

Ex 4.3

$$\delta_k(x) = \underset{(1^*)}{x^T \Sigma^{-1} \mu_k} - \frac{1}{2} \mu_k^T \Sigma^{-1} \mu_k + \log \pi_k \quad (*)$$

Now transforming X to \hat{Y} (and thus $\Sigma \rightarrow \hat{\Sigma}$, $\mu \rightarrow \hat{\mu}$) via linear regression, we wish to show that LDA using \hat{Y} is identical to LDA in the original space e.g.

$$\underset{(1)}{y^T \hat{\Sigma}^{-1} \hat{\mu}_k} - \frac{1}{2} \hat{\mu}_k^T \hat{\Sigma}^{-1} \hat{\mu}_k + \log \pi_k = (*)$$

We will begin with a few observations. Firstly noticing the change in dimensionality:

$$(a) X^T \rightarrow (P \times N) (1 \times P)$$

$$\hat{Y}^T = X^T \hat{\beta} \rightarrow (1 \times P)(P \times K) = (1 \times K)$$

\Rightarrow We expect $\hat{\Sigma}^{-1}$ ($P \times P$) to change such that we have $\hat{\Sigma}^{-1}$ ($K \times K$)

$$(b) \hat{\mu}_k = \sum_{g \in k} \frac{\hat{y}_i}{N_k} = \sum_{g \in k} \frac{\hat{\beta}^T x_i}{N_k}$$

Now we will use the fact that \hat{Y} acts as an indicator matrix where each column corresponds to a class $k \in K$, and acts as an indicator vector taking the value 1 if row i is of class k and 0 otherwise. Thus we may transform $\hat{\mu}_k$ from a $(K \times 1)$ mean vector for ~~class $k \in K$~~ ^{observation} to each class $\hat{\mu}$ matrix ($K \times K$) with each of the K columns corresponding to the mean vector $\hat{\mu}_k$ of that class.

$\Rightarrow \hat{\mu} = \hat{\beta}^T X^T Y D^{-1}$ Where D is a $(K \times K)$ diagonal matrix with each entry corresponding to the number of observations in class k , e.g.

$$D = \begin{pmatrix} N_1 & & \cdots & 0 \\ \vdots & N_2 & \ddots & \vdots \\ 0 & \cdots & \cdots & N_K \end{pmatrix}$$

(c) Part (b) can also be demonstrated by noticing that
 $\hat{\beta}^T X^T = \hat{Y}^T$, thus making it clear that
 $\hat{\mu}_K = \frac{\hat{Y}^T Y_K}{N_K}$ is simply the mean vector of $y_i = K$.

Similarly, we notice that $\frac{X^T Y_K}{N_K} = \mu_K$.

$$\begin{aligned}
 (d) \quad \hat{\Sigma} &= \frac{1}{N-K} \sum_{K=1}^K \sum_{g_i \in K} (y_i - \hat{\mu}_K)(\hat{y}_i - \hat{\mu}_K)^T \\
 &= \frac{1}{N-K} \sum_K \sum_{g_i} (\beta^T x_i - \hat{\mu}_K)(\beta^T x_i - \hat{\mu}_K)^T \\
 &= \frac{1}{N-K} \sum_K \sum_{g_i} (\beta^T x_i - \beta^T \mu_K)(\beta^T x_i - \beta^T \mu_K)^T \quad (\text{using (c)}) \\
 &= \beta^T \left[\frac{1}{N-K} \sum_K \sum_{g_i} (x_i - \mu_K)(x_i - \mu_K)^T \right] \beta \\
 &= \beta^T \Sigma \beta \quad (\text{by definition})
 \end{aligned}$$

Now, returning to the original problem we will attempt to prove that $1, 2, 3 = 1^*, 2^*, 3^*$ respectively.

① In fact, we will attempt to show $\hat{Y}^T \hat{\Sigma}^{-1} \hat{\mu} = X^T \Sigma^{-1} \mu$ as discussed in (b).

Now,

$$\hat{Y}^T \hat{\Sigma}^{-1} \hat{\mu} = (X\beta)(\beta^T \Sigma \beta)^{-1} (\beta^T X^T Y D^{-1}) \quad (\text{using (b) and (d)})$$

Now assuming β is full rank, then an inverse exists such that $\beta \beta^{-1} = I_p$

$$\begin{aligned}
 &= X \beta \beta^{-1} \Sigma^{-1} (\beta \beta^{-1})^T X^T Y D^{-1} \\
 &= X \Sigma^{-1} X^T Y D^{-1} \\
 &= X \Sigma^{-1} \mu \quad (\text{using (c)}) \\
 &\quad \text{as required.}
 \end{aligned}$$

$$\textcircled{2} \quad \text{W.T.S. } \hat{\mu}_k^T \hat{\Sigma}^{-1} \hat{\mu}_k = \mu_k^T \Sigma^{-1} \mu_k$$

$$\begin{aligned} &\Rightarrow \hat{\mu}_k^T \hat{\Sigma}^{-1} \hat{\mu}_k \\ &= (\beta^T \mu_k)^T (\beta^T \Sigma \beta)^{-1} (\beta^T \mu_k) \\ &= \hat{\mu}_k^T \Sigma^{-1} \mu_k \quad (\text{assuming } \beta \text{ Full Rank as in } \textcircled{1}) \end{aligned}$$

$$\textcircled{3} \quad \log \pi_k = \log \pi_k$$

Combining \textcircled{1}, \textcircled{2} and \textcircled{3} we have proved that transforming X to \hat{Y} gives identical results in LDA.