Multiple Linear Regression

The Hat Matrix and Ridge Regression

Math 392

Preamble

In the last lecture, we noted that the least squares estimate, $\hat{\beta} = (X'X)^{-1}X'Y$, is unique only if the matrix X'X is invertible.

Example: invert a 2 by 3 matrix.

Error in solve.default(t(X) %*% X): system is computationally s⁻¹

What if we transpose X so that it is tall rather than wide?

```
X \leftarrow t(X)
## [,1] [,2]
## [1,] 1 1
## [2,] 1 3
## [3,] 1 5
 solve(t(X) %*% X)
## [,1] [,2]
## [1,] 1.458333 -0.375
## [2,] -0.375000 0.125
```

A Very Interesting Matrix: $oldsymbol{H}$

Recall:

$$\hat{Y} = X\hat{eta} = X(X'X)^{-1}X'Y$$

The "hat matrix", H,

$$H = X(X'X)^{-1}X'$$

because it puts a hat on the Y.

Properties of H

H is *symmetric*.

$$H' = X(X'X)^{-1}X'$$

= $X''[(X'X)^{-1}]'X'$
= $X[(X'X)']^{-1}X'$
= $X(X'X)^{-1}X'$
= H

H is idempotent.

$$H^{2} = X(X'X)^{-1}X'X(X'X)^{-1}X'$$

= $X(X'X)^{-1}X'$
= H

We can easily express the vector of *residuals*, $\hat{\epsilon}$.

$$\hat{\epsilon} = Y - \hat{Y} = Y - HY = (I - H)Y$$

Leverage

Let $h_{i,j}$ denote the $(i,j)^{th}$ element of H. Then we can express the fitted value of the i^{th} observation as

$$\hat{{Y}_i} = h_{i,i}Y_i + \sum_{j
eq i} h_{i,j}Y_j$$

Leverage, cont.

In simple linear regression,

$$H = X(X'X)^{-1}X'$$

$$=$$

For the i^{th} observation:

$$h_{i,i}=rac{1}{n}+rac{(x_i-ar{x})^2}{SS_X}$$

High leverage:

$$h_{i,i} > 2 \operatorname{avg}(h_{i,i}) = 2 rac{p}{n}$$

Exercises

Using the hat matrix and its properties, find $E(\hat{\epsilon})$ and $Var(\hat{\epsilon})$.

$$egin{aligned} E(\hat{\epsilon}|X) &= E((I-H)Y|X) \ &= (I-H)E(Y|X) \ &= (I-H)Xeta \ &= xeta - X(X'X)^{-1}X'Xeta \ &= 0 \end{aligned} \ Var(\hat{\epsilon}|X) &= Var((I-H)Y) \ &= (I-H)Var(Y)(I-H)' \ &= (I-H)\sigma^2(I-H)' \ &= \sigma^2(II'-HI'-IH'+HH') \ &= \sigma^2(I-H-H+H) \ &= \sigma^2(I-H) \end{aligned}$$

Ridge Regression

Consider an alternative estimator for β .

$$\hat{eta}_{ridge} = \operatorname{argmin}(RSS(eta)) \quad ext{subject to} \quad c \geq \sum_{j=1}^{p-1} eta_j^2; \quad c \geq 0$$

This is equivalent to minimizing the penalized RSS (for the scalar λ):

$$PRSS(eta) = (Y - Xeta)'(Y - Xeta) + \lambdaeta'eta \ = Y'Y - eta'X'Y - Y'Xeta + eta'X'Xeta + \lambdaeta'eta \ = Y'Y - 2eta'X'Y.$$

To find the $\hat{\beta}_{ridge}$ that minimize this function, we take the derivative with respect to β , set to zero, and solve.

$$egin{align} rac{\partial PRSS}{\partial eta} &= 0 - 2X'Y + 2X'Xeta + 2\lambdaeta = 0 \ X'Y &= (X'X + \lambda I)eta \ \hat{eta}_{ridge} &= (X'X + \lambda I)^{-1}X'Y \end{aligned}$$

Preamble revisited

Let's apply this technique to solve the problem of invertibility that we encountered in the preamble.

```
X \leftarrow matrix(c(1, 1, 1, 1, 3, 5), byrow = TRUE, nrow = 2)
## [,1] [,2] [,3]
## [1,] 1 1 1
## [2,] 1 3 5
 I <- diag(ncol(X))</pre>
## [,1] [,2] [,3]
## [1,] 1 0 0
## [2,] 0 1 0
## [3,] 0 0 1
```

```
lambda <- .5
solve(t(X) %*% X + lambda * I)</pre>
```

```
## [,1] [,2] [,3]
## [1,] 1.02890173 -0.4624277 0.04624277
## [2,] -0.46242775 1.3988439 -0.73988439
## [3,] 0.04624277 -0.7398844 0.47398844
```

Ridge on a tall matrix

```
X \leftarrow t(X)
## [,1] [,2]
## [1,] 1 1
## [2,] 1 3
## [3,] 1 5
I <- diag(ncol(X))</pre>
## [,1] [,2]
## [1,] 1 0
## [2,] 0 1
```

```
solve(t(X) %*% X + lambda * I)
```

```
## [,1] [,2]
## [1,] 0.8208092 -0.20809249
## [2,] -0.2080925 0.08092486
```

Ridge on a tall matrix, cont.

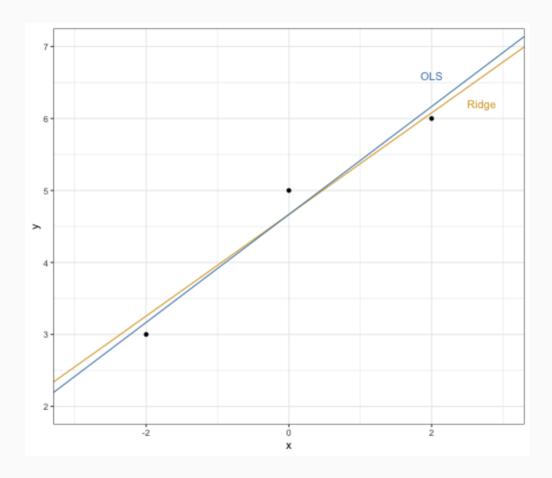
```
Y < -c(3, 5, 6)
# Ridge
solve(t(X) %*% X + lambda * I) %*% t(X) %*% Y
##
            [,1]
## [1,] 1.5028902
## [2,] 0.9710983
# OLS
solve(t(X) %*% X) %*% t(X) %*% Y
## [,1]
## [1,] 2.416667
## [2,] 0.750000
What's wrong?
```

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Excluding the intercept

```
X_{no_int} \leftarrow X[, -1]
X_no_int <- X_no_int - mean(X_no_int)</pre>
T <- 1
# Ridge
solve(t(X_no_int) %*% X_no_int + lambda * I) %*%
  t(X_no_int) %*% Y
##
              \lceil,1\rceil
## [1,] 0.7058824
# 01 S
solve(t(X_no_int) %*% X_no_int) %*% t(X_no_int) %*% Y
## [,1]
## [1,] 0.75
```

Ridge vs OLS



Properties of \hat{eta}_{ridge}

$$E(\hat{eta}_{ridge}|X) = E((X'X + \lambda I)^{-1}X'Y|X)$$

= $(X'X + \lambda I)^{-1}X'E(Y|X)$
= $(X'X + \lambda I)^{-1}X'X\beta$

Therefore the ridge estimates are biased for any $\lambda \neq 0$.

$$\begin{split} Var(\hat{\beta}_{ridge}|X) &= Var((X'X + \lambda I)^{-1}X'Y|X) \\ &= (X'X + \lambda I)^{-1}X'Var(Y|X)[(X'X + \lambda I)^{-1}X']' \\ &= (X'X + \lambda I)^{-1}X'\sigma^2I[(X'X + \lambda I)^{-1}X']' \\ &= \sigma^2(X'X + \lambda I)^{-1}X'X(X'X + \lambda I)^{-1} \end{split}$$

Note that as $\lambda \to \infty$, $Var(\hat{\beta}_{ridge}) \to 0$.