\beta \\ \text{Var}(Y) &= \text{Var}(X\beta + Zu + e) = \text{Var}(Zu + e)\end{aligned}$$

If we assume that u and e are independent then we can move forward easily

$V(A\underline{z}) = AV(z)A'$ $AV(z)A'$ is a $p \times p$ matrix

$$\text{Var}(Zu + e) = V(Z\underline{u}) + \text{Var}(\underline{e}) = Z\text{Var}(\underline{u})Z' + V(\underline{e}) = ZGZ' + R$$

*

*

February 26, 2013

*

*

	Block 1	Block 2
Treatments:	I (y_{11})	I (y_{12})
	II (y_{21})	II (y_{22})

Model in matrix form (Equation 3)

$$\begin{bmatrix} y_{11} \\ y_{12} \\ y_{21} \\ y_{22} \end{bmatrix} = \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \end{bmatrix} \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} + \begin{bmatrix} 1 & 0 \\ 0 & 1 \\ 1 & 0 \\ 0 & 1 \end{bmatrix} \begin{pmatrix} u_1 \\ u_2 \end{pmatrix} + \begin{bmatrix} e_{11} \\ e_{12} \\ e_{21} \\ e_{22} \end{bmatrix} \quad (3)$$

$$\begin{aligned}e_{ij} &\sim \text{iid } N(0, \sigma_e^2) \\ u_j &\sim \text{iid } N(0, \sigma_b^2)\end{aligned}$$

$$\begin{aligned}R &= \begin{bmatrix} \sigma_e^2 & 0 & 0 & 0 \\ 0 & \sigma_e^2 & 0 & 0 \\ 0 & 0 & \sigma_e^2 & 0 \\ 0 & 0 & 0 & \sigma_e^2 \end{bmatrix} \\ G &= \begin{pmatrix} \sigma_b^2 & 0 \\ 0 & \sigma_b^2 \end{pmatrix}\end{aligned}$$

$$\begin{aligned} & \begin{bmatrix} 1 & 0 \\ 1 & 0 \\ 0 & 1 \\ 0 & 1 \end{bmatrix} \begin{pmatrix} \sigma_b^2 & 0 \\ 0 & \sigma_b^2 \end{pmatrix} \begin{bmatrix} 1 & 0 & 1 & 0 \\ 0 & 1 & 0 & 1 \end{bmatrix} + \begin{bmatrix} \sigma_e^2 & 0 & 0 & 0 \\ 0 & \sigma_e^2 & 0 & 0 \\ 0 & 0 & \sigma_e^2 & 0 \\ 0 & 0 & 0 & \sigma_e^2 \end{bmatrix} \\ ZGZ' + R = & \begin{bmatrix} \sigma_e^2 + \sigma_b^2 & 0 & \sigma_b^2 & 0 \\ 0 & \sigma_e^2 + \sigma_b^2 & 0 & \sigma_b^2 \\ \sigma_b^2 & 0 & \sigma_e^2 + \sigma_b^2 & 0 \\ 0 & \sigma_b^2 & 0 & \sigma_e^2 + \sigma_b^2 \end{bmatrix} \end{aligned}$$

$$\rho_{y_{11}y_{21}} = \frac{\sigma_b^2}{\sqrt{\sigma_3^2 + \sigma_b^2} \sqrt{\sigma_3^2 + \sigma_b^2}} = \frac{\sigma_b^2}{\sigma_3^2 + \sigma_b^2}$$

```
twoway <- read.table("data/05twoway.dat", col.names = c("block", "seedType", "inoculate",
"yield"))
```

First model with multiple sources of variance, s2blk is a hyper-prior, or a prior on a prior

```
mdl <- 'model {
  for (i in 1:21) {
    yield[i] ~ dnorm(mu[i], 1/s2e)
    mu[i] <- aaron[cultn[i], inocn[i]] + u[block[i]]
  }

  for (i in 1:2) {
    for (j in 1:3) {
      aaron[i,j] ~ dnorm(30, .001);
    }
  }

  for (i in 1:4) {
    u[i] ~ dnorm(0, 1/s2blk)
  }

  s2e ~ dgamma(1.5, .1);
  s2blk ~ dgamma(1.5, .1);
}

writeLines(mdl, 'code/MixedModels.jags')

yield <- twoway[, 4]
cultn <- as.numeric(twoway$seedType)
inocn <- as.numeric(twoway$inoculate)
block <- twoway[, 1]

jags.data <- c("yield", "cultn", "inocn", "block")
jags.params <- c("aaron", "s2e", "s2blk")
set.seed(343)
mixedm.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/MixedModels.jags",
n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
##   Graph Size: 137
##
## Initializing model
```

Difference between 1 and 2

```

mixedm.sim <- as.mcmc(mixedm.jags)
cultivarA <- (mixedm.sim[, 1] + mixedm.sim[, 3] + mixedm.sim[, 5])/3
cultivarB <- (mixedm.sim[, 2] + mixedm.sim[, 4] + mixedm.sim[, 6])/3

mean(cultivarA)

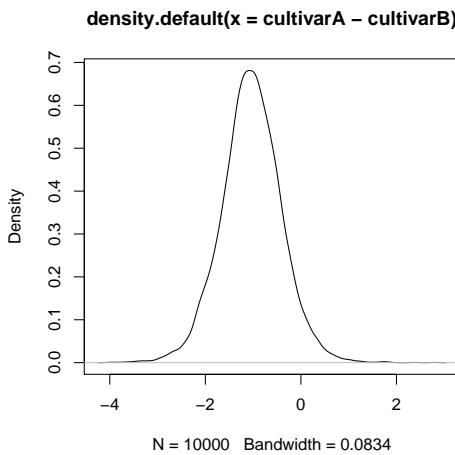
## [1] 30.09

mean(cultivarB)

## [1] 31.13

plot(density(cultivarA - cultivarB))

```



* * February 28, 2013 *

Simpler Mixed Model

```

bond <- read.table("data/07mixedmods.dat", header = TRUE)
metn <- rep(1:3, 21/3)
bond <- cbind(bond, metn)

mdl <- " model{\nfor (i in 1:21){\npressure[i] ~ dnorm(mu[i], 1/s2error);\nmu[i] <- gamma[metn[i]] + u[ingot[i]]\n}\ninvisible(\nwriteLines(mdl, \"code/SimpleMixedModel.jags\")

metn <- bond[, 4]
ingot <- bond[, 1]
pressure <- bond[, 3]

jags.data <- c("metn", "ingot", "pressure")
jags.params <- c("gamma", "s2error", "s2ing")

bond.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/SimpleMixedModel.jags",
n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

## Error:
## Error parsing model file:
## syntax error on line 6 near ""

bond.sim <- as.mcmc(bond.jags)

## Error: object 'bond.jags' not found

```

ICC or interclass corelation is a measure of reliability of measurements

$$ICC = \frac{\sigma_i^2}{\sigma_e^2 + \sigma_i^2} \quad (4)$$

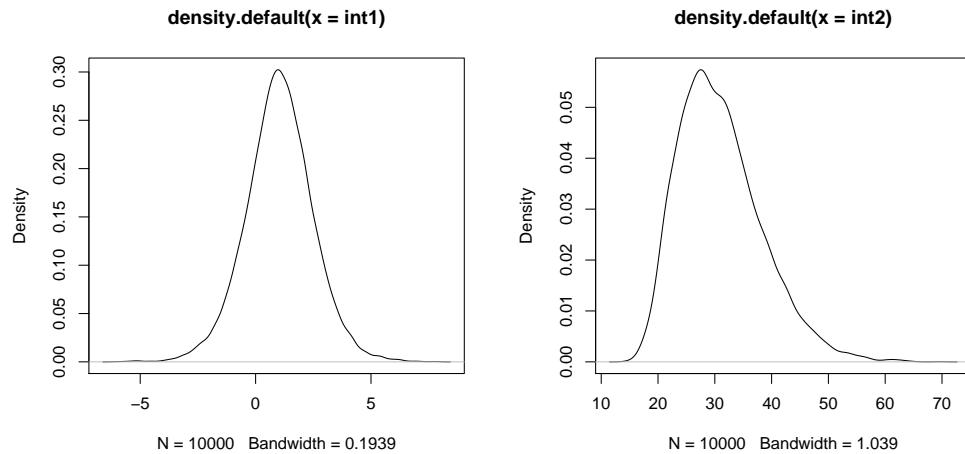
```
icc <- bond.sim[, 5]/(bond.sim[, 6] + bond.sim[, 5])
## Error: object 'bond.sim' not found
plot(density(icc))
## Error: object 'icc' not found

plot something
indif <- bond.sim[, 3] - bond.sim[, 2]
## Error: object 'bond.sim' not found
icdif <- bond.sim[, 3] - bond.sim[, 4]
## Error: object 'bond.sim' not found
```

Back to Complex

			inoc	
		con	dea	liv
cult	a	α_{11}	α_{12}	α_{13}
	b	α_{21}	α_{22}	α_{23}

```
int1 <- mixedm.sim[, 2] - mixedm.sim[, 4] - mixedm.sim[, 3] + mixedm.sim[, 5]
int2 <- mixedm.sim[, 4] - mixedm.sim[, 6] - mixedm.sim[, 5] + mixedm.sim[, 7]
plot(density(int1))
plot(density(int2))
```



Let's say that there's no interaction, because

* * March 5, 2013 *

11 Random Coefficients Model

Genetically identical seeds we treat as subjects

Sixty observations, but only ten wheat types, so there are really somewhere in the middle

$$\text{yield}_{ij} = \beta_0 + \beta_1 \cdot \text{moisture}_j + u_{1i} + u_{2i} \cdot \text{moisture}_j + \text{error}_{ij}$$

so the deviation u has both a slope and an intercept.

$$y = \begin{bmatrix} 1 & 10 \\ 1 & 17 \\ 1 & \vdots \\ \vdots & \ddots \end{bmatrix} \begin{pmatrix} \beta_0 \\ \beta_1 \end{pmatrix} \begin{bmatrix} 1 & 10 & 0 & 0 & 0 & 0 & \dots \\ 1 & 57 & 0 & 0 & 0 & 0 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \dots \\ 1 & 40 & 0 & 0 & 0 & 0 & \dots \\ 0 & 0 & 1 & 16 & 0 & 0 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \dots \\ 0 & 0 & 0 & 0 & 1 & 39 & \dots \\ \vdots & \vdots & \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix} \begin{pmatrix} u_{01} \\ u_{11} \\ u_{02} \\ u_{12} \\ u_{03} \\ u_{13} \\ u_{04} \\ u_{14} \\ u_{05} \\ u_{15} \\ u_{06} \\ u_{16} \end{pmatrix}$$

$$G = \begin{bmatrix} \sigma_{int}^2 & 0 & 0 & 0 & \dots \\ 0 & \sigma_{slp}^2 & 0 & 0 & \dots \\ 0 & 0 & \sigma_{int}^2 & 0 & \dots \\ 0 & 0 & 0 & \sigma_{slp}^2 & \dots \\ \vdots & \vdots & \vdots & \vdots & \ddots \end{bmatrix}$$

```
wheat <- read.table("data/10randomcoef.dat", header=TRUE)
yield <- wheat$yield
variety <- wheat$variety
moisture <- wheat$moisture
mdl <- ' model {
  for (i in 1:60) {
    yield[i] ~ dnorm(mu[i], 1/s2err)
    mu[i] <- b0 + b1 * moisture[i] + u0[variety[i]] + u1[variety[i]] * moisture[i]
  }
  b0 ~ dnorm(30, 0.001)
  b1 ~ dnorm(0, 0.1)

  for (i in 1:10) {
    u0[i] ~ dnorm(0, 1/s2int)
    u1[i] ~ dnorm(0, 1/s2slp)
  }

  s2err ~ dgamma(1.1, .5)
  s2int ~ dgamma(1.1, .1)
  s2slp ~ dgamma(1.1, 2)
}
'
writeLines(mdl, 'code/RandomCoeff.jags')
```

```

jags.data <- c("yield", "variety", "moisture")
jags.params <- c("b0", "b1", "u0", "u1", "s2err", "s2int", "s2slp")

rc1.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/RandomCoeff.jags",
  n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

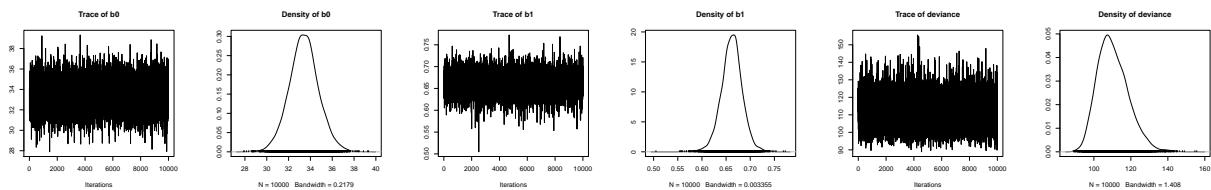
## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 366
##
## Initializing model

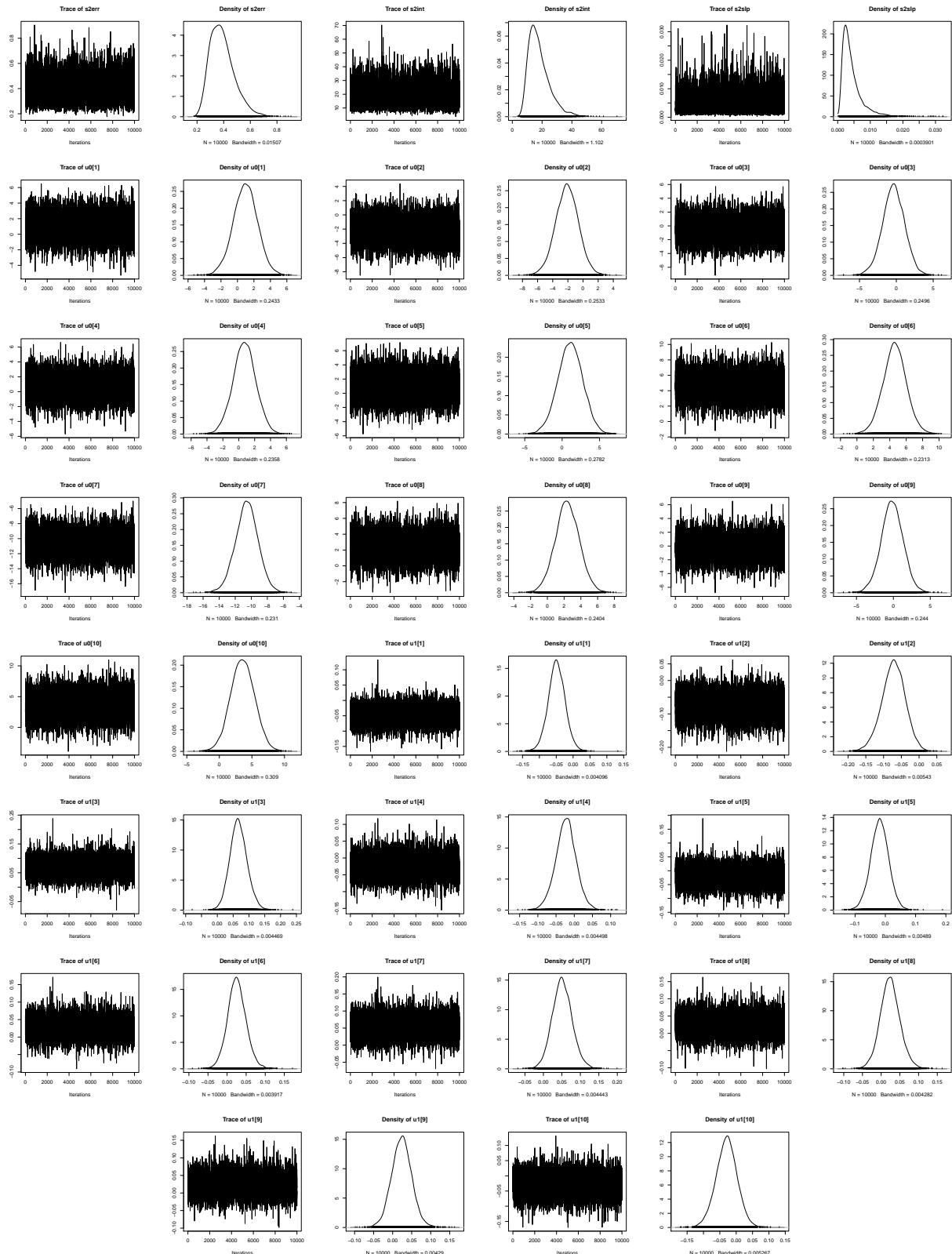
rc1.jags

## Inference for Bugs model at "code/RandomCoeff.jags", fit using jags,
## 1 chains, each with 12000 iterations (first 2000 discarded)
## n.sims = 10000 iterations saved
##      mu.vect sd.vect   2.5%    25%    50%    75%   97.5%
## b0      33.406  1.355 30.710 32.525 33.397 34.263 36.139
## b1       0.662  0.022  0.617  0.648  0.662  0.675  0.704
## s2err     0.389  0.092  0.248  0.322  0.378  0.442  0.603
## s2int    17.953  7.124  8.213 12.823 16.548 21.613 36.010
## s2slp     0.004  0.003  0.001  0.002  0.003  0.005  0.013
## u0[1]     0.958  1.472 -1.995 -0.016  0.964  1.925  3.839
## u0[2]    -2.114  1.518 -5.174 -3.128 -2.109 -1.107  0.894
## u0[3]    -0.371  1.501 -3.410 -1.352 -0.376  0.639  2.534
## u0[4]     0.716  1.447 -2.145 -0.225  0.713  1.656  3.551
## u0[5]     1.097  1.664 -2.225 -0.001  1.102  2.217  4.370
## u0[6]     4.602  1.434  1.760  3.684  4.604  5.529  7.451
## u0[7]   -10.563  1.420 -13.406 -11.479 -10.560 -9.637 -7.808
## u0[8]     2.380  1.449 -0.466  1.433  2.365  3.350  5.277
## u0[9]    -0.162  1.492 -3.148 -1.131 -0.170  0.815  2.783
## u0[10]    3.573  1.840 -0.041  2.339  3.551  4.809  7.153
## u1[1]    -0.049  0.026 -0.099 -0.065 -0.049 -0.032  0.002
## u1[2]    -0.072  0.033 -0.141 -0.093 -0.071 -0.050 -0.009
## u1[3]     0.067  0.028  0.016  0.049  0.066  0.085  0.124
## u1[4]    -0.023  0.028 -0.080 -0.041 -0.023 -0.005  0.033
## u1[5]    -0.019  0.030 -0.079 -0.038 -0.019  0.001  0.040
## u1[6]     0.026  0.025 -0.022  0.010  0.025  0.041  0.075
## u1[7]     0.051  0.027 -0.001  0.033  0.050  0.068  0.107
## u1[8]     0.024  0.027 -0.028  0.007  0.023  0.041  0.079
## u1[9]     0.024  0.027 -0.028  0.006  0.024  0.041  0.077
## u1[10]    -0.029  0.033 -0.095 -0.050 -0.029 -0.008  0.034
## deviance 110.389  8.401 96.576 104.372 109.459 115.604 128.815
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 35.3 and DIC = 145.7
## DIC is an estimate of expected predictive error (lower deviance is better).

rc1.sim <- as.mcmc(rc1.jags)
plot(rc1.sim, auto.layout = FALSE, ask = FALSE)

```





Find the variance of the y

```

g <- rbind(cbind(18.645, 0), cbind(0, 0.005))
r <- diag(0.392, 60)
z <- model.matrix(~-1 + as.factor(variety) + as.factor(variety):moisture, wheat)
Z <- z[, c(1, 11, 2, 12, 3, 13, 4, 14, 5, 15, 6, 16, 7, 17, 8, 18, 9, 19, 10, 20)] # need to rearrange so that the intercepts are at the top
G <- kronecker(diag(1, 10), g)
V <- Z %*% G %*% t(Z) + r

```

* * March 7, 2013 *

To get the slope or intercept for the particular variety, add the fixed effect and the random effect for the variety.

12 Hierarchical Models

$$y = \mu_0 + \mu_s \text{moisture} + e$$

$$\mu_0 \sim \mathcal{N}(\mu_0, \sigma_0^2), \mu_s \sim \mathcal{N}(\mu, \sigma_s^2), e \sim (0, \sigma_e^2)$$

```

wheat <- read.table("data/10randomcoef.dat", header=TRUE)
yield <- wheat$yield
variety <- wheat$variety
moisture <- wheat$moisture
mdl <- 'model {
    for (i in 1:60) {
        yield[i] ~ dnorm(mu[i], 1/s2err)
        mu[i] <- u0[variety[i]] + u1[variety[i]] * moisture[i]
    }
    b0 ~ dnorm(30, 0.001)
    b1 ~ dnorm(0, 0.1)

    for (i in 1:10) {
        u0[i] ~ dnorm(beta0, 1/s2int)
        u1[i] ~ dnorm(beta1, 1/s2slp)
    }

    beta0 ~ dnorm(30, 0.001)
    beta1 ~ dnorm(0, 0.1)

    s2err ~ dgamma(1.1, .5)
    s2int ~ dgamma(1.1, .1)
    s2slp ~ dgamma(1.1, 2)
}
'

writeLines(mdl, 'code/Hierarchical.jags')

jags.data <- c("yield", "variety", "moisture")
jags.params <- c("beta0", "beta1", "u0", "u1", "s2err", "s2int", "s2slp")

hier.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/Hierarchical.jags",
n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 334
##
## Initializing model

hier.jags

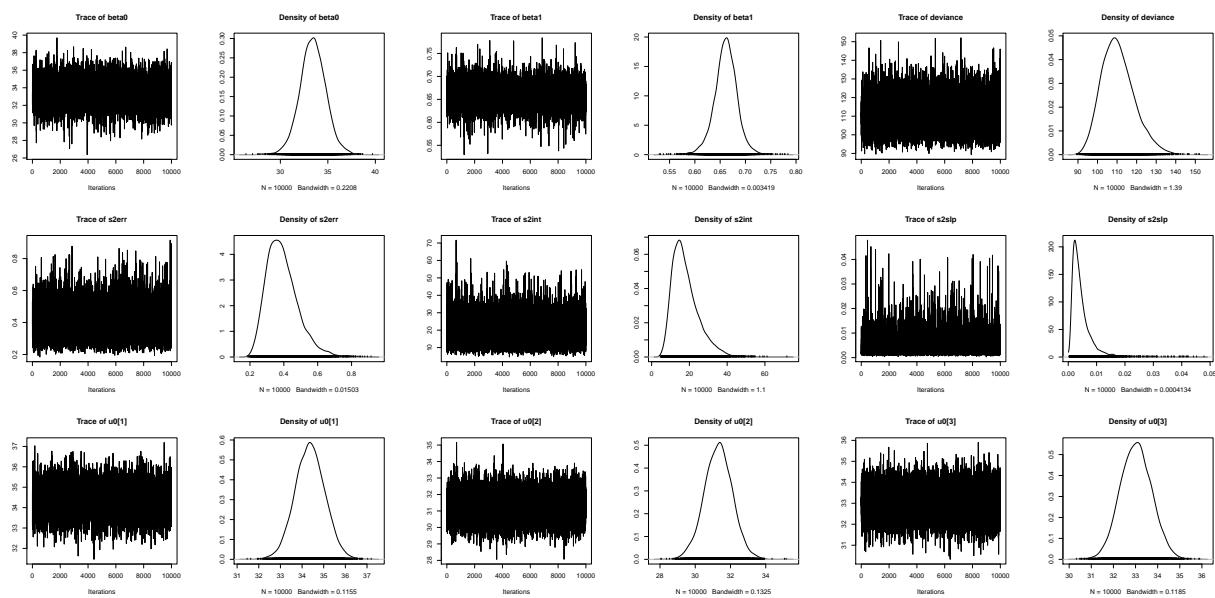
```

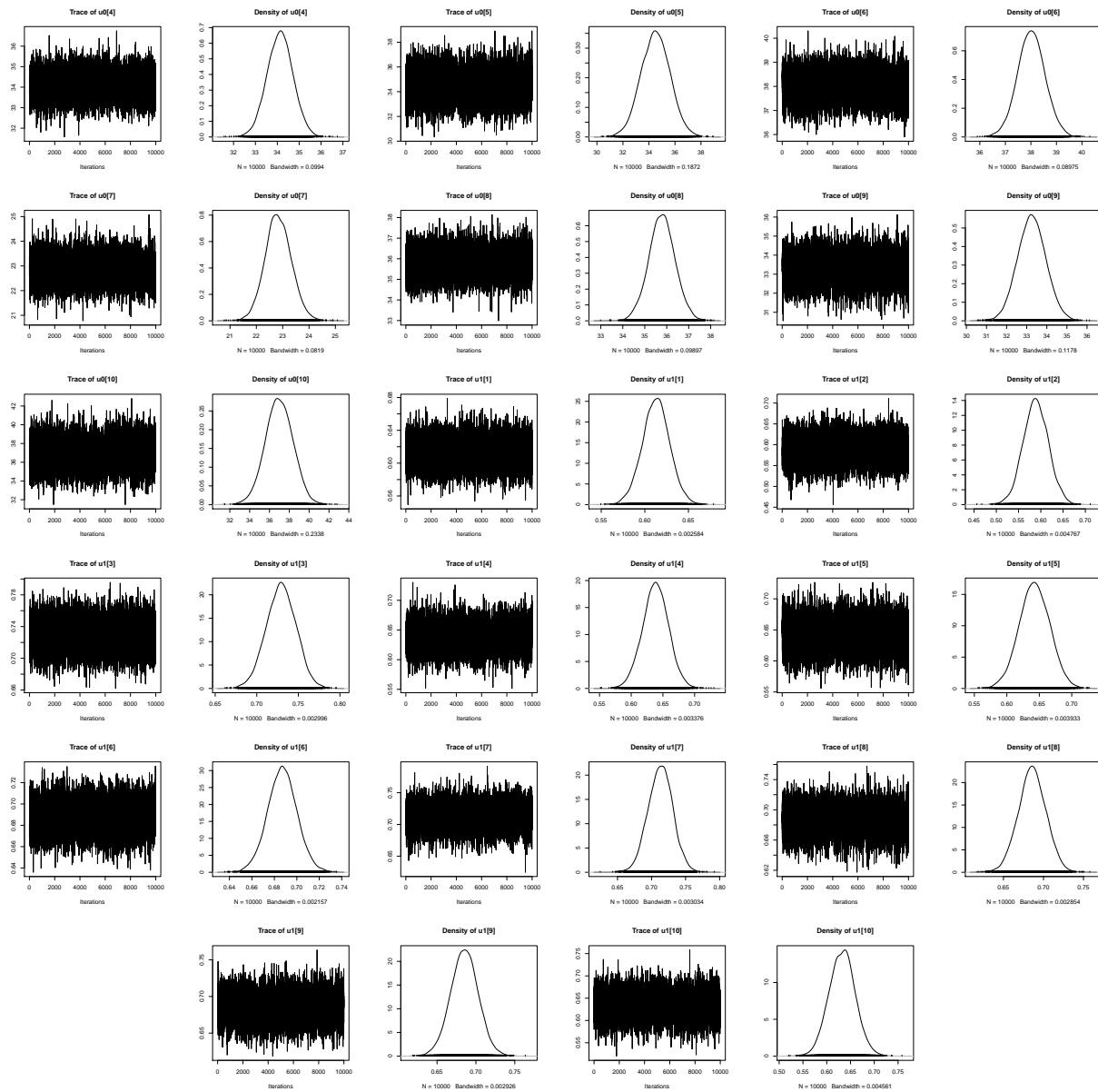
```

## Inference for Bugs model at "code/Hierarchical.jags", fit using jags,
## 1 chains, each with 12000 iterations (first 2000 discarded)
## n.sims = 10000 iterations saved
##          mu.vect    sd.vect      2.5%     25%     50%     75%   97.5%
## beta0    33.422    1.366 30.708 32.548 33.423 34.308 36.162
## beta1     0.661    0.023  0.613  0.648  0.662  0.675  0.706
## s2err     0.391    0.094  0.247  0.324  0.379  0.444  0.613
## s2int    18.019    7.236  8.118 12.861 16.541 21.633 35.702
## s2slp     0.005    0.004  0.001  0.002  0.004  0.006  0.015
## u0[1]    34.367    0.689 33.034 33.905 34.362 34.826 35.724
## u0[2]    31.309    0.794 29.711 30.783 31.322 31.840 32.870
## u0[3]    33.022    0.706 31.659 32.538 33.019 33.497 34.405
## u0[4]    34.126    0.592 32.972 33.725 34.129 34.518 35.286
## u0[5]    34.508    1.114 32.321 33.749 34.511 35.256 36.712
## u0[6]    38.019    0.539 36.958 37.661 38.017 38.376 39.098
## u0[7]    22.834    0.499 21.876 22.506 22.821 23.160 23.828
## u0[8]    35.783    0.599 34.601 35.388 35.780 36.178 36.983
## u0[9]    33.254    0.704 31.901 32.778 33.251 33.717 34.645
## u0[10]   36.978    1.391 34.238 36.047 36.969 37.916 39.672
## u1[1]     0.613    0.016  0.581  0.603  0.613  0.623  0.644
## u1[2]     0.589    0.028  0.533  0.570  0.589  0.608  0.645
## u1[3]     0.729    0.018  0.694  0.717  0.729  0.741  0.764
## u1[4]     0.638    0.021  0.598  0.625  0.638  0.652  0.679
## u1[5]     0.643    0.023  0.596  0.627  0.642  0.659  0.688
## u1[6]     0.687    0.013  0.661  0.678  0.687  0.696  0.713
## u1[7]     0.713    0.018  0.678  0.701  0.713  0.725  0.748
## u1[8]     0.686    0.017  0.652  0.675  0.686  0.697  0.719
## u1[9]     0.685    0.018  0.650  0.674  0.685  0.697  0.720
## u1[10]    0.632    0.027  0.579  0.614  0.633  0.650  0.686
## deviance 110.551   8.430 96.364 104.516 109.746 115.607 129.397
##
## DIC info (using the rule, pD = var(deviance)/2)
## pD = 35.5 and DIC = 146.1
## DIC is an estimate of expected predictive error (lower deviance is better).

hier.sim <- as.mcmc(hier.jags)
plot(hier.sim, auto.layout = inter_plots, ask = inter_plots)

```





Hierarchical models are virtually identical as random coefficients models, but are more pure Bayesian in thought process and methodology.

```

1 %let notedir = Z:\Dropbox\Active\STAT451\Notes\data;
filename data "&notedir./data/10randomcoef.dat";
filename outfile "&notedir./output/2013-03-07.html";
ods listing close;
5
data monkeyRatio;
  infile data;
  input obs variety yield moisture;
run;
10 ODS GRAPHICS on / imagename="2013-03-07-Plots";
  ods html body=outfile (url=none)

```

```

GPATH="&notedir.\figures\>;

15 proc mcmc data=wheat output=wheat_chains nmc=100000 nbi=20000 thin=10
   seed=1234 monitor=(_parms_);
   diag(ess autocorr r1) propcov=quanew;
   parms mu0 30 mul 0 s2 1;
   parms s2int 30 s2slp .1;
   random int ~ normal(mu0, var=s2int) subject=variety monitor=(int_1 int_2);
   random slp ~ normal(mul, var=s2slp) subject=variety;
   prior mu0 ~ normal(30, var=100);
   prior mul ~ normal(0, var=0.1);
   prior s2 ~ dunif(0,3);
   prior s2int ~ dunif(0,50);
   prior s2slp ~ dunif(0,.2);
   mu = int + slp*moisture;
   model yield ~ normal(mu, var=s2);
run;
ods html close;

```

*

*

March 12, 2013

*

*

13 Exam Debriefing

Covergence Diagnostics should have included:

- Traceplots
- Autocorrelation
- Effective sample size
- Raftery-lewis

Model should have been a model with six means and six variances

	Gain			
	1	2	3	
Duration	μ_1	μ_2	μ_3	$\frac{\mu_1+\mu_2+\mu_3}{3}$
	μ_4	μ_5	μ_6	$\frac{\mu_4+\mu_5+\mu_6}{3}$
	$\frac{\mu_1+\mu_4}{2}$	$\frac{\mu_2+\mu_5}{2}$	$\frac{\mu_3+\mu_6}{2}$	

Next exam will probably have

- a random effects and or hierachal model
- Non-normal likelihood

14 Logistic Regression

What we have done so far:

- ANOVA
- Regression
- ANCOVA — Both Categorical and Continuous predictors
- Mixed — More than one source of variability

		Response	
		Categorical	Continuous
Predictors	Categorical	ANOVA	
	Continuous	Regression	

$y = \text{Binomial}(n, \pi)$, with $0 \leq \pi \leq 1$

Odds: $\frac{\pi}{1-\pi}$

Sports “3 to 1” is 3 failures for every one success or $\pi = .25$

$Odds > 0$, 1 should be the midpoint

odds have a skewed distribution

probability	Odds Success-Failure	log(Odds)
		log(S/F)
$\frac{1}{5}$	1 - 1	0
$\frac{1}{5}$	1 - 4	-1.39
$\frac{4}{5}$	4 - 1	1.39

```
chd <- rbind(c(17, 274, 0, 0, 0), c(15, 122, 0, 1, 0), c(7, 59, 0, 0, 1), c(5, 32,
  0, 1, 1), c(1, 8, 1, 9, 9), c(9, 39, 1, 1, 0), c(3, 17, 1, 0, 1), c(14, 58, 1,
  1, 1))
colnames(chd) <- c("CHD", "nRisk", "Cat", "agegrp", "abECG")
tmt <- seq(8)
chd <- cbind(chd, tmt)
CHD <- chd[, 1]
nRisk <- chd[, 2]
tmt <- chd[, 6]
```

Cell means model

```
mdl <- ' model {
  for (i in 1:8) {
    CHD[i] ~ dbin(p[i], nRisk[i]);
    logit(p[i]) <- b[tmt[i]];
    b[i] ~ dnorm(0, .1);
  }
}
writeLines(mdl, 'code/LogitBinomial.jags')

jags.data <- c("CHD", "nRisk", "tmt")
jags.params <- c("b")

# innits <- list(list('p' <- runif(8,0,1)))

logit.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/LogitBinomial.jags",
```

```

n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1
## module glm loaded
## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
##   Graph Size: 42
##
## Initializing model

```

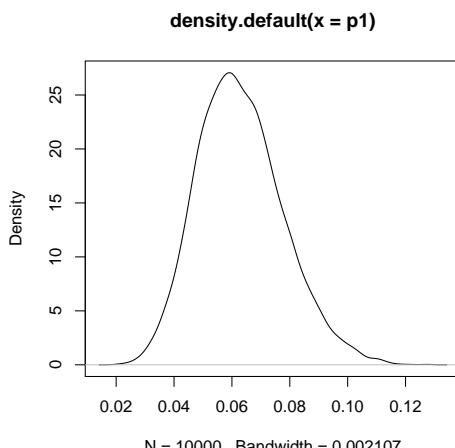
* * March 14, 2013 *

$$\begin{aligned}
\text{logit}(p_i) &= b_i \\
\log\left(\frac{p_i}{1-p_i}\right) &= b_i \\
\frac{p_i}{1-p_i} &= e^{b_i} \\
p_i &= e^{b_i} - pe^{b_i} \\
p_i + p_i e^{b_i} &= e^{b_i} \\
p_i &= \frac{e^{b_i}}{1+e^{b_i}} \frac{e^{-b_i}}{e^{-b_i}} \\
p_i &= \frac{1}{e^{-b_i}+1}
\end{aligned}$$

```

logit.sim <- as.mcmc(logit.jags)
p1 <- 1/(exp(-logit.sim[, 1]) + 1)
plot(density(p1))

```



```

jags.params <- c("b", "p")
logit.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/LogitBinomial.jags",
n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes

```

```
##      Graph Size: 42
##
## Initializing model
logit.sim <- as.mcmc(logit.jags)
```

factorial design

		abEUG	
		Y	O
Cat-N	abECG	N	μ_1
		Y	μ_3
		abEUG	
		Y	O
Cat-Y	abECG	N	μ_5
		Y	μ_7

To find the marginals for cat

$$\text{Cat-n} = \frac{\mu_1 + \mu_2 + \mu_3 + \mu_4}{4} \quad \text{Cat-y} = \frac{\mu_5 + \mu_6 + \mu_7 + \mu_8}{4} \quad (5)$$

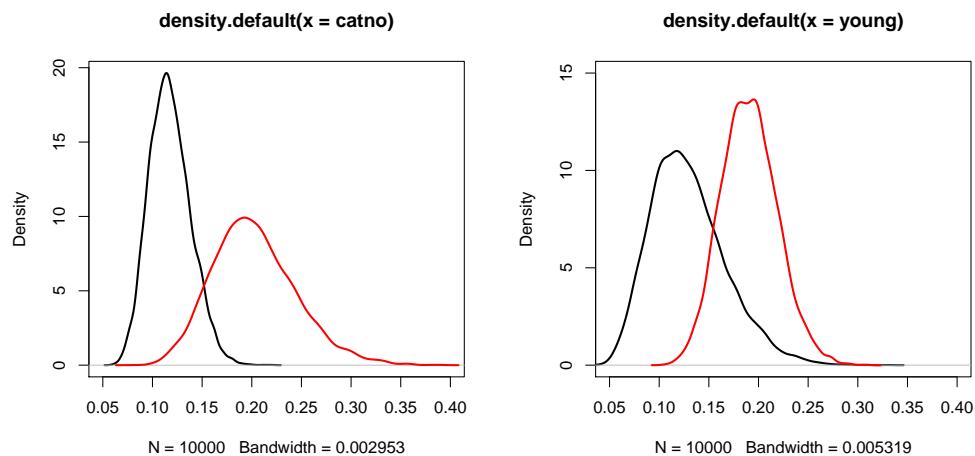
```
p1 <- logit.sim[, 10]
p2 <- logit.sim[, 11]
p3 <- logit.sim[, 12]
p4 <- logit.sim[, 13]
p5 <- logit.sim[, 14]
p6 <- logit.sim[, 15]
p7 <- logit.sim[, 16]
p8 <- logit.sim[, 17]
```

Find the marginals

```
catno <- (p1 + p2 + p3 + p4)/4
catyes <- (p5 + p6 + p7 + p8)/4
plot(density(catno), xlim = c(0.05, 0.4), lwd = 2)
lines(density(catyes), lwd = 2, col = "red")

young <- (p1 + p3 + p5 + p7)/4
old <- (p2 + p4 + p6 + p8)/4
plot(density(young), xlim = c(0.05, 0.4), ylim = c(0, 15), lwd = 2)
lines(density(old), lwd = 2, col = "red")

abNo <- (p1 + p2 + p5 + p6)/4
abYes <- (p3 + p4 + p7 + p8)/4
```



To compare old to young we use an odds ratio or a risk ratio

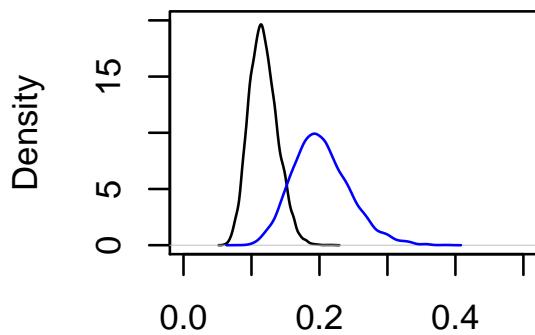
```
ORAge <- (old/(1 - old))/(young/(1 - young))
mean(ORAge > 1)
## [1] 0.9027

RRage <- old/young
mean(RRage > 1)
## [1] 0.9027

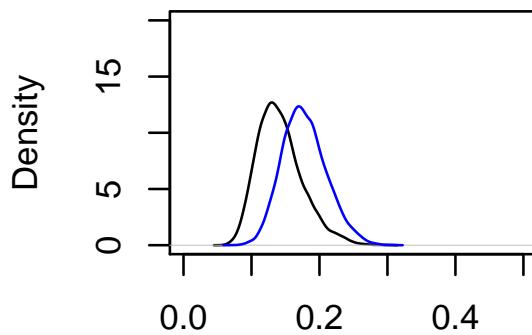
ORab <- (abYes/(1 - abYes))/(abNo/(1 - abNo))
ORCat <- (catyes/(1 - catyes))/(catno/(1 - catno))

mean(ORab)
## [1] 1.402
mean(ORCat)
## [1] 2.001

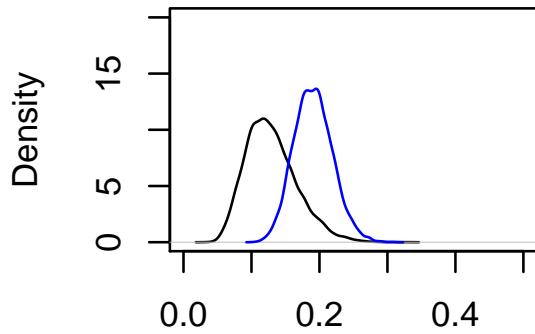
par(mfrow = c(2, 2))
xlims <- c(0, 0.5)
ylims <- c(0, 20)
plot(density(catno), xlim = xlims, ylim = ylims)
lines(density(catyces), col = "blue")
plot(density(abNo), xlim = xlims, ylim = ylims)
lines(density(abYes), col = "blue")
plot(density(young), xlim = xlims, ylim = ylims)
lines(density(old), col = "blue")
par(mfrow = c(1, 1))
```

density.default(x = catno)

N = 10000 Bandwidth = 0.002953

density.default(x = abNo)

N = 10000 Bandwidth = 0.004593

density.default(x = young)

N = 10000 Bandwidth = 0.005319

two way interaction for abnormal ECG and age averaged over catacolamines

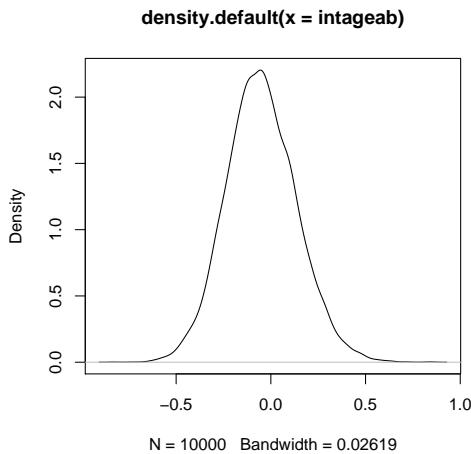
$$\mu_1 + \mu_5 - \mu_2 - \mu_6 - \mu_3 - \mu_7 + \mu_4 + \mu_8 = 0$$

to average across abnormal ECG, rewrite the tables abECG-N

	abEUG	
	Y	O
Cat	μ_1	μ_2
Y	μ_5	μ_6

$$\begin{array}{ccccccc}
 & & abEUG & & & & \\
 & & Y & O & & & \\
 abECG-Y & Cat & N & \mu_3 & \mu_4 & & \\
 & & Y & \mu_7 & \mu_8 & & \\
 & & & & & & \\
 & & \mu_1 + \mu_3 - \mu_2 - \mu_4 - \mu_5 - \mu_7 + \mu_6 + \mu_8 = 0 & & & &
 \end{array}$$

```
intageab <- p1 + p5 - p2 - p6 - p3 - p7 + p4 + p8  
plot(density(intageab))
```



* * March 19, 2013 *

14.1 Logistic Regression In SAS

JAGS freefloating, SAS constrained

Cat-N		Age			
ECG	N	$p_1, .07, .06$		$p_2, .12, .12$	
	Y	$p_3, .10, .12$		$p_4, .16, .16$	

Cat-Y		Age			
ECG	N	$p_5, .12, .15$		$p_6, .20, .23$	
	Y	$p_7, .17, .18$		$p_8, .26, .24$	

SAS:

$$\log \left(\frac{p}{1-p} \right) = \beta_0 + \beta_{\text{cat}} + \beta_{\text{age}} + \beta_{\text{ecg}}$$

$$p = \frac{1}{1 + e^{-x\beta}}$$

$$p_5 = \frac{1}{1 + e^{-\beta_0 - \beta_{\text{cat}}}}$$

This model has 4 degrees of freedom

$$X = \begin{bmatrix} 1 & 0 & 0 & 0 \\ 1 & 0 & 1 & 0 \\ 1 & 0 & 0 & 1 \\ 1 & 0 & 1 & 1 \\ 1 & 1 & 0 & 0 \\ 1 & 1 & 1 & 0 \\ 1 & 1 & 0 & 1 \\ 1 & 1 & 1 & 1 \end{bmatrix}$$

```
1/(1 + exp(2.6263 - 0.6199))
## [1] 0.1185
```

JAGS Model

$$\log\left(\frac{p_i}{1-p_i}\right) = b_i \quad i = \{1, \dots, 8\}$$

this model has 8 degrees of freedom

```
mdl <- 'model {
  for (i in 1:8) {
    CHD[i] ~ dbin(p[i], nRisk[i]);
    logit(p[i]) <- bint + bcat*cat[i] + bage*age[i] + becg*ecg[i];
  }

  bint ~ dnorm(0, .1);
  bcat ~ dnorm(0, .1);
  bage ~ dnorm(0, .1);
  becg ~ dnorm(0, .1);
}

writeLines(mdl, "code/LogitReducedModel.jags")'

CHD <- chd[, 1]
nRisk <- chd[, 2]
tmt <- chd[, 6]
cat <- chd[, 3]
age <- chd[, 4]
ecg <- chd[, 5]

jags.data <- c("CHD", "nRisk", "cat", "age", "ecg")
jags.params <- c("bint", "bcat", "bage", "becg", "p")

set.seed(12)
logit.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/LogitReducedModel.jags",
  n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 70
##
## Initializing model
```

SAS's Deviance is calculate as Dbar – Dmean

computing $pD = \frac{\text{Var}(\text{deviance})}{2}$ is a better way to calculate it in general

14.2 Random Effects

```
seeds <- read.table("data/seednew.dat", stringsAsFactors = FALSE, skip = 1, col.names = c("r",
  "n", "tmt", "nseed", "ntype"))
sum(seeds$n[seeds$tmt == 1])
## [1] 272
sum(seeds$r[seeds$tmt == 1])
## [1] 99
var(seeds$r[seeds$tmt == 1])
## [1] 40.7
```

Is the variability reasonable?

```
mdl <- 'model {
  for (i in 1:21) {
    r[i] ~ dbin(p[i], n[i])
    logit(p[i]) <- b[tmt[i]] + e[i]
  }

  for (i in 1:4) {
    b[i] ~ dnorm(0, 1)
  }

  for (i in 1:21) {
    e[i] ~ dnorm(0, prec)
  }
  s2 ~ dunif(0,2);
  prec <- 1/s2
}

writeLines(mdl, 'code/LogitMixedModel.jags')
mdl <- 'model {
  for (i in 1:21) {
    r[i] ~ dbin(p[i], n[i])
    logit(p[i]) <- b[tmt[i]]
  }

  for (i in 1:4) {
    b[i] ~ dnorm(0, 1)
  }
}
writeLines(mdl, 'code/Logit09Simplified.jags')

r <- seeds[, 1]
n <- seeds[, 2]
tmt <- seeds[, 3]
jags.data <- c("r", "n", "tmt")
jags.params <- c("b", "e")

seed.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/LogitMixedModel.jags",
  n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

## Compiling model graph
##   Resolving undeclared variables
##   Allocating nodes
##   Graph Size: 135
```

```

##  

## Initializing model  
  

jags.params <- c("b")  

seedsimple.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/Logit09Simplified.jags",  

n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)  
  

## Compiling model graph  

##   Resolving undeclared variables  

##   Allocating nodes  

##   Graph Size: 73  

##  

## Initializing model  
  

seed.jags  
  

## Inference for Bugs model at "code/LogitMixedModel.jags", fit using jags,  

## 1 chains, each with 12000 iterations (first 2000 discarded)  

## n.sims = 10000 iterations saved  

##          mu.vect sd.vect    2.5%     25%     50%     75%   97.5%  

## b[1]      -0.522  0.221 -0.948 -0.665 -0.525 -0.385 -0.068  

## b[2]       0.816  0.221  0.384  0.676  0.808  0.953  1.270  

## b[3]      -0.457  0.268 -1.008 -0.625 -0.446 -0.281  0.041  

## b[4]       0.017  0.254 -0.507 -0.143  0.028  0.186  0.503  

## e[1]      -0.287  0.298 -0.937 -0.466 -0.264 -0.082  0.232  

## e[2]      -0.016  0.265 -0.559 -0.177 -0.010  0.149  0.506  

## e[3]      -0.271  0.268 -0.846 -0.436 -0.252 -0.086  0.207  

## e[4]       0.338  0.289 -0.174  0.136  0.318  0.518  0.946  

## e[5]       0.144  0.282 -0.401 -0.035  0.134  0.314  0.728  

## e[6]       0.113  0.381 -0.585 -0.128  0.086  0.329  0.948  

## e[7]       0.068  0.261 -0.442 -0.099  0.061  0.230  0.605  

## e[8]       0.233  0.276 -0.264  0.045  0.213  0.403  0.820  

## e[9]      -0.171  0.275 -0.748 -0.347 -0.157  0.012  0.335  

## e[10]     -0.217  0.264 -0.767 -0.383 -0.204 -0.040  0.278  

## e[11]     0.128  0.350 -0.517 -0.093  0.103  0.326  0.919  

## e[12]     0.170  0.338 -0.451 -0.050  0.144  0.373  0.913  

## e[13]     -0.120  0.310 -0.766 -0.312 -0.108  0.073  0.494  

## e[14]     -0.210  0.319 -0.904 -0.406 -0.187 -0.002  0.375  

## e[15]     0.292  0.311 -0.258  0.079  0.265  0.480  0.978  

## e[16]     -0.221  0.413 -1.168 -0.433 -0.172  0.042  0.459  

## e[17]     -0.314  0.368 -1.148 -0.526 -0.275 -0.068  0.315  

## e[18]     0.088  0.292 -0.465 -0.098  0.076  0.266  0.706  

## e[19]     -0.010  0.296 -0.595 -0.192 -0.016  0.171  0.601  

## e[20]     0.302  0.301 -0.237  0.097  0.282  0.484  0.956  

## e[21]     -0.073  0.363 -0.828 -0.284 -0.062  0.148  0.641  

## deviance  98.860  5.637 88.953 94.822 98.517 102.524 110.858  

##  

## DIC info (using the rule, pD = var(deviance)/2)  

## pD = 15.9 and DIC = 114.7  

## DIC is an estimate of expected predictive error (lower deviance is better).  
  

seedsimple.jags  
  

## Inference for Bugs model at "code/Logit09Simplified.jags", fit using jags,  

## 1 chains, each with 12000 iterations (first 2000 discarded)  

## n.sims = 10000 iterations saved  

##          mu.vect sd.vect    2.5%     25%     50%     75%   97.5%  

## b[1]      -0.552  0.126 -0.800 -0.637 -0.551 -0.467 -0.308  

## b[2]       0.796  0.125  0.556  0.712  0.795  0.879  1.043  

## b[3]      -0.406  0.183 -0.766 -0.529 -0.404 -0.283 -0.056  

## b[4]       0.125  0.168 -0.207  0.010  0.126  0.238  0.454  

## deviance 111.575  2.873 108.077 109.515 110.875 112.916 118.965  

##  

## DIC info (using the rule, pD = var(deviance)/2)  

## pD = 4.1 and DIC = 115.7  

## DIC is an estimate of expected predictive error (lower deviance is better).

```

* * March 21, 2013 *

```

1 %let hwdir=Z:/Dropbox/Active/STAT451/Notes;
  filename data "&hwdir./pbib.dat";
  filename chains "&hwdir./MCMCChains.csv";
  filename html "&hwdir./MCMC_Output.html";
5
ods html close;
ODS GRAPHICS on / imangename="MCMCPlots";
ods html body=html (url=none)
      GPATH="&hwdir/figure/";
10
data seeds;
  infile "&hwdir./data/seeds.dat" firstobs=2;
  input seed $ type $ r n;
  if seed='a75' and type='bean' then tmt=1;
  if seed='a75' and type='cuc' then tmt=2;
15
  if seed='a73' and type='bean' then tmt=3;
  if seed='a73' and type='cuc' then tmt=4;
  plate=_N_;
run;
20
/* Our attempt */
proc mcmc data=seeds seed=12 nmc=200000 nbi=2000 thin=20 propcov=quanew
            statistics=(summary interval) diag=(rl ess) moniter=
            -parms_ p)
            outpost=chdout dic;
25
  array b[4];
  array e[21];
  array p[4];
  parms b p e s2;
30
  prior b: ~ normal(0, prec=1);
  random e ~ normal(0, var=s2) subject=plate;
  prior s2 ~ uniform(0,2);

  p[tmt] = logistic(b[tmt]);
35
  model chd ~ binomial(nrisk, p[tmt]);
run;

/* His Code */
40
proc mcmc data=seeds seed=12 nmc=200000 nbi=2000 thin=20 outpost=chdout
            statistics=(summary interval) diag=(rl autocorr ess)
            monitor=( p alpha s2plate )
            propcov=quanew dic;
  array alpha[4] alpha1-alpha4;
  array p[4] p1-p4;
45
  parms alpha: 0;
  parms s2plate 1;
  prior alpha: ~ normal(0, var=10); /* Fixed Effect */
  random b~ normal(0, var=s2plate) subject=plate monitor=(b); /* Random
    Effect: has hyper-prior */
  prior s2plate ~ gamma(1.5, scale=2);
50

```

```

p[tmt] = logistic(alpha[tmt]+b) ;
model r ~ binomial(n, p[tmt]) ;
run ;

```

15 Poisson Likelihood

11epi.dat Epilepsy

- “pid” — Patient ID
- “scount” — Seizure Count
- “tmt” — Treatment
- “base” — Base Effect
- “age” — Age
- “visit”

S is the number of seizures (scount)

```

epi <- read.table("data/11epi.dat", col.names = c("pid", "scount", "tmt", "base",
                                                 "age", "visit"))
N <- nrow(epi)
pid <- epi$pid
N_pid <- length(unique(pid))
scount <- epi$scount
tmt <- epi$tmt
base <- epi$base
age <- epi$age
visit <- epi$visit
epi$indic4 <- ifelse(visit == 4, 1, 0)
indic4 <- epi$indic4

```

$$S_i \sim \text{Poisson}(\lambda)$$

$$\lambda = a_0 + a_{\text{age}} * \text{age}_i + a_{\text{tmt}} * \text{tmt}_i + a_{\text{base}} * \text{base}_i$$

* * March 26, 2013 *

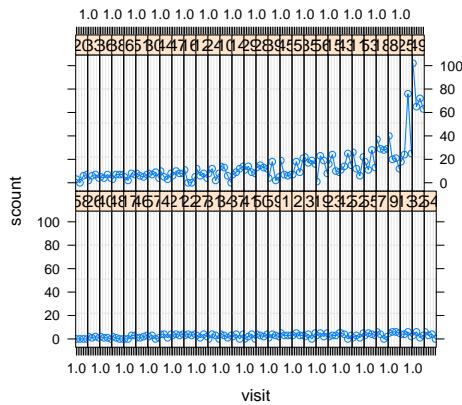
Plotting data

```

library(nlme)
g_epi <- groupedData(scount ~ visit | pid, epi)
plot(g_epi)
plot()

## Error: argument "x" is missing, with no default

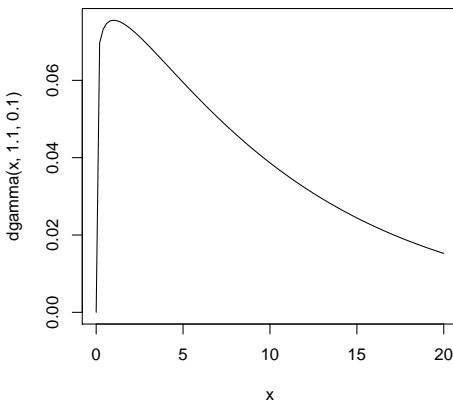
```



Things learned from this plot:

- There doesn't seem to be an effect due to time
- Subjects matter, so we need a random

```
curve(dgamma(x, 1.1, 0.1), from = 0, to = 20)
```



```
mdl <- "model {\nfor (i in 1:N) {\\nscount[i] ~ dpois(lambda[i])\\nlambda[i] <- a_0 + a_age*age[i] + a_base*base[i] + a_tmt*tmt[i]\nwriteLines(mdl, \"code/EPI_Interaction.jags\")\n\njags.data <- c(\"N\", \"pid\", \"scount\", \"age\", \"tmt\", \"base\")\njags.params <- c(\"a_0\", \"a_age\", \"a_base\", \"a_tmt\", \"int_age_base\", \"int_age_tmt\", \"int_base_tmt\", \"ll\")\nset.seed(3245)\nepi.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = \"code/EPI_Interaction.jags\", n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)\n\n## module glm loaded\n\n## Error:\n## Error parsing model file:\n## syntax error on line 6 near \"\"\n\nepi.jags\n\n## Error: object 'epi.jags' not found"
```

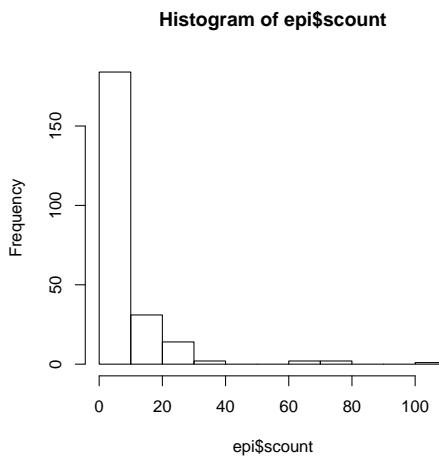
```

epi.sim <- as.mcmc(epi.jags)
## Error: object 'epi.jags' not found
plot(epi.sim, auto.layout = inter, ask = inter)
## Error: object 'epi.sim' not found
autocorr.plot(epi.sim, auto.layout = inter, ask = inter)
## Error: object 'inter' not found
rafthery.diag(epi.sim)
## Error: object 'epi.sim' not found

```

Posterior Predictive Plots

```
hist(epi$scount)
```



We can't do general poseritor predictives so let's look at particular values

```

hist(epi$scount[epi$tmt == 1 & epi$age >= 19 & epi$age <= 23 & 30 <= epi$base & epi$base <=
60])
plot(density(epi$age))
curve(dnorm(x, 30, 8), add = TRUE, col = "red")
plot(density(epi$base))
curve(dgamma(1.1, scale = 30), add = T, color = "red")

## Error: 'expr' must be a function, or a call or an expression containing 'x'

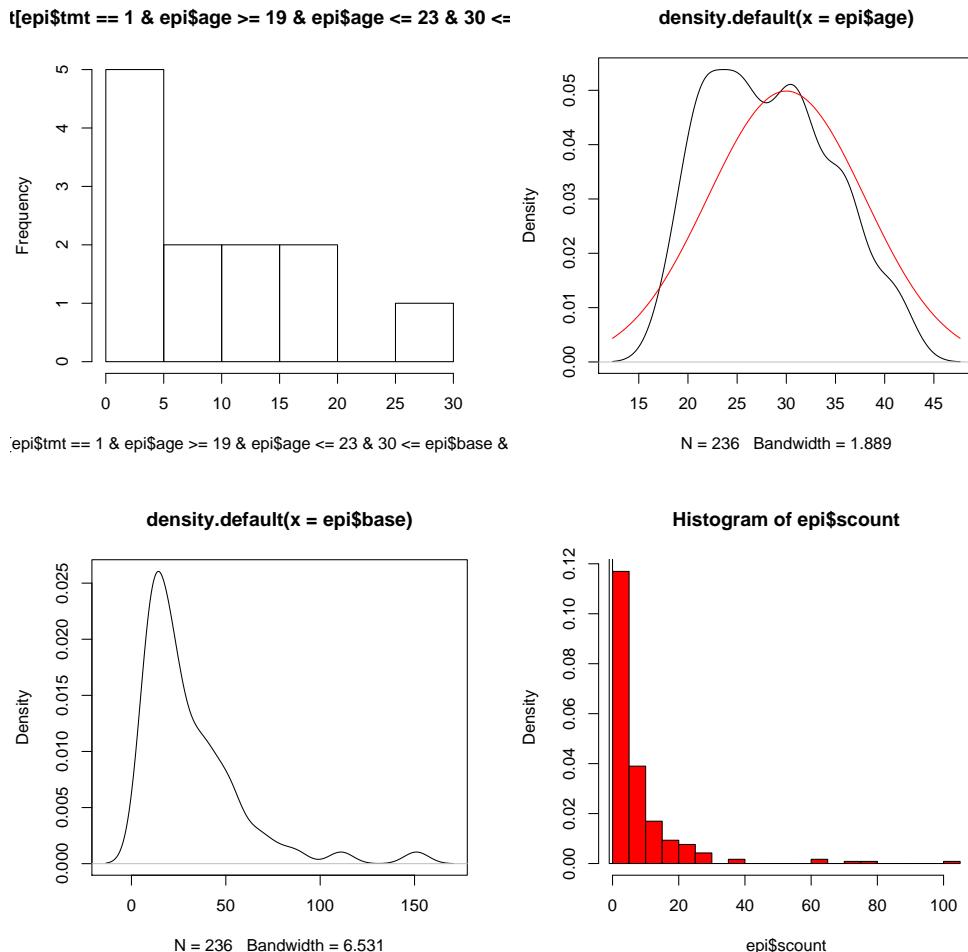
people <- cbind(age = rnorm(iter, 30, 8), )

## Error: object 'iter' not found

iter <- 10^4
pp <- numeric(iter)
for (i in 1:iter) {
  rage <- rnorm(1, 30, 8)
  rbase <- rgamma(1, 1.1, scale = 30)
  rtmt <- rbinom(1, 1, 0.5)
  x1 <- epi.sim[i, 1] + epi.sim[i, 2] * rage + epi.sim[, 3] * rbase + epi.sim[, 4] * rtmt +
    epi.sim[i, 6] * rage * rbase + epi.sim[i, 7] * rage * rtmt +
    epi.sim[i, 8] * rbase * rtmt + rexp(1, epi.sim[i, 9])
  pp[i] <- rpois(1, x1)
}
## Error: object 'epi.sim' not found

```

```
hist(epi$scount, breaks = 20, freq = F, col = "red")
hist(pp, breaks = 20, freq = F, add = T)
```



15.1 Log Poisson

* * March 28, 2013 *

```
1 count[i] <- dpois(lambda[i])
  log(lambda[i]) <- a_0 + a_age*log(age[i]) + a_base*log(base[i]/4) +
    a_tmt*tmt[i] +
    a_base_tmt*log(base[i]/4)+tmt[i] + a_4*indic4[i] +
    u[pid[i]] + aobs[i]
5 aobs[i] ~ dnorm(0,1/s2obs)
```

```
lbase4 <- log(base/4)
lage <- log(age)
mdl <-'model {
  for (i in 1:N) {
    scount[i] ~ dpois(lambda[i])
    log(lambda[i]) <- a_0 + a_age*lage[i] + a_base*lbase4[i] + a_tmt*tmt[i] +
      a_base_tmt*lbase4[i]*tmt[i] + a_4*indic4[i] +
```

```

        u_sub[pid[i]] + u_obs[i]
    u_obs[i] ~ dnorm(0, 1/s2obs)
}
for (i in 1:N_pid) {
    u_sub[i] ~ dnorm(0, 1/s2sub)
}

a_0      ~ dnorm(0, 0.0001)
a_base   ~ dnorm(0, 0.0001)
a_tmt    ~ dnorm(0, 0.0001)
a_base_tmt ~ dnorm(0, 0.0001)
a_age    ~ dnorm(0, 0.0001)
a_4      ~ dnorm(0, 0.0001)
s2obs    ~ dgamma(1.1, 1)
s2sub    ~ dgamma(1.1, 1)
}
'
writeLines(mdl, 'code/EPIPAPER.jags')

jags.data <- c("N", "N_pid", "pid", "scount", "lage", "tmt", "lbase4", "indic4")
jags.params <- c("a_0", "a_age", "a_base", "a_tmt", "a_base_tmt", "a_4", "s2obs",
                 "s2sub")

set.seed(3245)
epiP.jags <- jags(data = jags.data, parameters.to.save = jags.params, model.file = "code/EPIPAPER.jags",
                     n.iter = 12000, n.burnin = 2000, n.chains = 1, n.thin = 1)

## module glm loaded

## Compiling model graph
## Resolving undeclared variables
## Allocating nodes
## Graph Size: 2306
##
## Initializing model

```

If you have correlated X 's then you shift the X

* * April 2, 2013 *

15.2 Complex Model Posterior Predictive

```

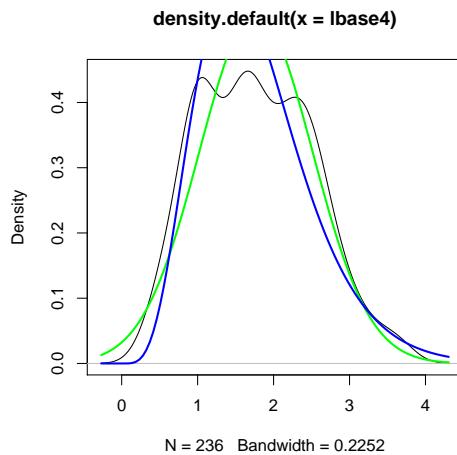
epiP.sim <- as.mcmc(epiP.jags)
mCol <- function(name, data = epiP.sim) {
  col <- which(attr(data, "dimnames")[[2]] == name)
  data[, col]
}
a_0 <- mCol("a_0")
a_age <- mCol("a_age")
a_base <- mCol("a_base")
a_tmt <- mCol("a_tmt")
a_base_tmt <- mCol("a_base_tmt")
a_4 <- mCol("a_4")
s2sub <- mCol("s2sub")
s2obs <- mCol("s2obs")

```

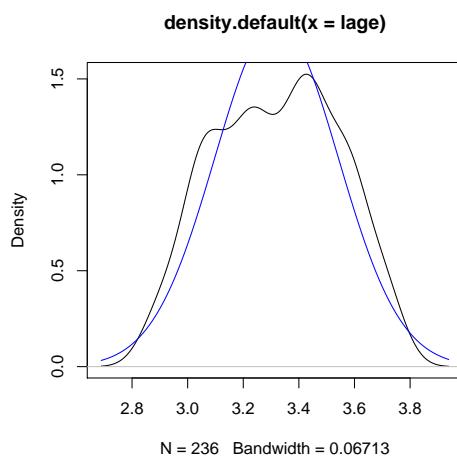
We might be able to sample better from age with a gamma

$$\begin{aligned}\mu &= \alpha\beta \\ v &= \alpha\beta^2 \\ \beta &= \frac{v}{\mu} = .31\end{aligned}$$

```
b <- var(lbase4)/mean(lbase4)
a <- mean(lbase4)/b
plot(density(lbase4))
curve(dnorm(x, mean(lbase4), sd(lbase4)), add = TRUE, col = "green", lwd = 2)
curve(dgamma(x, a, scale = b), add = TRUE, col = "blue", lwd = 2)
```



```
plot(density(lage))
curve(dnorm(x, mean(lage), sd(lage)), add = TRUE, col = "blue")
```

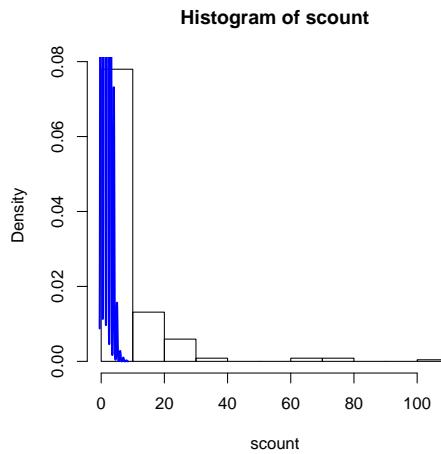


```
lmu <- numeric(10000)
mean_lage <- mean(lage)
sd_lage <- sd(lage)
for (i in 1:10000) {
  lb <- rgamma(1, a, scale = b)
```

```

tt <- rbinom(1, 1, 0.5)
la <- rnorm(1, mean_lage, sd_lage)
i4 <- rnorm(1, 1, 0.25)
lmu <- a_0[i] + a_age[i] * la + a_base[i] * lb + a_tmt[i] * tt + a_base_tmt[i] *
    lb * tt + a_4[i] * i4 + rnorm(1, 0, sqrt(s2sub[i])) + rnorm(1, 0, sqrt(s2obs[i]))
}
pp <- rpois(10000, exp(lmu))
hist(scount, freq = FALSE)
lines(density(pp), col = "blue", lwd = 2)

```



15.3 Non Diagonal Covariance

$$\begin{aligned}
 y &= x\beta + Zu + e \\
 e &\sim \mathcal{N}(0, R) \\
 u &\sim \mathcal{N}(0, G) \\
 V(y) &= ZGZ + R
 \end{aligned}$$

Compound Symmetric Structure

$$\begin{bmatrix}
 \sigma_e^2 + \sigma_u^2 & \sigma_u^2 & \sigma_u^2 & \sigma_u^2 \\
 \sigma_u^2 & \sigma_e^2 + \sigma_u^2 & \sigma_u^2 & \sigma_u^2 \\
 \sigma_u^2 & \sigma_u^2 & \sigma_e^2 + \sigma_u^2 & \sigma_u^2 \\
 \sigma_u^2 & \sigma_u^2 & \sigma_u^2 & \sigma_e^2 + \sigma_u^2
 \end{bmatrix}$$

$$\rho = \frac{\text{Cov}(1, 2)}{\sqrt{\text{Var}(1)\text{Var}(2)}} = \frac{\sigma_u^2}{\sqrt{\sigma_e^2\sigma_u^2}}$$

Three Groups in the experiment

- Control

- Reps
- Weight

ind1 = (50,50, 40, 60, 55, 60, 45)

Auto Regressive (AR(1))

$$V(Sub) = \begin{pmatrix} 1 & \rho & \rho^2 & \rho^3 & \rho^4 & \rho^5 & \rho^6 \\ \rho & 1 & \rho & \rho^2 & \rho^3 & \rho^4 & \rho^5 \\ \rho^2 & \rho & 1 & \rho & \rho^2 & \rho^3 & \rho^4 \\ \rho^3 & \rho^2 & \rho & 1 & \rho & \rho^2 & \rho^3 \\ \rho^4 & \rho^3 & \rho^2 & \rho & 1 & \rho & \rho^2 \\ \rho^5 & \rho^4 & \rho^3 & \rho^2 & \rho & 1 & \rho \\ \rho^6 & \rho^5 & \rho^4 & \rho^3 & \rho^2 & \rho & 1 \end{pmatrix}$$

```
weight <- read.table("data/weight.dat", header = TRUE)
weight$tz <- weight$time - 1
weight$st <- paste0(weight$tmt, "/", weight$subject)
weight$subj <- c(rep(1:20, times = 7), rep(21:36, times = 7), rep(37:57, times = 7))
weight$subj <- as.numeric(as.factor(paste(weight$tmt, sprintf("%02s", weight$subject))))
```

*

*

April 4, 2013

*

*

15.4 Exam 2 Debriefing

The first problem could be done with a simple beta binomial conjugate prior model.

You can't add an error term where you have only a single observation per subject.

To make it random, go back to the basic beta-binomial and then put priors and alpha and beta.

When you say $b_i + e_i$, and $e_i \sim \mathcal{N}(0, \sigma^2)$ is a noise term, we say that it is centered at 0, because it shouldn't be biased.