

Eventually an introduction to the aghq R package

Statistics and Epidemiology Group Meeting

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Motivation I

- People on ART A_i can be used as supplementary data for small-area estimation of HIV prevalence ρ_i

$$\begin{aligned}A_i &\sim \text{Bin}(N_i, \rho_i \alpha_i), \\ y_i &\sim \text{Bin}(m_i, \rho_i), \\ \text{logit}(\alpha_i) &\sim f(\vartheta_\alpha), \\ \text{logit}(\rho_i) &\sim g(\vartheta_\rho), \quad i = 1, \dots, n,\end{aligned}$$

- If f and g are Gaussian then model is almost, but not quite, a latent Gaussian model by the definition of Rue, Martino, and Chopin (2009)
 - This is due to small non-linearities (multiplying two latent Gaussian fields)
 - Each observation depends on more than one element of the latent field

Motivation II

- Previous slide is a simplified version of the Naomi evidence synthesis model
- Countries to fit the model using their own data (“in production”?)
 - Can’t run long MCMC on the cluster for weeks, as might be the case if this was one paper
- Can’t use R-INLA, require something more flexible
- Currently using Template Model Builder TMB (Kristensen et al. 2015)



Figure 1: A supermodel

Aside: common theme I

- Combining flawed (sparse, aggregated) gold standard (measuring the thing we want) data with other correlated (more available, high resolution) data (measuring not exactly what we want)
- Consistently resulting in models with multiple outcomes (evidence synthesis, multi-output)
- I think a lot of these are going to be not quite LGMs

Aside: common theme II

- Examples include
 - Naomi model: DHS data is “gold standard”, supported by ANC data from pregnant women
 - Sexual risk behaviour model: estimates of FSW population at national level, supported by DHS data
 - The national-level FSW estimates might be more like “bronze standard”
 - DHS approximately asks “have you received money or gifts in exchange for sex in past 12 months”
 - Loa loa prevalence and eyeworm history prevalence model: measuring eyeworm history is a cheap proxy for Loa loa (Amoah, Diggle, and Giorgi 2020)

Recap on latent Gaussian models

- Three-stage model

$$\text{(Observations)} \quad \mathbf{y} \sim p(\mathbf{y} \mid \mathbf{x}),$$

$$\text{(Latent field)} \quad \mathbf{x} \sim p(\mathbf{x} \mid \boldsymbol{\theta}),$$

$$\text{(Hyperparameters)} \quad \boldsymbol{\theta} \sim p(\boldsymbol{\theta}),$$

where $\mathbf{y} = (y_1, \dots, y_n)$, $\mathbf{x} = (x_1, \dots, x_n)$, $\boldsymbol{\theta} = (\theta_1, \dots, \theta_m)$

- Interested in learning both $(\boldsymbol{\theta}, \mathbf{x})$ from data \mathbf{y}
- Covers most of the models used in spatiotemporal statistics

Recap on Integrated Nested Laplace Approximation I

- Rue, Martino, and Chopin (2009) or e.g. Blangiardo and Cameletti (2015)
- Approximate Bayesian inference for **latent Gaussian models** (LGMs), which are three-stage models with middle layer

(Latent field) $p(\mathbf{x} \mid \boldsymbol{\theta}) = \mathcal{N}(\mathbf{x} \mid \boldsymbol{\mu}(\boldsymbol{\theta}), \mathbf{Q}(\boldsymbol{\theta})^{-1}).$

- R-INLA implementation takes advantage of sparsity properties of $\mathbf{Q}(\boldsymbol{\theta})$, i.e. if \mathbf{x} is a Gaussian Markov random field (GMRF)

Recap on Integrated Nested Laplace Approximation II

- Gives approximate **posterior marginals** $\{\tilde{p}(x_i | \mathbf{y})\}_{i=1}^n$ and $\{\tilde{p}(\theta_j | \mathbf{y})\}_{j=1}^m$
- To approximate posterior marginals below requires $\tilde{p}(\boldsymbol{\theta} | \mathbf{y})$ and $\tilde{p}(x_i | \boldsymbol{\theta}, \mathbf{y})$

$$p(x_i | \mathbf{y}) = \int p(x_i, \boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta} = \int p(x_i | \boldsymbol{\theta}, \mathbf{y}) p(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}, \quad i = 1, \dots, n, \quad (1)$$

$$p(\theta_j | \mathbf{y}) = \int p(\boldsymbol{\theta} | \mathbf{y}) d\boldsymbol{\theta}_{-j} \quad j = 1, \dots, m. \quad (2)$$

Recap on Integrated Nested Laplace Approximation III

- 1) First Laplace approximate hyperparameter posterior

$$\tilde{p}(\boldsymbol{\theta} | \mathbf{y}) \propto \frac{p(\mathbf{y}, \mathbf{x}, \boldsymbol{\theta})}{\tilde{p}_G(\mathbf{x} | \boldsymbol{\theta}, \mathbf{y})} \Big|_{\mathbf{x}=\boldsymbol{\mu}^*(\boldsymbol{\theta})} \quad (3)$$

which can be marginalised to get $\tilde{p}(\theta_j | \mathbf{y})$

- 2) In both (1) and (2) we want to integrate w.r.t. (3), so choose integration points and weights $\{\boldsymbol{\theta}^{(k)}, \Delta^{(k)}\}$
 - For low m INLA uses a grid-strategy which I illustrate in the next slide
 - For larger m this becomes too expensive and a CCD design is used

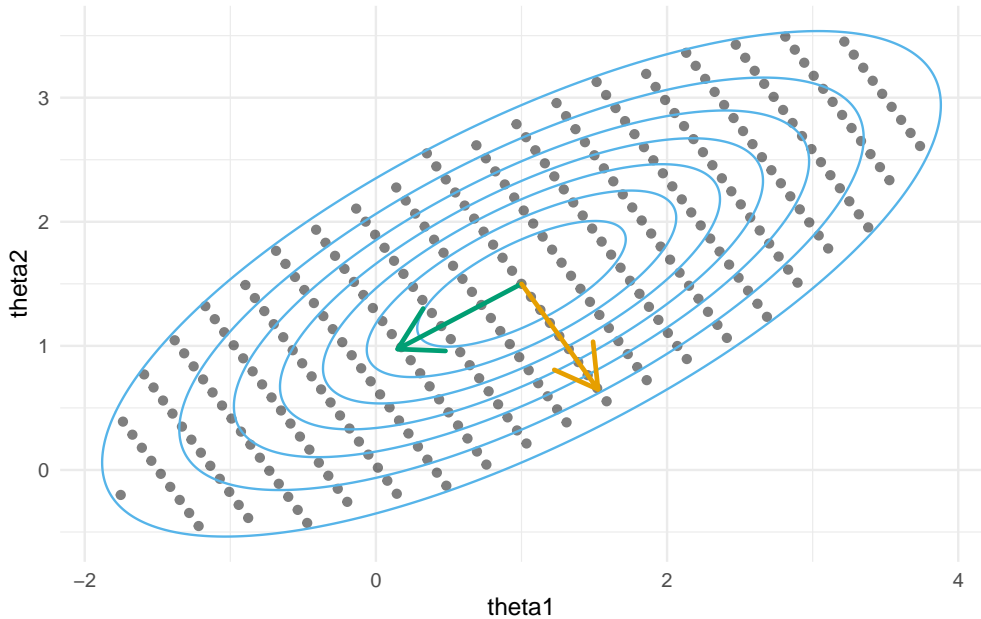


Figure 2: An illustration of the INLA grid method for selecting integration points using a toy

Recap on Integrated Nested Laplace Approximation IV

3) Choose approximation for $\tilde{p}(x_i | \boldsymbol{\theta}, \mathbf{y})$

- Simplest version (Rue and Martino 2007) is to marginalise the $p_G(\mathbf{x} | \boldsymbol{\theta}, \mathbf{y})$

$$\tilde{p}(x_i | \boldsymbol{\theta}, \mathbf{y}) = \mathcal{N}(x_i | \mu_i^*(\boldsymbol{\theta}), 1/q_i^*(\boldsymbol{\theta})) \quad (4)$$

- The above is referred to as “Gaussian” approximation, and confusingly there are two more complex ones called “simplified Laplace” and “Laplace”
- You can pick which one in R-INLA using the method option

4) Finally use quadrature to get

$$\tilde{p}(x_i | \mathbf{y}) = \sum_{k=1}^K \tilde{p}(x_i | \boldsymbol{\theta}^{(k)}, \mathbf{y}) \times \tilde{p}(\boldsymbol{\theta}^{(k)} | \mathbf{y}) \times \Delta^{(k)} \quad (5)$$

Template Model Builder I

- R package which implements the Laplace approximation for latent variable models using AD (via CppAD)
 - For more about AD see e.g. Griewank and Walther (2008)
 - Useful for getting the mode, Hessian
- Write an objective function $f(\mathbf{x}, \boldsymbol{\theta})$ in C++ (“user template”)
 - We select $f(\mathbf{x}, \boldsymbol{\theta}) = -\log p(\mathbf{y} | \mathbf{x}, \boldsymbol{\theta})p(\mathbf{x} | \boldsymbol{\theta})p(\boldsymbol{\theta})$

Template Model Builder II

```
#include <TMB.hpp>
```

```
template <class Type>
```

```
Type objective_function<Type>::operator>()() {
```

```
  // Define data e.g.
```

```
  DATA_VECTOR(y);
```

```
  // Define parameters e.g.
```

```
  PARAMETER(mu);
```

```
  // Calculate negative log-likelihood e.g.
```

```
  nll = Type(0.0);
```

```
  nll -= dnorm(y, mu, 1, true).sum()
```

```
  return(nll);
```

```
}
```

Template Model Builder III

- Performs the Laplace approximation $L_f(\boldsymbol{\theta}) \approx L_f^*(\boldsymbol{\theta})$ and use R to optimise this with respect to $\boldsymbol{\theta}$ to give $\hat{\boldsymbol{\theta}}$ (the central point in Figure 2)
 - This is done by specifying the `random` argument to be the parameters that you want to integrate out with a Laplace approximation (the latent field)
- MAP estimate of \mathbf{x} conditional on $\hat{\boldsymbol{\theta}}$
- Standard errors calculated using the δ -method (a Gaussian assumption)

Adaptive Gaussian Hermite Quadrature

- Recent work by Alex Stringer and coauthors on AGHQ
 - `aghq` R package and vignette (Stringer 2021)
 - Theory paper (Bilodeau, Stringer, and Tang 2021)
- Gauss-Hermite quadrature is a way of picking nodes and weights, and is based on the theory of polynomial interpolation
- The adaptive part means that it uses the location (mode) and curvature (Hessian) of the target (posterior)
- Use k quadrature points
 - If k is odd then they include the mode
 - If $k = 1$ then it's a Laplace approximation
 - In the vignette $k = 3$ (for each dimension, so 3^m total) is chosen quite often

Epil example I

- Epilepsy example from Section 5.2. of Rue, Martino, and Chopin (2009) (previously from BUGS):
 - Patients $i = 1, \dots, 59$ each either assigned treatment $\text{Trt}_i = 1$ or placebo $\text{Trt}_i = 0$ to help with seizures
 - Visits to clinics $j = 1, \dots, 4$ times with y_{ij} the number of seizures of the i th person in the two weeks proceeding their j th visit to the clinic
 - Covariates age Age_i , baseline seizure counts Base_i and an indicator for the final clinic visit V_4

Epil example II

This is what the model looks like (it's a Poisson GLMM):

$$y_{ij} \sim \text{Poisson}(\lambda_{ij}),$$

$$\lambda_{ij} = e^{\eta_{ij}},$$

$$\eta_{ij} = \beta_0 + \beta_{\text{Base}} \log(\text{Baseline}_j/4) + \beta_{\text{Trt}} \text{Trt}_i + \beta_{\text{Trt} \times \text{Base}} \text{Trt}_i \times \log(\text{Baseline}_j/4) \\ + \beta_{\text{Age}} \log(\text{Age}_i) + \beta_{\text{V}_4} \text{V}_{4j} + \epsilon_i + \nu_{ij}, \quad i = 1 : 59, \quad j = 1 : 4,$$

$$\beta \sim \mathcal{N}(0, 100^2), \quad \forall \beta,$$

$$\epsilon_i \sim \mathcal{N}(0, 1/\tau_\epsilon),$$

$$\nu_{ij} \sim \mathcal{N}(0, 1/\tau_\nu),$$

$$\tau_\epsilon \sim \Gamma(0.001, 0.001),$$

$$\tau_\nu \sim \Gamma(0.001, 0.001).$$

Epil example III

- aghq package interfaces really easily with TMB!
- This is the code I used to fit the model with TMB

```
obj <- MakeADFun(  
  data = dat,  
  parameters = param,  
  # These are the ones integrated out with a Laplace approximation  
  random = c("epsilon", "nu"),  
  DLL = "epil"  
)
```

Epil example IV

- Then to fit it with aghq it's only a very small modification

```
fit <- aghq::marginal_laplace_tmb(  
  obj,  
  k = 3,  
  startingvalue = c(param$beta, param$l_tau_epsilon, param$l_tau_nu)  
)
```

	Stan	INLA_G	INLA_SL	INLA_L	TMB	glmmTMB	tmbstan	aghq
beta_0	1.572	1.626	1.573	1.573	1.579	1.579	1.571	1.573
sd(beta_0)	0.076	0.077	0.078	0.078	0.073	0.073	0.081	0.076
beta_1	-0.968	-0.927	-0.954	-0.956	-0.949	-0.949	-0.961	-0.955
sd(beta_1)	0.420	0.419	0.419	0.419	0.396	0.396	0.426	0.411
beta_2	0.879	0.859	0.880	0.881	0.880	0.880	0.875	0.881
sd(beta_2)	0.136	0.138	0.138	0.138	0.129	0.129	0.135	0.135
beta_3	-0.103	-0.101	-0.103	-0.104	-0.103	-0.103	-0.105	-0.103
sd(beta_3)	0.087	0.086	0.086	0.086	0.086	0.086	0.086	0.086
beta_4	0.488	0.471	0.484	0.485	0.490	0.490	0.489	0.486
sd(beta_4)	0.352	0.364	0.364	0.364	0.342	0.342	0.375	0.357
beta_5	0.356	0.340	0.350	0.351	0.349	0.349	0.360	0.351
sd(beta_5)	0.212	0.213	0.213	0.213	0.200	0.200	0.215	0.209

Plan

- Test `aghq` for toy Naomi example
 - Do as above with the Epil example, testing versus a long MCMC run
 - Sometimes you have to look pretty hard for a node (element of the latent field) where there are differences. In the INLA paper they do this by computing a SKLD and ordering by maximum difference. Probably good to do here as well
- Extend `aghq` to replicate INLA functionality by adding the more complex versions of $\tilde{p}(x_i | \theta, \mathbf{y})$ then test that with Naomi
 - Håvard Rue philosophy: “do one thing and do it well”
 - R-INLA implementation of INLA based on sparsity of $\mathbf{Q}(\theta)$ that doesn't hold up for extended LGMs
 - Wood (2020) on how to still do it
- Try the INLA without R-INLA on other almost LGMs and see how far it can be pushed

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