

# Simple Regression

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# Introduction

## The problem

Let  $x$  and  $y$  be two statistical variables:

$$x = \begin{bmatrix} x_1 \\ \vdots \\ x_n \end{bmatrix} \quad \text{and} \quad y = \begin{bmatrix} y_1 \\ \vdots \\ y_n \end{bmatrix}.$$

- The correlation problem, in general, consists of looking for a possible relationship between  $x$  and  $y$ :  $y = f(x) + \text{"residual"}$  that makes the residual part the **smallest** possible.
- The variable  $x$  is called "predictor" or the "independent" variable,  $y$  is called "response" or "dependent" variable.
- If we look for "**linear**" functional " $f$ ", the correlation is hence of linear type.

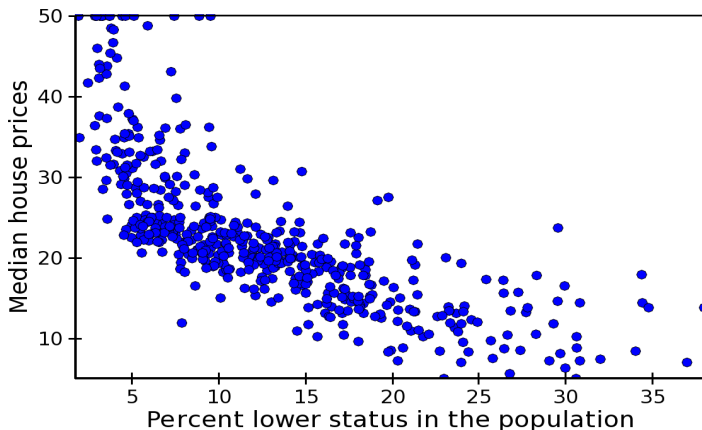
## Introduction

### The scattergram

The scattergram or "scatter-plot" is simply the set of points of coordinates

$$\{(x_i, y_i) : i = 1, \dots, n\}.$$

It gives a very useful a priori glance:

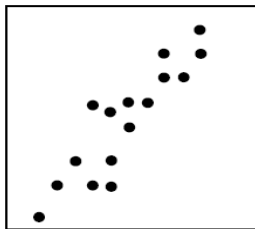


# Introduction

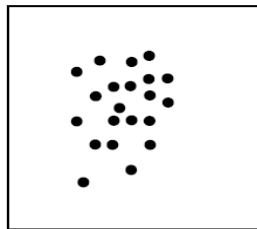
## Different types of correlation



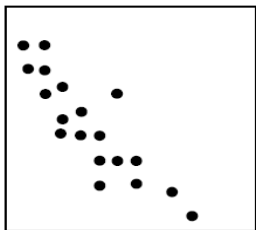
Strong positive correlation



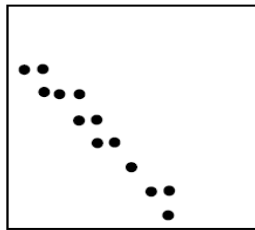
Moderate positive correlation



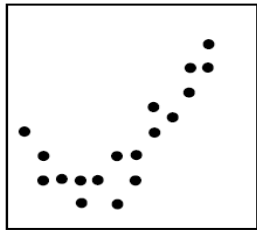
No correlation



Moderate negative correlation



Strong negative correlation



Curvilinear relationship

## Least square method

### The least square distance

Suppose our problem **reduced** to the linear case, then it can be reformulated as follows:

We look for  $\beta$  and  $\alpha$  such that if  $y_i^* = \beta x_i + \alpha$  then we have

$$\begin{aligned} \overline{y^*} &= \overline{y} \\ \sum_{i=1}^n (y_i^* - y_i)^2 &\text{ is minimal.} \end{aligned}$$

- 1 The line  $s = \beta t + \alpha$  is called "least square line" or "regression line".
- 2  $\beta$  is called "slope" and  $\alpha$  the "intercept" of the regression line.

## Least square method

### Vectorial formulation

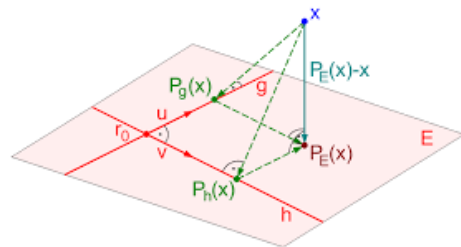
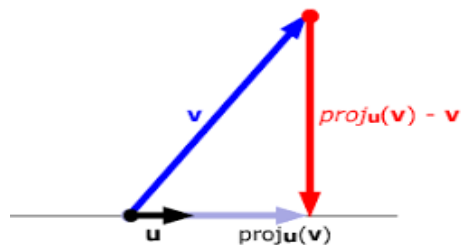
We look for a vector  $y^* \in \mathbb{R}^n$  spanned by  $\{x, \mathbf{1}\}$  i.e find out  $\beta$  and  $\alpha$  such that  $y^* = \beta x + \alpha \mathbf{1}$  minimizes the Euclidian norm  $\|y - y^*\|_n^2$ .

#### Theorem

$\|y - y^*\|_n^2$  is minimal if and only if

$y^* = \text{Orthogonal Projection of } y \text{ on the subspace } \text{span}\{x, \mathbf{1}\}.$

## Least square method





### Consequences

- $y - y^*$  is orthogonal to  $y^*$ .

$$y = y^* \oplus^\perp y - y^* \Rightarrow \|y\|_n^2 = \|y^*\|_n^2 + \|y - y^*\|_n^2 \quad (1)$$

- since  $\bar{y} = \overline{y^*}$  and (1) we obtain  
 $\Rightarrow \|y - \bar{y}\|_n^2 = \|y^* - \bar{y}\|_n^2 + \|y - y^*\|_n^2$
- *Total sum of squares (tss) = fitted sum of squares (fss) + residual sum of squares (rss)*
- $\Rightarrow s_y^2 = s_{y^*}^2 + s_{y-y^*}^2$   
*Total variance = Explained variance + Residual variance*

## Least square method

### Determination of the regression slope and intercept

We look for  $\beta$  and  $\alpha$  so that  $y - y^*$  is orthogonal to  $\text{span}\{x, \mathbf{1}\}$  i.e

i. On one hand

$$\begin{aligned}y - y^* \perp x &\Rightarrow {}^t x(y - y^*) = 0 \\&\Rightarrow {}^t xy - \beta {}^t xx + \alpha n\bar{x} = 0 \\&\Rightarrow |x|^2 \beta - n\bar{x}\alpha = {}^t xy\end{aligned}\tag{2}$$

ii. on the other hand

$$\begin{aligned}y - y^* \perp \mathbf{1} &\Rightarrow {}^t \mathbf{1}(y - y^*) = 0 \\&\Rightarrow n\bar{y} - n\beta\bar{x} + n\alpha = 0 \\&\Rightarrow \alpha = \beta\bar{x} - \bar{y}\end{aligned}\tag{3}$$

## Least square method

Finally

Combining (2) and (3)

$$\begin{cases} \beta = \frac{\overline{x * y} - \bar{x} * \bar{y}}{\overline{x^2} - \bar{x}^2} = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2} \\ \alpha = \bar{y} - \beta \bar{x} \end{cases}$$

## Interpretation

# Linear Determination coefficient vs Correlation Coefficient

- 1 Set the **Linear** Determination Coefficient as

$$L^2_{y|x} = \frac{fss}{tss} = \frac{\sum_i (y_i^* - \bar{y})^2}{\sum_i (y_i - \bar{y})^2} \in [0, 1]$$

- 2 It represents the strongness/strength of the linear correlation between  $y$  and  $x$ .
- 3 It can be read as the the rate of  $y$  **explained linearly** by  $x$ .

# Linear Determination Coefficient vs Correlation Coefficient

## Definition

The linear correlation coefficient of the statistical variables  $x$  and  $y$  is given by

$$r_{xy} = \frac{\text{Cov}(x, y)}{\sqrt{\sum_i (x_i - \bar{x})^2 \sum_i (y_i - \bar{y})^2}} \quad (4)$$

$$= \frac{\sum_i (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_i (x_i - \bar{x})^2 \sum_i (y_i - \bar{y})^2}} \quad (5)$$

$$= \frac{n}{n-1} \frac{\overline{x * y} - \bar{x} * \bar{y}}{s_x * s_y} \quad (6)$$

$$= \frac{\sum_i x_i y_i - \frac{(\sum_i x_i)(\sum_i y_i)}{n}}{\sqrt{\left(\sum_i x_i^2 - \frac{(\sum_i x_i)^2}{n}\right) \left(\sum_i y_i^2 - \frac{(\sum_i y_i)^2}{n}\right)}} \quad (7)$$

## Interpretation

### Very Important Remark

- ① The linear correlation coefficient can be seen as the **cosine** of the angle formed by the two vectors of  $\mathbb{R}^n$ ,  $x - \bar{x} * \mathbf{1}$  and  $y - \bar{y} * \mathbf{1}$ :

$$\begin{aligned}\cos(\widehat{x - \bar{x} * \mathbf{1}, y - \bar{y} * \mathbf{1}}) &= \frac{t(x - \bar{x} * \mathbf{1})(y - \bar{y} * \mathbf{1})}{\|x - \bar{x} * \mathbf{1}\| \|y - \bar{y} * \mathbf{1}\|} \\ &= \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2 \sum_{i=1}^n (y_i - \bar{y})^2}}\end{aligned}$$

- ② why?

$$r_{xy}^2 = L_{y|x}^2. \quad (8)$$

## Interpretation

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$$\begin{aligned}\cos(\widehat{x - \bar{x} * \mathbf{1}, y - \bar{y} * \mathbf{1}}) &= \frac{(x - \bar{x} * \mathbf{1})(y - \bar{y} * \mathbf{1})}{\|x - \bar{x} * \mathbf{1}\| \|y - \bar{y} * \mathbf{1}\|} \\ &= \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2 \sum_{i=1}^n (y_i - \bar{y})^2}}\end{aligned}$$

- ② why?

$$r_{xy}^2 = L_{y|x}^2. \quad (8)$$

- ① We have looked only for the best **linear** function of "x" that describes "y", it is not possible to look for all possible functions of "x", since we only have samples.

## Interpretation

### Consequences

- ① Since the correlation coefficient is a cosine it satisfies:

$$|r_{xy}| = \sqrt{L_{xy}^2} \leq 1.$$

- ② If  $|r_{xy}| = 1$  it means that we have a perfect linear correlation:

$$y_i = \beta x_i + \alpha, \text{ for all } i = 1, \dots, n.$$

- ③ If  $r_{xy} = 0$  it means that there is no linear relationship between  $x$  and  $y$ .

- ④ In all cases we have

$$\beta = r_{xy} \times \sqrt{\frac{\sum_{i=1}^n (y_i - \bar{y})^2}{\sum_{i=1}^n (x_i - \bar{x})^2}}$$



## Interpretation

We already know

- 1 If  $|r_{xy}| \in [0, 25\%[$ , there is no linear correlation !
- 2 If  $|r_{xy}| \in [25\%, 50\%[$ , there is a moderate linear correlation.
- 3 If  $|r_{xy}| \in [50\%, 75\%[$ , there is a fair linear correlation.
- 4 If  $|r_{xy}| \in [75\%, 100\%[$ , there is good linear correlation.

## Inferential study

Suppose that the two statistical series  $x$  and  $y$  came from two given random variables  $X$  and  $Y$  and let  $\{X_i : i = 1 \cdots n\}$  and  $\{Y_i : i = 1 \cdots n\}$  be two random samples of  $X$  and  $Y$ . Denote by

$$TSS = \sum_{i=1}^n (Y_i - \bar{Y})^2, \quad FSS = \sum_{i=1}^n (Y_i^* - \bar{Y})^2$$

and

$$RSS = \sum_{i=1}^n (Y_i - Y_i^*)^2$$

and consider the following statistic:

$$F = \frac{FSS}{RSS} (n - 2).$$

### A theoretical result: The linear Gaussian model

#### Theorem

Suppose that  $X$  and  $Y$  are normally distributed and that they are independent. Then

$$\frac{FSS}{RSS}(n-2) = \mathcal{F}(1, n-2)$$

where  $\mathcal{F}(1, n-2)$  is the Fisher distribution of degrees of freedom 1 and  $n-2$ .

#### Proposition and Definition

If  $A \sim \chi_{d_1}^2$  and  $B \sim \chi_{d_2}^2$  and are independent then

$$\frac{A/d_1}{B/d_2} \sim \mathcal{F}(d_1, d_2).$$

### So what: How to use the latter Theorem ?

We consider the test: ( $H_0$ ):  $\beta = 0$  (which means that there is no linear relationship between  $X$  and  $Y$ ) against ( $H_1$ ):  $\beta \neq 0$ :

- we reject the hypothesis ( $H_0$ ) and accept ( $H_1$ ) with a confidence level of 95%, if:

$$f^* = \frac{f_{ss}}{r_{ss}}(n - 2)$$

$$\mathbb{P}\{\mathcal{F}(1, n - 2) \leq f^*\} = 95\%$$

- or equivalently (and in general) we reject  $H_0$  and accept  $H_1$  when

$$p - value = \mathbb{P}\{\mathcal{F}(1, n - 2) \geq f^*\} \ll 1$$

### Confidence interval for the regression line

Let  $T_{n-2,\gamma/2}$  be such that  $\mathbb{P}\{|\mathcal{T}_{n-2}| > T_{n-2,\gamma/2}\} = \gamma/2$ .

- ① A confidence interval of level  $1 - \gamma$ , for the slope  $B$  is

$$\beta \pm T_{n-2,\gamma/2} \times \frac{s_y}{\sqrt{SSX}}.$$

- ② A confidence interval of level  $1 - \gamma$ , for the intercept  $A$  is

$$\alpha \pm T_{n-2,\gamma/2} \times s_y \sqrt{\frac{1}{n} + \frac{\bar{x}^2}{SSX}}.$$

- ③ The prediction interval for a new observation  $x_0$ , of level  $1 - \gamma$ , is

$$\alpha + \beta x_0 \pm T_{n-2,\gamma/2} \times s_y \sqrt{1 + \frac{1}{n} + \frac{(x_0 - \bar{x})^2}{SSX}}.$$

## Student Distribution

### Definition

Let  $T$  be a continuous random variable, it follows a t-student distribution with  $n$  degree of freedom if its density is given by

$$f(x) = \frac{\Gamma\left(\frac{n+1}{2}\right)}{\sqrt{n\pi}\Gamma\left(\frac{n}{2}\right)} \left(1 + \frac{t^2}{n}\right)^{-\frac{n+1}{2}}.$$

We denote  $T \sim \mathcal{T}_n$ .

### Proposition

Let  $Z \sim N(0, 1)$  and  $D \sim \chi_n^2$ , suppose more that  $Z$  and  $D$  are independent. Then

$$T = \frac{Z}{\sqrt{D/n}} \sim \mathcal{T}_n.$$

### Checking Gaussian hypothesis

To valid our results we need to check our Gaussian assumption about the residual:

- 1 We need to test if the residue comes from a normal distribution, for this we may use the QQ-plot graphical test.
- 2 We need to test if the residue and the fitted come from independent variables, for this we may use the scatterplot of the residue versus the fitted values as a graphical test.

### Non-Gaussian case

- All results above are based on the assumption of normality of the residue.
- One may ask whether we still have the same results if the normality assumption is not any more satisfied?
- This leads to the mathematical investigation of looking at the limit behavior of the distribution when the number of observations goes to infinity.
- The results are still approximately correct !
- What we should do when the model is non-gaussian? A suggestion is:
  - 1 add predictor variable(s).
  - 2 and look for non-linear model.
- This will be the topic of the next chapters.