Inference methods for extended latent Gaussian models

Waterloo SAS Student Seminar Series

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November 2022

Motivation

- Surveillance of the HIV epidemic in sub-Saharan Africa
- Want to estimate indicators used for monitoring and response, including:
 - Prevalence ρ : the proportion of people who are HIV positive
 - Incidence λ : the proportion of people newly infected
 - Treatment coverage α : the proportion of PLHIV on treatment
- We would like to provide them at a local, district-level
- This is a challenging task! Data is noisy, sparse and biased

A simple small-area model for prevalence

- Consider areas i = 1, ..., n
- Simple random sample household-survey taken in each area, with sample sizes $m_i^{\rm HS}$
- The number of people testing positive is y_i^{HS}
- Then we can use a binomial logistic regression of the form:

$$egin{aligned} y_i^{ ext{HS}} &\sim ext{Bin}(m_i^{ ext{HS}},
ho_i^{ ext{HS}}), \ ext{logit}(
ho_i^{ ext{HS}}) &\sim g(\vartheta^{ ext{HS}}), \quad i=1,\ldots,n, \end{aligned}$$

- If g is Gaussian then this is a latent Gaussian model in the sense of Rue, Martino, and Chopin (2009)
- We usually set up g as a spatial smoother, because the sample sizes in each area are too small to get reliable direct estimates
- One problem with the above model is that household surveys are expensive to run, so they only happen rarely

Latent Gaussian models

• Three-stage Bayesian hierarchical model

where
$$\mathbf{y} = (y_1, ..., y_n)$$
, $\mathbf{x} = (x_1, ..., x_n)$, $\mathbf{\theta} = (\theta_1, ..., \theta_m)$

- Interested in learning both (θ, \mathbf{x}) from data \mathbf{y}
- If the middle layer is Gaussian, then it's a latent Gaussian model

(Latent field)
$$p(\mathbf{x} \mid \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}(\boldsymbol{\theta}), \boldsymbol{Q}(\boldsymbol{\theta})^{-1}).$$

- Covers most of the models commonly used in spatiotemporal statistics
- Latent field is typically indexed by spatiotemporal location, such that n > m

Adding ANC surveillance

- Pregnant women attending antenatal care clinics are routinely tested for HIV. to avoid mother-to-child transmission
- This data source is more real-time than household surveys, but it's also more biased, because attendees are unlikely to be as representative of the population
- But perhaps this bias is consistent, in which case we can still make use of the ANC data to improve our model!

Adding ANC surveillance

• Suppose of m_i^{ANC} women attending ANC, y_i^{ANC} are HIV positive, then we can use another binomial logistic regression:

$$egin{aligned} y_i^{\mathsf{ANC}} &\sim \mathsf{Bin}(m_i^{\mathsf{ANC}},
ho_i^{\mathsf{ANC}}), \ \mathsf{logit}(
ho_i^{\mathsf{ANC}}) &= \mathsf{logit}(
ho_i^{\mathsf{HS}}) + b_i, \ b_i &\sim \mathcal{N}(eta_b, \sigma_b^2), \end{aligned}$$

• Similar to using ρ_i^{ANC} as a covariate in our model for household survey prevalence, but this way we allow it to vary rather than be fixed

Adding ART coverage

- Remember that we're also interested in what proportion of PLHIV are receiving treatment
- To estimate this, we need first to know the number of PLHIV

$$H_i = N_i \rho_i^{\mathsf{HS}}$$

Then

$$A_i \sim \mathsf{Bin}(N_i,
ho_i^{\mathsf{HS}} lpha_i), \ \mathsf{logit}(lpha_i) \sim \mathcal{N}(eta_lpha, \sigma_lpha^2),$$

Naomi evidence synthesis model

- Combining these three modules is the basis of the Naomi evidence synthesis model
- Used by countries (which provide their own data) to produce HIV estimates in a yearly process supported by UNAIDS
- Can't run long MCMC in this setting, requires fast, accurate, approximations
- It's a complicated model, and requires something more flexible than R-INLA
- Currently using Template Model Builder TMB (Kristensen et al. 2015)



Figure 1: A supermodel

Template Model Builder

- So what is TMB?
- R package which implements the Laplace approximation for latent variable models using automatic differentiation (via CppAD)
 - For more about AD see e.g. Griewank and Walther (2008)
 - Useful for getting the mode, Hessian
- Write an objective function $f(\mathbf{x}, \boldsymbol{\theta})$ in C++ ("user template")
 - We select $f(\mathbf{x}, \theta) = -\log p(\mathbf{y} | \mathbf{x}, \theta) p(\mathbf{x} | \theta) p(\theta)$

Template Model Builder

```
#include <TMB.hpp>
template <class Type>
Type objective_function<Type>::operator()() {
 // Define data e.g.
 DATA_VECTOR(v):
 // Define parameters e.g.
 PARAMETER(mu):
 // Calculate negative log-likelihood e.g.
 nll = Type(0.0);
 nll -= dnorm(y, mu, 1, true).sum()
 return(nll):
```

Template Model Builder

- Performs the Laplace approximation $L_f(\theta) \approx L_f^*(\theta)$ and use R to optimise this with respect to θ to give $\hat{\theta}$ (the central point in Figure 2)
 - This is done by specifying the random argument to be the parameters that you want to integrate out with a Laplace approximation (the latent field)
- MAP estimate of ${\bf x}$ conditional on $\hat{{m heta}}$
- Standard errors calculated using the δ -method (a Gaussian assumption)

- Suggested reading: the paper (Rue, Martino, and Chopin 2009) or a book e.g. Blangiardo and Cameletti (2015)
- Approximate Bayesian inference for latent Gaussian models (LGMs), which recall are three-stage models with middle layer

(Latent field)
$$p(\mathbf{x} \mid \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}(\boldsymbol{\theta}), \boldsymbol{Q}(\boldsymbol{\theta})^{-1}).$$

• R-INLA implementation takes advantage of sparsity properties of $Q(\theta)$, i.e. if \mathbf{x} is a Gaussian Markov random field (GMRF)

- Gives approximate posterior marginals $\{\tilde{p}(x_i \mid \mathbf{y})\}_{i=1}^n$ and $\{\tilde{p}(\theta_i \mid \mathbf{y})\}_{i=1}^m$
- To approximate posterior marginals below requires $\tilde{p}(\theta \mid \mathbf{y})$ and $\tilde{p}(x_i \mid \theta, \mathbf{y})$

$$p(x_i | \mathbf{y}) = \int p(x_i, \boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta} = \int p(x_i | \boldsymbol{\theta}, \mathbf{y}) p(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}, \quad i = 1, ..., n, \quad (1)$$

$$p(\theta_j | \mathbf{y}) = \int p(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}_{-j} \quad j = 1, ..., m. \quad (2)$$

1) First Laplace approximate hyperparameter posterior

$$\widetilde{p}(\theta \mid \mathbf{y}) \propto \frac{p(\mathbf{y}, \mathbf{x}, \theta)}{\widetilde{p}_{G}(\mathbf{x} \mid \theta, \mathbf{y})}\Big|_{\mathbf{x} = \mu^{\star}(\theta)}$$
(3)

which can be marginalised to get $\tilde{p}(\theta_i | \mathbf{y})$

• Note: this involves integrating out a Gaussian approximation to the latent field, which we denote by $p_G(\mathbf{x} \mid \boldsymbol{\theta}, \mathbf{y})$

- 2) In both (1) and (2) we want to integrate w.r.t. (3), so choose integration points and weights $\{\theta^{(k)}, \Delta^{(k)}\}$
- For low *m* R-INLA uses a grid-strategy (illustrated in the next slide)
- For larger m this becomes too expensive and R-INLA uses a CCD design is used
- Other approaches, like adaptive Gaussian Hermite quadrature (AGHQ)
 have recently been shown to have theoretical guarantees (Bilodeau,
 Stringer, and Tang 2021) and may be preferable

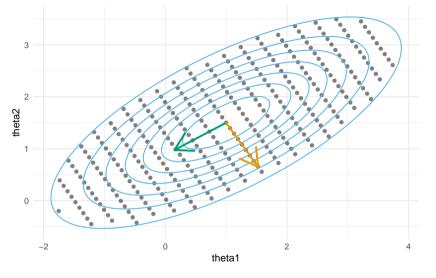


Figure 2: An illustration of the R-INLA grid method for selecting integration points using a toy bivariate Gaussian distribution for θ . Start at the mode and work outwards along the eigenvectors until the density drops sufficiently low.

Adaptive Gaussian Hermite Quadrature

- aghq R package and vignette (Stringer 2021)
- Gauss-Hermite quadrature is a way of picking nodes and weights, and is based on the theory of polynomial interpolation
- The adaptive part means that it uses the location (mode) and curvature (Hessian) of the target (posterior)
- Use k quadrature points
 - If *k* is odd then they include the mode
 - If k = 1 then it's a Laplace approximation
 - In the vignette k = 3 (for each dimension, so 3^m total) is chosen quite often

- 3) Choose approximation for $\tilde{p}(x_i | \theta, \mathbf{y})$
- Simplest version (Rue and Martino 2007) is to marginalise the $p_G(\mathbf{x} \mid \boldsymbol{\theta}, \mathbf{y})$

$$\tilde{p}(x_i \mid \boldsymbol{\theta}, \mathbf{y}) = \mathcal{N}(x_i \mid \mu_i^{\star}(\boldsymbol{\theta}), 1/q_i^{\star}(\boldsymbol{\theta})) \tag{4}$$

- The above is referred to as method = "gaussian" in R-INLA
- There are two better, more complex approximations
 - Confusingly called "simplified laplace" and "laplace"

- 4) Finally (!) use quadrature combining
 - our approximation $p(\theta \mid \mathbf{y})$ from step 1),
 - our integration points and weights $\{\theta^{(k)}, \Delta^{(k)}\}$ from step 2),
 - and our approximation $\tilde{p}(x_i | \theta, \mathbf{y})$ from step 3) to give

$$\tilde{p}(x_i \mid \mathbf{y}) = \sum_{k=1}^K \tilde{p}(x_i \mid \boldsymbol{\theta}^{(k)}, \mathbf{y}) \times \tilde{p}(\boldsymbol{\theta}^{(k)} \mid \mathbf{y}) \times \Delta^{(k)}$$
(5)

Experiments

- We would like to find an accurate way to do inference for the Naomi model
- It also needs to be fast enough to use in production!

Experiments

- We wrote a simplified version of the Naomi model up in TMB
- This allowed us to test three inference methods all using precisely the same model and C++ code
- 1. A direct Gaussian approximation via TMB
- 2. Adaptive Gaussian Hermite quadrature via aghq
- 3. No-U-Turn Sampling (NUTS a type of Hamiltonian Monte Carlo) via tmbstan
- Note: using different software it is often very difficult to ensure the model is precisely the same, so we're very fortunate here

Comparison approach

- You could look at the summaries like the mean and standard deviation of each of the posterior marginals
 - Any approximation method should be pretty good at getting the mean right
 - Gaussian approximations should be good at getting the second moment right
- It's probably better to compare the whole posterior distributions
- One way to do this is via Kolmogorov-Smirnov statistics, which give the maximum difference between two empirical CDFs

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