

Module 11: Lesson 2 Lab

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Background

Suppose that we have data on incubation periods $y = (y_1, \dots, y_n)$. We assume that the data are *independent* draws from a Gamma distribution with shape α and rate β , i.e.

$$y_i \sim \Gamma(\alpha, \beta) \quad \text{i.i.d., } i = 1, \dots, n$$

where the Gamma distribution has probability density function:

$$f(x|\alpha, \beta) = \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} \exp(-\beta x), \quad \alpha > 0, \beta > 0.$$

The goal is to make (sampling-based) Bayesian inference for the parameters α and β .

Exercise 1

Assume that α and β are *a priori* independent and assign Gamma distributions with parameters $\lambda\alpha$ and $\nu\alpha$, and $\lambda\beta$ and $\nu\beta$, respectively. Write down (i.e. express in mathematical terms) the posterior density of interest, i.e.

$$\pi(\alpha, \beta|y) \propto \pi(y|\alpha, \beta)\pi(\alpha)\pi(\beta)$$

where $\pi(y|\alpha, \beta)$ denotes the likelihood function.

$$\begin{aligned}
y_i &\sim \Gamma(\alpha, \beta) \\
f(x|\alpha, \beta) &= \frac{\beta^\alpha}{\Gamma(\alpha)} x^{\alpha-1} \exp(-\beta x) \\
\alpha &\sim \Gamma(\lambda_\alpha, \nu_\alpha) \\
\beta &\sim \Gamma(\lambda_\beta, \nu_\beta)
\end{aligned}$$

$$\begin{aligned}
\pi(\alpha, \beta|y) &\propto \pi(y|\alpha, \beta)\pi(\alpha)\pi(\beta) \\
\pi(y|\alpha, \beta) &= \frac{\beta^{n\alpha} \prod_k y_k^{\alpha-1} \exp(-\beta \sum y_k)}{\{\Gamma(\alpha)\}^n} \\
\pi(\alpha, \beta|y) &\propto \frac{\beta^{n\alpha} \prod_k y_k^{\alpha-1} \exp(-\beta \sum y_k)}{\{\Gamma(\alpha)\}^n} \times \pi(\alpha)\pi(\beta) \\
&\propto \frac{\beta^{n\alpha} \prod_k y_k^{\alpha-1} \exp(-\beta \sum y_k)}{\{\Gamma(\alpha)\}^n} \times \left(\frac{\nu_\alpha^{\lambda_\alpha}}{\Gamma(\lambda_\alpha)} \alpha^{\lambda_\alpha-1} \exp(-\nu_\alpha \alpha) \right) \times \left(\frac{\nu_\beta^{\lambda_\beta}}{\Gamma(\lambda_\beta)} \beta^{\lambda_\beta-1} \exp(-\nu_\beta \beta) \right) \\
&\propto \left(\frac{\beta^{n\alpha} \prod_k y_k^{\alpha-1} \exp(-\beta \sum y_k)}{\{\Gamma(\alpha)\}^n} \right) \times (\alpha^{\lambda_\alpha-1} \exp(-\nu_\alpha \alpha)) \times (\beta^{\lambda_\beta-1} \exp(-\nu_\beta \beta))
\end{aligned}$$

Exercise 2

Having obtained the joint posterior density in Exercise 1, first derive the densities of the posterior distribution of the parameters α and β up to proportionality, i.e. $\pi(\alpha|y)$ and $\pi(\beta|y)$.

Is any of these densities of a known form (i.e. the density of a standard/well-known distribution, e.g. Gamma, Normal etc)?

Full conditional density of α

$$\pi(\alpha|\beta, y) \propto \left(\frac{\beta^{n\alpha} \prod_k y_k^{\alpha-1}}{\{\Gamma(\alpha)\}^n} \right) \times (\alpha^{\lambda_\alpha-1} \exp(-\nu_\alpha \alpha))$$

This doesn't look like a standard well-known distribution, so we'll need to use M-H algorithm

Full conditional density of β

$$\begin{aligned}
\pi(\alpha, \beta|y) &\propto \left(\beta^{n\alpha} \prod_k \exp(-\beta y_k) \right) \times (\beta^{\lambda_\beta-1} \exp(-\nu_\beta \beta)) \\
&\propto \left(\beta^{n\alpha} \exp\left(-\beta \sum_{k=1}^n y_k\right) \right) \times (\beta^{\lambda_\beta-1} \exp(-\nu_\beta \beta)) \\
&\propto \beta^{n\alpha+\lambda_\beta-1} \exp\left[\beta \left(\sum_{k=1}^n y_k + \nu_\beta\right)\right] \\
\beta|\alpha, y &\sim \Gamma\left(n\alpha + \lambda_\beta, \sum_{k=1}^n y_k + \nu_\beta\right)
\end{aligned}$$

We can see that this looks like a Gamma distribution with the associated shape parameters, therefore, for β we can sample directly from its conditional distribution.

Exercise 3

Write a R function to implement an MCMC algorithm (from scratch!) which samples from the joint posterior distribution $\pi(\alpha, \beta|y)$.

First, think about what the input (i.e. arguments) the function should have, e.g. a vector corresponding to data \mathbf{y} , hyper-parameter values $\lambda\alpha$, etc. Then also think about what the output should be; surely the function should output the posterior samples of α and β . Anything else?

To get you started, a pseudo-code of this MCMC algorithm is given below:

1. Choose initial values for α and β ;
2. Update α by sampling from $\pi(\alpha|\beta, y)$ using a Gaussian random-walk;
3. Update β by sampling from $\pi(\beta|\alpha, y)$ directly.
4. Go to Step (b)

Look at the trace plots and convince yourself that the chain has reached stationarity and is mixing well.

Exercise 4

Test your algorithm is producing sensible results.

One way to do this is to first simulate a dataset from a Gamma distribution with some specific values for α and β , eg. `y <- rgamma(100, 4, 2)` will simulate a vector of 100 draws from a $\text{Ga}(4, 2)$.

Then use your MCMC algorithm (in Exercise 3) to sample from the joint posterior distribution of α and β assuming vague priors, e.g. $\lambda\alpha = \lambda\beta = 1$ and $\nu\alpha = \nu\beta = 10 - 3$.

In principle, your marginal posterior distributions ($\pi(\alpha|y)$ and $\pi(\beta|y)$) should be centered around the values you have chosen to simulate the data from (e.g. 4 and 2 in this case).

Exercise 5

If you have convinced yourself that your MCMC algorithm is doing what is supposed to be doing, then use it to fit the model to the *Campylobacter* data (24 observations in total) from Evans et al. 1996) by `y <- c(rep(2, 2), rep(3, 6), rep(4, 11), rep(5, 3), rep(7, 7))`

Does the model offer a good fit? How would you go assessing that?

Exercise 6* [Optional]

Now that you have got a working algorithm, can you improve its mixing?

One way to do this is to update both α and β at the same time (in a block). This can be done, for example, and as discussed in the lecture, by proposing values drawn from some (bivariate distribution) with density $q(\alpha, \beta)$ and accept/reject according to the Metropolis-Hastings ratio. There are many choices here, but not all of them will lead to an efficient sampler.

Here are a few choices to try:

- Gaussian random-walk in a block, i.e. propose values for α and β drawn from a bivariate Normal distribution with mean the current values of α and β and some (2×2) variance-covariance matrix, and then accept/reject using the Metropolis-Hastings ratio.

- Gaussian approximation to the posterior distribution $\pi(\alpha, \beta|y)$. (Hint: use the command `optim` to find the values of α and β that maximise the posterior density).